

Package ‘VGAM’

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Title Vector Generalized Linear and Additive Models

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Depends R (>= 4.1.0), methods, stats, stats4, splines

Suggests VGAMextra, MASS, mgcv

Enhances VGAMdata

Description An implementation of about 6 major classes of statistical regression models. The central algorithm is Fisher scoring and iterative reweighted least squares. At the heart of this package are the vector generalized linear and additive model (VGLM/VGAM) classes. VGLMs can be loosely thought of as multivariate GLMs. VGAMs are data-driven VGLMs that use smoothing. The book ``Vector Generalized Linear and Additive Models: With an Implementation in R'' (Yee, 2015) <[DOI:10.1007/978-1-4939-2818-7](https://doi.org/10.1007/978-1-4939-2818-7)> gives details of the statistical framework and the package. Currently only fixed-effects models are implemented. Many (100+) models and distributions are estimated by maximum likelihood estimation (MLE) or penalized MLE. The other classes are RR-VGLMs (reduced-rank VGLMs), quadratic RR-VGLMs, doubly constrained RR-VGLMs, quadratic RR-VGLMs, reduced-rank VGAMs, RCIMs (row-column interaction models)---these classes perform constrained and unconstrained quadratic ordination (CQO/UQO) models in ecology, as well as constrained additive ordination (CAO). Hauck-Donner effect detection is implemented. Note that these functions are subject to change; see the NEWS and ChangeLog files for latest changes.

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URL <https://www.stat.auckland.ac.nz/~yee/VGAM/>

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Contents

VGAM-package	13
A1A2A3	17
AA.Aa.aa	18
AB.Ab.aB.ab	20
ABO	21
acat	22
add1.vglm	24
AICvlm	25
alogitlink	27
altered	29
amlbinomial	31
amlexponential	33
amlnormal	35
amlpoisson	38
anova.vglm	40
AR1	42
AR1EIM	45
asinlink	49
auuc	51
aux.posbernoulli.t	52
backPain	53
beggs	54
bell	55
Benford	56
Benini	57
benini1	59
Betabinom	60
betabinomial	64
betabinomialff	67
betaff	70
Betageom	72
betageometric	73
betaII	75
Betanorm	76
betaprime	78
betaR	79
Biamhcop	81
biamhcop	82

Biclaytoncop	84
biclaytoncop	85
BICvlm	87
Bifgmcop	88
bifgmcop	89
bifgmexp	90
bifrankcop	92
bigumbellexp	93
bilogis	95
biogistic	96
Binom2.or	98
binom2.or	100
Binom2.rho	103
binom2.rho	105
binomialff	107
Binorm	110
binormal	111
binormalcop	113
Binormcop	115
Biplackett	116
biplackettcop	117
biplot-methods	119
Bisa	119
bisa	121
Bistudentt	123
bistudentt	124
bmi.nz	125
borel.tanner	126
Bort	128
Brat	129
brat	130
bratt	132
budworm	134
calibrate	135
calibrate-methods	137
calibrate.qrrvglm	137
calibrate.qrrvglm.control	141
calibrate.rrvglm	142
calibrate.rrvglm.control	144
cao	145
cao.control	149
Card	152
cardioid	154
cauchitlink	156
cauchy	158
cdf.lmscreg	160
cens.gumbel	161
cens.normal	163

cens.poisson	165
cfibrosis	167
cgo	168
chest.nz	169
chinese.nz	170
chisq	171
clo	172
clogloglink	174
CM.equid	176
coalminers	177
Coef	178
Coef.qrrvglm	179
Coef.qrrvglm-class	181
Coef.rrvglm	183
Coef.rrvglm-class	184
Coef.vlm	185
coefvgam	186
coefvlm	187
CommonVGAMffArguments	188
concoef	197
concoef-methods	198
confintvglm	199
constraints	201
cops	203
corbet	204
cqo	205
crashes	212
cratio	214
cumulative	216
Dagum	220
dagum	222
dARI	223
deermice	225
deplot.lmscreg	226
depvar	228
dextlogF	229
df.residual	230
dgaitdplot	231
dhuber	234
Diffzeta	236
diffzeta	237
dirichlet	238
dirmul.old	240
dirmultinomial	242
dlogF	244
double.cens.normal	245
double.expbinomial	247
ducklings	250

eCDF	250
enzyme	251
erf	252
erlang	253
Expectiles-Exponential	255
Expectiles-Normal	256
Expectiles-sc.t2	258
Expectiles-Uniform	259
expexpff	262
expexpff1	264
expgeom	266
expgeometric	267
expint	268
explink	270
explog	271
explogff	272
exponential	274
exppois	275
exppoisson	277
Extbetabinom	278
extbetabinomial	280
extlogF1	282
familynome	285
Felix	286
felix	287
fff	288
fill1	290
finney44	293
fisherzlink	294
Fisk	295
fisk	296
fittedvln	298
fix.crossing	300
flourbeetle	301
Foldnorm	302
foldnormal	303
formulavln	305
Frank	306
Frechet	308
frechet	309
freund61	311
Gaitdbinom	313
Gaitdlog	316
gaitdlog	318
Gaitdnbinom	321
gaitdnbinomial	323
Gaitdpois	327
gaitdpoisson	331

Gaitdzeta	337
gaitdzeta	340
gamma1	342
gamma2	344
gammaff.mm	346
gammahyperbola	348
gammaR	349
garma	351
GenbetaII	353
genbetaII	354
gengamma.stacy	356
gengammaUC	358
Genpois0	360
Genpois1	362
genpoisson0	363
genpoisson1	366
genpoisson2	367
genray	369
genrayleigh	370
Gensh	372
gensh	373
geometric	375
get.smart	376
get.smart.prediction	377
gev	378
gevUC	381
gew	383
goffset	384
Gompertz	386
gompertz	387
gpd	389
gpdUC	392
grain.us	394
grc	395
gumbel	400
Gumbel-II	403
gumbelIII	404
gumbelUC	406
guplot	408
has.interceptvlm	409
hatvalues	410
hdiff	412
hdiffsev	416
hormone	418
hspider	420
huber2	422
Huggins89.t1	423
hunua	425

Hurea	427
hurea	428
hyperg	430
hypersecant	431
Hzeta	433
hzeta	434
iam	435
identitylink	437
Influence	438
inv.binomial	439
Inv.gaussian	441
inv.gaussianff	442
Inv.lomax	444
inv.lomax	445
Inv.paralogistic	447
inv.paralogistic	448
is.buggy	450
is.crossing	451
is.parallel	452
is.smart	453
is.zero	454
kendall.tau	455
KLD	456
Kumar	458
kumar	459
lakeO	460
lambertW	462
laplace	463
laplaceUC	465
latvar	466
leipnik	468
lerch	469
leukemia	471
levy	472
lgamma1	473
lgammaUC	475
Lindley	476
lindley	478
linkfun	479
Links	480
Lino	483
lino	485
lirat	487
lms.bcg	488
lms.bcn	490
lms.yjn	493
Log	496
log1mexp	497

loglink	498
logF	499
logff	501
logistic	502
logitlink	504
logitoffsetlink	507
logLik.vlm	508
loglinb2	509
loglinb3	511
loglink	512
logloglink	514
lognormal	515
logofflink	517
Lomax	518
lomax	520
lpossoms	521
lqnorm	522
lrt.stat	524
lrtest	526
lvplot	527
lvplot.qrrvglm	529
lvplot.rrvglm	533
machinists	536
Makeham	537
makeham	538
margeff	540
marital.nz	542
Max	543
Maxwell	545
maxwell	546
mccullagh89	547
meangaitd	549
melbmaxtemp	550
meplot	551
micmen	553
mills.ratio	555
mix2exp	556
mix2normal	558
mix2poisson	560
MNSs	562
model.framevlm	564
model.matrixqrrvglm	565
model.matrixvlm	566
moffset	568
multilogitlink	570
multinomial	571
N1binom	576
N1binomial	577

Nlpois	579
Nlpoisson	580
Nakagami	582
nakagami	583
nbcnlink	585
negbinomial	587
negbinomial.size	594
normal.vcm	596
nparam.vlm	599
olympics	601
Opt	602
ordpoisson	603
ordsup	606
oxtemp	607
Paralogistic	608
paralogistic	609
Pareto	611
paretoff	612
ParetoIV	614
paretoIV	616
Perks	619
perks	620
perspqrvglm	622
pgamma.deriv	624
pgamma.deriv.unscaled	626
plotdeplot.lmscreg	627
plotdgaitd.vglm	629
plotqrrvglm	630
plotqtplot.lmscreg	631
plotrcim0	633
plotvgam	635
plotvgam.control	638
plotvglm	639
pneumo	641
poisson.points	642
poissonff	643
PoissonPoints	645
Polono	646
posbernoulli.b	648
posbernoulli.t	651
posbernoulli.tb	654
posbernUC	657
posbinomial	659
Posgeom	661
posnegbinomial	662
Posnorm	665
posnormal	667
pospoisson	669

powerlink	670
prats	671
predictqrrvglm	673
predictvglm	674
prentice74	676
prinia	678
probitlink	679
profilevglm	681
propodds	682
prplot	683
put.smart	685
qrrvglm.control	686
qtplot.gumbel	691
qtplot.lmscreg	693
Qvar	694
qvar	698
R2latvar	699
Rank	700
Rayleigh	701
rayleigh	702
Rcim	704
rcqo	706
rdiric	710
rec.exp1	711
rec.normal	712
reciprocallink	714
residualsvglm	715
rhobitlink	717
Rice	719
riceff	720
rigff	722
rlplot.gevff	723
rootogram4	725
round2	727
rrar	728
rrvglm	730
rrvglm-class	734
rrvglm.control	736
rrvglm.optim.control	739
ruge	741
s	742
sc.studentt2	744
score.stat	745
seglines	747
Select	748
seq2binomial	751
setup.smart	752
Simplex	754

simplex	755
simulate.vlm	756
Sinmad	758
sinmad	759
Skellam	761
skellam	763
skewnorm	764
skewnormal	766
Slash	767
slash	769
sloglink	770
sm.os	772
sm.ps	775
smart.expression	778
smart.mode.is	778
smartpred	779
specials	781
spikeplot	782
sqrtlink	784
sratio	786
step4	789
studentt	790
summary.drrvglm	792
summarypvgam	795
summaryvgam	796
summaryvglm	797
SURff	800
SurvS4	802
SurvS4-class	804
TIC	805
Tobit	806
tobit	808
Tol	812
Topple	814
topple	815
toxop	817
Triangle	818
trim.constraints	819
Trinorm	821
trinormal	823
trplot	824
trplot.qrrvglm	826
Trunc	828
Truncpareto	830
truncweibull	831
ucberk	833
uninormal	834
UtilitiesVGAM	836

V1	838
V2	839
vcovvlm	840
venice	842
vgam	843
vgam-class	848
vgam.control	850
vglm	853
vglm-class	859
vglm.control	862
vglmff-class	866
vonmises	869
vplot.profile	871
vsmooth.spline	872
waitakere	874
wald.stat	876
waldff	878
weibull.mean	879
weibullR	881
weightsvglm	884
wine	885
wrapup.smart	886
yeo.johnson	887
Yules	888
yulesimon	890
Zabinom	891
zabinomial	892
Zageom	894
zageometric	895
Zanegbin	897
zanegbinomial	899
Zapois	902
zapoisson	903
zero	905
Zeta	907
zeta	908
zetaff	910
Zibinom	912
zibinomial	913
Zigeom	916
zigeometric	917
Zinegbin	919
zinegbinomial	921
zipebcom	924
Zipf	927
zipf	928
Zipfmb	930
Zipois	931

zipoisson	933
Zoabeta	937
zoabetaR	939

Index	941
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VGAM-package	<i>Vector Generalized Linear and Additive Models and Other Associated Models</i>
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Description

VGAM provides functions for fitting vector generalized linear and additive models (VGLMs and VGAMs), and associated models (Reduced-rank VGLMs or RR-VGLMs, Doubly constrained RR-VGLMs (DRR-VGLMs), Quadratic RR-VGLMs, Reduced-rank VGAMs). This package fits many models and distributions by maximum likelihood estimation (MLE) or penalized MLE, under this statistical framework. Also fits constrained ordination models in ecology such as constrained quadratic ordination (CQO).

Details

This package centers on the *iteratively reweighted least squares* (IRLS) algorithm. Other key words include Fisher scoring, additive models, reduced-rank regression, penalized likelihood, and constrained ordination. The central modelling functions are `vglm`, `vgam`, `rrvglm`, `rcim`, `cqo`, `cao`. Function `vglm` operates very similarly to `glm` but is much more general, and many methods functions such as `coef` and `predict` are available. The package uses S4 (see `methods-package`).

Some notable companion packages: (1) **VGAMdata** mainly contains data sets useful for illustrating **VGAM**. Some of the big ones were initially from **VGAM**. Recently, some older **VGAM** family functions have been shifted into this package. (2) **VGAMextra** written by Victor Miranda has some additional **VGAM** family and link functions, with a bent towards time series models. (3) **svyVGAM** provides design-based inference, e.g., to survey sampling settings. This is because the weights argument of `vglm` can be assigned any positive values including survey weights.

Compared to other similar packages, such as **gamlss** and **mgecv**, **VGAM** has more models implemented (150+ of them) and they are not restricted to a location-scale-shape framework or (largely) the 1-parameter exponential family. The general statistical framework behind it all, once grasped, makes regression modelling unified. Some features of the package are: (i) many family functions handle multiple responses; (ii) reduced-rank regression is available by operating on latent variables (optimal linear combinations of the explanatory variables); (iii) basic automatic smoothing parameter selection is implemented for VGAMs (`sm.os` and `sm.ps` with a call to `magic`), although it has to be refined; (iv) *smart* prediction allows correct prediction of nested terms in the formula provided smart functions are used.

The GLM and GAM classes are special cases of VGLMs and VGAMs. The VGLM/VGAM framework is intended to be very general so that it encompasses as many distributions and models as possible. VGLMs are limited only by the assumption that the regression coefficients enter through a set of linear predictors. The VGLM class is very large and encompasses a wide range of multivariate response types and models, e.g., it includes univariate and multivariate distributions, categorical data analysis, extreme values, correlated binary data, quantile and expectile regression, time series

problems. Potentially, it can handle generalized estimating equations, survival analysis, bioassay data and nonlinear least-squares problems.

Crudely, VGAMs are to VGLMs what GAMs are to GLMs. Two types of VGAMs are implemented: 1st-generation VGAMs with `s` use vector backfitting, while 2nd-generation VGAMs with `sm.os` and `sm.ps` use O-splines and P-splines so have a direct solution (hence avoids backfitting) and have automatic smoothing parameter selection. The former is older and is based on Yee and Wild (1996). The latter is more modern (Yee, Somchit and Wild, 2024) but it requires a reasonably large number of observations to work well because it is based on optimizing over a predictive criterion rather than using a Bayesian approach.

An important feature of the framework is that of *constraint matrices*. They apportion the regression coefficients according to each explanatory variable. For example, since each parameter has a link function applied to it to turn it into a linear or additive predictor, does a covariate have an equal effect on each parameter? Or no effect? Arguments such as zero, parallel and exchangeable, are merely easy ways to have them constructed internally. Users may input them explicitly using the constraint argument, and `CM.symm0` etc. can make this easier.

Another important feature is implemented by `xi.j`. It allows different linear/additive predictors to have a different values of the same explanatory variable, e.g., `multinomial` for the conditional logit model and the like.

VGLMs with dimension reduction form the class of RR-VGLMs. This is achieved by reduced rank regression. Here, a subset of the constraint matrices are estimated rather than being known and prespecified. Optimal linear combinations of the explanatory variables are taken (creating latent variables) which are used for fitting a VGLM. Thus the regression can be thought of as being in two stages. The class of DRR-VGLMs provides further structure to RR-VGLMs by allowing constraint matrices to be specified for each column of **A** and row of **C**. Thus the reduced rank regression can be fitted with greater control.

This package is the first to check for the *Hauck-Donner effect* (HDE) in regression models; see `hdeff`. This is an aberration of the Wald statistics when the parameter estimates are too close to the boundary of the parameter space. When present the p-value of a regression coefficient is biased upwards so that a highly significant variable might be deemed nonsignificant. Thus the HDE can create havoc for variable selection!

Somewhat related to the previous paragraph, hypothesis testing using the likelihood ratio test, Rao's score test (Lagrange multiplier test) and (modified) Wald's test are all available; see `summaryvglm`. For all regression coefficients of a model, taken one at a time, all three methods require further IRLS iterations to obtain new values of the other regression coefficients after one of the coefficients has had its value set (usually to 0). Hence the computation load is overall significant.

For a complete list of this package, use `library(help = "VGAM")`. New **VGAM** family functions are continually being written and added to the package.

Warning

This package is undergoing continual development and improvement, therefore users should treat many things as subject to change. This includes the family function names, argument names, many of the internals, moving some functions to **VGAMdata**, the use of link functions, and slot names. For example, many link functions were renamed in 2019 so that they all end in "link", e.g., `loglink()` instead of `loge()`. Some future pain can be avoided by using good programming techniques, e.g., using extractor functions such as `coef()`, `weights()`, `vcov()`, `predict()`. Although

changes are now less frequent, please expect changes in all aspects of the package. See the NEWS file for a list of changes from version to version.

Author(s)

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Yee, T. W. and Ma, C. (2024). Generally altered, inflated, truncated and deflated regression. *Statistical Science*, **39** (in press).

Yee, T. W. (2022). On the Hauck-Donner effect in Wald tests: Detection, tipping points and parameter space characterization, *Journal of the American Statistical Association*, **117**, 1763–1774. doi:10.1080/01621459.2021.1886936.

Yee, T. W. and Somchit, C. and Wild, C. J. (2024). Penalized vector generalized additive models. Manuscript in preparation.

The website for the **VGAM** package and book is <https://www.stat.auckland.ac.nz/~yee/>. There are some resources there, especially as relating to my book and new features added to **VGAM**.

Some useful background reference for the package include:

Chambers, J. and Hastie, T. (1991). *Statistical Models in S*. Wadsworth & Brooks/Cole.

Green, P. J. and Silverman, B. W. (1994). *Nonparametric Regression and Generalized Linear Models: A Roughness Penalty Approach*. Chapman and Hall.

Hastie, T. J. and Tibshirani, R. J. (1990). *Generalized Additive Models*. Chapman and Hall.

See Also

[vglm](#), [vgam](#), [rrvglm](#), [rcim](#), [cqo](#), [TypicalVGAMfamilyFunction](#), [CommonVGAMffArguments](#), [Links](#), [hdeff](#), [glm](#), [lm](#), <https://CRAN.R-project.org/package=VGAM>.

Examples

```
# Example 1; proportional odds model
pneumo <- transform(pneumo, let = log(exposure.time))
(fit1 <- vglm(cbind(normal, mild, severe) ~ let, propodds, data = pneumo))
depvar(fit1) # Better than using fit1@y; dependent variable (response)
weights(fit1, type = "prior") # Number of observations
coef(fit1, matrix = TRUE) # p.179, in McCullagh and Nelder (1989)
constraints(fit1) # Constraint matrices
summary(fit1) # HDE could affect these results
summary(fit1, lrt0 = TRUE, score0 = TRUE, wald0 = TRUE) # No HDE
hdeff(fit1) # Check for any Hauck-Donner effect

# Example 2; zero-inflated Poisson model
zdata <- data.frame(x2 = runif(nn <- 2000))
zdata <- transform(zdata, pstr0 = logitlink(-0.5 + 1*x2, inverse = TRUE),
                  lambda = loglink( 0.5 + 2*x2, inverse = TRUE))
zdata <- transform(zdata, y = rzipois(nn, lambda, pstr0 = pstr0))
with(zdata, table(y))
fit2 <- vglm(y ~ x2, zipoisson, data = zdata, trace = TRUE)
coef(fit2, matrix = TRUE) # These should agree with the above values

# Example 3; fit a two species GAM simultaneously
fit3 <- vgam(cbind(agaaus, kniexc) ~ s(altitude, df = c(2, 3)),
            binomialff(multiple.responses = TRUE), data = hunua)
coef(fit3, matrix = TRUE) # Not really interpretable
## Not run: plot(fit3, se = TRUE, overlay = TRUE, lcol = 3:4, scol = 3:4)

ooo <- with(hunua, order(altitude))
with(hunua, matplot(altitude[ooo], fitted(fit3)[ooo, ], type = "l",
                  lwd = 2, col = 3:4,
                  xlab = "Altitude (m)", ylab = "Probability of presence", las = 1,
                  main = "Two plant species' response curves", ylim = c(0, 0.8)))
with(hunua, rug(altitude))
## End(Not run)

# Example 4; LMS quantile regression
fit4 <- vgam(BMI ~ s(age, df = c(4, 2)), lms.bcn(zero = 1),
            data = bmi.nz, trace = TRUE)
head(predict(fit4))
head(fitted(fit4))
head(bmi.nz) # Person 1 is near the lower quartile among people his age
head(cdf(fit4))

## Not run: par(mfrow = c(1,1), bty = "l", mar = c(5,4,4,3)+0.1, xpd=TRUE)
qtplot(fit4, percentiles = c(5,50,90,99), main = "Quantiles", las = 1,
```



```

      xlim = c(15, 90), ylab = "BMI", lwd=2, lcol=4) # Quantile plot

ygrid <- seq(15, 43, len = 100) # BMI ranges
par(mfrow = c(1, 1), lwd = 2) # Density plot
aa <- deplot(fit4, x0 = 20, y = ygrid, xlab = "BMI", col = "black",
  main = "Density functions at Age=20 (black), 42 (red) and 55 (blue)")
aa
aa <- deplot(fit4, x0 = 42, y = ygrid, add = TRUE, llty = 2, col = "red")
aa <- deplot(fit4, x0 = 55, y = ygrid, add = TRUE, llty = 4, col = "blue",
  Attach = TRUE)
aa$post$deplot # Contains density function values

## End(Not run)

# Example 5; GEV distribution for extremes
(fit5 <- vglm(maxtemp ~ 1, gevff, data = oxtemp, trace = TRUE))
head(fitted(fit5))
coef(fit5, matrix = TRUE)
Coef(fit5)
vcov(fit5)
vcov(fit5, untransform = TRUE)
sqrt(diag(vcov(fit5))) # Approximate standard errors
## Not run: rplot(fit5)

```

A1A2A3

*The A1A2A3 Blood Group System***Description**

Estimates the three independent parameters of the the A1A2A3 blood group system.

Usage

```
A1A2A3(link = "logitlink", inbreeding = FALSE, ip1 = NULL, ip2 = NULL, iF = NULL)
```

Arguments

link	Link function applied to p1, p2 and f. See Links for more choices.
inbreeding	Logical. Is there inbreeding?
ip1, ip2, iF	Optional initial value for p1, p2 and f.

Details

The parameters p1 and p2 are probabilities, so that $p_3=1-p_1-p_2$ is the third probability. The parameter f is the third independent parameter if inbreeding = TRUE. If inbreeding = FALSE then $f = 0$ and Hardy-Weinberg Equilibrium (HWE) is assumed.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Note

The input can be a 6-column matrix of counts, with columns corresponding to A1A1, A1A2, A1A3, A2A2, A2A3, A3A3 (in order). Alternatively, the input can be a 6-column matrix of proportions (so each row adds to 1) and the `weights` argument is used to specify the total number of counts for each row.

Author(s)

T. W. Yee

References

Lange, K. (2002). *Mathematical and Statistical Methods for Genetic Analysis*, 2nd ed. New York: Springer-Verlag.

See Also

[AA.Aa.aa](#), [AB.Ab.aB.ab](#), [ABO](#), [MNSs](#).

Examples

```
yamat <- cbind(108, 196, 429, 143, 513, 559)
fit <- vglm(yamat ~ 1, A1A2A3(link = probitlink), trace = TRUE, crit = "coef")
fit <- vglm(yamat ~ 1, A1A2A3(link = logitlink, ip1 = 0.3, ip2 = 0.3, iF = 0.02),
           trace = TRUE, crit = "coef")
Coef(fit) # Estimated p1 and p2
rbind(yamat, sum(yamat) * fitted(fit))
sqrt(diag(vcov(fit)))
```

AA.Aa.aa

The AA-Aa-aa Blood Group System

Description

Estimates the parameter of the AA-Aa-aa blood group system, with or without Hardy Weinberg equilibrium.

Usage

```
AA.Aa.aa(linkp = "logitlink", linkf = "logitlink", inbreeding = FALSE,
         ipA = NULL, ifp = NULL, zero = NULL)
```

Arguments

linkp, linkf	Link functions applied to p_A and f . See Links for more choices.
ipA, ifp	Optional initial values for p_A and f .
inbreeding	Logical. Is there inbreeding?
zero	See CommonVGAMffArguments for information.

Details

This one or two parameter model involves a probability called p_A . The probability of getting a count in the first column of the input (an AA) is $p_A * p_A$. When `inbreeding = TRUE`, an additional parameter f is used. If `inbreeding = FALSE` then $f = 0$ and Hardy-Weinberg Equilibrium (HWE) is assumed. The EIM is used if `inbreeding = FALSE`.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Warning

Setting `inbreeding = FALSE` makes estimation difficult with non-intercept-only models. Currently, this code seems to work with intercept-only models.

Note

The input can be a 3-column matrix of counts, where the columns are AA, Ab and aa (in order). Alternatively, the input can be a 3-column matrix of proportions (so each row adds to 1) and the `weights` argument is used to specify the total number of counts for each row.

Author(s)

T. W. Yee

References

Weir, B. S. (1996). *Genetic Data Analysis II: Methods for Discrete Population Genetic Data*, Sunderland, MA: Sinauer Associates, Inc.

See Also

[AB.Ab.aB.ab](#), [ABO](#), [A1A2A3](#), [MNSs](#).

Examples

```
y <- cbind(53, 95, 38)
fit1 <- vglm(y ~ 1, AA.Aa.aa, trace = TRUE)
fit2 <- vglm(y ~ 1, AA.Aa.aa(inbreeding = TRUE), trace = TRUE)
rbind(y, sum(y) * fitted(fit1))
Coef(fit1) # Estimated pA
```

```
Coef(fit2) # Estimated pA and f
summary(fit1)
```

AB.Ab.aB.ab

The AB-Ab-aB-ab Blood Group System

Description

Estimates the parameter of the AB-Ab-aB-ab blood group system.

Usage

```
AB.Ab.aB.ab(link = "logitlink", init.p = NULL)
```

Arguments

`link` Link function applied to p. See [Links](#) for more choices.
`init.p` Optional initial value for p.

Details

This one parameter model involves a probability called p.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Note

The input can be a 4-column matrix of counts, where the columns are AB, Ab, aB and ab (in order). Alternatively, the input can be a 4-column matrix of proportions (so each row adds to 1) and the `weights` argument is used to specify the total number of counts for each row.

Author(s)

T. W. Yee

References

Lange, K. (2002). *Mathematical and Statistical Methods for Genetic Analysis*, 2nd ed. New York: Springer-Verlag.

See Also

[AA.Aa.aa](#), [ABO](#), [A1A2A3](#), [MNSS](#).

Examples

```

ymat <- cbind(AB=1997, Ab=906, aB=904, ab=32) # Data from Fisher (1925)
fit <- vglm(ymat ~ 1, AB.Ab.aB.ab(link = "identitylink"), trace = TRUE)
fit <- vglm(ymat ~ 1, AB.Ab.aB.ab, trace = TRUE)
rbind(ymat, sum(ymat)*fitted(fit))
Coef(fit) # Estimated p
p <- sqrt(4*(fitted(fit)[, 4]))
p*p
summary(fit)

```

ABO

*The ABO Blood Group System***Description**

Estimates the two independent parameters of the the ABO blood group system.

Usage

```

ABO(link.pA = "logitlink", link.pB = "logitlink", ipA = NULL, ipB = NULL,
    ipO = NULL, zero = NULL)

```

Arguments

link.pA, link.pB	Link functions applied to pA and pB. See Links for more choices.
ipA, ipB, ipO	Optional initial value for pA and pB and pO. A NULL value means values are computed internally.
zero	Details at CommonVGAMffArguments .

Details

The parameters pA and pB are probabilities, so that $pO=1-pA-pB$ is the third probability. The probabilities pA and pB correspond to A and B respectively, so that pO is the probability for O. It is easier to make use of initial values for pO than for pB. In documentation elsewhere I sometimes use $pA=p$, $pB=q$, $pO=r$.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Note

The input can be a 4-column matrix of counts, where the columns are A, B, AB, O (in order). Alternatively, the input can be a 4-column matrix of proportions (so each row adds to 1) and the weights argument is used to specify the total number of counts for each row.

Author(s)

T. W. Yee

References

Lange, K. (2002). *Mathematical and Statistical Methods for Genetic Analysis*, 2nd ed. New York: Springer-Verlag.

See Also

[AA.Aa.aa](#), [AB.Ab.aB.ab](#), [A1A2A3](#), [MNSs](#).

Examples

```
yamat <- cbind(A = 725, B = 258, AB = 72, O = 1073) # Order matters, not the name
fit <- vglm(yamat ~ 1, ABO(link.pA = "identitylink",
                        link.pB = "identitylink"), trace = TRUE,
           crit = "coef")
coef(fit, matrix = TRUE)
Coef(fit) # Estimated pA and pB
rbind(yamat, sum(yamat) * fitted(fit))
sqrt(diag(vcov(fit)))
```

acat

*Ordinal Regression with Adjacent Categories Probabilities***Description**

Fits an adjacent categories regression model to an ordered (preferably) factor response.

Usage

```
acat(link = "loglink", parallel = FALSE, reverse = FALSE,
     zero = NULL, ynames = FALSE, Thresh = NULL, Trev = reverse,
     Tref = if (Trev) "M" else 1, whitespace = FALSE)
```

Arguments

link	Link function applied to the ratios of the adjacent categories probabilities. See Links for more choices.
parallel	A logical, or formula specifying which terms have equal/unequal coefficients.
reverse	Logical. By default, the linear/additive predictors used are $\eta_j = \log(P[Y = j + 1]/P[Y = j])$ for $j = 1, \dots, M$. If reverse is TRUE then $\eta_j = \log(P[Y = j]/P[Y = j + 1])$ will be used.
ynames	See multinomial for information.

- zero An integer-valued vector specifying which linear/additive predictors are modelled as intercepts only. The values must be from the set $\{1, 2, \dots, M\}$. See [CommonVGAMffArguments](#) for more information.
- Thresh, Trev, Tref See [cumulative](#) for information. These arguments apply to ordinal categorical regression models.
- whitespace See [CommonVGAMffArguments](#) for information.

Details

In this help file the response Y is assumed to be a factor with ordered values $1, 2, \dots, M + 1$, so that M is the number of linear/additive predictors η_j . By default, the log link is used because the ratio of two probabilities is positive.

Internally, [deriv3](#) is called to perform symbolic differentiation and consequently this family function will struggle if M becomes too large. If this occurs, try combining levels so that M is effectively reduced. One idea is to aggregate levels with the fewest observations in them first.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), [rrvglm](#) and [vgam](#).

Warning

No check is made to verify that the response is ordinal if the response is a matrix; see [ordered](#).

Note

The response should be either a matrix of counts (with row sums that are all positive), or an ordered factor. In both cases, the y slot returned by [vglm/vgam/rrvglm](#) is the matrix of counts.

For a nominal (unordered) factor response, the multinomial logit model ([multinomial](#)) is more appropriate.

Here is an example of the usage of the `parallel` argument. If there are covariates x_1 , x_2 and x_3 , then `parallel = TRUE ~ x1 + x2 - 1` and `parallel = FALSE ~ x3` are equivalent. This would constrain the regression coefficients for x_1 and x_2 to be equal; those of the intercepts and x_3 would be different.

Author(s)

Thomas W. Yee

References

Agresti, A. (2013). *Categorical Data Analysis*, 3rd ed. Hoboken, NJ, USA: Wiley.
 Tutz, G. (2012). *Regression for Categorical Data*, Cambridge: Cambridge University Press.
 Yee, T. W. (2010). The **VGAM** package for categorical data analysis. *Journal of Statistical Software*, **32**, 1–34. doi:10.18637/jss.v032.i10.

See Also

[cumulative](#), [cratio](#), [sratio](#), [multinomial](#), [CM.equid](#), [CommonVGAMffArguments](#), [margeff](#), [pneumo](#), [budworm](#), [deriv3](#).

Examples

```
pneumo <- transform(pneumo, let = log(exposure.time))
(fit <- vglm(cbind(normal, mild, severe) ~ let, acat, pneumo))
coef(fit, matrix = TRUE)
constraints(fit)
model.matrix(fit)
```

add1.vglm

Add or Drop All Possible Single Terms to/from a Model

Description

Compute all the single terms in the scope argument that can be added to or dropped from the model, fit those models and compute a table of the changes in fit.

Usage

```
## S3 method for class 'vglm'
add1(object, scope, test = c("none", "LRT"), k = 2, ...)
## S3 method for class 'vglm'
drop1(object, scope, test = c("none", "LRT"), k = 2, ...)
```

Arguments

object	a fitted vglm model object.
scope, k	See drop1.glm .
test	Same as drop1.glm but with fewer choices.
...	further arguments passed to or from other methods.

Details

These functions are a direct adaptation of [add1.glm](#) and [drop1.glm](#) for [vglm-class](#) objects. For [drop1](#) methods, a missing scope is taken to be all terms in the model. The hierarchy is respected when considering terms to be added or dropped: all main effects contained in a second-order interaction must remain, and so on. In a scope formula `.` means ‘what is already there’.

Compared to [add1.glm](#) and [drop1.glm](#) these functions are simpler, e.g., there is no C_p , F and Rao (score) tests, x and scale arguments. Most models do not have a deviance, however twice the log-likelihood differences are used to test the significance of terms.

The default output table gives AIC, defined as minus twice log likelihood plus $2p$ where p is the rank of the model (the number of effective parameters). This is only defined up to an additive constant (like log-likelihoods).

Value

An object of class "anova" summarizing the differences in fit between the models.

Warning

In general, the same warnings in `add1.glm` and `drop1.glm` apply here. Furthermore, these functions have not been rigorously tested for all models, so treat the results cautiously and please report any bugs.

Care is needed to check that the constraint matrices of added terms are correct. Also, if object is of the form `vglm(..., constraints = list(x1 = cm1, x2 = cm2))` then `add1.vglm` may fail because the `constraints` argument needs to have the constraint matrices for *all* terms.

Note

Most **VGAM** family functions do not compute a deviance, but instead the likelihood function is evaluated at the MLE. Hence a column name "Deviance" only appears for a few models; and almost always there is a column labelled "logLik".

See Also

[step4vglm](#), [vglm](#), [extractAIC.vglm](#), [trim.constraints](#), [anova.vglm](#), [backPain2](#), [update](#).

Examples

```
data("backPain2", package = "VGAM")
summary(backPain2)
fit1 <- vglm(pain ~ x2 + x3 + x4, propodds, data = backPain2)
coef(fit1)
add1(fit1, scope = ~ x2 * x3 * x4, test = "LRT")
drop1(fit1, test = "LRT")
fit2 <- vglm(pain ~ x2 * x3 * x4, propodds, data = backPain2)
drop1(fit2)
```

AICv1m

Akaike's Information Criterion

Description

Calculates the Akaike information criterion for a fitted model object for which a log-likelihood value has been obtained.

Usage

```
AICv1m(object, ..., corrected = FALSE, k = 2)
AICvgam(object, ..., k = 2)
AICrrvglm(object, ..., k = 2)
AICdrrvglm(object, ..., k = 2)
AICqrrvglm(object, ..., k = 2)
AICrrvgam(object, ..., k = 2)
```

Arguments

object	Some VGAM object, for example, having class <code>vglm-class</code> .
...	Other possible arguments fed into <code>logLik</code> in order to compute the log-likelihood.
corrected	Logical, perform the finite sample correction?
k	Numeric, the penalty per parameter to be used; the default is the classical AIC.

Details

The following formula is used for VGLMs: $-2\log\text{-likelihood} + kn_{par}$, where n_{par} represents the number of parameters in the fitted model, and $k = 2$ for the usual AIC. One could assign $k = \log(n)$ (n the number of observations) for the so-called BIC or SBC (Schwarz's Bayesian criterion). This is the function `AICvIm()`.

This code relies on the log-likelihood being defined, and computed, for the object. When comparing fitted objects, the smaller the AIC, the better the fit. The log-likelihood and hence the AIC is only defined up to an additive constant.

Any estimated scale parameter (in GLM parlance) is used as one parameter.

For VGAMs and CAO the nonlinear effective degrees of freedom for each smoothed component is used. This formula is heuristic. These are the functions `AICvgam()` and `AICcao()`.

The finite sample correction is usually recommended when the sample size is small or when the number of parameters is large. When the sample size is large their difference tends to be negligible. The correction is described in Hurvich and Tsai (1989), and is based on a (univariate) linear model with normally distributed errors.

Value

Returns a numeric value with the corresponding AIC (or BIC, or ..., depending on `k`).

Warning

This code has not been double-checked. The general applicability of AIC for the VGLM/VGAM classes has not been developed fully. In particular, AIC should not be run on some **VGAM** family functions because of violation of certain regularity conditions, etc.

Note

AIC has not been defined for QRR-VGLMs, yet.

Using AIC to compare `posbinomial` models with, e.g., `posbernoulli.tb` models, requires `posbinomial(omit.constant = TRUE)`. See `posbinomial` for an example. A warning is given if it suspects a wrong `omit.constant` value was used.

Where defined, `AICc(...)` is the same as `AIC(..., corrected = TRUE)`.

Author(s)

T. W. Yee.

References

Hurvich, C. M. and Tsai, C.-L. (1989). Regression and time series model selection in small samples, *Biometrika*, **76**, 297–307.

See Also

VGLMs are described in [vglm-class](#); VGAMs are described in [vgam-class](#); RR-VGLMs are described in [rrvglm-class](#); [AIC](#), [BICv1m](#), [TICv1m](#), [drop1.vglm](#), [extractAIC.vglm](#).

Examples

```
pneumo <- transform(pneumo, let = log(exposure.time))
(fit1 <- vglm(cbind(normal, mild, severe) ~ let,
             cumulative(parallel = TRUE, reverse = TRUE), data = pneumo))
coef(fit1, matrix = TRUE)
AIC(fit1)
AICc(fit1) # Quick way
AIC(fit1, corrected = TRUE) # Slow way
(fit2 <- vglm(cbind(normal, mild, severe) ~ let,
             cumulative(parallel = FALSE, reverse = TRUE), data = pneumo))
coef(fit2, matrix = TRUE)
AIC(fit2)
AICc(fit2)
AIC(fit2, corrected = TRUE)
```

alogitlink

Arcsine–Logit Link Mixtures

Description

Computes some arcsine–logit mixture link transformations, including their inverse and the first few derivatives.

Usage

```
alogitlink(theta, bvalue = NULL, taumix.logit = 1,
           tol = 1e-13, nmax = 99, inverse = FALSE, deriv = 0,
           short = TRUE, tag = FALSE, c10 = c(4, -pi))
lcalogitlink(theta, bvalue = NULL, pmix.logit = 0.01,
            tol = 1e-13, nmax = 99, inverse = FALSE, deriv = 0,
            short = TRUE, tag = FALSE, c10 = c(4, -pi))
```

Arguments

theta Numeric or character. See below for further details.
bvalue See [Links](#).

taumix.logit	Numeric, of length 1. Mixing parameter assigned to logitlink . Then $1 - \exp(-\text{taumix.log} * \text{theta})$ is used to weight asinlink . Thus a 0 value will result in logitlink and a very large numeric such as 1e4 should be roughly equivalent to asinlink over almost all of the parameter space.
pmix.logit	Numeric, of length 1. Mixing probability assigned to logitlink . Then $1 - \text{pmix.logit}$ is used to weight asinlink . Thus a 0 value will result in asinlink and 1 is equivalent to logitlink .
tol, nmax inverse, deriv, short, tag	Arguments fed into a function implementing a vectorized bisection method. Details at Links .
c10	See asinlink and logitlink .

Details

[lcalogitlink](#) is a *linear combination* (LC) of [asinlink](#) and [logitlink](#).

Value

The following holds for the LC variant. For $\text{deriv} \geq 0$, $(1 - \text{pmix.logit}) * \text{asinlink}(p, \text{deriv} = \text{deriv}) + \text{pmix.logit} * \text{logitlink}(p, \text{deriv} = \text{deriv})$ when $\text{inverse} = \text{FALSE}$, and if $\text{inverse} = \text{TRUE}$ then a nonlinear equation is solved for the probability, given eta. For $\text{deriv} = 1$, then the function returns $d \text{ eta} / d \text{ theta}$ as a function of theta if $\text{inverse} = \text{FALSE}$, else if $\text{inverse} = \text{TRUE}$ then it returns the reciprocal.

Warning

The default values for `taumix.logit` and `pmix.logit` may change in the future. The name and order of the arguments may change too.

Author(s)

Thomas W. Yee

References

Hauck, J. W. W. and A. Donner (1977). Wald's test as applied to hypotheses in logit analysis. *Journal of the American Statistical Association*, **72**, 851–853.

See Also

[asinlink](#), [logitlink](#), [Links](#), [probitlink](#), [clogloglink](#), [cauchitlink](#), [binomialff](#), [sloglink](#), [hdeff](#), <https://www.cia.gov/index.html>.

Examples

```
p <- seq(0.01, 0.99, length= 10)
alogitlink(p)
max(abs(alogitlink(alogitlink(p), inv = TRUE) - p)) # 0?
```

```

## Not run:
par(mfrow = c(2, 2), lwd = (mylwd <- 2))
y <- seq(-4, 4, length = 100)
p <- seq(0.01, 0.99, by = 0.01)

for (d in 0:1) {
  matplot(p, cbind(logitlink(p, deriv = d), probitlink(p, deriv = d)),
           type = "n", col = "blue", ylab = "transformation",
           las = 1, main = if (d == 0) "Some probability link functions"
           else "First derivative")
  lines(p, logitlink(p, deriv = d), col = "green")
  lines(p, probitlink(p, deriv = d), col = "blue")
  lines(p, clogloglink(p, deriv = d), col = "tan")
  lines(p, alogitlink(p, deriv = d), col = "red3")
  if (d == 0) {
    abline(v = 0.5, h = 0, lty = "dashed")
    legend(0, 4.5, c("logitlink", "probitlink", "clogloglink",
                   "alogitlink"), lwd = mylwd,
           col = c("green", "blue", "tan", "red3"))
  } else
    abline(v = 0.5, lwd = 0.5, col = "gray")
}

for (d in 0) {
  matplot(y, cbind(logitlink(y, deriv = d, inverse = TRUE),
                  probitlink(y, deriv = d, inverse = TRUE)),
           type = "n", col = "blue", xlab = "transformation", ylab = "p",
           main = if (d == 0) "Some inverse probability link functions"
           else "First derivative", las=1)
  lines(y, logitlink(y, deriv = d, inverse = TRUE), col = "green")
  lines(y, probitlink(y, deriv = d, inverse = TRUE), col = "blue")
  lines(y, clogloglink(y, deriv = d, inverse = TRUE), col = "tan")
  lines(y, alogitlink(y, deriv = d, inverse = TRUE), col = "red3")
  if (d == 0) {
    abline(h = 0.5, v = 0, lwd = 0.5, col = "gray")
    legend(-4, 1, c("logitlink", "probitlink", "clogloglink",
                   "alogitlink"), lwd = mylwd,
           col = c("green", "blue", "tan", "red3"))
  }
}
par(lwd = 1)

## End(Not run)

```

altered

*Altered, Inflated, Truncated and Deflated Values in GAITD Regression***Description**

Return the altered, inflated, truncated and deflated values in a GAITD regression object, else test whether the model is altered, inflated, truncated or deflated.

Usage

```
altered(object, ...)
inflated(object, ...)
truncated(object, ...)
is.altered(object, ...)
is.deflated(object, ...)
is.inflated(object, ...)
is.truncated(object, ...)
```

Arguments

`object` an object of class "vglm". Currently only a GAITD regression object returns valid results of these functions.

`...` any additional arguments, to future-proof this function.

Details

Yee and Ma (2023) propose GAITD regression where values from four (or seven since there are parametric and nonparametric forms) disjoint sets are referred to as *special*. These extractor functions return one set each; they are the alter, inflate, truncate, deflate (and sometimes `max.support`) arguments from the family function.

Value

Returns one type of 'special' sets associated with GAITD regression. This is a vector, else a list for truncation. All three sets are returned by `specialsvglm`.

Warning

Some of these functions are subject to change. Only family functions beginning with "gaitd" will work with these functions, hence `zipoisson` fits will return FALSE or empty values.

References

Yee, T. W. and Ma, C. (2024). Generally altered, inflated, truncated and deflated regression. *Statistical Science*, **39** (in press).

See Also

[vglm](#), [vglm-class](#), [specialsvglm](#), [gaitdpoisson](#), [gaitdlog](#), [gaitdzeta](#), [Gaitdpois](#).

Examples

```
## Not run:
abdata <- data.frame(y = 0:7, w = c(182, 41, 12, 2, 2, 0, 0, 1))
fit1 <- vglm(y ~ 1, gaitdpoisson(a.mix = 0),
            data = abdata, weight = w, subset = w > 0)
specials(fit1) # All three sets
altered(fit1) # Subject to change
inflated(fit1) # Subject to change
```

```
truncated(fit1) # Subject to change
is.altered(fit1)
is.inflated(fit1)
is.truncated(fit1)
## End(Not run)
```

amlbinomial	<i>Binomial Logistic Regression by Asymmetric Maximum Likelihood Estimation</i>
-------------	---

Description

Binomial quantile regression estimated by maximizing an asymmetric likelihood function.

Usage

```
amlbinomial(w.aml = 1, parallel = FALSE, digw = 4, link = "logitlink")
```

Arguments

w.aml	Numeric, a vector of positive constants controlling the percentiles. The larger the value the larger the fitted percentile value (the proportion of points below the “w-regression plane”). The default value of unity results in the ordinary maximum likelihood (MLE) solution.
parallel	If w.aml has more than one value then this argument allows the quantile curves to differ by the same amount as a function of the covariates. Setting this to be TRUE should force the quantile curves to not cross (although they may not cross anyway). See CommonVGAMffArguments for more information.
digw	Passed into Round as the digits argument for the w.aml values; used cosmetically for labelling.
link	See binomialff .

Details

The general methodology behind this **VGAM** family function is given in Efron (1992) and full details can be obtained there. This model is essentially a logistic regression model (see [binomialff](#)) but the usual deviance is replaced by an asymmetric squared error loss function; it is multiplied by *w.aml* for positive residuals. The solution is the set of regression coefficients that minimize the sum of these deviance-type values over the data set, weighted by the *weights* argument (so that it can contain frequencies). Newton-Raphson estimation is used here.

Value

An object of class “*vglmff*” (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Warning

If `w.aml` has more than one value then the value returned by deviance is the sum of all the (weighted) deviances taken over all the `w.aml` values. See Equation (1.6) of Efron (1992).

Note

On fitting, the extra slot has list components "w.aml" and "percentile". The latter is the percent of observations below the "w-regression plane", which is the fitted values. Also, the individual deviance values corresponding to each element of the argument `w.aml` is stored in the extra slot.

For `amlbinomial` objects, methods functions for the generic functions `qtplot` and `cdf` have not been written yet.

See [amlpoisson](#) about comments on the jargon, e.g., *expectiles* etc.

In this documentation the word *quantile* can often be interchangeably replaced by *expectile* (things are informal here).

Author(s)

Thomas W. Yee

References

Efron, B. (1992). Poisson overdispersion estimates based on the method of asymmetric maximum likelihood. *Journal of the American Statistical Association*, **87**, 98–107.

See Also

[amlpoisson](#), [amlexponential](#), [amlnormal](#), [extlogF1](#), [alaplace1](#), [denorm](#).

Examples

```
# Example: binomial data with lots of trials per observation
set.seed(1234)
sizevec <- rep(100, length = (nn <- 200))
mydat <- data.frame(x = sort(runif(nn)))
mydat <- transform(mydat,
                  prob = logitlink(-0 + 2.5*x + x^2, inverse = TRUE))
mydat <- transform(mydat, y = rbinom(nn, size = sizevec, prob = prob))
(fit <- vgam(cbind(y, sizevec - y) ~ s(x, df = 3),
            amlbinomial(w = c(0.01, 0.2, 1, 5, 60)),
            mydat, trace = TRUE))

fit@extra

## Not run:
par(mfrow = c(1,2))
# Quantile plot
with(mydat, plot(x, jitter(y), col = "blue", las = 1, main =
  paste(paste(round(fit@extra$percentile, digits = 1), collapse = ", "),
    "percentile-expectile curves")))
with(mydat, matlines(x, 100 * fitted(fit), lwd = 2, col = "blue", lty=1))
```



```

# Compare the fitted expectiles with the quantiles
with(mydat, plot(x, jitter(y), col = "blue", las = 1, main =
  paste(paste(round(fit@extra$percentile, digits = 1), collapse = ", "),
    "percentile curves are red")))
with(mydat, matlines(x, 100 * fitted(fit), lwd = 2, col = "blue", lty = 1))

for (ii in fit@extra$percentile)
  with(mydat, matlines(x, 100 *
    qbinom(p = ii/100, size = sizevec, prob = prob) / sizevec,
      col = "red", lwd = 2, lty = 1))

## End(Not run)

```

amlexponential	<i>Exponential Regression by Asymmetric Maximum Likelihood Estimation</i>
----------------	---

Description

Exponential expectile regression estimated by maximizing an asymmetric likelihood function.

Usage

```
amlexponential(w.aml = 1, parallel = FALSE, imethod = 1, digw = 4,
  link = "loglink")
```

Arguments

w.aml	Numeric, a vector of positive constants controlling the expectiles. The larger the value the larger the fitted expectile value (the proportion of points below the “w-regression plane”). The default value of unity results in the ordinary maximum likelihood (MLE) solution.
parallel	If w.aml has more than one value then this argument allows the quantile curves to differ by the same amount as a function of the covariates. Setting this to be TRUE should force the quantile curves to not cross (although they may not cross anyway). See CommonVGAMffArguments for more information.
imethod	Integer, either 1 or 2 or 3. Initialization method. Choose another value if convergence fails.
digw	Passed into Round as the digits argument for the w.aml values; used cosmetically for labelling.
link	See exponential and the warning below.

Details

The general methodology behind this **VGAM** family function is given in Efron (1992) and full details can be obtained there. This model is essentially an exponential regression model (see [exponential](#)) but the usual deviance is replaced by an asymmetric squared error loss function; it is multiplied by *w.aml* for positive residuals. The solution is the set of regression coefficients that minimize the sum of these deviance-type values over the data set, weighted by the weights argument (so that it can contain frequencies). Newton-Raphson estimation is used here.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Warning

Note that the link argument of [exponential](#) and [amlexponential](#) are currently different: one is the rate parameter and the other is the mean (expectile) parameter.

If *w.aml* has more than one value then the value returned by deviance is the sum of all the (weighted) deviances taken over all the *w.aml* values. See Equation (1.6) of Efron (1992).

Note

On fitting, the extra slot has list components "w.aml" and "percentile". The latter is the percent of observations below the "w-regression plane", which is the fitted values. Also, the individual deviance values corresponding to each element of the argument *w.aml* is stored in the extra slot.

For [amlexponential](#) objects, methods functions for the generic functions `qtplot` and `cdf` have not been written yet.

See [amlpoisson](#) about comments on the jargon, e.g., *expectiles* etc.

In this documentation the word *quantile* can often be interchangeably replaced by *expectile* (things are informal here).

Author(s)

Thomas W. Yee

References

Efron, B. (1992). Poisson overdispersion estimates based on the method of asymmetric maximum likelihood. *Journal of the American Statistical Association*, **87**, 98–107.

See Also

[exponential](#), [amlbinomial](#), [amlpoisson](#), [amlnormal](#), [extlogF1](#), [alaplacel](#), [lms.bcg](#), [deexp](#).

Examples

```

nn <- 2000
mydat <- data.frame(x = seq(0, 1, length = nn))
mydat <- transform(mydat,
                  mu = loglink(-0 + 1.5*x + 0.2*x^2, inverse = TRUE))
mydat <- transform(mydat, mu = loglink(0 - sin(8*x), inverse = TRUE))
mydat <- transform(mydat, y = rexp(nn, rate = 1/mu))
(fit <- vgam(y ~ s(x, df=5), amlexponential(w=c(0.001, 0.1, 0.5, 5, 60)),
            mydat, trace = TRUE))
fit@extra

## Not run: # These plots are against the sqrt scale (to increase clarity)
par(mfrow = c(1,2))
# Quantile plot
with(mydat, plot(x, sqrt(y), col = "blue", las = 1, main =
                paste(paste(round(fit@extra$percentile, digits = 1), collapse=" ",
                             "percentile-expectile curves"))))
with(mydat, matlines(x, sqrt(fitted(fit)), lwd = 2, col = "blue", lty=1))

# Compare the fitted expectiles with the quantiles
with(mydat, plot(x, sqrt(y), col = "blue", las = 1, main =
                paste(paste(round(fit@extra$percentile, digits = 1), collapse=" ",
                             "percentile curves are orange"))))
with(mydat, matlines(x, sqrt(fitted(fit)), lwd = 2, col = "blue", lty=1))

for (ii in fit@extra$percentile)
  with(mydat, matlines(x, sqrt(qexp(p = ii/100, rate = 1/mu)),
                      col = "orange"))

## End(Not run)

```

amlnormal

*Asymmetric Least Squares Quantile Regression***Description**

Asymmetric least squares, a special case of maximizing an asymmetric likelihood function of a normal distribution. This allows for expectile/quantile regression using asymmetric least squares error loss.

Usage

```
amlnormal(w.aml = 1, parallel = FALSE, lexpectile = "identitylink",
          iexpectile = NULL, imethod = 1, digw = 4)
```

Arguments

w.aml Numeric, a vector of positive constants controlling the percentiles. The larger the value the larger the fitted percentile value (the proportion of points below the “w-regression plane”). The default value of unity results in the ordinary least squares (OLS) solution.

parallel	If <code>w.aml</code> has more than one value then this argument allows the quantile curves to differ by the same amount as a function of the covariates. Setting this to be TRUE should force the quantile curves to not cross (although they may not cross anyway). See CommonVGAMffArguments for more information.
lexpectile, iexpectile	See CommonVGAMffArguments for more information.
imethod	Integer, either 1 or 2 or 3. Initialization method. Choose another value if convergence fails.
digw	Passed into Round as the <code>digits</code> argument for the <code>w.aml</code> values; used cosmetically for labelling.

Details

This is an implementation of Efron (1991) and full details can be obtained there. Equation numbers below refer to that article. The model is essentially a linear model (see [lm](#)), however, the asymmetric squared error loss function for a residual r is r^2 if $r \leq 0$ and wr^2 if $r > 0$. The solution is the set of regression coefficients that minimize the sum of these over the data set, weighted by the weights argument (so that it can contain frequencies). Newton-Raphson estimation is used here.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Note

On fitting, the extra slot has list components "w.aml" and "percentile". The latter is the percent of observations below the "w-regression plane", which is the fitted values.

One difficulty is finding the `w.aml` value giving a specified percentile. One solution is to fit the model within a root finding function such as [uniroot](#); see the example below.

For `amlnormal` objects, methods functions for the generic functions `qtplot` and `cdf` have not been written yet.

See the note in [amlpoisson](#) on the jargon, including *expectiles* and *regression quantiles*.

The deviance slot computes the total asymmetric squared error loss (2.5). If `w.aml` has more than one value then the value returned by the slot is the sum taken over all the `w.aml` values.

This **VGAM** family function could well be renamed `amlnormal()` instead, given the other function names [amlpoisson](#), [amlbinomial](#), etc.

In this documentation the word *quantile* can often be interchangeably replaced by *expectile* (things are informal here).

Author(s)

Thomas W. Yee

References

Efron, B. (1991). Regression percentiles using asymmetric squared error loss. *Statistica Sinica*, **1**, 93–125.

See Also

[amlpoisson](#), [amlbinomial](#), [amlexponential](#), [bmi.nz](#), [extlogF1](#), [alaplacel](#), [denorm](#), [lms.bcn](#) and similar variants are alternative methods for quantile regression.

Examples

```
## Not run:
# Example 1
ooo <- with(bmi.nz, order(age))
bmi.nz <- bmi.nz[ooo, ] # Sort by age
(fit <- vglm(BMI ~ sm.bs(age), amlnormal(w.aml = 0.1), bmi.nz))
fit@extra # Gives the w value and the percentile
coef(fit, matrix = TRUE)

# Quantile plot
with(bmi.nz, plot(age, BMI, col = "blue", main =
  paste(round(fit@extra$percentile, digits = 1),
    "expectile-percentile curve")))
with(bmi.nz, lines(age, c(fitted(fit)), col = "black"))

# Example 2
# Find the w values that give the 25, 50 and 75 percentiles
find.w <- function(w, percentile = 50) {
  fit2 <- vglm(BMI ~ sm.bs(age), amlnormal(w = w), data = bmi.nz)
  fit2@extra$percentile - percentile
}
# Quantile plot
with(bmi.nz, plot(age, BMI, col = "blue", las = 1, main =
  "25, 50 and 75 expectile-percentile curves"))
for (myp in c(25, 50, 75)) {
# Note: uniroot() can only find one root at a time
  bestw <- uniroot(f = find.w, interval = c(1/10^4, 10^4),
    percentile = myp)
  fit2 <- vglm(BMI ~ sm.bs(age), amlnormal(w = bestw$root), bmi.nz)
  with(bmi.nz, lines(age, c(fitted(fit2)), col = "orange"))
}

# Example 3; this is Example 1 but with smoothing splines and
# a vector w and a parallelism assumption.
ooo <- with(bmi.nz, order(age))
bmi.nz <- bmi.nz[ooo, ] # Sort by age
fit3 <- vgam(BMI ~ s(age, df = 4), data = bmi.nz, trace = TRUE,
  amlnormal(w = c(0.1, 1, 10), parallel = TRUE))
fit3@extra # The w values, percentiles and weighted deviances

# The linear components of the fit; not for human consumption:
coef(fit3, matrix = TRUE)

# Quantile plot
with(bmi.nz, plot(age, BMI, col="blue", main =
  paste(paste(round(fit3@extra$percentile, digits = 1), collapse = ", "),
    "expectile-percentile curves")))
```

```
with(bmi.nz, matlines(age, fitted(fit3), col = 1:fit3@extra$M, lwd = 2))
with(bmi.nz, lines(age, c(fitted(fit )), col = "black")) # For comparison

## End(Not run)
```

amlpoisson

*Poisson Regression by Asymmetric Maximum Likelihood Estimation***Description**

Poisson quantile regression estimated by maximizing an asymmetric likelihood function.

Usage

```
amlpoisson(w.aml = 1, parallel = FALSE, imethod = 1, digw = 4,
           link = "loglink")
```

Arguments

w.aml	Numeric, a vector of positive constants controlling the percentiles. The larger the value the larger the fitted percentile value (the proportion of points below the “w-regression plane”). The default value of unity results in the ordinary maximum likelihood (MLE) solution.
parallel	If w.aml has more than one value then this argument allows the quantile curves to differ by the same amount as a function of the covariates. Setting this to TRUE should force the quantile curves to not cross (although they may not cross anyway). See CommonVGAMffArguments for more information.
imethod	Integer, either 1 or 2 or 3. Initialization method. Choose another value if convergence fails.
digw	Passed into Round as the digits argument for the w.aml values; used cosmetically for labelling.
link	See poissonff .

Details

This method was proposed by Efron (1992) and full details can be obtained there. The model is essentially a Poisson regression model (see [poissonff](#)) but the usual deviance is replaced by an asymmetric squared error loss function; it is multiplied by *w.aml* for positive residuals. The solution is the set of regression coefficients that minimize the sum of these deviance-type values over the data set, weighted by the weights argument (so that it can contain frequencies). Newton-Raphson estimation is used here.

Value

An object of class “vglmff” (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Warning

If `w.aml` has more than one value then the value returned by deviance is the sum of all the (weighted) deviances taken over all the `w.aml` values. See Equation (1.6) of Efron (1992).

Note

On fitting, the extra slot has list components "w.aml" and "percentile". The latter is the percent of observations below the "w-regression plane", which is the fitted values. Also, the individual deviance values corresponding to each element of the argument `w.aml` is stored in the extra slot.

For `amlpoisson` objects, methods functions for the generic functions `qtpplot` and `cdf` have not been written yet.

About the jargon, Newey and Powell (1987) used the name *expectiles* for regression surfaces obtained by asymmetric least squares. This was deliberate so as to distinguish them from the original *regression quantiles* of Koenker and Bassett (1978). Efron (1991) and Efron (1992) use the general name *regression percentile* to apply to all forms of asymmetric fitting. Although the asymmetric maximum likelihood method very nearly gives regression percentiles in the strictest sense for the normal and Poisson cases, the phrase *quantile regression* is used loosely in this **VGAM** documentation.

In this documentation the word *quantile* can often be interchangeably replaced by *expectile* (things are informal here).

Author(s)

Thomas W. Yee

References

- Efron, B. (1991). Regression percentiles using asymmetric squared error loss. *Statistica Sinica*, **1**, 93–125.
- Efron, B. (1992). Poisson overdispersion estimates based on the method of asymmetric maximum likelihood. *Journal of the American Statistical Association*, **87**, 98–107.
- Koenker, R. and Bassett, G. (1978). Regression quantiles. *Econometrica*, **46**, 33–50.
- Newey, W. K. and Powell, J. L. (1987). Asymmetric least squares estimation and testing. *Econometrica*, **55**, 819–847.

See Also

[amlnormal](#), [amlbinomial](#), [extlogF1](#), [alaplace1](#).

Examples

```
set.seed(1234)
mydat <- data.frame(x = sort(runif(nn <- 200)))
mydat <- transform(mydat, y = rpois(nn, exp(0 - sin(8*x))))
(fit <- vgam(y ~ s(x), fam = amlpoisson(w.aml = c(0.02, 0.2, 1, 5, 50)),
           mydat, trace = TRUE))
fit@extra
```

```
## Not run:
# Quantile plot
with(mydat, plot(x, jitter(y), col = "blue", las = 1, main =
  paste(paste(round(fit@extra$percentile, digits = 1), collapse = ", "),
    "percentile-expectile curves")))
with(mydat, matlines(x, fitted(fit), lwd = 2))
## End(Not run)
```

anova.vglm

*Analysis of Deviance for Vector Generalized Linear Model Fits***Description**

Compute an analysis of deviance table for one or more vector generalized linear model fits.

Usage

```
## S3 method for class 'vglm'
anova(object, ..., type = c("II", "I", "III", 2, 1, 3),
  test = c("LRT", "none"), trydev = TRUE, silent = TRUE)
```

Arguments

object, ...	objects of class <code>vglm</code> , typically the result of a call to <code>vglm</code> , or a list of objects for the <code>"vglm"</code> method. Each model must have an intercept term. If <code>"vglm"</code> is used then <code>type = 1</code> or <code>type = "I"</code> must be specified.
type	character or numeric; any one of the (effectively three) choices given. Note that <code>anova.glm</code> has 1 or "I" as its default; and that <code>Anova.glm()</code> in car (that is, the car package) has 2 or "II" as its default (and allows for <code>type = "III"</code>), so one can think of this function as a combination of <code>anova.glm</code> and <code>Anova.glm()</code> in car , but with the default of the latter. See Details below for more information.
test	a character string, (partially) matching one of "LRT" and "none". In the future it is hoped that "Rao" be also supported, to conduct score tests. The first value is the default.
trydev	logical; if TRUE then the deviance is used if possible. Note that only a few VGAM family functions have a deviance that is defined and implemented. Setting it FALSE means the log-likelihood will be used.
silent	logical; if TRUE then any warnings will be suppressed. These may arise by IRLS iterations not converging during the fitting of submodels. Setting it FALSE means that any warnings are given.

Details

`anova.vglm` is intended to be similar to `anova.glm` so specifying a single object and `type = 1` gives a *sequential* analysis of deviance table for that fit. By *analysis of deviance*, it is meant loosely that if the deviance of the model is not defined or implemented, then twice the difference between the log-likelihoods of two nested models remains asymptotically chi-squared distributed with degrees

of freedom equal to the difference in the number of parameters of the two models. Of course, the usual regularity conditions are assumed to hold. For Type I, the analysis of deviance table has the reductions in the residual deviance as each term of the formula is added in turn are given in as the rows of a table, plus the residual deviances themselves. *Type I* or sequential tests (as in [anova.glm](#)). are computationally the easiest of the three methods. For this, the order of the terms is important, and the each term is added sequentially from first to last.

The `Anova()` function in **car** allows for testing *Type II* and *Type III* (SAS jargon) hypothesis tests, although the definitions used are *not* precisely that of SAS. As **car** notes, *Type I* rarely test interesting hypotheses in unbalanced designs. Type III enter each term *last*, keeping all the other terms in the model.

Type II tests, according to SAS, add the term after all other terms have been added to the model except terms that contain the effect being tested; an effect is contained in another effect if it can be derived by deleting variables from the latter effect. Type II tests are currently the default.

As in [anova.glm](#), but not as `Anova.glm()` in **car**, if more than one object is specified, then the table has a row for the residual degrees of freedom and deviance for each model. For all but the first model, the change in degrees of freedom and deviance is also given. (This only makes statistical sense if the models are nested.) It is conventional to list the models from smallest to largest, but this is up to the user. It is necessary to have `type = 1` with more than one objects are specified.

See [anova.glm](#) for more details and warnings. The **VGAM** package now implements full likelihood models only, therefore no dispersion parameters are estimated.

Value

An object of class "anova" inheriting from class "data.frame".

Warning

See [anova.glm](#). Several **VGAM** family functions implement distributions which do not satisfying the usual regularity conditions needed for the LRT to work. No checking or warning is given for these.

As **car** says, be careful of Type III tests because they violate marginality. Type II tests (the default) do not have this problem.

Note

It is possible for this function to `stop` when `type = 2` or `3`, e.g., `anova(vglm(cans ~ myfactor, poissonff, data = boxcar))` where `myfactor` is a factor.

The code was adapted directly from [anova.glm](#) and `Anova.glm()` in **car** by T. W. Yee. Hence the Type II and Type III tests do *not* correspond precisely with the SAS definition.

See Also

[anova.glm](#), [stat.anova](#), `stats:::print.anova`, `Anova.glm()` in **car** if **car** is installed, [vglm](#), [lrtest](#), [add1.vglm](#), [drop1.vglm](#), [lrt.stat.vlm](#), [score.stat.vlm](#), [wald.stat.vlm](#), [backPain2](#), [update](#).

Examples

```
# Example 1: a proportional odds model fitted to pneumo.
set.seed(1)
pneumo <- transform(pneumo, let = log(exposure.time), x3 = runif(8))
fit1 <- vglm(cbind(normal, mild, severe) ~ let, propodds, pneumo)
fit2 <- vglm(cbind(normal, mild, severe) ~ let + x3, propodds, pneumo)
fit3 <- vglm(cbind(normal, mild, severe) ~ let + x3, cumulative, pneumo)
anova(fit1, fit2, fit3, type = 1) # Remember to specify 'type'!!
anova(fit2)
anova(fit2, type = "I")
anova(fit2, type = "III")

# Example 2: a proportional odds model fitted to backPain2.
data("backPain2", package = "VGAM")
summary(backPain2)
fitlogit <- vglm(pain ~ x2 * x3 * x4, propodds, data = backPain2)
coef(fitlogit)
anova(fitlogit)
anova(fitlogit, type = "I")
anova(fitlogit, type = "III")
```

 AR1

Autoregressive Process with Order-1 Family Function

Description

Maximum likelihood estimation of the three-parameter AR-1 model

Usage

```
AR1(ldrift = "identitylink", lsd = "loglink", lvar = "loglink", lrho = "rhitlink",
    idrift = NULL, isd = NULL, ivar = NULL, irho = NULL, imethod = 1,
    ishrinkage = 0.95, type.likelihood = c("exact", "conditional"),
    type.EIM = c("exact", "approximate"), var.arg = FALSE, nodrift = FALSE,
    print.EIM = FALSE, zero = c(if (var.arg) "var" else "sd", "rho"))
```

Arguments

`ldrift`, `lsd`, `lvar`, `lrho`

Link functions applied to the scaled mean, standard deviation or variance, and correlation parameters. The parameter `drift` is known as the *drift*, and it is a scaled mean. See [Links](#) for more choices.

`idrift`, `isd`, `ivar`, `irho`

Optional initial values for the parameters. If failure to converge occurs then try different values and monitor convergence by using `trace = TRUE`. For a S -column response, these arguments can be of length S , and they are recycled by the columns first. A value `NULL` means an initial value for each response is computed internally.

<code>ishrinkage, imethod, zero</code>	See CommonVGAMffArguments for more information. The default for <code>zero</code> assumes there is a drift parameter to be estimated (the default for that argument), so if a drift parameter is suppressed and there are covariates, then <code>zero</code> will need to be assigned the value 1 or 2 or <code>NULL</code> .
<code>var.arg</code>	Same meaning as uninormal .
<code>nodrift</code>	Logical, for determining whether to estimate the drift parameter. The default is to estimate it. If <code>TRUE</code> , the drift parameter is set to 0 and not estimated.
<code>type.EIM</code>	What type of expected information matrix (EIM) is used in Fisher scoring. By default, this family function calls AR1EIM , which recursively computes the exact EIM for the AR process with Gaussian white noise. See Porat and Friedlander (1986) for further details on the exact EIM. If <code>type.EIM = "approximate"</code> then approximate expression for the EIM of Autoregressive processes is used; this approach holds when the number of observations is large enough. Succinct details about the approximate EIM are delineated at Porat and Friedlander (1987).
<code>print.EIM</code>	Logical. If <code>TRUE</code> , then the first few EIMs are printed. Here, the result shown is the sum of each EIM.
<code>type.likelihood</code>	What type of likelihood function is maximized. The first choice (default) is the sum of the marginal likelihood and the conditional likelihood. Choosing the conditional likelihood means that the first observation is effectively ignored (this is handled internally by setting the value of the first prior weight to be some small positive number, e.g., $1.0e-6$). See the note below.

Details

The AR-1 model implemented here has

$$Y_1 \sim N(\mu, \sigma^2 / (1 - \rho^2)),$$

and

$$Y_i = \mu^* + \rho Y_{i-1} + e_i,$$

where the e_i are i.i.d. $\text{Normal}(0, \text{sd} = \sigma)$ random variates.

Here are a few notes: (1). A test for weak stationarity might be to verify whether $1/\rho$ lies outside the unit circle. (2). The mean of all the Y_i is $\mu^*/(1 - \rho)$ and these are returned as the fitted values. (3). The correlation of all the Y_i with Y_{i-1} is ρ . (4). The default link function ensures that $-1 < \rho < 1$.

Value

An object of class `"vglmff"` (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Warning

Monitoring convergence is urged, i.e., set `trace = TRUE`.

Moreover, if the exact EIMs are used, set `print.EIM = TRUE` to compare the computed exact to the approximate EIM.

Under the VGLM/VGAM approach, parameters can be modelled in terms of covariates. Particularly, if the standard deviation of the white noise is modelled in this way, then `type.EIM = "exact"` may certainly lead to unstable results. The reason is that white noise is a stationary process, and consequently, its variance must remain as a constant. Consequently, the use of variates to model this parameter contradicts the assumption of stationary random components to compute the exact EIMs proposed by Porat and Friedlander (1987).

To prevent convergence issues in such cases, this family function internally verifies whether the variance of the white noise remains as a constant at each Fisher scoring iteration. If this assumption is violated and `type.EIM = "exact"` is set, then AR1 automatically shifts to `type.EIM = "approximate"`. Also, a warning is accordingly displayed.

Note

Multiple responses are handled. The mean is returned as the fitted values.

Author(s)

Victor Miranda (exact method) and Thomas W. Yee (approximate method).

References

Porat, B. and Friedlander, B. (1987). The Exact Cramer-Rao Bond for Gaussian Autoregressive Processes. *IEEE Transactions on Aerospace and Electronic Systems*, **AES-23(4)**, 537–542.

See Also

[AR1EIM](#), [vglm.control](#), [dAR1](#), [arma.sim](#).

Examples

```
## Not run:
### Example 1: using arma.sim() to generate a 0-mean stationary time series.
nn <- 500
tsdata <- data.frame(x2 = runif(nn))
ar.coef.1 <- rhobitlink(-1.55, inverse = TRUE) # Approx -0.65
ar.coef.2 <- rhobitlink( 1.0, inverse = TRUE) # Approx  0.50
set.seed(1)
tsdata <- transform(tsdata,
                    index = 1:nn,
                    TS1 = arma.sim(nn, model = list(ar = ar.coef.1),
                                       sd = exp(1.5)),
                    TS2 = arma.sim(nn, model = list(ar = ar.coef.2),
                                       sd = exp(1.0 + 1.5 * x2)))

### An autoregressive intercept--only model. ###
### Using the exact EIM, and "nodrift = TRUE" ###
fit1a <- vglm(TS1 ~ 1, data = tsdata, trace = TRUE,
              AR1(var.arg = FALSE, nodrift = TRUE,
                  type.EIM = "exact",
                  print.EIM = FALSE),
              crit = "coefficients")
```

```

Coef(fit1a)
summary(fit1a)

### Two responses. Here, the white noise standard deviation of TS2   ###
### is modelled in terms of 'x2'. Also, 'type.EIM = exact'.   ###
fit1b <- vglm(cbind(TS1, TS2) ~ x2,
              AR1(zero = NULL, nodrift = TRUE,
                  var.arg = FALSE,
                  type.EIM = "exact"),
              constraints = list("(Intercept)" = diag(4),
                                "x2" = rbind(0, 0, 1, 0)),
              data = tsdata, trace = TRUE, crit = "coefficients")
coef(fit1b, matrix = TRUE)
summary(fit1b)

### Example 2: another stationary time series
nn <- 500
my.rho <- rhobitlink(1.0, inverse = TRUE)
my.mu <- 1.0
my.sd <- exp(1)
tsdata <- data.frame(index = 1:nn, TS3 = runif(nn))

set.seed(2)
for (ii in 2:nn)
  tsdata$TS3[ii] <- my.mu/(1 - my.rho) +
    my.rho * tsdata$TS3[ii-1] + rnorm(1, sd = my.sd)
tsdata <- tsdata[-(1:ceiling(nn/5)), ] # Remove the burn-in data:

### Fitting an AR(1). The exact EIMs are used.
fit2a <- vglm(TS3 ~ 1, AR1(type.likelihood = "exact", # "conditional",
                          type.EIM = "exact"),
              data = tsdata, trace = TRUE, crit = "coefficients")

Coef(fit2a)
summary(fit2a) # SEs are useful to know

Coef(fit2a)["rho"] # Estimate of rho, for intercept-only models
my.rho # The 'truth' (rho)
Coef(fit2a)["drift"] # Estimate of drift, for intercept-only models
my.mu / (1 - my.rho) # The 'truth' (drift)

## End(Not run)

```

Description

Computation of the exact Expected Information Matrix of the Autoregressive process of order-1 (AR(1)) with Gaussian white noise and stationary random components.

Usage

```
AR1EIM(x = NULL, var.arg = NULL, p.drift = NULL,
       WNs = NULL, ARcoeff1 = NULL, eps.porat = 1e-2)
```

Arguments

x	A vector of quantiles. The gaussian time series for which the EIMs are computed. If multiple time series are being analyzed, then x must be a matrix where each column allocates a response. That is, the number of columns (denoted as <i>NOS</i>) must match the number of responses.
var.arg	Logical. Same as with AR1 .
p.drift	A numeric vector with the <i>scaled mean(s)</i> (commonly referred as <i>drift</i>) of the AR process(es) in turn. Its length matches the number of responses.
WNsd, ARcoeff1	Matrices. The standard deviation of the white noise, and the correlation (coefficient) of the AR(1) model, for each observation. That is, the dimension for each matrix is $N \times NOS$, where N is the number of observations and NOS is the number of responses. Else, these arguments are recycled.
eps.porat	A very small positive number to test whether the standar deviation (WNsd) is close enough to its value estimated in this function. See below for further details.

Details

This function implements the algorithm of Porat and Friedlander (1986) to *recursively* compute the exact expected information matrix (EIM) of Gaussian time series with stationary random components.

By default, when the VGLM/VGAM family function [AR1](#) is used to fit an AR(1) model via [vglm](#), Fisher scoring is executed using the **approximate** EIM for the AR process. However, this model can also be fitted using the **exact** EIMs computed by AR1EIM.

Given N consecutive data points, y_0, y_1, \dots, y_{N-1} with probability density $f(\mathbf{y})$, the Porat and Friedlander algorithm calculates the EIMs $[J_{n-1}(\boldsymbol{\theta})]$, for all $1 \leq n \leq N$. This is done based on the Levinson-Durbin algorithm for computing the orthogonal polynomials of a Toeplitz matrix. In particular, for the AR(1) model, the vector of parameters to be estimated under the VGAM/VGLM approach is

$$\boldsymbol{\eta} = (\mu^*, \log(\sigma^2), rhobit(\rho)),$$

where σ^2 is the variance of the white noise and μ^* is the drift parameter (See [AR1](#) for further details on this).

Consequently, for each observation $n = 1, \dots, N$, the EIM, $J_n(\boldsymbol{\theta})$, has dimension 3×3 , where the diagonal elements are:

$$J_{[n,1,1]} = E[-\partial^2 \log f(\mathbf{y}) / \partial(\mu^*)^2],$$

$$J_{[n,2,2]} = E[-\partial^2 \log f(\mathbf{y}) / \partial(\sigma^2)^2],$$

and

$$J_{[n,3,3]} = E[-\partial^2 \log f(\mathbf{y})/\partial(\rho)^2].$$

As for the off-diagonal elements, one has the usual entries, i.e.,

$$J_{[n,1,2]} = J_{[n,2,1]} = E[-\partial^2 \log f(\mathbf{y})/\partial\sigma^2\partial\rho],$$

etc.

If `var.arg = FALSE`, then σ instead of σ^2 is estimated. Therefore, $J_{[n,2,2]}$, $J_{[n,1,2]}$, etc., are correspondingly replaced.

Once these expected values are internally computed, they are returned in an array of dimension $N \times 1 \times 6$, of the form

$$J[, 1,] = [J_{[1,1,1]}, J_{[1,2,2]}, J_{[1,3,3]}, J_{[1,1,2]}, J_{[1,2,3]}, J_{[1,1,3]}].$$

AR1EIM handles multiple time series, say NOS . If this happens, then it accordingly returns an array of dimension $N \times NOS \times 6$. Here, $J[, k,]$, for $k = 1, \dots, NOS$, is a matrix of dimension $N \times 6$, which stores the EIMs for the k^{th} response, as above, i.e.,

$$J[, k,] = [J_{[1,1,1]}, J_{[1,2,2]}, J_{[1,3,3]}, \dots],$$

the *bandwith* form, as per required by AR1.

Value

An array of dimension $N \times NOS \times 6$, as above.

This array stores the EIMs calculated from the joint density as a function of

$$\boldsymbol{\theta} = (\mu^*, \sigma^2, \rho).$$

Nevertheless, note that, under the VGAM/VGLM approach, the EIMs must be correspondingly calculated in terms of the linear predictors, $\boldsymbol{\eta}$.

Asymptotic behaviour of the algorithm

For large enough n , the EIMs, $J_n(\boldsymbol{\theta})$, become approximately linear in n . That is, for some n_0 ,

$$J_n(\boldsymbol{\theta}) \equiv J_{n_0}(\boldsymbol{\theta}) + (n - n_0)\bar{J}(\boldsymbol{\theta}), \quad (**)$$

where $\bar{J}(\boldsymbol{\theta})$ is a constant matrix.

This relationship is internally considered if a proper value of n_0 is determined. Different ways can be adopted to find n_0 . In AR1EIM, this is done by checking the difference between the internally estimated variances and the entered ones at `WNSd`. If this difference is less than `eps.porat` at some iteration, say at iteration n_0 , then AR1EIM takes $\bar{J}(\boldsymbol{\theta})$ as the last computed increment of $J_n(\boldsymbol{\theta})$, and extrapolates $J_k(\boldsymbol{\theta})$, for all $k \geq n_0$ using (*). Else, the algorithm will complete the iterations for $1 \leq n \leq N$.

Finally, note that the rate of convergence reasonably decreases if the asymptotic relationship (*) is used to compute $J_k(\boldsymbol{\theta})$, $k \geq n_0$. Normally, the number of operations involved on this algorithm is proportional to N^2 .

See Porat and Friedlander (1986) for full details on the asymptotic behaviour of the algorithm.

Warning

Arguments `WNsd`, and `ARcoeff1` are matrices of dimension $N \times NOS$. Else, these arguments are accordingly recycled.

Note

For simplicity, one can assume that the time series analyzed has a 0-mean. Consequently, where the family function [AR1](#) calls `AR1EIM` to compute the EIMs, the argument `p.drift` is internally set to zero-vector, whereas `x` is *centered* by subtracting its mean value.

Author(s)

V. Miranda and T. W. Yee.

References

Porat, B. and Friedlander, B. (1986). Computation of the Exact Information Matrix of Gaussian Time Series with Stationary Random Components. *IEEE Transactions on Acoustics, Speech, and Signal Processing*, **54(1)**, 118–130.

See Also

[AR1](#).

Examples

```
set.seed(1)
nn <- 500
ARcoeff1 <- c(0.3, 0.25)      # Will be recycled.
WNsd      <- c(exp(1), exp(1.5)) # Will be recycled.
p.drift   <- c(0, 0)         # Zero-mean gaussian time series.

### Generate two (zero-mean) AR(1) processes ###
ts1 <- p.drift[1]/(1 - ARcoeff1[1]) +
      arima.sim(model = list(ar = ARcoeff1[1]), n = nn,
                sd = WNsd[1])
ts2 <- p.drift[2]/(1 - ARcoeff1[2]) +
      arima.sim(model = list(ar = ARcoeff1[2]), n = nn,
                sd = WNsd[2])

ARdata <- matrix(cbind(ts1, ts2), ncol = 2)

### Compute the exact EIMs: TWO responses. ###
ExactEIM <- AR1EIM(x = ARdata, var.arg = FALSE, p.drift = p.drift,
                  WNsd = WNsd, ARcoeff1 = ARcoeff1)

### For response 1:
head(ExactEIM[, 1, ])      # NOTICE THAT THIS IS A (nn x 6) MATRIX!

### For response 2:
```



```
head(ExactEIM[, 2 ,]) # NOTICE THAT THIS IS A (nn x 6) MATRIX!
```

asinlink	<i>Arcsine Link Function</i>
----------	------------------------------

Description

Computes the arcsine link, including its inverse and the first few derivatives.

Usage

```
asinlink(theta, bvalue = NULL, inverse = FALSE,
         deriv = 0, short = TRUE, tag = FALSE, c10 = c(4, -pi))
```

Arguments

theta	Numeric or character. See below for further details.
bvalue	See Links .
inverse, deriv, short, tag	Details at Links .
c10	Similar to sqrtlink . The default is intended to match lcalogitlink for binomialff at binomial probabilities (theta) equal to 0.5.

Details

Function [alogitlink](#) gives some motivation for this link. However, the problem with this link is that it is bounded by default between $(-\pi, \pi)$ so that it can be unsuitable for regression. This link is a scaled and centred CDF of the arcsine distribution. The centring is chosen so that $\text{asinlink}(0.5)$ is 0, and the scaling is chosen so that $\text{asinlink}(0.5, \text{deriv} = 1)$ and $\text{logitlink}(0.5, \text{deriv} = 1)$ are equal (the value 4 actually), hence this link will operate similar to the [logitlink](#) when close to 0.5.

Value

Similar to [logitlink](#) but using different formulas.

Warning

It is possible that the scaling might change in the future.

Author(s)

Thomas W. Yee

See Also

[logitlink](#), [alogitlink](#), [Links](#), [probitlink](#), [clogloglink](#), [cauchitlink](#), [binomialff](#), [sloglink](#), [hdeff](#).

Examples

```

p <- seq(0.01, 0.99, length= 10)
asinlink(p)
max(abs(asinlink(asinlink(p), inv = TRUE) - p)) # 0?

## Not run:
par(mfrow = c(2, 2), lwd = (mylwd <- 2))
y <- seq(-4, 4, length = 100)
p <- seq(0.01, 0.99, by = 0.01)

for (d in 0:1) {
  matplot(p, cbind(logitlink(p, deriv = d), probitlink(p, deriv = d)),
          type = "n", col = "blue", ylab = "transformation",
          log = ifelse(d == 1, "y", ""),
          las = 1, main = if (d == 0) "Some probability link functions"
          else "First derivative")
  lines(p, logitlink(p, deriv = d), col = "green")
  lines(p, probitlink(p, deriv = d), col = "blue")
  lines(p, clogloglink(p, deriv = d), col = "tan")
  lines(p, asinlink(p, deriv = d), col = "red3")
  if (d == 0) {
    abline(v = 0.5, h = 0, lty = "dashed")
    legend(0, 4.5, c("logitlink", "probitlink", "clogloglink",
                  "asinlink"), lwd = mylwd,
          col = c("green", "blue", "tan", "red3"))
  } else
    abline(v = 0.5, lwd = 0.5, col = "gray")
}

for (d in 0) {
  matplot(y, cbind(logitlink(y, deriv = d, inverse = TRUE),
                  probitlink(y, deriv = d, inverse = TRUE)),
          type = "n", col = "blue", xlab = "transformation", ylab = "p",
          main = if (d == 0) "Some inverse probability link functions"
          else "First derivative", las=1)
  lines(y, logitlink(y, deriv = d, inverse = TRUE), col = "green")
  lines(y, probitlink(y, deriv = d, inverse = TRUE), col = "blue")
  lines(y, clogloglink(y, deriv = d, inverse = TRUE), col = "tan")
  lines(y, asinlink(y, deriv = d, inverse = TRUE), col = "red3")
  if (d == 0) {
    abline(h = 0.5, v = 0, lwd = 0.5, col = "gray")
    legend(-4, 1, c("logitlink", "probitlink", "clogloglink",
                  "asinlink"), lwd = mylwd,
          col = c("green", "blue", "tan", "red3"))
  }
}
par(lwd = 1)

## End(Not run)

```

auuc

Auckland University Undergraduate Counts Data

Description

Undergraduate student enrolments at the University of Auckland in 1990.

Usage

```
data(auuc)
```

Format

A data frame with 4 observations on the following 5 variables.

Commerce a numeric vector of counts.

Arts a numeric vector of counts.

SciEng a numeric vector of counts.

Law a numeric vector of counts.

Medicine a numeric vector of counts.

Details

Each student is cross-classified by their colleges (Science and Engineering have been combined) and the socio-economic status (SES) of their fathers (1 = highest, down to 4 = lowest).

Source

Dr Tony Morrison.

References

Wild, C. J. and Seber, G. A. F. (2000). *Chance Encounters: A First Course in Data Analysis and Inference*, New York: Wiley.

Examples

```
auuc
## Not run:
round(fitted(grc(auuc)))
round(fitted(grc(auuc, Rank = 2)))

## End(Not run)
```

aux.posbernoulli.t *Auxiliary Function for the Positive Bernoulli Family Function with Time Effects*

Description

Returns behavioural effects indicator variables from a capture history matrix.

Usage

```
aux.posbernoulli.t(y, check.y = FALSE, rename = TRUE, name = "bei")
```

Arguments

y	Capture history matrix. Rows are animals, columns are sampling occasions, and values should be 0s and 1s only.
check.y	Logical, if TRUE then some basic checking is performed.
rename, name	If rename = TRUE then the behavioural effects indicator are named using the value of name as the prefix. If FALSE then use the same column names as y.

Details

This function can help fit certain capture–recapture models (commonly known as M_{tb} or M_{tbb} (no prefix h means it is an intercept-only model) in the literature). See [posbernoulli.t](#) for details.

Value

A list with the following components.

cap.hist1 A matrix the same dimension as y. In any particular row there are 0s up to the first capture. Then there are 1s thereafter.

cap1 A vector specifying which time occasion the animal was first captured.

y0i Number of noncaptures before the first capture.

yr0i Number of noncaptures after the first capture.

yr1i Number of recaptures after the first capture.

See Also

[posbernoulli.t](#), [deermice](#).

Examples

```
# Fit a M_tbh model to the deermice data:
(pdata <- aux.posbernoulli.t(with(deermice,
                                cbind(y1, y2, y3, y4, y5, y6))))

deermice <- data.frame(deermice,
                      bei = 0, # Add this
                      pdata$cap.hist1) # Incorporate these
head(deermice) # Augmented with behavioural effect indicator variables
tail(deermice)
```

backPain

*Data on Back Pain Prognosis, from Anderson (1984)***Description**

Data from a study of patients suffering from back pain. Prognostic variables were recorded at presentation and progress was categorised three weeks after treatment.

Usage

```
data(backPain)
```

Format

A data frame with 101 observations on the following 4 variables.

x2 length of previous attack.

x3 pain change.

x4 lordosis.

pain an ordered factor describing the progress of each patient with levels worse < same < slight.improvement < moderate.improvement < marked.improvement < complete.relief.

Source

<http://ideas.repec.org/c/boc/bocode/s419001.html>

The data set and this help file was copied from **gnm** so that a vignette in **VGAM** could be run; the analysis is described in Yee (2010).

The data frame backPain2 is a modification of backPain where the variables have been renamed (x1 becomes x2, x2 becomes x3, x3 becomes x4) and converted into factors.

References

Anderson, J. A. (1984). Regression and Ordered Categorical Variables. *J. R. Statist. Soc. B*, **46(1)**, 1-30.

Yee, T. W. (2010). The **VGAM** package for categorical data analysis. *Journal of Statistical Software*, **32**, 1–34. doi:10.18637/jss.v032.i10.

Examples

```
summary(backPain)
summary(backPain2)
```

beggs

Bacon and Eggs Data

Description

Purchasing of bacon and eggs.

Usage

```
data(beggs)
```

Format

Data frame of a two way table.

b0, b1, b2, b3, b4 The b refers to bacon. The number of times bacon was purchased was 0, 1, 2, 3, or 4.

e0, e1, e2, e3, e4 The e refers to eggs. The number of times eggs was purchased was 0, 1, 2, 3, or 4.

Details

The data is from Information Resources, Inc., a consumer panel based in a large US city [see Bell and Lattin (1998) for further details]. Starting in June 1991, the purchases in the bacon and fresh eggs product categories for a sample of 548 households over four consecutive store trips was tracked. Only those grocery shopping trips with a total basket value of at least five dollars was considered. For each household, the total number of bacon purchases in their four eligible shopping trips and the total number of egg purchases (usually a package of eggs) for the same trips, were counted.

Source

Bell, D. R. and Lattin, J. M. (1998) Shopping Behavior and Consumer Preference for Store Price Format: Why 'Large Basket' Shoppers Prefer EDLP. *Marketing Science*, **17**, 66–88.

References

Danaher, P. J. and Hardie, B. G. S. (2005). Bacon with Your Eggs? Applications of a New Bivariate Beta-Binomial Distribution. *American Statistician*, **59**(4), 282–286.

See Also

[rrvglm](#), [rcim](#), [grc](#).

Examples

```
beggs
colSums(beggs)
rowSums(beggs)
```

bell

The Bell Series of Integers

Description

Returns the values of the Bell series.

Usage

```
bell(n)
```

Arguments

n Vector of non-negative integers. Values greater than 218 return an Inf. Non-integers or negative values return a NaN.

Details

The Bell numbers emerge from a series expansion of $\exp(e^x - 1)$ for real x . The first few values are $B_0 = 1$, $B_1 = 1$, $B_2 = 2$, $B_3 = 5$, $B_4 = 15$. The series increases quickly so that overflow occurs when its argument is more than 218.

Value

This function returns B_n .

Author(s)

T. W. Yee

References

Bell, E. T. (1934). Exponential polynomials. *Ann. Math.*, **35**, 258–277.

Bell, E. T. (1934). Exponential numbers. *Amer. Math. Monthly*, **41**, 411–419.

See Also

[bellff](#), [rbell](#).

Examples

```
## Not run:
plot(0:10, bell(0:10), log = "y", type = "h", col = "blue")

## End(Not run)
```

Benford

*Benford's Distribution***Description**

Density, distribution function, quantile function, and random generation for Benford's distribution.

Usage

```
dbenf(x, ndigits = 1, log = FALSE)
pbenf(q, ndigits = 1, lower.tail = TRUE, log.p = FALSE)
qbenf(p, ndigits = 1, lower.tail = TRUE, log.p = FALSE)
rbenf(n, ndigits = 1)
```

Arguments

<code>x, q</code>	Vector of quantiles. See <code>ndigits</code> .
<code>p</code>	vector of probabilities.
<code>n</code>	number of observations. A single positive integer. Else if <code>length(n) > 1</code> then the length is taken to be the number required.
<code>ndigits</code>	Number of leading digits, either 1 or 2. If 1 then the support of the distribution is $\{1, \dots, 9\}$, else $\{10, \dots, 99\}$.
<code>log, log.p</code>	Logical. If <code>log.p = TRUE</code> then all probabilities <code>p</code> are given as $\log(p)$.
<code>lower.tail</code>	Same meaning as in pnorm or qnorm .

Details

Benford's Law (aka *the significant-digit law*) is the empirical observation that in many naturally occurring tables of numerical data, the leading significant (nonzero) digit is not uniformly distributed in $\{1, 2, \dots, 9\}$. Instead, the leading significant digit ($= D$, say) obeys the law

$$P(D = d) = \log_{10} \left(1 + \frac{1}{d} \right)$$

for $d = 1, \dots, 9$. This means the probability the first significant digit is 1 is approximately 0.301, etc.

Benford's Law was apparently first discovered in 1881 by astronomer/mathematician S. Newcombe. It started by the observation that the pages of a book of logarithms were dirtiest at the beginning and progressively cleaner throughout. In 1938, a General Electric physicist called F. Benford re-discovered the law on this same observation. Over several years he collected data from different

sources as different as atomic weights, baseball statistics, numerical data from *Reader's Digest*, and drainage areas of rivers.

Applications of Benford's Law has been as diverse as to the area of fraud detection in accounting and the design computers.

Benford's distribution has been called "a" logarithmic distribution; see [logff](#).

Value

`dbenf` gives the density, `pbenf` gives the distribution function, and `qbenf` gives the quantile function, and `rbenf` generates random deviates.

Author(s)

T. W. Yee and Kai Huang

References

Benford, F. (1938). The Law of Anomalous Numbers. *Proceedings of the American Philosophical Society*, **78**, 551–572.

Newcomb, S. (1881). Note on the Frequency of Use of the Different Digits in Natural Numbers. *American Journal of Mathematics*, **4**, 39–40.

Examples

```
dbenf(x <- c(0:10, NA, NaN, -Inf, Inf))
pbenf(x)

## Not run:
xx <- 1:9
barplot(dbenf(xx), col = "lightblue", xlab = "Leading digit",
        ylab = "Probability", names.arg = as.character(xx),
        main = "Benford's distribution", las = 1)

hist(rbenf(1000), border = "blue", prob = TRUE,
     main = "1000 random variates from Benford's distribution",
     xlab = "Leading digit", sub="Red is the true probability",
     breaks = 0:9 + 0.5, ylim = c(0, 0.35), xlim = c(0, 10.0))
lines(xx, dbenf(xx), col = "red", type = "h")
points(xx, dbenf(xx), col = "red")

## End(Not run)
```

Description

Density, distribution function, quantile function and random generation for the Benini distribution with parameter shape.

Usage

```

dbenini(x, y0, shape, log = FALSE)
pbenini(q, y0, shape, lower.tail = TRUE, log.p = FALSE)
qbenini(p, y0, shape, lower.tail = TRUE, log.p = FALSE)
rbenini(n, y0, shape)

```

Arguments

x, q	vector of quantiles.
p	vector of probabilities.
n	number of observations. Same as runif .
y0	the scale parameter y_0 .
shape	the positive shape parameter b .
log	Logical. If log = TRUE then the logarithm of the density is returned.
lower.tail, log.p	Same meaning as in pnorm or qnorm .

Details

See [benini1](#), the **VGAM** family function for estimating the parameter s by maximum likelihood estimation, for the formula of the probability density function and other details.

Value

dbenini gives the density, pbenini gives the distribution function, qbenini gives the quantile function, and rbenini generates random deviates.

Author(s)

T. W. Yee and Kai Huang

References

Kleiber, C. and Kotz, S. (2003). *Statistical Size Distributions in Economics and Actuarial Sciences*, Hoboken, NJ, USA: Wiley-Interscience.

See Also

[benini1](#).

Examples

```

## Not run:
y0 <- 1; shape <- exp(1)
xx <- seq(0.0, 4, len = 101)
plot(xx, dbenini(xx, y0 = y0, shape = shape), col = "blue",
      main = "Blue is density, orange is the CDF", type = "l",
      sub = "Purple lines are the 10,20,...,90 percentiles",

```

```

      ylim = 0:1, las = 1, ylab = "", xlab = "x")
abline(h = 0, col = "blue", lty = 2)
lines(xx, pbenini(xx, y0 = y0, shape = shape), col = "orange")
probs <- seq(0.1, 0.9, by = 0.1)
Q <- qbenini(probs, y0 = y0, shape = shape)
lines(Q, dbenini(Q, y0 = y0, shape = shape),
      col = "purple", lty = 3, type = "h")
pbenini(Q, y0 = y0, shape = shape) - probs # Should be all zero

## End(Not run)

```

benini1

*Benini Distribution Family Function***Description**

Estimating the 1-parameter Benini distribution by maximum likelihood estimation.

Usage

```

benini1(y0 = stop("argument 'y0' must be specified"),
        lshape = "loglink", ishape = NULL, imethod = 1,
        zero = NULL, parallel = FALSE,
        type.fitted = c("percentiles", "Qlink"),
        percentiles = 50)

```

Arguments

<code>y0</code>	Positive scale parameter.
<code>lshape</code>	Parameter link function and extra argument of the parameter b , which is the shape parameter. See Links for more choices. A log link is the default because b is positive.
<code>ishape</code>	Optional initial value for the shape parameter. The default is to compute the value internally.
<code>imethod, zero, parallel</code>	Details at CommonVGAMffArguments .
<code>type.fitted, percentiles</code>	See CommonVGAMffArguments for information. Using "Qlink" is for quantile-links in VGAMextra .

Details

The Benini distribution has a probability density function that can be written

$$f(y) = 2s \exp(-s[(\log(y/y_0))^2]) \log(y/y_0)/y$$

for $0 < y_0 < y$, and shape $s > 0$. The cumulative distribution function for Y is

$$F(y) = 1 - \exp(-s[(\log(y/y_0))^2]).$$

Here, Newton-Raphson and Fisher scoring coincide. The median of Y is now returned as the fitted values, by default. This **VGAM** family function can handle a multiple responses, which is inputted as a matrix.

On fitting, the extra slot has a component called y_0 which contains the value of the y_0 argument.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Note

Yet to do: the 2-parameter Benini distribution estimates another shape parameter a too. Hence, the code may change in the future.

Author(s)

T. W. Yee

References

Kleiber, C. and Kotz, S. (2003). *Statistical Size Distributions in Economics and Actuarial Sciences*, Hoboken, NJ, USA: Wiley-Interscience.

See Also

[Benini](#).

Examples

```
y0 <- 1; nn <- 3000
bdata <- data.frame(y = rbenini(nn, y0 = y0, shape = exp(2)))
fit <- vglm(y ~ 1, benini1(y0 = y0), data = bdata, trace = TRUE)
coef(fit, matrix = TRUE)
Coef(fit)
fit@extra$y0
c(head(fitted(fit), 1), with(bdata, median(y))) # Should be equal
```

Description

Density, distribution function, and random generation for the beta-binomial distribution and the inflated beta-binomial distribution.

Usage

```

dbetabinom(x, size, prob, rho = 0, log = FALSE)
pbetabinom(q, size, prob, rho = 0, log.p = FALSE)
rbetabinom(n, size, prob, rho = 0)
dbetabinom.ab(x, size, shape1, shape2, log = FALSE,
              Inf.shape = exp(20), limit.prob = 0.5)
pbetabinom.ab(q, size, shape1, shape2, limit.prob = 0.5,
              log.p = FALSE)
rbetabinom.ab(n, size, shape1, shape2, limit.prob = 0.5,
              .dontuse.prob = NULL)
dzoibetabinom(x, size, prob, rho = 0, pstr0 = 0, pstrsize = 0,
              log = FALSE)
pzoibetabinom(q, size, prob, rho, pstr0 = 0, pstrsize = 0,
              lower.tail = TRUE, log.p = FALSE)
rzoibetabinom(n, size, prob, rho = 0, pstr0 = 0, pstrsize = 0)
dzoibetabinom.ab(x, size, shape1, shape2, pstr0 = 0, pstrsize = 0,
                 log = FALSE)
pzoibetabinom.ab(q, size, shape1, shape2, pstr0 = 0, pstrsize = 0,
                 lower.tail = TRUE, log.p = FALSE)
rzoibetabinom.ab(n, size, shape1, shape2, pstr0 = 0, pstrsize = 0)

```

Arguments

x, q	vector of quantiles.
size	number of trials.
n	number of observations. Same as <code>runif</code> .
prob	the probability of success μ . Must be in the unit closed interval $[0, 1]$.
rho	the correlation parameter ρ , which should be in the interval $[0, 1)$. The default value of 0 corresponds to the usual binomial distribution with probability <code>prob</code> . Setting <code>rho = 1</code> would set both shape parameters equal to 0, and the ratio $0/0$, which is actually NaN, is interpreted by <code>Beta</code> as 0.5. See the warning below.
shape1, shape2	the two (positive) shape parameters of the standard beta distribution. They are called a and b in <code>beta</code> respectively. Note that <code>shape1 = prob*(1-rho)/rho</code> and <code>shape2 = (1-prob)*(1-rho)/rho</code> is an important relationship between the parameters, so that the shape parameters are infinite by default because <code>rho = 0</code> ; hence <code>limit.prob = prob</code> is used to obtain the behaviour of the usual binomial distribution.
log, log.p, lower.tail	Same meaning as <code>runif</code> .
Inf.shape	Numeric. A large value such that, if <code>shape1</code> or <code>shape2</code> exceeds this, then special measures are taken, e.g., calling <code>dbinom</code> . Also, if <code>shape1</code> or <code>shape2</code> is less than its reciprocal, then special measures are also taken. This feature/approximation is needed to avoid numerical problem with catastrophic cancellation of multiple <code>lbeta</code> calls.
limit.prob	Numerical vector; recycled if necessary. If either shape parameters are Inf then the binomial limit is taken, with <code>shape1 / (shape1 + shape2)</code> as the probability

of success. In the case where both are `Inf` this probability will be a `NaN = Inf/Inf`, however, the value `limit.prob` is used instead. Hence the default for `dbetabinom.ab()` is to assume that both shape parameters are equal as the limit is taken (indeed, `Beta` uses 0.5). Note that for `[dpr]betabinom()`, because `rho = 0` by default, then `limit.prob = prob` so that the beta-binomial distribution behaves like the ordinary binomial distribution with respect to arguments `size` and `prob`.

<code>.dontuse.prob</code>	An argument that should be ignored and <i>not</i> used.
<code>pstr0</code>	Probability of a structural zero (i.e., ignoring the beta-binomial distribution). The default value of <code>pstr0</code> corresponds to the response having a beta-binomial distribution inflated only at <code>size</code> .
<code>pstrsize</code>	Probability of a structural maximum value <code>size</code> . The default value of <code>pstrsize</code> corresponds to the response having a beta-binomial distribution inflated only at 0.

Details

The beta-binomial distribution is a binomial distribution whose probability of success is not a constant but it is generated from a beta distribution with parameters `shape1` and `shape2`. Note that the mean of this beta distribution is $\mu = \text{shape1}/(\text{shape1}+\text{shape2})$, which therefore is the mean or the probability of success.

See `betabinomial` and `betabinomialff`, the **VGAM** family functions for estimating the parameters, for the formula of the probability density function and other details.

For the inflated beta-binomial distribution, the probability mass function is

$$P(Y = y) = (1 - pstr0 - pstrsize) \times BB(y) + pstr0 \times I[y = 0] + pstrsize \times I[y = size]$$

where $BB(y)$ is the probability mass function of the beta-binomial distribution with the same shape parameters (`pbetabinom.ab`), `pstr0` is the inflated probability at 0 and `pstrsize` is the inflated probability at 1. The default values of `pstr0` and `pstrsize` mean that these functions behave like the ordinary `Betabinom` when only the essential arguments are inputted.

Value

`dbetabinom` and `dbetabinom.ab` give the density, `pbetabinom` and `pbetabinom.ab` give the distribution function, and `rbetabinom` and `rbetabinom.ab` generate random deviates.

`dzoibetabinom` and `dzoibetabinom.ab` give the inflated density, `pzoibetabinom` and `pzoibetabinom.ab` give the inflated distribution function, and `rzoibetabinom` and `rzoibetabinom.ab` generate random inflated deviates.

Warning

Setting `rho = 1` is not recommended, however the code may be modified in the future to handle this special case.

Note

pzoibetabinom, pzoibetabinom.ab, pbetabinom and pbetabinom.ab can be particularly slow. The functions here ending in .ab are called from those functions which don't. The simple transformations $\mu = \alpha/(\alpha + \beta)$ and $\rho = 1/(1 + \alpha + \beta)$ are used, where α and β are the two shape parameters.

Author(s)

T. W. Yee and Xiangjie Xue

See Also

[Extbetabinom](#), [betabinomial](#), [betabinomialfff](#), [Zoabeta](#), [Beta](#).

Examples

```
set.seed(1); rbetabinom(10, 100, prob = 0.5)
set.seed(1); rbinom(10, 100, prob = 0.5) # The same as rho = 0

## Not run: N <- 9; xx <- 0:N; s1 <- 2; s2 <- 3
dy <- dbetabinom.ab(xx, size = N, shape1 = s1, shape2 = s2)
barplot(rbind(dy, dbinom(xx, size = N, prob = s1 / (s1+s2))),
        beside = TRUE, col = c("blue", "green"), las = 1,
        main = paste("Beta-binomial (size=", N, ", shape1=", s1,
                    ", shape2=", s2, ") (blue) vs\n",
                    " Binomial(size=", N, ", prob=", s1/(s1+s2), ") (green)",
                    sep = ""),
        names.arg = as.character(xx), cex.main = 0.8)
sum(dy * xx) # Check expected values are equal
sum(dbinom(xx, size = N, prob = s1 / (s1+s2)) * xx)
# Should be all 0:
cumsum(dy) - pbetabinom.ab(xx, N, shape1 = s1, shape2 = s2)

y <- rbetabinom.ab(n = 1e4, size = N, shape1 = s1, shape2 = s2)
ty <- table(y)
barplot(rbind(dy, ty / sum(ty)),
        beside = TRUE, col = c("blue", "orange"), las = 1,
        main = paste("Beta-binomial (size=", N, ", shape1=", s1,
                    ", shape2=", s2, ") (blue) vs\n",
                    " Random generated beta-binomial(size=", N, ", prob=",
                    s1/(s1+s2), ") (orange)", sep = ""), cex.main = 0.8,
        names.arg = as.character(xx))

N <- 1e5; size <- 20; pstr0 <- 0.2; pstrsize <- 0.2
kk <- rzoibetabinom.ab(N, size, s1, s2, pstr0, pstrsize)
hist(kk, probability = TRUE, border = "blue", ylim = c(0, 0.25),
     main = "Blue/green = inflated; orange = ordinary beta-binomial",
     breaks = -0.5 : (size + 0.5))
sum(kk == 0) / N # Proportion of 0
sum(kk == size) / N # Proportion of size
lines(0 : size,
     dbetabinom.ab(0 : size, size, s1, s2), col = "orange")
```

```
lines(0 : size, col = "green", type = "b",
      dzoibetabinom.ab(0 : size, size, s1, s2, pstr0, pstrsize))

## End(Not run)
```

betabinomial

*Beta-binomial Distribution Family Function***Description**

Fits a beta-binomial distribution by maximum likelihood estimation. The two parameters here are the mean and correlation coefficient.

Usage

```
betabinomial(lmu = "logitlink", lrho = "logitlink",
             irho = NULL, imethod = 1,
             ishrinkage = 0.95, nsimEIM = NULL, zero = "rho")
```

Arguments

lmu, lrho	Link functions applied to the two parameters. See Links for more choices. The defaults ensure the parameters remain in $(0, 1)$, however, see the warning below. For lrho, log1plink (with an offset $\log(\text{size} - 1)$ for η_2) and cloglink may be very good choices.
irho	Optional initial value for the correlation parameter. If given, it must be in $(0, 1)$, and is recycled to the necessary length. Assign this argument a value if a convergence failure occurs. Having <code>irho = NULL</code> means an initial value is obtained internally, though this can give unsatisfactory results.
imethod	An integer with value 1 or 2 or \dots , which specifies the initialization method for μ . If failure to converge occurs try the another value and/or else specify a value for <code>irho</code> .
zero	Specifyies which linear/additive predictor is to be modelled as an intercept only. If assigned, the single value can be either 1 or 2. The default is to have a single correlation parameter. To model both parameters as functions of the covariates assign <code>zero = NULL</code> . See CommonVGAMffArguments for more information.
ishrinkage, nsimEIM	See CommonVGAMffArguments for more information. The argument <code>ishrinkage</code> is used only if <code>imethod = 2</code> . Using the argument <code>nsimEIM</code> may offer large advantages for large values of N and/or large data sets.

Details

There are several parameterizations of the beta-binomial distribution. This family function directly models the mean and correlation parameter, i.e., the probability of success. The model can be written $T|P = p \sim \text{Binomial}(N, p)$ where P has a beta distribution with shape parameters α and

β . Here, N is the number of trials (e.g., litter size), $T = NY$ is the number of successes, and p is the probability of a success (e.g., a malformation). That is, Y is the *proportion* of successes. Like [binomialff](#), the fitted values are the estimated probability of success (i.e., $E[Y]$ and not $E[T]$) and the prior weights N are attached separately on the object in a slot.

The probability function is

$$P(T = t) = \binom{N}{t} \frac{Be(\alpha + t, \beta + N - t)}{Be(\alpha, \beta)}$$

where $t = 0, 1, \dots, N$, and Be is the [beta](#) function with shape parameters α and β . Recall $Y = T/N$ is the real response being modelled.

The default model is $\eta_1 = \text{logit}(\mu)$ and $\eta_2 = \text{logit}(\rho)$ because both parameters lie between 0 and 1. The mean (of Y) is $p = \mu = \alpha/(\alpha + \beta)$ and the variance (of Y) is $\mu(1 - \mu)(1 + (N - 1)\rho)/N$. Here, the correlation ρ is given by $1/(1 + \alpha + \beta)$ and is the correlation between the N individuals within a litter. A *litter effect* is typically reflected by a positive value of ρ . It is known as the *over-dispersion parameter*.

This family function uses Fisher scoring. Elements of the second-order expected derivatives with respect to α and β are computed numerically, which may fail for large α , β , N or else take a long time.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#).

Suppose `fit` is a fitted beta-binomial model. Then `depvar(fit)` are the sample proportions y , `fitted(fit)` returns estimates of $E(Y)$, and `weights(fit, type = "prior")` returns the number of trials N .

Warning

If the estimated rho parameter is close to 0 then a good solution is to use [extbetabinomial](#). Or you could try `lrho = "rhobitlink"`.

This family function is prone to numerical difficulties due to the expected information matrices not being positive-definite or ill-conditioned over some regions of the parameter space. If problems occur try setting `irho` to some numerical value, `nsimEIM = 100`, say, or else use `etastart` argument of [vglm](#), etc.

Note

This function processes the input in the same way as [binomialff](#). But it does not handle the case $N = 1$ very well because there are two parameters to estimate, not one, for each row of the input. Cases where $N = 1$ can be omitted via the `subset` argument of [vglm](#).

The *extended* beta-binomial distribution of Prentice (1986) implemented by [extbetabinomial](#) is the preferred **VGAM** family function for BBD regression.

Author(s)

T. W. Yee

References

Moore, D. F. and Tsiatis, A. (1991). Robust estimation of the variance in moment methods for extra-binomial and extra-Poisson variation. *Biometrics*, **47**, 383–401.

See Also

[extbetabinomial](#), [betabinomialfff](#), [Betabinom](#), [binomialfff](#), [betafff](#), [dirmultinomial](#), [log1plink](#), [cloglink](#), [lirat](#), [simulate.vlm](#).

Examples

```
# Example 1
bdata <- data.frame(N = 10, mu = 0.5, rho = 0.8)
bdata <- transform(bdata,
  y = rbetabinom(100, size = N, prob = mu, rho = rho))
fit <- vglm(cbind(y, N-y) ~ 1, betabinomial, bdata, trace = TRUE)
coef(fit, matrix = TRUE)
Coef(fit)
head(cbind(depvar(fit), weights(fit, type = "prior")))
```

```
# Example 2
fit <- vglm(cbind(R, N-R) ~ 1, betabinomial, lirat,
  trace = TRUE, subset = N > 1)
coef(fit, matrix = TRUE)
Coef(fit)
t(fitted(fit))
t(depvar(fit))
t(weights(fit, type = "prior"))
```

```
# Example 3, which is more complicated
lirat <- transform(lirat, fgrp = factor(grp))
summary(lirat) # Only 5 litters in group 3
fit2 <- vglm(cbind(R, N-R) ~ fgrp + hb, betabinomial(zero = 2),
  data = lirat, trace = TRUE, subset = N > 1)
coef(fit2, matrix = TRUE)
## Not run: with(lirat, plot(hb[N > 1], fit2@misc$rho,
  xlab = "Hemoglobin", ylab = "Estimated rho",
  pch = as.character(grp[N > 1]), col = grp[N > 1]))
## End(Not run)
## Not run: # cf. Figure 3 of Moore and Tsiatis (1991)
with(lirat, plot(hb, R / N, pch = as.character(grp), col = grp,
  xlab = "Hemoglobin level", ylab = "Proportion Dead",
  main = "Fitted values (lines)", las = 1))
smalldf <- with(lirat, lirat[N > 1, ])
for (gp in 1:4) {
  xx <- with(smalldf, hb[grp == gp])
  yy <- with(smalldf, fitted(fit2)[grp == gp])
  ooo <- order(xx)
  lines(xx[ooo], yy[ooo], col = gp, lwd = 2)
}
```

```
## End(Not run)
```

betabinomialff	<i>Beta-binomial Distribution Family Function</i>
----------------	---

Description

Fits a beta-binomial distribution by maximum likelihood estimation. The two parameters here are the shape parameters of the underlying beta distribution.

Usage

```
betabinomialff(lshape1 = "loglink", lshape2 = "loglink",
  ishape1 = 1, ishape2 = NULL, imethod = 1, ishrinkage = 0.95,
  nsimEIM = NULL, zero = NULL)
```

Arguments

lshape1, lshape2	Link functions for the two (positive) shape parameters of the beta distribution. See Links for more choices.
ishape1, ishape2	Initial value for the shape parameters. The first must be positive, and is recycled to the necessary length. The second is optional. If a failure to converge occurs, try assigning a different value to ishape1 and/or using ishape2.
zero	Can be an integer specifying which linear/additive predictor is to be modelled as an intercept only. If assigned, the single value should be either 1 or 2. The default is to model both shape parameters as functions of the covariates. If a failure to converge occurs, try zero = 2. See CommonVGAMffArguments for more information.
ishrinkage, nsimEIM, imethod	See CommonVGAMffArguments for more information. The argument ishrinkage is used only if imethod = 2. Using the argument nsimEIM may offer large advantages for large values of N and/or large data sets.

Details

There are several parameterizations of the beta-binomial distribution. This family function directly models the two shape parameters of the associated beta distribution rather than the probability of success (however, see **Note** below). The model can be written $T|P = p \sim \text{Binomial}(N, p)$ where P has a beta distribution with shape parameters α and β . Here, N is the number of trials (e.g., litter size), $T = NY$ is the number of successes, and p is the probability of a success (e.g., a malformation). That is, Y is the *proportion* of successes. Like [binomialff](#), the fitted values are the estimated probability of success (i.e., $E[Y]$ and not $E[T]$) and the prior weights N are attached separately on the object in a slot.

The probability function is

$$P(T = t) = \binom{N}{t} \frac{B(\alpha + t, \beta + N - t)}{B(\alpha, \beta)}$$

where $t = 0, 1, \dots, N$, and B is the beta function with shape parameters α and β . Recall $Y = T/N$ is the real response being modelled.

The default model is $\eta_1 = \log(\alpha)$ and $\eta_2 = \log(\beta)$ because both parameters are positive. The mean (of Y) is $p = \mu = \alpha/(\alpha + \beta)$ and the variance (of Y) is $\mu(1 - \mu)(1 + (N - 1)\rho)/N$. Here, the correlation ρ is given by $1/(1 + \alpha + \beta)$ and is the correlation between the N individuals within a litter. A *litter effect* is typically reflected by a positive value of ρ . It is known as the *over-dispersion parameter*.

This family function uses Fisher scoring. The two diagonal elements of the second-order expected derivatives with respect to α and β are computed numerically, which may fail for large α , β , N or else take a long time.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#).

Suppose `fit` is a fitted beta-binomial model. Then `fit@y` (better: `depvar(fit)`) contains the sample proportions y , `fitted(fit)` returns estimates of $E(Y)$, and `weights(fit, type = "prior")` returns the number of trials N .

Warning

This family function is prone to numerical difficulties due to the expected information matrices not being positive-definite or ill-conditioned over some regions of the parameter space. If problems occur try setting `ishape1` to be some other positive value, using `ishape2` and/or setting `zero = 2`.

This family function may be renamed in the future. See the warnings in [betabinomial](#).

Note

This function processes the input in the same way as [binomialff](#). But it does not handle the case $N = 1$ very well because there are two parameters to estimate, not one, for each row of the input. Cases where $N = 1$ can be omitted via the `subset` argument of [vglm](#).

Although the two linear/additive predictors given above are in terms of α and β , basic algebra shows that the default amounts to fitting a logit link to the probability of success; subtracting the second linear/additive predictor from the first gives that logistic regression linear/additive predictor. That is, $\text{logit}(p) = \eta_1 - \eta_2$. This is illustrated in one of the examples below.

The *extended* beta-binomial distribution of Prentice (1986) implemented by [extbetabinomial](#) is the preferred **VGAM** family function for BBD regression.

Author(s)

T. W. Yee

References

Moore, D. F. and Tsiatis, A. (1991). Robust estimation of the variance in moment methods for extra-binomial and extra-Poisson variation. *Biometrics*, **47**, 383–401.

Prentice, R. L. (1986). Binary regression using an extended beta-binomial distribution, with discussion of correlation induced by covariate measurement errors. *Journal of the American Statistical Association*, **81**, 321–327.

See Also

[extbetabinomial](#), [betabinomial](#), [Betabinom](#), [binomialff](#), [betaff](#), [dirmultinomial](#), [lirat](#), [simulate.vlm](#).

Examples

```
# Example 1
N <- 10; s1 <- exp(1); s2 <- exp(2)
y <- rbetabinom.ab(n = 100, size = N, shape1 = s1, shape2 = s2)
fit <- vglm(cbind(y, N-y) ~ 1, betabinomialff, trace = TRUE)
coef(fit, matrix = TRUE)
Coef(fit)
head(fit@misc$rho) # The correlation parameter
head(cbind(depvar(fit), weights(fit, type = "prior"))

# Example 2
fit <- vglm(cbind(R, N-R) ~ 1, betabinomialff, data = lirat,
            trace = TRUE, subset = N > 1)
coef(fit, matrix = TRUE)
Coef(fit)
fit@misc$rho # The correlation parameter
t(fitted(fit))
t(depvar(fit))
t(weights(fit, type = "prior"))
# A "loglink" link for the 2 shape params is a logistic regression:
all.equal(c(fitted(fit)),
          as.vector(logitlink(predict(fit)[, 1] -
                               predict(fit)[, 2], inverse = TRUE)))

# Example 3, which is more complicated
lirat <- transform(lirat, fgrp = factor(grp))
summary(lirat) # Only 5 litters in group 3
fit2 <- vglm(cbind(R, N-R) ~ fgrp + hb, betabinomialff(zero = 2),
            data = lirat, trace = TRUE, subset = N > 1)
coef(fit2, matrix = TRUE)
coef(fit2, matrix = TRUE)[, 1] -
coef(fit2, matrix = TRUE)[, 2] # logitlink(p)
## Not run: with(lirat, plot(hb[N > 1], fit2@misc$rho,
  xlab = "Hemoglobin", ylab = "Estimated rho",
  pch = as.character(grp[N > 1]), col = grp[N > 1]))
## End(Not run)
```

```
## Not run: # cf. Figure 3 of Moore and Tsiatis (1991)
with(lirat, plot(hb, R / N, pch = as.character(grp), col = grp,
  xlab = "Hemoglobin level", ylab = "Proportion Dead", las = 1,
  main = "Fitted values (lines)"))

smalldf <- with(lirat, lirata[N > 1, ])
for (gp in 1:4) {
  xx <- with(smalldf, hb[grp == gp])
  yy <- with(smalldf, fitted(fit2)[grp == gp])
  ooo <- order(xx)
  lines(xx[ooo], yy[ooo], col = gp, lwd = 2)
}
## End(Not run)
```

betaff

*The Two-parameter Beta Distribution Family Function***Description**

Estimation of the mean and precision parameters of the beta distribution.

Usage

```
betaff(A = 0, B = 1, lmu = "logitlink", lphi = "loglink",
  imu = NULL, iphi = NULL,
  gprobs.y = ppoints(8), gphi = exp(-3:5)/4, zero = NULL)
```

Arguments

A, B Lower and upper limits of the distribution. The defaults correspond to the *standard beta distribution* where the response lies between 0 and 1.

lmu, lphi Link function for the mean and precision parameters. The values *A* and *B* are extracted from the min and max arguments of [extlogitlink](#). Consequently, only [extlogitlink](#) is allowed.

imu, iphi Optional initial value for the mean and precision parameters respectively. A NULL value means a value is obtained in the initialize slot.

gprobs.y, gphi, zero See [CommonVGAMffArguments](#) for more information.

Details

The two-parameter beta distribution can be written $f(y) =$

$$(y - A)^{\mu_1 \phi - 1} \times (B - y)^{(1 - \mu_1) \phi - 1} / [\text{beta}(\mu_1 \phi, (1 - \mu_1) \phi) \times (B - A)^{\phi - 1}]$$

for $A < y < B$, and $\text{beta}(\cdot, \cdot)$ is the beta function (see [beta](#)). The parameter μ_1 satisfies $\mu_1 = (\mu - A) / (B - A)$ where μ is the mean of Y . That is, μ_1 is the mean of a standard beta distribution:

$E(Y) = A + (B - A) \times \mu_1$, and these are the fitted values of the object. Also, ϕ is positive and $A < \mu < B$. Here, the limits A and B are *known*.

Another parameterization of the beta distribution involving the raw shape parameters is implemented in [betaR](#).

For general A and B , the variance of Y is $(B - A)^2 \times \mu_1 \times (1 - \mu_1)/(1 + \phi)$. Then ϕ can be interpreted as a *precision* parameter in the sense that, for fixed μ , the larger the value of ϕ , the smaller the variance of Y . Also, $\mu_1 = \text{shape1}/(\text{shape1} + \text{shape2})$ and $\phi = \text{shape1} + \text{shape2}$. Fisher scoring is implemented.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Note

The response must have values in the interval (A, B) . The user currently needs to manually choose `lmu` to match the input of arguments `A` and `B`, e.g., with [extlogitlink](#); see the example below.

Author(s)

Thomas W. Yee

References

Ferrari, S. L. P. and Francisco C.-N. (2004). Beta regression for modelling rates and proportions. *Journal of Applied Statistics*, **31**, 799–815.

See Also

[betaR](#), [Beta](#), [dzoabeta](#), [genbetaII](#), [betaII](#), [betabinomialff](#), [betageometric](#), [betaprime](#), [rbetageom](#), [rbetanorm](#), [kumar](#), [extlogitlink](#), [simulate.vlm](#).

Examples

```
bdata <- data.frame(y = rbeta(nn <- 1000, shape1 = exp(0),
                    shape2 = exp(1)))
fit1 <- vglm(y ~ 1, betaff, data = bdata, trace = TRUE)
coef(fit1, matrix = TRUE)
Coef(fit1) # Useful for intercept-only models

# General A and B, and with a covariate
bdata <- transform(bdata, x2 = runif(nn))
bdata <- transform(bdata, mu = logitlink(0.5 - x2, inverse = TRUE),
                  prec = exp(3.0 + x2)) # prec == phi
bdata <- transform(bdata, shape2 = prec * (1 - mu),
                  shape1 = mu * prec)
bdata <- transform(bdata,
                  y = rbeta(nn, shape1 = shape1, shape2 = shape2))
bdata <- transform(bdata, Y = 5 + 8 * y) # From 5--13, not 0--1
```

```
fit <- vglm(Y ~ x2, data = bdata, trace = TRUE,
  betaff(A = 5, B = 13, lmu = extlogitlink(min = 5, max = 13)))
coef(fit, matrix = TRUE)
```

 Betageom

The Beta-Geometric Distribution

Description

Density, distribution function, and random generation for the beta-geometric distribution.

Usage

```
dbetageom(x, shape1, shape2, log = FALSE)
pbetageom(q, shape1, shape2, log.p = FALSE)
rbetageom(n, shape1, shape2)
```

Arguments

x, q	vector of quantiles.
n	number of observations. Same as runif .
shape1, shape2	the two (positive) shape parameters of the standard beta distribution. They are called a and b in beta respectively.
log, log.p	Logical. If TRUE then all probabilities p are given as log(p).

Details

The beta-geometric distribution is a geometric distribution whose probability of success is not a constant but it is generated from a beta distribution with parameters shape1 and shape2. Note that the mean of this beta distribution is $\text{shape1}/(\text{shape1}+\text{shape2})$, which therefore is the mean of the probability of success.

Value

dbetageom gives the density, pbetageom gives the distribution function, and rbetageom generates random deviates.

Note

pbetageom can be particularly slow.

Author(s)

T. W. Yee

See Also

[geometric](#), [betaff](#), [Beta](#).

Examples

```
## Not run:
shape1 <- 1; shape2 <- 2; y <- 0:30
proby <- dbetageom(y, shape1, shape2, log = FALSE)
plot(y, proby, type = "h", col = "blue", ylab = "P[Y=y]", main = paste0(
  "Y ~ Beta-geometric(shape1=", shape1,", shape2=", shape2, ")"))
sum(proby)

## End(Not run)
```

betageometric

*Beta-geometric Distribution Family Function***Description**

Maximum likelihood estimation for the beta-geometric distribution.

Usage

```
betageometric(lprob = "logitlink", lshape = "loglink",
  iprob = NULL,   ishape = 0.1,
  moreSummation = c(2, 100), tolerance = 1.0e-10, zero = NULL)
```

Arguments

lprob, lshape	Parameter link functions applied to the parameters p and ϕ (called prob and shape below). The former lies in the unit interval and the latter is positive. See Links for more choices.
iprob, ishape	Numeric. Initial values for the two parameters. A NULL means a value is computed internally.
moreSummation	Integer, of length 2. When computing the expected information matrix a series summation from 0 to $\text{moreSummation}[1] \cdot \max(y) + \text{moreSummation}[2]$ is made, in which the upper limit is an approximation to infinity. Here, y is the response.
tolerance	Positive numeric. When all terms are less than this then the series is deemed to have converged.
zero	An integer-valued vector specifying which linear/additive predictors are modelled as intercepts only. If used, the value must be from the set $\{1,2\}$. See CommonVGAMffArguments for more information.

Details

A random variable Y has a 2-parameter beta-geometric distribution if $P(Y = y) = p(1 - p)^y$ for $y = 0, 1, 2, \dots$ where p are generated from a standard beta distribution with shape parameters shape1 and shape2. The parameterization here is to focus on the parameters p and $\phi =$

$1/(shape1 + shape2)$, where ϕ is shape. The default link functions for these ensure that the appropriate range of the parameters is maintained. The mean of Y is $E(Y) = shape2/(shape1 - 1) = (1 - p)/(p - \phi)$ if $shape1 > 1$, and if so, then this is returned as the fitted values.

The geometric distribution is a special case of the beta-geometric distribution with $\phi = 0$ (see [geometric](#)). However, fitting data from a geometric distribution may result in numerical problems because the estimate of $\log(\phi)$ will 'converge' to $-\text{Inf}$.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Note

The first iteration may be very slow; if practical, it is best for the weights argument of [vglm](#) etc. to be used rather than inputting a very long vector as the response, i.e., `vglm(y ~ 1, ..., weights = wts)` is to be preferred over `vglm(rep(y, wts) ~ 1, ...)`. If convergence problems occur try inputting some values of argument `ishape`.

If an intercept-only model is fitted then the `misc` slot of the fitted object has list components `shape1` and `shape2`.

Author(s)

T. W. Yee

References

Paul, S. R. (2005). Testing goodness of fit of the geometric distribution: an application to human fecundability data. *Journal of Modern Applied Statistical Methods*, **4**, 425–433.

See Also

[geometric](#), [betaff](#), [rbetageom](#).

Examples

```
## Not run:
bdata <- data.frame(y = 0:11,
                   wts = c(227,123,72,42,21,31,11,14,6,4,7,28))
fitb <- vglm(y ~ 1, betageometric, bdata, weight = wts, trace = TRUE)
fitg <- vglm(y ~ 1, geometric, bdata, weight = wts, trace = TRUE)
coef(fitb, matrix = TRUE)
Coef(fitb)
sqrt(diag(vcov(fitb, untransform = TRUE)))
fitb@misc$shape1
fitb@misc$shape2
# Very strong evidence of a beta-geometric:
pchisq(2 * (logLik(fitb) - logLik(fitg)), df = 1, lower.tail = FALSE)

## End(Not run)
```

betaII	<i>Beta Distribution of the Second Kind</i>
--------	---

Description

Maximum likelihood estimation of the 3-parameter beta II distribution.

Usage

```
betaII(lscale = "loglink", lshape2.p = "loglink",
       lshape3.q = "loglink", iscale = NULL, ishape2.p = NULL,
       ishape3.q = NULL, imethod = 1,
       gscale = exp(-5:5), gshape2.p = exp(-5:5),
       gshape3.q = seq(0.75, 4, by = 0.25),
       probs.y = c(0.25, 0.5, 0.75), zero = "shape")
```

Arguments

`lscale`, `lshape2.p`, `lshape3.q`
 Parameter link functions applied to the (positive) parameters `scale`, `p` and `q`.
 See [Links](#) for more choices.

`iscale`, `ishape2.p`, `ishape3.q`, `imethod`, `zero`
 See [CommonVGAMffArguments](#) for information.

`gscale`, `gshape2.p`, `gshape3.q`
 See [CommonVGAMffArguments](#) for information.

`probs.y`
 See [CommonVGAMffArguments](#) for information.

Details

The 3-parameter beta II is the 4-parameter *generalized* beta II distribution with shape parameter $a = 1$. It is also known as the Pearson VI distribution. Other distributions which are special cases of the 3-parameter beta II include the Lomax ($p = 1$) and inverse Lomax ($q = 1$). More details can be found in Kleiber and Kotz (2003).

The beta II distribution has density

$$f(y) = y^{p-1} / [b^p B(p, q) \{1 + y/b\}^{p+q}]$$

for $b > 0$, $p > 0$, $q > 0$, $y \geq 0$. Here, b is the scale parameter `scale`, and the others are shape parameters. The mean is

$$E(Y) = b \Gamma(p+1) \Gamma(q-1) / (\Gamma(p) \Gamma(q))$$

provided $q > 1$; these are returned as the fitted values. This family function handles multiple responses.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Note

See the notes in [genbetaII](#).

Author(s)

T. W. Yee

References

Kleiber, C. and Kotz, S. (2003). *Statistical Size Distributions in Economics and Actuarial Sciences*, Hoboken, NJ, USA: Wiley-Interscience.

See Also

[betaff](#), [genbetaII](#), [dagum](#), [sinmad](#), [fisk](#), [inv.lomax](#), [lomax](#), [paralogistic](#), [inv.paralogistic](#).

Examples

```
bdata <- data.frame(y = rsinmad(2000, shape1.a = 1,
                             shape3.q = exp(2), scale = exp(1))) # Not genuine data!
# fit <- vglm(y ~ 1, betaII, data = bdata, trace = TRUE)
fit <- vglm(y ~ 1, betaII(ishape2.p = 0.7, ishape3.q = 0.7),
           data = bdata, trace = TRUE)
coef(fit, matrix = TRUE)
Coef(fit)
summary(fit)
```

Betanorm

The Beta-Normal Distribution

Description

Density, distribution function, quantile function and random generation for the univariate beta-normal distribution.

Usage

```
dbetanorm(x, shape1, shape2, mean = 0, sd = 1, log = FALSE)
pbetanorm(q, shape1, shape2, mean = 0, sd = 1,
          lower.tail = TRUE, log.p = FALSE)
qbetanorm(p, shape1, shape2, mean = 0, sd = 1,
          lower.tail = TRUE, log.p = FALSE)
rbetanorm(n, shape1, shape2, mean = 0, sd = 1)
```

Arguments

x, q	vector of quantiles.
p	vector of probabilities.
n	number of observations. Same as <code>runif</code> .
shape1, shape2	the two (positive) shape parameters of the standard beta distribution. They are called a and b respectively in <code>beta</code> .
mean, sd	the mean and standard deviation of the univariate normal distribution (<code>Normal</code>).
log, log.p	Logical. If TRUE then all probabilities p are given as log(p).
lower.tail	Logical. If TRUE then the upper tail is returned, i.e., one minus the usual answer.

Details

The function `betauninormal`, the **VGAM** family function for estimating the parameters, has not yet been written.

Value

`dbetanorm` gives the density, `pbetanorm` gives the distribution function, `qbetanorm` gives the quantile function, and `rbetanorm` generates random deviates.

Author(s)

T. W. Yee

References

Gupta, A. K. and Nadarajah, S. (2004). *Handbook of Beta Distribution and Its Applications*, pp.146–152. New York: Marcel Dekker.

Examples

```
## Not run:
shape1 <- 0.1; shape2 <- 4; m <- 1
x <- seq(-10, 2, len = 501)
plot(x, dbetanorm(x, shape1, shape2, m = m), type = "l",
      ylim = 0:1, las = 1,
      ylab = paste0("betanorm(", shape1, ", ", shape2, ", m=", m, ", sd=1)"),
      main = "Blue is density, orange is the CDF",
      sub = "Gray lines are the 10,20,...,90 percentiles", col = "blue")
lines(x, pbetanorm(x, shape1, shape2, m = m), col = "orange")
abline(h = 0, col = "black")
probs <- seq(0.1, 0.9, by = 0.1)
Q <- qbetanorm(probs, shape1, shape2, m = m)
lines(Q, dbetanorm(Q, shape1, shape2, m = m),
      col = "gray50", lty = 2, type = "h")
lines(Q, pbetanorm(Q, shape1, shape2, m = m),
      col = "gray50", lty = 2, type = "h")
abline(h = probs, col = "gray50", lty = 2)
pbetanorm(Q, shape1, shape2, m = m) - probs # Should be all 0
```

```
## End(Not run)
```

 betaprime

The Beta-Prime Distribution

Description

Estimation of the two shape parameters of the beta-prime distribution by maximum likelihood estimation.

Usage

```
betaprime(lshape = "loglink", ishape1 = 2, ishape2 = NULL,
          zero = NULL)
```

Arguments

`lshape` Parameter link function applied to the two (positive) shape parameters. See [Links](#) for more choices.

`ishape1`, `ishape2`, `zero` See [CommonVGAMffArguments](#) for more information.

Details

The beta-prime distribution is given by

$$f(y) = y^{\text{shape1}-1}(1+y)^{-\text{shape1}-\text{shape2}}/B(\text{shape1}, \text{shape2})$$

for $y > 0$. The shape parameters are positive, and here, B is the beta function. The mean of Y is $\text{shape1}/(\text{shape2} - 1)$ provided $\text{shape2} > 1$; these are returned as the fitted values.

If Y has a $Beta(\text{shape1}, \text{shape2})$ distribution then $Y/(1-Y)$ and $(1-Y)/Y$ have a $Betaprime(\text{shape1}, \text{shape2})$ and $Betaprime(\text{shape2}, \text{shape1})$ distribution respectively. Also, if Y_1 has a $gamma(\text{shape1})$ distribution and Y_2 has a $gamma(\text{shape2})$ distribution then Y_1/Y_2 has a $Betaprime(\text{shape1}, \text{shape2})$ distribution.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), [rrvglm](#) and [vgam](#).

Note

The response must have positive values only.

The beta-prime distribution is also known as the *beta distribution of the second kind* or the *inverted beta distribution*.

Author(s)

Thomas W. Yee

References

Johnson, N. L. and Kotz, S. and Balakrishnan, N. (1995). Chapter 25 of: *Continuous Univariate Distributions*, 2nd edition, Volume 2, New York: Wiley.

See Also

[betaff](#), [Beta](#).

Examples

```
nn <- 1000
bdata <- data.frame(shape1 = exp(1), shape2 = exp(3))
bdata <- transform(bdata, yb = rbeta(nn, shape1, shape2))
bdata <- transform(bdata, y1 = (1-yb) / yb,
                  y2 = yb / (1-yb),
                  y3 = rgamma(nn, exp(3)) / rgamma(nn, exp(2)))

fit1 <- vglm(y1 ~ 1, betaprime, data = bdata, trace = TRUE)
coef(fit1, matrix = TRUE)

fit2 <- vglm(y2 ~ 1, betaprime, data = bdata, trace = TRUE)
coef(fit2, matrix = TRUE)

fit3 <- vglm(y3 ~ 1, betaprime, data = bdata, trace = TRUE)
coef(fit3, matrix = TRUE)

# Compare the fitted values
with(bdata, mean(y3))
head(fitted(fit3))
Coef(fit3) # Useful for intercept-only models
```

betaR

The Two-parameter Beta Distribution Family Function

Description

Estimation of the shape parameters of the two-parameter beta distribution.

Usage

```
betaR(lshape1 = "loglink", lshape2 = "loglink",
      i1 = NULL, i2 = NULL, trim = 0.05,
      A = 0, B = 1, parallel = FALSE, zero = NULL)
```

Arguments

lshape1, lshape2, i1, i2	Details at CommonVGAMffArguments . See Links for more choices.
trim	An argument which is fed into <code>mean()</code> ; it is the fraction (0 to 0.5) of observations to be trimmed from each end of the response <code>y</code> before the mean is computed. This is used when computing initial values, and guards against outliers.
A, B	Lower and upper limits of the distribution. The defaults correspond to the <i>standard beta distribution</i> where the response lies between 0 and 1.
parallel, zero	See CommonVGAMffArguments for more information.

Details

The two-parameter beta distribution is given by $f(y) =$

$$(y - A)^{shape1-1} \times (B - y)^{shape2-1} / [Beta(shape1, shape2) \times (B - A)^{shape1+shape2-1}]$$

for $A < y < B$, and $Beta(.,.)$ is the beta function (see [beta](#)). The shape parameters are positive, and here, the limits A and B are known. The mean of Y is $E(Y) = A + (B - A) \times shape1 / (shape1 + shape2)$, and these are the fitted values of the object.

For the standard beta distribution the variance of Y is $shape1 \times shape2 / [(1 + shape1 + shape2) \times (shape1 + shape2)^2]$. If $\sigma^2 = 1 / (1 + shape1 + shape2)$ then the variance of Y can be written $\sigma^2 \mu(1 - \mu)$ where $\mu = shape1 / (shape1 + shape2)$ is the mean of Y .

Another parameterization of the beta distribution involving the mean and a precision parameter is implemented in [betaff](#).

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), [rrvglm](#) and [vgam](#).

Note

The response must have values in the interval (A, B) . **VGAM 0.7-4** and prior called this function [betaff](#).

Author(s)

Thomas W. Yee

References

Johnson, N. L. and Kotz, S. and Balakrishnan, N. (1995). Chapter 25 of: *Continuous Univariate Distributions*, 2nd edition, Volume 2, New York: Wiley.

Gupta, A. K. and Nadarajah, S. (2004). *Handbook of Beta Distribution and Its Applications*, New York: Marcel Dekker.

See Also

[betaff](#), [Beta](#), [genbetaII](#), [betaII](#), [betabinomialff](#), [betageometric](#), [betaprime](#), [rbetageom](#), [rbetanorm](#), [kumar](#), [simulate.vlm](#).

Examples

```
bdata <- data.frame(y = rbeta(1000, shape1 = exp(0), shape2 = exp(1)))
fit <- vglm(y ~ 1, betaR(lshape1 = "identitylink",
  lshape2 = "identitylink"), bdata, trace = TRUE, crit = "coef")
fit <- vglm(y ~ 1, betaR, data = bdata, trace = TRUE, crit = "coef")
coef(fit, matrix = TRUE)
Coef(fit) # Useful for intercept-only models

bdata <- transform(bdata, Y = 5 + 8 * y) # From 5 to 13, not 0 to 1
fit <- vglm(Y ~ 1, betaR(A = 5, B = 13), data = bdata, trace = TRUE)
Coef(fit)
c(meanY = with(bdata, mean(Y)), head(fitted(fit),2))
```

Biamhcop

*Ali-Mikhail-Haq Bivariate Distribution***Description**

Density, distribution function, and random generation for the (one parameter) bivariate Ali-Mikhail-Haq distribution.

Usage

```
dbiamhcop(x1, x2, apar, log = FALSE)
pbiamhcop(q1, q2, apar)
rbiamhcop(n, apar)
```

Arguments

`x1, x2, q1, q2` vector of quantiles.
`n` number of observations. Same as [runif](#)
`apar` the association parameter.
`log` Logical. If TRUE then the logarithm is returned.

Details

See [biamhcop](#), the **VGAM** family functions for estimating the parameter by maximum likelihood estimation, for the formula of the cumulative distribution function and other details.

Value

`dbiamhcop` gives the density, `pbiamhcop` gives the distribution function, and `rbiamhcop` generates random deviates (a two-column matrix).

Author(s)

T. W. Yee and C. S. Chee

See Also

[biamhcop](#).

Examples

```
x <- seq(0, 1, len = (N <- 101)); apar <- 0.7
ox <- expand.grid(x, x)
zedd <- dbiamhcop(ox[, 1], ox[, 2], apar = apar)
## Not run:
contour(x, x, matrix(zedd, N, N), col = "blue")
zedd <- pbiamhcop(ox[, 1], ox[, 2], apar = apar)
contour(x, x, matrix(zedd, N, N), col = "blue")

plot(r <- rbiamhcop(n = 1000, apar = apar), col = "blue")
par(mfrow = c(1, 2))
hist(r[, 1]) # Should be uniform
hist(r[, 2]) # Should be uniform

## End(Not run)
```

biamhcop

Ali-Mikhail-Haq Distribution Family Function

Description

Estimate the association parameter of Ali-Mikhail-Haq's bivariate distribution by maximum likelihood estimation.

Usage

```
biamhcop(lapar = "rhopitlink", iapar = NULL, imethod = 1,
         nsimEIM = 250)
```

Arguments

lapar	Link function applied to the association parameter α , which is real and $-1 < \alpha < 1$. See Links for more choices.
iapar	Numeric. Optional initial value for α . By default, an initial value is chosen internally. If a convergence failure occurs try assigning a different value. Assigning a value will override the argument imethod.
imethod	An integer with value 1 or 2 which specifies the initialization method. If failure to converge occurs try the other value, or else specify a value for iapar.
nsimEIM	See CommonVGAMffArguments for more information.

Details

The cumulative distribution function is

$$P(Y_1 \leq y_1, Y_2 \leq y_2) = y_1 y_2 / (1 - \alpha(1 - y_1)(1 - y_2))$$

for $-1 < \alpha < 1$. The support of the function is the unit square. The marginal distributions are the standard uniform distributions. When $\alpha = 0$ the random variables are independent. This is an Archimedean copula.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Note

The response must be a two-column matrix. Currently, the fitted value is a matrix with two columns and values equal to 0.5. This is because each marginal distribution corresponds to a standard uniform distribution.

Author(s)

T. W. Yee and C. S. Chee

References

Balakrishnan, N. and Lai, C.-D. (2009). *Continuous Bivariate Distributions*, 2nd ed. New York: Springer.

See Also

[rbiamhcop](#), [bifgmcop](#), [bigumbelIexp](#), [rbilogis](#), [simulate.vlm](#).

Examples

```
ymat <- rbiamhcop(1000, apar = rhobitlink(2, inverse = TRUE))
fit <- vglm(ymat ~ 1, biamhcop, trace = TRUE)
coef(fit, matrix = TRUE)
Coef(fit)
```

Biclaytoncop*Clayton Copula (Bivariate) Distribution*

Description

Density and random generation for the (one parameter) bivariate Clayton copula distribution.

Usage

```
dbiclaytoncop(x1, x2, apar = 0, log = FALSE)
rbiclaytoncop(n, apar = 0)
```

Arguments

<code>x1, x2</code>	vector of quantiles. The <code>x1</code> and <code>x2</code> should both be in the interval $(0, 1)$.
<code>n</code>	number of observations. Same as rnorm .
<code>apar</code>	the association parameter. Should be in the interval $[0, \infty)$. The default corresponds to independence.
<code>log</code>	Logical. If TRUE then the logarithm is returned.

Details

See [biclaytoncop](#), the **VGAM** family functions for estimating the parameter by maximum likelihood estimation, for the formula of the cumulative distribution function and other details.

Value

`dbiclaytoncop` gives the density at point $(x1,x2)$, `rbiclaytoncop` generates random deviates (a two-column matrix).

Note

`dbiclaytoncop()` does not yet handle $x1 = 0$ and/or $x2 = 0$.

Author(s)

R. Feyter and T. W. Yee

References

Clayton, D. (1982). A model for association in bivariate survival data. *Journal of the Royal Statistical Society, Series B, Methodological*, **44**, 414–422.

See Also

[biclaytoncop](#), [binormalcop](#), [binormal](#).

Examples

```
## Not run: edge <- 0.01 # A small positive value
N <- 101; x <- seq(edge, 1.0 - edge, len = N); Rho <- 0.7
ox <- expand.grid(x, x)
zedd <- dbiclaytoncop(ox[, 1], ox[, 2], apar = Rho, log = TRUE)
par(mfrow = c(1, 2))
contour(x, matrix(zedd, N, N), col = 4, labcex = 1.5, las = 1)
plot(rbiclaytoncop(1000, 2), col = 4, las = 1)
## End(Not run)
```

biclaytoncop

*Clayton Copula (Bivariate) Family Function***Description**

Estimate the correlation parameter of the (bivariate) Clayton copula distribution by maximum likelihood estimation.

Usage

```
biclaytoncop(lapar = "loglink", iapar = NULL, imethod = 1,
             parallel = FALSE, zero = NULL)
```

Arguments

lapar, iapar, imethod

Details at [CommonVGAMffArguments](#). See [Links](#) for more link function choices.

parallel, zero Details at [CommonVGAMffArguments](#). If parallel = TRUE then the constraint is also applied to the intercept.

Details

The cumulative distribution function is

$$P(u_1, u_2; \alpha) = (u_1^{-\alpha} + u_2^{-\alpha} - 1)^{-1/\alpha}$$

for $0 \leq \alpha$. Here, α is the association parameter. The support of the function is the interior of the unit square; however, values of 0 and/or 1 are not allowed (currently). The marginal distributions are the standard uniform distributions. When $\alpha = 0$ the random variables are independent.

This **VGAM** family function can handle multiple responses, for example, a six-column matrix where the first 2 columns is the first out of three responses, the next 2 columns being the next response, etc.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Note

The response matrix must have a multiple of two-columns. Currently, the fitted value is a matrix with the same number of columns and values equal to 0.5. This is because each marginal distribution corresponds to a standard uniform distribution.

This **VGAM** family function is fragile; each response must be in the interior of the unit square.

Author(s)

R. Feyter and T. W. Yee

References

Clayton, D. (1982). A model for association in bivariate survival data. *Journal of the Royal Statistical Society, Series B, Methodological*, **44**, 414–422.

Schepsmeier, U. and Stober, J. (2014). Derivatives and Fisher information of bivariate copulas. *Statistical Papers* **55**, 525–542.

See Also

[rbiclaytoncop](#), [dbiclaytoncop](#), [kendall.tau](#).

Examples

```
ymat <- rbiclaytoncop(n = (nn <- 1000), apar = exp(2))
bdata <- data.frame(y1 = ymat[, 1], y2 = ymat[, 2],
                  y3 = ymat[, 1], y4 = ymat[, 2], x2 = runif(nn))
summary(bdata)
## Not run: plot(ymat, col = "blue")
fit1 <-
  vglm(cbind(y1, y2, y3, y4) ~ 1, # 2 responses, e.g., (y1,y2) is the 1st
       biclaytoncop, data = bdata,
       trace = TRUE, crit = "coef") # Sometimes a good idea
coef(fit1, matrix = TRUE)
Coef(fit1)
head(fitted(fit1))
summary(fit1)

# Another example; apar is a function of x2
bdata <- transform(bdata, apar = exp(-0.5 + x2))
ymat <- rbiclaytoncop(n = nn, apar = with(bdata, apar))
bdata <- transform(bdata, y5 = ymat[, 1], y6 = ymat[, 2])
fit2 <- vgam(cbind(y5, y6) ~ s(x2), data = bdata,
            biclaytoncop(lapar = "loglink"), trace = TRUE)
## Not run: plot(fit2, lcol = "blue", scol = "orange", se = TRUE)
```

`BICv1m`*Bayesian Information Criterion*

Description

Calculates the Bayesian information criterion (BIC) for a fitted model object for which a log-likelihood value has been obtained.

Usage

```
BICv1m(object, ..., k = log(nobs(object)))
```

Arguments

<code>object, ...</code>	Same as AICv1m .
<code>k</code>	Numeric, the penalty per parameter to be used; the default is $\log(n)$ where n is the number of observations).

Details

The so-called BIC or SBC (Schwarz's Bayesian criterion) can be computed by calling [AICv1m](#) with a different `k` argument. See [AICv1m](#) for information and caveats.

Value

Returns a numeric value with the corresponding BIC, or ..., depending on `k`.

Warning

Like [AICv1m](#), this code has not been double-checked. The general applicability of BIC for the VGLM/VGAM classes has not been developed fully. In particular, BIC should not be run on some **VGAM** family functions because of violation of certain regularity conditions, etc.

Many **VGAM** family functions such as [cumulative](#) can have the number of observations absorbed into the prior weights argument (e.g., `weights` in [vglm](#)), either before or after fitting. Almost all **VGAM** family functions can have the number of observations defined by the `weights` argument, e.g., as an observed frequency. BIC simply uses the number of rows of the model matrix, say, as defining n , hence the user must be very careful of this possible error. Use at your own risk!!

Note

BIC, AIC and other ICs can have many additive constants added to them. The important thing are the differences since the minimum value corresponds to the best model.

BIC has not been defined for QRR-VGLMs yet.

Author(s)

T. W. Yee.

See Also

[AICvglm](#), VGLMs are described in [vglm-class](#); VGAMs are described in [vgam-class](#); RR-VGLMs are described in [rrvglm-class](#); [BIC](#), [AIC](#).

Examples

```
pneumo <- transform(pneumo, let = log(exposure.time))
(fit1 <- vglm(cbind(normal, mild, severe) ~ let,
             cumulative(parallel = TRUE, reverse = TRUE), data = pneumo))
coef(fit1, matrix = TRUE)
BIC(fit1)
(fit2 <- vglm(cbind(normal, mild, severe) ~ let,
             cumulative(parallel = FALSE, reverse = TRUE), data = pneumo))
coef(fit2, matrix = TRUE)
BIC(fit2)
```

Bifgmcop

Farlie-Gumbel-Morgenstern's Bivariate Distribution

Description

Density, distribution function, and random generation for the (one parameter) bivariate Farlie-Gumbel-Morgenstern's distribution.

Usage

```
dbifgmcop(x1, x2, apar, log = FALSE)
pbifgmcop(q1, q2, apar)
rbifgmcop(n, apar)
```

Arguments

x1, x2, q1, q2	vector of quantiles.
n	number of observations. Same as in runif .
apar	the association parameter.
log	Logical. If TRUE then the logarithm is returned.

Details

See [bifgmcop](#), the **VGAM** family functions for estimating the parameter by maximum likelihood estimation, for the formula of the cumulative distribution function and other details.

Value

`dbifgmcop` gives the density, `pbifgmcop` gives the distribution function, and `rbifgmcop` generates random deviates (a two-column matrix).

Author(s)

T. W. Yee

See Also[bifgmcop](#).**Examples**

```
## Not run: N <- 101; x <- seq(0.0, 1.0, len = N); apar <- 0.7
ox <- expand.grid(x, x)
zedd <- dbifgmcop(ox[, 1], ox[, 2], apar = apar)
contour(x, x, matrix(zedd, N, N), col = "blue")
zedd <- pbifgmcop(ox[, 1], ox[, 2], apar = apar)
contour(x, x, matrix(zedd, N, N), col = "blue")

plot(r <- rbifgmcop(n = 3000, apar = apar), col = "blue")
par(mfrow = c(1, 2))
hist(r[, 1]) # Should be uniform
hist(r[, 2]) # Should be uniform

## End(Not run)
```

bifgmcop

Farlie-Gumbel-Morgenstern's Bivariate Distribution Family Function

Description

Estimate the association parameter of Farlie-Gumbel-Morgenstern's bivariate distribution by maximum likelihood estimation.

Usage

```
bifgmcop(lapar = "rhobitlink", iapar = NULL, imethod = 1)
```

Arguments

lapar, iapar, imethod

Details at [CommonVGAMffArguments](#). See [Links](#) for more link function choices.

Details

The cumulative distribution function is

$$P(Y_1 \leq y_1, Y_2 \leq y_2) = y_1 y_2 (1 + \alpha(1 - y_1)(1 - y_2))$$

for $-1 < \alpha < 1$. The support of the function is the unit square. The marginal distributions are the standard uniform distributions. When $\alpha = 0$ the random variables are independent.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Note

The response must be a two-column matrix. Currently, the fitted value is a matrix with two columns and values equal to 0.5. This is because each marginal distribution corresponds to a standard uniform distribution.

Author(s)

T. W. Yee

References

- Castillo, E., Hadi, A. S., Balakrishnan, N. and Sarabia, J. S. (2005). *Extreme Value and Related Models with Applications in Engineering and Science*, Hoboken, NJ, USA: Wiley-Interscience.
- Smith, M. D. (2007). Invariance theorems for Fisher information. *Communications in Statistics—Theory and Methods*, **36**(12), 2213–2222.

See Also

[rbifgmcop](#), [bifrankcop](#), [bifgmexp](#), [simulate.vlm](#).

Examples

```
yamat <- rbifgmcop(1000, apar = rhobitlink(3, inverse = TRUE))
## Not run: plot(yamat, col = "blue")
fit <- vglm(yamat ~ 1, fam = bifgmcop, trace = TRUE)
coef(fit, matrix = TRUE)
Coef(fit)
head(fitted(fit))
```

bifgmexp

Bivariate Farlie-Gumbel-Morgenstern Exponential Distribution Family Function

Description

Estimate the association parameter of FGM bivariate exponential distribution by maximum likelihood estimation.

Usage

```
bifgmexp(lapar = "rhobitlink", iapar = NULL, tola0 = 0.01,
         imethod = 1)
```

Arguments

lapar	Link function for the association parameter α , which lies between -1 and 1 . See Links for more choices and other information.
iapar	Numeric. Optional initial value for α . By default, an initial value is chosen internally. If a convergence failure occurs try assigning a different value. Assigning a value will override the argument imethod.
tolao	Positive numeric. If the estimate of α has an absolute value less than this then it is replaced by this value. This is an attempt to fix a numerical problem when the estimate is too close to zero.
imethod	An integer with value 1 or 2 which specifies the initialization method. If failure to converge occurs try the other value, or else specify a value for ia.

Details

The cumulative distribution function is

$$P(Y_1 \leq y_1, Y_2 \leq y_2) = e^{-y_1 - y_2} (1 + \alpha[1 - e^{-y_1}][1 - e^{-y_2}]) + 1 - e^{-y_1} - e^{-y_2}$$

for α between -1 and 1 . The support of the function is for $y_1 > 0$ and $y_2 > 0$. The marginal distributions are an exponential distribution with unit mean. When $\alpha = 0$ then the random variables are independent, and this causes some problems in the estimation process since the distribution no longer depends on the parameter.

A variant of Newton-Raphson is used, which only seems to work for an intercept model. It is a very good idea to set `trace = TRUE`.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Note

The response must be a two-column matrix. Currently, the fitted value is a matrix with two columns and values equal to 1. This is because each marginal distribution corresponds to a exponential distribution with unit mean.

This **VGAM** family function should be used with caution.

Author(s)

T. W. Yee

References

Castillo, E., Hadi, A. S., Balakrishnan, N. and Sarabia, J. S. (2005). *Extreme Value and Related Models with Applications in Engineering and Science*, Hoboken, NJ, USA: Wiley-Interscience.

See Also

[bifgmcop](#), [bigumbelIexp](#).

Examples

```
N <- 1000; mdata <- data.frame(y1 = rexp(N), y2 = rexp(N))
## Not run: plot(yamat)
fit <- vglm(cbind(y1, y2) ~ 1, bifgmexp, data = mdata, trace = TRUE)
fit <- vglm(cbind(y1, y2) ~ 1, bifgmexp, data = mdata, # May fail
           trace = TRUE, crit = "coef")
coef(fit, matrix = TRUE)
Coef(fit)
head(fitted(fit))
```

bifrankcop

Frank's Bivariate Distribution Family Function

Description

Estimate the association parameter of Frank's bivariate distribution by maximum likelihood estimation.

Usage

```
bifrankcop(lapar = "loglink", iapar = 2, nsimEIM = 250)
```

Arguments

lapar	Link function applied to the (positive) association parameter α . See Links for more choices.
iapar	Numeric. Initial value for α . If a convergence failure occurs try assigning a different value.
nsimEIM	See CommonVGAMffArguments .

Details

The cumulative distribution function is

$$P(Y_1 \leq y_1, Y_2 \leq y_2) = H_\alpha(y_1, y_2) = \log_\alpha[1 + (\alpha^{y_1} - 1)(\alpha^{y_2} - 1)/(\alpha - 1)]$$

for $\alpha \neq 1$. Note the logarithm here is to base α . The support of the function is the unit square.

When $0 < \alpha < 1$ the probability density function $h_\alpha(y_1, y_2)$ is symmetric with respect to the lines $y_2 = y_1$ and $y_2 = 1 - y_1$. When $\alpha > 1$ then $h_\alpha(y_1, y_2) = h_{1/\alpha}(1 - y_1, y_2)$.

$\alpha = 1$ then $H(y_1, y_2) = y_1 y_2$, i.e., uniform on the unit square. As α approaches 0 then $H(y_1, y_2) = \min(y_1, y_2)$. As α approaches infinity then $H(y_1, y_2) = \max(0, y_1 + y_2 - 1)$.

The default is to use Fisher scoring implemented using [rbifrankcop](#). For intercept-only models an alternative is to set `nsimEIM=NULL` so that a variant of Newton-Raphson is used.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Note

The response must be a two-column matrix. Currently, the fitted value is a matrix with two columns and values equal to a half. This is because the marginal distributions correspond to a standard uniform distribution.

Author(s)

T. W. Yee

References

Genest, C. (1987). Frank's family of bivariate distributions. *Biometrika*, **74**, 549–555.

See Also

[rbifrankcop](#), [bifgmcop](#), [simulate.vlm](#).

Examples

```
## Not run:
ymat <- rbifrankcop(n = 2000, apar = exp(4))
plot(ymat, col = "blue")
fit <- vglm(ymat ~ 1, fam = bifrankcop, trace = TRUE)
coef(fit, matrix = TRUE)
Coef(fit)
vcov(fit)
head(fitted(fit))
summary(fit)

## End(Not run)
```

bigumbelIexp

Gumbel's Type I Bivariate Distribution Family Function

Description

Estimate the association parameter of Gumbel's Type I bivariate distribution by maximum likelihood estimation.

Usage

```
bigumbelIexp(lapar = "identitylink", iapar = NULL, imethod = 1)
```

Arguments

lpar	Link function applied to the association parameter α . See Links for more choices.
ipar	Numeric. Optional initial value for α . By default, an initial value is chosen internally. If a convergence failure occurs try assigning a different value. Assigning a value will override the argument imethod.
imethod	An integer with value 1 or 2 which specifies the initialization method. If failure to converge occurs try the other value, or else specify a value for ipar.

Details

The cumulative distribution function is

$$P(Y_1 \leq y_1, Y_2 \leq y_2) = e^{-y_1 - y_2 + \alpha y_1 y_2} + 1 - e^{-y_1} - e^{-y_2}$$

for real α . The support of the function is for $y_1 > 0$ and $y_2 > 0$. The marginal distributions are an exponential distribution with unit mean.

A variant of Newton-Raphson is used, which only seems to work for an intercept model. It is a very good idea to set trace=TRUE.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Note

The response must be a two-column matrix. Currently, the fitted value is a matrix with two columns and values equal to 1. This is because each marginal distribution corresponds to a exponential distribution with unit mean.

This **VGAM** family function should be used with caution.

Author(s)

T. W. Yee

References

Gumbel, E. J. (1960). Bivariate Exponential Distributions. *Journal of the American Statistical Association*, **55**, 698–707.

See Also

[bifgmexp](#).

Examples

```

nn <- 1000
gdata <- data.frame(y1 = rexp(nn), y2 = rexp(nn))
## Not run: with(gdata, plot(cbind(y1, y2)))
fit <- vglm(cbind(y1, y2) ~ 1, bigumbelIexp, gdata, trace = TRUE)
coef(fit, matrix = TRUE)
Coef(fit)
head(fitted(fit))

```

bilogis

Bivariate Logistic Distribution

Description

Density, distribution function, quantile function and random generation for the 4-parameter bivariate logistic distribution.

Usage

```

dbilogis(x1, x2, loc1 = 0, scale1 = 1, loc2 = 0, scale2 = 1,
         log = FALSE)
pbilogis(q1, q2, loc1 = 0, scale1 = 1, loc2 = 0, scale2 = 1)
rbilogis(n, loc1 = 0, scale1 = 1, loc2 = 0, scale2 = 1)

```

Arguments

x1, x2, q1, q2	vector of quantiles.
n	number of observations. Same as rlogis .
loc1, loc2	the location parameters l_1 and l_2 .
scale1, scale2	the scale parameters s_1 and s_2 .
log	Logical. If log = TRUE then the logarithm of the density is returned.

Details

See [bilogis](#), the **VGAM** family function for estimating the four parameters by maximum likelihood estimation, for the formula of the cumulative distribution function and other details.

Value

dbilogis gives the density, pbilogis gives the distribution function, and rbilogis generates random deviates (a two-column matrix).

Note

Gumbel (1961) proposed two bivariate logistic distributions with logistic distribution marginals, which he called Type I and Type II. The Type I is this one. The Type II belongs to the Morgenstern type. The [biamhcop](#) distribution has, as a special case, this distribution, which is when the random variables are independent.

Author(s)

T. W. Yee

References

Gumbel, E. J. (1961). Bivariate logistic distributions. *Journal of the American Statistical Association*, **56**, 335–349.

See Also

[bilogistic](#), [biamhcop](#).

Examples

```
## Not run: par(mfrow = c(1, 3))
ymat <- rbilogis(n = 2000, loc1 = 5, loc2 = 7, scale2 = exp(1))
myxlim <- c(-2, 15); myylim <- c(-10, 30)
plot(ymat, xlim = myxlim, ylim = myylim)

N <- 100
x1 <- seq(myxlim[1], myxlim[2], len = N)
x2 <- seq(myylim[1], myylim[2], len = N)
ox <- expand.grid(x1, x2)
z <- dbilogis(ox[,1], ox[,2], loc1 = 5, loc2 = 7, scale2 = exp(1))
contour(x1, x2, matrix(z, N, N), main = "density")
z <- pbilogis(ox[,1], ox[,2], loc1 = 5, loc2 = 7, scale2 = exp(1))
contour(x1, x2, matrix(z, N, N), main = "cdf")
## End(Not run)
```

 bilogistic

Bivariate Logistic Distribution Family Function

Description

Estimates the four parameters of the bivariate logistic distribution by maximum likelihood estimation.

Usage

```
bilogistic(llocation = "identitylink", lscale = "loglink",
           iloc1 = NULL, iscale1 = NULL, iloc2 = NULL, iscale2 =
           NULL, imethod = 1, nsimEIM = 250, zero = NULL)
```


Arguments

llocation	Link function applied to both location parameters l_1 and l_2 . See Links for more choices.
lscale	Parameter link function applied to both (positive) scale parameters s_1 and s_2 . See Links for more choices.
iloc1, iloc2	Initial values for the location parameters. By default, initial values are chosen internally using <code>imethod</code> . Assigning values here will override the argument <code>imethod</code> .
iscale1, iscale2	Initial values for the scale parameters. By default, initial values are chosen internally using <code>imethod</code> . Assigning values here will override the argument <code>imethod</code> .
imethod	An integer with value 1 or 2 which specifies the initialization method. If failure to converge occurs try the other value.
nsimEIM, zero	See CommonVGAMffArguments for details.

Details

The four-parameter bivariate logistic distribution has a density that can be written as

$$f(y_1, y_2; l_1, s_1, l_2, s_2) = 2 \frac{\exp[-(y_1 - l_1)/s_1 - (y_2 - l_2)/s_2]}{s_1 s_2 (1 + \exp[-(y_1 - l_1)/s_1] + \exp[-(y_2 - l_2)/s_2])^3}$$

where $s_1 > 0$ and $s_2 > 0$ are the scale parameters, and l_1 and l_2 are the location parameters. Each of the two responses are unbounded, i.e., $-\infty < y_j < \infty$. The mean of Y_1 is l_1 etc. The fitted values are returned in a 2-column matrix. The cumulative distribution function is

$$F(y_1, y_2; l_1, s_1, l_2, s_2) = (1 + \exp[-(y_1 - l_1)/s_1] + \exp[-(y_2 - l_2)/s_2])^{-1}$$

The marginal distribution of Y_1 is

$$P(Y_1 \leq y_1) = F(y_1; l_1, s_1) = (1 + \exp[-(y_1 - l_1)/s_1])^{-1}.$$

By default, $\eta_1 = l_1$, $\eta_2 = \log(s_1)$, $\eta_3 = l_2$, $\eta_4 = \log(s_2)$ are the linear/additive predictors.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), [rrvglm](#) and [vgam](#).

Author(s)

T. W. Yee

References

- Gumbel, E. J. (1961). Bivariate logistic distributions. *Journal of the American Statistical Association*, **56**, 335–349.
- Castillo, E., Hadi, A. S., Balakrishnan, N. and Sarabia, J. S. (2005). *Extreme Value and Related Models with Applications in Engineering and Science*, Hoboken, NJ, USA: Wiley-Interscience.

See Also

[logistic](#), [rbilogis](#).

Examples

```
## Not run:
ymat <- rbilogis(n <- 50, loc1 = 5, loc2 = 7, scale2 = exp(1))
plot(ymat)
bfit <- vglm(ymat ~ 1, family = bilogistic, trace = TRUE)
coef(bfit, matrix = TRUE)
Coef(bfit)
head(fitted(bfit))
vcov(bfit)
head(weights(bfit, type = "work"))
summary(bfit)

## End(Not run)
```

 Binom2.or

Bivariate Odds Ratio Model

Description

Density and random generation for a bivariate binary regression model using an odds ratio as the measure of dependency.

Usage

```
rbinom2.or(n, mu1,
  mu2 = if (exchangeable) mu1 else
  stop("argument 'mu2' not specified"),
  oratio = 1, exchangeable = FALSE, tol = 0.001,
  twoCols = TRUE, colnames = if (twoCols) c("y1","y2") else
  c("00", "01", "10", "11"),
  ErrorCheck = TRUE)
dbinom2.or(mu1, mu2 = if (exchangeable) mu1 else
  stop("'mu2' not specified"),
  oratio = 1, exchangeable = FALSE, tol = 0.001,
  colnames = c("00", "01", "10", "11"), ErrorCheck = TRUE)
```

Arguments

n	number of observations. Same as in runif . The arguments mu1, mu2, oratio are recycled to this value.
mu1, mu2	The marginal probabilities. Only mu1 is needed if exchangeable = TRUE. Values should be between 0 and 1.

oratio	Odds ratio. Must be numeric and positive. The default value of unity means the responses are statistically independent.
exchangeable	Logical. If TRUE, the two marginal probabilities are constrained to be equal.
twoCols	Logical. If TRUE, then a $n \times 2$ matrix of 1s and 0s is returned. If FALSE, then a $n \times 4$ matrix of 1s and 0s is returned.
colnames	The <code>dimnames</code> argument of <code>matrix</code> is assigned <code>list(NULL, colnames)</code> .
tol	Tolerance for testing independence. Should be some small positive numerical value.
ErrorCheck	Logical. Do some error checking of the input parameters?

Details

The function `rbinom2.or` generates data coming from a bivariate binary response model. The data might be fitted with the **VGAM** family function `binom2.or`.

The function `dbinom2.or` does not really compute the density (because that does not make sense here) but rather returns the four joint probabilities.

Value

The function `rbinom2.or` returns either a 2 or 4 column matrix of 1s and 0s, depending on the argument `twoCols`.

The function `dbinom2.or` returns a 4 column matrix of joint probabilities; each row adds up to unity.

Author(s)

T. W. Yee

See Also

[binom2.or](#).

Examples

```
nn <- 1000 # Example 1
ymat <- rbinom2.or(nn, mu1 = logitlink(1, inv = TRUE),
                 oratio = exp(2), exch = TRUE)
(mytab <- table(ymat[, 1], ymat[, 2], dnn = c("Y1", "Y2")))
(myor <- mytab["0","0"] * mytab["1","1"] / (mytab["1","0"] *
mytab["0","1"]))
fit <- vglm(ymat ~ 1, binom2.or(exch = TRUE))
coef(fit, matrix = TRUE)

bdata <- data.frame(x2 = sort(runif(nn))) # Example 2
bdata <- transform(bdata,
                 mu1 = logitlink(-2 + 4 * x2, inverse = TRUE),
                 mu2 = logitlink(-1 + 3 * x2, inverse = TRUE))
dmat <- with(bdata, dbinom2.or(mu1 = mu1, mu2 = mu2,
                             oratio = exp(2)))
```

```

ymat <- with(bdata, rbinom2.or(n = nn, mu1 = mu1, mu2 = mu2,
                             oratio = exp(2)))
fit2 <- vglm(ymat ~ x2, binom2.or, data = bdata)
coef(fit2, matrix = TRUE)
## Not run:
matplot(with(bdata, x2), dmat, lty = 1:4, col = 1:4,
        main = "Joint probabilities", ylim = 0:1, type = "l",
        ylab = "Probabilities", xlab = "x2", las = 1)
legend("top", lty = 1:4, col = 1:4,
      legend = c("1 = (y1=0, y2=0)", "2 = (y1=0, y2=1)",
                "3 = (y1=1, y2=0)", "4 = (y1=1, y2=1)"))

## End(Not run)

```

binom2.or

*Bivariate Binary Regression with an Odds Ratio (Family Function)***Description**

Fits a Palmgren (bivariate odds-ratio model, or bivariate logistic regression) model to two binary responses. Actually, a bivariate logistic/probit/cloglog/cauchit model can be fitted. The odds ratio is used as a measure of dependency.

Usage

```

binom2.or(lmu = "logitlink", lmu1 = lmu, lmu2 = lmu, loratio = "loglink",
         imu1 = NULL, imu2 = NULL, ioratio = NULL, zero = "oratio",
         exchangeable = FALSE, tol = 0.001, more.robust = FALSE)

```

Arguments

lmu	Link function applied to the two marginal probabilities. See Links for more choices. See the note below.
lmu1, lmu2	Link function applied to the first and second of the two marginal probabilities.
loratio	Link function applied to the odds ratio. See Links for more choices.
imu1, imu2, ioratio	Optional initial values for the marginal probabilities and odds ratio. See CommonVGAMffArguments for more details. In general good initial values are often required so use these arguments if convergence failure occurs.
zero	Which linear/additive predictor is modelled as an intercept only? The default is for the odds ratio. A NULL means none. See CommonVGAMffArguments for more details.
exchangeable	Logical. If TRUE, the two marginal probabilities are constrained to be equal.
tol	Tolerance for testing independence. Should be some small positive numerical value.
more.robust	Logical. If TRUE then some measures are taken to compute the derivatives and working weights more robustly, i.e., in an attempt to avoid numerical problems. Currently this feature is not debugged if set TRUE.

Details

Also known informally as the *Palmgren model*, the bivariate logistic model is a full-likelihood based model defined as two logistic regressions plus $\log(\text{oratio}) = \eta_3$ where η_3 is the third linear/additive predictor relating the odds ratio to explanatory variables. Explicitly, the default model is

$$\text{logit}[P(Y_j = 1)] = \eta_j, \quad j = 1, 2$$

for the marginals, and

$$\log[P(Y_{00} = 1)P(Y_{11} = 1)/(P(Y_{01} = 1)P(Y_{10} = 1))] = \eta_3,$$

specifies the dependency between the two responses. Here, the responses equal 1 for a success and a 0 for a failure, and the odds ratio is often written $\psi = p_{00}p_{11}/(p_{10}p_{01})$. The model is fitted by maximum likelihood estimation since the full likelihood is specified. The two binary responses are independent if and only if the odds ratio is unity, or equivalently, the log odds ratio is 0. Fisher scoring is implemented.

The default models η_3 as a single parameter only, i.e., an intercept-only model, but this can be circumvented by setting `zero = NULL` in order to model the odds ratio as a function of all the explanatory variables. The function `binom2.or()` can handle other probability link functions such as `probitlink`, `clogloglink` and `cauchitlink` links as well, so is quite general. In fact, the two marginal probabilities can each have a different link function. A similar model is the *bivariate probit model* (`binom2.rho`), which is based on a standard bivariate normal distribution, but the bivariate probit model is less interpretable and flexible.

The exchangeable argument should be used when the error structure is exchangeable, e.g., with eyes or ears data.

Value

An object of class "vglmff" (see `vglmff-class`). The object is used by modelling functions such as `vglm` and `vgam`.

When fitted, the `fitted.values` slot of the object contains the four joint probabilities, labelled as $(Y_1, Y_2) = (0,0), (0,1), (1,0), (1,1)$, respectively. These estimated probabilities should be extracted with the `fitted` generic function.

Note

At present we call `binom2.or` families a *bivariate odds-ratio model*. The response should be either a 4-column matrix of counts (whose columns correspond to $(Y_1, Y_2) = (0,0), (0,1), (1,0), (1,1)$ respectively), or a two-column matrix where each column has two distinct values, or a factor with four levels. The function `rbinom2.or` may be used to generate such data. Successful convergence requires at least one case of each of the four possible outcomes.

By default, a constant odds ratio is fitted because `zero = 3`. Set `zero = NULL` if you want the odds ratio to be modelled as a function of the explanatory variables; however, numerical problems are more likely to occur.

The argument `lmu`, which is actually redundant, is used for convenience and for upward compatibility: specifying `lmu` only means the link function will be applied to `lmu1` and `lmu2`. Users who want a different link function for each of the two marginal probabilities should use the `lmu1` and `lmu2`

arguments, and the argument `lmu` is then ignored. It doesn't make sense to specify `exchangeable = TRUE` and have different link functions for the two marginal probabilities.

Regarding Yee and Dirnbock (2009), the `xij` (see [vglm.control](#)) argument enables environmental variables with different values at the two time points to be entered into an exchangeable [binom2.or](#) model. See the author's webpage for sample code.

Author(s)

Thomas W. Yee

References

- McCullagh, P. and Nelder, J. A. (1989). *Generalized Linear Models*, 2nd ed. London: Chapman & Hall.
- le Cessie, S. and van Houwelingen, J. C. (1994). Logistic regression for correlated binary data. *Applied Statistics*, **43**, 95–108.
- Palmgren, J. (1989). *Regression Models for Bivariate Binary Responses*. Technical Report no. 101, Department of Biostatistics, University of Washington, Seattle.
- Yee, T. W. and Dirnbock, T. (2009). Models for analysing species' presence/absence data at two time points. *Journal of Theoretical Biology*, **259**(4), 684–694.

See Also

[rbinom2.or](#), [binom2.rho](#), [loglinb2](#), [loglinb3](#), [zipebcom](#), [coalminers](#), [binomialff](#), [logitlink](#), [probitlink](#), [clogloglink](#), [cauchitlink](#).

Examples

```
# Fit the model in Table 6.7 in McCullagh and Nelder (1989)
coalminers <- transform(coalminers, Age = (age - 42) / 5)
fit <- vglm(cbind(nBnW, nBW, BnW, BW) ~ Age,
            binom2.or(zero = NULL), data = coalminers)
fitted(fit)
summary(fit)
coef(fit, matrix = TRUE)
c(weights(fit, type = "prior")) * fitted(fit) # Table 6.8

## Not run: with(coalminers, matplot(Age, fitted(fit), type = "l", las = 1,
                                   xlab = "(age - 42) / 5", lwd = 2))
with(coalminers, matpoints(Age, depvar(fit), col=1:4))
legend(x = -4, y = 0.5, lty = 1:4, col = 1:4, lwd = 2,
       legend = c("1 = (Breathlessness=0, Wheeze=0)",
                  "2 = (Breathlessness=0, Wheeze=1)",
                  "3 = (Breathlessness=1, Wheeze=0)",
                  "4 = (Breathlessness=1, Wheeze=1)"))

## End(Not run)

# Another model: pet ownership
## Not run: data(xs.nz, package = "VGAMdata")
```

```

# More homogeneous:
petdata <- subset(xsns.nz, ethnicity == "European" & age < 70 &
                 sex == "M")
petdata <- na.omit(petdata[, c("cat", "dog", "age")])
summary(petdata)
with(petdata, table(cat, dog)) # Can compute the odds ratio

fit <- vgam(cbind((1-cat) * (1-dog), (1-cat) * dog,
                 cat * (1-dog), cat * dog) ~ s(age, df = 5),
           binom2.or(zero = 3), data = petdata, trace = TRUE)
colSums(depvar(fit))
coef(fit, matrix = TRUE)

## End(Not run)

## Not run: # Plot the estimated probabilities
ooo <- order(with(petdata, age))
matplot(with(petdata, age)[ooo], fitted(fit)[ooo, ], type = "l",
        xlab = "Age", ylab = "Probability", main = "Pet ownership",
        ylim = c(0, max(fitted(fit))), las = 1, lwd = 1.5)
legend("topleft", col=1:4, lty = 1:4, leg = c("no cat or dog ",
        "dog only", "cat only", "cat and dog"), lwd = 1.5)
## End(Not run)

```

Binom2.rho

Bivariate Probit Model

Description

Density and random generation for a bivariate probit model. The correlation parameter ρ is the measure of dependency.

Usage

```

rbinom2.rho(n, mu1,
            mu2 = if (exchangeable) mu1 else stop("'mu2' not specified"),
            rho = 0, exchangeable = FALSE, twoCols = TRUE,
            colnames = if (twoCols) c("y1", "y2") else c("00", "01", "10", "11"),
            ErrorCheck = TRUE)
dbinom2.rho(mu1,
            mu2 = if (exchangeable) mu1 else stop("'mu2' not specified"),
            rho = 0, exchangeable = FALSE,
            colnames = c("00", "01", "10", "11"), ErrorCheck = TRUE)

```

Arguments

n number of observations. Same as in `runif`. The arguments `mu1`, `mu2`, `rho` are recycled to this value.

mu1, mu2	The marginal probabilities. Only mu1 is needed if exchangeable = TRUE. Values should be between 0 and 1.
rho	The correlation parameter. Must be numeric and lie between -1 and 1. The default value of zero means the responses are uncorrelated.
exchangeable	Logical. If TRUE, the two marginal probabilities are constrained to be equal.
twoCols	Logical. If TRUE, then a $n \times 2$ matrix of 1s and 0s is returned. If FALSE, then a $n \times 4$ matrix of 1s and 0s is returned.
colnames	The dimnames argument of <code>matrix</code> is assigned <code>list(NULL, colnames)</code> .
ErrorCheck	Logical. Do some error checking of the input parameters?

Details

The function `rbinom2.rho` generates data coming from a bivariate probit model. The data might be fitted with the **VGAM** family function `binom2.rho`.

The function `dbinom2.rho` does not really compute the density (because that does not make sense here) but rather returns the four joint probabilities.

Value

The function `rbinom2.rho` returns either a 2 or 4 column matrix of 1s and 0s, depending on the argument `twoCols`.

The function `dbinom2.rho` returns a 4 column matrix of joint probabilities; each row adds up to unity.

Author(s)

T. W. Yee

See Also

[binom2.rho](#).

Examples

```
(myrho <- rhobitlink(2, inverse = TRUE)) # Example 1
nn <- 2000
ymat <- rbinom2.rho(nn, mu1 = 0.8, rho = myrho, exch = TRUE)
(mytab <- table(ymat[, 1], ymat[, 2], dnn = c("Y1", "Y2")))
fit <- vglm(ymat ~ 1, binom2.rho(exch = TRUE))
coef(fit, matrix = TRUE)

bdata <- data.frame(x2 = sort(runif(nn))) # Example 2
bdata <- transform(bdata, mu1 = probitlink(-2+4*x2, inv = TRUE),
                  mu2 = probitlink(-1+3*x2, inv = TRUE))
dmat <- with(bdata, dbinom2.rho(mu1, mu2, myrho))
ymat <- with(bdata, rbinom2.rho(nn, mu1, mu2, myrho))
fit2 <- vglm(ymat ~ x2, binom2.rho, data = bdata)
coef(fit2, matrix = TRUE)
## Not run: matplot(with(bdata, x2), dmat, lty = 1:4, col = 1:4,
```



```

      type = "l", main = "Joint probabilities",
      ylim = 0:1, lwd = 2, ylab = "Probability")
legend(x = 0.25, y = 0.9, lty = 1:4, col = 1:4, lwd = 2,
      legend = c("1 = (y1=0, y2=0)", "2 = (y1=0, y2=1)",
      "3 = (y1=1, y2=0)", "4 = (y1=1, y2=1)"))
## End(Not run)

```

binom2.rho

Bivariate Probit Regression

Description

Fits a bivariate probit model to two binary responses. The correlation parameter ρ is the measure of dependency.

Usage

```

binom2.rho(lmu = "probitlink", lrho = "rhobitlink",
  imu1 = NULL, imu2 = NULL,
  irho = NULL, imethod = 1, zero = "rho",
  exchangeable = FALSE, grho = seq(-0.95, 0.95, by = 0.05),
  nsimEIM = NULL)
binom2.Rho(rho = 0, imu1 = NULL, imu2 = NULL,
  exchangeable = FALSE, nsimEIM = NULL)

```

Arguments

<code>lmu</code>	Link function applied to the marginal probabilities. Should be left alone.
<code>lrho</code>	Link function applied to the ρ association parameter. See Links for more choices.
<code>imu1, imu2</code>	Optional initial values for the two marginal probabilities. May be a vector.
<code>irho</code>	Optional initial value for ρ . If given, this should lie between -1 and 1 . See below for more comments.
<code>zero</code>	Specifies which linear/additive predictors are modelled as intercept-only. A NULL means none. Numerically, the ρ parameter is easiest modelled as an intercept only, hence the default. See CommonVGAMffArguments for more information.
<code>exchangeable</code>	Logical. If TRUE, the two marginal probabilities are constrained to be equal.
<code>imethod, nsimEIM, grho</code>	See CommonVGAMffArguments for more information. A value of at least 100 for <code>nsimEIM</code> is recommended; the larger the value the better.
<code>rho</code>	Numeric vector. Values are recycled to the needed length, and ought to be in range, which is $(-1, 1)$.

Details

The *bivariate probit model* was one of the earliest regression models to handle two binary responses jointly. It has a probit link for each of the two marginal probabilities, and models the association between the responses by the ρ parameter of a standard bivariate normal distribution (with zero means and unit variances). One can think of the joint probabilities being $\Phi(\eta_1, \eta_2; \rho)$ where Φ is the cumulative distribution function of a standard bivariate normal distribution.

Explicitly, the default model is

$$\text{probit}[P(Y_j = 1)] = \eta_j, \quad j = 1, 2$$

for the marginals, and

$$\text{rhobit}[\rho] = \eta_3.$$

The joint probability $P(Y_1 = 1, Y_2 = 1) = \Phi(\eta_1, \eta_2; \rho)$, and from these the other three joint probabilities are easily computed. The model is fitted by maximum likelihood estimation since the full likelihood is specified. Fisher scoring is implemented.

The default models η_3 as a single parameter only, i.e., an intercept-only model for ρ , but this can be circumvented by setting `zero = NULL` in order to model ρ as a function of all the explanatory variables.

The bivariate probit model should not be confused with a *bivariate logit model* with a probit link (see [binom2.or](#)). The latter uses the odds ratio to quantify the association. Actually, the bivariate logit model is recommended over the bivariate probit model because the odds ratio is a more natural way of measuring the association between two binary responses.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

When fitted, the `fitted.values` slot of the object contains the four joint probabilities, labelled as $(Y_1, Y_2) = (0,0), (0,1), (1,0), (1,1)$, respectively.

Note

See [binom2.or](#) about the form of input the response should have.

By default, a constant ρ is fitted because `zero = "rho"`. Set `zero = NULL` if you want the ρ parameter to be modelled as a function of the explanatory variables. The value ρ lies in the interval $(-1, 1)$, therefore a [rhobitlink](#) link is default.

Converge problems can occur. If so, assign `irho` a range of values and monitor convergence (e.g., set `trace = TRUE`). Else try `imethod`. Practical experience shows that local solutions can occur, and that `irho` needs to be quite close to the (global) solution. Also, `imu1` and `imu2` may be used.

This help file is mainly about `binom2.rho()`. `binom2.Rho()` fits a bivariate probit model with *known* ρ . The inputted ρ is saved in the `misc` slot of the fitted object, with `rho` as the component name.

In some econometrics applications (e.g., Freedman 2010, Freedman and Sekhon 2010) one response is used as an explanatory variable, e.g., a *recursive* binomial probit model. Such will not work here. Historically, the bivariate probit model was the first VGAM I ever wrote, based on Ashford and Sowden (1970). I don't think they ever thought of it either! Hence the criticisms raised go beyond the use of what was originally intended.

Author(s)

Thomas W. Yee

References

- Ashford, J. R. and Sowden, R. R. (1970). Multi-variate probit analysis. *Biometrics*, **26**, 535–546.
- Freedman, D. A. (2010). *Statistical Models and Causal Inference: a Dialogue with the Social Sciences*, Cambridge: Cambridge University Press.
- Freedman, D. A. and Sekhon, J. S. (2010). Endogeneity in probit response models. *Political Analysis*, **18**, 138–150.

See Also

[rbinom2.rho](#), [rhobitlink](#), [pbinorm](#), [binom2.or](#), [loglinb2](#), [coalminers](#), [binomialff](#), [rhobitlink](#), [fisherzlink](#).

Examples

```
coalminers <- transform(coalminers, Age = (age - 42) / 5)
fit <- vglm(cbind(nBnW, nBW, BnW, BW) ~ Age,
            binom2.rho, data = coalminers, trace = TRUE)
summary(fit)
coef(fit, matrix = TRUE)
```

binomialff

Binomial Family Function

Description

Family function for fitting generalized linear models to binomial responses

Usage

```
binomialff(link = "logitlink", multiple.responses = FALSE,
           parallel = FALSE, zero = NULL, bred = FALSE, earg.link = FALSE)
```

Arguments

- link** Link function; see [Links](#) and [CommonVGAMffArguments](#) for more information.
- multiple.responses** Multivariate response? If TRUE, then the response is interpreted as M independent binary responses, where M is the number of columns of the response matrix. In this case, the response matrix should have Q columns consisting of counts (successes), and the **weights** argument should have Q columns consisting of the number of trials (successes plus failures).
If FALSE and the response is a (2-column) matrix, then the number of successes is given in the first column, and the second column is the number of failures.

<code>parallel</code>	A logical or formula. Used only if <code>multiple.responses</code> is TRUE. This argument allows for the parallelism assumption whereby the regression coefficients for a variable is constrained to be equal over the M linear/additive predictors. If <code>parallel = TRUE</code> then the constraint is not applied to the intercepts.
<code>zero</code>	An integer-valued vector specifying which linear/additive predictors are modelled as intercepts only. The values must be from the set $\{1, 2, \dots, M\}$, where M is the number of columns of the matrix response. See CommonVGAMffArguments for more information.
<code>earg.link</code>	Details at CommonVGAMffArguments .
<code>bred</code>	Details at CommonVGAMffArguments . Setting <code>bred = TRUE</code> should work for multiple responses (<code>multiple.responses = TRUE</code>) and all VGAM link functions; it has been tested for <code>logitlink</code> only (and it gives similar results to <code>brglm</code> but not identical), and further testing is required. One result from fitting bias reduced binary regression is that finite regression coefficients occur when the data is separable (see example below). Currently <code>hdeff.vglm</code> does not work when <code>bred = TRUE</code> .

Details

This function is largely to mimic [binomial](#), however there are some differences.

When used with `cqo` and `cao`, it may be preferable to use the `clogloglink` link.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as `vglm`, `vgam`, `rrvglm`, `cqo`, and `cao`.

Warning

See the above note regarding `bred`.

The maximum likelihood estimate will not exist if the data is *completely separable* or *quasi-completely separable*. See Chapter 10 of Altman et al. (2004) for more details, and **safeBinaryRegression** and `hdeff.vglm`. Yet to do: add a `sepcheck = TRUE`, say, argument to further detect this problem and give an appropriate warning.

Note

If `multiple.responses` is FALSE (default) then the response can be of one of two formats: a factor (first level taken as failure), or a 2-column matrix (first column = successes) of counts. The argument `weights` in the modelling function can also be specified as any vector of positive values. In general, 1 means success and 0 means failure (to check, see the `y` slot of the fitted object). Note that a general vector of proportions of success is no longer accepted.

The notation M is used to denote the number of linear/additive predictors.

If `multiple.responses` is TRUE, then the matrix response can only be of one format: a matrix of 1's and 0's (1 = success).

Fisher scoring is used. This can sometimes fail to converge by oscillating between successive iterations (Ridout, 1990). See the example below.

Author(s)

Thomas W. Yee

References

McCullagh, P. and Nelder, J. A. (1989). *Generalized Linear Models*, 2nd ed. London: Chapman & Hall.

Altman, M. and Gill, J. and McDonald, M. P. (2004). *Numerical Issues in Statistical Computing for the Social Scientist*, Hoboken, NJ, USA: Wiley-Interscience.

Ridout, M. S. (1990). Non-convergence of Fisher's method of scoring—a simple example. *GLIM Newsletter*, 20(6).

See Also

[hdeff.vglm](#), [Links](#), [alogitlink](#), [asinlink](#), [N1binomial](#), [rrvglm](#), [cqo](#), [cao](#), [betabinomial](#), [posbinomial](#), [zibinomial](#), [double.expbinomial](#), [seq2binomial](#), [amlbinomial](#), [simplex](#), [binomial](#), [simulate.vlm](#), [safeBinaryRegression](#), [residualsvglm](#).

Examples

```
shunua <- hunua[sort.list(with(hunua, altitude)), ] # Sort by altitude
fit <- vglm(agaau ~ poly(altitude, 2), binomialff(link = clogloglink),
           data = shunua)
## Not run:
plot(agaau ~ jitter(altitude), shunua, ylab = "Pr(Agaau = 1)",
     main = "Presence/absence of Agathis australis", col = 4, las = 1)
with(shunua, lines(altitude, fitted(fit), col = "orange", lwd = 2))
## End(Not run)

# Fit two species simultaneously
fit2 <- vgam(cbind(agaau, kniexc) ~ s(altitude),
            binomialff(multiple.responses = TRUE), data = shunua)
## Not run:
with(shunua, matplot(altitude, fitted(fit2), type = "l",
                    main = "Two species response curves", las = 1))
## End(Not run)

# Shows that Fisher scoring can sometime fail. See Ridout (1990).
ridout <- data.frame(v = c(1000, 100, 10), r = c(4, 3, 3), n = rep(5, 3))
(ridout <- transform(ridout, logv = log(v)))
# The iterations oscillates between two local solutions:
glm.fail <- glm(r / n ~ offset(logv) + 1, weight = n,
               binomial(link = 'cloglog'), ridout, trace = TRUE)
coef(glm.fail)
# vglm()'s half-stepping ensures the MLE of -5.4007 is obtained:
vglm.ok <- vglm(cbind(r, n-r) ~ offset(logv) + 1,
               binomialff(link = clogloglink), ridout, trace = TRUE)
coef(vglm.ok)

# Separable data
```

```

set.seed(123)
threshold <- 0
bdata <- data.frame(x2 = sort(rnorm(nn <- 100)))
bdata <- transform(bdata, y1 = ifelse(x2 < threshold, 0, 1))
fit <- vglm(y1 ~ x2, binomialff(bred = TRUE),
            data = bdata, criter = "coef", trace = TRUE)
coef(fit, matrix = TRUE) # Finite!!
summary(fit)
## Not run: plot(depvar(fit) ~ x2, data = bdata, col = "blue", las = 1)
lines(fitted(fit) ~ x2, data = bdata, col = "orange")
abline(v = threshold, col = "gray", lty = "dashed")
## End(Not run)

```

Binorm

*Bivariate Normal Distribution Cumulative Distribution Function***Description**

Density, cumulative distribution function and random generation for the bivariate normal distribution distribution.

Usage

```

dbinorm(x1, x2, mean1 = 0, mean2 = 0, var1 = 1, var2 = 1, cov12 = 0,
        log = FALSE)
pbinorm(q1, q2, mean1 = 0, mean2 = 0, var1 = 1, var2 = 1, cov12 = 0)
rbinorm(n, mean1 = 0, mean2 = 0, var1 = 1, var2 = 1, cov12 = 0)
pnorm2(x1, x2, mean1 = 0, mean2 = 0, var1 = 1, var2 = 1, cov12 = 0)

```

Arguments

`x1, x2, q1, q2` vector of quantiles.
`mean1, mean2, var1, var2, cov12` vector of means, variances and the covariance.
`n` number of observations. Same as `rnorm`.
`log` Logical. If `log = TRUE` then the logarithm of the density is returned.

Details

The default arguments correspond to the standard bivariate normal distribution with correlation parameter $\rho = 0$. That is, two independent standard normal distributions. Let `sd1` (say) be `sqrt(var1)` and written σ_1 , etc. Then the general formula for the correlation coefficient is $\rho = cov/(\sigma_1\sigma_2)$ where `cov` is argument `cov12`. Thus if arguments `var1` and `var2` are left alone then `cov12` can be inputted with ρ .

One can think of this function as an extension of `pnorm` to two dimensions, however note that the argument names have been changed for **VGAM** 0.9-1 onwards.

Value

dbinorm gives the density, pbinorm gives the cumulative distribution function, rbinorm generates random deviates (n by 2 matrix).

Warning

Being based on an approximation, the results of pbinorm() may be negative! Also, pnorm2() should be withdrawn soon; use pbinorm() instead because it is identical.

Note

For rbinorm(), if the i th variance-covariance matrix is not positive-definite then the i th row is all NAs.

References

pbinorm() is based on Donnelly (1973), the code was translated from FORTRAN to ratfor using struct, and then from ratfor to C manually. The function was originally called bivnor, and TWY only wrote a wrapper function.

Donnelly, T. G. (1973). Algorithm 462: Bivariate Normal Distribution. *Communications of the ACM*, **16**, 638.

See Also

[pnorm](#), [binormal](#), [uninormal](#).

Examples

```
yvec <- c(-5, -1.96, 0, 1.96, 5)
ymat <- expand.grid(yvec, yvec)
cbind(ymat, pbinorm(ymat[, 1], ymat[, 2]))

## Not run: rhovec <- seq(-0.95, 0.95, by = 0.01)
plot(rhovec, pbinorm(0, 0, cov12 = rhovec),
     xlab = expression(rho), lwd = 2,
     type = "l", col = "blue", las = 1)
abline(v = 0, h = 0.25, col = "gray", lty = "dashed")
## End(Not run)
```

binormal

Bivariate Normal Distribution Family Function

Description

Maximum likelihood estimation of the five parameters of a bivariate normal distribution.

Usage

```
binormal(lmean1 = "identitylink", lmean2 = "identitylink",
        lsd1 = "loglink", lsd2 = "loglink",
        lrho = "rhobitlink",
        imean1 = NULL, imean2 = NULL,
        isd1 = NULL, isd2 = NULL,
        irho = NULL, imethod = 1,
        eq.mean = FALSE, eq.sd = FALSE,
        zero = c("sd", "rho"), rho.arg = NA)
```

Arguments

`lmean1`, `lmean2`, `lsd1`, `lsd2`, `lrho`
 Link functions applied to the means, standard deviations and rho parameters. See [Links](#) for more choices. Being positive quantities, a log link is the default for the standard deviations.

`imean1`, `imean2`, `isd1`, `isd2`, `irho`, `imethod`, `zero`
 See [CommonVGAMffArguments](#) for more information.

`eq.mean`, `eq.sd` Logical or formula. Constrains the means or the standard deviations to be equal.

`rho.arg` If ρ is known then this argument may be assigned the (scalar) value lying in $(-1, 1)$. The default is to estimate that parameter so that $M = 5$. If known, then other arguments such as `lrho` and `irho` are ignored, and "rho" is removed from zero.

Details

For the bivariate normal distribution, this fits a linear model (LM) to the means, and by default, the other parameters are intercept-only. The response should be a two-column matrix. The correlation parameter is rho, which lies between -1 and 1 (thus the [rhobitlink](#) link is a reasonable choice). The fitted means are returned as the fitted values, which is in the form of a two-column matrix. Fisher scoring is implemented.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Warning

This function may be renamed to `normal2()` or something like that at a later date.

Note

If both equal means and equal standard deviations are desired then use something like `constraints = list("(Intercept)" = matrix(c(1,1,0,0,0, 0,0,1,1,0, 0,0,0,0,1), 5, 3))` and maybe `zero = NULL` etc.

Author(s)

T. W. Yee

See Also[uninormal](#), [trinormal](#), [pbinorm](#), [bistudentt](#), [rhobitlink](#).**Examples**

```

set.seed(123); nn <- 1000
bdata <- data.frame(x2 = runif(nn), x3 = runif(nn))
bdata <- transform(bdata, y1 = rnorm(nn, 1 + 2 * x2),
                  y2 = rnorm(nn, 3 + 4 * x2))
fit1 <- vglm(cbind(y1, y2) ~ x2,
            binormal(eq.sd = TRUE), bdata, trace = TRUE)
coef(fit1, matrix = TRUE)
constraints(fit1)
summary(fit1)

# Estimated P(Y1 <= y1, Y2 <= y2) under the fitted model
var1 <- loglink(2 * predict(fit1)[, "loglink(sd1)"], inv = TRUE)
var2 <- loglink(2 * predict(fit1)[, "loglink(sd2)"], inv = TRUE)
cov12 <- rhobitlink(predict(fit1)[, "rhobitlink(rho)"], inv = TRUE)
head(with(bdata, pbinorm(y1, y2,
                        mean1 = predict(fit1)[, "mean1"],
                        mean2 = predict(fit1)[, "mean2"],
                        var1 = var1, var2 = var2, cov12 = cov12)))

```

binormalcop

*Gaussian Copula (Bivariate) Family Function***Description**

Estimate the correlation parameter of the (bivariate) Gaussian copula distribution by maximum likelihood estimation.

Usage

```
binormalcop(lrho = "rhobitlink", irho = NULL,
           imethod = 1, parallel = FALSE, zero = NULL)
```

Arguments

`lrho`, `irho`, `imethod` Details at [CommonVGAMffArguments](#). See [Links](#) for more link function choices.

`parallel`, `zero` Details at [CommonVGAMffArguments](#). If `parallel = TRUE` then the constraint is applied to the intercept too.

Details

The cumulative distribution function is

$$P(Y_1 \leq y_1, Y_2 \leq y_2) = \Phi_2(\Phi^{-1}(y_1), \Phi^{-1}(y_2); \rho)$$

for $-1 < \rho < 1$, Φ_2 is the cumulative distribution function of a standard bivariate normal (see [pbinorm](#)), and Φ is the cumulative distribution function of a standard univariate normal (see [pnorm](#)).

The support of the function is the interior of the unit square; however, values of 0 and/or 1 are not allowed. The marginal distributions are the standard uniform distributions. When $\rho = 0$ the random variables are independent.

This **VGAM** family function can handle multiple responses, for example, a six-column matrix where the first 2 columns is the first out of three responses, the next 2 columns being the next response, etc.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Note

The response matrix must have a multiple of two-columns. Currently, the fitted value is a matrix with the same number of columns and values equal to 0.5. This is because each marginal distribution corresponds to a standard uniform distribution.

This **VGAM** family function is fragile; each response must be in the interior of the unit square. Setting `crit = "coef"` is sometimes a good idea because inaccuracies in [pbinorm](#) might mean unnecessary half-stepping will occur near the solution.

Author(s)

T. W. Yee

References

Schepsmeier, U. and Stober, J. (2014). Derivatives and Fisher information of bivariate copulas. *Statistical Papers* **55**, 525–542.

See Also

[rbinormcop](#), [rhobitlink](#), [pnorm](#), [kendall.tau](#).

Examples

```
nn <- 1000
ymat <- rbinormcop(nn, rho = rhobitlink(-0.9, inverse = TRUE))
bdata <- data.frame(y1 = ymat[, 1], y2 = ymat[, 2],
                  y3 = ymat[, 1], y4 = ymat[, 2],
                  x2 = runif(nn))
summary(bdata)
```

```
## Not run: plot(yamat, col = "blue")
fit1 <- # 2 responses, e.g., (y1,y2) is the 1st
  vglm(cbind(y1, y2, y3, y4) ~ 1, fam = binormalcop,
        crit = "coef", # Sometimes a good idea
        data = bdata, trace = TRUE)
coef(fit1, matrix = TRUE)
Coef(fit1)
head(fitted(fit1))
summary(fit1)

# Another example; rho is a linear function of x2
bdata <- transform(bdata, rho = -0.5 + x2)
ymat <- rbinormcop(n = nn, rho = with(bdata, rho))
bdata <- transform(bdata, y5 = ymat[, 1], y6 = ymat[, 2])
fit2 <- vgam(cbind(y5, y6) ~ s(x2), data = bdata,
             binormalcop(lrho = "identitylink"), trace = TRUE)
## Not run: plot(fit2, lcol = "blue", scol = "orange", se = TRUE)
```

Binormcop

*Gaussian Copula (Bivariate) Distribution***Description**

Density, distribution function, and random generation for the (one parameter) bivariate Gaussian copula distribution.

Usage

```
dbinormcop(x1, x2, rho = 0, log = FALSE)
pbinormcop(q1, q2, rho = 0)
rbinormcop(n, rho = 0)
```

Arguments

x1, x2, q1, q2	vector of quantiles. The x1 and x2 should be in the interval (0, 1). Ditto for q1 and q2.
n	number of observations. Same as rnorm .
rho	the correlation parameter. Should be in the interval (-1, 1).
log	Logical. If TRUE then the logarithm is returned.

Details

See [binormalcop](#), the **VGAM** family functions for estimating the parameter by maximum likelihood estimation, for the formula of the cumulative distribution function and other details.

Value

`dbinormcop` gives the density, `pbinormcop` gives the distribution function, and `rbinormcop` generates random deviates (a two-column matrix).

Note

Yettodo: allow x_1 and/or x_2 to have values 1, and to allow any values for x_1 and/or x_2 to be outside the unit square.

Author(s)

T. W. Yee

See Also

[binormalcop](#), [binormal](#).

Examples

```
## Not run: edge <- 0.01 # A small positive value
N <- 101; x <- seq(edge, 1.0 - edge, len = N); Rho <- 0.7
ox <- expand.grid(x, x)
zedd <- dbinormcop(ox[, 1], ox[, 2], rho = Rho, log = TRUE)
contour(x, x, matrix(zedd, N, N), col = "blue", labcex = 1.5)
zedd <- pbinormcop(ox[, 1], ox[, 2], rho = Rho)
contour(x, x, matrix(zedd, N, N), col = "blue", labcex = 1.5)

## End(Not run)
```

Biplackett

Plackett's Bivariate Copula

Description

Density, distribution function, and random generation for the (one parameter) bivariate Plackett copula.

Usage

```
dbiplackcop(x1, x2, oratio, log = FALSE)
pbiplackcop(q1, q2, oratio)
rbiplackcop(n, oratio)
```

Arguments

x_1, x_2, q_1, q_2	vector of quantiles.
n	number of observations. Same as in runif .
$oratio$	the positive odds ratio ψ .
log	Logical. If TRUE then the logarithm is returned.

Details

See [biplackettcop](#), the **VGAM** family functions for estimating the parameter by maximum likelihood estimation, for the formula of the cumulative distribution function and other details.

Value

`dbiplackcop` gives the density, `pbiplackcop` gives the distribution function, and `rbiplackcop` generates random deviates (a two-column matrix).

Author(s)

T. W. Yee

References

Mardia, K. V. (1967). Some contributions to contingency-type distributions. *Biometrika*, **54**, 235–249.

See Also

[biplackettcop](#), [bifrankcop](#).

Examples

```
## Not run: N <- 101; oratio <- exp(1)
x <- seq(0.0, 1.0, len = N)
ox <- expand.grid(x, x)
zedd <- dbiplackcop(ox[, 1], ox[, 2], oratio = oratio)
contour(x, x, matrix(zedd, N, N), col = "blue")
zedd <- pbiplackcop(ox[, 1], ox[, 2], oratio = oratio)
contour(x, x, matrix(zedd, N, N), col = "blue")

plot(rr <- rbiplackcop(n = 3000, oratio = oratio))
par(mfrow = c(1, 2))
hist(rr[, 1]) # Should be uniform
hist(rr[, 2]) # Should be uniform

## End(Not run)
```

biplackettcop

Plackett's Bivariate Copula Family Function

Description

Estimate the association parameter of Plackett's bivariate distribution (copula) by maximum likelihood estimation.

Usage

```
biplackettcop(link = "loglink", ioratio = NULL, imethod = 1,
              nsimEIM = 200)
```

Arguments

link Link function applied to the (positive) odds ratio ψ . See [Links](#) for more choices and information.

ioratio Numeric. Optional initial value for ψ . If a convergence failure occurs try assigning a value or a different value.

imethod, nsimEIM See [CommonVGAMffArguments](#).

Details

The defining equation is

$$\psi = H \times (1 - y_1 - y_2 + H) / ((y_1 - H) \times (y_2 - H))$$

where $P(Y_1 \leq y_1, Y_2 \leq y_2) = H_\psi(y_1, y_2)$ is the cumulative distribution function. The density function is $h_\psi(y_1, y_2) =$

$$\psi[1 + (\psi - 1)(y_1 + y_2 - 2y_1y_2)] / ([1 + (\psi - 1)(y_1 + y_2)]^2 - 4\psi(\psi - 1)y_1y_2)^{3/2}$$

for $\psi > 0$. Some writers call ψ the *cross product ratio* but it is called the *odds ratio* here. The support of the function is the unit square. The marginal distributions here are the standard uniform although it is commonly generalized to other distributions.

If $\psi = 1$ then $h_\psi(y_1, y_2) = y_1y_2$, i.e., independence. As the odds ratio tends to infinity one has $y_1 = y_2$. As the odds ratio tends to 0 one has $y_2 = 1 - y_1$.

Fisher scoring is implemented using [rbiplackcop](#). Convergence is often quite slow.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Note

The response must be a two-column matrix. Currently, the fitted value is a 2-column matrix with 0.5 values because the marginal distributions correspond to a standard uniform distribution.

Author(s)

T. W. Yee

References

Plackett, R. L. (1965). A class of bivariate distributions. *Journal of the American Statistical Association*, **60**, 516–522.

See Also

[rbiplackcop](#), [bifrankcop](#).

Examples

```
## Not run:
ymat <- rbiplackcop(n = 2000, oratio = exp(2))
plot(ymat, col = "blue")
fit <- vglm(ymat ~ 1, fam = biplackettcop, trace = TRUE)
coef(fit, matrix = TRUE)
Coef(fit)
vcov(fit)
head(fitted(fit))
summary(fit)

## End(Not run)
```

biplot-methods

Biplot of Constrained Regression Models

Description

biplot is a generic function applied to RR-VGLMs and QRR-VGLMs etc. These apply to rank-1 and rank-2 models of these only. For RR-VGLMs these plot the second latent variable scores against the first latent variable scores.

Methods

x The object from which the latent variables are extracted and/or plotted.

Note

See [lvplot](#) which is very much related to biplots.

Bisa

The Birnbaum-Saunders Distribution

Description

Density, distribution function, and random generation for the Birnbaum-Saunders distribution.

Usage

```
dbisa(x, scale = 1, shape, log = FALSE)
pbisa(q, scale = 1, shape, lower.tail = TRUE, log.p = FALSE)
qbisa(p, scale = 1, shape, lower.tail = TRUE, log.p = FALSE)
rbisa(n, scale = 1, shape)
```

Arguments

<code>x, q</code>	vector of quantiles.
<code>p</code>	vector of probabilities.
<code>n</code>	Same as in runif .
<code>scale, shape</code>	the (positive) scale and shape parameters.
<code>log</code>	Logical. If TRUE then the logarithm of the density is returned.
<code>lower.tail, log.p</code>	Same meaning as in pnorm or qnorm .

Details

The Birnbaum-Saunders distribution is a distribution which is used in survival analysis. See [bisa](#), the **VGAM** family function for estimating the parameters, for more details.

Value

`dbisa` gives the density, `pbisa` gives the distribution function, and `qbisa` gives the quantile function, and `rbisa` generates random deviates.

Author(s)

T. W. Yee and Kai Huang

See Also

[bisa](#).

Examples

```
## Not run:
x <- seq(0, 6, len = 400)
plot(x, dbisa(x, shape = 1), type = "l", col = "blue",
      ylab = "Density", lwd = 2, ylim = c(0,1.3), lty = 3,
      main = "X ~ Birnbaum-Saunders(shape, scale = 1)")
lines(x, dbisa(x, shape = 2), col = "orange", lty = 2, lwd = 2)
lines(x, dbisa(x, shape = 0.5), col = "green", lty = 1, lwd = 2)
legend(x = 3, y = 0.9, legend = paste("shape = ", c(0.5, 1, 2)),
       col = c("green", "blue", "orange"), lty = 1:3, lwd = 2)

shape <- 1; x <- seq(0.0, 4, len = 401)
plot(x, dbisa(x, shape = shape), type = "l", col = "blue",
      main = "Blue is density, orange is the CDF", las = 1,
      sub = "Red lines are the 10,20,...,90 percentiles",
      ylab = "", ylim = 0:1)
abline(h = 0, col = "blue", lty = 2)
lines(x, pbisa(x, shape = shape), col = "orange")
probs <- seq(0.1, 0.9, by = 0.1)
Q <- qbisa(probs, shape = shape)
lines(Q, dbisa(Q, shape = shape), col = "red", lty = 3, type = "h")
```



```
pbisa(Q, shape = shape) - probs # Should be all zero
abline(h = probs, col = "red", lty = 3)
lines(Q, pbisa(Q, shape = shape), col = "red", lty = 3, type = "h")

## End(Not run)
```

bisa

*Birnbaum-Saunders Regression Family Function***Description**

Estimates the shape and scale parameters of the Birnbaum-Saunders distribution by maximum likelihood estimation.

Usage

```
bisa(lscale = "loglink", lshape = "loglink", iscale = 1,
     ishape = NULL, imethod = 1, zero = "shape", nowarning = FALSE)
```

Arguments

nowarning	Logical. Suppress a warning? Ignored for VGAM 0.9-7 and higher.
lscale, lshape	Parameter link functions applied to the shape and scale parameters (a and b below). See Links for more choices. A log link is the default for both because they are positive.
iscale, ishape	Initial values for a and b . A NULL means an initial value is chosen internally using imethod.
imethod	An integer with value 1 or 2 or 3 which specifies the initialization method. If failure to converge occurs try the other value, or else specify a value for ishape and/or iscale.
zero	Specifies which linear/additive predictor is modelled as intercept-only. If used, choose one value from the set {1,2}. See CommonVGAMffArguments for more details.

Details

The (two-parameter) Birnbaum-Saunders distribution has a cumulative distribution function that can be written as

$$F(y; a, b) = \Phi[\xi(y/b)/a]$$

where $\Phi(\cdot)$ is the cumulative distribution function of a standard normal (see [pnorm](#)), $\xi(t) = \sqrt{t} - 1/\sqrt{t}$, $y > 0$, $a > 0$ is the shape parameter, $b > 0$ is the scale parameter. The mean of Y (which is the fitted value) is $b(1 + a^2/2)$. and the variance is $a^2b^2(1 + \frac{5}{4}a^2)$. By default, $\eta_1 = \log(a)$ and $\eta_2 = \log(b)$ for this family function.

Note that a and b are orthogonal, i.e., the Fisher information matrix is diagonal. This family function implements Fisher scoring, and it is unnecessary to compute any integrals numerically.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Author(s)

T. W. Yee

References

Lemonte, A. J. and Cribari-Neto, F. and Vasconcellos, K. L. P. (2007). Improved statistical inference for the two-parameter Birnbaum-Saunders distribution. *Computational Statistics & Data Analysis*, **51**, 4656–4681.

Birnbaum, Z. W. and Saunders, S. C. (1969). A new family of life distributions. *Journal of Applied Probability*, **6**, 319–327.

Birnbaum, Z. W. and Saunders, S. C. (1969). Estimation for a family of life distributions with applications to fatigue. *Journal of Applied Probability*, **6**, 328–347.

Engelhardt, M. and Bain, L. J. and Wright, F. T. (1981). Inferences on the parameters of the Birnbaum-Saunders fatigue life distribution based on maximum likelihood estimation. *Technometrics*, **23**, 251–256.

Johnson, N. L. and Kotz, S. and Balakrishnan, N. (1995). *Continuous Univariate Distributions*, 2nd edition, Volume 2, New York: Wiley.

See Also

[pbisa](#), [inv.gaussianff](#), [CommonVGAMffArguments](#).

Examples

```
bdata1 <- data.frame(x2 = runif(nn <- 1000))
bdata1 <- transform(bdata1, shape = exp(-0.5 + x2),
                  scale = exp(1.5))
bdata1 <- transform(bdata1, y = rbisa(nn, scale, shape))
fit1 <- vglm(y ~ x2, bisa(zero = 1), data = bdata1, trace = TRUE)
coef(fit1, matrix = TRUE)

## Not run:
bdata2 <- data.frame(shape = exp(-0.5), scale = exp(0.5))
bdata2 <- transform(bdata2, y = rbisa(nn, scale, shape))
fit <- vglm(y ~ 1, bisa, data = bdata2, trace = TRUE)
with(bdata2, hist(y, prob = TRUE, ylim = c(0, 0.5),
                col = "lightblue"))
coef(fit, matrix = TRUE)
with(bdata2, mean(y))
head(fitted(fit))
x <- with(bdata2, seq(0, max(y), len = 200))
lines(dbisa(x, Coef(fit)[1], Coef(fit)[2]) ~ x, data = bdata2,
      col = "orange", lwd = 2)
## End(Not run)
```

Bistudentt*Bivariate Student-t Distribution Density Function*

Description

Density for the bivariate Student-t distribution.

Usage

```
dbistudentt(x1, x2, df, rho = 0, log = FALSE)
```

Arguments

x1, x2	vector of quantiles.
df, rho	vector of degrees of freedom and correlation parameter. For df, a value Inf is currently not working.
log	Logical. If log = TRUE then the logarithm of the density is returned.

Details

One can think of this function as an extension of [dt](#) to two dimensions. See [bistudentt](#) for more information.

Value

dbistudentt gives the density.

See Also

[bistudentt](#), [dt](#).

Examples

```
## Not run: N <- 101; x <- seq(-4, 4, len = N); Rho <- 0.7
mydf <- 10; ox <- expand.grid(x, x)
zedd <- dbistudentt(ox[, 1], ox[, 2], df = mydf,
                    rho = Rho, log = TRUE)
contour(x, x, matrix(zedd, N, N), col = "blue", labcex = 1.5)

## End(Not run)
```

bistudentt

*Bivariate Student-t Family Function***Description**

Estimate the degrees of freedom and correlation parameters of the (bivariate) Student-t distribution by maximum likelihood estimation.

Usage

```
bistudentt(ldf = "logloglink", lrho = "rhobitlink",
            idf = NULL, irho = NULL, imethod = 1,
            parallel = FALSE, zero = "rho")
```

Arguments

ldf, lrho, idf, irho, imethod

Details at [CommonVGAMffArguments](#). See [Links](#) for more link function choices.

parallel, zero Details at [CommonVGAMffArguments](#).

Details

The density function is

$$f(y_1, y_2; \nu, \rho) = \frac{1}{2\pi\sqrt{1-\rho^2}} (1 + (y_1^2 + y_2^2 - 2\rho y_1 y_2)/(\nu(1-\rho^2)))^{-(\nu+2)/2}$$

for $-1 < \rho < 1$, and real y_1 and y_2 .

This **VGAM** family function can handle multiple responses, for example, a six-column matrix where the first 2 columns is the first out of three responses, the next 2 columns being the next response, etc.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Warning

The working weight matrices have not been fully checked.

Note

The response matrix must have a multiple of two-columns. Currently, the fitted value is a matrix with the same number of columns and values equal to 0.0.

Author(s)

T. W. Yee, with help from Thibault Vatter.

References

Schepsmeier, U. and Stober, J. (2014). Derivatives and Fisher information of bivariate copulas. *Statistical Papers* **55**, 525–542.

See Also

[dbistudentt](#), [binormal](#), [pt](#).

Examples

```
nn <- 1000
mydof <- logloglink(1, inverse = TRUE)
ymat <- cbind(rt(nn, df = mydof), rt(nn, df = mydof))
bdata <- data.frame(y1 = ymat[, 1], y2 = ymat[, 2],
                   y3 = ymat[, 1], y4 = ymat[, 2],
                   x2 = runif(nn))
summary(bdata)
## Not run: plot(ymat, col = "blue")
fit1 <- # 2 responses, e.g., (y1,y2) is the 1st
  vglm(cbind(y1, y2, y3, y4) ~ 1,
       bistudentt, # crit = "coef", # Sometimes a good idea
       data = bdata, trace = TRUE)
coef(fit1, matrix = TRUE)
Coef(fit1)
head(fitted(fit1))
summary(fit1)
```

bmi.nz

Body Mass Index of New Zealand Adults Data

Description

The body mass indexes and ages from an approximate random sample of 700 New Zealand adults.

Usage

```
data(bmi.nz)
```

Format

A data frame with 700 observations on the following 2 variables.

age a numeric vector; their age (years).

BMI a numeric vector; their body mass indexes, which is their weight divided by the square of their height (kg / m^2).

Details

They are a random sample from the Fletcher Challenge/Auckland Heart and Health survey conducted in the early 1990s.

There are some outliers in the data set.

A variable gender would be useful, and may be added later.

Source

Formerly the Clinical Trials Research Unit, University of Auckland, New Zealand.

References

MacMahon, S., Norton, R., Jackson, R., Mackie, M. J., Cheng, A., Vander Hoorn, S., Milne, A., McCulloch, A. (1995) Fletcher Challenge-University of Auckland Heart & Health Study: design and baseline findings. *New Zealand Medical Journal*, **108**, 499–502.

Examples

```
## Not run: with(bmi.nz, plot(age, BMI, col = "blue"))
fit <- vgam(BMI ~ s(age, df = c(2, 4, 2)), lms.yjn,
           data = bmi.nz, trace = TRUE)
qtplot(fit, pcol = "blue", tcol = "brown", lcol = "brown")
## End(Not run)
```

borel.tanner

Borel-Tanner Distribution Family Function

Description

Estimates the parameter of a Borel-Tanner distribution by maximum likelihood estimation.

Usage

```
borel.tanner(Qsize = 1, link = "logitlink", imethod = 1)
```

Arguments

Qsize	A positive integer. It is called Q below and is the initial queue size. The default value $Q = 1$ corresponds to the Borel distribution.
link	Link function for the parameter; see Links for more choices and for general information.
imethod	See CommonVGAMffArguments . Valid values are 1, 2, 3 or 4.

Details

The Borel-Tanner distribution (Tanner, 1953) describes the distribution of the total number of customers served before a queue vanishes given a single queue with random arrival times of customers (at a constant rate r per unit time, and each customer taking a constant time b to be served). Initially the queue has Q people and the first one starts to be served. The two parameters appear in the density only in the form of the product rb , therefore we use $a = rb$, say, to denote the single parameter to be estimated. The density function is

$$f(y; a) = \frac{Q}{(y - Q)!} y^{y-Q-1} a^{y-Q} \exp(-ay)$$

where $y = Q, Q + 1, Q + 2, \dots$. The case $Q = 1$ corresponds to the *Borel* distribution (Borel, 1942). For the $Q = 1$ case it is necessary for $0 < a < 1$ for the distribution to be proper. The Borel distribution is a basic Lagrangian distribution of the first kind. The Borel-Tanner distribution is an Q -fold convolution of the Borel distribution.

The mean is $Q/(1 - a)$ (returned as the fitted values) and the variance is $Qa/(1 - a)^3$. The distribution has a very long tail unless a is small. Fisher scoring is implemented.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Author(s)

T. W. Yee

References

- Tanner, J. C. (1953). A problem of interference between two queues. *Biometrika*, **40**, 58–69.
- Borel, E. (1942). Sur l'emploi du theoreme de Bernoulli pour faciliter le calcul d'une infinite de coefficients. Application au probleme de l'attente a un guichet. *Comptes Rendus, Academie des Sciences, Paris, Series A*, **214**, 452–456.
- Johnson N. L., Kemp, A. W. and Kotz S. (2005). *Univariate Discrete Distributions*, 3rd edition, p.328. Hoboken, New Jersey: Wiley.
- Consul, P. C. and Famoye, F. (2006). *Lagrangian Probability Distributions*, Boston, MA, USA: Birkhauser.

See Also

[rbort](#), [poissonff](#), [felix](#).

Examples

```
bdata <- data.frame(y = rbort(n <- 200))
fit <- vglm(y ~ 1, borel.tanner, bdata, trace = TRUE, crit = "c")
coef(fit, matrix = TRUE)
Coef(fit)
summary(fit)
```

Description

Density and random generation for the Borel-Tanner distribution.

Usage

```
dbort(x, Qsize = 1, a = 0.5, log = FALSE)
rbort(n, Qsize = 1, a = 0.5)
```

Arguments

x	vector of quantiles.
n	number of observations. Must be a positive integer of length 1.
Qsize, a	See borel.tanner .
log	Logical. If log = TRUE then the logarithm of the density is returned.

Details

See [borel.tanner](#), the **VGAM** family function for estimating the parameter, for the formula of the probability density function and other details.

Value

dbort gives the density, rbort generates random deviates.

Warning

Looping is used for [rbort](#), therefore values of a close to 1 will result in long (or infinite!) computational times. The default value of a is subjective.

Author(s)

T. W. Yee

See Also

[borel.tanner](#).

Examples

```
## Not run: qsize <- 1; a <- 0.5; x <- qsize:(qsize+10)
plot(x, dbort(x, qsize, a), type = "h", las = 1, col = "blue",
      ylab = paste("fbort(qsize=", qsize, ", a=", a, ")"),
      log = "y", main = "Borel-Tanner density function")
## End(Not run)
```

 Brat

Inputting Data to fit a Bradley Terry Model

Description

Takes in a square matrix of counts and outputs them in a form that is accessible to the [brat](#) and [bratt](#) family functions.

Usage

```
Brat(mat, ties = 0 * mat, string = c(">", "=="), whitespace = FALSE)
```

Arguments

mat	Matrix of counts, which is considered M by M in dimension when there are ties, and $M + 1$ by $M + 1$ when there are no ties. The rows are winners and the columns are losers, e.g., the 2-1 element is now many times Competitor 2 has beaten Competitor 1. The matrices are best labelled with the competitors' names.
ties	Matrix of counts. This should be the same dimension as mat. By default, there are no ties. The matrix must be symmetric, and the diagonal should contain NAs.
string	Character. The matrices are labelled with the first value of the descriptor, e.g., "NZ > Oz" 'means' NZ beats Australia in rugby. Suggested alternatives include " beats " or " wins against ". The second value is used to handle ties.
whitespace	Logical. If TRUE then a white space is added before and after string; it generally enhances readability. See CommonVGAMffArguments for some similar-type information.

Details

In the **VGAM** package it is necessary for each matrix to be represented as a single row of data by [brat](#) and [bratt](#). Hence the non-diagonal elements of the $M + 1$ by $M + 1$ matrix are concatenated into $M(M + 1)$ values (no ties), while if there are ties, the non-diagonal elements of the M by M matrix are concatenated into $M(M - 1)$ values.

Value

A matrix with 1 row and either $M(M + 1)$ or $M(M - 1)$ columns.

Note

This is a data preprocessing function for [brat](#) and [bratt](#).

Yet to do: merge InverseBrat into brat.

Author(s)

T. W. Yee

References

Agresti, A. (2013). *Categorical Data Analysis*, 3rd ed. Hoboken, NJ, USA: Wiley.

See Also

[brat](#), [bratt](#), [InverseBrat](#).

Examples

```
journal <- c("Biometrika", "Comm Statist", "JASA", "JRSS-B")
mat <- matrix(c( NA, 33, 320, 284, 730, NA, 813, 276,
               498, 68, NA, 325, 221, 17, 142, NA), 4, 4)
dimnames(mat) <- list(winner = journal, loser = journal)
Brat(mat) # Less readable
Brat(mat, whitespace = TRUE) # More readable
vglm(Brat(mat, whitespace = TRUE) ~ 1, brat, trace = TRUE)
```

brat

Bradley Terry Model

Description

Fits a Bradley Terry model (intercept-only model) by maximum likelihood estimation.

Usage

```
brat(refgp = "last", refvalue = 1, ialpha = 1)
```

Arguments

refgp	Integer whose value must be from the set $\{1, \dots, M + 1\}$, where there are $M + 1$ competitors. The default value indicates the last competitor is used—but don't input a character string, in general.
refvalue	Numeric. A positive value for the reference group.
ialpha	Initial values for the α s. These are recycled to the appropriate length.

Details

The Bradley Terry model involves $M + 1$ competitors who either win or lose against each other (no draws/ties allowed in this implementation—see [bratt](#) if there are ties). The probability that Competitor i beats Competitor j is $\alpha_i / (\alpha_i + \alpha_j)$, where all the α s are positive. Loosely, the α s can be thought of as the competitors' 'abilities'. For identifiability, one of the α_i is set to a known value `refvalue`, e.g., 1. By default, this function chooses the last competitor to have this reference value. The data can be represented in the form of a $M + 1$ by $M + 1$ matrix of counts, where winners are the rows and losers are the columns. However, this is not the way the data should be inputted (see below).

Excluding the reference value/group, this function chooses $\log(\alpha_j)$ as the M linear predictors. The log link ensures that the α s are positive.

The Bradley Terry model can be fitted by logistic regression, but this approach is not taken here. The Bradley Terry model can be fitted with covariates, e.g., a home advantage variable, but unfortunately, this lies outside the VGLM theoretical framework and therefore cannot be handled with this code.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#).

Warning

Presently, the residuals are wrong, and the prior weights are not handled correctly. Ideally, the total number of counts should be the prior weights, after the response has been converted to proportions. This would make it similar to family functions such as [multinomial](#) and [binomialff](#).

Note

The function [Brat](#) is useful for coercing a $M + 1$ by $M + 1$ matrix of counts into a one-row matrix suitable for [brat](#). Diagonal elements are skipped, and the usual S order of `c(a.matrix)` of elements is used. There should be no missing values apart from the diagonal elements of the square matrix. The matrix should have winners as the rows, and losers as the columns. In general, the response should be a 1-row matrix with $M(M + 1)$ columns.

Only an intercept model is recommended with [brat](#). It doesn't make sense really to include covariates because of the limited VGLM framework.

Notationally, note that the **VGAM** family function [brat](#) has $M + 1$ contestants, while [bratt](#) has M contestants.

Author(s)

T. W. Yee

References

Agresti, A. (2013). *Categorical Data Analysis*, 3rd ed. Hoboken, NJ, USA: Wiley.

Stigler, S. (1994). Citation patterns in the journals of statistics and probability. *Statistical Science*, **9**, 94–108.

The **BradleyTerry2** package has more comprehensive capabilities than this function.

See Also

[bratt](#), [Brat](#), [multinomial](#), [binomialff](#).

Examples

```
# Citation statistics: being cited is a 'win'; citing is a 'loss'
journal <- c("Biometrika", "Comm.Statist", "JASA", "JRSS-B")
mat <- matrix(c( NA, 33, 320, 284,
                730, NA, 813, 276,
                498, 68, NA, 325,
                221, 17, 142, NA), 4, 4)
dimnames(mat) <- list(winner = journal, loser = journal)
fit <- vglm(Brat(mat) ~ 1, brat(refgp = 1), trace = TRUE)
fit <- vglm(Brat(mat) ~ 1, brat(refgp = 1), trace = TRUE, crit = "coef")
summary(fit)
c(0, coef(fit)) # Log-abilities (in order of "journal")
c(1, Coef(fit)) # Abilities (in order of "journal")
fitted(fit)     # Probabilities of winning in awkward form
(check <- InverseBrat(fitted(fit))) # Probabilities of winning
check + t(check) # Should be 1's in the off-diagonals
```

bratt

*Bradley Terry Model With Ties***Description**

Fits a Bradley Terry model with ties (intercept-only model) by maximum likelihood estimation.

Usage

```
bratt(refgp = "last", refvalue = 1, ialpha = 1, i0 = 0.01)
```

Arguments

refgp	Integer whose value must be from the set $\{1, \dots, M\}$, where there are M competitors. The default value indicates the last competitor is used—but don't input a character string, in general.
refvalue	Numeric. A positive value for the reference group.
ialpha	Initial values for the α s. These are recycled to the appropriate length.
i0	Initial value for α_0 . If convergence fails, try another positive value.

Details

There are several models that extend the ordinary Bradley Terry model to handle ties. This family function implements one of these models. It involves M competitors who either win or lose or tie against each other. (If there are no draws/ties then use `brat`). The probability that Competitor i beats Competitor j is $\alpha_i/(\alpha_i + \alpha_j + \alpha_0)$, where all the α s are positive. The probability that Competitor i ties with Competitor j is $\alpha_0/(\alpha_i + \alpha_j + \alpha_0)$. Loosely, the α s can be thought of as the competitors' 'abilities', and α_0 is an added parameter to model ties. For identifiability, one of the α_i is set to a known value `refvalue`, e.g., 1. By default, this function chooses the last competitor to have this reference value. The data can be represented in the form of a M by M matrix of counts,

where winners are the rows and losers are the columns. However, this is not the way the data should be inputted (see below).

Excluding the reference value/group, this function chooses $\log(\alpha_j)$ as the first $M - 1$ linear predictors. The log link ensures that the α s are positive. The last linear predictor is $\log(\alpha_0)$.

The Bradley Terry model can be fitted with covariates, e.g., a home advantage variable, but unfortunately, this lies outside the VGLM theoretical framework and therefore cannot be handled with this code.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#).

Note

The function [Brat](#) is useful for coercing a M by M matrix of counts into a one-row matrix suitable for [bratt](#). Diagonal elements are skipped, and the usual S order of `c(a.matrix)` of elements is used. There should be no missing values apart from the diagonal elements of the square matrix. The matrix should have winners as the rows, and losers as the columns. In general, the response should be a matrix with $M(M - 1)$ columns.

Also, a symmetric matrix of ties should be passed into [Brat](#). The diagonal of this matrix should be all NAs.

Only an intercept model is recommended with [bratt](#). It doesn't make sense really to include covariates because of the limited VGLM framework.

Notationally, note that the **VGAM** family function [brat](#) has $M + 1$ contestants, while [bratt](#) has M contestants.

Author(s)

T. W. Yee

References

Torsney, B. (2004). Fitting Bradley Terry models using a multiplicative algorithm. In: Antoch, J. (ed.) *Proceedings in Computational Statistics COMPSTAT 2004*, Physica-Verlag: Heidelberg. Pages 513–526.

See Also

[brat](#), [Brat](#), [binomialff](#).

Examples

```
# citation statistics: being cited is a 'win'; citing is a 'loss'
journal <- c("Biometrika", "Comm.Statist", "JASA", "JRSS-B")
mat <- matrix(c( NA, 33, 320, 284,
                730, NA, 813, 276,
                498, 68, NA, 325,
                221, 17, 142, NA), 4, 4)
```

```

dimnames(mat) <- list(winner = journal, loser = journal)

# Add some ties. This is fictional data.
ties <- 5 + 0 * mat
ties[2, 1] <- ties[1,2] <- 9

# Now fit the model
fit <- vglm(Brat(mat, ties) ~ 1, bratt(refgp = 1), trace = TRUE,
           crit = "coef")

summary(fit)
c(0, coef(fit)) # Log-abilities (last is log(alpha0))
c(1, Coef(fit)) # Abilities (last is alpha0)

fit@misc$alpha # alpha_1,...,alpha_M
fit@misc$alpha0 # alpha_0

fitted(fit) # Probabilities of winning and tying, in awkward form
predict(fit)
(check <- InverseBrat(fitted(fit))) # Probabilities of winning
qprob <- attr(fitted(fit), "probtie") # Probabilities of a tie
qprobmat <- InverseBrat(c(qprob), NCo = nrow(ties)) # Pr(tie)
check + t(check) + qprobmat # Should be 1s in the off-diagonals

```

budworm

Western Spuce Budworm

Description

Counts of western spuce budworm (*Choristoneura freemani*) across seven developmental stages (five larval instars, pupae, and adults) on 12 sampling occasions.

Usage

```
data(budworm)
```

Format

A data frame with the following variables.

ddeg Degree days.

total Sum of stages 1–7.

stage1, stage2, stage3, stage4 Successive stages.

stage5, stage6, stage7 Successive stages.

Details

This data concerns the development of a defoliating moth widespread in western North America (i.e., north of Mexico). According to Boersch-Supan (2021), the insect passes through successive stages $j = 1, \dots, r$, delimited by $r - 1$ moults. The data was originally used in a 1986 publication but has been corrected for two sampling occasions; the data appears in Candy (1990) and was analyzed in Boersch-Supan (2021). See the latter for more references.

Source

Candy, S. G. (1990). *Biology of the mountain pinhole borer, Platypus subgranosus Scheld, in Tasmania*. MA thesis, University of Tasmania, Australia. <https://eprints.utas.edu.au/18864/>.

References

Boersch-Supan, P. H. (2021). Modeling insect phenology using ordinal regression and continuation ratio models. *ReScience C*, **7.1**, 1–14.

Examples

```
budworm
summary(budworm)
```

 calibrate

Model Calibrations

Description

calibrate is a generic function used to produce calibrations from various model fitting functions. The function invokes particular ‘methods’ which depend on the ‘class’ of the first argument.

Usage

```
calibrate(object, ...)
```

Arguments

object	An object for which a calibration is desired.
...	Additional arguments affecting the calibration produced. Usually the most important argument in ... is newdata which, for calibrate, contains new <i>response</i> data, Y , say.

Details

Given a regression model with explanatory variables **X** and response **Y**, calibration involves estimating **X** from **Y** using the regression model. It can be loosely thought of as the opposite of `predict` (which takes an **X** and returns a **Y** of some sort.) In general, the central algorithm is maximum likelihood calibration.

Value

In general, given a new response **Y**, some function of the explanatory variables **X** are returned. For example, for constrained ordination models such as CQO and CAO models, it is usually not possible to return **X**, so the latent variables are returned instead (they are linear combinations of the **X**). See the specific calibrate methods functions to see what they return.

Note

This function was not called `predictx` because of the inability of constrained ordination models to return **X**; they can only return the latent variable values (also known as site scores) instead.

Author(s)

T. W. Yee

References

ter Braak, C. J. F. and van Dam, H. (1989). Inferring pH from diatoms: a comparison of old and new calibration methods. *Hydrobiologia*, **178**, 209–223.

See Also

[predict](#), [calibrate.rrvglm](#), [calibrate.qrrvglm](#).

Examples

```
## Not run:
hspider[, 1:6] <- scale(hspider[, 1:6]) # Stdzed environmental vars
set.seed(123)
pcao1 <- cao(cbind(Pardlugu, Pardmont, Pardnigr, Pardpull, Zoraspin) ~
             WaterCon + BareSand + FallTwig + CoveMoss + CoveHerb + ReflLux,
             family = poissonff, data = hspider, Rank = 1, Bestof = 3,
             df1.nl = c(Zoraspin = 2, 1.9), Crowlpositive = TRUE)

siteNos <- 1:2 # Calibrate these sites
cpcao1 <- calibrate(pcao1, trace = TRUE,
                  newdata = data.frame(depvar(pcao1)[siteNos, ],
                                       model.matrix(pcao1)[siteNos, ]))

# Graphically compare the actual site scores with their calibrated values
persp(pcao1, main = "Site scores: solid=actual, dashed=calibrated",
      label = TRUE, col = "blue", las = 1)
abline(v = latvar(pcao1)[siteNos], col = seq(siteNos)) # Actual scores
abline(v = cpcao1, lty = 2, col = seq(siteNos)) # Calibrated values

## End(Not run)
```

calibrate-methods	<i>Calibration for Constrained Regression Models</i>
-------------------	--

Description

calibrate is a generic function applied to RR-VGLMs, QRR-VGLMs and RR-VGAMs, etc.

Methods

object The object from which the calibration is performed.

calibrate.qrrvglm	<i>Calibration for CQO and CAO models</i>
-------------------	---

Description

Performs maximum likelihood calibration for constrained quadratic and additive ordination models (CQO and CAO models are better known as QRR-VGLMs and RR-VGAMs respectively).

Usage

```
calibrate.qrrvglm(object, newdata = NULL,
  type = c("latvar", "predictors", "response", "vcov", "everything"),
  lr.confint = FALSE, cf.confint = FALSE,
  level = 0.95, initial.vals = NULL, ...)
```

Arguments

object	The fitted CQO/CAO model.
newdata	A data frame with new response data, such as new species data. The default is to use the original data used to fit the model; however, the calibration may take a long time to compute because the computations are expensive. Note that the creation of the model frame associated with newdata is fragile. Factors may not be created properly. If a variable is binary then its best for it to be straightforward and have only 0 and 1 as values.
type	What type of result to be returned. The first are the calibrated latent variables or site scores. This is always computed. The "predictors" are the linear/quadratic or additive predictors evaluated at the calibrated latent variables or site scores. The "response" are the fitted values (usually means) evaluated at the calibrated latent variables or site scores. The "vcov" are the Wald-type estimated variance-covariance matrices of the calibrated latent variables or site scores. The "everything" is for all of them, i.e., all types. Note that for CAO models, the "vcov" type is unavailable.

<code>lr.confint, level</code>	Compute <i>approximate</i> likelihood ratio based confidence intervals? If TRUE then <code>level</code> is the confidence level required and one should have <code>type = "latvar"</code> or <code>type = "everything"</code> ; and currently only rank-1 models are supported. This option works for CLO and CQO models and not for CAO models. The function <code>uniroot</code> is called to solve for the root of a nonlinear equation to obtain each confidence limit, and this is not entirely reliable. It is assumed that the likelihood function is unimodal about its MLE because only one root is returned if there is more than one. One root is found on each side of the MLE. Technically, the default is to find the value of the latent variable whose difference in deviance (or twice the difference in log-likelihoods) from the optimal model is equal to <code>qchisq(level, df = 1)</code> . The intervals are not true profile likelihood intervals because it is not possible to estimate the regression coefficients of the QRR-VGLM/RR-VGLM based on one response vector. See <code>confint</code> to get the flavour of these two arguments in general.
<code>cf.confint</code>	Compute <i>approximate</i> characteristic function based confidence intervals? If TRUE then <code>level</code> is the confidence level required and one should have <code>type = "latvar"</code> or <code>type = "everything"</code> ; and currently only rank-1 models are supported. This option works for <code>binomialff</code> and <code>poissonff</code> CLO and CQO models and not for CAO models. The function <code>uniroot</code> is called to solve for the root of a nonlinear equation to obtain each confidence limit, and this is not entirely reliable. It is assumed that the likelihood function is unimodal because only one root is returned if there is more than one. Technically, the CDF of a normalized score statistic is obtained by Gauss–Hermite numerical integration of a complex-valued integrand, and this is based on the inversion formula described in Abate and Witt (1992).
<code>initial.vals</code>	Initial values for the search. For rank-1 models, this should be a vector having length equal to <code>nrow(newdata)</code> , and for rank-2 models this should be a two-column matrix with the number of rows equalling the number of rows in <code>newdata</code> . The default is a grid defined by arguments in <code>calibrate.qrrvglm.control</code> .
<code>...</code>	Arguments that are fed into <code>calibrate.qrrvglm.control</code> .

Details

Given a fitted regression CQO/CAO model, maximum likelihood calibration is theoretically easy and elegant. However, the method assumes that all the responses are independent, which is often not true in practice. More details and references are given in Yee (2018) and ch.6 of Yee (2015).

The function `optim` is used to search for the maximum likelihood solution. Good initial values are needed, and arguments in `calibrate.qrrvglm.control` allows the user some control over the choice of these.

Value

Several methods are implemented to obtain confidence intervals/regions for the calibration estimates. One method is when `lr.confint = TRUE`, then a 4-column matrix is returned with the confidence limits being the final 2 columns (if `type = "everything"` then the matrix is returned in the `lr.confint` list component). Another similar method is when `cf.confint = TRUE`. There may be

some redundancy in whatever is returned. Other methods are returned by using `type` and they are described as follows.

The argument `type` determines what is returned. If `type = "everything"` then all the type values are returned in a list, with the following components. Each component has length `nrow(newdata)`.

<code>latvar</code>	Calibrated latent variables or site scores (the default). This may have the attribute <code>"objectiveFunction"</code> which is usually the log-likelihood or the deviance.
<code>predictors</code>	linear/quadratic or additive predictors. For example, for Poisson families, this will be on a log scale, and for binomial families, this will be on a logit scale.
<code>response</code>	Fitted values of the response, evaluated at the calibrated latent variables.
<code>vcov</code>	Wald-type estimated variance-covariance matrices of the calibrated latent variables or site scores. Actually, these are stored in a 3-D array whose dimension is <code>c(Rank(object), Rank(object), nrow(newdata))</code> . This type has only been implemented for <code>binomialff</code> and <code>poissonff</code> models with canonical links and <code>noRRR = ~ 1</code> and, for CQOs, <code>I.tolerances = TRUE</code> or <code>eq.tolerances = TRUE</code> .

Warning

This function is computationally expensive. Setting `trace = TRUE` to get a running log can be a good idea. This function has been tested but not extensively.

Note

Despite the name of this function, CAO models are handled as well to a certain extent. Some combinations of parameters are not handled, e.g., `lr.confint = TRUE` only works for rank-1, `type = "vcov"` only works for `binomialff` and `poissonff` models with canonical links and `noRRR = ~ 1`, and higher-order rank models need `eq.tolerances = TRUE` or `I.tolerances = TRUE` as well. For rank-1 objects, `lr.confint = TRUE` is recommended above `type = "vcov"` in terms of accuracy and overall generality. For class `"qrrvglm"` objects it is necessary that all response' tolerance matrices are positive-definite which correspond to bell-shaped response curves/surfaces.

For `binomialff` and `poissonff` models the deviance slot is used for the optimization rather than the loglikelihood slot, therefore one can calibrate using real-valued responses. (If the loglikelihood slot were used then functions such as `dpois` would be used with `log = TRUE` and then would be restricted to feed in integer-valued response values.)

Maximum likelihood calibration for Gaussian logit regression models may be performed by `ri-oja` but this applies to a single environmental variable such as pH in `data("SWAP", package = "rioja")`. In **VGAM** `calibrate()` estimates values of the *latent variable* rather than individual explanatory variables, hence the setting is more on ordination.

Author(s)

T. W. Yee. Recent work on the standard errors by David Zucker and Sam Oman at HUJI is gratefully acknowledged—these are returned in the `vcov` component and provided inspiration for `lr.confint` and `cf.confint`. A joint publication is being prepared on this subject.

References

Abate, J. and Whitt, W. (1992). The Fourier-series method for inverting transforms of probability distributions. *Queueing Systems*, **10**, 5–88.

ter Braak, C. J. F. (1995). Calibration. In: *Data Analysis in Community and Landscape Ecology* by Jongman, R. H. G., ter Braak, C. J. F. and van Tongeren, O. F. R. (Eds.) Cambridge University Press, Cambridge.

See Also

[calibrate.qrrvglm.control](#), [calibrate.rrvglm](#), [calibrate](#), [cqo](#), [cao](#), [optim](#), [uniroot](#).

Examples

```
## Not run:
hspider[, 1:6] <- scale(hspider[, 1:6]) # Stdze environmental variables
set.seed(123)
siteNos <- c(1, 5) # Calibrate these sites
pet1 <- cqo(cbind(Pardlugu, Pardmont, Pardnigr, Pardpull, Zoraspin) ~
  WaterCon + BareSand + FallTwig + CoveMoss + CoveHerb + ReflLux,
  trace = FALSE,
  data = hspider[-siteNos, ], # Sites not in fitted model
  family = poissonff, I.toler = TRUE, Crowlpositive = TRUE)
y0 <- hspider[siteNos, colnames(depvar(pet1))] # Species counts
(cpet1 <- calibrate(pet1, trace = TRUE, newdata = data.frame(y0)))
(clrpet1 <- calibrate(pet1, lr.confint = TRUE, newdata = data.frame(y0)))
(ccfpet1 <- calibrate(pet1, cf.confint = TRUE, newdata = data.frame(y0)))
(cp1wald <- calibrate(pet1, newdata = y0, type = "everything"))

## End(Not run)

## Not run:
# Graphically compare the actual site scores with their calibrated
# values. 95 percent likelihood-based confidence intervals in green.
persp(pet1, main = "Site scores: solid=actual, dashed=calibrated",
  label = TRUE, col = "gray50", las = 1)
# Actual site scores:
xvars <- rownames(concoef(pet1)) # Variables comprising the latvar
est.latvar <- as.matrix(hspider[siteNos, xvars]) %*% concoef(pet1)
abline(v = est.latvar, col = seq(siteNos))
abline(v = cpet1, lty = 2, col = seq(siteNos)) # Calibrated values
arrows(clrpet1[, 3], c(60, 60), clrpet1[, 4], c(60, 60), # Add CIs
  length = 0.08, col = "orange", angle = 90, code = 3, lwd = 2)
arrows(ccfpet1[, 3], c(70, 70), ccfpet1[, 4], c(70, 70), # Add CIs
  length = 0.08, col = "limegreen", angle = 90, code = 3, lwd = 2)
arrows(cp1wald$latvar - 1.96 * sqrt(cp1wald$vcov), c(65, 65),
  cp1wald$latvar + 1.96 * sqrt(cp1wald$vcov), c(65, 65), # Wald CIs
  length = 0.08, col = "blue", angle = 90, code = 3, lwd = 2)
legend("topright", lwd = 2,
  leg = c("CF interval", "Wald interval", "LR interval"),
  col = c("limegreen", "blue", "orange"), lty = 1)
```

```
## End(Not run)
```

```
calibrate.qrrvglm.control
```

Control Function for CQO/CAO Calibration

Description

Algorithmic constants and parameters for running `calibrate.qrrvglm` are set using this function.

Usage

```
calibrate.qrrvglm.control(object, trace = FALSE, method.optim = "BFGS",
  gridSize = ifelse(Rank == 1, 21, 9), varI.latvar = FALSE, ...)
```

Arguments

<code>object</code>	The fitted CQO/CAO model. The user should ignore this argument.
<code>trace</code>	Logical indicating if output should be produced for each iteration. It is a good idea to set this argument to be TRUE since the computations are expensive.
<code>method.optim</code>	Character. Fed into the method argument of <code>optim</code> .
<code>gridSize</code>	Numeric, recycled to length Rank. Controls the resolution of the grid used for initial values. For each latent variable, an equally spaced grid of length <code>gridSize</code> is cast from the smallest site score to the largest site score. Then the likelihood function is evaluated on the grid, and the best fit is chosen as the initial value. Thus increasing the value of <code>gridSize</code> increases the chance of obtaining the global solution, however, the computing time increases proportionately.
<code>varI.latvar</code>	Logical. For CQO objects only, this argument is fed into <code>Coef.qrrvglm</code> .
<code>...</code>	Avoids an error message for extraneous arguments.

Details

Most CQO/CAO users will only need to make use of `trace` and `gridSize`. These arguments should be used inside their call to `calibrate.qrrvglm`, not this function directly.

Value

A list which with the following components.

<code>trace</code>	Numeric (even though the input can be logical).
<code>gridSize</code>	Positive integer.
<code>varI.latvar</code>	Logical.

Note

Despite the name of this function, CAO models are handled as well.

References

Yee, T. W. (2020). On constrained and unconstrained quadratic ordination. *Manuscript in preparation*.

See Also

[calibrate.qrrvglm](#), [Coef.qrrvglm](#).

Examples

```
## Not run: hspider[, 1:6] <- scale(hspider[, 1:6]) # Needed for I.tol=TRUE
set.seed(123)
p1 <- cqo(cbind(Alopacce, Alopcone, Pardlugu, Pardnigr,
               Pardpull, Trocterr, Zoraspin) ~
          WaterCon + BareSand + FallTwig + CoveMoss + CoveHerb + ReflLux,
          family = poissonff, data = hspider, I.tol = TRUE)
sort(deviance(p1, history = TRUE)) # A history of all the iterations
siteNos <- 3:4 # Calibrate these sites
cp1 <- calibrate(p1, trace = TRUE,
                new = data.frame(depvar(p1)[siteNos, ]))

## End(Not run)
## Not run:
# Graphically compare the actual site scores with their calibrated values
persp(p1, main = "Site scores: solid=actual, dashed=calibrated",
      label = TRUE, col = "blue", las = 1)
abline(v = latvar(p1)[siteNos], col = seq(siteNos)) # Actual site scores
abline(v = cp1, lty = 2, col = seq(siteNos)) # Calibrated values

## End(Not run)
```

calibrate.rrvglm

Calibration for CLO models (RR-VGLMs)

Description

Performs maximum likelihood calibration for constrained linear ordination models (CLO models are better known as RR-VGLMs).

Usage

```
calibrate.rrvglm(object, newdata = NULL,
                 type = c("latvar", "predictors", "response", "vcov", "everything"),
                 lr.confint = FALSE, cf.confint = FALSE,
                 level = 0.95, initial.vals = NULL, ...)
```

Arguments

object	The fitted <code>rrvglm</code> model. Note that object should be fitted with corner constraints.
newdata	See <code>calibrate.qrrvglm</code> .
type	See <code>calibrate.qrrvglm</code> . If <code>type = "vcov"</code> then object should have been fitted using <code>binomialff</code> or <code>poissonff</code> with canonical links, and have <code>noRRR = ~ 1</code> .
lr.confint, cf.confint, level	Same as <code>calibrate.qrrvglm</code> .
initial.vals	Same as <code>calibrate.qrrvglm</code> . The default is a grid defined by arguments in <code>calibrate.rrvglm.control</code> .
...	Arguments that are fed into <code>calibrate.rrvglm.control</code> .

Details

Given a fitted regression CLO model, maximum likelihood calibration is theoretically easy and elegant. However, the method assumes that all responses are independent. More details and references are given in Yee (2015).

Calibration requires *grouped* or *non-sparse* data as the response. For example, if the family function is `multinomial` then one cannot usually calibrate y_0 if it is a vector of 0s except for one 1. Instead, the response vector should be from grouped data so that there are few 0s. Indeed, it is found empirically that the stereotype model (also known as a reduced-rank `multinomial` logit model) calibrates well only with grouped data, and if the response vector is all 0s except for one 1 then the MLE will probably be at $-\text{Inf}$ or $+\text{Inf}$. As another example, if the family function is `poissonff` then y_0 must not be a vector of all 0s; instead, the response vector should have few 0s ideally. In general, you can use simulation to see what type of data calibrates acceptably.

Internally, this function is a simplification of `calibrate.qrrvglm` and users should look at that function for details. Good initial values are needed, and a grid is constructed to obtain these. The function `calibrate.rrvglm.control` allows the user some control over the choice of these.

Value

See `calibrate.qrrvglm`. Of course, the quadratic term in the latent variables vanishes for RR-VGLMs, so the model is simpler.

Warning

See `calibrate.qrrvglm`.

Note

See `calibrate.qrrvglm` about, e.g., calibration using real-valued responses.

Author(s)

T. W. Yee

See Also

[calibrate.qrrvglm](#), [calibrate](#), [rrvglm](#), [weightsvglm](#), [optim](#), [uniroot](#).

Examples

```
## Not run: # Example 1
nona.xs.nz <- na.omit(xs.nz) # Overkill!! (Data in VGAMdata package)
nona.xs.nz$dmd <- with(nona.xs.nz, round(drinkmaxday))
nona.xs.nz$feethr <- with(nona.xs.nz, round(feethour))
nona.xs.nz$sleephr <- with(nona.xs.nz, round(sleep))
nona.xs.nz$beats <- with(nona.xs.nz, round(pulse))

p2 <- rrvglm(cbind(dmd, feethr, sleephr, beats) ~ age + smokenow +
  depressed + embarrassed + fedup + hurt + miserable + # 11 psychological
  nofriend + moody + nervous + tense + worry + worrier, # variables
  noRRR = ~ age + smokenow, trace = FALSE, poissonff, data = nona.xs.nz,
  Rank = 2)
cp2 <- calibrate(p2, newdata = head(nona.xs.nz, 9), trace = TRUE)
cp2

two.cases <- nona.xs.nz[1:2, ] # Another calibration example
two.cases$dmd <- c(4, 10)
two.cases$feethr <- c(4, 7)
two.cases$sleephr <- c(7, 8)
two.cases$beats <- c(62, 71)
(cp2b <- calibrate(p2, newdata = two.cases))

# Example 2
p1 <- rrvglm(cbind(dmd, feethr, sleephr, beats) ~ age + smokenow +
  depressed + embarrassed + fedup + hurt + miserable + # 11 psychological
  nofriend + moody + nervous + tense + worry + worrier, # variables
  noRRR = ~ age + smokenow, trace = FALSE, poissonff, data = nona.xs.nz,
  Rank = 1)
(cp1c <- calibrate(p1, newdata = two.cases, lr.confint = TRUE))

## End(Not run)
```

calibrate.rrvglm.control

Control Function for CLO (RR-VGLM) Calibration

Description

Algorithmic constants and parameters for running [calibrate.rrvglm](#) are set using this function.

Usage

```
calibrate.rrvglm.control(object, trace = FALSE, method.optim = "BFGS",
  gridSize = ifelse(Rank == 1, 17, 9), ...)
```


Arguments

object	The fitted <code>rrvglm</code> model. The user should ignore this argument.
trace, method, optim	Same as <code>calibrate.qrrvglm.control</code> .
gridSize	Same as <code>calibrate.qrrvglm.control</code> .
...	Avoids an error message for extraneous arguments.

Details

Most CLO users will only need to make use of `trace` and `gridSize`. These arguments should be used inside their call to `calibrate.rrvglm`, not this function directly.

Value

Similar to `calibrate.qrrvglm.control`.

See Also

`calibrate.rrvglm`, `Coef.rrvglm`.

 cao

Fitting Constrained Additive Ordination (CAO)

Description

A constrained additive ordination (CAO) model is fitted using the *reduced-rank vector generalized additive model* (RR-VGAM) framework.

Usage

```
cao(formula, family = stop("argument 'family' needs to be assigned"),
    data = list(),
    weights = NULL, subset = NULL, na.action = na.fail,
    etastart = NULL, mustart = NULL, coefstart = NULL,
    control = cao.control(...), offset = NULL,
    method = "cao.fit", model = FALSE, x.arg = TRUE, y.arg = TRUE,
    contrasts = NULL, constraints = NULL,
    extra = NULL, qr.arg = FALSE, smart = TRUE, ...)
```

Arguments

formula	a symbolic description of the model to be fit. The RHS of the formula is used to construct the latent variables, upon which the smooths are applied. All the variables in the formula are used for the construction of latent variables except for those specified by the argument <code>noRRR</code> , which is itself a formula. The LHS of the formula contains the response variables, which should be a matrix with each column being a response (species).
---------	--

family	a function of class "vglmff" (see vglmff-class) describing what statistical model is to be fitted. This is called a "VGAM family function". See CommonVGAMffArguments for general information about many types of arguments found in this type of function. See cqq for a list of those presently implemented.
data	an optional data frame containing the variables in the model. By default the variables are taken from <code>environment(formula)</code> , typically the environment from which <code>cao</code> is called.
weights	an optional vector or matrix of (prior) weights to be used in the fitting process. For <code>cao</code> , this argument currently should not be used.
subset	an optional logical vector specifying a subset of observations to be used in the fitting process.
na.action	a function which indicates what should happen when the data contain NAs. The default is set by the <code>na.action</code> setting of options , and is <code>na.fail</code> if that is unset. The "factory-fresh" default is <code>na.omit</code> .
etastart	starting values for the linear predictors. It is a M -column matrix. If $M = 1$ then it may be a vector. For <code>cao</code> , this argument currently should not be used.
mustart	starting values for the fitted values. It can be a vector or a matrix. Some family functions do not make use of this argument. For <code>cao</code> , this argument currently should not be used.
coefstart	starting values for the coefficient vector. For <code>cao</code> , this argument currently should not be used.
control	a list of parameters for controlling the fitting process. See cao.control for details.
offset	a vector or M -column matrix of offset values. These are <i>a priori</i> known and are added to the linear predictors during fitting. For <code>cao</code> , this argument currently should not be used.
method	the method to be used in fitting the model. The default (and presently only) method <code>cao.fit</code> uses iteratively reweighted least squares (IRLS) within FORTRAN code called from optim .
model	a logical value indicating whether the <i>model frame</i> should be assigned in the model slot.
x.arg, y.arg	logical values indicating whether the model matrix and response vector/matrix used in the fitting process should be assigned in the x and y slots. Note the model matrix is the linear model (LM) matrix.
contrasts	an optional list. See the <code>contrasts.arg</code> of model.matrix.default .
constraints	an optional list of constraint matrices. For <code>cao</code> , this argument currently should not be used. The components of the list must be named with the term it corresponds to (and it must match in character format). Each constraint matrix must have M rows, and be of full-column rank. By default, constraint matrices are the M by M identity matrix unless arguments in the family function itself override these values. If <code>constraints</code> is used it must contain <i>all</i> the terms; an incomplete list is not accepted.
extra	an optional list with any extra information that might be needed by the family function. For <code>cao</code> , this argument currently should not be used.

<code>qr.arg</code>	For cao, this argument currently should not be used.
<code>smart</code>	logical value indicating whether smart prediction (<code>smartpred</code>) will be used.
<code>...</code>	further arguments passed into <code>cao.control</code> .

Details

The arguments of `cao` are a mixture of those from `vgam` and `cqo`, but with some extras in `cao.control`. Currently, not all of the arguments work properly.

CAO can be loosely be thought of as the result of fitting generalized additive models (GAMs) to several responses (e.g., species) against a very small number of latent variables. Each latent variable is a linear combination of the explanatory variables; the coefficients \mathbf{C} (called C below) are called *constrained coefficients* or *canonical coefficients*, and are interpreted as weights or loadings. The \mathbf{C} are estimated by maximum likelihood estimation. It is often a good idea to apply `scale` to each explanatory variable first.

For each response (e.g., species), each latent variable is smoothed by a cubic smoothing spline, thus CAO is data-driven. If each smooth were a quadratic then CAO would simplify to *constrained quadratic ordination* (CQO; formerly called *canonical Gaussian ordination* or CGO). If each smooth were linear then CAO would simplify to *constrained linear ordination* (CLO). CLO can theoretically be fitted with `cao` by specifying `df1.n1=0`, however it is more efficient to use `rrvglm`.

Currently, only Rank=1 is implemented, and only noRRR = ~1 models are handled.

With binomial data, the default formula is

$$\text{logit}(P[Y_s = 1]) = \eta_s = f_s(\nu), \quad s = 1, 2, \dots, S$$

where x_2 is a vector of environmental variables, and $\nu = C^T x_2$ is a R -vector of latent variables. The η_s is an additive predictor for species s , and it models the probabilities of presence as an additive model on the logit scale. The matrix C is estimated from the data, as well as the smooth functions f_s . The argument `noRRR = ~ 1` specifies that the vector x_1 , defined for RR-VGLMs and QRR-VGLMs, is simply a 1 for an intercept. Here, the intercept in the model is absorbed into the functions. A `clogloglink` link may be preferable over a `logitlink` link.

With Poisson count data, the formula is

$$\log(E[Y_s]) = \eta_s = f_s(\nu)$$

which models the mean response as an additive models on the log scale.

The fitted latent variables (site scores) are scaled to have unit variance. The concept of a tolerance is undefined for CAO models, but the optimums and maximums are defined. The generic functions `Max` and `Opt` should work for CAO objects, but note that if the maximum occurs at the boundary then `Max` will return a NA. Inference for CAO models is currently undeveloped.

Value

An object of class "cao" (this may change to "rrvgam" in the future). Several generic functions can be applied to the object, e.g., `Coef`, `concoef`, `lvplot`, `summary`.

Warning

CAO is very costly to compute. With version 0.7-8 it took 28 minutes on a fast machine. I hope to look at ways of speeding things up in the future.

Use `set.seed` just prior to calling `cao()` to make your results reproducible. The reason for this is finding the optimal CAO model presents a difficult optimization problem, partly because the log-likelihood function contains many local solutions. To obtain the (global) solution the user is advised to try *many* initial values. This can be done by setting `Bestof` some appropriate value (see `cao.control`). Trying many initial values becomes progressively more important as the nonlinear degrees of freedom of the smooths increase.

Note

CAO models are computationally expensive, therefore setting `trace = TRUE` is a good idea, as well as running it on a simple random sample of the data set instead.

Sometimes the IRLS algorithm does not converge within the FORTRAN code. This results in warnings being issued. In particular, if an error code of 3 is issued, then this indicates the IRLS algorithm has not converged. One possible remedy is to increase or decrease the nonlinear degrees of freedom so that the curves become more or less flexible, respectively.

Author(s)

T. W. Yee

References

Yee, T. W. (2006). Constrained additive ordination. *Ecology*, **87**, 203–213.

See Also

`cao.control`, `Coef.cao`, `cqo`, `latvar`, `Opt`, `Max`, `calibrate.qrrvglm`, `persp.cao`, `poissonff`, `binomialff`, `negbinomial`, `gamma2`, `set.seed`, `gam()` in **gam**, `trap0`.

Examples

```
## Not run:
hspider[, 1:6] <- scale(hspider[, 1:6]) # Stdzd environmental vars
set.seed(149) # For reproducible results
ap1 <- cao(cbind(Pardlugu, Pardmont, Pardnigr, Pardpull) ~
  WaterCon + BareSand + FallTwig + CoveMoss + CoveHerb + ReflLux,
  family = poissonff, data = hspider, Rank = 1,
  df1.nl = c(Pardpull= 2.7, 2.5),
  Bestof = 7, Crow1positive = FALSE)
sort(deviance(ap1, history = TRUE)) # A history of all the iterations

Coef(ap1)
concoef(ap1)

par(mfrow = c(2, 2))
plot(ap1) # All the curves are unimodal; some quite symmetric
```

```

par(mfrow = c(1, 1), las = 1)
index <- 1:ncol(depvar(ap1))
lvplot(ap1, lcol = index, pcol = index, y = TRUE)

trplot(ap1, label = TRUE, col = index)
abline(a = 0, b = 1, lty = 2)

trplot(ap1, label = TRUE, col = "blue", log = "xy", which.sp = c(1, 3))
abline(a = 0, b = 1, lty = 2)

persp(ap1, col = index, lwd = 2, label = TRUE)
abline(v = Opt(ap1), lty = 2, col = index)
abline(h = Max(ap1), lty = 2, col = index)

## End(Not run)

```

 cao.control

Control Function for RR-VGAMs (CAO)

Description

Algorithmic constants and parameters for a constrained additive ordination (CAO), by fitting a *reduced-rank vector generalized additive model* (RR-VGAM), are set using this function. This is the control function for [cao](#).

Usage

```

cao.control(Rank = 1, all.knots = FALSE, criterion = "deviance", Cinit = NULL,
  Crow1positive = TRUE, epsilon = 1.0e-05, Etamat.colmax = 10,
  GradientFunction = FALSE, iKvector = 0.1, iShape = 0.1,
  noRRR = ~ 1, Norrr = NA,
  SmallNo = 5.0e-13, Use.Init.Poisson.Q0 = TRUE,
  Bestof = if (length(Cinit)) 1 else 10, maxitl = 10,
  imethod = 1, bf.epsilon = 1.0e-7, bf.maxit = 10,
  Maxit.optim = 250, optim.maxit = 20, sd.sitescores = 1.0,
  sd.Cinit = 0.02, suppress.warnings = TRUE,
  trace = TRUE, df1.nl = 2.5, df2.nl = 2.5,
  spar1 = 0, spar2 = 0, ...)

```

Arguments

Rank	The numerical rank R of the model, i.e., the number of latent variables. Currently only Rank = 1 is implemented.
all.knots	Logical indicating if all distinct points of the smoothing variables are to be used as knots. Assigning the value FALSE means fewer knots are chosen when the number of distinct points is large, meaning less computational expense. See vgam.control for details.

criterion	Convergence criterion. Currently, only one is supported: the deviance is minimized.
Cinit	Optional initial C matrix which may speed up convergence.
Crow1positive	Logical vector of length Rank (recycled if necessary): are the elements of the first row of C positive? For example, if Rank is 4, then specifying <code>Crow1positive = c(FALSE, TRUE)</code> will force C [1,1] and C [1,3] to be negative, and C [1,2] and C [1,4] to be positive.
epsilon	Positive numeric. Used to test for convergence for GLMs fitted in FORTRAN. Larger values mean a loosening of the convergence criterion.
Etamat.colmax	Positive integer, no smaller than Rank. Controls the amount of memory used by <code>.Init.Poisson.QO()</code> . It is the maximum number of columns allowed for the pseudo-response and its weights. In general, the larger the value, the better the initial value. Used only if <code>Use.Init.Poisson.QO = TRUE</code> .
GradientFunction	Logical. Whether <code>optim</code> 's argument <code>gr</code> is used or not, i.e., to compute gradient values. Used only if <code>FastAlgorithm</code> is <code>TRUE</code> . Currently, this argument must be set to <code>FALSE</code> .
iKvector, iShape	See qrrvglm.control .
noRRR	Formula giving terms that are <i>not</i> to be included in the reduced-rank regression (or formation of the latent variables). The default is to omit the intercept term from the latent variables. Currently, only <code>noRRR = ~ 1</code> is implemented.
Norrr	Defunct. Please use <code>noRRR</code> . Use of <code>Norrr</code> will become an error soon.
SmallNo	Positive numeric between <code>.Machine\$double.eps</code> and <code>0.0001</code> . Used to avoid under- or over-flow in the IRLS algorithm.
Use.Init.Poisson.QO	Logical. If <code>TRUE</code> then the function <code>.Init.Poisson.QO</code> is used to obtain initial values for the canonical coefficients C . If <code>FALSE</code> then random numbers are used instead.
Bestof	Integer. The best of <code>Bestof</code> models fitted is returned. This argument helps guard against local solutions by (hopefully) finding the global solution from many fits. The argument works only when the function generates its own initial value for C , i.e., when C are <i>not</i> passed in as initial values. The default is only a convenient minimal number and users are urged to increase this value.
maxitl	Positive integer. Maximum number of Newton-Raphson/Fisher-scoring/local-scoring iterations allowed.
imethod	See qrrvglm.control .
bf.epsilon	Positive numeric. Tolerance used by the modified vector backfitting algorithm for testing convergence.
bf.maxit	Positive integer. Number of backfitting iterations allowed in the compiled code.
Maxit.optim	Positive integer. Number of iterations given to the function <code>optim</code> at each of the <code>optim.maxit</code> iterations.
optim.maxit	Positive integer. Number of times <code>optim</code> is invoked.

sd.sitescores	Numeric. Standard deviation of the initial values of the site scores, which are generated from a normal distribution. Used when Use.Init.Poisson.Q0 is FALSE.
sd.Cinit	Standard deviation of the initial values for the elements of C. These are normally distributed with mean zero. This argument is used only if Use.Init.Poisson.Q0 = FALSE.
suppress.warnings	Logical. Suppress warnings?
trace	Logical indicating if output should be produced for each iteration. Having the value TRUE is a good idea for large data sets.
df1.nl, df2.nl	Numeric and non-negative, recycled to length S . Nonlinear degrees of freedom for smooths of the first and second latent variables. A value of 0 means the smooth is linear. Roughly, a value between 1.0 and 2.0 often has the approximate flexibility of a quadratic. The user should not assign too large a value to this argument, e.g., the value 4.0 is probably too high. The argument df1.nl is ignored if spar1 is assigned a positive value or values. Ditto for df2.nl.
spar1, spar2	Numeric and non-negative, recycled to length S . Smoothing parameters of the smooths of the first and second latent variables. The larger the value, the more smooth (less wiggly) the fitted curves. These arguments are an alternative to specifying df1.nl and df2.nl. A value 0 (the default) for spar1 means that df1.nl is used. Ditto for spar2. The values are on a scaled version of the latent variables. See Green and Silverman (1994) for more information.
...	Ignored at present.

Details

Many of these arguments are identical to `qrrvglm.control`. Here, R is the Rank, M is the number of additive predictors, and S is the number of responses (species). Thus $M = S$ for binomial and Poisson responses, and $M = 2S$ for the negative binomial and 2-parameter gamma distributions.

Allowing the smooths too much flexibility means the CAO optimization problem becomes more difficult to solve. This is because the number of local solutions increases as the nonlinearity of the smooths increases. In situations of high nonlinearity, many initial values should be used, so that Bestof should be assigned a larger value. In general, there should be a reasonable value of df1.nl somewhere between 0 and about 3 for most data sets.

Value

A list with the components corresponding to its arguments, after some basic error checking.

Note

The argument df1.nl can be inputted in the format `c(spp1 = 2, spp2 = 3, 2.5)`, say, meaning the default value is 2.5, but two species have alternative values.

If `spar1 = 0` and `df1.nl = 0` then this represents fitting linear functions (CLO). Currently, this is handled in the awkward manner of setting df1.nl to be a small positive value, so that the smooth is almost linear but not quite. A proper fix to this special case should be done in the short future.

Author(s)

T. W. Yee

References

- Yee, T. W. (2006). Constrained additive ordination. *Ecology*, **87**, 203–213.
- Green, P. J. and Silverman, B. W. (1994). *Nonparametric Regression and Generalized Linear Models: A Roughness Penalty Approach*, London: Chapman & Hall.

See Also[cao](#).**Examples**

```
## Not run:
hspider[,1:6] <- scale(hspider[,1:6]) # Standardized environmental vars
set.seed(123)
ap1 <- cao(cbind(Pardlugu, Pardmont, Pardnigr, Pardpull, Zoraspin) ~
           WaterCon + BareSand + FallTwig +
           CoveMoss + CoveHerb + ReflLux,
           family = poissonff, data = hspider,
           df1.nl = c(Zoraspin = 2.3, 2.1),
           Bestof = 10, Crowlpositive = FALSE)
sort(deviance(ap1, history = TRUE)) # A history of all the iterations

Coef(ap1)

par(mfrow = c(2, 3)) # All or most of the curves are unimodal; some are
plot(ap1, lcol = "blue") # quite symmetric. Hence a CQO model should be ok

par(mfrow = c(1, 1), las = 1)
index <- 1:ncol(depvar(ap1)) # lvplot is jagged because only 28 sites
lvplot(ap1, lcol = index, pcol = index, y = TRUE)

trplot(ap1, label = TRUE, col = index)
abline(a = 0, b = 1, lty = 2)

persp(ap1, label = TRUE, col = 1:4)

## End(Not run)
```

Card

Cardioid Distribution

Description

Density, distribution function, quantile function and random generation for the cardioid distribution.

Usage

```

dcard(x, mu, rho, log = FALSE)
pcard(q, mu, rho, lower.tail = TRUE, log.p = FALSE)
qcard(p, mu, rho, tolerance = 1e-07, maxits = 500,
      lower.tail = TRUE, log.p = FALSE)
rcard(n, mu, rho, ...)

```

Arguments

<code>x, q</code>	vector of quantiles.
<code>p</code>	vector of probabilities.
<code>n</code>	number of observations. Same as in runif .
<code>mu, rho</code>	See cardioid for more information.
<code>tolerance, maxits, ...</code>	The first two are control parameters for the algorithm used to solve for the roots of a nonlinear system of equations; <code>tolerance</code> controls for the accuracy and <code>maxits</code> is the maximum number of iterations. <code>rcard</code> calls <code>qcard</code> so the <code>...</code> can be used to vary the two arguments.
<code>log</code>	Logical. If <code>log = TRUE</code> then the logarithm of the density is returned.
<code>lower.tail, log.p</code>	Same meaning as in pnorm or qnorm .

Details

See [cardioid](#), the **VGAM** family function for estimating the two parameters by maximum likelihood estimation, for the formula of the probability density function and other details.

Value

`dcard` gives the density, `pcard` gives the distribution function, `qcard` gives the quantile function, and `rcard` generates random deviates.

Note

Convergence problems might occur with `rcard`.

Author(s)

Thomas W. Yee and Kai Huang

See Also

[cardioid](#).

Examples

```
## Not run:
mu <- 4; rho <- 0.4; x <- seq(0, 2*pi, len = 501)
plot(x, dcard(x, mu, rho), type = "l", las = 1, ylim = c(0, 1),
      ylab = paste("[dp]card(mu=", mu, ", rho=", rho, ")"),
      main = "Blue is density, orange is the CDF", col = "blue",
      sub = "Purple lines are the 10,20,...,90 percentiles")
lines(x, pcard(x, mu, rho), col = "orange")

probs <- seq(0.1, 0.9, by = 0.1)
Q <- qcard(probs, mu, rho)
lines(Q, dcard(Q, mu, rho), col = "purple", lty = 3, type = "h")
lines(Q, pcard(Q, mu, rho), col = "purple", lty = 3, type = "h")
abline(h = c(0,probs, 1), v = c(0, 2*pi), col = "purple", lty = 3)
max(abs(pcard(Q, mu, rho) - probs)) # Should be 0

## End(Not run)
```

cardioid

*Cardioid Distribution Family Function***Description**

Estimates the two parameters of the cardioid distribution by maximum likelihood estimation.

Usage

```
cardioid(lmu = extlogitlink(min = 0, max = 2*pi),
         lrho = extlogitlink(min = -0.5, max = 0.5),
         imu = NULL, irho = 0.3, nsimEIM = 100, zero = NULL)
```

Arguments

`lmu, lrho` Parameter link functions applied to the μ and ρ parameters, respectively. See [Links](#) for more choices.

`imu, irho` Initial values. A NULL means an initial value is chosen internally. See [CommonVGAMffArguments](#) for more information.

`nsimEIM, zero` See [CommonVGAMffArguments](#) for more information.

Details

The two-parameter cardioid distribution has a density that can be written as

$$f(y; \mu, \rho) = \frac{1}{2\pi} (1 + 2\rho \cos(y - \mu))$$

where $0 < y < 2\pi$, $0 < \mu < 2\pi$, and $-0.5 < \rho < 0.5$ is the concentration parameter. The default link functions enforce the range constraints of the parameters.

For positive ρ the distribution is unimodal and symmetric about μ . The mean of Y (which make up the fitted values) is $\pi + (\rho/\pi)((2\pi - \mu) \sin(2\pi - \mu) + \cos(2\pi - \mu) - \mu \sin(\mu) - \cos(\mu))$.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), [rrvglm](#) and [vgam](#).

Warning

Numerically, this distribution can be difficult to fit because of a log-likelihood having multiple maximums. The user is therefore encouraged to try different starting values, i.e., make use of `imu` and `irho`.

Note

Fisher scoring using simulation is used.

Author(s)

T. W. Yee

References

Jammalamadaka, S. R. and SenGupta, A. (2001). *Topics in Circular Statistics*, Singapore: World Scientific.

See Also

[rcard](#), [extlogitlink](#), [vonmises](#).

CircStats and **circular** currently have a lot more R functions for circular data than the **VGAM** package.

Examples

```
## Not run:
cdata <- data.frame(y = rcard(n = 1000, mu = 4, rho = 0.45))
fit <- vglm(y ~ 1, cardioid, data = cdata, trace = TRUE)
coef(fit, matrix=TRUE)
Coef(fit)
c(with(cdata, mean(y)), head(fitted(fit), 1))
summary(fit)

## End(Not run)
```

`cauchitlink`*Cauchit Link Function*

Description

Computes the cauchit (tangent) link transformation, including its inverse and the first two derivatives.

Usage

```
cauchitlink(theta, bvalue = .Machine$double.eps,  
            inverse = FALSE, deriv = 0, short = TRUE, tag = FALSE)
```

Arguments

`theta` Numeric or character. See below for further details.
`bvalue` See [Links](#).
`inverse, deriv, short, tag`
Details at [Links](#).

Details

This link function is an alternative link function for parameters that lie in the unit interval. This type of link bears the same relation to the Cauchy distribution as the probit link bears to the Gaussian. One characteristic of this link function is that the tail is heavier relative to the other links (see examples below).

Numerical values of `theta` close to 0 or 1 or out of range result in `Inf`, `-Inf`, `NA` or `NaN`.

Value

For `deriv = 0`, the tangent of `theta`, i.e., $\tan(\pi * (\theta - 0.5))$ when `inverse = FALSE`, and if `inverse = TRUE` then $0.5 + \text{atan}(\theta) / \pi$.

For `deriv = 1`, then the function returns $d \eta / d \theta$ as a function of `theta` if `inverse = FALSE`, else if `inverse = TRUE` then it returns the reciprocal.

Note

Numerical instability may occur when `theta` is close to 1 or 0. One way of overcoming this is to use `bvalue`.

As mentioned above, in terms of the threshold approach with cumulative probabilities for an ordinal response this link function corresponds to the Cauchy distribution (see [cauchy1](#)).

Author(s)

Thomas W. Yee

References

McCullagh, P. and Nelder, J. A. (1989). *Generalized Linear Models*, 2nd ed. London: Chapman & Hall.

See Also

[logitlink](#), [probitlink](#), [clogloglink](#), [loglink](#), [cauchy](#), [cauchy1](#), [Cauchy](#).

Examples

```
p <- seq(0.01, 0.99, by = 0.01)
cauchitlink(p)
max(abs(cauchitlink(cauchitlink(p), inverse = TRUE) - p)) # Should be 0

p <- c(seq(-0.02, 0.02, by=0.01), seq(0.97, 1.02, by = 0.01))
cauchitlink(p) # Has no NAs

## Not run:
par(mfrow = c(2, 2), lwd = (mylwd <- 2))
y <- seq(-4, 4, length = 100)
p <- seq(0.01, 0.99, by = 0.01)

for (d in 0:1) {
  matplot(p, cbind(logitlink(p, deriv = d), probitlink(p, deriv = d)),
           type = "n", col = "purple", ylab = "transformation",
           las = 1, main = if (d == 0) "Some probability link functions"
           else "First derivative")
  lines(p, logitlink(p, deriv = d), col = "limegreen")
  lines(p, probitlink(p, deriv = d), col = "purple")
  lines(p, clogloglink(p, deriv = d), col = "chocolate")
  lines(p, cauchitlink(p, deriv = d), col = "tan")
  if (d == 0) {
    abline(v = 0.5, h = 0, lty = "dashed")
    legend(0, 4.5, c("logitlink", "probitlink", "clogloglink",
                    "cauchitlink"), lwd = mylwd,
           col = c("limegreen", "purple", "chocolate", "tan"))
  } else
    abline(v = 0.5, lty = "dashed")
}

for (d in 0) {
  matplot(y, cbind(logitlink(y, deriv = d, inverse = TRUE),
                  probitlink(y, deriv = d, inverse = TRUE)),
           type = "n", col = "purple", xlab = "transformation", ylab = "p",
           main = if (d == 0) "Some inverse probability link functions"
           else "First derivative", las=1)
  lines(y, logitlink(y, deriv = d, inverse = TRUE), col = "limegreen")
  lines(y, probitlink(y, deriv = d, inverse = TRUE), col = "purple")
  lines(y, clogloglink(y, deriv = d, inverse = TRUE), col = "chocolate")
  lines(y, cauchitlink(y, deriv = d, inverse = TRUE), col = "tan")
  if (d == 0) {
    abline(h = 0.5, v = 0, lty = "dashed")
  }
}
```

```

      legend(-4, 1, c("logitlink", "probitlink", "clogloglink",
                    "cauchitlink"), lwd = mylwd,
            col = c("limegreen", "purple", "chocolate", "tan"))
    }
  }
  par(lwd = 1)

## End(Not run)

```

cauchy

*Cauchy Distribution Family Function***Description**

Estimates either the location parameter or both the location and scale parameters of the Cauchy distribution by maximum likelihood estimation.

Usage

```

cauchy(llocation = "identitylink", lscale = "loglink",
       imethod = 1, ilocation = NULL, iscale = NULL,
       gprobs.y = ppoints(19), gscale.mux = exp(-3:3), zero = "scale")
cauchy1(scale.arg = 1, llocation = "identitylink", ilocation = NULL,
        imethod = 1, gprobs.y = ppoints(19), zero = NULL)

```

Arguments

<code>llocation, lscale</code>	Parameter link functions for the location parameter a and the scale parameter b . See Links for more choices.
<code>ilocation, iscale</code>	Optional initial value for a and b . By default, an initial value is chosen internally for each.
<code>imethod</code>	Integer, either 1 or 2 or 3. Initial method, three algorithms are implemented. The user should try all possible values to help avoid converging to a local solution. Also, choose the another value if convergence fails, or use <code>ilocation</code> and/or <code>iscale</code> .
<code>gprobs.y, gscale.mux, zero</code>	See CommonVGAMffArguments for information.
<code>scale.arg</code>	Known (positive) scale parameter, called b below.

Details

The Cauchy distribution has density function

$$f(y; a, b) = \{\pi b[1 + ((y - a)/b)^2]\}^{-1}$$

where y and a are real and finite, and $b > 0$. The distribution is symmetric about a and has a heavy tail. Its median and mode are a , but the mean does not exist. The fitted values are the estimates of a . Fisher scoring is used.

If the scale parameter is known (`cauchy1`) then there may be multiple local maximum likelihood solutions for the location parameter. However, if both location and scale parameters are to be estimated (`cauchy`) then there is a unique maximum likelihood solution provided $n > 2$ and less than half the data are located at any one point.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as `vglm`, and `vgam`.

Warning

It is well-known that the Cauchy distribution may have local maximums in its likelihood function; make full use of `imethod`, `ilocation`, `iscale` etc.

Note

Good initial values are needed. By default `cauchy` searches for a starting value for a and b on a 2-D grid. Likewise, by default, `cauchy1` searches for a starting value for a on a 1-D grid. If convergence to the global maximum is not achieved then it also pays to select a wide range of initial values via the `ilocation` and/or `iscale` and/or `imethod` arguments.

Author(s)

T. W. Yee

References

- Forbes, C., Evans, M., Hastings, N. and Peacock, B. (2011). *Statistical Distributions*, Hoboken, NJ, USA: John Wiley and Sons, Fourth edition.
- Barnett, V. D. (1966). Evaluation of the maximum-likelihood estimator where the likelihood equation has multiple roots. *Biometrika*, **53**, 151–165.
- Copas, J. B. (1975). On the unimodality of the likelihood for the Cauchy distribution. *Biometrika*, **62**, 701–704.
- Efron, B. and Hinkley, D. V. (1978). Assessing the accuracy of the maximum likelihood estimator: Observed versus expected Fisher information. *Biometrika*, **65**, 457–481.

See Also

[Cauchy](#), [cauchit](#), [studentt](#), [simulate.vlm](#).

Examples

```
# Both location and scale parameters unknown
set.seed(123)
cdata <- data.frame(x2 = runif(nn <- 1000))
cdata <- transform(cdata, loc = exp(1 + 0.5 * x2), scale = exp(1))
cdata <- transform(cdata, y2 = rcauchy(nn, loc, scale))
fit2 <- vglm(y2 ~ x2, cauchy(lloc = "loglink"), data = cdata)
coef(fit2, matrix = TRUE)
head(fitted(fit2)) # Location estimates
summary(fit2)

# Location parameter unknown
cdata <- transform(cdata, scale1 = 0.4)
cdata <- transform(cdata, y1 = rcauchy(nn, loc, scale1))
fit1 <- vglm(y1 ~ x2, cauchy1(scale = 0.4), data = cdata, trace = TRUE)
coef(fit1, matrix = TRUE)
```

cdf.lmscreg

*Cumulative Distribution Function for LMS Quantile Regression***Description**

Computes the cumulative distribution function (CDF) for observations, based on a LMS quantile regression.

Usage

```
cdf.lmscreg(object, newdata = NULL, ...)
```

Arguments

object	A VGAM quantile regression model, i.e., an object produced by modelling functions such as <code>vglm</code> and <code>vgam</code> with a family function beginning with "lms."
newdata	Data frame where the predictions are to be made. If missing, the original data is used.
...	Parameters which are passed into functions such as <code>cdf.lms.yjn</code> .

Details

The CDFs returned here are values lying in [0,1] giving the relative probabilities associated with the quantiles `newdata`. For example, a value near 0.75 means it is close to the upper quartile of the distribution.

Value

A vector of CDF values lying in [0,1].

Note

The data are treated like quantiles, and the percentiles are returned. The opposite is performed by [qtplot.lmscreg](#).

The CDF values of the model have been placed in `@post$cdf` when the model was fitted.

Author(s)

Thomas W. Yee

References

Yee, T. W. (2004). Quantile regression via vector generalized additive models. *Statistics in Medicine*, **23**, 2295–2315.

See Also

[deplot.lmscreg](#), [qtplot.lmscreg](#), [lms.bcn](#), [lms.bcg](#), [lms.yjn](#), [CommonVGAMffArguments](#).

Examples

```
fit <- vgam(BMI ~ s(age, df=c(4, 2)), lms.bcn(zero = 1), data = bmi.nz)
head(fit@post$cdf)
head(cdf(fit)) # Same
head(depvar(fit))
head(fitted(fit))

cdf(fit, data.frame(age = c(31.5, 39), BMI = c(28.4, 24)))
```

cens.gumbel

Censored Gumbel Distribution

Description

Maximum likelihood estimation of the 2-parameter Gumbel distribution when there are censored observations. A matrix response is not allowed.

Usage

```
cens.gumbel(llocation = "identitylink", lscale = "loglink",
            iscale = NULL, mean = TRUE, percentiles = NULL,
            zero = "scale")
```

Arguments

llocation, lscale	Character. Parameter link functions for the location and (positive) <i>scale</i> parameters. See Links for more choices.
iscale	Numeric and positive. Initial value for <i>scale</i> . Recycled to the appropriate length. In general, a larger value is better than a smaller value. The default is to choose the value internally.
mean	Logical. Return the mean? If TRUE then the mean is returned, otherwise percentiles given by the percentiles argument.
percentiles	Numeric with values between 0 and 100. If mean=FALSE then the fitted values are percentiles which must be specified by this argument.
zero	An integer-valued vector specifying which linear/additive predictors are modelled as intercepts only. The value (possibly values) must be from the set {1,2} corresponding respectively to <i>location</i> and <i>scale</i> . If zero=NULL then all linear/additive predictors are modelled as a linear combination of the explanatory variables. The default is to fit the shape parameter as an intercept only. See CommonVGAMffArguments for more information.

Details

This **VGAM** family function is like [gumbel](#) but handles observations that are left-censored (so that the true value would be less than the observed value) else right-censored (so that the true value would be greater than the observed value). To indicate which type of censoring, input `extra = list(leftcensored = vec1, rightcensored = vec2)` where `vec1` and `vec2` are logical vectors the same length as the response. If the two components of this list are missing then the logical values are taken to be FALSE. The fitted object has these two components stored in the `extra` slot.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Warning

Numerical problems may occur if the amount of censoring is excessive.

Note

See [gumbel](#) for details about the Gumbel distribution. The initial values are based on assuming all uncensored observations, therefore could be improved upon.

Author(s)

T. W. Yee

References

Coles, S. (2001). *An Introduction to Statistical Modeling of Extreme Values*. London: Springer-Verlag.

See Also

[gumbel](#), [gumbelff](#), [rgumbel](#), [guplot](#), [gev](#), [venice](#).

Examples

```
# Example 1
ystar <- venice[["r1"]] # Use the first order statistic as the response
nn <- length(ystar)
L <- runif(nn, 100, 104) # Lower censoring points
U <- runif(nn, 130, 135) # Upper censoring points
y <- pmax(L, ystar) # Left censored
y <- pmin(U, y) # Right censored
extra <- list(leftcensored = ystar < L, rightcensored = ystar > U)
fit <- vglm(y ~ scale(year), data = venice, trace = TRUE, extra = extra,
            fam = cens.gumbel(mean = FALSE, perc = c(5, 25, 50, 75, 95)))
coef(fit, matrix = TRUE)
head(fitted(fit))
fit@extra

# Example 2: simulated data
nn <- 1000
ystar <- rgumbel(nn, loc = 1, scale = exp(0.5)) # The uncensored data
L <- runif(nn, -1, 1) # Lower censoring points
U <- runif(nn, 2, 5) # Upper censoring points
y <- pmax(L, ystar) # Left censored
y <- pmin(U, y) # Right censored
## Not run: par(mfrow = c(1, 2)); hist(ystar); hist(y);
extra <- list(leftcensored = ystar < L, rightcensored = ystar > U)
fit <- vglm(y ~ 1, trace = TRUE, extra = extra, fam = cens.gumbel)
coef(fit, matrix = TRUE)
```

cens.normal

Censored Normal Distribution

Description

Maximum likelihood estimation for the normal distribution with left and right censoring.

Usage

```
cens.normal(lmu = "identitylink", lsd = "loglink", imethod = 1,
            zero = "sd")
```

Arguments

`lmu`, `lsd` Parameter link functions applied to the mean and standard deviation parameters. See [Links](#) for more choices. The standard deviation is a positive quantity, therefore a log link is the default.

imethod	Initialization method. Either 1 or 2, this specifies two methods for obtaining initial values for the parameters.
zero	A vector, e.g., containing the value 1 or 2; if so, the mean or standard deviation respectively are modelled as an intercept only. Setting zero = NULL means both linear/additive predictors are modelled as functions of the explanatory variables. See CommonVGAMffArguments for more information.

Details

This function is like [uninormal](#) but handles observations that are left-censored (so that the true value would be less than the observed value) else right-censored (so that the true value would be greater than the observed value). To indicate which type of censoring, input `extra = list(leftcensored = vec1, rightcensored = vec2)` where `vec1` and `vec2` are logical vectors the same length as the response. If the two components of this list are missing then the logical values are taken to be FALSE. The fitted object has these two components stored in the extra slot.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Note

This function, which is an alternative to [tobit](#), cannot handle a matrix response and uses different working weights. If there are no censored observations then [uninormal](#) is recommended instead.

Author(s)

T. W. Yee

See Also

[tobit](#), [uninormal](#), [double.cens.normal](#).

Examples

```
## Not run:
cdata <- data.frame(x2 = runif(nn <- 1000)) # ystar are true values
cdata <- transform(cdata, ystar = rnorm(nn, m = 100 + 15 * x2, sd = exp(3)))
with(cdata, hist(ystar))
cdata <- transform(cdata, L = runif(nn, 80, 90), # Lower censoring points
                  U = runif(nn, 130, 140)) # Upper censoring points
cdata <- transform(cdata, y = pmax(L, ystar)) # Left censored
cdata <- transform(cdata, y = pmin(U, y))    # Right censored
with(cdata, hist(y))
Extra <- list(leftcensored = with(cdata, ystar < L),
             rightcensored = with(cdata, ystar > U))
fit1 <- vglm(y ~ x2, cens.normal, data = cdata, crit = "c", extra = Extra)
fit2 <- vglm(y ~ x2, tobit(Lower = with(cdata, L), Upper = with(cdata, U)),
            data = cdata, crit = "c", trace = TRUE)
coef(fit1, matrix = TRUE)
```

```

max(abs(coef(fit1, matrix = TRUE) -
        coef(fit2, matrix = TRUE))) # Should be 0
names(fit1@extra)

## End(Not run)

```

cens.poisson

*Censored Poisson Family Function***Description**

Family function for a censored Poisson response.

Usage

```

cens.poisson(link = "loglink", imu = NULL,
             biglambda = 10, smallno = 1e-10)

```

Arguments

link	Link function applied to the mean; see Links for more choices.
imu	Optional initial value; see CommonVGAMffArguments for more information.
biglambda, smallno	Used to help robustify the code when lambda is very large and the ppois value is so close to 0 that the first derivative is computed to be a NA or NaN. When this occurs mills.ratio is called.

Details

Often a table of Poisson counts has an entry $J+$ meaning $\geq J$. This family function is similar to [poissonff](#) but handles such censored data. The input requires [SurvS4](#). Only a univariate response is allowed. The Newton-Raphson algorithm is used.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Warning

As the response is discrete, care is required with [Surv](#) (the old class because of `setOldClass(c("SurvS4", "Surv"))`; see [setOldClass](#)), especially with "interval" censored data because of the (start, end] format. See the examples below. The examples have $y < L$ as left censored and $y \geq U$ (formatted as $U+$) as right censored observations, therefore $L \leq y < U$ is for uncensored and/or interval censored observations. Consequently the input must be tweaked to conform to the (start, end] format.

A bit of attention has been directed to try robustify the code when lambda is very large, however this currently works for left and right censored data only, not interval censored data. Sometime the fix involves an approximation, hence it is a good idea to set `trace = TRUE`.

Note

The function `poissonff` should be used when there are no censored observations. Also, NAs are not permitted with `SurvS4`, nor is `type = "counting"`.

Author(s)

Thomas W. Yee

References

See **survival** for background.

See Also

[SurvS4](#), [poissonff](#), [Links](#), [mills.ratio](#).

Examples

```
# Example 1: right censored data
set.seed(123); U <- 20
cdata <- data.frame(y = rpois(N <- 100, exp(3)))
cdata <- transform(cdata, cy = pmin(U, y),
                  rcensored = (y >= U))
cdata <- transform(cdata, status = ifelse(rcensored, 0, 1))
with(cdata, table(cy))
with(cdata, table(rcensored))
with(cdata, table(print(SurvS4(cy, status)))) # Check; U+ means >= U
fit <- vglm(SurvS4(cy, status) ~ 1, cens.poisson, data = cdata,
           trace = TRUE)
coef(fit, matrix = TRUE)
table(print(depvar(fit))) # Another check; U+ means >= U

# Example 2: left censored data
L <- 15
cdata <- transform(cdata,
                  cY = pmax(L, y),
                  lcensored = y < L) # Note y < L, not cY == L or y <= L
cdata <- transform(cdata, status = ifelse(lcensored, 0, 1))
with(cdata, table(cY))
with(cdata, table(lcensored))
with(cdata, table(print(SurvS4(cY, status, type = "left")))) # Check
fit <- vglm(SurvS4(cY, status, type = "left") ~ 1, cens.poisson,
           data = cdata, trace = TRUE)
coef(fit, matrix = TRUE)

# Example 3: interval censored data
cdata <- transform(cdata, Lvec = rep(L, len = N),
                  Uvec = rep(U, len = N))
cdata <-
  transform(cdata,
           icensored = Lvec <= y & y < Uvec) # Not lcensored or rcensored
with(cdata, table(icensored))
```

```

cdata <- transform(cdata, status = rep(3, N)) # 3 == interval censored
cdata <- transform(cdata,
  status = ifelse(rcensored, 0, status)) # 0 means right censored
cdata <- transform(cdata,
  status = ifelse(lcensored, 2, status)) # 2 means left censored
# Have to adjust Lvec and Uvec because of the (start, end] format:
cdata$Lvec[with(cdata, icensored)] <- cdata$Lvec[with(cdata, icensored)]-1
cdata$Uvec[with(cdata, icensored)] <- cdata$Uvec[with(cdata, icensored)]-1
# Unchanged:
cdata$Lvec[with(cdata, lcensored)] <- cdata$Lvec[with(cdata, lcensored)]
cdata$Uvec[with(cdata, rcensored)] <- cdata$Uvec[with(cdata, rcensored)]
with(cdata, # Check
  table(ii <- print(SurvS4(Lvec, Uvec, status, type = "interval"))))
fit <- vglm(SurvS4(Lvec, Uvec, status, type = "interval") ~ 1,
  cens.poisson, data = cdata, trace = TRUE)
coef(fit, matrix = TRUE)
table(print(depvar(fit))) # Another check

# Example 4: Add in some uncensored observations
index <- (1:N)[with(cdata, icensored)]
index <- head(index, 4)
cdata$status[index] <- 1 # actual or uncensored value
cdata$Lvec[index] <- cdata$y[index]
with(cdata, table(ii <- print(SurvS4(Lvec, Uvec, status,
  type = "interval")))) # Check
fit <- vglm(SurvS4(Lvec, Uvec, status, type = "interval") ~ 1,
  cens.poisson, data = cdata, trace = TRUE, crit = "c")
coef(fit, matrix = TRUE)
table(print(depvar(fit))) # Another check

```

cfibrosis

Cystic Fibrosis Data

Description

This data frame concerns families data and cystic fibrosis.

Usage

```
data(cfibrosis)
```

Format

A data frame with 24 rows on the following 4 variables.

siblings, affected, ascertained, families Over ascertained families, the k th ascertained family has s_k siblings of whom r_k are affected and a_k are ascertained.

Details

The data set allows a classical segregation analysis to be performed. In particular, to test Mendelian segregation ratios in nuclear family data. The likelihood has similarities with [seq2binomial](#).

Source

The data is originally from Crow (1965) and appears as Table 2.3 of Lange (2002).

Crow, J. F. (1965) Problems of ascertainment in the analysis of family data. *Epidemiology and Genetics of Chronic Disease*. Public Health Service Publication 1163, Neel J. V., Shaw M. W., Schull W. J., editors, Department of Health, Education, and Welfare, Washington, DC, USA.

Lange, K. (2002) *Mathematical and Statistical Methods for Genetic Analysis*. Second Edition. Springer-Verlag: New York, USA.

Examples

```
cfibrosis
summary(cfibrosis)
```

cgo

Redirects the user to cgo

Description

Redirects the user to the function [cgo](#).

Usage

```
cgo(...)
```

Arguments

... Ignored.

Details

The former function `cgo` has been renamed `cgo` because CGO (for *canonical Gaussian ordination*) is a confusing and inaccurate name. CQO (for *constrained quadratic ordination*) is better. This new nomenclature described in Yee (2006).

Value

Nothing is returned; an error message is issued.

Warning

The code, therefore, in Yee (2004) will not run without changing the "g" to a "q".

Author(s)

Thomas W. Yee

References

Yee, T. W. (2004). A new technique for maximum-likelihood canonical Gaussian ordination. *Ecological Monographs*, **74**, 685–701.

Yee, T. W. (2006). Constrained additive ordination. *Ecology*, **87**, 203–213.

See Also

[cgo](#).

Examples

```
## Not run:  
cgo()  
  
## End(Not run)
```

chest.nz

Chest Pain in NZ Adults Data

Description

Presence/absence of chest pain in 10186 New Zealand adults.

Usage

```
data(chest.nz)
```

Format

A data frame with 73 rows and the following 5 variables.

age a numeric vector; age (years).

nolnor a numeric vector of counts; no pain on LHS or RHS.

nolr a numeric vector of counts; no pain on LHS but pain on RHS.

lnor a numeric vector of counts; no pain on RHS but pain on LHS.

lr a numeric vector of counts; pain on LHS and RHS of chest.

Details

Each adult was asked their age and whether they experienced any pain or discomfort in their chest over the last six months. If yes, they indicated whether it was on their LHS and/or RHS of their chest.

Source

MacMahon, S., Norton, R., Jackson, R., Mackie, M. J., Cheng, A., Vander Hoorn, S., Milne, A., McCulloch, A. (1995) Fletcher Challenge-University of Auckland Heart & Health Study: design and baseline findings. *New Zealand Medical Journal*, **108**, 499–502.

Examples

```
## Not run:
fit <- vgam(cbind(nolnor, nolr, lnor, lr) ~ s(age, c(4, 3)),
           binom2.or(exchan = TRUE, zero = NULL), data = chest.nz)
coef(fit, matrix = TRUE)

## End(Not run)
## Not run: plot(fit, which.cf = 2, se = TRUE)
```

chinese.nz

*Chinese Population in New Zealand 1867–2001 Data***Description**

The Chinese population in New Zealand from 1867 to 2001, along with the whole of the New Zealand population.

Usage

```
data(chinese.nz)
```

Format

A data frame with 27 observations on the following 4 variables.

year Year.

male Number of Chinese males.

female Number of Chinese females.

nz Total number in the New Zealand population.

Details

Historically, there was a large exodus of Chinese from the Guangdong region starting in the mid-1800s to the gold fields of South Island of New Zealand, California (a region near Mexico), and southern Australia, etc. Discrimination then meant that only men were allowed entry, to hinder permanent settlement. In the case of New Zealand, the government relaxed its immigration laws after WWII to allow wives of Chinese already in NZ to join them because China had been among the Allied powers. Gradual relaxation in the immigration and an influx during the 1980s meant the Chinese population became increasingly demographically normal over time.

The NZ total for the years 1867 and 1871 exclude the Maori population. Three modifications have been made to the female column to make the data internally consistent with the original table.

References

Page 6 of *Aliens At My Table: Asians as New Zealanders See Them* by M. Ip and N. Murphy, (2005). Penguin Books. Auckland, New Zealand.

Examples

```
## Not run: par(mfrow = c(1, 2))
plot(female / (male + female) ~ year, chinese.nz, type = "b",
     ylab = "Proportion", col = "blue", las = 1,
     cex = 0.015 * sqrt(male + female),
#   cex = 0.10 * sqrt((male + female)^1.5 / sqrt(female) / sqrt(male)),
     main = "Proportion of NZ Chinese that are female")
abline(h = 0.5, lty = "dashed", col = "gray")

fit1.cnz <- vglm(cbind(female, male) ~ year,          binomialff,
                data = chinese.nz)
fit2.cnz <- vglm(cbind(female, male) ~ sm.poly(year, 2), binomialff,
                data = chinese.nz)
fit4.cnz <- vglm(cbind(female, male) ~ sm.bs(year, 5), binomialff,
                data = chinese.nz)

lines(fitted(fit1.cnz) ~ year, chinese.nz, col = "purple", lty = 1)
lines(fitted(fit2.cnz) ~ year, chinese.nz, col = "green", lty = 2)
lines(fitted(fit4.cnz) ~ year, chinese.nz, col = "orange", lwd = 2, lty = 1)
legend("bottomright", col = c("purple", "green", "orange"),
      lty = c(1, 2, 1), leg = c("linear", "quadratic", "B-spline"))

plot(100*(male+female)/nz ~ year, chinese.nz, type = "b", ylab = "Percent",
     ylim = c(0, max(100*(male+female)/nz)), col = "blue", las = 1,
     main = "Percent of NZers that are Chinese")
abline(h = 0, lty = "dashed", col = "gray")
## End(Not run)
```

chisq

Chi-squared and Chi Distributions

Description

Maximum likelihood estimation of the degrees of freedom for a chi-squared distribution. Also fits the chi distribution.

Usage

```
chisq(link = "loglink", zero = NULL, squared = TRUE)
```

Arguments

link, zero See [CommonVGAMffArguments](#) for information.
squared Logical. Set FALSE for the chi distribution.

Details

The degrees of freedom is treated as a real parameter to be estimated and not as an integer. Being positive, a log link is used by default. Fisher scoring is used.

If a random variable has a chi-squared distribution then the square root of the random variable has a chi distribution. For both distributions, the fitted value is the mean.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Note

Multiple responses are permitted. There may be convergence problems if the degrees of freedom is very large or close to zero.

Author(s)

T. W. Yee

References

Forbes, C., Evans, M., Hastings, N. and Peacock, B. (2011). *Statistical Distributions*, Hoboken, NJ, USA: John Wiley and Sons, Fourth edition.

See Also

[Chisquare](#). [uninormal](#).

Examples

```
cdata <- data.frame(x2 = runif(nn <- 1000))
cdata <- transform(cdata, y1 = rchisq(nn, df = exp(1 - 1 * x2)),
                  y2 = rchisq(nn, df = exp(2 - 2 * x2)))
fit <- vglm(cbind(y1, y2) ~ x2, chisq, data = cdata, trace = TRUE)
coef(fit, matrix = TRUE)
```

clo

Redirects the User to rrvglm()

Description

Redirects the user to the function [rrvglm](#).

Usage

```
clo(...)
```

Arguments

... Ignored.

Details

CLO stands for *constrained linear ordination*, and is fitted with a statistical class of models called *reduced-rank vector generalized linear models* (RR-VGLMs). It allows for generalized reduced-rank regression in that response types such as Poisson counts and presence/absence data can be handled.

Currently in the **VGAM** package, [rrvglm](#) is used to fit RR-VGLMs. However, the Author's opinion is that linear responses to a latent variable (composite environmental gradient) is not as common as unimodal responses, therefore [cqo](#) is often more appropriate.

The new CLO/CQO/CAO nomenclature described in Yee (2006).

Value

Nothing is returned; an error message is issued.

Author(s)

Thomas W. Yee

References

Yee, T. W. (2006). Constrained additive ordination. *Ecology*, **87**, 203–213.

Yee, T. W. and Hastie, T. J. (2003). Reduced-rank vector generalized linear models. *Statistical Modelling*, **3**, 15–41.

See Also

[rrvglm](#), [cqo](#).

Examples

```
## Not run:  
clo()  
  
## End(Not run)
```

clogloglink

Complementary Log-log Link Function

Description

Computes the complementary log-log transformation, including its inverse and the first two derivatives. The complementary log transformation is also computed.

Usage

```
clogloglink(theta, bvalue = NULL, inverse = FALSE, deriv = 0,
             short = TRUE, tag = FALSE)
cloglink(theta, bvalue = NULL, inverse = FALSE, deriv = 0,
          short = TRUE, tag = FALSE)
```

Arguments

theta Numeric or character. See below for further details.

bvalue See [Links](#) for general information about links.

inverse, deriv, short, tag
 Details at [Links](#).

Details

The complementary log-log link function is commonly used for parameters that lie in the unit interval. But unlike [logitlink](#), [probitlink](#) and [cauchitlink](#), this link is not symmetric. It is the inverse CDF of the extreme value (or Gumbel or log-Weibull) distribution. Numerical values of theta close to 0 or 1 or out of range result in Inf, -Inf, NA or NaN.

The complementary log link function is the same as the complementary log-log but the outer log is omitted. This link is suitable for lrho in [betabinomial](#) because it handles probability-like parameters but also allows slight negative values in theory. In particular, [cloglink](#) safeguards against parameters exceeding unity.

Value

For deriv = 0, the complimentary log-log of theta, i.e., $\log(-\log(1 - \theta))$ when inverse = FALSE, and if inverse = TRUE then $1 - \exp(-\exp(\theta))$.

For deriv = 1, then the function returns $d \text{ eta} / d \text{ theta}$ as a function of theta if inverse = FALSE, else if inverse = TRUE then it returns the reciprocal.

Here, all logarithms are natural logarithms, i.e., to base e .

Note

Numerical instability may occur when theta is close to 1 or 0. One way of overcoming this is to use `bvalue`.

Changing 1s to 0s and 0s to 1s in the response means that effectively a loglog link is fitted. That is, transform y by $1 - y$. That's why only one of `clogloglink` and `logloglink` is written.

With constrained ordination (e.g., `cqo` and `cao`) used with `binomialff`, a complementary log-log link function is preferred over the default `logitlink`, for a good reason. See the example below.

In terms of the threshold approach with cumulative probabilities for an ordinal response this link function corresponds to the extreme value distribution.

Author(s)

Thomas W. Yee

References

McCullagh, P. and Nelder, J. A. (1989). *Generalized Linear Models*, 2nd ed. London: Chapman & Hall.

See Also

[Links](#), [logitoffsetlink](#), [logitlink](#), [probitlink](#), [cauchitlink](#), [pgumbel](#).

Examples

```
p <- seq(0.01, 0.99, by = 0.01)
clogloglink(p)
max(abs(clogloglink(clogloglink(p), inverse = TRUE) - p)) # Should be 0

p <- c(seq(-0.02, 0.02, by = 0.01), seq(0.97, 1.02, by = 0.01))
clogloglink(p) # Has NAs
clogloglink(p, bvalue = .Machine$double.eps) # Has no NAs

## Not run:
p <- seq(0.01, 0.99, by = 0.01)
plot(p, logitlink(p), type = "l", col = "limegreen", lwd = 2, las = 1,
     main = "Some probability link functions", ylab = "transformation")
lines(p, probitlink(p), col = "purple", lwd = 2)
lines(p, clogloglink(p), col = "chocolate", lwd = 2)
lines(p, cauchitlink(p), col = "tan", lwd = 2)
abline(v = 0.5, h = 0, lty = "dashed")
legend(0.1, 4, c("logitlink", "probitlink", "clogloglink", "cauchitlink"),
      col = c("limegreen", "purple", "chocolate", "tan"), lwd = 2)

## End(Not run)

## Not run:
# This example shows that clogloglink is preferred over logitlink
n <- 500; p <- 5; S <- 3; Rank <- 1 # Species packing model:
mydata <- rcqo(n, p, S, eq.tol = TRUE, es.opt = TRUE, eq.max = TRUE,
```

```

        family = "binomial", hi.abundance = 5, seed = 123,
        Rank = Rank)
fitc <- cqo(attr(mydata, "formula"), I.tol = TRUE, data = mydata,
        fam = binomialfff(multiple.responses = TRUE, link = "cloglog"),
        Rank = Rank)
fitl <- cqo(attr(mydata, "formula"), I.tol = TRUE, data = mydata,
        fam = binomialfff(multiple.responses = TRUE, link = "logitlink"),
        Rank = Rank)

# Compare the fitted models (cols 1 and 3) with the truth (col 2)
cbind(concoef(fitc), attr(mydata, "concoefficients"), concoef(fitl))

## End(Not run)

```

CM.equid

Constraint Matrices for Symmetry, Order, Parallelism, etc.

Description

Given M linear/additive predictors, construct the constraint matrices to allow symmetry, (linear and normal) ordering, etc. in terms such as the intercept.

Usage

```

CM.equid(M, Trev = FALSE, Tref = 1)
CM.free(M, Trev = FALSE, Tref = 1)
CM.ones(M, Trev = FALSE, Tref = 1)
CM.symm0(M, Trev = FALSE, Tref = 1)
CM.symm1(M, Trev = FALSE, Tref = 1)
CM.qnorm(M, Trev = FALSE, Tref = 1)

```

Arguments

M	Number of linear/additive predictors, usually > 1 .
Tref	Reference level for the threshold, this should be a single value from $1:M$. This argument is ignored by some of the above functions.
Trev	Logical. Apply reverse direction for the thresholds direction? This argument is ignored by some of the above functions.

Details

A constraint matrix is $M \times R$ where R is its rank and usually the elements are 0, 1 or -1 . There is a constraint matrix for each column of the LM matrix used to fit the `vglm`. They are used to apportion the regression coefficients to the linear predictors, e.g., parallelism, exchangeability, etc. The functions described here are intended to construct constraint matrices easily for symmetry constraints and linear ordering etc. They are potentially useful for categorical data analysis (e.g., `cumulative`, `multinomial`), especially for the intercept term. When applied to `cumulative`, they are sometimes called *structured thresholds*, e.g., **ordinal**.

One example is the stereotype model proposed by Anderson (1984) (see [multinomial](#) and [rrvglm](#)) where the elements of the **A** matrix are ordered. This is not fully possible in **VGAM** but some special cases can be fitted, e.g., use [CM.equid](#) to create a linear ordering. And [CM.symm1](#) might result in fully ordered estimates too, etc.

[CM.free](#) creates *free* or unconstrained estimates. It is almost always the case for VGLMs, and is simply `diag(M)`.

[CM.ones](#) creates *equal* estimates, which is also known as the *parallelism* assumption in models such as [cumulative](#). It gets its name because the constraint matrix is simply `matrix(1, M, 1)`.

[CM.equid](#) creates *equidistant* estimates. This is a linear scaling, and the direction and origin are controlled by `Treverse` and `Tref` respectively.

[CM.qnorm](#) and [CM.qlogis](#) are based on [qnorm](#) and [qlogis](#). For example, `CM.qnorm(M)` is essentially `cbind(qnorm(seq(M) / (M + 1)))`. This might be useful with a model with [probitlink](#) applied to multiple intercepts.

Further details can be found at [cumulative](#) and [CommonVGAMffArguments](#),

Value

A constraint matrix.

See Also

[CommonVGAMffArguments](#), [cumulative](#), [acat](#), [cratio](#), [sratio](#), [multinomial](#).

Examples

```
CM.equid(4)
CM.equid(4, Trev = TRUE, Tref = 3)
CM.symm1(5)
CM.symm0(5)
CM.qnorm(5)
```

coalminers

Breathlessness and Wheeze Amongst Coalminers Data

Description

Coalminers who are smokers without radiological pneumoconiosis, classified by age, breathlessness and wheeze.

Usage

```
data(coalminers)
```

Format

A data frame with 9 age groups with the following 5 columns.

BW Counts with breathlessness and wheeze.

BnW Counts with breathlessness but no wheeze.

nBW Counts with no breathlessness but wheeze.

nBnW Counts with neither breathlessness or wheeze.

age Age of the coal miners (actually, the midpoints of the 5-year category ranges).

Details

The data were published in Ashford and Sowden (1970). A more recent analysis is McCullagh and Nelder (1989, Section 6.6).

Source

Ashford, J. R. and Sowden, R. R. (1970) Multi-variate probit analysis. *Biometrics*, **26**, 535–546.

References

McCullagh, P. and Nelder, J. A. (1989). *Generalized Linear Models*. 2nd ed. London: Chapman & Hall.

Examples

```
str(coalminers)
```

Coef

Computes Model Coefficients and Quantities

Description

Coef is a generic function which computes model coefficients from objects returned by modelling functions. It is an auxiliary function to `coef` that enables extra capabilities for some specific models.

Usage

```
Coef(object, ...)
```

Arguments

`object` An object for which the computation of other types of model coefficients or quantities is meaningful.

`...` Other arguments fed into the specific methods function of the model.

Details

This function can often be useful for `vglm` objects with just an intercept term in the RHS of the formula, e.g., $y \sim 1$. Then often this function will apply the inverse link functions to the parameters. See the example below.

For reduced-rank VGLMs, this function can return the **A**, **C** matrices, etc.

For quadratic and additive ordination models, this function can return ecological meaningful quantities such as tolerances, optimums, maximums.

Value

The value returned depends specifically on the methods function invoked.

Warning

This function may not work for *all* **VGAM** family functions. You should check your results on some artificial data before applying it to models fitted to real data.

Author(s)

Thomas W. Yee

References

Yee, T. W. and Hastie, T. J. (2003). Reduced-rank vector generalized linear models. *Statistical Modelling*, **3**, 15–41.

See Also

[coef](#), [Coef.vlm](#), [Coef.rrvglm](#), [Coef.qrrvglm](#), [depvar](#).

Examples

```
nn <- 1000
bdata <- data.frame(y = rbeta(nn, shape1 = 1, shape2 = 3)) # Original scale
fit <- vglm(y ~ 1, betaR, data = bdata, trace = TRUE) # Intercept-only model
coef(fit, matrix = TRUE) # Both on a log scale
Coef(fit) # On the original scale
```

Coef.qrrvglm

Returns Important Matrices etc. of a QO Object

Description

This methods function returns important matrices etc. of a QO object.

Usage

```
Coef.qrrvglm(object, varI.latvar = FALSE, refResponse = NULL, ...)
```

Arguments

object	A CQO object. The former has class "qrrvglm".
varI.latvar	Logical indicating whether to scale the site scores (latent variables) to have variance-covariance matrix equal to the rank- R identity matrix. All models have uncorrelated site scores (latent variables), and this option stretches or shrinks the ordination axes if TRUE. See below for further details.
refResponse	Integer or character. Specifies the <i>reference response</i> or <i>reference species</i> . By default, the reference species is found by searching sequentially starting from the first species until a positive-definite tolerance matrix is found. Then this tolerance matrix is transformed to the identity matrix. Then the sites scores (latent variables) are made uncorrelated. See below for further details.
...	Currently unused.

Details

If `I.tolerances=TRUE` or `eq.tolerances=TRUE` (and its estimated tolerance matrix is positive-definite) then all species' tolerances are unity by transformation or by definition, and the spread of the site scores can be compared to them. Vice versa, if one wishes to compare the tolerances with the sites score variability then setting `varI.latvar=TRUE` is more appropriate.

For rank-2 QRR-VGLMs, one of the species can be chosen so that the angle of its major axis and minor axis is zero, i.e., parallel to the ordination axes. This means the effect on the latent vars is independent on that species, and that its tolerance matrix is diagonal. The argument `refResponse` allows one to choose which is the reference species, which must have a positive-definite tolerance matrix, i.e., is bell-shaped. If `refResponse` is not specified, then the code will try to choose some reference species starting from the first species. Although the `refResponse` argument could possibly be offered as an option when fitting the model, it is currently available after fitting the model, e.g., in the functions `Coef.qrrvglm` and `lvplot.qrrvglm`.

Value

The **A**, **B1**, **C**, **T**, **D** matrices/arrays are returned, along with other slots. The returned object has class "Coef.qrrvglm" (see `Coef.qrrvglm-class`).

Note

Consider an equal-tolerances Poisson/binomial CQO model with $\text{noRRR} = \sim 1$. For $R = 1$ it has about $2S + p_2$ parameters. For $R = 2$ it has about $3S + 2p_2$ parameters. Here, S is the number of species, and $p_2 = p - 1$ is the number of environmental variables making up the latent variable. For an unequal-tolerances Poisson/binomial CQO model with $\text{noRRR} = \sim 1$, it has about $3S - 1 + p_2$ parameters for $R = 1$, and about $6S - 3 + 2p_2$ parameters for $R = 2$. Since the total number of data points is nS , where n is the number of sites, it pays to divide the number of data points by the number of parameters to get some idea about how much information the parameters contain.

Author(s)

Thomas W. Yee

References

Yee, T. W. (2004). A new technique for maximum-likelihood canonical Gaussian ordination. *Ecological Monographs*, **74**, 685–701.

Yee, T. W. (2006). Constrained additive ordination. *Ecology*, **87**, 203–213.

See Also

[cqo](#), [Coef.qrrvglm-class](#), [print.Coef.qrrvglm](#), [lvplot.qrrvglm](#).

Examples

```
set.seed(123)
x2 <- rnorm(n <- 100)
x3 <- rnorm(n)
x4 <- rnorm(n)
latvar1 <- 0 + x3 - 2*x4
lambda1 <- exp(3 - 0.5 * (latvar1-0)^2)
lambda2 <- exp(2 - 0.5 * (latvar1-1)^2)
lambda3 <- exp(2 - 0.5 * ((latvar1+4)/2)^2) # Unequal tolerances
y1 <- rpois(n, lambda1)
y2 <- rpois(n, lambda2)
y3 <- rpois(n, lambda3)
set.seed(111)
# vvv p1 <- cqo(cbind(y1, y2, y3) ~ x2 + x3 + x4, poissonff, trace = FALSE)
## Not run: lvplot(p1, y = TRUE, lcol = 1:3, pch = 1:3, pcol = 1:3)

# vvv Coef(p1)
# vvv print(Coef(p1), digits=3)
```

Coef.qrrvglm-class *Class "Coef.qrrvglm"*

Description

The most pertinent matrices and other quantities pertaining to a QRR-VGLM (CQO model).

Objects from the Class

Objects can be created by calls of the form `Coef(object, ...)` where `object` is an object of class `"qrrvglm"` (created by [cqo](#)).

In this document, R is the *rank*, M is the number of linear predictors and n is the number of observations.

Slots

- A:** Of class "matrix", **A**, which are the linear 'coefficients' of the matrix of latent variables. It is M by R .
- B1:** Of class "matrix", **B1**. These correspond to terms of the argument noRRR.
- C:** Of class "matrix", **C**, the canonical coefficients. It has R columns.
- Constrained:** Logical. Whether the model is a constrained ordination model.
- D:** Of class "array", $D[, , j]$ is an order-Rank matrix, for $j = 1, \dots, M$. Ideally, these are negative-definite in order to make the response curves/surfaces bell-shaped.
- Rank:** The rank (dimension, number of latent variables) of the RR-VGLM. Called R .
- latvar:** n by R matrix of latent variable values.
- latvar.order:** Of class "matrix", the permutation returned when the function `order` is applied to each column of `latvar`. This enables each column of `latvar` to be easily sorted.
- Maximum:** Of class "numeric", the M maximum fitted values. That is, the fitted values at the optimums for noRRR = ~ 1 models. If noRRR is not ~ 1 then these will be NAs.
- NOS:** Number of species.
- Optimum:** Of class "matrix", the values of the latent variables where the optimums are. If the curves are not bell-shaped, then the value will be NA or NaN.
- Optimum.order:** Of class "matrix", the permutation returned when the function `order` is applied to each column of `Optimum`. This enables each row of `Optimum` to be easily sorted.
- bellshaped:** Vector of logicals: is each response curve/surface bell-shaped?
- dispersion:** Dispersion parameter(s).
- Dzero:** Vector of logicals, is each of the response curves linear in the latent variable(s)? It will be if and only if $D[, , j]$ equals **O**, for $j = 1, \dots, M$.
- Tolerance:** Object of class "array", $Tolerance[, , j]$ is an order-Rank matrix, for $j = 1, \dots, M$, being the matrix of tolerances (squared if on the diagonal). These are denoted by **T** in Yee (2004). Ideally, these are positive-definite in order to make the response curves/surfaces bell-shaped. The tolerance matrices satisfy $T_s = -\frac{1}{2}D_s^{-1}$.

Author(s)

Thomas W. Yee

References

Yee, T. W. (2004). A new technique for maximum-likelihood canonical Gaussian ordination. *Ecological Monographs*, **74**, 685–701.

See Also

`Coef.qrrvglm`, `cqo`, `print.Coef.qrrvglm`.

Examples

```

x2 <- rnorm(n <- 100)
x3 <- rnorm(n)
x4 <- rnorm(n)
latvar1 <- 0 + x3 - 2*x4
lambda1 <- exp(3 - 0.5 * ( latvar1-0)^2)
lambda2 <- exp(2 - 0.5 * ( latvar1-1)^2)
lambda3 <- exp(2 - 0.5 * ((latvar1+4)/2)^2)
y1 <- rpois(n, lambda1)
y2 <- rpois(n, lambda2)
y3 <- rpois(n, lambda3)
yy <- cbind(y1, y2, y3)
# vvv p1 <- cqq(yy ~ x2 + x3 + x4, fam = poissonff, trace = FALSE)
## Not run:
lvplot(p1, y = TRUE, lcol = 1:3, pch = 1:3, pcol = 1:3)

## End(Not run)
# vvv print(Coef(p1), digits = 3)

```

Coef.rrvglm

Returns Important Matrices etc. of a RR-VGLM Object

Description

This methods function returns important matrices etc. of a RR-VGLM object.

Usage

```
Coef.rrvglm(object, ...)
```

Arguments

object	An object of class "rrvglm".
...	Currently unused.

Details

The **A**, **B1**, **C** matrices are returned, along with other slots. See [rrvglm](#) for details about RR-VGLMs.

Value

An object of class "Coef.rrvglm" (see [Coef.rrvglm-class](#)).

Note

This function is an alternative to `coef.rrvglm`.

Author(s)

Thomas W. Yee

References

Yee, T. W. and Hastie, T. J. (2003). Reduced-rank vector generalized linear models. *Statistical Modelling*, **3**, 15–41.

See Also

[Coef.rrvglm-class](#), [print.Coeff.rrvglm](#), [rrvglm](#).

Examples

```
# Rank-1 stereotype model of Anderson (1984)
pneumo <- transform(pneumo, let = log(exposure.time), x3 = runif(nrow(pneumo)))
fit <- rrvglm(cbind(normal, mild, severe) ~ let + x3, multinomial, data = pneumo)
coef(fit, matrix = TRUE)
Coef(fit)
```

Coef.rrvglm-class *Class "Coef.rrvglm"*

Description

The most pertinent matrices and other quantities pertaining to a RR-VGLM.

Objects from the Class

Objects can be created by calls of the form `Coef(object, ...)` where `object` is an object of class `rrvglm` (see [rrvglm-class](#)).

In this document, M is the number of linear predictors and n is the number of observations.

Slots

A: Of class "matrix", **A**.

B1: Of class "matrix", **B1**.

C: Of class "matrix", **C**.

Rank: The rank of the RR-VGLM.

colx1.index: Index of the columns of the "vlm"-type model matrix corresponding to the variables in **x1**. These correspond to **B1**.

colx2.index: Index of the columns of the "vlm"-type model matrix corresponding to the variables in **x2**. These correspond to the reduced-rank regression.

Atilde: Object of class "matrix", the **A** matrix with the corner rows removed. Thus each of the elements have been estimated. This matrix is returned only if corner constraints were used.

Author(s)

Thomas W. Yee

References

Yee, T. W. and Hastie, T. J. (2003). Reduced-rank vector generalized linear models. *Statistical Modelling*, **3**, 15–41.

See Also

[Coef.rrvglm](#), [rrvglm](#), [rrvglm-class](#), [print.Coeff.rrvglm](#).

Examples

```
# Rank-1 stereotype model of Anderson (1984)
pneumo <- transform(pneumo, let = log(exposure.time), x3 = runif(nrow(pneumo)))
fit <- rrvglm(cbind(normal, mild, severe) ~ let + x3, multinomial, data = pneumo)
coef(fit, matrix = TRUE)
Coef(fit)
# print(Coef(fit), digits = 3)
```

Coef.vlm

Extract Model Coefficients for VLM Objects

Description

Amongst other things, this function applies inverse link functions to the parameters of intercept-only VGLMs.

Usage

```
Coef.vlm(object, ...)
```

Arguments

object	A fitted model.
...	Arguments which may be passed into coef .

Details

Most **VGAM** family functions apply a link function to the parameters, e.g., positive parameter are often have a log link, parameters between 0 and 1 have a logit link. This function can back-transform the parameter estimate to the original scale.

Value

For intercept-only models (e.g., formula is $y \sim 1$) the back-transformed parameter estimates can be returned.

Warning

This function may not work for *all* **VGAM** family functions. You should check your results on some artificial data before applying it to models fitted to real data.

Author(s)

Thomas W. Yee

References

Yee, T. W. and Hastie, T. J. (2003). Reduced-rank vector generalized linear models. *Statistical Modelling*, **3**, 15–41.

See Also

[Coef](#), [coef](#).

Examples

```
set.seed(123); nn <- 1000
bdata <- data.frame(y = rbeta(nn, shape1 = 1, shape2 = 3))
fit <- vglm(y ~ 1, betaff, data = bdata, trace = TRUE) # intercept-only model
coef(fit, matrix = TRUE) # log scale
Coef(fit) # On the original scale
```

coefvgam

Extract Model Coefficients of a vgam() Object

Description

Extracts the estimated coefficients from vgam() objects.

Usage

```
coefvgam(object, type = c("linear", "nonlinear"), ...)
```

Arguments

object	A vgam object.
type	Character. The default is the first choice.
...	Optional arguments fed into coefvglm .

Details

For VGAMs, because modified backfitting is performed, each fitted function is decomposed into a linear and nonlinear (smooth) part. The argument type is used to return which one is wanted.

Value

A vector if `type = "linear"`. A list if `type = "nonlinear"`, and each component of this list corresponds to an `s` term; the component contains an S4 object with slot names such as "Bcoefficients", "knots", "xmin", "xmax".

Author(s)

Thomas W. Yee

See Also

[vgam](#), [coefvlm](#), [coef](#).

Examples

```
fit <- vgam(agaaus ~ s(altitude, df = 2), binomialff, data = hunua)
coef(fit) # Same as coef(fit, type = "linear")
(ii <- coef(fit, type = "nonlinear"))
is.list(ii)
names(ii)
slotNames(ii[[1]])
```

coefvlm

Extract Model Coefficients

Description

Extracts the estimated coefficients from VLM objects such as VGLMs.

Usage

```
coefvlm(object, matrix.out = FALSE, label = TRUE, colon = FALSE, ...)
```

Arguments

<code>object</code>	An object for which the extraction of coefficients is meaningful. This will usually be a vglm object.
<code>matrix.out</code>	Logical. If TRUE then a matrix is returned. The explanatory variables are the rows. The linear/additive predictors are the columns. The constraint matrices are used to compute this matrix.
<code>label</code>	Logical. If FALSE then the names of the vector of coefficients are set to NULL.
<code>colon</code>	Logical. Explanatory variables which appear in more than one linear/additive predictor are labelled with a colon, e.g., <code>age:1</code> , <code>age:2</code> . However, if it only appears in one linear/additive predictor then the <code>:1</code> is omitted by default. Then setting <code>colon = TRUE</code> will add the <code>:1</code> .
<code>...</code>	Currently unused.

Details

This function works in a similar way to applying `coef()` to a `lm` or `glm` object. However, for VGLMs, there are more options available.

Value

A vector usually. A matrix if `matrix.out = TRUE`.

Author(s)

Thomas W. Yee

References

Yee, T. W. and Hastie, T. J. (2003). Reduced-rank vector generalized linear models. *Statistical Modelling*, **3**, 15–41.

See Also

[vglm](#), [coefvgam](#), [coef](#).

Examples

```
zdata <- data.frame(x2 = runif(nn <- 200))
zdata <- transform(zdata, pstr0 = logitlink(-0.5 + 1*x2, inverse = TRUE),
                  lambda = loglink( 0.5 + 2*x2, inverse = TRUE))
zdata <- transform(zdata, y2 = rzipois(nn, lambda, pstr0 = pstr0))

fit2 <- vglm(y2 ~ x2, zipoisson(zero = 1), data = zdata, trace = TRUE)
coef(fit2, matrix = TRUE) # Always a good idea
coef(fit2)
coef(fit2, colon = TRUE)
```

CommonVGAMffArguments *Common VGAM Family Function Arguments*

Description

Here is a description of some common and typical arguments found in many **VGAM** family functions, e.g., `zero`, `lsigma`, `isigma`, `gsigma`, `eq.mean`, `nsimEI` and `parallel`.

Usage

```
TypicalVGAMfamilyFunction(lsigma = "loglink",
                          isigma = NULL,
                          zero = NULL, gsigma = exp(-5:5),
                          eq.mean = FALSE,
                          parallel = TRUE,
```

```

imethod = 1,
vfl = FALSE, Form2 = NULL,
type.fitted = c("mean", "quantiles", "Qlink",
               "pobs0", "pstr0", "onempstr0"),
percentiles = c(25, 50, 75),
probs.x = c(0.15, 0.85),
probs.y = c(0.25, 0.50, 0.75),
multiple.responses = FALSE, earg.link = FALSE,
ishrinkage = 0.95, nointercept = NULL,
whitespace = FALSE, bred = FALSE, lss = TRUE,
oim = FALSE, nsimEIM = 100, byrow.arg = FALSE,
link.list = list("(Default)" = "identitylink",
                 x2          = "loglink",
                 x3          = "logofflink",
                 x4          = "multilogitlink",
                 x5          = "multilogitlink"),
earg.list = list("(Default)" = list(),
                 x2          = list(),
                 x3          = list(offset = -1),
                 x4          = list(),
                 x5          = list()),
Thresh = NULL, nrfs = 1)

```

Arguments

- lsigma** Character. Link function applied to a parameter and not necessarily a mean. See [Links](#) for a selection of choices. If there is only one parameter then this argument is often called link.
- isigma** Optional initial values can often be inputted using an argument beginning with "i". For example, "isigma" and "ilocation", or just "init" if there is one parameter. A value of NULL means a value is computed internally, i.e., a *self-starting* VGAM family function. If a failure to converge occurs make use of these types of arguments.
- zero** An important argument, either an integer vector, or a vector of character strings. If an integer, then it specifies which linear/additive predictor is modelled as *intercept-only*. That is, the regression coefficients are set to zero for all covariates except for the intercept. If zero is specified then it may be a vector with values from the set $\{1, 2, \dots, M\}$. The value zero = NULL means model *all* linear/additive predictors as functions of the explanatory variables. Here, M is the number of linear/additive predictors. Technically, if zero contains the value j then the j th row of every constraint matrix (except for the intercept) consists of all 0 values.
- Some VGAM family functions allow the zero argument to accept negative values; if so then its absolute value is recycled over each (usual) response. For example, zero = -2 for the two-parameter negative binomial distribution would mean, for each response, the second linear/additive predictor is modelled as intercepts-only. That is, for all the k parameters in [negbinomial](#) (this VGAM family function can handle a matrix of responses).

Suppose `zero = zerovec` where `zerovec` is a vector of negative values. If G is the usual M value for a univariate response then the actual values for argument `zero` are all values in `c(abs(zerovec), G + abs(zerovec), 2*G + abs(zerovec), ...)` lying in the integer range 1 to M . For example, setting `zero = -c(2, 3)` for a matrix response of 4 columns with `zinegbinomial` (which usually has $G = M = 3$ for a univariate response) would be equivalent to `zero = c(2, 3, 5, 6, 8, 9, 11, 12)`. This example has $M = 12$. Note that if `zerovec` contains negative values then their absolute values should be elements from the set `1:G`.

Note: `zero` may have positive and negative values, for example, setting `zero = c(-2, 3)` in the above example would be equivalent to `zero = c(2, 3, 5, 8, 11)`.

The argument `zero` also accepts a character vector (for **VGAM** 1.0-1 onwards). Each value is fed into `grep` with `fixed = TRUE`, meaning that wildcards `"*"` are not useful. See the example below—all the variants work; those with `LOCAT` issue a warning that that value is unmatched. Importantly, the parameter names are `c("location1", "scale1", "location2", "scale2")` because there are 2 responses. Yee (2015) described `zero` for only numerical input. Allowing character input is particularly important when the number of parameters cannot be determined without having the actual data first. For example, with time series data, an $ARMA(p, q)$ process might have parameters $\theta_1, \dots, \theta_p$ which should be intercept-only by default. Then specifying a numerical default value for `zero` would be too difficult (there are the drift and scale parameters too). However, it is possible with the character representation: `zero = "theta"` would achieve this. In the future, most **VGAM** family functions might be converted to the character representation—the advantage being that it is more readable. When programming a **VGAM** family function that allows character input, the variable `predictors.names` must be assigned correctly.

If the `constraints` argument is used then the family function's `zero` argument (if it exists) needs to be set to `NULL`. This avoids what could be a probable contradiction. Sometimes setting other arguments related to constraint matrices to `FALSE` is also a good idea, e.g., `parallel = FALSE`, `exchangeable = FALSE`.

`gsigma`

Grid-search initial values can be inputted using an argument beginning with `"g"`, e.g., `"gsigma"`, `"gshape"` and `"gscale"`. If argument `isigma` is inputted then that has precedence over `gsigma`, etc. If the grid search is 2-dimensional then it is advisable not to make the vectors too long as a nested for loop may be used. Ditto for 3-dimensions etc. Sometimes a `".mux"` is added as a suffix, e.g., `gshape.mux`; this means that the grid is created relatively and not absolutely, e.g., its values are multiplied by some single initial estimate of the parameter in order to create the grid on an absolute scale.

Some family functions have an argument called `gprobs.y`. This is fed into the `probs` argument of `quantile` in order to obtain some values of central tendency of the response, i.e., some spread of values in the middle. when `imethod = 1` to obtain an initial value for the mean. Some family functions have an argument called `iprobs.y`, and if so, then these values can overwrite `gprobs.y`.

`eq.mean`

Logical. Constrain all the means to be equal? This type of argument is simpler than `parallel` because only a single `TRUE` or `FALSE` can be assigned and not a formula. Thus if `TRUE` then it will be enforced over all variables.

- parallel** A logical, or a simple formula specifying which terms have equal/unequal coefficients. The formula must be simple, i.e., additive with simple main effects terms. Interactions and nesting etc. are not handled. To handle complex formulas use the `constraints` argument (of `vglm` etc.); however, there is a lot more setting up involved and things will not be as convenient.
- Here are some examples. 1. `parallel = TRUE ~ x2 + x5` means the parallelism assumption is only applied to X_2 , X_5 and the intercept. 2. `parallel = TRUE ~ -1` and `parallel = TRUE ~ 0` mean the parallelism assumption is applied to *no* variables at all. Similarly, `parallel = FALSE ~ -1` and `parallel = FALSE ~ 0` mean the parallelism assumption is applied to *all* the variables including the intercept. 3. `parallel = FALSE ~ x2 - 1` and `parallel = FALSE ~ x2 + 0` applies the parallelism constraint to all terms (including the intercept) except for X_2 . 4. `parallel = FALSE ~ x2 * x3` probably will not work. Instead, expand it out manually to get `parallel = FALSE ~ x2 + x3 + x2:x3`, and that should work. That's because the main formula processes or expands the "*" operator but `parallel` does not. 5. To check whether `parallel` has done what was expected, type `coef(fit, matrix = TRUE)` or `constraints(fit)` for confirmation.
- This argument is common in **VGAM** family functions for categorical responses, e.g., `cumulative`, `acat`, `cratio`, `sratio`. For the proportional odds model (`cumulative`) having parallel constraints applied to each explanatory variable (except for the intercepts) means the fitted probabilities do not become negative or greater than 1. However this parallelism or proportional-odds assumption ought to be checked.
- nsimEIM** Some **VGAM** family functions use simulation to obtain an approximate expected information matrix (EIM). For those that do, the `nsimEIM` argument specifies the number of random variates used per observation; the mean of `nsimEIM` random variates is taken. Thus `nsimEIM` controls the accuracy and a larger value may be necessary if the EIMs are not positive-definite. For intercept-only models ($y \sim 1$) the value of `nsimEIM` can be smaller (since the common value used is also then taken as the mean over the observations), especially if the number of observations is large.
- Some **VGAM** family functions provide two algorithms for estimating the EIM. If applicable, set `nsimEIM = NULL` to choose the other algorithm.
- imethod** An integer with value 1 or 2 or 3 or ... which specifies the initialization method for some parameters or a specific parameter. If failure to converge occurs try the next higher value, and continue until success. For example, `imethod = 1` might be the method of moments, and `imethod = 2` might be another method. If no value of `imethod` works then it will be necessary to use arguments such as `isigma`. For many **VGAM** family functions it is advisable to try this argument with all possible values to safeguard against problems such as converging to a local solution. **VGAM** family functions with this argument usually correspond to a model or distribution that is relatively hard to fit successfully, therefore care is needed to ensure the global solution is obtained. So using all possible values that this argument supplies is a good idea.
- VGAM** family functions such `genpoisson2` recycle `imethod` to be of length 2 corresponding to the 2 parameters. In the future, this feature will be extended to other family functions to confer more flexibility.

Form2	<p>Formula. Using applied to models with $M = 2$. Specifies the terms for η_2 and the other terms belong to η_1. It is a way to partition the \mathbf{X} matrix into two sets of covariates, where they are assigned to each η_j separately. This argument sets up constraint matrices <code>rbind(0, 1)</code> for terms in Form2 and <code>rbind(1, 0)</code> for <code>setdiff(formula, Form2)</code> so to speak. Note that sometimes this argument is only accessed if <code>vf1 = TRUE</code>. Arguments such as Form1 and Form3 are also possible in VGAM family functions because the η_j which is likely to be modelled more simply is chosen for convenience.</p>
vf1	<p>A single logical. This stands for <i>variance–variance factored loglinear</i> (VFL) model. If TRUE then usually some other argument such as Form2 or <code>parallel</code> is used to partition the main <code>vglm</code> formula into two sets of covariates. For some families such as <code>negbinomial</code> this enables overdispersion to be modelled conveniently via a loglinear model, given the mean. It is necessary to read the online help regarding each VGAM family function because each one may differ from others. To fit some VFL models it is necessary to make a copy of existing covariates by creating new variable names and then adding it to the main formula.</p> <p>A good question is: why is <code>vf1</code> necessary? Wouldn't Form2 be sufficient? Setting <code>vf1 = TRUE</code> enables some internal checking such as avoiding conflicts. For example, it is often necessary to set <code>zero = NULL</code> and <code>parallel = FALSE</code>, otherwise there would be contradictions.</p>
type.fitted	<p>Character. Type of fitted value returned by the <code>fitted()</code> methods function. The first choice is always the default. The available choices depends on what kind of family function it is. Using the first few letters of the chosen choice is okay. See <code>fittedvglm</code> for more details.</p> <p>The choice "Qlink" refers to quantile-links, which was introduced in December 2018 in VGAMextra 0.0-2 for several 1-parameter distributions. Here, either the <code>loglink</code> or <code>logitlink</code> or <code>identitylink</code> of the quantile is the link function (and the choice is dependent on the support of the distribution), and link functions end in "Qlink". A limited amount of support is provided for such links, e.g., <code>fitted(fit)</code> are the fitted quantiles, which is the same as <code>predict(fit, type = "response")</code>. However, <code>fitted(fit, percentiles = 77)</code> will not work.</p>
percentiles	<p>Numeric vector, with values between 0 and 100 (although it is not recommended that exactly 0 or 100 be inputted). Used only if <code>type.fitted = "quantiles"</code> or <code>type.fitted = "percentiles"</code>, then this argument specifies the values of these quantiles. The argument name tries to reinforce that the values lie between 0 and 100. See <code>fittedvglm</code> for more details.</p>
probs.x, probs.y	<p>Numeric, with values in (0, 1). The probabilities that define quantiles with respect to some vector, usually an x or y of some sort. This is used to create two subsets of data corresponding to 'low' and 'high' values of x or y. Each value is separately fed into the <code>probs</code> argument of <code>quantile</code>. If the data set size is small then it may be necessary to increase/decrease slightly the first/second values respectively.</p>
lss	<p>Logical. This stands for the ordering: location, scale and shape. Should the ordering of the parameters be in this order? Almost all VGAM family functions</p>

have this order by default, but in order to match the arguments of existing R functions, one might need to set `lss = FALSE`. For example, the arguments of `weibullR` are scale and shape, whereas `rweibull` are shape and scale. As a temporary measure (from **VGAM** 0.9-7 onwards but prior to version 1.0-0), some family functions such as `sinmad` have an `lss` argument without a default. For these, setting `lss = FALSE` will work. Later, `lss = TRUE` will be the default. Be careful for the `dpqr`-type functions, e.g., `rsinmad`.

Thresh	<i>Thresholds</i> is another name for the intercepts, e.g., for categorical models. They may be constrained by functions such as <code>CM.equid</code> and <code>CM.qnorm</code> . The string "CM." and the Thresh argument is pasted and then that function is called to obtain the constraint matrix for the intercept term. So <code>Thresh = "free"</code> , <code>Thresh = "equid"</code> , <code>Thresh = "qnorm"</code> , <code>Thresh = "symm0"</code> , <code>Thresh = "symm1"</code> etc. are possibilities. Families that use this include <code>multinomial</code> , <code>cratio</code> , <code>sratio</code> , <code>cumulative</code> , <code>acat</code> .
whitespace	Logical. Should white spaces (" ") be used in the labelling of the linear/additive predictors? Setting <code>TRUE</code> usually results in more readability but it occupies more columns of the output.
oim	Logical. Should the observed information matrices (OIMs) be used for the working weights? In general, setting <code>oim = TRUE</code> means the Newton-Raphson algorithm, and <code>oim = FALSE</code> means Fisher-scoring. The latter uses the EIM, and is usually recommended. If <code>oim = TRUE</code> then <code>nsimEIM</code> is ignored.
nrfs	Numeric, a value in $[0, 1]$. Experimental argument for allowing a mixture of Newton-Raphson and Fisher scoring. The working weights are taken as a linear combination of the two. If <code>nrfs = 0</code> then Newton-Raphson is used, i.e., the OIM is wholly used. If <code>nrfs = 1</code> then Fisher scoring is used, i.e., the EIM is wholly used. If convergence is successful then one might expect Newton-Raphson to be faster than Fisher scoring because the former has an order-2 convergence rate while the latter has a linear rate.
,	
multiple.responses	Logical. Some VGAM family functions allow a multivariate or vector response. If so, then usually the response is a matrix with columns corresponding to the individual response variables. They are all fitted simultaneously. Arguments such as <code>parallel</code> may then be useful to allow for relationships between the regressions of each response variable. If <code>multiple.responses = TRUE</code> then sometimes the response is interpreted differently, e.g., <code>posbinomial</code> chooses the first column of a matrix response as success and combines the other columns as failure, but when <code>multiple.responses = TRUE</code> then each column of the response matrix is the number of successes and the <code>weights</code> argument is of the same dimension as the response and contains the number of trials.
earg.link	This argument should be generally ignored.
byrow.arg	Logical. Some VGAM family functions that handle multiple responses have arguments that allow input to be fed in which affect all the responses, e.g., <code>imu</code> for initializing a μ parameter. In such cases it is sometime more convenient to input one value per response by setting <code>byrow.arg = TRUE</code> ; then values are recycled in order to form a matrix of the appropriate dimension. This argument

	matches byrow in <code>matrix</code> ; in fact it is fed into such using <code>matrix(..., byrow = byrow.arg)</code> . This argument has no effect when there is one response.
<code>ishrinkage</code>	Shrinkage factor s used for obtaining initial values. Numeric, between 0 and 1. In general, the formula used is something like $s\mu + (1 - s)y$ where μ is a measure of central tendency such as a weighted mean or median, and y is the response vector. For example, the initial values are slight perturbations of the mean towards the actual data. For many types of models this method seems to work well and is often reasonably robust to outliers in the response. Often this argument is only used if the argument <code>imethod</code> is assigned a certain value.
<code>nointercept</code>	An integer-valued vector specifying which linear/additive predictors have no intercepts. Any values must be from the set $\{1, 2, \dots, M\}$. A value of <code>NULL</code> means no such constraints.
<code>bred</code>	Logical. Some VGAM family functions will allow bias-reduction based on the work by Kosmidis and Firth. Sometimes half-stepping is a good idea; set <code>stepsize = 0.5</code> and monitor convergence by setting <code>trace = TRUE</code> .
<code>link.list, earg.list</code>	Some VGAM family functions (such as <code>normal.vcm</code>) implement models with potentially lots of parameter link functions. These two arguments allow many such links and extra arguments to be inputted more easily. One has something like <code>link.list = list("Default" = "identitylink", x2 = "loglink", x3 = "logofflink")</code> and <code>earg.list = list("Default" = list(), x2 = list(), x3 = "list(offset = -1)")</code> . Then any unnamed terms will have the default link with its corresponding extra argument. Note: the <code>multilogitlink</code> link is also possible, and if so, at least two instances of it are necessary. Then the last term is the baseline/reference group.

Details

Full details will be given in documentation yet to be written, at a later date! A general recommendation is to set `trace = TRUE` whenever any model fitting is done, since monitoring convergence is usually very informative.

Value

An object of class "vglmff" (see `vglmff-class`). The object is used by modelling functions such as `vglm` and `vgam`.

Warning

The `zero` argument is supplied for convenience but conflicts can arise with other arguments, e.g., the `constraints` argument of `vglm` and `vgam`. See Example 5 below for an example. If not sure, use, e.g., `constraints(fit)` and `coef(fit, matrix = TRUE)` to check the result of a fit `fit`.

The arguments `zero` and `nointercept` can be inputted with values that fail. For example, `multinomial(zero = 2, nointercept = 1:3)` means the second linear/additive predictor is identically zero, which will cause a failure.

Be careful about the use of other potentially contradictory constraints, e.g., `multinomial(zero = 2, parallel = TRUE ~ x3)`. If in doubt, apply `constraints()` to the fitted object to check.

VGAM family functions with the `nsimEIM` may have inaccurate working weight matrices. If so, then the standard errors of the regression coefficients may be inaccurate. Thus output from `summary(fit)`, `vcov(fit)`, etc. may be misleading.

Changes relating to the `lss` argument have very important consequences and users must beware. Good programming style is to rely on the argument names and not on the order.

Author(s)

T. W. Yee

References

Yee, T. W. (2015). *Vector Generalized Linear and Additive Models: With an Implementation in R*. New York, USA: *Springer*.

Kosmidis, I. and Firth, D. (2009). Bias reduction in exponential family nonlinear models. *Biometrika*, **96**, 793–804.

Miranda-Soberanis, V. F. and Yee, T. W. (2019). New link functions for distribution-specific quantile regression based on vector generalized linear and additive models. *Journal of Probability and Statistics*, **5**, 1–11.

See Also

[Links](#), [vglm](#), [vgam](#), [vglmff-class](#), [UtilitiesVGAM](#), [multilogitlink](#), [multinomial](#), [VGAMextra](#).

Examples

```
# Example 1
cumulative()
cumulative(link = "probitlink", reverse = TRUE, parallel = TRUE)

# Example 2
wdata <- data.frame(x2 = runif(nn <- 1000))
wdata <- transform(wdata,
  y = rweibull(nn, shape = 2 + exp(1 + x2), scale = exp(-0.5)))
fit <- vglm(y ~ x2, weibullR(lshape = logofflink(offset = -2), zero = 2),
  data = wdata)
coef(fit, mat = TRUE)

# Example 3; multivariate (multiple) response
## Not run:
ndata <- data.frame(x = runif(nn <- 500))
ndata <- transform(ndata,
  y1 = rnbinom(nn, exp(1), mu = exp(3+x)), # k is size
  y2 = rnbinom(nn, exp(0), mu = exp(2-x)))
fit <- vglm(cbind(y1, y2) ~ x, negbinomial(zero = -2), ndata)
coef(fit, matrix = TRUE)

## End(Not run)
# Example 4
```

```

## Not run:
# fit1 and fit2 are equivalent
fit1 <- vglm(ymatrix ~ x2 + x3 + x4 + x5,
            cumulative(parallel = FALSE ~ 1 + x3 + x5), cdata)
fit2 <- vglm(ymatrix ~ x2 + x3 + x4 + x5,
            cumulative(parallel = TRUE ~ x2 + x4), cdata)

## End(Not run)

# Example 5
udata <- data.frame(x2 = rnorm(nn <- 200))
udata <- transform(udata,
                  x1copy = 1, # Copy of the intercept
                  x3 = runif(nn),
                  y1 = rnorm(nn, 1 - 3*x2, sd = exp(1 + 0.2*x2)),
                  y2 = rnorm(nn, 1 - 3*x2, sd = exp(1)))
args(uninormal)
fit1 <- vglm(y1 ~ x2, uninormal, udata) # This is okay
fit2 <- vglm(y2 ~ x2, uninormal(zero = 2), udata) # This is okay
fit4 <- vglm(y2 ~ x2 + x1copy + x3,
            uninormal(zero = NULL, vf1 = TRUE,
                    Form2 = ~ x1copy + x3 - 1), udata)
coef(fit4, matrix = TRUE) # VFL model

# This creates potential conflict
clist <- list("(Intercept)" = diag(2), "x2" = diag(2))
fit3 <- vglm(y2 ~ x2, uninormal(zero = 2), data = udata,
            constraints = clist) # Conflict!
coef(fit3, matrix = TRUE) # Shows that clist[["x2"]] was overwritten,
constraints(fit3) # i.e., 'zero' seems to override the 'constraints' arg

# Example 6 ('whitespace' argument)
pneumo <- transform(pneumo, let = log(exposure.time))
fit1 <- vglm(cbind(normal, mild, severe) ~ let,
            sratio(whitespace = FALSE, parallel = TRUE), pneumo)
fit2 <- vglm(cbind(normal, mild, severe) ~ let,
            sratio(whitespace = TRUE, parallel = TRUE), pneumo)
head(predict(fit1), 2) # No white spaces
head(predict(fit2), 2) # Uses white spaces

# Example 7 ('zero' argument with character input)
set.seed(123); n <- 1000
ldata <- data.frame(x2 = runif(n))
ldata <- transform(ldata, y1 = rlogis(n, loc = 5*x2, scale = exp(2)))
ldata <- transform(ldata, y2 = rlogis(n, loc = 5*x2, scale = exp(1*x2)))
ldata <- transform(ldata, w1 = runif(n))
ldata <- transform(ldata, w2 = runif(n))
fit7 <- vglm(cbind(y1, y2) ~ x2,
            # logistic(zero = "location1"), # location1 is intercept-only
            # logistic(zero = "location2"),
            # logistic(zero = "location*"), # Not okay... all is unmatched
            # logistic(zero = "scale1"),
            # logistic(zero = "scale2"),

```

```

#      logistic(zero = "scale"), # Both scale parameters are matched
#      logistic(zero = c("location", "scale2")), # All but scale1
#      logistic(zero = c("LOCAT", "scale2")), # Only scale2 is matched
#      logistic(zero = c("LOCAT")), # Nothing is matched
#      trace = TRUE,
#      weights = cbind(w1, w2),
#      weights = w1,
#      data = ldata)
coef(fit7, matrix = TRUE)

```

concoef

*Extract Model Constrained/Canonical Coefficients***Description**

concoef is a generic function which extracts the constrained (canonical) coefficients from objects returned by certain modelling functions.

Usage

```
concoef(object, ...)
```

Arguments

`object` An object for which the extraction of canonical coefficients is meaningful.
`...` Other arguments fed into the specific methods function of the model.

Details

For constrained quadratic and ordination models, *canonical coefficients* are the elements of the **C** matrix used to form the latent variables. They are highly interpretable in ecology, and are looked at as weights or loadings.

They are also applicable for reduced-rank VGLMs.

Value

The value returned depends specifically on the methods function invoked.

Warning

[concoef](#) replaces `ccoef`; the latter is deprecated.

For QO models, there is a direct inverse relationship between the scaling of the latent variables (site scores) and the tolerances. One normalization is for the latent variables to have unit variance. Another normalization is for all the species' tolerances to be unit (provided `eq.tolerances` is TRUE). These two normalizations cannot simultaneously hold in general. For rank R models with $R > 1$ it becomes more complicated because the latent variables are also uncorrelated. An important argument when fitting quadratic ordination models is whether `eq.tolerances` is TRUE or FALSE. See Yee (2004) for details.

Author(s)

Thomas W. Yee

References

Yee, T. W. and Hastie, T. J. (2003). Reduced-rank vector generalized linear models. *Statistical Modelling*, **3**, 15–41.

Yee, T. W. (2004). A new technique for maximum-likelihood canonical Gaussian ordination. *Ecological Monographs*, **74**, 685–701.

Yee, T. W. (2006). Constrained additive ordination. *Ecology*, **87**, 203–213.

See Also

[concoef-method](#), [concoef.qrrvglm](#), [concoef.cao](#), [coef](#).

Examples

```
## Not run: set.seed(111) # This leads to the global solution
hspider[,1:6] <- scale(hspider[,1:6]) # Standardized environmental vars
p1 <- cqc(cbind(Alopacce, Alopcone, Alopfabr, Arctlute, Arctperi,
               Auloalbi, Pardlugu, Pardmont, Pardnigr, Pardpull,
               Trocterr, Zoraspin) ~
          WaterCon + BareSand + FallTwig + CoveMoss + CoveHerb + ReflLux,
          family = poissonff, data = hspider, Crow1positive = FALSE)
concoef(p1)

## End(Not run)
```

concoef-methods

Constrained (Canonical) Coefficients

Description

concoef is a generic function used to return the constrained (canonical) coefficients of a constrained ordination model. The function invokes particular methods which depend on the class of the first argument.

Methods

object The object from which the constrained coefficients are extracted.

confintvglm

*Confidence Intervals for Parameters of VGLMs***Description**

Computes confidence intervals (CIs) for one or more parameters in a fitted model. Currently the object must be a "vglm" object.

Usage

```
confintvglm(object, parm, level = 0.95, method = c("wald", "profile"),
            trace = NULL, ...)
```

Arguments

object	A fitted model object.
parm, level, ...	Same as <code>confint</code> .
method	Character. The default is the first method. Abbreviations are allowed. Currently "profile" is basically working; and it is likely to be more accurate especially for small samples, as it is based on a profile log likelihood, however it is computationally intensive.
trace	Logical. If TRUE then one can monitor the computation as it progresses (because it is expensive). The default is the original model's trace value (see <code>vglm.control</code>). Setting FALSE suppresses all intermediate output.

Details

The default for this methods function is based on `confint.default` and assumes asymptotic normality. In particular, the `coef` and `vcov` methods functions are used for `vglm-class` objects.

When `method = "profile"` the function `profilevglm` is called to do the profiling. The code is very heavily based on `profile.glm` which was originally written by D. M. Bates and W. N. Venables (For S in 1996) and subsequently corrected by B. D. Ripley. Sometimes the profiling method can give problems, for example, `cumulative` requires the M linear predictors not to intersect in the data cloud. Such numerical problems are less common when `method = "wald"`, however, it is well-known that inference based on profile likelihoods is generally more accurate than Wald, especially when the sample size is small. The deviance (`deviance(object)`) is used if possible, else the difference $2 * (\logLik(object) - e11)$ is computed, where `e11` are the values of the loglikelihood on a grid.

For Wald CIs and `rrvglm-class` objects, currently an error message is produced because I haven't gotten around to write the methods function; it's not too hard, but am too busy! An interim measure is to coerce the object into a "vglm" object, but then the confidence intervals will tend to be too narrow because the estimated constraint matrices are treated as known.

For Wald CIs and `vgam-class` objects, currently an error message is produced because the theory is undeveloped.

Value

Same as [confint](#).

Note

The order of the values of argument `method` may change in the future without notice. The functions `plot.profile.glm` and `pairs.profile.glm` from **MASS** appear to work with output from this function.

Author(s)

Thomas Yee adapted [confint.lm](#) to handle "vglm" objects, for Wald-type confidence intervals. Also, [profile.glm](#) was originally written by D. M. Bates and W. N. Venables (For S in 1996) and subsequently corrected by B. D. Ripley. This function effectively calls `confint.profile.glm()` in **MASS**.

See Also

[vcovvglm](#), [summaryvglm](#), [confint](#), [profile.glm](#), [lrt.stat.vlm](#), [wald.stat](#), [plot.profile.glm](#), [pairs.profile.glm](#).

Examples

```
# Example 1: this is based on a glm example
counts <- c(18,17,15,20,10,20,25,13,12)
outcome <- gl(3, 1, 9); treatment <- gl(3, 3)
glm.D93 <- glm(counts ~ outcome + treatment, family = poisson())
vglm.D93 <- vglm(counts ~ outcome + treatment, family = poissonff)
confint(glm.D93) # needs MASS to be present on the system
confint.default(glm.D93) # based on asymptotic normality
confint(vglm.D93)
confint(vglm.D93) - confint(glm.D93) # Should be all 0s
confint(vglm.D93) - confint.default(glm.D93) # based on asympt. normality

# Example 2: simulated negative binomial data with multiple responses
ndata <- data.frame(x2 = runif(nn <- 100))
ndata <- transform(ndata, y1 = rnbinom(nn, mu = exp(3+x2), size = exp(1)),
                   y2 = rnbinom(nn, mu = exp(2-x2), size = exp(0)))
fit1 <- vglm(cbind(y1, y2) ~ x2, negbinomial, data = ndata, trace = TRUE)
coef(fit1)
coef(fit1, matrix = TRUE)
confint(fit1)
confint(fit1, "x2:1") # This might be improved to "x2" some day...
## Not run:
confint(fit1, method = "profile") # Computationally expensive
confint(fit1, "x2:1", method = "profile", trace = FALSE)

## End(Not run)

fit2 <- rrvglm(y1 ~ x2, negbinomial(zero = NULL), data = ndata)
confint(as(fit2, "vglm")) # Too narrow (SEs are biased downwards)
```


constraints

*Constraint Matrices***Description**

Extractor function for the *constraint matrices* of objects in the **VGAM** package.

Usage

```
constraints(object, ...)
constraints.vlm(object, type = c("lm", "term"), all = TRUE, which,
               matrix.out = FALSE, colnames.arg = TRUE,
               rownames.arg = TRUE, ...)
```

Arguments

object	Some VGAM object, for example, having class <code>vglmff-class</code> .
type	Character. Whether LM- or term-type constraints are to be returned. The number of such matrices returned is equal to <code>nvar(object, type = "lm")</code> and the number of terms, respectively.
all, which	If <code>all = FALSE</code> then which gives the integer index or a vector of logicals specifying the selection.
matrix.out	Logical. If TRUE then the constraint matrices are <code>cbind()</code> ed together. The result is usually more compact because the default is a list of constraint matrices.
colnames.arg, rownames.arg	Logical. If TRUE then column and row names are assigned corresponding to the variables.
...	Other possible arguments such as type.

Details

Constraint matrices describe the relationship of coefficients/component functions of a particular explanatory variable between the linear/additive predictors in VGLM/VGAM models. For example, they may be all different (constraint matrix is the identity matrix) or all the same (constraint matrix has one column and has unit values).

VGLMs and VGAMs have constraint matrices which are *known*. The class of RR-VGLMs have constraint matrices which are *unknown* and are to be estimated.

Value

The extractor function `constraints()` returns a list comprising of constraint matrices—usually one for each column of the VLM model matrix, and in that order. The list is labelled with the variable names. Each constraint matrix has M rows, where M is the number of linear/additive predictors, and whose rank is equal to the number of columns. A model with no constraints at all has an order M identity matrix as each variable's constraint matrix.

For `vglm` and `vgam` objects, feeding in `type = "term"` constraint matrices back into the same model should work and give an identical model. The default are the "lm"-type constraint matrices; this is a list with one constraint matrix per column of the LM matrix. See the `constraints` argument of `vglm`, and the example below.

Note

In all **VGAM** family functions `zero = NULL` means none of the linear/additive predictors are modelled as intercepts-only. Other arguments found in certain **VGAM** family functions which affect constraint matrices include `parallel` and `exchangeable`.

The `constraints` argument in `vglm` and `vgam` allows constraint matrices to be inputted. If so, then `constraints(fit, type = "lm")` can be fed into the `constraints` argument of the same object to get the same model.

The `xij` argument does not affect constraint matrices; rather, it allows each row of the constraint matrix to be multiplied by a specified vector.

Author(s)

T. W. Yee

References

Yee, T. W. and Wild, C. J. (1996). Vector generalized additive models. *Journal of the Royal Statistical Society, Series B, Methodological*, **58**, 481–493.

Yee, T. W. and Hastie, T. J. (2003). Reduced-rank vector generalized linear models. *Statistical Modelling*, **3**, 15–41.

See Also

`is.parallel`, `is.zero`, `trim.constraints`. VGLMs are described in `vglm-class`; RR-VGLMs are described in `rrvglm-class`.

Arguments such as `zero` and `parallel` found in many **VGAM** family functions are a way of creating/modifying constraint matrices conveniently, e.g., see `zero`. See `CommonVGAMffArguments` for more information.

Examples

```
# Fit the proportional odds model:
pneumo <- transform(pneumo, let = log(exposure.time))
(fit1 <- vglm(cbind(normal, mild, severe) ~ sm.bs(let, 3),
             cumulative(parallel = TRUE, reverse = TRUE), data = pneumo))
coef(fit1, matrix = TRUE)
constraints(fit1) # Parallel assumption results in this
constraints(fit1, type = "term") # Same as the default ("vglm"-type)
is.parallel(fit1)

# An equivalent model to fit1 (needs the type "term" constraints):
clist.term <- constraints(fit1, type = "term") # "term"-type constraints
# cumulative() has no 'zero' argument to set to NULL (a good idea
```

```
# when using the 'constraints' argument):
(fit2 <- vglm(cbind(normal, mild, severe) ~ sm.bs(let, 3), data = pneumo,
             cumulative(reverse = TRUE), constraints = clist.term))
abs(max(coef(fit1, matrix = TRUE) -
        coef(fit2, matrix = TRUE))) # Should be zero

# Fit a rank-1 stereotype (RR-multinomial logit) model:
fit <- rrvglm(Country ~ Width + Height + HP, multinomial, data = car.all)
constraints(fit) # All except the first are the estimated A matrix
```

cops

Centre of the Parameter Space

Description

Returns a vector similar to `coefs()` comprising the centre of the parameter space (COPS) values, given a fitted VGLM regression.

Usage

```
cops(object, ...)
copsvglm(object, beta.range = c(-5, 6),
         tol = .Machine$double.eps^0.25,
         dointercepts = TRUE, trace. = FALSE,
         slowtrain = FALSE, ...)
```

Arguments

<code>object</code>	A <code>vglm</code> object. However, this function will not work for all such objects.
<code>beta.range</code>	Numeric. Interval for the numerical search. After a little scaling, it is effectively fed into <code>interval</code> in <code>optimize</code> . Convergence failure may occur if this argument is too wide so it is a good idea to vary this argument. In fact, it is strongly recommended that several values be fed into this argument to help ensure the proper solution is obtained.
<code>tol</code>	Numeric. Fed into <code>tol</code> in <code>optimize</code> .
<code>dointercepts</code>	Logical. Compute the COPS for the intercepts? This should be set to <code>FALSE</code> for models such as <code>propodds</code> and <code>cumulative</code> .
<code>trace.</code>	Logical. Print a running log? This may or may not work properly.
<code>slowtrain</code>	Logical. If <code>TRUE</code> then all columns of a matrix is computed. If <code>FALSE</code> then only one column of a matrix is computed, and this is the only column needed.
<code>...</code>	currently unused but may be used in the future for further arguments passed into the other methods functions.

Details

For many models, some COPS values will be Inf or -Inf so that manual checking is needed, for example, `poissonff`. Each value returned may be effectively that of `beta.range` or NA. The answers returned by this function only make sense if the COPSs are in the interior of the parameter space. This function was written specifically for logistic regression but has much wider applicability. Currently the result returned depends critically on `beta.range` so that the answer should be checked after several values are fed into that argument.

Value

A named vector, similar to `coefvglm`. If `trace`, then a list is returned, having a component comprising a matrix of function evaluations used by `optimize`.

Note

This function is experimental and can be made to run more efficiently in the future.

Author(s)

Thomas W. Yee.

References

Yee, T. W. (2024). Musings and new results on the parameter space. *Under review*.

See Also

`hdeff`, `coefvglm`, `coef`.

Examples

```
## Not run: data("xs.nz", package = "VGAMdata")
data1 <- na.omit(xs.nz[, c("age", "cancer", "sex")])
fit1 <- vglm(cancer ~ age + sex, binomialff, data1)
cops(fit1) # 'beta.range' is okay here

## End(Not run)
```

corbet

Corbet's Butterfly Data

Description

About 3300 individual butterflies were caught in Malaya by naturalist Corbet trapping butterflies. They were classified to about 500 species.

Usage

```
data(corbet)
```

Format

A data frame with 24 observations on the following 2 variables.

species Number of species.

ofreq Observed frequency of individual butterflies of that species.

Details

In the early 1940s Corbet spent two years trapping butterflies in Malaya. Of interest was the total number of species. Some species were so rare (e.g., 118 species had only one specimen) that it was thought likely that there were many unknown species.

Actually, 119 species had over 24 observed frequencies, so this could/should be appended to the data set. Hence there are 620 species in total in a sample size of $n = 9031$ individuals.

References

Fisher, R. A., Corbet, A. S. and Williams, C. B. (1943). The Relation Between the Number of Species and the Number of Individuals in a Random Sample of an Animal Population. *Journal of Animal Ecology*, **12**, 42–58.

Examples

```
summary(corbet)
```

 cqo

Fitting Constrained Quadratic Ordination (CQO)

Description

A *constrained quadratic ordination* (CQO; formerly called *canonical Gaussian ordination* or CGO) model is fitted using the *quadratic reduced-rank vector generalized linear model* (QRR-VGLM) framework.

Usage

```
cqo(formula, family = stop("argument 'family' needs to be assigned"),
     data = list(), weights = NULL, subset = NULL,
     na.action = na.fail, etastart = NULL, mustart = NULL,
     coefstart = NULL, control = qrrvglm.control(...), offset = NULL,
     method = "cqo.fit", model = FALSE, x.arg = TRUE, y.arg = TRUE,
     contrasts = NULL, constraints = NULL, extra = NULL,
     smart = TRUE, ...)
```

Arguments

formula	a symbolic description of the model to be fit. The RHS of the formula is applied to each linear predictor. Different variables in each linear predictor can be chosen by specifying constraint matrices.
family	a function of class "vglmff" (see vglmff-class) describing what statistical model is to be fitted. This is called a "VGAM family function". See CommonVGAMffArguments for general information about many types of arguments found in this type of function. Currently the following families are supported: poissonff , binomialff (logitlink and clogloglink links available), negbinomial , gamma2 . Sometimes special arguments are required for <code>cqo()</code> , e.g., <code>binomialff(multiple.responses = TRUE)</code> .
data	an optional data frame containing the variables in the model. By default the variables are taken from <code>environment(formula)</code> , typically the environment from which <code>cqo</code> is called.
weights	an optional vector or matrix of (prior) weights to be used in the fitting process. Currently, this argument should not be used.
subset	an optional logical vector specifying a subset of observations to be used in the fitting process.
na.action	a function which indicates what should happen when the data contain NAs. The default is set by the <code>na.action</code> setting of options , and is <code>na.fail</code> if that is unset. The "factory-fresh" default is <code>na.omit</code> .
etastart	starting values for the linear predictors. It is a M -column matrix. If $M = 1$ then it may be a vector. Currently, this argument probably should not be used.
mustart	starting values for the fitted values. It can be a vector or a matrix. Some family functions do not make use of this argument. Currently, this argument probably should not be used.
coefstart	starting values for the coefficient vector. Currently, this argument probably should not be used.
control	a list of parameters for controlling the fitting process. See qrrvglm.control for details.
offset	This argument must not be used.
method	the method to be used in fitting the model. The default (and presently only) method <code>cqo.fit</code> uses <i>iteratively reweighted least squares</i> (IRLS).
model	a logical value indicating whether the <i>model frame</i> should be assigned in the model slot.
x.arg, y.arg	logical values indicating whether the model matrix and response matrix used in the fitting process should be assigned in the x and y slots. Note the model matrix is the LM model matrix.
contrasts	an optional list. See the <code>contrasts.arg</code> of <code>model.matrix.default</code> .
constraints	an optional list of constraint matrices. The components of the list must be named with the term it corresponds to (and it must match in character format). Each constraint matrix must have M rows, and be of full-column rank. By default, constraint matrices are the M by M identity matrix unless arguments in the

	family function itself override these values. If <code>constraints</code> is used it must contain <i>all</i> the terms; an incomplete list is not accepted. Constraint matrices for x_2 variables are taken as the identity matrix.
<code>extra</code>	an optional list with any extra information that might be needed by the family function.
<code>smart</code>	logical value indicating whether smart prediction (<code>smartpred</code>) will be used.
<code>...</code>	further arguments passed into <code>qrrvglm.control</code> .

Details

QRR-VGLMs or *constrained quadratic ordination* (CQO) models are estimated here by maximum likelihood estimation. Optimal linear combinations of the environmental variables are computed, called *latent variables* (these appear as `latvar` for $R = 1$ else `latvar1`, `latvar2`, etc. in the output). Here, R is the *rank* or the number of ordination axes. Each species' response is then a regression of these latent variables using quadratic polynomials on a transformed scale (e.g., log for Poisson counts, logit for presence/absence responses). The solution is obtained iteratively in order to maximize the log-likelihood function, or equivalently, minimize the deviance.

The central formula (for Poisson and binomial species data) is given by

$$\eta = B_1^T x_1 + A\nu + \sum_{m=1}^M (\nu^T D_m \nu) e_m$$

where x_1 is a vector (usually just a 1 for an intercept), x_2 is a vector of environmental variables, $\nu = C^T x_2$ is a R -vector of latent variables, e_m is a vector of 0s but with a 1 in the m th position. The η are a vector of linear/additive predictors, e.g., the m th element is $\eta_m = \log(E[Y_m])$ for the m th species. The matrices B_1 , A , C and D_m are estimated from the data, i.e., contain the regression coefficients. The tolerance matrices satisfy $T_s = -\frac{1}{2} D_s^{-1}$. Many important CQO details are directly related to arguments in `qrrvglm.control`, e.g., the argument `noRRR` specifies which variables comprise x_1 .

Theoretically, the four most popular **VGAM** family functions to be used with `cgo` correspond to the Poisson, binomial, normal, and negative binomial distributions. The latter is a 2-parameter model. All of these are implemented, as well as the 2-parameter gamma.

For initial values, the function `.Init.Poisson.QO` should work reasonably well if the data is Poisson with species having equal tolerances. It can be quite good on binary data too. Otherwise the `Cinit` argument in `qrrvglm.control` can be used.

It is possible to relax the quadratic form to an additive model. The result is a data-driven approach rather than a model-driven approach, so that CQO is extended to *constrained additive ordination* (CAO) when $R = 1$. See `cao` for more details.

In this documentation, M is the number of linear predictors, S is the number of responses (species). Then $M = S$ for Poisson and binomial species data, and $M = 2S$ for negative binomial and gamma distributed species data.

Incidentally, *Unconstrained quadratic ordination* (UQO) may be performed by, e.g., fitting a Goodman's RC association model; see `uqo` and the Yee and Hadi (2014) referenced there. For UQO, the response is the usual site-by-species matrix and there are no environmental variables; the site scores are free parameters. UQO can be performed under the assumption that all species have the same tolerance matrices.

Value

An object of class "qrrvglm".

Warning

Local solutions are not uncommon when fitting CQO models. To increase the chances of obtaining the global solution, increase the value of the argument `Bestof` in `qrrvglm.control`. For reproducibility of the results, it pays to set a different random number seed before calling `cgo` (the function `set.seed` does this). The function `cgo` chooses initial values for **C** using `.Init.Poisson.QO()` if `Use.Init.Poisson.QO = TRUE`, else random numbers.

Unless `I.tolerances = TRUE` or `eq.tolerances = FALSE`, CQO is computationally expensive with memory and time. It pays to keep the rank down to 1 or 2. If `eq.tolerances = TRUE` and `I.tolerances = FALSE` then the cost grows quickly with the number of species and sites (in terms of memory requirements and time). The data needs to conform quite closely to the statistical model, and the environmental range of the data should be wide in order for the quadratics to fit the data well (bell-shaped response surfaces). If not, RR-VGLMs will be more appropriate because the response is linear on the transformed scale (e.g., log or logit) and the ordination is called *constrained linear ordination* or CLO.

Like many regression models, CQO is sensitive to outliers (in the environmental and species data), sparse data, high leverage points, multicollinearity etc. For these reasons, it is necessary to examine the data carefully for these features and take corrective action (e.g., omitting certain species, sites, environmental variables from the analysis, transforming certain environmental variables, etc.). Any optimum lying outside the convex hull of the site scores should not be trusted. Fitting a CAO is recommended first, then upon transformations etc., possibly a CQO can be fitted.

For binary data, it is necessary to have 'enough' data. In general, the number of sites n ought to be much larger than the number of species S , e.g., at least 100 sites for two species. Compared to count (Poisson) data, numerical problems occur more frequently with presence/absence (binary) data. For example, if `Rank = 1` and if the response data for each species is a string of all absences, then all presences, then all absences (when enumerated along the latent variable) then infinite parameter estimates will occur. In general, setting `I.tolerances = TRUE` may help.

This function was formerly called `cgo`. It has been renamed to reinforce a new nomenclature described in Yee (2006).

Note

The input requires care, preparation and thought—a *lot more* than other ordination methods. Here is a partial **checklist**.

- (1) The number of species should be kept reasonably low, e.g., 12 max. Feeding in 100+ species wholesale is a recipe for failure. Choose a few species carefully. Using 10 well-chosen species is better than 100+ species thrown in willy-nilly.
- (2) Each species should be screened individually first, e.g., for presence/absence is the species totally absent or totally present at all sites? For presence/absence data `sort(colMeans(data))` can help avoid such species.
- (3) The number of explanatory variables should be kept low, e.g., 7 max.

- (4) Each explanatory variable should be screened individually first, e.g., is it heavily skewed or are there outliers? They should be plotted and then transformed where needed. They should not be too highly correlated with each other.
- (5) Each explanatory variable should be scaled, e.g., to mean 0 and unit variance. This is especially needed for `I.tolerance = TRUE`.
- (6) Keep the rank low. Only if the data is very good should a rank-2 model be attempted. Usually a rank-1 model is all that is practically possible even after a lot of work. The rank-1 model should always be attempted first. Then might be clever and try use this for initial values for a rank-2 model.
- (7) If the number of sites is large then choose a random sample of them. For example, choose a maximum of 500 sites. This will reduce the memory and time expense of the computations.
- (8) Try `I.tolerance = TRUE` or `eq.tolerance = FALSE` if the inputted data set is large, so as to reduce the computational expense. That's because the default, `I.tolerance = FALSE` and `eq.tolerance = TRUE`, is very memory hungry.

By default, a rank-1 equal-tolerances QRR-VGLM model is fitted (see `qrrvglm.control` for the default control parameters). If `Rank > 1` then the latent variables are always transformed so that they are uncorrelated. By default, the argument `trace` is `TRUE` meaning a running log is printed out while the computations are taking place. This is because the algorithm is computationally expensive, therefore users might think that their computers have frozen if `trace = FALSE`!

The argument `Bestof` in `qrrvglm.control` controls the number of models fitted (each uses different starting values) to the data. This argument is important because convergence may be to a *local* solution rather than the *global* solution. Using more starting values increases the chances of finding the global solution. Always plot an ordination diagram (use the generic function `lvplot`) and see if it looks sensible. Local solutions arise because the optimization problem is highly nonlinear, and this is particularly true for CAO.

Many of the arguments applicable to `cgo` are common to `vglm` and `rvvglm.control`. The most important arguments are `Rank`, `noRRR`, `Bestof`, `I.tolerances`, `eq.tolerances`, `isd.latvar`, and `MUXfactor`.

When fitting a 2-parameter model such as the negative binomial or gamma, it pays to have `eq.tolerances = TRUE` and `I.tolerances = FALSE`. This is because numerical problems can occur when fitting the model far away from the global solution when `I.tolerances = TRUE`. Setting the two arguments as described will slow down the computation considerably, however it is numerically more stable.

In Example 1 below, an unequal-tolerances rank-1 QRR-VGLM is fitted to the hunting spiders dataset, and Example 2 is the equal-tolerances version. The latter is less likely to have convergence problems compared to the unequal-tolerances model. In Example 3 below, an equal-tolerances rank-2 QRR-VGLM is fitted to the hunting spiders dataset. The numerical difficulties encountered in fitting the rank-2 model suggests a rank-1 model is probably preferable. In Example 4 below, constrained binary quadratic ordination (in old nomenclature, constrained Gaussian logit ordination) is fitted to some simulated data coming from a species packing model. With multivariate binary responses, one must use `multiple.responses = TRUE` to indicate that the response (matrix) is multivariate. Otherwise, it is interpreted as a single binary response variable. In Example 5 below, the deviance residuals are plotted for each species. This is useful as a diagnostic plot. This is done by (re)regressing each species separately against the latent variable.

Sometime in the future, this function might handle input of the form `cgo(x, y)`, where `x` and `y` are matrices containing the environmental and species data respectively.

Author(s)

Thomas W. Yee. Thanks to Alvin Sou for converting a lot of the original FORTRAN code into C.

References

Yee, T. W. (2004). A new technique for maximum-likelihood canonical Gaussian ordination. *Ecological Monographs*, **74**, 685–701.

ter Braak, C. J. F. and Prentice, I. C. (1988). A theory of gradient analysis. *Advances in Ecological Research*, **18**, 271–317.

Yee, T. W. (2006). Constrained additive ordination. *Ecology*, **87**, 203–213.

See Also

[qrrvglm.control](#), [Coef.qrrvglm](#), [predictqrrvglm](#), [calibrate.qrrvglm](#), [model.matrixqrrvglm](#), [vcovqrrvglm](#), [rcqo](#), [cao](#), [uqo](#), [rrvglm](#), [poissonff](#), [binomialff](#), [negbinomial](#), [gamma2](#), [lvplot.qrrvglm](#), [persprrvglm](#), [trplot.qrrvglm](#), [vglm](#), [set.seed](#), [hspider](#), [trap0](#).

Examples

```
## Not run:
# Example 1; Fit an unequal tolerances model to the hunting spiders data
hspider[,1:6] <- scale(hspider[,1:6]) # Standardized environmental variables
set.seed(1234) # For reproducibility of the results
p1ut <- cgo(cbind(Alopacce, Alopcone, Alopfabr, Arctlute, Arctperi,
                Auloalbi, Pardlugu, Pardmont, Pardnigr, Pardpull,
                Trocterr, Zoraspin) ~
            WaterCon + BareSand + FallTwig + CoveMoss + CoveHerb + ReflLux,
            fam = poissonff, data = hspider, Crowlpositive = FALSE,
            eq.tol = FALSE)
sort(deviance(p1ut, history = TRUE)) # A history of all the iterations
if (deviance(p1ut) > 1177) warning("suboptimal fit obtained")

S <- ncol(depvar(p1ut)) # Number of species
clr <- (1:(S+1))[-7] # Omits yellow
lvplot(p1ut, y = TRUE, lcol = clr, pch = 1:S, pcol = clr,
       las = 1) # Ordination diagram
legend("topright", leg = colnames(depvar(p1ut)), col = clr,
       pch = 1:S, merge = TRUE, bty = "n", lty = 1:S, lwd = 2)
(cp <- Coef(p1ut))

(a <- latvar(cp)[cp@latvar.order]) # Ordered site scores along the gradient
# Names of the ordered sites along the gradient:
rownames(latvar(cp))[cp@latvar.order]
(aa <- Opt(cp)[, cp@Optimum.order]) # Ordered optimums along the gradient
aa <- aa[!is.na(aa)] # Delete the species that is not unimodal
names(aa) # Names of the ordered optimums along the gradient

trplot(p1ut, which.species = 1:3, log = "xy", type = "b", lty = 1, lwd = 2,
       col = c("blue", "red", "green"), label = TRUE) -> ii # Trajectory plot
legend(0.00005, 0.3, paste(ii$species[, 1], ii$species[, 2], sep = " and "),
       lwd = 2, lty = 1, col = c("blue", "red", "green"))
```

```

abline(a = 0, b = 1, lty = "dashed")

S <- ncol(depvar(p1ut)) # Number of species
clr <- (1:(S+1))[-7] # Omits yellow
persp(p1ut, col = clr, label = TRUE, las = 1) # Perspective plot

# Example 2; Fit an equal tolerances model. Less numerically fraught.
set.seed(1234)
p1et <- cqo(cbind(Alopacce, Alopcone, Alopfabr, Arctlute, Arctperi,
                 Auloalbi, Pardlugu, Pardmont, Pardnigr, Pardpull,
                 Trocterr, Zoraspin) ~
            WaterCon + BareSand + FallTwig + CoveMoss + CoveHerb + ReflLux,
            poissonff, data = hspider, Crowlpositive = FALSE)
sort(deviance(p1et, history = TRUE)) # A history of all the iterations
if (deviance(p1et) > 1586) warning("suboptimal fit obtained")
S <- ncol(depvar(p1et)) # Number of species
clr <- (1:(S+1))[-7] # Omits yellow
persp(p1et, col = clr, label = TRUE, las = 1)

# Example 3: A rank-2 equal tolerances CQO model with Poisson data
# This example is numerically fraught... need I.toler = TRUE too.
set.seed(555)
p2 <- cqo(cbind(Alopacce, Alopcone, Alopfabr, Arctlute, Arctperi,
                 Auloalbi, Pardlugu, Pardmont, Pardnigr, Pardpull,
                 Trocterr, Zoraspin) ~
            WaterCon + BareSand + FallTwig + CoveMoss + CoveHerb + ReflLux,
            poissonff, data = hspider, Crowlpositive = FALSE,
            I.toler = TRUE, Rank = 2, Bestof = 3, isd.latvar = c(2.1, 0.9))
sort(deviance(p2, history = TRUE)) # A history of all the iterations
if (deviance(p2) > 1127) warning("suboptimal fit obtained")
lvplot(p2, ellips = FALSE, label = TRUE, xlim = c(-3,4),
        C = TRUE, Ccol = "brown", sites = TRUE, scol = "grey",
        pcol = "blue", pch = "+", chull = TRUE, ccol = "grey")

# Example 4: species packing model with presence/absence data
set.seed(2345)
n <- 200; p <- 5; S <- 5
mydata <- rcqo(n, p, S, fam = "binomial", hi.abundance = 4,
              eq.tol = TRUE, es.opt = TRUE, eq.max = TRUE)
myform <- attr(mydata, "formula")
set.seed(1234)
b1et <- cqo(myform, binomialfff(multiple.responses = TRUE, link = "clogloglink"),
            data = mydata)
sort(deviance(b1et, history = TRUE)) # A history of all the iterations
lvplot(b1et, y = TRUE, lcol = 1:S, pch = 1:S, pcol = 1:S, las = 1)
Coef(b1et)

# Compare the fitted model with the 'truth'
cbind(truth = attr(mydata, "concoefficients"), fitted = concoef(b1et))

```

```

# Example 5: Plot the deviance residuals for diagnostic purposes
set.seed(1234)
p1et <- cqa(cbind(Alopacce, Alopcone, Alopfabr, Arctlute, Arctperi,
                 Auloalbi, Pardlugu, Pardmont, Pardnigr, Pardpull,
                 Trocterr, Zoraspin) ~
           WaterCon + BareSand + FallTwig + CoveMoss + CoveHerb + ReflLux,
           poissonff, data = hspider, eq.tol = TRUE, trace = FALSE)
sort(deviance(p1et, history = TRUE)) # A history of all the iterations
if (deviance(p1et) > 1586) warning("suboptimal fit obtained")
S <- ncol(depvar(p1et))
par(mfrow = c(3, 4))
for (ii in 1:S) {
  tempdata <- data.frame(latvar1 = c(latvar(p1et)),
                        sppCounts = depvar(p1et)[, ii])
  tempdata <- transform(tempdata, myOffset = -0.5 * latvar1^2)

  # For species ii, refit the model to get the deviance residuals
  fit1 <- vglm(sppCounts ~ offset(myOffset) + latvar1, poissonff,
              data = tempdata, trace = FALSE)

  # For checking: this should be 0
  # print("max(abs(c(Coef(p1et)@B1[1,ii],Coef(p1et)@A[ii,1])-coef(fit1)))")
  # print( max(abs(c(Coef(p1et)@B1[1,ii],Coef(p1et)@A[ii,1])-coef(fit1))) )

  # Plot the deviance residuals
  devresid <- resid(fit1, type = "deviance")
  predvalues <- predict(fit1) + fit1@offset
  ooo <- with(tempdata, order(latvar1))
  plot(predvalues + devresid ~ latvar1, data = tempdata, col = "red",
       xlab = "latvar1", ylab = "", main = colnames(depvar(p1et))[ii])
  with(tempdata, lines(latvar1[ooo], predvalues[ooo], col = "blue"))
}

## End(Not run)

```

crashes

Crashes on New Zealand Roads in 2009

Description

A variety of reported crash data cross-classified by time (hour of the day) and day of the week, accumulated over 2009. These include fatalities and injuries (by car), trucks, motor cycles, bicycles and pedestrians. There are some alcohol-related data too.

Usage

```

data(crashi)
data(crashf)
data(crasht)

```

```
data(crashmc)
data(crashbc)
data(crashp)
data(alcoff)
data(alclevels)
```

Format

Data frames with hourly times as rows and days of the week as columns. The `alclevels` dataset has hourly times and alcohol levels.

Mon, Tue, Wed, Thu, Fri, Sat, Sun Day of the week.

0-30, 31-50, 51-80, 81-100, 101-120, 121-150, 151-200, 201-250, 251-300, 301-350, 350+ Blood alcohol level (milligrams alcohol per 100 millilitres of blood).

Details

Each cell is the aggregate number of crashes reported at each hour-day combination, over the 2009 calendar year. The rownames of each data frame is the start time (hourly from midnight onwards) on a 24 hour clock, e.g., 21 means 9.00pm to 9.59pm.

For crashes, `chrashi` are the number of injuries by car, `crashf` are the number of fatalities by car (not included in `chrashi`), `crashtr` are the number of crashes involving trucks, `crashmc` are the number of crashes involving motorcyclists, `crashbc` are the number of crashes involving bicycles, and `crashp` are the number of crashes involving pedestrians. For alcohol-related offences, `alcoff` are the number of alcohol offenders from breath screening drivers, and `alclevels` are the blood alcohol levels of fatally injured drivers.

Source

<http://www.transport.govt.nz/research/Pages/Motor-Vehicle-Crashes-in-New-Zealand-2009.aspx>. Thanks to Warwick Goold and Alfian F. Hadi for assistance.

References

Motor Vehicles Crashes in New Zealand 2009; Statistical Statement Calendar Year 2009. Ministry of Transport, NZ Government; Yearly Report 2010. ISSN: 1176-3949

See Also

[rrvglm](#), [rcim](#), [grc](#).

Examples

```
## Not run: plot(unlist(alcoff), type = "l", frame.plot = TRUE,
  axes = FALSE, col = "blue", bty = "o",
  main = "Alcoholic offenders on NZ roads, aggregated over 2009",
  sub = "Vertical lines at midnight (purple) and noon (orange)",
  xlab = "Day/hour", ylab = "Number of offenders")
axis(1, at = 1 + (0:6) * 24 + 12, labels = colnames(alcoff))
axis(2, las = 1)
```

```

axis(3:4, labels = FALSE, tick = FALSE)
abline(v = sort(1 + c((0:7) * 24, (0:6) * 24 + 12)), lty = "dashed",
       col = c("purple", "orange"))
## End(Not run)

# Goodmans RC models
## Not run:
fitgrc1 <- grc(alcoff) # Rank-1 model
fitgrc2 <- grc(alcoff, Rank = 2, Corner = FALSE, Uncor = TRUE)
Coef(fitgrc2)

## End(Not run)
## Not run: biplot(fitgrc2, scaleA = 2.3, Ccol = "blue", Acol = "orange",
                  Clabels = as.character(1:23), xlim = c(-1.3, 2.3),
                  ylim = c(-1.2, 1))
## End(Not run)

```

cratio

Ordinal Regression with Continuation Ratios

Description

Fits a continuation ratio logit/probit/cloglog/cauchit/... regression model to an ordered (preferably) factor response.

Usage

```

cratio(link = "logitlink", parallel = FALSE, reverse = FALSE,
       zero = NULL, ynames = FALSE, Thresh = NULL, Trev = reverse,
       Tref = if (Trev) "M" else 1, whitespace = FALSE)

```

Arguments

link	Link function applied to the M continuation ratio probabilities. See Links for more choices.
parallel	A logical, or formula specifying which terms have equal/unequal coefficients.
reverse	Logical. By default, the continuation ratios used are $\eta_j = \text{logit}(P[Y > j Y \geq j])$ for $j = 1, \dots, M$. If reverse is TRUE, then $\eta_j = \text{logit}(P[Y < j + 1 Y \leq j + 1])$ will be used.
ynames	See multinomial for information.
zero	An integer-valued vector specifying which linear/additive predictors are modelled as intercepts only. The values must be from the set $\{1, 2, \dots, M\}$. The default value means none are modelled as intercept-only terms. See CommonVGAMffArguments for more information.
Thresh, Trev, Tref	See cumulative for information. These arguments apply to ordinal categorical regression models.
whitespace	See CommonVGAMffArguments for information.

Details

In this help file the response Y is assumed to be a factor with ordered values $1, 2, \dots, M + 1$, so that M is the number of linear/additive predictors η_j .

There are a number of definitions for the *continuation ratio* in the literature. To make life easier, in the **VGAM** package, we use *continuation* ratios and *stopping* ratios (see [sratio](#)). Stopping ratios deal with quantities such as $\text{logitlink}(P[Y=j|Y \geq j])$.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), [rrvglm](#) and [vgam](#).

Warning

No check is made to verify that the response is ordinal if the response is a matrix; see [ordered](#). Boersch-Supan (2021) looks at sparse data and the numerical problems that result; see [sratio](#).

Note

The response should be either a matrix of counts (with row sums that are all positive), or a factor. In both cases, the y slot returned by [vglm/vgam/rrvglm](#) is the matrix of counts.

For a nominal (unordered) factor response, the multinomial logit model ([multinomial](#)) is more appropriate.

Here is an example of the usage of the `parallel` argument. If there are covariates x_1 , x_2 and x_3 , then `parallel = TRUE ~ x1 + x2 - 1` and `parallel = FALSE ~ x3` are equivalent. This would constrain the regression coefficients for x_1 and x_2 to be equal; those of the intercepts and x_3 would be different.

Author(s)

Thomas W. Yee

References

See [sratio](#).

See Also

[sratio](#), [acat](#), [cumulative](#), [multinomial](#), [CM.equid](#), [CommonVGAMffArguments](#), [margeff](#), [pneumo](#), [budworm](#), [logitlink](#), [probitlink](#), [clogloglink](#), [cauchitlink](#).

Examples

```
pneumo <- transform(pneumo, let = log(exposure.time))
(fit <- vglm(cbind(normal, mild, severe) ~ let,
            cratio(parallel = TRUE), data = pneumo))
coef(fit, matrix = TRUE)
constraints(fit)
predict(fit)
```

```
predict(fit, untransform = TRUE)
margeff(fit)
```

cumulative

*Ordinal Regression with Cumulative Probabilities***Description**

Fits a cumulative link regression model to a (preferably ordered) factor response.

Usage

```
cumulative(link = "logitlink", parallel = FALSE,
           reverse = FALSE, multiple.responses = FALSE,
           ynames = FALSE, Thresh = NULL, Trev = reverse,
           Tref = if (Trev) "M" else 1, whitespace = FALSE)
```

Arguments

link	Link function applied to the J cumulative probabilities. See Links for more choices, e.g., for the cumulative probitlink/clogloglink/... models.
parallel	A logical or formula specifying which terms have equal/unequal coefficients. See below for more information about the parallelism assumption. The default results in what some people call the <i>generalized ordered logit model</i> to be fitted. If <code>parallel = TRUE</code> then it does not apply to the intercept. The <i>partial proportional odds model</i> can be fitted by assigning this argument something like <code>parallel = TRUE ~ -1 + x3 + x5</code> so that there is one regression coefficient for x_3 and x_5 . Equivalently, setting <code>parallel = FALSE ~ 1 + x2 + x4</code> means M regression coefficients for the intercept and x_2 and x_4 . It is important that the intercept is never parallel. See CommonVGAMffArguments for more information.
reverse	Logical. By default, the cumulative probabilities used are $P(Y \leq 1)$, $P(Y \leq 2)$, ..., $P(Y \leq J)$. If <code>reverse</code> is <code>TRUE</code> then $P(Y \geq 2)$, $P(Y \geq 3)$, ..., $P(Y \geq J + 1)$ are used.
ynames	See multinomial for information.
multiple.responses	Logical. Multiple responses? If <code>TRUE</code> then the input should be a matrix with values $1, 2, \dots, L$, where $L = J + 1$ is the number of levels. Each column of the matrix is a response, i.e., multiple responses. A suitable matrix can be obtained from <code>Cut</code> .
Thresh	Character. The choices concern constraint matrices applied to the intercepts. They can be constrained to be equally-spaced (<i>equidistant</i>) etc. See CommonVGAMffArguments and constraints for general information. Basically, the choice is pasted to the end of "CM." and that function is called. This means users can easily write their own CM.-type function.

If equally-spaced then the direction and the reference level are controlled by `Trev` and `Tref`, and the constraint matrix will be M by 2, with the second column corresponding to the distance between the thresholds.

If "symm1" then the fitted intercepts are *symmetric* about the median (M odd) intercept. If M is even then the median is the mean of the two most inner and adjacent intercepts. For this, `CM.symm1` is used to construct the appropriate constraint matrix.

If "symm0" then the median intercept is 0 by definition and the symmetry occurs about the origin. Thus the intercepts comprise pairs that differ by sign only. The appropriate constraint matrix is as with "symm1" but with the first column deleted. The choices "symm1" and "symm0" are effectively equivalent to "symmetric" and "symmetric2" respectively in **ordinal**.

For "qnorm" then `CM.qnorm` uses the `qnorm((1:M)/(M+1))` quantiles of the standard normal.

<code>Trev, Tref</code>	Support arguments for <code>Thresh</code> for equally-spaced intercepts. The logical argument <code>Trev</code> is applied first to give the direction (i.e., ascending or descending) before row <code>Tref</code> (ultimately numeric) of the first (intercept) constraint matrix is set to the reference level. See constraints for information.
<code>whitespace</code>	See CommonVGAMffArguments for information.

Details

In this help file the response Y is assumed to be a factor with ordered values $1, 2, \dots, J + 1$. Hence M is the number of linear/additive predictors η_j ; for `cumulative()` one has $M = J$.

This **VGAM** family function fits the class of *cumulative link models* to (hopefully) an ordinal response. By default, the *non-parallel* cumulative logit model is fitted, i.e.,

$$\eta_j = \text{logit}(P[Y \leq j])$$

where $j = 1, 2, \dots, M$ and the η_j are not constrained to be parallel. This is also known as the *non-proportional odds model*. If the logit link is replaced by a complementary log-log link ([clogloglink](#)) then this is known as the *proportional-hazards model*.

In almost all the literature, the constraint matrices associated with this family of models are known. For example, setting `parallel = TRUE` will make all constraint matrices (except for the intercept) equal to a vector of M 1's. If the constraint matrices are equal, unknown and to be estimated, then this can be achieved by fitting the model as a reduced-rank vector generalized linear model (RR-VGLM; see [rrvglm](#)). Currently, reduced-rank vector generalized additive models (RR-VGAMs) have not been implemented here.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Warning

No check is made to verify that the response is ordinal if the response is a matrix; see [ordered](#). Boersch-Supan (2021) looks at sparse data and the numerical problems that result; see [sratio](#).

Note

The response should be either a matrix of counts (with row sums that are all positive), or a factor. In both cases, the `y` slot returned by `vglm/vgam/rrvglm` is the matrix of counts. The formula must contain an intercept term. Other **VGAM** family functions for an ordinal response include `acat`, `cratio`, `sratio`. For a nominal (unordered) factor response, the multinomial logit model (`multinomial`) is more appropriate.

With the logit link, setting `parallel = TRUE` will fit a proportional odds model. Note that the `TRUE` here does not apply to the intercept term. In practice, the validity of the *proportional odds assumption* needs to be checked, e.g., by a likelihood ratio test (LRT). If acceptable on the data, then numerical problems are less likely to occur during the fitting, and there are less parameters. Numerical problems occur when the linear/additive predictors cross, which results in probabilities outside of $(0, 1)$; setting `parallel = TRUE` will help avoid this problem.

Here is an example of the usage of the `parallel` argument. If there are covariates `x2`, `x3` and `x4`, then `parallel = TRUE ~ x2 + x3 - 1` and `parallel = FALSE ~ x4` are equivalent. This would constrain the regression coefficients for `x2` and `x3` to be equal; those of the intercepts and `x4` would be different.

If the data is inputted in *long* format (not *wide* format, as in `pneumo` below) and the self-starting initial values are not good enough then try using `mustart`, `coefstart` and/or `etatstart`. See the example below.

To fit the proportional odds model one can use the **VGAM** family function `propodds`. Note that `propodds(reverse)` is equivalent to `cumulative(parallel = TRUE, reverse = reverse)` (which is equivalent to `cumulative(parallel = TRUE, reverse = reverse, link = "logitlink")`). It is for convenience only. A call to `cumulative()` is preferred since it reminds the user that a parallelism assumption is made, as well as being a lot more flexible.

Category specific effects may be modelled using the `xij`-facility; see `vglm.control` and `fill1`.

With most Threshold choices, the first few fitted regression coefficients need care in their interpretation. For example, some values could be the distance away from the median intercept. Typing something like `constraints(fit)[[1]]` gives the constraint matrix of the intercept term.

Author(s)

Thomas W. Yee

References

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Yee, T. W. and Wild, C. J. (1996). Vector generalized additive models. *Journal of the Royal Statistical Society, Series B, Methodological*, **58**, 481–493.

See Also

[propodds](#), [constraints](#), [CM.ones](#), [CM.equid](#), [R2latvar](#), [ordsup](#), [prplot](#), [margeff](#), [acat](#), [cratio](#), [sratio](#), [multinomial](#), [CommonVGAMffArguments](#), [pneumo](#), [budworm](#), [Links](#), [hdeff.vglm](#), [logitlink](#), [probitlink](#), [clogloglink](#), [cauchitlink](#), [logistic1](#).

Examples

```
# Proportional odds model (p.179) of McCullagh and Nelder (1989)
pneumo <- transform(pneumo, let = log(exposure.time))
(fit <- vglm(cbind(normal, mild, severe) ~ let,
            cumulative(parallel = TRUE, reverse = TRUE), pneumo))
depvar(fit) # Sample proportions (good technique)
fit@y      # Sample proportions (bad technique)
weights(fit, type = "prior") # Number of observations
coef(fit, matrix = TRUE)
constraints(fit) # Constraint matrices
apply(fitted(fit), 1, which.max) # Classification
apply(predict(fit, newdata = pneumo, type = "response"),
      1, which.max) # Classification
R2latvar(fit)

# Check that the model is linear in let -----
fit2 <- vgam(cbind(normal, mild, severe) ~ s(let, df = 2),
            cumulative(reverse = TRUE), data = pneumo)
## Not run:
plot(fit2, se = TRUE, overlay = TRUE, lcol = 1:2, scol = 1:2)
## End(Not run)

# Check the proportional odds assumption with a LRT -----
(fit3 <- vglm(cbind(normal, mild, severe) ~ let,
            cumulative(parallel = FALSE, reverse = TRUE), pneumo))
pchisq(2 * (logLik(fit3) - logLik(fit)), df =
      length(coef(fit3)) - length(coef(fit)), lower.tail = FALSE)
lrtest(fit3, fit) # More elegant

# A factor() version of fit -----
# This is in long format (cf. wide format above)
Nobs <- round(depvar(fit) * c(weights(fit, type = "prior")))
sumNobs <- colSums(Nobs) # apply(Nobs, 2, sum)

pneumo.long <-
  data.frame(symptoms = ordered(rep(rep(colnames(Nobs), nrow(Nobs)),
                                times = c(t(Nobs))),
                                levels = colnames(Nobs)),
            let = rep(rep(with(pneumo, let), each = ncol(Nobs)),
                    times = c(t(Nobs))))
with(pneumo.long, table(let, symptoms)) # Should be same as pneumo
```

```
(fit.long1 <- vglm(symptoms ~ let, data = pneumo.long, trace = TRUE,
  cumulative(parallel = TRUE, reverse = TRUE)))
coef(fit.long1, matrix = TRUE) # cf. coef(fit, matrix = TRUE)
# Could try using mustart if fit.long1 failed to converge.
mymustart <- matrix(sumNobs / sum(sumNobs),
  nrow(pneumo.long), ncol(Nobs), byrow = TRUE)
fit.long2 <- vglm(symptoms ~ let, mustart = mymustart,
  cumulative(parallel = TRUE, reverse = TRUE),
  data = pneumo.long, trace = TRUE)
coef(fit.long2, matrix = TRUE) # cf. coef(fit, matrix = TRUE)
```

Dagum

The Dagum Distribution

Description

Density, distribution function, quantile function and random generation for the Dagum distribution with shape parameters a and p , and scale parameter $scale$.

Usage

```
ddagum(x, scale = 1, shape1.a, shape2.p, log = FALSE)
pdagum(q, scale = 1, shape1.a, shape2.p, lower.tail = TRUE,
  log.p = FALSE)
qdagum(p, scale = 1, shape1.a, shape2.p, lower.tail = TRUE,
  log.p = FALSE)
rdagum(n, scale = 1, shape1.a, shape2.p)
```

Arguments

<code>x, q</code>	vector of quantiles.
<code>p</code>	vector of probabilities.
<code>n</code>	number of observations. If <code>length(n) > 1</code> , the length is taken to be the number required.
<code>shape1.a, shape2.p</code>	shape parameters.
<code>scale</code>	scale parameter.
<code>log</code>	Logical. If <code>log = TRUE</code> then the logarithm of the density is returned.
<code>lower.tail, log.p</code>	Same meaning as in pnorm or qnorm .

Details

See [dagum](#), which is the **VGAM** family function for estimating the parameters by maximum likelihood estimation.

Value

ddagum gives the density, pdagum gives the distribution function, qdagum gives the quantile function, and rdagum generates random deviates.

Note

The Dagum distribution is a special case of the 4-parameter generalized beta II distribution.

Author(s)

T. W. Yee and Kai Huang

References

Kleiber, C. and Kotz, S. (2003). *Statistical Size Distributions in Economics and Actuarial Sciences*, Hoboken, NJ, USA: Wiley-Interscience.

See Also

[dagum](#), [genbetaII](#).

Examples

```
probs <- seq(0.1, 0.9, by = 0.1)
shape1.a <- 1; shape2.p <- 2
# Should be 0:
max(abs(pdagum(qdagum(probs, shape1.a = shape1.a, shape2.p =
  shape2.p), shape1.a = shape1.a, shape2.p = shape2.p) - probs))

## Not run: par(mfrow = c(1, 2))
x <- seq(-0.01, 5, len = 401)
plot(x, dexp(x), type = "l", col = "black",
  ylab = "", las = 1, ylim = c(0, 1),
  main = "Black is std exponential, others are ddagum(x, ...)")
lines(x, ddagum(x, shape1.a = shape1.a, shape2.p = 1), col = "orange")
lines(x, ddagum(x, shape1.a = shape1.a, shape2.p = 2), col = "blue")
lines(x, ddagum(x, shape1.a = shape1.a, shape2.p = 5), col = "green")
legend("topright", col = c("orange", "blue", "green"),
  lty = rep(1, len = 3), legend = paste("shape1.a =", shape1.a,
  ", shape2.p =", c(1, 2, 5)))

plot(x, pexp(x), type = "l", col = "black", ylab = "", las = 1,
  main = "Black is std exponential, others are pdagum(x, ...)")
lines(x, pdagum(x, shape1.a = shape1.a, shape2.p = 1), col = "orange")
lines(x, pdagum(x, shape1.a = shape1.a, shape2.p = 2), col = "blue")
lines(x, pdagum(x, shape1.a = shape1.a, shape2.p = 5), col = "green")
legend("bottomright", col = c("orange", "blue", "green"),
  lty = rep(1, len = 3), legend = paste("shape1.a =", shape1.a,
  ", shape2.p =", c(1, 2, 5)))

## End(Not run)
```

dagum

*Dagum Distribution Family Function***Description**

Maximum likelihood estimation of the 3-parameter Dagum distribution.

Usage

```
dagum(lscale = "loglink", lshape1.a = "loglink", lshape2.p =
      "loglink", iscale = NULL, ishape1.a = NULL, ishape2.p =
      NULL, imethod = 1, lss = TRUE, gscale = exp(-5:5), gshape1.a =
      seq(0.75, 4, by = 0.25), gshape2.p = exp(-5:5), probs.y =
      c(0.25, 0.5, 0.75), zero = "shape")
```

Arguments

`lss` See [CommonVGAMffArguments](#) for important information.

`lshape1.a, lscale, lshape2.p` Parameter link functions applied to the (positive) parameters `a`, `scale`, and `p`. See [Links](#) for more choices.

`iscale, ishape1.a, ishape2.p, imethod, zero` See [CommonVGAMffArguments](#) for information. For `imethod = 2` a good initial value for `ishape2.p` is needed to obtain a good estimate for the other parameter.

`gscale, gshape1.a, gshape2.p` See [CommonVGAMffArguments](#) for information.

`probs.y` See [CommonVGAMffArguments](#) for information.

Details

The 3-parameter Dagum distribution is the 4-parameter generalized beta II distribution with shape parameter $q = 1$. It is known under various other names, such as the Burr III, inverse Burr, beta-K, and 3-parameter kappa distribution. It can be considered a generalized log-logistic distribution. Some distributions which are special cases of the 3-parameter Dagum are the inverse Lomax ($a = 1$), Fisk ($p = 1$), and the inverse paralogistic ($a = p$). More details can be found in Kleiber and Kotz (2003).

The Dagum distribution has a cumulative distribution function

$$F(y) = [1 + (y/b)^{-a}]^{-p}$$

which leads to a probability density function

$$f(y) = apy^{ap-1} / [b^{ap} \{1 + (y/b)^a\}^{p+1}]$$

for $a > 0$, $b > 0$, $p > 0$, $y \geq 0$. Here, b is the scale parameter `scale`, and the others are shape parameters. The mean is

$$E(Y) = b\Gamma(p + 1/a)\Gamma(1 - 1/a)/\Gamma(p)$$

provided $-ap < 1 < a$; these are returned as the fitted values. This family function handles multiple responses.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Note

See the notes in [genbetaII](#).

From Kleiber and Kotz (2003), the MLE is rather sensitive to isolated observations located sufficiently far from the majority of the data. Reliable estimation of the scale parameter require $n > 7000$, while estimates for a and p can be considered unbiased for $n > 2000$ or 3000.

Author(s)

T. W. Yee

References

Kleiber, C. and Kotz, S. (2003). *Statistical Size Distributions in Economics and Actuarial Sciences*, Hoboken, NJ, USA: Wiley-Interscience.

See Also

[Dagum](#), [genbetaII](#), [betaII](#), [sinmad](#), [fisk](#), [inv.lomax](#), [lomax](#), [paralogistic](#), [inv.paralogistic](#), [simulate.vlm](#).

Examples

```
## Not run:
ddata <- data.frame(y = rdagum(n = 3000, scale = exp(2),
                             shape1 = exp(1), shape2 = exp(1)))
fit <- vglm(y ~ 1, dagum(lss = FALSE), data = ddata, trace = TRUE)
fit <- vglm(y ~ 1, dagum(lss = FALSE, ishape1.a = exp(1)),
           data = ddata, trace = TRUE)
coef(fit, matrix = TRUE)
Coef(fit)
summary(fit)

## End(Not run)
```

Description

Density for the AR-1 model.

Usage

```
dAR1(x, drift = 0, var.error = 1, ARcoef1 = 0.0,
      type.likelihood = c("exact", "conditional"), log = FALSE)
```

Arguments

x vector of quantiles.

drift the scaled mean (also known as the *drift* parameter), μ^* . Note that the mean is $\mu^*/(1 - \rho)$. The default corresponds to observations that have mean 0.

log Logical. If TRUE then the logarithm of the density is returned.

type.likelihood, var.error, ARcoef1
See [AR1](#). The argument `ARcoef1` is ρ . The argument `var.error` is the variance of the i.i.d. random noise, i.e., σ^2 . If `type.likelihood = "conditional"` then the first element or row of the result is currently assigned NA—this is because the density of the first observation is effectively ignored.

Details

Most of the background to this function is given in [AR1](#). All the arguments are converted into matrices, and then all their dimensions are obtained. They are then coerced into the same size: the number of rows is the maximum of all the single rows, and ditto for the number of columns.

Value

dAR1 gives the density.

Author(s)

T. W. Yee and Victor Miranda

See Also

[AR1](#).

Examples

```
## Not run:
nn <- 100; set.seed(1)
tdata <- data.frame(index = 1:nn,
                    TS1 = arima.sim(nn, model = list(ar = -0.50),
                                       sd = exp(1)))
fit1 <- vglm(TS1 ~ 1, AR1, data = tdata, trace = TRUE)
rhitlink(-0.5)
coef(fit1, matrix = TRUE)
(Cfit1 <- Coef(fit1))
summary(fit1) # SEs are useful to know
logLik(fit1)
sum(dAR1(depvar(fit1), drift = Cfit1[1], var.error = (Cfit1[2])^2,
          ARcoef1 = Cfit1[3], log = TRUE))
```



```

fit2 <- vglm(TS1 ~ 1, AR1(type.likelihood = "cond"), data = tdata, trace = TRUE)
(Cfit2 <- Coef(fit2)) # Okay for intercept-only models
logLik(fit2)
head(keep <- dAR1(depvar(fit2), drift = Cfit2[1], var.error = (Cfit2[2])^2,
                 ARcoef1 = Cfit2[3], type.likelihood = "cond", log = TRUE))
sum(keep[-1])

## End(Not run)

```

deermice

*Captures of *Peromyscus maniculatus* (Also Known as Deer Mice).*

Description

Captures of *Peromyscus maniculatus* collected at East Stuart Gulch, Colorado, USA.

Usage

```
data(deermice)
```

Format

The format is a data frame.

Details

Peromyscus maniculatus is a rodent native to North America. The deer mouse is small in size, only about 8 to 10 cm long, not counting the length of the tail.

Originally, the columns of this data frame represent the sex (m or f), the ages (y: young, sa: semi-adult, a: adult), the weights in grams, and the capture histories of 38 individuals over 6 trapping occasions (1: captured, 0: not captured).

The data set was collected by V. Reid and distributed with the **CAPTURE** program of Otis et al. (1978).

deermice has 38 deermice whereas Perom had 36 deermice (Perom has been withdrawn.) In deermice the two semi-adults have been classified as adults. The sex variable has 1 for female, and 0 for male.

References

Huggins, R. M. (1991). Some practical aspects of a conditional likelihood approach to capture experiments. *Biometrics*, **47**, 725–732.

Otis, D. L. et al. (1978). Statistical inference from capture data on closed animal populations, *Wildlife Monographs*, **62**, 3–135.

See Also

[posbernoulli.b](#), [posbernoulli.t](#), [fill1](#).

Examples

```
head(deermice)
## Not run:
fit1 <- vglm(cbind(y1, y2, y3, y4, y5, y6) ~ sex + age,
             posbernoulli.t(parallel.t = TRUE), deermice, trace = TRUE)
coef(fit1)
coef(fit1, matrix = TRUE)

## End(Not run)
```

deplot.lmscreg	<i>Density Plot for LMS Quantile Regression</i>
----------------	---

Description

Plots a probability density function associated with a LMS quantile regression.

Usage

```
deplot.lmscreg(object, newdata = NULL, x0, y.arg, show.plot =
               TRUE, ...)
```

Arguments

object	A VGAM quantile regression model, i.e., an object produced by modelling functions such as <code>vglm</code> and <code>vgam</code> with a family function beginning with "lms.", e.g., <code>lms.yjn</code> .
newdata	Optional data frame containing secondary variables such as sex. It should have a maximum of one row. The default is to use the original data.
x0	Numeric. The value of the primary variable at which to make the 'slice'.
y.arg	Numerical vector. The values of the response variable at which to evaluate the density. This should be a grid that is fine enough to ensure the plotted curves are smooth.
show.plot	Logical. Plot it? If FALSE no plot will be done.
...	Graphical parameter that are passed into <code>plotdeplot.lmscreg</code> .

Details

This function calls, e.g., `deplot.lms.yjn` in order to compute the density function.

Value

The original object but with a list placed in the slot `post`, called `@post$deplot`. The list has components

<code>newdata</code>	The argument <code>newdata</code> above, or a one-row data frame constructed out of the <code>x0</code> argument.
<code>y</code>	The argument <code>y.arg</code> above.
<code>density</code>	Vector of the density function values evaluated at <code>y.arg</code> .

Note

[plotdeplot.lmscreg](#) actually does the plotting.

Author(s)

Thomas W. Yee

References

Yee, T. W. (2004). Quantile regression via vector generalized additive models. *Statistics in Medicine*, **23**, 2295–2315.

See Also

[plotdeplot.lmscreg](#), [qtplot.lmscreg](#), [lms.bcn](#), [lms.bcg](#), [lms.yjn](#).

Examples

```
## Not run:
fit <- vgam(BMI ~ s(age, df = c(4, 2)), lms.bcn(zero = 1), bmi.nz)
ygrid <- seq(15, 43, by = 0.25)
deplot(fit, x0 = 20, y = ygrid, xlab = "BMI", col = "green", llwd = 2,
       main = "BMI distribution at ages 20 (green), 40 (blue), 60 (red)")
deplot(fit, x0 = 40, y = ygrid, add = TRUE, col = "blue", llwd = 2)
deplot(fit, x0 = 60, y = ygrid, add = TRUE, col = "red", llwd = 2) -> a

names(a@post$deplot)
a@post$deplot$newdata
head(a@post$deplot$y)
head(a@post$deplot$density)

## End(Not run)
```

`depvar`*Response Variable Extracted*

Description

A generic function that extracts the response/dependent variable from objects.

Usage

```
depvar(object, ...)
```

Arguments

<code>object</code>	An object that has some response/dependent variable.
<code>...</code>	Other arguments fed into the specific methods function of the model. In particular, sometimes <code>type = c("lm", "lm2")</code> is available, in which case the first one is chosen if the user does not input a value. The latter value corresponds to argument <code>form2</code> , and sometimes a response for that is optional.

Details

By default this function is preferred to calling `fit@y`, say.

Value

The response/dependent variable, usually as a matrix or vector.

Author(s)

Thomas W. Yee

See Also

[model.matrix](#), [vglm](#).

Examples

```
pneumo <- transform(pneumo, let = log(exposure.time))
(fit <- vglm(cbind(normal, mild, severe) ~ let, propodds, pneumo))
fit@y      # Sample proportions (not recommended)
depvar(fit) # Better than using fit@y
weights(fit, type = "prior") # Number of observations
```

dextlogF*Extended log-F Distribution*

Description

Density for the extended log-F distribution.

Usage

```
dextlogF(x, lambda, tau, location = 0, scale = 1, log = FALSE)
```

Arguments

x	Vector of quantiles.
lambda, tau	See extlogF1 .
location, scale	See extlogF1 .
log	If TRUE then the log density is returned, else the density.

Details

The details are given in [extlogF1](#).

Value

dextlogF gives the density.

Author(s)

T. W. Yee

See Also

[extlogF1](#), [dalap](#).

Examples

```
## Not run: x <- seq(-2, 8, by = 0.1); mytau <- 0.25; mylambda <- 0.2
plot(x, dextlogF(x, mylambda, tau = mytau), type = "l",
     las = 1, col = "blue", ylab = "PDF (log-scale)", log = "y",
     main = "Extended log-F density function is blue",
     sub = "Asymmetric Laplace is orange dashed")
lines(x, dalap(x, tau = mytau, scale = 3.5), col = "orange", lty = 2)
abline(v = 0, col = "gray", lty = 2)
## End(Not run)
```

df.residual	<i>Residual Degrees-of-Freedom</i>
-------------	------------------------------------

Description

Returns the residual degrees-of-freedom extracted from a fitted VGLM object.

Usage

```
df.residual_vlm(object, type = c("vlm", "lm"), ...)
```

Arguments

object	an object for which the degrees-of-freedom are desired, e.g., a vglm object.
type	the type of residual degrees-of-freedom wanted. In some applications the 'usual' LM-type value may be more appropriate. The default is the first choice.
...	additional optional arguments.

Details

When a VGLM is fitted, a *large* (VLM) generalized least squares (GLS) fit is done at each IRLS iteration. To do this, an ordinary least squares (OLS) fit is performed by transforming the GLS using Cholesky factors. The number of rows is M times the 'ordinary' number of rows of the LM-type model: nM . Here, M is the number of linear/additive predictors. So the formula for the VLM-type residual degrees-of-freedom is $nM - p^*$ where p^* is the number of columns of the 'big' VLM matrix. The formula for the LM-type residual degrees-of-freedom is $n - p_j$ where p_j is the number of columns of the 'ordinary' LM matrix corresponding to the j th linear/additive predictor.

Value

The value of the residual degrees-of-freedom extracted from the object. When `type = "vlm"` this is a single integer, and when `type = "lm"` this is a M -vector of integers.

See Also

[vglm](#), [deviance](#), [lm](#), [anova.vglm](#),

Examples

```
pneumo <- transform(pneumo, let = log(exposure.time))
(fit <- vglm(cbind(normal, mild, severe) ~ let, propodds, pneumo))
head(model.matrix(fit, type = "vlm"))
head(model.matrix(fit, type = "lm"))

df.residual(fit, type = "vlm") # n * M - p_VLM
nobs(fit, type = "vlm") # n * M
nvar(fit, type = "vlm") # p_VLM
```

```
df.residual(fit, type = "lm") # n - p_LM(j)
nobs(fit, type = "lm") # n
nvar(fit, type = "lm") # p_LM
nvar_vlm(fit, type = "lm") # p_LM(j) (<= p_LM elementwise)
```

dgaitdplot

Plotting the GAITD Combo Density

Description

Plots a 1- or 2-parameter GAITD combo probability mass function.

Usage

```
dgaitdplot(theta.p, fam = "pois", a.mix = NULL, i.mix = NULL,
  d.mix = NULL, a.mlm = NULL, i.mlm = NULL,
  d.mlm = NULL, truncate = NULL, max.support = Inf,
  pobs.mix = 0, pobs.mlm = 0,
  pstr.mix = 0, pstr.mlm = 0,
  pdip.mix = 0, pdip.mlm = 0, byrow.aid = FALSE,
  theta.a = theta.p, theta.i = theta.p, theta.d = theta.p,
  deflation = FALSE, plot.it = TRUE, new.plot = TRUE,
  offset.x = ifelse(new.plot, 0, 0.25), type.plot = "h",
  xlim = c(0, min(100, max.support + 2)),
  ylim = NULL, xlab = "", ylab = "Probability", main = "",
  cex.main = 1.2, posn.main = NULL,
  all.col = NULL, all.lty = NULL, all.lwd = NULL,
  lty.p = "solid", lty.a.mix = "longdash", lty.a.mlm = "longdash",
  lty.i.mix = "dashed", lty.i.mlm = "dashed",
  lty.d.mix = "solid", lty.d.mlm = "solid", lty.d.dip = "dashed",
  col.p = "pink2",
  col.a.mix = artichoke.col, col.a.mlm = asparagus.col,
  col.i.mix = indigo.col, col.i.mlm = iris.col,
  col.d.mix = deer.col, col.d.mlm = dirt.col, col.d.dip = desire.col,
  col.t = turquoise.col, cex.p = 1, lwd.p = NULL, lwd.a = NULL,
  lwd.i = NULL, lwd.d = NULL, iontop = TRUE, dontop = TRUE,
  las = 0, lend = "round", axes.x = TRUE, axes.y = TRUE,
  Plot.trunc = TRUE, cex.t = 1, pch.t = 1,
  baseparams.argnames = NULL, nparams = 1, flip.args = FALSE, ...)
```

Arguments

theta.p	Numeric, usually scalar but may have length 2. This matches with, e.g., lambda.p for Gaitdpois . A length 2 example is c(size.p, munb.p) for Gaitdnbinom , in which case fam = "nbinom". Another length 2 example is c(mean.p, dispind.p) for Gaitgenpois1 , in which case fam = "genpois1".
fam	Character, paste0("dgait", fam) should be a d-type function returning the PMF. The default is for the GAITD Poisson combo.

- `a.mix`, `i.mix`, `a.mlm`, `i.mlm`
See [Gaitdpois](#) and [gaitdpoisson](#).
- `d.mix`, `d.mlm` See [Gaitdpois](#) and [gaitdpoisson](#).
- `truncate`, `max.support`
See [Gaitdpois](#) and [gaitdpoisson](#).
- `pobs.mix`, `pobs.mlm`, `byrow.aid`
See [Gaitdpois](#) and [gaitdpoisson](#).
- `pstr.mix`, `pstr.mlm`, `pdip.mix`, `pdip.mlm`
See [Gaitdpois](#) and [gaitdpoisson](#).
- `theta.a`, `theta.i`, `theta.d`
Similar to `theta.p`, and they should have the same length too.
- `deflation` Logical. Plot the deflation (dip) probabilities?
- `plot.it` Logical. Plot the PMF?
- `new.plot`, `offset.x`
If `new.plot` then `plot` is called. If multiple plots are desired then use `offset.x` to shift the lines.
- `xlim`, `ylim`, `xlab`, `ylab`
See [par](#) and [plot.default](#). Argument `xlim` should be integer-valued.
- `main`, `cex.main`, `posn.main`
Character, size and position of `main` for the title. See [title](#), [par](#) and [plot.default](#). The position is used if it is a 2-vector.
- `all.col`, `all.lty`, `all.lwd`
These arguments allow all the colours, line types and line widths arguments to be assigned to these values, i.e., so that they are the same for all values of the support. For example, if `all.lwd = 2` then this sets `lwd.p`, `lwd.a`, `lwd.i` and `lwd.d` all equal to 2.
- `lty.p`, `lty.a.mix`, `lty.a.mlm`, `lty.i.mix`, `lty.i.mlm`
Line type for parent, altered and inflated. See [par](#) and [plot.default](#).
- `col.p`, `col.a.mix`, `col.a.mlm`, `col.i.mix`, `col.i.mlm`
Line colour for parent (nonspecial), altered, inflated, truncated and deflated values. See [par](#) and [plot.default](#). Roughly, by default and currently, the parent is pink-like, the altered are greenish, the inflated are purplish/violet, the truncated are light blue, and the deflated are brownish with the dip probabilities being reddish. The proper colour names are similar to being acrostic. For each operator, the colours of "mix" vs "mlm" are similar but different—this is intentional. Warning: the default colours might change, depending on style!
- `lty.d.mix`, `lty.d.mlm`, `lty.d.dip`
Similar to above. Used when `deflation = TRUE`.
- `col.d.mix`, `col.d.mlm`, `col.d.dip`
Similar to above. Used when `deflation = TRUE`. The website <https://www.spycolor.com> was used to choose some of the default colours; the first two are also called "dirt" and "deer" respectively, which are both brownish.
- `col.t` Point colour for truncated values, the default is "tan".
- `type.plot`, `cex.p`
The former matches 'type' argument in [plot.default](#). The latter is the size of the point if `type.plot = "p"` or `type.plot = "b"`, etc.

<code>lwd.p, lwd.a, lwd.i, lwd.d</code>	Line width for parent, altered and inflated. See par and plot.default . By default <code>par()\\$lwd</code> is used for all of them.
<code>las, lend</code>	See par .
<code>iontop, dontop</code>	Logicals. Draw the inflated and deflated bars on top? The default is to draw the spikes on top, but if FALSE then the spikes are drawn from the bottom—this makes it easier to see their distribution. Likewise, if <code>deflation = TRUE</code> then <code>dontop</code> is used to position the deflation (dip) probabilities.
<code>axes.x, axes.y</code>	Logical. Plot axes? See par and plot.default .
<code>Plot.trunc, cex.t, pch.t</code>	Logical. Plot the truncated values? If so, then specify the size and plotting character. See par and plot.default .
<code>baseparams.argnames</code>	Character string specifying the argument name for the generic parameter <code>theta</code> , e.g., "lambda" for gaitdpoisson . By appending <code>.p</code> , there is an argument called <code>lambda.p</code> in dgaitdpois . Another example is for gaitdlog : "shape" appended with <code>.p</code> means that dgaitdlog should have an argument called <code>shape.p</code> . This argument is optional and increases the reliability of the <code>do.call</code> call internally.
<code>nparams, flip.args</code>	Not for use by the user. It is used internally to handle the NBD.
<code>...</code>	Currently unused but there is provision for passing graphical arguments in the future; see par .

Details

This is meant to be a crude function to plot the PMF of the GAITD combo model. Some flexibility is offered via many graphical arguments, but there are still many improvements that could be done.

Value

A list is returned invisibly. The components are:

<code>x</code>	The integer values between the values of <code>xlim</code> .
<code>pmf.z</code>	The value of the PMF, by calling the d-type function with all the arguments fed in.
<code>sc.parent</code>	The same level as the scaled parent distribution. Thus for inflated values, the value where the spikes begin. And for deflated values, the value at the top of the dips. This is a convenient way to obtain them as it is quite cumbersome to compute them manually. For any nonspecial value, such as non-inflated and non-deflated values, they are equal to <code>pmf.z</code> .
<code>unsc.parent</code>	Unscaled parent distribution. If there is no alteration, inflation, deflation and truncation then this is the basic PMF stipulated by the parent distribution only. Usually this is FYI only.

Note

This utility function may change a lot in the future. Because this function is called by a **shiny** app, if any parameter values lie outside the parameter space then `stop` will be called. For example, too much deflation results in NaN values returned by `dgaitdnbinom`.

Author(s)

T. W. Yee.

See Also

[plotdgaitd](#), [spikeplot](#), [meangaitd](#), [Gaitdpois](#), [gaitdpoisson](#), [Gaitdnbinom](#), [multilogitlink](#).

Examples

```
## Not run:  i.mix <- seq(0, 25, by = 5)
mean.p <- 10; size.p <- 8
dgaitdplot(c(size.p, mean.p), fam = "nbinom", xlim = c(0, 25),
  a.mix = i.mix + 1, i.mix = i.mix, pobs.mix = 0.1,
  pstr.mix = 0.1, lwd.i = 2, lwd.p = 2, lwd.a = 2)

## End(Not run)
```

dhuber

Huber's Least Favourable Distribution

Description

Density, distribution function, quantile function and random generation for Huber's least favourable distribution, see Huber and Ronchetti (2009).

Usage

```
dhuber(x, k = 0.862, mu = 0, sigma = 1, log = FALSE)
edhuber(x, k = 0.862, mu = 0, sigma = 1, log = FALSE)
rhuber(n, k = 0.862, mu = 0, sigma = 1)
qhuber(p, k = 0.862, mu = 0, sigma = 1, lower.tail = TRUE,
  log.p = FALSE)
phuber(q, k = 0.862, mu = 0, sigma = 1, lower.tail = TRUE,
  log.p = FALSE)
```

Arguments

x, q	numeric vector, vector of quantiles.
p	vector of probabilities.
n	number of random values to be generated. If <code>length(n) > 1</code> then the length is taken to be the number required.

k	numeric. Borderline value of central Gaussian part of the distribution. This is known as the tuning constant, and should be positive. For example, $k = 0.862$ refers to a 20% contamination neighborhood of the Gaussian distribution. If $k = 1.40$ then this is 5% contamination.
mu	numeric. distribution mean.
sigma	numeric. Distribution scale ($\sigma = 1$ defines the distribution in standard form, with standard Gaussian centre).
log	Logical. If <code>log = TRUE</code> then the logarithm of the result is returned.
lower.tail, log.p	Same meaning as in pnorm or qnorm .

Details

Details are given in [huber2](#), the **VGAM** family function for estimating the parameters `mu` and `sigma`.

Value

`dhuber` gives out a vector of density values.

`edhuber` gives out a list with components `val` (density values) and `eps` (contamination proportion).

`rhuber` gives out a vector of random numbers generated by Huber's least favourable distribution.

`phuber` gives the distribution function, `qhuber` gives the quantile function.

Author(s)

Christian Hennig wrote `[d,ed,r]huber()` (from **smoothest**) and slight modifications were made by T. W. Yee to replace looping by vectorization and addition of the `log` argument. Arash Ardalan wrote `[pq]huber()`, and two arguments for these were implemented by Kai Huang. This helpfile was adapted from **smoothest**.

See Also

[huber2](#).

Examples

```
set.seed(123456)
edhuber(1:5, k = 1.5)
rhuber(5)

## Not run: mu <- 3; xx <- seq(-2, 7, len = 100) # Plot CDF and PDF
plot(xx, dhuber(xx, mu = mu), type = "l", col = "blue", las = 1,
     main = "blue is density, orange is the CDF", ylab = "",
     sub = "Purple lines are the 10,20,...,90 percentiles",
     ylim = 0:1)
abline(h = 0, col = "blue", lty = 2)
lines(xx, phuber(xx, mu = mu), type = "l", col = "orange")
probs <- seq(0.1, 0.9, by = 0.1)
Q <- qhuber(probs, mu = mu)
```

```

lines(Q, dhuber(Q, mu = mu), col = "purple", lty = 3, type = "h")
lines(Q, phuber(Q, mu = mu), col = "purple", lty = 3, type = "h")
abline(h = probs, col = "purple", lty = 3)
phuber(Q, mu = mu) - probs # Should be all 0s

## End(Not run)

```

Diffzeta

Differenced Zeta Distribution

Description

Density, distribution function, quantile function, and random generation for the differenced zeta distribution.

Usage

```

ddiffzeta(x, shape, start = 1, log = FALSE)
pdiffzeta(q, shape, start = 1, lower.tail = TRUE)
qdiffzeta(p, shape, start = 1)
rdiffzeta(n, shape, start = 1)

```

Arguments

x, q, p, n Same as in [runif](#).
shape, start Details at [diffzeta](#).
log, lower.tail Same as in [runif](#).

Details

This distribution appears to work well on the distribution of English words in such texts. Some more details are given in [diffzeta](#).

Value

`ddiffzeta` gives the density, `pdiffzeta` gives the distribution function, `qdiffzeta` gives the quantile function, and `rdiffzeta` generates random deviates.

Note

Given some response data, the **VGAM** family function [diffzeta](#) estimates the parameter shape. Function `pdiffzeta()` suffers from the problems that [plog](#) sometimes has, i.e., when p is very close to 1.

Author(s)

T. W. Yee

See Also

[diffzeta](#), [zetaff](#), [zipf](#), [Oizeta](#).

Examples

```
ddiffzeta(1:20, 0.5, start = 2)
rdiffzeta(20, 0.5)

## Not run: shape <- 0.8; x <- 1:10
plot(x, ddiffzeta(x, sh = shape), type = "h", ylim = 0:1, las = 1,
     sub = "shape=0.8", col = "blue", ylab = "Probability",
     main = "Differenced zeta distribution: blue=PMF; orange=CDF")
lines(x + 0.1, pdiffzeta(x, shape = shape), col = "orange",
     lty = 3, type = "h")
## End(Not run)
```

diffzeta

Differenced Zeta Distribution Family Function

Description

Estimates the parameter of the differenced zeta distribution.

Usage

```
diffzeta(start = 1, lshape = "loglink", ishape = NULL)
```

Arguments

`lshape`, `ishape` Same as [zetaff](#).
`start` Smallest value of the support of the distribution. Must be a positive integer.

Details

The PMF is

$$P(Y = y) = (a/y)^s - (a/(1+y))^s, \quad s > 0, \quad y = a, a + 1, \dots,$$

where s is the positive shape parameter, and a is `start`. According to Moreno-Sanchez et al. (2016), this model fits quite well to about 40 percent of all the English books in the Project Gutenberg data base (about 30,000 texts). Multiple responses are handled.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Author(s)

T. W. Yee

References

Moreno-Sanchez, I., Font-Clos, F. and Corral, A. (2016). Large-Scale Analysis of Zipf's Law in English Texts, *PLoS ONE*, **11**(1), 1–19.

See Also

[Diffzeta](#), [zetaff](#), [zeta](#), [zipf](#), [zipf](#).

Examples

```
odata <- data.frame(x2 = runif(nn <- 1000)) # Artificial data
odata <- transform(odata, shape = loglink(-0.25 + x2, inv = TRUE))
odata <- transform(odata, y1 = rdifzeta(nn, shape))
with(odata, table(y1))
ofit <- vglm(y1 ~ x2, difzeta, odata, trace = TRUE)
coef(ofit, matrix = TRUE)
```

dirichlet

Fitting a Dirichlet Distribution

Description

Fits a Dirichlet distribution to a matrix of compositions.

Usage

```
dirichlet(link = "loglink", parallel = FALSE, zero = NULL,
          imethod = 1)
```

Arguments

link Link function applied to each of the M (positive) shape parameters α_j . See [Links](#) for more choices. The default gives $\eta_j = \log(\alpha_j)$.

parallel, zero, imethod

See [CommonVGAMffArguments](#) for more information.

Details

In this help file the response is assumed to be a M -column matrix with positive values and whose rows each sum to unity. Such data can be thought of as compositional data. There are M linear/additive predictors η_j .

The Dirichlet distribution is commonly used to model compositional data, including applications in genetics. Suppose $(Y_1, \dots, Y_M)^T$ is the response. Then it has a Dirichlet distribution if $(Y_1, \dots, Y_{M-1})^T$ has density

$$\frac{\Gamma(\alpha_+)}{\prod_{j=1}^M \Gamma(\alpha_j)} \prod_{j=1}^M y_j^{\alpha_j - 1}$$

where $\alpha_+ = \alpha_1 + \dots + \alpha_M$, $\alpha_j > 0$, and the density is defined on the unit simplex

$$\Delta_M = \left\{ (y_1, \dots, y_M)^T : y_1 > 0, \dots, y_M > 0, \sum_{j=1}^M y_j = 1 \right\}.$$

One has $E(Y_j) = \alpha_j / \alpha_+$, which are returned as the fitted values. For this distribution Fisher scoring corresponds to Newton-Raphson.

The Dirichlet distribution can be motivated by considering the random variables $(G_1, \dots, G_M)^T$ which are each independent and identically distributed as a gamma distribution with density $f(g_j) = g_j^{\alpha_j - 1} e^{-g_j} / \Gamma(\alpha_j)$. Then the Dirichlet distribution arises when $Y_j = G_j / (G_1 + \dots + G_M)$.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), [rrvglm](#) and [vgam](#).

When fitted, the `fitted.values` slot of the object contains the M -column matrix of means.

Note

The response should be a matrix of positive values whose rows each sum to unity. Similar to this is count data, where probably a multinomial logit model ([multinomial](#)) may be appropriate. Another similar distribution to the Dirichlet is the Dirichlet-multinomial (see [dirmultinomial](#)).

Author(s)

Thomas W. Yee

References

- Lange, K. (2002). *Mathematical and Statistical Methods for Genetic Analysis*, 2nd ed. New York: Springer-Verlag.
- Forbes, C., Evans, M., Hastings, N. and Peacock, B. (2011). *Statistical Distributions*, Hoboken, NJ, USA: John Wiley and Sons, Fourth edition.

See Also

[rdiric](#), [dirmultinomial](#), [multinomial](#), [simplex](#).

Examples

```

ddata <- data.frame(rdiric(1000,
                          shape = exp(c(y1 = -1, y2 = 1, y3 = 0)))
fit <- vglm(cbind(y1, y2, y3) ~ 1, dirichlet,
            data = ddata, trace = TRUE, crit = "coef")
Coef(fit)
coef(fit, matrix = TRUE)
head(fitted(fit))

```

dirmul.old

*Fitting a Dirichlet-Multinomial Distribution***Description**

Fits a Dirichlet-multinomial distribution to a matrix of non-negative integers.

Usage

```

dirmul.old(link = "loglink", ialpha = 0.01, parallel = FALSE,
           zero = NULL)

```

Arguments

link	Link function applied to each of the M (positive) shape parameters α_j for $j = 1, \dots, M$. See Links for more choices. Here, M is the number of columns of the response matrix.
ialpha	Numeric vector. Initial values for the alpha vector. Must be positive. Recycled to length M .
parallel	A logical, or formula specifying which terms have equal/unequal coefficients.
zero	An integer-valued vector specifying which linear/additive predictors are modelled as intercepts only. The values must be from the set $\{1, 2, \dots, M\}$. See CommonVGAMffArguments for more information.

Details

The Dirichlet-multinomial distribution, which is somewhat similar to a Dirichlet distribution, has probability function

$$P(Y_1 = y_1, \dots, Y_M = y_M) = \binom{2y_*}{y_1, \dots, y_M} \frac{\Gamma(\alpha_+)}{\Gamma(2y_* + \alpha_+)} \prod_{j=1}^M \frac{\Gamma(y_j + \alpha_j)}{\Gamma(\alpha_j)}$$

for $\alpha_j > 0$, $\alpha_+ = \alpha_1 + \dots + \alpha_M$, and $2y_* = y_1 + \dots + y_M$. Here, $\binom{a}{b}$ means “ a choose b ” and refers to combinations (see [choose](#)). The (posterior) mean is

$$E(Y_j) = (y_j + \alpha_j) / (2y_* + \alpha_+)$$

for $j = 1, \dots, M$, and these are returned as the fitted values as a M -column matrix.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), [rrvglm](#) and [vgam](#).

Note

The response should be a matrix of non-negative values. Convergence seems to slow down if there are zero values. Currently, initial values can be improved upon.

This function is almost defunct and may be withdrawn soon. Use [dirmultinomial](#) instead.

Author(s)

Thomas W. Yee

References

Lange, K. (2002). *Mathematical and Statistical Methods for Genetic Analysis*, 2nd ed. New York: Springer-Verlag.

Forbes, C., Evans, M., Hastings, N. and Peacock, B. (2011). *Statistical Distributions*, Hoboken, NJ, USA: John Wiley and Sons, Fourth edition.

Paul, S. R., Balasooriya, U. and Banerjee, T. (2005). Fisher information matrix of the Dirichlet-multinomial distribution. *Biometrical Journal*, **47**, 230–236.

Tvedebrink, T. (2010). Overdispersion in allelic counts and θ -correction in forensic genetics. *Theoretical Population Biology*, **78**, 200–210.

See Also

[dirmultinomial](#), [dirichlet](#), [betabinomialff](#), [multinomial](#).

Examples

```
# Data from p.50 of Lange (2002)
alleleCounts <- c(2, 84, 59, 41, 53, 131, 2, 0,
                 0, 50, 137, 78, 54, 51, 0, 0,
                 0, 80, 128, 26, 55, 95, 0, 0,
                 0, 16, 40, 8, 68, 14, 7, 1)
dim(alleleCounts) <- c(8, 4)
alleleCounts <- data.frame(t(alleleCounts))
dimnames(alleleCounts) <- list(c("White", "Black", "Chicano", "Asian"),
                               paste("Allele", 5:12, sep = ""))

set.seed(123) # @initialize uses random numbers
fit <- vglm(cbind(Allele5, Allele6, Allele7, Allele8, Allele9,
                 Allele10, Allele11, Allele12) ~ 1, dirmul.old,
           trace = TRUE, crit = "c", data = alleleCounts)

(sfit <- summary(fit))
vcov(sfit)
round(eta2theta(coef(fit),
```

```

fit@misc$link,
fit@misc$earg), digits = 2) # not preferred
round(Coef(fit), digits = 2) # preferred
round(t(fitted(fit)), digits = 4) # 2nd row of Lange (2002, Table 3.5)
coef(fit, matrix = TRUE)

pfit <- vglm(cbind(Allele5,Allele6,Allele7,Allele8,Allele9,
Allele10,Allele11,Allele12) ~ 1,
dirmul.old(parallel = TRUE), trace = TRUE,
data = alleleCounts)
round(eta2theta(coef(pfit, matrix = TRUE), pfit@misc$link,
pfit@misc$earg), digits = 2) # 'Right' answer
round(Coef(pfit), digits = 2) # 'Wrong' due to parallelism constraint

```

dirmultinomial

Fitting a Dirichlet-Multinomial Distribution

Description

Fits a Dirichlet-multinomial distribution to a matrix response.

Usage

```
dirmultinomial(lphi = "logitlink", iphi = 0.10, parallel = FALSE,
zero = "M")
```

Arguments

lphi	Link function applied to the ϕ parameter, which lies in the open unit interval $(0, 1)$. See Links for more choices.
iphi	Numeric. Initial value for ϕ . Must be in the open unit interval $(0, 1)$. If a failure to converge occurs then try assigning this argument a different value.
parallel	A logical (formula not allowed here) indicating whether the probabilities π_1, \dots, π_{M-1} are to be equal via equal coefficients. Note π_M will generally be different from the other probabilities. Setting <code>parallel = TRUE</code> will only work if you also set <code>zero = NULL</code> because of interference between these arguments (with respect to the intercept term).
zero	An integer-valued vector specifying which linear/additive predictors are modelled as intercepts only. The values must be from the set $\{1, 2, \dots, M\}$. If the character "M" then this means the numerical value M , which corresponds to linear/additive predictor associated with ϕ . Setting <code>zero = NULL</code> means none of the values from the set $\{1, 2, \dots, M\}$. See CommonVGAMffArguments for more information.

Details

The Dirichlet-multinomial distribution arises from a multinomial distribution where the probability parameters are not constant but are generated from a multivariate distribution called the Dirichlet distribution. The Dirichlet-multinomial distribution has probability function

$$P(Y_1 = y_1, \dots, Y_M = y_M) = \binom{N_*}{y_1, \dots, y_M} \frac{\prod_{j=1}^M \prod_{r=1}^{y_j} (\pi_j(1 - \phi) + (r - 1)\phi)}{\prod_{r=1}^{N_*} (1 - \phi + (r - 1)\phi)}$$

where ϕ is the *over-dispersion* parameter and $N_* = y_1 + \dots + y_M$. Here, $\binom{a}{b}$ means “*a* choose *b*” and refers to combinations (see [choose](#)). The above formula applies to each row of the matrix response. In this **VGAM** family function the first $M - 1$ linear/additive predictors correspond to the first $M - 1$ probabilities via

$$\eta_j = \log(P[Y = j]/P[Y = M]) = \log(\pi_j/\pi_M)$$

where η_j is the j th linear/additive predictor ($\eta_M = 0$ by definition for $P[Y = M]$ but not for ϕ) and $j = 1, \dots, M - 1$. The M th linear/additive predictor corresponds to `lphi` applied to ϕ .

Note that $E(Y_j) = N_*\pi_j$ but the probabilities (returned as the fitted values) π_j are bundled together as a M -column matrix. The quantities N_* are returned as the prior weights.

The beta-binomial distribution is a special case of the Dirichlet-multinomial distribution when $M = 2$; see [betabinomial](#). It is easy to show that the first shape parameter of the beta distribution is $shape1 = \pi(1/\phi - 1)$ and the second shape parameter is $shape2 = (1 - \pi)(1/\phi - 1)$. Also, $\phi = 1/(1 + shape1 + shape2)$, which is known as the *intra-cluster correlation* coefficient.

Value

An object of class “`vglmff`” (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), [rrvglm](#) and [vgam](#).

If the model is an intercept-only model then `@misc` (which is a list) has a component called `shape` which is a vector with the M values $\pi_j(1/\phi - 1)$.

Warning

This **VGAM** family function is prone to numerical problems, especially when there are covariates.

Note

The response can be a matrix of non-negative integers, or else a matrix of sample proportions and the total number of counts in each row specified using the `weights` argument. This dual input option is similar to [multinomial](#).

To fit a ‘parallel’ model with the ϕ parameter being an intercept-only you will need to use the `constraints` argument.

Currently, Fisher scoring is implemented. To compute the expected information matrix a for loop is used; this may be very slow when the counts are large. Additionally, convergence may be slower than usual due to round-off error when computing the expected information matrices.

Author(s)

Thomas W. Yee

References

- Paul, S. R., Balasooriya, U. and Banerjee, T. (2005). Fisher information matrix of the Dirichlet-multinomial distribution. *Biometrical Journal*, **47**, 230–236.
- Tvedebrink, T. (2010). Overdispersion in allelic counts and θ -correction in forensic genetics. *Theoretical Population Biology*, **78**, 200–210.
- Yu, P. and Shaw, C. A. (2014). An Efficient Algorithm for Accurate Computation of the Dirichlet-Multinomial Log-Likelihood Function. *Bioinformatics*, **30**, 1547–54.

See Also

[dirmul.old](#), [betabinomial](#), [betabinomialff](#), [dirichlet](#), [multinomial](#).

Examples

```
nn <- 5; M <- 4; set.seed(1)
ydata <- data.frame(round(matrix(runif(nn * M, max = 100), nn, M)))
colnames(ydata) <- paste("y", 1:M, sep = "") # Integer counts

fit <- vglm(cbind(y1, y2, y3, y4) ~ 1, dirmultinomial,
            data = ydata, trace = TRUE)
head(fitted(fit))
devar(fit) # Sample proportions
weights(fit, type = "prior", matrix = FALSE) # Total counts per row

## Not run:
ydata <- transform(ydata, x2 = runif(nn))
fit <- vglm(cbind(y1, y2, y3, y4) ~ x2, dirmultinomial,
            data = ydata, trace = TRUE)

Coef(fit)
coef(fit, matrix = TRUE)
(sfit <- summary(fit))
vcov(sfit)

## End(Not run)
```

dlogF

log F Distribution

Description

Density for the log F distribution.

Usage

```
dlogF(x, shape1, shape2, log = FALSE)
```

Arguments

x Vector of quantiles.
 shape1, shape2 Positive shape parameters.
 log if TRUE then the log density is returned, else the density.

Details

The details are given in [logF](#).

Value

dlogF gives the density.

Author(s)

T. W. Yee

See Also

[hypersecant](#), [dextlogF](#).

Examples

```
## Not run: shape1 <- 1.5; shape2 <- 0.5; x <- seq(-5, 8, length = 1001)
plot(x, dlogF(x, shape1, shape2), type = "l",
     las = 1, col = "blue", ylab = "pdf",
     main = "log F density function")

## End(Not run)
```

double.cens.normal *Univariate Normal Distribution with Double Censoring*

Description

Maximum likelihood estimation of the two parameters of a univariate normal distribution when there is double censoring.

Usage

```
double.cens.normal(r1 = 0, r2 = 0, lmu = "identitylink", lsd =
  "loglink", imu = NULL, isd = NULL, zero = "sd")
```

Arguments

r1, r2 Integers. Number of smallest and largest values censored, respectively.
 lmu, lsd Parameter link functions applied to the mean and standard deviation. See [Links](#) for more choices.
 imu, isd, zero See [CommonVGAMffArguments](#) for more information.

Details

This family function uses the Fisher information matrix given in Harter and Moore (1966). The matrix is not diagonal if either r_1 or r_2 are positive.

By default, the mean is the first linear/additive predictor and the log of the standard deviation is the second linear/additive predictor.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Note

This family function only handles a vector or one-column matrix response. The weights argument, if used, are interpreted as frequencies, therefore it must be a vector with positive integer values.

With no censoring at all (the default), it is better (and equivalent) to use [uninormal](#).

Author(s)

T. W. Yee

References

Harter, H. L. and Moore, A. H. (1966). Iterative maximum-likelihood estimation of the parameters of normal populations from singly and doubly censored samples. *Biometrika*, **53**, 205–213.

See Also

[uninormal](#), [cens.normal](#), [tobit](#).

Examples

```
## Not run: # Repeat the simulations of Harter & Moore (1966)
SIMS <- 100 # Number of simulations (change this to 1000)
mu.save <- sd.save <- rep(NA, len = SIMS)
r1 <- 0; r2 <- 4; nn <- 20
for (sim in 1:SIMS) {
  y <- sort(rnorm(nn))
  y <- y[(1+r1):(nn-r2)] # Delete r1 smallest and r2 largest
  fit <- vglm(y ~ 1, double.cens.normal(r1 = r1, r2 = r2))
  mu.save[sim] <- predict(fit)[1, 1]
  sd.save[sim] <- exp(predict(fit)[1, 2]) # Assumes a log link & ~ 1
}
c(mean(mu.save), mean(sd.save)) # Should be c(0,1)
c(sd(mu.save), sd(sd.save))

## End(Not run)

# Data from Sarhan & Greenberg (1962); MLEs are mu=9.2606, sd=1.3754
strontium90 <- data.frame(y = c(8.2, 8.4, 9.1, 9.8, 9.9))
```

```
fit <- vglm(y ~ 1, double.cens.normal(r1 = 2, r2 = 3, isd = 6),
           data = strontium90, trace = TRUE)
coef(fit, matrix = TRUE)
Coef(fit)
```

double.expbinomial *Double Exponential Binomial Distribution Family Function*

Description

Fits a double exponential binomial distribution by maximum likelihood estimation. The two parameters here are the mean and dispersion parameter.

Usage

```
double.expbinomial(lmean = "logitlink", ldispersion = "logitlink",
                  idispersion = 0.25, zero = "dispersion")
```

Arguments

lmean, ldispersion	Link functions applied to the two parameters, called μ and θ respectively below. See Links for more choices. The defaults cause the parameters to be restricted to $(0, 1)$.
idispersion	Initial value for the dispersion parameter. If given, it must be in range, and is recycled to the necessary length. Use this argument if convergence failure occurs.
zero	A vector specifying which linear/additive predictor is to be modelled as intercept-only. If assigned, the single value can be either 1 or 2. The default is to have a single dispersion parameter value. To model both parameters as functions of the covariates assign zero = NULL. See CommonVGAMffArguments for more details.

Details

This distribution provides a way for handling overdispersion in a binary response. The double exponential binomial distribution belongs the family of double exponential distributions proposed by Efron (1986). Below, equation numbers refer to that original article. Briefly, the idea is that an ordinary one-parameter exponential family allows the addition of a second parameter θ which varies the dispersion of the family without changing the mean. The extended family behaves like the original family with sample size changed from n to $n\theta$. The extended family is an exponential family in μ when n and θ are fixed, and an exponential family in θ when n and μ are fixed. Having $0 < \theta < 1$ corresponds to overdispersion with respect to the binomial distribution. See Efron (1986) for full details.

This **VGAM** family function implements an *approximation* (2.10) to the exact density (2.4). It replaces the normalizing constant by unity since the true value nearly equals 1. The default model fitted is $\eta_1 = \text{logit}(\mu)$ and $\eta_2 = \text{logit}(\theta)$. This restricts both parameters to lie between 0 and 1, although the dispersion parameter can be modelled over a larger parameter space by assigning the arguments ldispersion and edispersion.

Approximately, the mean (of Y) is μ . The *effective sample size* is the dispersion parameter multiplied by the original sample size, i.e., $n\theta$. This family function uses Fisher scoring, and the two estimates are asymptotically independent because the expected information matrix is diagonal.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#).

Warning

Numerical difficulties can occur; if so, try using `idispersion`.

Note

This function processes the input in the same way as [binomialff](#), however multiple responses are not allowed (`binomialff(multiple.responses = FALSE)`).

Author(s)

T. W. Yee

References

Efron, B. (1986). Double exponential families and their use in generalized linear regression. *Journal of the American Statistical Association*, **81**, 709–721.

See Also

[binomialff](#), [toxop](#), [CommonVGAMffArguments](#).

Examples

```
# This example mimics the example in Efron (1986).
# The results here differ slightly.

# Scale the variables
toxop <- transform(toxop,
                  phat = positive / ssize,
                  srainfall = scale(rainfall), # (6.1)
                  sN = scale(ssize))         # (6.2)

# A fit similar (should be identical) to Sec.6 of Efron (1986).
# But does not use poly(), and M = 1.25 here, as in (5.3)
cmlist <- list("Intercept" = diag(2),
              "I(srainfall)" = rbind(1, 0),
              "I(srainfall^2)" = rbind(1, 0),
              "I(srainfall^3)" = rbind(1, 0),
              "I(sN)" = rbind(0, 1),
              "I(sN^2)" = rbind(0, 1))

fit <-
  vglm(cbind(phat, 1 - phat) * ssize ~
```



```

I(srainfall) + I(srainfall^2) + I(srainfall^3) +
I(sN) + I(sN^2),
double.expbinoimial(ldisp = extlogitlink(min = 0, max = 1.25),
                    idisp = 0.2, zero = NULL),
                    toxop, trace = TRUE, constraints = cmlist)

# Now look at the results
coef(fit, matrix = TRUE)
head(fitted(fit))
summary(fit)
vcov(fit)
sqrt(diag(vcov(fit))) # Standard errors

# Effective sample size (not quite the last column of Table 1)
head(predict(fit))
Dispersion <- extlogitlink(predict(fit)[,2], min = 0, max = 1.25,
                           inverse = TRUE)
c(round(weights(fit, type = "prior") * Dispersion, digits = 1))

# Ordinary logistic regression (gives same results as (6.5))
ofit <- vglm(cbind(phat, 1 - phat) * ssize ~
            I(srainfall) + I(srainfall^2) + I(srainfall^3),
            binomialff, toxop, trace = TRUE)

# Same as fit but it uses poly(), and can be plotted (cf. Fig.1)
cmlist2 <- list("(Intercept)" = diag(2),
               "poly(srainfall, degree = 3)" = rbind(1, 0),
               "poly(sN, degree = 2)" = rbind(0, 1))
fit2 <-
  vglm(cbind(phat, 1 - phat) * ssize ~
        poly(srainfall, degree = 3) + poly(sN, degree = 2),
        double.expbinoimial(ldisp = extlogitlink(min = 0, max = 1.25),
                            idisp = 0.2, zero = NULL),
        toxop, trace = TRUE, constraints = cmlist2)
## Not run: par(mfrow = c(1, 2)) # Cf. Fig.1
plot(as(fit2, "vgam"), se = TRUE, lcol = "blue", scol = "orange")

# Cf. Figure 1(a)
par(mfrow = c(1,2))
ooo <- with(toxop, sort.list(rainfall))
with(toxop, plot(rainfall[ooo], fitted(fit2)[ooo], type = "l",
                col = "blue", las = 1, ylim = c(0.3, 0.65)))
with(toxop, points(rainfall[ooo], fitted(ofit)[ooo],
                  col = "orange", type = "b", pch = 19))

# Cf. Figure 1(b)
ooo <- with(toxop, sort.list(ssize))
with(toxop, plot(ssize[ooo], Dispersion[ooo], type = "l",
                col = "blue", las = 1, xlim = c(0, 100)))
## End(Not run)

```

ducklings

Relative Frequencies of Serum Proteins in White Pekin Ducklings

Description

Relative frequencies of serum proteins in white Pekin ducklings as determined by electrophoresis.

Usage

```
data(ducklings)
```

Format

The format is: chr "ducklings"

Details

Columns p1, p2, p3 stand for pre-albumin, albumin, globulins respectively. These were collected from 3-week old white Pekin ducklings. Let Y_1 be proportional to the total milligrams of pre-albumin in the blood serum of a duckling. Similarly, let Y_2 and Y_3 be directly proportional to the same factor as Y_1 to the total milligrams respectively of albumin and globulins in its blood serum. The proportion of pre-albumin is given by $Y_1/(Y_1 + Y_2 + Y_3)$, and similarly for the others.

Source

Mosimann, J. E. (1962) On the compound multinomial distribution, the multivariate β -distribution, and correlations among proportions, *Biometrika*, **49**, 65–82.

See Also

[dirichlet](#).

Examples

```
print(ducklings)
```

eCDF

Empirical Cumulative Distribution Function

Description

Returns the desired quantiles of quantile regression object such as an `extlogF1()` or `lms.bcn()` VGLM object

Usage

```
eCDF.vglm(object, all = FALSE, ...)
```

Arguments

object an object such as a `vglm` object with family function `extlogF1` or `lms.bcn`.
 all Logical. Return all other information? If true, the empirical CDF is returned.
 ... additional optional arguments. Currently unused.

Details

This function was specifically written for a `vglm` object with family function `extlogF1` or `lms.bcn`. It returns the proportion of data lying below each of the fitted quantiles, and optionally the desired quantiles (arguments `tau` or `percentiles / 100` in the family function). The output is coerced to be comparable between family functions by calling the columns by the same names.

Value

A vector with each value lying in (0, 1). If `all = TRUE` then a 2-column matrix with the second column being the tau values or equivalent.

See Also

`extlogF1`, `lms.bcn`, `vglm`.

Examples

```
fit1 <- vglm(BMI ~ ns(age, 4), extlogF1, data = bmi.nz) # trace = TRUE
eCDF(fit1)
eCDF(fit1, all = TRUE)
```

 enzyme

Enzyme Data

Description

Enzyme velocity and substrate concentration.

Usage

```
data(enzyme)
```

Format

A data frame with 12 observations on the following 2 variables.

conc a numeric explanatory vector; substrate concentration

velocity a numeric response vector; enzyme velocity

Details

Sorry, more details need to be included later.

Source

Sorry, more details need to be included later.

References

Watts, D. G. (1981). An introduction to nonlinear least squares. In: L. Endrenyi (Ed.), *Kinetic Data Analysis: Design and Analysis of Enzyme and Pharmacokinetic Experiments*, pp.1–24. New York: Plenum Press.

See Also

[micmen](#).

Examples

```
## Not run:
fit <- vglm(velocity ~ 1, micmen, data = enzyme, trace = TRUE,
           form2 = ~ conc - 1, crit = "crit")
summary(fit)

## End(Not run)
```

erf

Error Function, and variants

Description

Computes the error function, or its inverse, based on the normal distribution. Also computes the complement of the error function, or its inverse,

Usage

```
erf(x, inverse = FALSE)
erfc(x, inverse = FALSE)
```

Arguments

x	Numeric.
inverse	Logical. Of length 1.

Details

$Erf(x)$ is defined as

$$Erf(x) = \frac{2}{\sqrt{\pi}} \int_0^x \exp(-t^2) dt$$

so that it is closely related to [pnorm](#). The inverse function is defined for x in $(-1, 1)$.

Value

Returns the value of the function evaluated at x .

Note

Some authors omit the term $2/\sqrt{\pi}$ from the definition of $Erf(x)$. Although defined for complex arguments, this function only works for real arguments.

The *complementary error function* $erfc(x)$ is defined as $1 - erf(x)$, and is implemented by `erfc`. Its inverse function is defined for x in $(0, 2)$.

Author(s)

T. W. Yee

References

Abramowitz, M. and Stegun, I. A. (1972). *Handbook of Mathematical Functions with Formulas, Graphs, and Mathematical Tables*, New York: Dover Publications Inc.

See Also

[pnorm](#).

Examples

```
## Not run:
curve(erf, -3, 3, col = "orange", ylab = "", las = 1)
curve(pnorm, -3, 3, add = TRUE, col = "blue", lty = "dotted", lwd = 2)
abline(v = 0, h = 0, lty = "dashed")
legend("topleft", c("erf(x)", "pnorm(x)"), col = c("orange", "blue"),
      lty = c("solid", "dotted"), lwd = 1:2)
## End(Not run)
```

erlang

Erlang Distribution

Description

Estimates the scale parameter of the Erlang distribution by maximum likelihood estimation.

Usage

```
erlang(shape.arg, lscale = "loglink", imethod = 1, zero = NULL)
```

Arguments

shape.arg	The shape parameters. The user must specify a positive integer, or integers for multiple responses. They are recycled by <code>.row = TRUE</code> according to matrix .
lscale	Link function applied to the (positive) <i>scale</i> parameter. See Links for more choices.
imethod, zero	See CommonVGAMffArguments for more details.

Details

The Erlang distribution is a special case of the gamma distribution with *shape* that is a positive integer. If `shape.arg = 1` then it simplifies to the exponential distribution. As illustrated in the example below, the Erlang distribution is the distribution of the sum of `shape.arg` independent and identically distributed exponential random variates.

The probability density function of the Erlang distribution is given by

$$f(y) = \exp(-y/scale)y^{shape-1}scale^{-shape}/\Gamma(shape)$$

for known positive integer *shape*, unknown $scale > 0$ and $y > 0$. Here, $\Gamma(shape)$ is the gamma function, as in [gamma](#). The mean of Y is $\mu = shape \times scale$ and its variance is $shape \times scale^2$. The linear/additive predictor, by default, is $\eta = \log(scale)$.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Note

Multiple responses are permitted. The rate parameter found in [gammaR](#) is $1/scale$ here—see also [rgamma](#).

Author(s)

T. W. Yee

References

Most standard texts on statistical distributions describe this distribution, e.g.,

Forbes, C., Evans, M., Hastings, N. and Peacock, B. (2011). *Statistical Distributions*, Hoboken, NJ, USA: John Wiley and Sons, Fourth edition.

See Also

[gammaR](#), [exponential](#), [simulate.vlm](#).

Examples

```

rate <- exp(2); myshape <- 3
edata <- data.frame(y = rep(0, nn <- 1000))
for (ii in 1:myshape)
  edata <- transform(edata, y = y + rexp(nn, rate = rate))
fit <- vglm(y ~ 1, erlang(shape = myshape), edata, trace = TRUE)
coef(fit, matrix = TRUE)
Coef(fit) # Answer = 1/rate
1/rate
summary(fit)

```

Expectiles-Exponential

Expectiles of the Exponential Distribution

Description

Density function, distribution function, and expectile function and random generation for the distribution associated with the expectiles of an exponential distribution.

Usage

```

deexp(x, rate = 1, log = FALSE)
peexp(q, rate = 1, lower.tail = TRUE, log.p = FALSE)
qeexp(p, rate = 1, Maxit.nr = 10, Tol.nr = 1.0e-6,
      lower.tail = TRUE, log.p = FALSE)
reexp(n, rate = 1)

```

Arguments

x, p, q	See deunif .
n, rate, log	See rexp .
lower.tail, log.p	Same meaning as in pexp or qeexp .
Maxit.nr, Tol.nr	See deunif .

Details

General details are given in [deunif](#) including a note regarding the terminology used. Here, `exp` corresponds to the distribution of interest, F , and `eexp` corresponds to G . The addition of “e” is for the ‘other’ distribution associated with the parent distribution. Thus `deexp` is for g , `peexp` is for G , `qeexp` is for the inverse of G , `reexp` generates random variates from g .

For `qeexp` the Newton-Raphson algorithm is used to solve for y satisfying $p = G(y)$. Numerical problems may occur when values of p are very close to 0 or 1.

Value

`deexp(x)` gives the density function $g(x)$. `peexp(q)` gives the distribution function $G(q)$. `qeexp(p)` gives the expectile function: the value y such that $G(y) = p$. `reexp(n)` gives n random variates from G .

Author(s)

T. W. Yee and Kai Huang

See Also

[deunif](#), [denorm](#), [dexp](#).

Examples

```
my.p <- 0.25; y <- rexp(nn <- 1000)
(myexp <- qeexp(my.p))
sum(myexp - y[y <= myexp]) / sum(abs(myexp - y)) # Should be my.p

## Not run: par(mfrow = c(2,1))
yy <- seq(-0, 4, len = nn)
plot(yy, deexp(yy), col = "blue", ylim = 0:1, xlab = "y", ylab = "g(y)",
      type = "l", main = "g(y) for Exp(1); dotted green is f(y) = dexp(y)")
lines(yy, dexp(yy), col = "green", lty = "dotted", lwd = 2) # 'original'

plot(yy, peexp(yy), type = "l", col = "blue", ylim = 0:1,
      xlab = "y", ylab = "G(y)", main = "G(y) for Exp(1)")
abline(v = 1, h = 0.5, col = "red", lty = "dashed")
lines(yy, pexp(yy), col = "green", lty = "dotted", lwd = 2)
## End(Not run)
```

Expectiles-Normal

Expectiles of the Normal Distribution

Description

Density function, distribution function, and expectile function and random generation for the distribution associated with the expectiles of a normal distribution.

Usage

```
denorm(x, mean = 0, sd = 1, log = FALSE)
penorm(q, mean = 0, sd = 1, lower.tail = TRUE, log.p = FALSE)
qenorm(p, mean = 0, sd = 1, Maxit.nr = 10, Tol.nr = 1.0e-6,
       lower.tail = TRUE, log.p = FALSE)
renorm(n, mean = 0, sd = 1)
```


Arguments

`x, p, q` See [deunif](#).
`n, mean, sd, log` See [rnorm](#).
`lower.tail, log.p`
 Same meaning as in [pnorm](#) or [qnorm](#).
`Maxit.nr, Tol.nr`
 See [deunif](#).

Details

General details are given in [deunif](#) including a note regarding the terminology used. Here, `norm` corresponds to the distribution of interest, F , and `enorm` corresponds to G . The addition of “e” is for the ‘other’ distribution associated with the parent distribution. Thus `denorm` is for g , `penorm` is for G , `qenorm` is for the inverse of G , `renorm` generates random variates from g .

For `qenorm` the Newton-Raphson algorithm is used to solve for y satisfying $p = G(y)$. Numerical problems may occur when values of p are very close to 0 or 1.

Value

`denorm(x)` gives the density function $g(x)$. `penorm(q)` gives the distribution function $G(q)$. `qenorm(p)` gives the expectile function: the value y such that $G(y) = p$. `renorm(n)` gives n random variates from G .

Author(s)

T. W. Yee and Kai Huang

See Also

[deunif](#), [deexp](#), [dnorm](#), [amlnormal](#), [lms.bcn](#).

Examples

```
my.p <- 0.25; y <- rnorm(nn <- 1000)
(myexp <- qenorm(my.p))
sum(myexp - y[y <= myexp]) / sum(abs(myexp - y)) # Should be my.p

# Non-standard normal
my.mean <- 1; my.sd <- 2
yy <- rnorm(nn, my.mean, my.sd)
(myexp <- qenorm(my.p, my.mean, my.sd))
sum(myexp - yy[yy <= myexp]) / sum(abs(myexp - yy)) # Should be my.p
penorm(-Inf, my.mean, my.sd) # Should be 0
penorm( Inf, my.mean, my.sd) # Should be 1
penorm(mean(yy), my.mean, my.sd) # Should be 0.5
abs(qenorm(0.5, my.mean, my.sd) - mean(yy)) # Should be 0
abs(penorm(myexp, my.mean, my.sd) - my.p) # Should be 0
integrate(f = denorm, lower = -Inf, upper = Inf,
          my.mean, my.sd) # Should be 1
```

```
## Not run:
par(mfrow = c(2, 1))
yy <- seq(-3, 3, len = nn)
plot(yy, denorm(yy), type = "l", col="blue", xlab = "y", ylab = "g(y)",
      main = "g(y) for N(0,1); dotted green is f(y) = dnorm(y)")
lines(yy, dnorm(yy), col = "green", lty = "dotted", lwd = 2) # 'original'

plot(yy, penorm(yy), type = "l", col = "blue", ylim = 0:1,
      xlab = "y", ylab = "G(y)", main = "G(y) for N(0,1)")
abline(v = 0, h = 0.5, col = "red", lty = "dashed")
lines(yy, pnorm(yy), col = "green", lty = "dotted", lwd = 2)
## End(Not run)
```

Expectiles-sc.t2

*Expectiles/Quantiles of the Scaled Student t Distribution with 2 Df***Description**

Density function, distribution function, and quantile/expectile function and random generation for the scaled Student t distribution with 2 degrees of freedom.

Usage

```
dsc.t2(x, location = 0, scale = 1, log = FALSE)
psc.t2(q, location = 0, scale = 1, lower.tail = TRUE, log.p = FALSE)
qsc.t2(p, location = 0, scale = 1, lower.tail = TRUE, log.p = FALSE)
rsc.t2(n, location = 0, scale = 1)
```

Arguments

`x, q` Vector of expectiles/quantiles. See the terminology note below.
`p` Vector of probabilities. These should lie in $(0, 1)$.
`n, log` See [runif](#).
`location, scale` Location and scale parameters. The latter should have positive values. Values of these vectors are recycled.
`lower.tail, log.p` Same meaning as in [pt](#) or [qt](#).

Details

A Student-t distribution with 2 degrees of freedom and a scale parameter of $\sqrt{2}$ is equivalent to the standard form of this distribution (called Koenker's distribution below). Further details about this distribution are given in [sc.studentt2](#).

Value

`dsc.t2(x)` gives the density function. `psc.t2(q)` gives the distribution function. `qsc.t2(p)` gives the expectile and quantile function. `rsc.t2(n)` gives n random variates.

Author(s)

T. W. Yee and Kai Huang

See Also[dt](#), [sc.studentt2](#).**Examples**

```

my.p <- 0.25; y <- rsc.t2(nn <- 5000)
(myexp <- qsc.t2(my.p))
sum(myexp - y[y <= myexp]) / sum(abs(myexp - y)) # Should be my.p
# Equivalently:
I1 <- mean(y <= myexp) * mean(myexp - y[y <= myexp])
I2 <- mean(y > myexp) * mean(-myexp + y[y > myexp])
I1 / (I1 + I2) # Should be my.p
# Or:
I1 <- sum(myexp - y[y <= myexp])
I2 <- sum(-myexp + y[y > myexp])

# Non-standard Koenker distribution
myloc <- 1; myscale <- 2
yy <- rsc.t2(nn, myloc, myscale)
(myexp <- qsc.t2(my.p, myloc, myscale))
sum(myexp - yy[y <= myexp]) / sum(abs(myexp - yy)) # Should be my.p
psc.t2(mean(yy), myloc, myscale) # Should be 0.5
abs(qsc.t2(0.5, myloc, myscale) - mean(yy)) # Should be 0
abs(psc.t2(myexp, myloc, myscale) - my.p) # Should be 0
integrate(f = dsc.t2, lower = -Inf, upper = Inf,
          locat = myloc, scale = myscale) # Should be 1

y <- seq(-7, 7, len = 201)
max(abs(dsc.t2(y) - dt(y / sqrt(2), df = 2) / sqrt(2))) # Should be 0
## Not run: plot(y, dsc.t2(y), type = "l", col = "blue", las = 1,
               ylim = c(0, 0.4), main = "Blue = Koenker; orange = N(0, 1)")
lines(y, dnorm(y), type = "l", col = "orange")
abline(h = 0, v = 0, lty = 2)
## End(Not run)

```

Description

Density function, distribution function, and expectile function and random generation for the distribution associated with the expectiles of a uniform distribution.

Usage

```
deunif(x, min = 0, max = 1, log = FALSE)
peunif(q, min = 0, max = 1, lower.tail = TRUE, log.p = FALSE)
qeunif(p, min = 0, max = 1, Maxit.nr = 10, Tol.nr = 1.0e-6,
       lower.tail = TRUE, log.p = FALSE)
reunif(n, min = 0, max = 1)
```

Arguments

<code>x, q</code>	Vector of expectiles. See the terminology note below.
<code>p</code>	Vector of probabilities. These should lie in (0, 1).
<code>n, min, max, log</code>	See <code>runif</code> .
<code>lower.tail, log.p</code>	Same meaning as in <code>punif</code> or <code>qunif</code> .
<code>Maxit.nr</code>	Numeric. Maximum number of Newton-Raphson iterations allowed. A warning is issued if convergence is not obtained for all <code>p</code> values.
<code>Tol.nr</code>	Numeric. Small positive value specifying the tolerance or precision to which the expectiles are computed.

Details

Jones (1994) elucidated on the property that the expectiles of a random variable X with distribution function $F(x)$ correspond to the quantiles of a distribution $G(x)$ where G is related by an explicit formula to F . In particular, let y be the p -expectile of F . Then y is the p -quantile of G where

$$p = G(y) = (P(y) - yF(y)) / (2[P(y) - yF(y)] + y - \mu),$$

and μ is the mean of X . The derivative of G is

$$g(y) = (\mu F(y) - P(y)) / (2[P(y) - yF(y)] + y - \mu)^2.$$

Here, $P(y)$ is the partial moment $\int_{-\infty}^y xf(x) dx$ and $0 < p < 1$. The 0.5-expectile is the mean μ and the 0.5-quantile is the median.

A note about the terminology used here. Recall in the S language there are the `dpqr`-type functions associated with a distribution, e.g., `dunif`, `punif`, `qunif`, `runif`, for the uniform distribution. Here, `unif` corresponds to F and `eunif` corresponds to G . The addition of “e” (for *expectile*) is for the ‘other’ distribution associated with the parent distribution. Thus `deunif` is for g , `peunif` is for G , `qeunif` is for the inverse of G , `reunif` generates random variates from g .

For `qeunif` the Newton-Raphson algorithm is used to solve for y satisfying $p = G(y)$. Numerical problems may occur when values of `p` are very close to 0 or 1.

Value

`deunif(x)` gives the density function $g(x)$. `peunif(q)` gives the distribution function $G(q)$. `qeunif(p)` gives the expectile function: the expectile y such that $G(y) = p$. `reunif(n)` gives n random variates from G .

Author(s)

T. W. Yee and Kai Huang

References

Jones, M. C. (1994). Expectiles and M-quantiles are quantiles. *Statistics and Probability Letters*, **20**, 149–153.

See Also

[deexp](#), [denorm](#), [dunif](#), [dsc.t2](#).

Examples

```
my.p <- 0.25; y <- runif(nn <- 1000)
(myexp <- qeunif(my.p))
sum(myexp - y[y <= myexp]) / sum(abs(myexp - y)) # Should be my.p
# Equivalently:
I1 <- mean(y <= myexp) * mean(myexp - y[y <= myexp])
I2 <- mean(y > myexp) * mean(-myexp + y[y > myexp])
I1 / (I1 + I2) # Should be my.p
# Or:
I1 <- sum(myexp - y[y <= myexp])
I2 <- sum(-myexp + y[y > myexp])

# Non-standard uniform
mymin <- 1; mymax <- 8
yy <- runif(nn, mymin, mymax)
(myexp <- qeunif(my.p, mymin, mymax))
sum(myexp - yy[yy <= myexp]) / sum(abs(myexp - yy)) # Should be my.p
peunif(mymin, mymin, mymax) # Should be 0
peunif(mymax, mymin, mymax) # Should be 1
peunif(mean(yy), mymin, mymax) # Should be 0.5
abs(qeunif(0.5, mymin, mymax) - mean(yy)) # Should be 0
abs(qeunif(0.5, mymin, mymax) - (mymin+mymax)/2) # Should be 0
abs(peunif(myexp, mymin, mymax) - my.p) # Should be 0
integrate(f = deunif, lower = mymin - 3, upper = mymax + 3,
          min = mymin, max = mymax) # Should be 1

## Not run:
par(mfrow = c(2,1))
yy <- seq(0.0, 1.0, len = nn)
plot(yy, deunif(yy), type = "l", col = "blue", ylim = c(0, 2),
      xlab = "y", ylab = "g(y)", main = "g(y) for Uniform(0,1)")
lines(yy, dunif(yy), col = "green", lty = "dotted", lwd = 2) # 'original'

plot(yy, peunif(yy), type = "l", col = "blue", ylim = 0:1,
      xlab = "y", ylab = "G(y)", main = "G(y) for Uniform(0,1)")
abline(a = 0.0, b = 1.0, col = "green", lty = "dotted", lwd = 2)
abline(v = 0.5, h = 0.5, col = "red", lty = "dashed")
## End(Not run)
```

expexpff

*Exponentiated Exponential Distribution***Description**

Estimates the two parameters of the exponentiated exponential distribution by maximum likelihood estimation.

Usage

```
expexpff(lrate = "loglink", lshape = "loglink",
         irate = NULL, ishape = 1.1, tolerance = 1.0e-6, zero = NULL)
```

Arguments

lshape, lrate	Parameter link functions for the α and λ parameters. See Links for more choices. The defaults ensure both parameters are positive.
ishape	Initial value for the α parameter. If convergence fails try setting a different value for this argument.
irate	Initial value for the λ parameter. By default, an initial value is chosen internally using ishape.
tolerance	Numeric. Small positive value for testing whether values are close enough to 1 and 2.
zero	An integer-valued vector specifying which linear/additive predictors are modelled as intercepts only. The default is none of them. If used, choose one value from the set {1,2}. See CommonVGAMffArguments for more information.

Details

The exponentiated exponential distribution is an alternative to the Weibull and the gamma distributions. The formula for the density is

$$f(y; \lambda, \alpha) = \alpha \lambda (1 - \exp(-\lambda y))^{\alpha-1} \exp(-\lambda y)$$

where $y > 0$, $\lambda > 0$ and $\alpha > 0$. The mean of Y is $(\psi(\alpha + 1) - \psi(1))/\lambda$ (returned as the fitted values) where ψ is the digamma function. The variance of Y is $(\psi'(1) - \psi'(\alpha + 1))/\lambda^2$ where ψ' is the trigamma function.

This distribution has been called the two-parameter generalized exponential distribution by Gupta and Kundu (2006). A special case of the exponentiated exponential distribution: $\alpha = 1$ is the exponential distribution.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Warning

Practical experience shows that reasonably good initial values really helps. In particular, try setting different values for the `ishape` argument if numerical problems are encountered or failure to convergence occurs. Even if convergence occurs try perturbing the initial value to make sure the global solution is obtained and not a local solution. The algorithm may fail if the estimate of the shape parameter is too close to unity.

Note

Fisher scoring is used, however, convergence is usually very slow. This is a good sign that there is a bug, but I have yet to check that the expected information is correct. Also, I have yet to implement Type-I right censored data using the results of Gupta and Kundu (2006).

Another algorithm for fitting this model is implemented in [expexpff1](#).

Author(s)

T. W. Yee

References

Gupta, R. D. and Kundu, D. (2001). Exponentiated exponential family: an alternative to gamma and Weibull distributions, *Biometrical Journal*, **43**, 117–130.

Gupta, R. D. and Kundu, D. (2006). On the comparison of Fisher information of the Weibull and GE distributions, *Journal of Statistical Planning and Inference*, **136**, 3130–3144.

See Also

[expexpff1](#), [gammaR](#), [weibullR](#), [CommonVGAMffArguments](#).

Examples

```
## Not run:
# A special case: exponential data
edata <- data.frame(y = rexp(n <- 1000))
fit <- vglm(y ~ 1, fam = expexpff, data = edata, trace = TRUE, maxit = 99)
coef(fit, matrix = TRUE)
Coef(fit)

# Ball bearings data (number of million revolutions before failure)
edata <- data.frame(bbearings = c(17.88, 28.92, 33.00, 41.52, 42.12, 45.60,
48.80, 51.84, 51.96, 54.12, 55.56, 67.80, 68.64, 68.64,
68.88, 84.12, 93.12, 98.64, 105.12, 105.84, 127.92,
128.04, 173.40))
fit <- vglm(bbearings ~ 1, fam = expexpff(irate = 0.05, ish = 5),
           trace = TRUE, maxit = 300, data = edata)
coef(fit, matrix = TRUE)
Coef(fit) # Authors get c(rate=0.0314, shape=5.2589)
logLik(fit) # Authors get -112.9763
```

```
# Failure times of the airconditioning system of an airplane
eedata <- data.frame(acplane = c(23, 261, 87, 7, 120, 14, 62, 47,
225, 71, 246, 21, 42, 20, 5, 12, 120, 11, 3, 14,
71, 11, 14, 11, 16, 90, 1, 16, 52, 95))
fit <- vglm(acplane ~ 1, fam = expexpff(ishape = 0.8, irate = 0.15),
            trace = TRUE, maxit = 99, data = eedata)
coef(fit, matrix = TRUE)
Coef(fit) # Authors get c(rate=0.0145, shape=0.8130)
logLik(fit) # Authors get log-lik -152.264

## End(Not run)
```

expexpff1

Exponentiated Exponential Distribution

Description

Estimates the two parameters of the exponentiated exponential distribution by maximizing a profile (concentrated) likelihood.

Usage

```
expexpff1(lrate = "loglink", irate = NULL, ishape = 1)
```

Arguments

lrate	Parameter link function for the (positive) λ parameter. See Links for more choices.
irate	Initial value for the λ parameter. By default, an initial value is chosen internally using ishape.
ishape	Initial value for the α parameter. If convergence fails try setting a different value for this argument.

Details

See [expexpff](#) for details about the exponentiated exponential distribution. This family function uses a different algorithm for fitting the model. Given λ , the MLE of α can easily be solved in terms of λ . This family function maximizes a profile (concentrated) likelihood with respect to λ . Newton-Raphson is used, which compares with Fisher scoring with [expexpff](#).

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Warning

The standard errors produced by a summary of the model may be wrong.

Note

This family function works only for intercept-only models, i.e., $y \sim 1$ where y is the response.

The estimate of α is attached to the `misc` slot of the object, which is a list and contains the component shape.

As Newton-Raphson is used, the working weights are sometimes negative, and some adjustment is made to these to make them positive.

Like `expexpff`, good initial values are needed. Convergence may be slow.

Author(s)

T. W. Yee

References

Gupta, R. D. and Kundu, D. (2001). Exponentiated exponential family: an alternative to gamma and Weibull distributions, *Biometrical Journal*, **43**, 117–130.

See Also

[expexpff](#), [CommonVGAMffArguments](#).

Examples

```
# Ball bearings data (number of million revolutions before failure)
edata <- data.frame(bbearings = c(17.88, 28.92, 33.00, 41.52, 42.12, 45.60,
48.80, 51.84, 51.96, 54.12, 55.56, 67.80, 68.64, 68.64,
68.88, 84.12, 93.12, 98.64, 105.12, 105.84, 127.92,
128.04, 173.40))
fit <- vglm(bbearings ~ 1, expexpff1(ishape = 4), trace = TRUE,
           maxit = 250, checkwz = FALSE, data = edata)
coef(fit, matrix = TRUE)
Coef(fit) # Authors get c(0.0314, 5.2589) with log-lik -112.9763
logLik(fit)
fit@misc$shape # Estimate of shape
```

```
# Failure times of the airconditioning system of an airplane
eedata <- data.frame(acplane = c(23, 261, 87, 7, 120, 14, 62, 47,
225, 71, 246, 21, 42, 20, 5, 12, 120, 11, 3, 14,
71, 11, 14, 11, 16, 90, 1, 16, 52, 95))
fit <- vglm(acplane ~ 1, expexpff1(ishape = 0.8), trace = TRUE,
           maxit = 50, checkwz = FALSE, data = eedata)
coef(fit, matrix = TRUE)
Coef(fit) # Authors get c(0.0145, 0.8130) with log-lik -152.264
logLik(fit)
fit@misc$shape # Estimate of shape
```

expgeom

The Exponential Geometric Distribution

Description

Density, distribution function, quantile function and random generation for the exponential geometric distribution.

Usage

```
dexpgeom(x, scale = 1, shape, log = FALSE)
pexpgeom(q, scale = 1, shape)
qexpgeom(p, scale = 1, shape)
rexpgeom(n, scale = 1, shape)
```

Arguments

x, q	vector of quantiles.
p	vector of probabilities.
n	number of observations. If <code>length(n) > 1</code> then the length is taken to be the number required.
scale, shape	positive scale and shape parameters.
log	Logical. If <code>log = TRUE</code> then the logarithm of the density is returned.

Details

See [expgeometric](#), the **VGAM** family function for estimating the parameters, for the formula of the probability density function and other details.

Value

`dexpgeom` gives the density, `pexpgeom` gives the distribution function, `qexpgeom` gives the quantile function, and `rexpgeom` generates random deviates.

Note

We define `scale` as the reciprocal of the scale parameter used by Adamidis and Loukas (1998).

Author(s)

J. G. Lauder and T. W. Yee

See Also

[expgeometric](#), [exponential](#), [geometric](#).

Examples

```
## Not run:
shape <- 0.5; scale <- 1; nn <- 501
x <- seq(-0.10, 3.0, len = nn)
plot(x, dexpgeom(x, scale, shape), type = "l", las = 1, ylim = c(0, 2),
      ylab = paste("[dp]expgeom(shape = ", shape, ", scale = ", scale, ")"),
      col = "blue", cex.main = 0.8,
      main = "Blue is density, red is cumulative distribution function",
      sub = "Purple lines are the 10,20,...,90 percentiles")
lines(x, pexpgeom(x, scale, shape), col = "red")
probs <- seq(0.1, 0.9, by = 0.1)
Q <- qexpgeom(probs, scale, shape)
lines(Q, dexpgeom(Q, scale, shape), col = "purple", lty = 3, type = "h")
lines(Q, pexpgeom(Q, scale, shape), col = "purple", lty = 3, type = "h")
abline(h = probs, col = "purple", lty = 3)
max(abs(pexpgeom(Q, scale, shape) - probs)) # Should be 0

## End(Not run)
```

expgeometric

*Exponential Geometric Distribution Family Function***Description**

Estimates the two parameters of the exponential geometric distribution by maximum likelihood estimation.

Usage

```
expgeometric(lscale = "loglink", lshape = "logitlink",
             iscale = NULL,  ishape = NULL,
             tol12 = 1e-05, zero = 1, nsimEIM = 400)
```

Arguments

`lscale, lshape` Link function for the two parameters. See [Links](#) for more choices.
`iscale, ishape` Numeric. Optional initial values for the scale and shape parameters.
`tol12` Numeric. Tolerance for testing whether a parameter has value 1 or 2.
`zero, nsimEIM` See [CommonVGAMffArguments](#).

Details

The exponential geometric distribution has density function

$$f(y; c = \text{scale}, s = \text{shape}) = (1/c)(1 - s)e^{-y/c}(1 - se^{-y/c})^{-2}$$

where $y > 0$, $c > 0$ and $s \in (0, 1)$. The mean, $(c(s - 1)/s) \log(1 - s)$ is returned as the fitted values. Note the median is $c \log(2 - s)$. Simulated Fisher scoring is implemented.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Note

We define scale as the reciprocal of the scale parameter used by Adamidis and Loukas (1998).

Author(s)

J. G. Lauder and T. W. Yee

References

Adamidis, K., Loukas, S. (1998). A lifetime distribution with decreasing failure rate. *Statistics and Probability Letters*, **39**, 35–42.

See Also

[dexpgeom](#), [exponential](#), [geometric](#).

Examples

```
## Not run:
Scale <- exp(2); shape = logitlink(-1, inverse = TRUE);
edata <- data.frame(y = rexpgeom(n = 2000, scale = Scale, shape = shape))
fit <- vglm(y ~ 1, expgeometric, edata, trace = TRUE)
c(with(edata, mean(y)), head(fitted(fit), 1))
coef(fit, matrix = TRUE)
Coef(fit)
summary(fit)

## End(Not run)
```

expint

The Exponential Integral and Variants

Description

Computes the exponential integral $Ei(x)$ for real values, as well as $\exp(-x) \times Ei(x)$ and $E_1(x)$ and their derivatives (up to the 3rd derivative).

Usage

```
expint(x, deriv = 0)
expexpint(x, deriv = 0)
expint.E1(x, deriv = 0)
```

Arguments

x	Numeric. Ideally a vector of positive reals.
deriv	Integer. Either 0, 1, 2 or 3.

Details

The exponential integral $Ei(x)$ function is the integral of $\exp(t)/t$ from 0 to x , for positive real x . The function $E_1(x)$ is the integral of $\exp(-t)/t$ from x to infinity, for positive real x .

Value

Function `expint(x, deriv = n)` returns the n th derivative of $Ei(x)$ (up to the 3rd), function `expexpint(x, deriv = n)` returns the n th derivative of $\exp(-x) \times Ei(x)$ (up to the 3rd), function `expint.E1(x, deriv = n)` returns the n th derivative of $E_1(x)$ (up to the 3rd).

Warning

These functions have not been tested thoroughly.

Author(s)

T. W. Yee has simply written a small wrapper function to call the NETLIB FORTRAN code. Xiangjie Xue modified the functions to calculate derivatives. Higher derivatives can actually be calculated—please let me know if you need it.

References

<https://netlib.org/specfun/ei>.

See Also

[log, exp](#). There is also a package called `expint`.

Examples

```
## Not run:
par(mfrow = c(2, 2))
curve(expint, 0.01, 2, xlim = c(0, 2), ylim = c(-3, 5),
      las = 1, col = "orange")
abline(v = (-3):5, h = (-4):5, lwd = 2, lty = "dotted", col = "gray")
abline(h = 0, v = 0, lty = "dashed", col = "blue")

curve(expexpint, 0.01, 2, xlim = c(0, 2), ylim = c(-3, 2),
      las = 1, col = "orange")
abline(v = (-3):2, h = (-4):5, lwd = 2, lty = "dotted", col = "gray")
abline(h = 0, v = 0, lty = "dashed", col = "blue")

curve(expint.E1, 0.01, 2, xlim = c(0, 2), ylim = c(0, 5),
      las = 1, col = "orange")
abline(v = (-3):2, h = (-4):5, lwd = 2, lty = "dotted", col = "gray")
```

```
abline(h = 0, v = 0, lty = "dashed", col = "blue")

## End(Not run)
```

 explink

Exponential Link Function

Description

Computes the exponential transformation, including its inverse and the first two derivatives.

Usage

```
explink(theta, bvalue = NULL, inverse = FALSE, deriv = 0,
        short = TRUE, tag = FALSE)
```

Arguments

theta Numeric or character. See below for further details.
 bvalue See [clogloglink](#).
 inverse, deriv, short, tag
 Details at [Links](#).

Details

The exponential link function is potentially suitable for parameters that are positive. Numerical values of theta close to negative or positive infinity may result in 0, Inf, -Inf, NA or NaN.

Value

For explink with deriv = 0, the exponential of theta, i.e., $\exp(\text{theta})$ when inverse = FALSE. And if inverse = TRUE then $\log(\text{theta})$; if theta is not positive then it will return NaN.

For deriv = 1, then the function returns $d \text{ eta} / d \text{ theta}$ as a function of theta if inverse = FALSE, else if inverse = TRUE then it returns the reciprocal.

Here, all logarithms are natural logarithms, i.e., to base e .

Note

This function has particular use for computing quasi-variances when used with [rcim](#) and [uinormal](#).

Numerical instability may occur when theta is close to negative or positive infinity. One way of overcoming this (one day) is to use bvalue.

Author(s)

Thomas W. Yee

See Also

[Links](#), [loglink](#), [rcim](#), [Qvar](#), [uninormal](#).

Examples

```
theta <- rnorm(30)
explink(theta)
max(abs(explink(explink(theta), inverse = TRUE) - theta)) # 0?
```

 explog

The Exponential Logarithmic Distribution

Description

Density, distribution function, quantile function and random generation for the exponential logarithmic distribution.

Usage

```
dexplog(x, scale = 1, shape, log = FALSE)
pexplog(q, scale = 1, shape)
qexplog(p, scale = 1, shape)
rexplog(n, scale = 1, shape)
```

Arguments

x, q	vector of quantiles.
p	vector of probabilities.
n	number of observations. If <code>length(n) > 1</code> then the length is taken to be the number required.
scale, shape	positive scale and shape parameters.
log	Logical. If <code>log = TRUE</code> then the logarithm of the density is returned.

Details

See [explogff](#), the **VGAM** family function for estimating the parameters, for the formula of the probability density function and other details.

Value

`dexplog` gives the density, `pexplog` gives the distribution function, `qexplog` gives the quantile function, and `rexplog` generates random deviates.

Note

We define `scale` as the reciprocal of the scale parameter used by Tahmasabi and Rezaei (2008).

Author(s)

J. G. Lauder and T. W. Yee

See Also

[explogff](#), [exponential](#).

Examples

```
## Not run:
shape <- 0.5; scale <- 2; nn <- 501
x <- seq(-0.50, 6.0, len = nn)
plot(x, dexplog(x, scale, shape), type = "l", las = 1, ylim = c(0, 1.1),
      ylab = paste("[dp]explog(shape = ", shape, ", scale = ", scale, ")"),
      col = "blue", cex.main = 0.8,
      main = "Blue is density, orange is cumulative distribution function",
      sub = "Purple lines are the 10,20,...,90 percentiles")
lines(x, pexplog(x, scale, shape), col = "orange")
probs <- seq(0.1, 0.9, by = 0.1)
Q <- qexplog(probs, scale, shape = shape)
lines(Q, dexplog(Q, scale, shape = shape), col = "purple", lty = 3, type = "h")
lines(Q, pexplog(Q, scale, shape = shape), col = "purple", lty = 3, type = "h")
abline(h = probs, col = "purple", lty = 3)
max(abs(pexplog(Q, scale, shape = shape) - probs)) # Should be 0

## End(Not run)
```

explogff

Exponential Logarithmic Distribution Family Function

Description

Estimates the two parameters of the exponential logarithmic distribution by maximum likelihood estimation.

Usage

```
explogff(lscale = "loglink", lshape = "logitlink",
         iscale = NULL,  ishape = NULL,
         tol12 = 1e-05, zero = 1, nsimEIM = 400)
```

Arguments

lscale, lshape See [CommonVGAMffArguments](#) for information.
 tol12 Numeric. Tolerance for testing whether a parameter has value 1 or 2.
 iscale, ishape, zero, nsimEIM See [CommonVGAMffArguments](#).

Details

The exponential logarithmic distribution has density function

$$f(y; c, s) = (1/(-\log p))(((1/c)(1-s)e^{-y/c})/(1 - (1-s)e^{-y/c}))$$

where $y > 0$, scale parameter $c > 0$, and shape parameter $s \in (0, 1)$. The mean, $(-polylog(2, 1 - p)c)/\log(s)$ is *not* returned as the fitted values. Note the median is $c \log(1 + \sqrt{s})$ and it is *currently* returned as the fitted values. Simulated Fisher scoring is implemented.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Note

We define scale as the reciprocal of the rate parameter used by Tahmasabi and Sadegh (2008).

Yet to do: find a `polylog()` function.

Author(s)

J. G. Lauder and T. W. Yee

References

Tahmasabi, R., Sadegh, R. (2008). A two-parameter lifetime distribution with decreasing failure rate. *Computational Statistics and Data Analysis*, **52**, 3889–3901.

See Also

[dexplot](#), [exponential](#),

Examples

```
## Not run: Scale <- exp(2); shape <- logitlink(-1, inverse = TRUE)
edata <- data.frame(y = rexplog(n = 2000, scale = Scale, shape = shape))
fit <- vglm(y ~ 1, explogff, data = edata, trace = TRUE)
c(with(edata, median(y)), head(fitted(fit), 1))
coef(fit, matrix = TRUE)
Coef(fit)
summary(fit)

## End(Not run)
```

 exponential

Exponential Distribution

Description

Maximum likelihood estimation for the exponential distribution.

Usage

```
exponential(link = "loglink", location = 0, expected = TRUE,
            type.fitted = c("mean", "percentiles", "Qlink"),
            percentiles = 50,
            ishrinkage = 0.95, parallel = FALSE, zero = NULL)
```

Arguments

link	Parameter link function applied to the positive parameter <i>rate</i> . See Links for more choices.
location	Numeric of length 1, the known location parameter, <i>A</i> , say.
expected	Logical. If TRUE Fisher scoring is used, otherwise Newton-Raphson. The latter is usually faster.
ishrinkage, parallel, zero	See CommonVGAMffArguments for information.
type.fitted, percentiles	See CommonVGAMffArguments for information.

Details

The family function assumes the response Y has density

$$f(y) = \lambda \exp(-\lambda(y - A))$$

for $y > A$, where A is the known location parameter. By default, $A = 0$. Then $E(Y) = A + 1/\lambda$ and $Var(Y) = 1/\lambda^2$.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Note

Suppose $A = 0$. For a fixed time interval, the number of events is Poisson with mean λ if the time between events has a geometric distribution with mean λ^{-1} . The argument *rate* in [exponential](#) is the same as [rexp](#) etc. The argument *lambda* in [rpois](#) is somewhat the same as *rate* here.

Author(s)

T. W. Yee

References

Forbes, C., Evans, M., Hastings, N. and Peacock, B. (2011). *Statistical Distributions*, Hoboken, NJ, USA: John Wiley and Sons, Fourth edition.

See Also

[amlexponential](#), [gpd](#), [laplace](#), [expgeometric](#), [explogff](#), [poissonff](#), [mix2exp](#), [freund61](#), [simulate.vlm](#), [Exponential](#).

Examples

```

edata <- data.frame(x2 = runif(nn <- 100) - 0.5)
edata <- transform(edata, x3 = runif(nn) - 0.5)
edata <- transform(edata, eta = 0.2 - 0.7 * x2 + 1.9 * x3)
edata <- transform(edata, rate = exp(eta))
edata <- transform(edata, y = rexp(nn, rate = rate))
with(edata, stem(y))

fit.slow <- vglm(y ~ x2 + x3, exponential, data = edata, trace = TRUE)
fit.fast <- vglm(y ~ x2 + x3, exponential(exp = FALSE), data = edata,
               trace = TRUE, crit = "coef")
coef(fit.slow, mat = TRUE)
summary(fit.slow)

# Compare results with a GPD. Has a threshold.
threshold <- 0.5
gdata <- data.frame(y1 = threshold + rexp(n = 3000, rate = exp(1.5)))

fit.exp <- vglm(y1 ~ 1, exponential(location = threshold), data = gdata)
coef(fit.exp, matrix = TRUE)
Coef(fit.exp)
logLik(fit.exp)

fit.gpd <- vglm(y1 ~ 1, gpd(threshold = threshold), data = gdata)
coef(fit.gpd, matrix = TRUE)
Coef(fit.gpd)
logLik(fit.gpd)

```

Description

Density, distribution function, quantile function and random generation for the exponential poisson distribution.

Usage

```
dexppois(x, rate = 1, shape, log = FALSE)
pexppois(q, rate = 1, shape, lower.tail = TRUE, log.p = FALSE)
qexppois(p, rate = 1, shape, lower.tail = TRUE, log.p = FALSE)
rexppois(n, rate = 1, shape)
```

Arguments

x, q	vector of quantiles.
p	vector of probabilities.
n	number of observations. If <code>length(n) > 1</code> then the length is taken to be the number required.
shape, rate	positive parameters.
log	Logical. If <code>log = TRUE</code> then the logarithm of the density is returned.
lower.tail, log.p	Same meaning as in pnorm or qnorm .

Details

See [exppoisson](#), the **VGAM** family function for estimating the parameters, for the formula of the probability density function and other details.

Value

`dexppois` gives the density, `pexppois` gives the distribution function, `qexppois` gives the quantile function, and `rexppois` generates random deviates.

Author(s)

Kai Huang and J. G. Lauder

See Also

[exppoisson](#).

Examples

```
## Not run: rate <- 2; shape <- 0.5; nn <- 201
x <- seq(-0.05, 1.05, len = nn)
plot(x, dexppois(x, rate = rate, shape), type = "l", las = 1, ylim = c(0, 3),
     ylab = paste("fexppoisson(rate = ", rate, ", shape = ", shape, ")"),
     col = "blue", cex.main = 0.8,
     main = "Blue is the density, orange the cumulative distribution function",
     sub = "Purple lines are the 10,20,...,90 percentiles")
lines(x, pexppois(x, rate = rate, shape), col = "orange")
probs <- seq(0.1, 0.9, by = 0.1)
Q <- qexppois(probs, rate = rate, shape)
lines(Q, dexppois(Q, rate = rate, shape), col = "purple", lty = 3, type = "h")
lines(Q, pexppois(Q, rate = rate, shape), col = "purple", lty = 3, type = "h")
```

```
abline(h = probs, col = "purple", lty = 3); abline(h = 0, col = "gray50")
max(abs(pexppois(Q, rate = rate, shape) - probs)) # Should be 0

## End(Not run)
```

exppoisson

Exponential Poisson Distribution Family Function

Description

Estimates the two parameters of the exponential Poisson distribution by maximum likelihood estimation.

Usage

```
exppoisson(lrate = "loglink", lshape = "loglink",
           irate = 2, ishape = 1.1, zero = NULL)
```

Arguments

lshape, lrate	Link function for the two positive parameters. See Links for more choices.
ishape, irate	Numeric. Initial values for the shape and rate parameters. Currently this function is not intelligent enough to obtain better initial values.
zero	See CommonVGAMffArguments .

Details

The exponential Poisson distribution has density function

$$f(y; \beta = rate, \lambda = shape) = \frac{\lambda\beta}{1 - e^{-\lambda}} e^{-\lambda - \beta y + \lambda \exp(-\beta y)}$$

where $y > 0$, and the parameters shape, λ , and rate, β , are positive. The distribution implies a population facing discrete hazard rates which are multiples of a base hazard. This **VGAM** family function requires the hypergeo package (to use their genhypergeo function). The median is returned as the fitted value.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Warning

This **VGAM** family function does not work properly!

Author(s)

J. G. Lauder, jamesglaunder@gmail.com

References

Kus, C., (2007). A new lifetime distribution. *Computational Statistics and Data Analysis*, **51**, 4497–4509.

See Also

[dexppois](#), [exponential](#), [poisson](#).

Examples

```
## Not run:
shape <- exp(1); rate <- exp(2)
rdata <- data.frame(y = rexppois(n = 1000, rate = rate, shape = shape))
library("hypergeo") # Required!
fit <- vglm(y ~ 1, expoisson, data = rdata, trace = FALSE, maxit = 1200)
c(with(rdata, median(y)), head(fitted(fit), 1))
coef(fit, matrix = TRUE)
Coef(fit)
summary(fit)

## End(Not run)
```

Extbetabinom

The Beta-Binomial Distribution

Description

Density, distribution function, quantile function and random generation for the extended beta-binomial distribution.

Usage

```
dextbetabinom(x, size, prob, rho = 0,
              log = FALSE, forbycol = TRUE)
pextbetabinom(q, size, prob, rho = 0,
              lower.tail = TRUE, forbycol = TRUE)
qextbetabinom(p, size, prob, rho = 0,
              forbycol = TRUE)
rextbetabinom(n, size, prob, rho = 0)
```

Arguments

x, q	vector of quantiles.
p	vector of probabilities.
size	number of trials.
n	number of observations. Same as runif .
prob	the probability of success μ . Must be in the unit closed interval $[0, 1]$.

rho	the correlation parameter ρ , which may be negative for underdispersion or else be in the interval $[0, 1)$ for overdispersion. The default value of 0 corresponds to the usual binomial distribution with probability prob.
log, lower.tail	Same meaning as <code>runif</code> .
forbycol	Logical. A for loop cycles over either the rows or columns and this argument determines which. The rows are $1:\text{length}(x)$ and the columns are $0:\text{max}(\text{size})$. The best choice is data set dependent.

Details

The *extended* beta-binomial distribution allows for a slightly negative correlation parameter between binary responses within a cluster (e.g., a litter). An exchangeable error structure with correlation ρ is assumed.

Value

`dextbetabinom` gives the density, `pextbetabinom` gives the distribution function, `qextbetabinom` gives the quantile function and `rextbetabinom` generates random deviates.

Warning

Setting `rho = 1` is not recommended as NaN is returned, however the code may be modified in the future to handle this special case.

Note

Currently most of the code is quite slow. Speed improvements are a future project. Use `forbycol` optimally.

See Also

[extbetabinomial](#), [Betabinom](#), [Binomial](#).

Examples

```
set.seed(1); rextbetabinom(10, 100, 0.5)
set.seed(1);          rbinom(10, 100, 0.5) # Same

## Not run: N <- 9; xx <- 0:N; prob <- 0.5; rho <- -0.02
dy <- dextbetabinom(xx, N, prob, rho)
barplot(rbind(dy, dbinom(xx, size = N, prob)),
        beside = TRUE, col = c("blue", "green"), las = 1,
        main = paste0("Beta-binom(size=", N,
        ", prob=", prob, ", rho=", rho, ") (blue) vs\n",
        " Binom(size=", N, ", prob=", prob, ") (green)"),
        names.arg = as.character(xx), cex.main = 0.8)
sum(dy * xx) # Check expected values are equal
sum(dbinom(xx, size = N, prob = prob) * xx)
cumsum(dy) - pextbetabinom(xx, N, prob, rho) # 0?

## End(Not run)
```

extbetabinomial *Extended Beta-binomial Distribution Family Function*

Description

Fits an extended beta-binomial distribution by maximum likelihood estimation. The two parameters here are the mean and correlation coefficient.

Usage

```
extbetabinomial(lmu = "logitlink", lrho = "cloglink",
  zero = "rho", irho = 0, grho = c(0, 0.05, 0.1, 0.2),
  vf1 = FALSE, Form2 = NULL,
  imethod = 1, ishrinkage = 0.95)
```

Arguments

lmu, lrho	Link functions applied to the two parameters. See Links for more choices. The first default ensure the mean remain in $(0, 1)$, while the second allows for a slightly negative correlation parameter: you could say it lies in $(\max(-\mu/(N - \mu - 1), -(1 - \mu)/(N - (1 - \mu) - 1)), 1)$ where μ is the mean (probability) and N is size. See below for details. For lrho, cloglink is a good choice because it handles parameter values from 1 downwards. Other good choices include logofflink(offset = 1) and rhobitlink .
irho, grho	The first is similar to betabinomial and it is a good idea to use this argument because to conduct a grid search based on grho is expensive. The default is effectively a binomial distribution. Set irho = NULL to perform a grid search which is more reliable but slow.
imethod	Similar to betabinomial .
zero	Similar to betabinomial . Also, see CommonVGAMffArguments for more information. Modelling rho with covariates requires large samples.
ishrinkage	See betabinomial and CommonVGAMffArguments for information.
vf1, Form2	See CommonVGAMffArguments . If vf1 = TRUE then Form2 should be a formula specifying the terms for η_2 and all others are used for μ . It is similar to uninormal . If these arguments are used then <code>cbind(0, log(size1 / (size1 - 1)))</code> should be used as an offset, and set zero = NULL too.

Details

The *extended* beta-binomial distribution (EBBD) proposed by Prentice (1986) allows for a slightly negative correlation parameter whereas the ordinary BBD [betabinomial](#) only allows values in $(0, 1)$ so it handles overdispersion only. When negative, the data is underdispersed relative to an ordinary binomial distribution.

Argument rho is used here for the δ used in Prentice (1986) because it is the correlation between the (almost) Bernoulli trials. (They are actually simple binary variates.) We use here N for the number

of trials (e.g., litter size), $T = NY$ is the number of successes, and p (or μ) is the probability of a success (e.g., a malformation). That is, Y is the *proportion* of successes. Like `binomialfff`, the fitted values are the estimated probability of success (i.e., $E[Y]$ and not $E[T]$) and the prior weights N are attached separately on the object in a slot.

The probability function is difficult to write but it involves three series of products. Recall $Y = T/N$ is the real response being modelled, where T is the (total) sum of N correlated (almost) Bernoulli trials.

The default model is $\eta_1 = \text{logit}(\mu)$ and $\eta_2 = \text{clog}(\rho)$ because the first parameter lies between 0 and 1. The second link is `cloglink`. The mean of Y is $p = \mu$ and the variance of Y is $\mu(1 - \mu)(1 + (N - 1)\rho)/N$. Here, the correlation ρ may be slightly negative and is the correlation between the N individuals within a litter. A *litter effect* is typically reflected by a positive value of ρ and corresponds to *overdispersion* with respect to the binomial distribution. Thus an *exchangeable* error structure is assumed between units within a litter for the EBBD.

This family function uses Fisher scoring. Elements of the second-order expected derivatives are computed numerically, which may fail for models very near the boundary of the parameter space. Usually, the computations are expensive for large N because of a for loop, so it may take a long time.

Value

An object of class "vglmfff" (see `vglmfff-class`). The object is used by modelling functions such as `vglm`.

Suppose `fit` is a fitted EBB model. Then `depvar(fit)` are the sample proportions y , `fitted(fit)` returns estimates of $E(Y)$, and `weights(fit, type = "prior")` returns the number of trials N .

Warning

Modelling rho using covariates well requires much data so it is usually best to leave zero alone. It is good to set `trace = TRUE` and play around with `irho` if there are problems achieving convergence. Convergence problems will occur when the estimated rho is close to the lower bound, i.e., the underdispersion is almost too severe for the EBB to cope.

Note

This function is recommended over `betabinomial` and `betabinomialfff`. It processes the input in the same way as `binomialfff`. But it does not handle the case $N \leq 2$ very well because there are two parameters to estimate, not one, for each row of the input. Cases where $N > 2$ can be selected via the `subset` argument of `vglm`.

Author(s)

T. W. Yee

References

Prentice, R. L. (1986). Binary regression using an extended beta-binomial distribution, with discussion of correlation induced by covariate measurement errors. *Journal of the American Statistical Association*, **81**, 321–327.

See Also

[Extbetabinom](#), [betabinomial](#), [betabinomialff](#), [binomialff](#), [dirmultinomial](#), [cloglink](#), [lirat](#).

Examples

```
# Example 1
edata <- data.frame(N = 10, mu = 0.5, rho = 0.1)
edata <- transform(edata,
  y = rextbetabinom(100, N, mu, rho = rho))
fit1 <- vglm(cbind(y, N-y) ~ 1, extbetabinomial, edata, trace = TRUE)
coef(fit1, matrix = TRUE)
Coef(fit1)
head(cbind(depvar(fit1), weights(fit1, type = "prior"))

# Example 2: VFL model
## Not run: N <- size1 <- 10; nn <- 2000; set.seed(1)
edata <- # Generate the data set. Expensive.
  data.frame(x2 = runif(nn),
    ooo = log(size1 / (size1 - 1)))
edata <- transform(edata, x1copy = 1, x2copy = x2,
  y2 = rextbetabinom(nn, size1, # Expensive
    logitlink(1 + x2, inverse = TRUE),
    cloglink(ooo + 1 - 0.5 * x2, inv = TRUE)))
fit2 <- vglm(data = edata,
  cbind(y2, N - y2) ~ x2 + x1copy + x2copy,
  extbetabinomial(zero = NULL, vfl = TRUE,
    Form2 = ~ x1copy + x2copy - 1),
  offset = cbind(0, ooo), trace = TRUE)
coef(fit2, matrix = TRUE)
wald.stat(fit2, values0 = c(1, 1, -0.5))

## End(Not run)
```

extlogF1

Extended log-F Distribution Family Function

Description

Maximum likelihood estimation of the 1-parameter extended log-F distribution.

Usage

```
extlogF1(tau = c(0.25, 0.5, 0.75), parallel = TRUE ~ 0,
  seppar = 0, tol0 = -0.001,
  llocation = "identitylink", ilocation = NULL,
  lambda.arg = NULL, scale.arg = 1, ishrinkage = 0.95,
  digt = 4, idf.mu = 3, imethod = 1)
```

Arguments

- `tau` Numeric, the desired quantiles. A strictly increasing sequence, each value must be in $(0, 1)$. The default values are the three quartiles, matching `lms.bcn`.
- `parallel` Similar to `alaplance1`, applying to the location parameters. One can try fix up the quantile-crossing problem after fitting the model by calling `fix.crossing`. Use `is.crossing` to see if there is a problem. The default for `parallel` is totally FALSE, i.e., FALSE for every variable including the intercept. Quantile-crossing can occur when values of `tau` are too close, given the data. How the quantiles are modelled with respect to the covariates also has a big effect, e.g., if they are too flexible or too inflexible then the problem is likely to occur. For example, using `bs` with `df = 10` is likely to create problems.
- Setting `parallel = TRUE` results in a totally parallel model; *all* quantiles are parallel and this assumption can be too strong for some data sets. Instead, `fix.crossing` only repairs the quantiles that cross. So one must carefully choose values of `tau` for fitting the original fit.
- `separ, tol0` Numeric, both of unit length and nonnegative, the separation and shift parameters. If `separ` is positive then any crossing quantile is penalized by the difference cubed multiplied by `separ`. The log-likelihood subtracts the penalty. The shift parameter ensures that the result is strictly noncrossing when `separ` is large enough; otherwise if `tol0 = 0` and `separ` is large then the crossing quantiles remain crossed even though the offending amount becomes small but never exactly 0. Informally, `tol0` pushes the adjustment enough so that `is.crossing` should return FALSE.
- If `tol0` is positive then that is the shift in absolute terms. But `tol0` may be assigned a negative value, in which case it is interpreted multiplicatively *relative* to the midspread of the response; `tol0 <- abs(tol0) * midspread`. Regardless, `fit@extra$tol0` is the amount in absolute terms.
- If avoiding the quantile crossing problem is of concern to you, try increasing `separ` to decrease the amount of crossing. Probably it is best to choose the smallest value of `separ` so that `is.crossing` returns FALSE. Increasing `tol0` relatively or absolutely means the fitted quantiles are allowed to move apart more. However, `tau` must be considered when choosing `tol0`.
- `llocation, ilocation` See [Links](#) for more choices and [CommonVGAMffArguments](#) for more information. Choosing `loglink` should usually be good for counts. And choosing `logitlink` should be a reasonable for proportions. However, avoid choosing `tau` values close to the boundary, for example, if p_0 is the proportion of 0s then choose $p_0 \ll \tau$. For proportions grouped data is much better than ungrouped data, and the bigger the groups the more the granularity so that the empirical proportion can approximate `tau` more closely.
- `lambda.arg` Positive tuning parameter which controls the sharpness of the cusp. The limit as it approaches 0 is probably very similar to `dalap`. The default is to choose the value internally. If `scale.arg` increases, then probably `lambda.arg` needs to increase accordingly. If `lambda.arg` is too large then the empirical quantiles may not be very close to `tau`. If `lambda.arg` is too close to 0 then the convergence behaviour will not be good and local solutions found, as well as

	numerical problems in general. Monitoring convergence is recommended when varying <code>lambda.arg</code> .
<code>scale.arg</code>	Positive scale parameter and sometimes called <code>scale</code> . The transformation used is $(y - \text{location}) / \text{scale}$. This function should be okay for response variables having a moderate range (0–100, say), but if very different from this then experimenting with this argument will be a good idea.
<code>ishrinkage</code> , <code>idf.mu</code> , <code>dig</code>	Similar to alaplace1 .
<code>imethod</code>	Initialization method. Either the value 1, 2, or . . . See CommonVGAMffArguments for more information.

Details

This is an experimental family function for quantile regression. Fasiolo et al. (2020) propose an *extended* log-F distribution (ELF) however this family function only estimates the location parameter. The distribution has a scale parameter which can be inputted (default value is unity). One location parameter is estimated for each τ value and these are the estimated quantiles. For quantile regression it is not necessary to estimate the scale parameter since the log-likelihood function is triangle shaped.

The ELF is used as an approximation of the asymmetric Laplace distribution (ALD). The latter cannot be estimated properly using Fisher scoring/IRLS but the ELF holds promise because it has continuous derivatives and therefore fewer problems with the regularity conditions. Because the ELF is fitted to data to obtain an empirical result the convergence behaviour may not be gentle and smooth. Hence there is a function-specific control function called `extlogF1.control` which has something like `stepsize = 0.5` and `maxits = 100`. It has been found that slowing down the rate of convergence produces greater stability during the estimation process. Regardless, convergence should be monitored carefully always.

This function accepts a vector response but not a matrix response.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Note

Changes will occur in the future to fine-tune things. In general setting `trace = TRUE` is strongly encouraged because it is needful to check that convergence occurs properly.

If `seppar > 0` then `logLik(fit)` will return the penalized log-likelihood.

Author(s)

Thomas W. Yee

References

Fasiolo, M., Wood, S. N., Zaffran, M., Nedellec, R. and Goude, Y. (2020). Fast calibrated additive quantile regression. *J. Amer. Statist. Assoc.*, in press.

Yee, T. W. (2020). On quantile regression based on the 1-parameter extended log-F distribution. *In preparation*.

See Also

[dextlogF](#), [is.crossing](#), [fix.crossing](#), [eCDF](#), [vglm.control](#), [logF](#), [alaplance1](#), [dalap](#), [lms.bcn](#).

Examples

```
## Not run:
nn <- 1000; mytau <- c(0.25, 0.75)
edata <- data.frame(x2 = sort(rnorm(nn)))
edata <- transform(edata, y1 = 1 + x2 + rnorm(nn, sd = exp(-1)),
                    y2 = cos(x2) / (1 + abs(x2)) + rnorm(nn, sd = exp(-1)))
fit1 <- vglm(y1 ~ x2, extlogF1(tau = mytau), data = edata) # trace = TRUE
fit2 <- vglm(y2 ~ bs(x2, 6), extlogF1(tau = mytau), data = edata)
coef(fit1, matrix = TRUE)
fit2@extra$percentile # Empirical percentiles here
summary(fit2)
c(is.crossing(fit1), is.crossing(fit2))
head(fitted(fit1))
plot(y2 ~ x2, edata, col = "blue")
matlines(with(edata, x2), fitted(fit2), col="orange", lty = 1, lwd = 2)
## End(Not run)
```

familyname

Family Function Name

Description

Extractor function for the name of the family function of an object in the **VGAM** package.

Usage

```
familyname(object, ...)
familyname.vlm(object, all = FALSE, ...)
```

Arguments

object	Some VGAM object, for example, having class vglmff-class .
all	If all = TRUE then all of the vfamily slot is returned; this contains subclasses the object might have. The default is the return the first value only.
...	Other possible arguments for the future.

Details

Currently **VGAM** implements over 150 family functions. This function returns the name of the function assigned to the family argument, for modelling functions such as `vglm` and `vgam`. Sometimes a slightly different answer is returned, e.g., `propodds` really calls `cumulative` with some arguments set, hence the output returned by this function is "cumulative" (note that one day this might change, however).

Value

A character string or vector.

Note

Arguments used in the invocation are not included. Possibly this is something to be done in the future.

See Also

[vglmff-class](#), [vglm-class](#).

Examples

```
pneumo <- transform(pneumo, let = log(exposure.time))
fit1 <- vglm(cbind(normal, mild, severe) ~ let,
            cumulative(parallel = TRUE, reverse = TRUE), data = pneumo)
familyname(fit1)
familyname(fit1, all = TRUE)
familyname(propodds()) # "cumulative"
```

Felix

The Felix Distribution

Description

Density for the Felix distribution.

Usage

```
dfelix(x, rate = 0.25, log = FALSE)
```

Arguments

<code>x</code>	vector of quantiles.
<code>rate</code>	See felix .
<code>log</code>	Logical. If <code>log = TRUE</code> then the logarithm of the density is returned.

Details

See [felix](#), the **VGAM** family function for estimating the parameter, for the formula of the probability density function and other details.

Value

`dfelix` gives the density.

Warning

The default value of `rate` is subjective.

Author(s)

T. W. Yee

See Also

[felix](#).

Examples

```
## Not run:
rate <- 0.25; x <- 1:15
plot(x, dfelix(x, rate), type = "h", las = 1, col = "blue",
      ylab = paste("dfelix(rate=", rate, ")"),
      main = "Felix density function")

## End(Not run)
```

felix

Felix Distribution Family Function

Description

Estimates the parameter of a Felix distribution by maximum likelihood estimation.

Usage

```
felix(lrate = extlogitlink(min = 0, max = 0.5), imethod = 1)
```

Arguments

`lrate` Link function for the parameter, called a below; see [Links](#) for more choices and for general information.

`imethod` See [CommonVGAMffArguments](#). Valid values are 1, 2, 3 or 4.

Details

The Felix distribution is an important basic Lagrangian distribution. The density function is

$$f(y; a) = \frac{1}{((y-1)/2)!} y^{(y-3)/2} a^{(y-1)/2} \exp(-ay)$$

where $y = 1, 3, 5, \dots$ and $0 < a < 0.5$. The mean is $1/(1 - 2a)$ (returned as the fitted values). Fisher scoring is implemented.

Value

An object of class "vglmfff" (see [vglmfff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Author(s)

T. W. Yee

References

Consul, P. C. and Famoye, F. (2006). *Lagrangian Probability Distributions*, Boston, USA: Birkhauser.

See Also

[dfelix](#), [borel.tanner](#).

Examples

```
fdata <- data.frame(y = 2 * rpois(n = 200, 1) + 1) # Not real data!
fit <- vglm(y ~ 1, felix, data = fdata, trace = TRUE, crit = "coef")
coef(fit, matrix = TRUE)
Coef(fit)
summary(fit)
```

fff

F Distribution Family Function

Description

Maximum likelihood estimation of the (2-parameter) F distribution.

Usage

```
fff(link = "loglink", idf1 = NULL, idf2 = NULL, nsimEIM = 100,
    imethod = 1, zero = NULL)
```


Arguments

link	Parameter link function for both parameters. See Links for more choices. The default keeps the parameters positive.
idf1, idf2	Numeric and positive. Initial value for the parameters. The default is to choose each value internally.
nsimEIM, zero	See CommonVGAMffArguments for more information.
imethod	Initialization method. Either the value 1 or 2. If both fail try setting values for idf1 and idf2.

Details

The F distribution is named after Fisher and has a density function that has two parameters, called $df1$ and $df2$ here. This function treats these degrees of freedom as *positive reals* rather than integers. The mean of the distribution is $df2/(df2 - 2)$ provided $df2 > 2$, and its variance is $2df2^2(df1 + df2 - 2)/(df1(df2 - 2)^2(df2 - 4))$ provided $df2 > 4$. The estimated mean is returned as the fitted values. Although the F distribution can be defined to accommodate a non-centrality parameter ncp , it is assumed zero here. Actually it shouldn't be too difficult to handle any known ncp ; something to do in the short future.

Value

An object of class "vg1mff" (see [vg1mff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Warning

Numerical problems will occur when the estimates of the parameters are too low or too high.

Author(s)

T. W. Yee

References

Forbes, C., Evans, M., Hastings, N. and Peacock, B. (2011). *Statistical Distributions*, Hoboken, NJ, USA: John Wiley and Sons, Fourth edition.

See Also

[FDist](#).

Examples

```
## Not run:
fdata <- data.frame(x2 = runif(nn <- 2000))
fdata <- transform(fdata, df1 = exp(2+0.5*x2),
                  df2 = exp(2-0.5*x2))
fdata <- transform(fdata, y = rf(nn, df1, df2))
fit <- vglm(y ~ x2, fff, data = fdata, trace = TRUE)
```

```
coef(fit, matrix = TRUE)

## End(Not run)
```

fill1

Creates a Matrix of Appropriate Dimension

Description

A support function for the argument `xij`, it generates a matrix of an appropriate dimension.

Usage

```
fill1(x, values = 0, ncolx = ncol(x))
```

Arguments

<code>x</code>	A vector or matrix which is used to determine the dimension of the answer, in particular, the number of rows. After converting <code>x</code> to a matrix if necessary, the answer is a matrix of values <code>values</code> , of dimension <code>nrow(x)</code> by <code>ncolx</code> .
<code>values</code>	Numeric. The answer contains these values, which are recycled <i>columnwise</i> if necessary, i.e., as <code>matrix(values, ..., byrow=TRUE)</code> .
<code>ncolx</code>	The number of columns of the returned matrix. The default is the number of columns of <code>x</code> .

Details

The `xij` argument for `vglm` allows the user to input variables specific to each linear/additive predictor. For example, consider the bivariate logit model where the first/second linear/additive predictor is the logistic regression of the first/second binary response respectively. The third linear/additive predictor is $\log(\text{OR}) = \eta_3$, where OR is the odds ratio. If one has ocular pressure as a covariate in this model then `xij` is required to handle the ocular pressure for each eye, since these will be different in general. [This contrasts with a variable such as age, the age of the person, which has a common value for both eyes.] In order to input these data into `vglm` one often finds that functions `fill1`, `fill2`, etc. are useful.

All terms in the `xij` and `formula` arguments in `vglm` must appear in the `form2` argument too.

Value

`matrix(values, nrow=nrow(x), ncol=ncolx)`, i.e., a matrix consisting of values `values`, with the number of rows matching `x`, and the default number of columns is the number of columns of `x`.

Note

The effect of the `xij` argument is after other arguments such as `exchangeable` and `zero`. Hence `xij` does not affect constraint matrices.

Additionally, there are currently 3 other identical `fill1` functions, called `fill2`, `fill3` and `fill4`; if you need more then assign `fill5 = fill6 = fill1` etc. The reason for this is that if more than one `fill1` function is needed then they must be unique. For example, if $M = 4$ then `xij = list(op ~ lop + rop + fill1(mop) + fill1(mop))` would reduce to `xij = list(op ~ lop + rop + fill1(mop))`, whereas `xij = list(op ~ lop + rop + fill1(mop) + fill2(mop))` would retain all M terms, which is needed.

In Examples 1 to 3 below, the `xij` argument illustrates covariates that are specific to a linear predictor. Here, `lop/rop` are the ocular pressures of the left/right eye in an artificial dataset, and `mop` is their mean. Variables `leye` and `reye` might be the presence/absence of a particular disease on the LHS/RHS eye respectively.

In Example 3, the `xij` argument illustrates fitting the (exchangeable) model where there is a common smooth function of the ocular pressure. One should use regression splines since `s` in `vgam` does not handle the `xij` argument. However, regression splines such as `bs` and `ns` need to have the same basis functions here for both functions, and Example 3 illustrates a trick involving a function `BS` to obtain this, e.g., same knots. Although regression splines create more than a single column per term in the model matrix, `fill1(BS(lop, rop))` creates the required (same) number of columns.

Author(s)

T. W. Yee

See Also

[vglm.control](#), [vglm](#), [multinomial](#), [Select](#).

Examples

```
fill1(runif(5))
fill1(runif(5), ncol = 3)
fill1(runif(5), val = 1, ncol = 3)

# Generate (independent) eyes data for the examples below; OR=1.
## Not run: nn <- 1000 # Number of people
eyesdata <- data.frame(lop = round(runif(nn), 2),
                      rop = round(runif(nn), 2),
                      age = round(rnorm(nn, 40, 10)))
eyesdata <- transform(eyesdata,
  mop = (lop + rop) / 2,      # Mean ocular pressure
  op  = (lop + rop) / 2,      # Value unimportant unless plotting
  # op = lop,                # Choose this if plotting
  eta1 = 0 - 2*lop + 0.04*age, # Linear predictor for left eye
  eta2 = 0 - 2*rop + 0.04*age) # Linear predictor for right eye
eyesdata <- transform(eyesdata,
  leye = rbinom(nn, size=1, prob = logitlink(eta1, inverse = TRUE)),
  reye = rbinom(nn, size=1, prob = logitlink(eta2, inverse = TRUE)))
```



```

          y1 + y2 + y3 + y4 + y5),
form2 = ~ sex + weight + Lag1 +
          fill1(y1) + fill1(y2) + fill1(y3) + fill1(y4) +
          fill1(y5) + fill1(y6) +
          y1 + y2 + y3 + y4 + y5 + y6,
data = deermice, trace = TRUE)
coef(M.tbh.lag1)
## End(Not run)

```

 finney44

Toxicity trial for insects

Description

A data frame of a toxicity trial.

Usage

```
data(finney44)
```

Format

A data frame with 6 observations on the following 3 variables.

pconc a numeric vector, percent concentration of pyrethrins.

hatched number of eggs that hatched.

unhatched number of eggs that did not hatch.

Details

Finney (1944) describes a toxicity trial of five different concentrations of pyrethrins (percent) plus a control that were administered to eggs of *Ephestia kuhniella*. The natural mortality rate is large, and a common adjustment is to use Abbott's formula.

References

Finney, D. J. (1944). The application of the probit method to toxicity test data adjusted for mortality in the controls. *Annals of Applied Biology*, **31**, 68–74.

Abbott, W. S. (1925). A method of computing the effectiveness of an insecticide. *Journal of Economic Entomology*, 18, 265–7.

Examples

```

data(finney44)
transform(finney44, mortality = unhatched / (hatched + unhatched))

```

fisherzlink

*Fisher's Z Link Function***Description**

Computes the Fisher Z transformation, including its inverse and the first two derivatives.

Usage

```
fisherzlink(theta, bminvalue = NULL, bmaxvalue = NULL,
            inverse = FALSE, deriv = 0, short = TRUE, tag = FALSE)
```

Arguments

`theta` Numeric or character. See below for further details.

`bminvalue, bmaxvalue` Optional boundary values. Values of `theta` which are less than or equal to -1 can be replaced by `bminvalue` before computing the link function value. Values of `theta` which are greater than or equal to 1 can be replaced by `bmaxvalue` before computing the link function value. See [Links](#).

`inverse, deriv, short, tag` Details at [Links](#).

Details

The fisherz link function is commonly used for parameters that lie between -1 and 1 . Numerical values of `theta` close to -1 or 1 or out of range result in `Inf`, `-Inf`, `NA` or `NaN`.

Value

For `deriv = 0`, $0.5 * \log((1+\text{theta})/(1-\text{theta}))$ (same as `atanh(theta)`) when `inverse = FALSE`, and if `inverse = TRUE` then $(\exp(2*\text{theta})-1)/(\exp(2*\text{theta})+1)$ (same as `tanh(theta)`).

For `deriv = 1`, then the function returns $d \text{ eta} / d \text{ theta}$ as a function of `theta` if `inverse = FALSE`, else if `inverse = TRUE` then it returns the reciprocal.

Here, all logarithms are natural logarithms, i.e., to base e .

Note

Numerical instability may occur when `theta` is close to -1 or 1 . One way of overcoming this is to use, e.g., `bminvalue`.

The link function `rhobitlink` is very similar to `fisherzlink`, e.g., just twice the value of `fisherzlink`. This link function may be renamed to `atanhlink` in the near future.

Author(s)

Thomas W. Yee

References

McCullagh, P. and Nelder, J. A. (1989). *Generalized Linear Models*, 2nd ed. London: Chapman & Hall.

See Also

[Links](#), [rhobitlink](#), [logitlink](#).

Examples

```
theta <- seq(-0.99, 0.99, by = 0.01)
y <- fisherzlink(theta)
## Not run: plot(theta, y, type = "l", las = 1, ylab = "",
  main = "fisherzlink(theta)", col = "blue")
abline(v = (-1):1, h = 0, lty = 2, col = "gray")
## End(Not run)

x <- c(seq(-1.02, -0.98, by = 0.01), seq(0.97, 1.02, by = 0.01))
fisherzlink(x) # Has NAs
fisherzlink(x, bminvalue = -1 + .Machine$double.eps,
  bmaxvalue = 1 - .Machine$double.eps) # Has no NAs
```

Fisk

The Fisk Distribution

Description

Density, distribution function, quantile function and random generation for the Fisk distribution with shape parameter a and scale parameter $scale$.

Usage

```
dfisk(x, scale = 1, shape1.a, log = FALSE)
pfisk(q, scale = 1, shape1.a, lower.tail = TRUE, log.p = FALSE)
qfisk(p, scale = 1, shape1.a, lower.tail = TRUE, log.p = FALSE)
rfisk(n, scale = 1, shape1.a)
```

Arguments

<code>x, q</code>	vector of quantiles.
<code>p</code>	vector of probabilities.
<code>n</code>	number of observations. If <code>length(n) > 1</code> then the length is taken to be the number required.
<code>shape1.a</code>	shape parameter.
<code>scale</code>	scale parameter.
<code>log</code>	Logical. If <code>log = TRUE</code> then the logarithm of the density is returned.
<code>lower.tail, log.p</code>	Same meaning as in pnorm or qnorm .

Details

See [fisk](#), which is the **VGAM** family function for estimating the parameters by maximum likelihood estimation.

Value

`dfisk` gives the density, `pfisk` gives the distribution function, `qfisk` gives the quantile function, and `rfisk` generates random deviates.

Note

The Fisk distribution is a special case of the 4-parameter generalized beta II distribution.

Author(s)

T. W. Yee and Kai Huang

References

Kleiber, C. and Kotz, S. (2003). *Statistical Size Distributions in Economics and Actuarial Sciences*, Hoboken, NJ, USA: Wiley-Interscience.

See Also

[fisk](#), [genbetaII](#).

Examples

```
fdata <- data.frame(y = rfisk(1000, shape = exp(1), scale = exp(2)))
fit <- vglm(y ~ 1, fisk(lss = FALSE), data = fdata, trace = TRUE)
coef(fit, matrix = TRUE)
Coef(fit)
```

fisk

Fisk Distribution family function

Description

Maximum likelihood estimation of the 2-parameter Fisk distribution.

Usage

```
fisk(lscale = "loglink", lshape1.a = "loglink", iscale = NULL,
     ishape1.a = NULL, imethod = 1, lss = TRUE,
     gscale = exp(-5:5), gshape1.a = seq(0.75, 4, by = 0.25),
     probs.y = c(0.25, 0.5, 0.75), zero = "shape")
```


Arguments

lss	See CommonVGAMffArguments for important information.
lshape1.a, lscale	Parameter link functions applied to the (positive) parameters a and scale. See Links for more choices.
iscale, ishape1.a, imethod, zero	See CommonVGAMffArguments for information. For <code>imethod = 2</code> a good initial value for <code>iscale</code> is needed to obtain a good estimate for the other parameter.
gscale, gshape1.a	See CommonVGAMffArguments for information.
probs.y	See CommonVGAMffArguments for information.

Details

The 2-parameter Fisk (aka log-logistic) distribution is the 4-parameter generalized beta II distribution with shape parameter $q = p = 1$. It is also the 3-parameter Singh-Maddala distribution with shape parameter $q = 1$, as well as the Dagum distribution with $p = 1$. More details can be found in Kleiber and Kotz (2003).

The Fisk distribution has density

$$f(y) = ay^{a-1}/[b^a\{1 + (y/b)^a\}^2]$$

for $a > 0$, $b > 0$, $y \geq 0$. Here, b is the scale parameter `scale`, and a is a shape parameter. The cumulative distribution function is

$$F(y) = 1 - [1 + (y/b)^a]^{-1} = [1 + (y/b)^{-a}]^{-1}.$$

The mean is

$$E(Y) = b\Gamma(1 + 1/a)\Gamma(1 - 1/a)$$

provided $a > 1$; these are returned as the fitted values. This family function handles multiple responses.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Note

See the notes in [genbetaII](#).

Author(s)

T. W. Yee

References

Kleiber, C. and Kotz, S. (2003). *Statistical Size Distributions in Economics and Actuarial Sciences*, Hoboken, NJ, USA: Wiley-Interscience.

See Also

[Fisk](#), [genbetaII](#), [betaII](#), [dagum](#), [sinmad](#), [inv.lomax](#), [lomax](#), [paralogistic](#), [inv.paralogistic](#), [simulate.vlm](#).

Examples

```
fdata <- data.frame(y = rfisk(200, shape = exp(1), exp(2)))
fit <- vglm(y ~ 1, fisk(lss = FALSE), data = fdata, trace = TRUE)
fit <- vglm(y ~ 1, fisk(ishape1.a = exp(2)), fdata, trace = TRUE)
coef(fit, matrix = TRUE)
Coef(fit)
summary(fit)
```

fittedvlm

Fitted Values of a VLM object

Description

Extractor function for the fitted values of a model object that inherits from a *vector linear model* (VLM), e.g., a model of class "vglm".

Usage

```
fittedvlm(object, drop = FALSE, type.fitted = NULL,
          percentiles = NULL, ...)
```

Arguments

object	a model object that inherits from a VLM.
drop	Logical. If FALSE then the answer is a matrix. If TRUE then the answer is a vector.
type.fitted	Character. Some VGAM family functions have a type.fitted argument. If so then a different type of fitted value can be returned. It is recomputed from the model after convergence. Note: this is an experimental feature and not all VGAM family functions have this implemented yet. See CommonVGAMffArguments for more details.
percentiles	See CommonVGAMffArguments for details.
...	Currently unused.

Details

The “fitted values” usually corresponds to the mean response, however, because the **VGAM** package fits so many models, this sometimes refers to quantities such as quantiles. The mean may even not exist, e.g., for a Cauchy distribution.

Note that the fitted value is output from the @linkinv slot of the **VGAM** family function, where the eta argument is the $n \times M$ matrix of linear predictors.

Value

The fitted values evaluated at the final IRLS iteration.

Note

This function is one of several extractor functions for the **VGAM** package. Others include `coef`, `deviance`, `weights` and `constraints` etc. This function is equivalent to the `methods` function for the generic function `fitted.values`.

If `fit` is a VLM or VGLM then `fitted(fit)` and `predict(fit, type = "response")` should be equivalent (see [predictvglm](#)). The latter has the advantage in that it handles a `newdata` argument so that the fitted values can be computed for a different data set.

Author(s)

Thomas W. Yee

References

Chambers, J. M. and T. J. Hastie (eds) (1992). *Statistical Models in S*. Wadsworth & Brooks/Cole.

See Also

[fitted](#), [predictvglm](#), [vglmff-class](#).

Examples

```
# Categorical regression example 1
pneumo <- transform(pneumo, let = log(exposure.time))
(fit1 <- vglm(cbind(normal, mild, severe) ~ let, propodds, pneumo))
fitted(fit1)

# LMS quantile regression example 2
fit2 <- vgam(BMI ~ s(age, df = c(4, 2)),
             lms.bcn(zero = 1), data = bmi.nz, trace = TRUE)
head(predict(fit2, type = "response")) # Equals to both these:
head(fitted(fit2))
predict(fit2, type = "response", newdata = head(bmi.nz))

# Zero-inflated example 3
zdata <- data.frame(x2 = runif(nn <- 1000))
zdata <- transform(zdata,
                  pstr0.3 = logitlink(-0.5, inverse = TRUE),
                  lambda.3 = loglink(-0.5 + 2*x2, inverse = TRUE))
zdata <- transform(zdata,
                  y1 = rzipois(nn, lambda = lambda.3, pstr0 = pstr0.3))
fit3 <- vglm(y1 ~ x2, zipoisson(zero = NULL), zdata, trace = TRUE)
head(fitted(fit3, type.fitted = "mean" )) # E(Y) (the default)
head(fitted(fit3, type.fitted = "pobs0")) # Pr(Y = 0)
head(fitted(fit3, type.fitted = "pstr0")) # Prob of a structural 0
head(fitted(fit3, type.fitted = "onempstr0")) # 1 - Pr(structural 0)
```

 fix.crossing

Fixing a Quantile Regression having Crossing

Description

Returns a similar object fitted with columns of the constraint matrices amalgamated so it is a partially parallel VGLM object. The columns combined correspond to certain crossing quantiles. This applies especially to an `extlogF1()` VGLM object.

Usage

```
fix.crossing.vglm(object, maxit = 100, trace = FALSE, ...)
```

Arguments

<code>object</code>	an object such as a <code>vglm</code> object with family function <code>extlogF1</code> .
<code>maxit, trace</code>	values for overwriting components in <code>vglm.control</code> . Setting these to <code>NULL</code> will mean the values in <code>vglm.control</code> on <code>object</code> will be retained.
<code>...</code>	additional optional arguments. Currently unused.

Details

The quantile crossing problem has been described as *disturbing* and *embarrassing*. This function was specifically written for a `vglm` with family function `extlogF1`. It examines the fitted quantiles of `object` to see if any cross. If so, then a pair of columns is combined to make those two quantiles parallel. After fitting the submodel it then repeats testing for crossing quantiles and repairing them, until there is no more quantile crossing detected. Note that it is possible that the quantiles cross in some subset of the covariate space not covered by the data—see `is.crossing`.

This function is fragile and likely to change in the future. For `extlogF1` models, it is assumed that argument `data` has been assigned a data frame, and that the default values of the argument `parallel` has been used; this means that the second constraint matrix is $\text{diag}(M)$. The constraint matrix of the intercept term remains unchanged as $\text{diag}(M)$.

Value

An object very similar to the original object, but with possibly different constraint matrices (partially parallel) so as to remove any quantile crossing.

See Also

`extlogF1`, `is.crossing`, `lms.bcn.vglm`.

Examples

```
## Not run: ooo <- with(bmi.nz, order(age))
bmi.nz <- bmi.nz[ooo, ] # Sort by age
with(bmi.nz, plot(age, BMI, col = "blue"))
mytau <- c(50, 93, 95, 97) / 100 # Some quantiles are quite close
fit1 <- vglm(BMI ~ ns(age, 7), extlogF1(mytau), bmi.nz, trace = TRUE)
plot(BMI ~ age, bmi.nz, col = "blue", las = 1,
     main = "Partially parallel (darkgreen) & nonparallel quantiles",
     sub = "Crossing quantiles are orange")
fix.crossing(fit1)
matlines(with(bmi.nz, age), fitted(fit1), lty = 1, col = "orange")
fit2 <- fix.crossing(fit1) # Some quantiles have been fixed
constraints(fit2)
matlines(with(bmi.nz, age), fitted(fit2), lty = "dashed",
         col = "darkgreen", lwd = 2)
## End(Not run)
```

 flourbeetle

Mortality of Flour Beetles from Carbon Disulphide

Description

The flourbeetle data frame has 8 rows and 4 columns. Two columns are explanatory, the other two are responses.

Usage

```
data(flourbeetle)
```

Format

This data frame contains the following columns:

logdose `log10` applied to CS2mgL.

CS2mgL a numeric vector, the concentration of gaseous carbon disulphide in mg per litre.

exposed a numeric vector, counts; the number of beetles exposed to the poison.

killed a numeric vector, counts; the numbers killed.

Details

These data were originally given in Table IV of Bliss (1935) and are the combination of two series of toxicological experiments involving *Tribolium confusum*, also known as the flour beetle. Groups of such adult beetles were exposed for 5 hours of gaseous carbon disulphide at different concentrations, and their mortality measured.

Source

Bliss, C.I., 1935. The calculation of the dosage-mortality curve. *Annals of Applied Biology*, **22**, 134–167.

See Also

[binomialff](#), [probitlink](#).

Examples

```
fit1 <- vglm(cbind(killed, exposed - killed) ~ logdose,
            binomialff(link = probitlink), flourbeetle, trace = TRUE)
summary(fit1)
```

Foldnorm

*The Folded-Normal Distribution***Description**

Density, distribution function, quantile function and random generation for the (generalized) folded-normal distribution.

Usage

```
dfoldnorm(x, mean = 0, sd = 1, a1 = 1, a2 = 1, log = FALSE)
pfoldnorm(q, mean = 0, sd = 1, a1 = 1, a2 = 1,
          lower.tail = TRUE, log.p = FALSE)
qfoldnorm(p, mean = 0, sd = 1, a1 = 1, a2 = 1,
          lower.tail = TRUE, log.p = FALSE, ...)
rfoldnorm(n, mean = 0, sd = 1, a1 = 1, a2 = 1)
```

Arguments

x, q	vector of quantiles.
p	vector of probabilities.
n	number of observations. Same as rnorm .
mean, sd	see rnorm .
a1, a2	see foldnormal .
log	Logical. If TRUE then the log density is returned.
lower.tail, log.p	Same meaning as in pnorm or qnorm .
...	Arguments that can be passed into uniroot .

Details

See [foldnormal](#), the **VGAM** family function for estimating the parameters, for the formula of the probability density function and other details.

Value

`dfoldnorm` gives the density, `pfoldnorm` gives the distribution function, `qfoldnorm` gives the quantile function, and `rfoldnorm` generates random deviates.

Author(s)

T. W. Yee and Kai Huang. Suggestions from Mauricio Romero led to improvements in `qfoldnorm()`.

See Also

[foldnormal](#), [uniroot](#).

Examples

```
## Not run:
m <- 1.5; SD <- exp(0)
x <- seq(-1, 4, len = 501)
plot(x, dfoldnorm(x, m = m, sd = SD), type = "l", ylim = 0:1,
      ylab = paste("foldnorm(m = ", m, ", sd = ",
                  round(SD, digits = 3), ")"), las = 1,
      main = "Blue is density, orange is CDF", col = "blue",
      sub = "Purple lines are the 10,20,...,90 percentiles")
abline(h = 0, col = "gray50")
lines(x, pfoldnorm(x, m = m, sd = SD), col = "orange")
probs <- seq(0.1, 0.9, by = 0.1)
Q <- qfoldnorm(probs, m = m, sd = SD)
lines(Q, dfoldnorm(Q, m, SD), col = "purple", lty = 3, type = "h")
lines(Q, pfoldnorm(Q, m, SD), col = "purple", lty = 3, type = "h")
abline(h = probs, col = "purple", lty = 3)
max(abs(pfoldnorm(Q, m = m, sd = SD) - probs)) # Should be 0

## End(Not run)
```

foldnormal

Folded Normal Distribution Family Function

Description

Fits a (generalized) folded (univariate) normal distribution.

Usage

```
foldnormal(lmean = "identitylink", lsd = "loglink", imean = NULL,
           isd = NULL, a1 = 1, a2 = 1, nsimEIM = 500, imethod = 1,
           zero = NULL)
```

Arguments

<code>lmean</code> , <code>lsd</code>	Link functions for the mean and standard deviation parameters of the usual univariate normal distribution. They are μ and σ respectively. See Links for more choices.
<code>imean</code> , <code>isd</code>	Optional initial values for μ and σ . A NULL means a value is computed internally. See CommonVGAMffArguments .

`a1, a2` Positive weights, called a_1 and a_2 below. Each must be of length 1.
`nsimEIM, imethod, zero`
 See [CommonVGAMffArguments](#).

Details

If a random variable has an ordinary univariate normal distribution then the absolute value of that random variable has an ordinary *folded normal distribution*. That is, the sign has not been recorded; only the magnitude has been measured.

More generally, suppose X is normal with mean `mean` and standard deviation `sd`. Let $Y = \max(a_1X, -a_2X)$ where a_1 and a_2 are positive weights. This means that $Y = a_1X$ for $X > 0$, and $Y = a_2X$ for $X < 0$. Then Y is said to have a *generalized folded normal distribution*. The ordinary folded normal distribution corresponds to the special case $a_1 = a_2 = 1$.

The probability density function of the ordinary folded normal distribution can be written `dnorm(y, mean, sd) + dnorm(y, -mean, sd)` for $y \geq 0$. By default, `mean` and `log(sd)` are the linear/additive predictors. Having `mean=0` and `sd=1` results in the *half-normal* distribution. The mean of an ordinary folded normal distribution is

$$E(Y) = \sigma \sqrt{2/\pi} \exp(-\mu^2/(2\sigma^2)) + \mu[1 - 2\Phi(-\mu/\sigma)]$$

and these are returned as the fitted values. Here, $\Phi()$ is the cumulative distribution function of a standard normal ([pnorm](#)).

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Warning

Under- or over-flow may occur if the data is ill-conditioned. It is recommended that several different initial values be used to help avoid local solutions.

Note

The response variable for this family function is the same as [uninormal](#) except positive values are required. Reasonably good initial values are needed. Fisher scoring using simulation is implemented.

See [CommonVGAMffArguments](#) for general information about many of these arguments.

Yet to do: implement the results of Johnson (1962) which gives expressions for the EIM, albeit, under a different parameterization. Also, one element of the EIM appears to require numerical integration.

Author(s)

Thomas W. Yee

References

- Lin, P. C. (2005). Application of the generalized folded-normal distribution to the process capability measures. *International Journal of Advanced Manufacturing Technology*, **26**, 825–830.
- Johnson, N. L. (1962). The folded normal distribution: accuracy of estimation by maximum likelihood. *Technometrics*, **4**, 249–256.

See Also

[rfoldnorm](#), [uninormal](#), [dnorm](#), [skewnormal](#).

Examples

```
## Not run: m <- 2; SD <- exp(1)
fdata <- data.frame(y = rfoldnorm(n <- 1000, m = m, sd = SD))
hist(with(fdata, y), prob = TRUE, main = paste("foldnormal(m = ",
      m, ", sd = ", round(SD, 2), ")"))
fit <- vglm(y ~ 1, foldnormal, data = fdata, trace = TRUE)
coef(fit, matrix = TRUE)
(Cfit <- Coef(fit))
# Add the fit to the histogram:
mygrid <- with(fdata, seq(min(y), max(y), len = 200))
lines(mygrid, dfoldnorm(mygrid, Cfit[1], Cfit[2]), col = "orange")

## End(Not run)
```

formulavlm

Model Formulae and Term Names for VGLMs

Description

The methods function for formula to extract the formula from a fitted object, as well as a methods function to return the names of the terms in the formula.

Usage

```
## S3 method for class 'vlm'
formula(x, ...)
formulavlm(x, form.number = 1, ...)
term.names(model, ...)
term.namesvml(model, form.number = 1, ...)
```

Arguments

x, model	A fitted model object.
form.number	Formula number, is 1 or 2. which correspond to the arguments formula and form2 respectively.
...	Same as formula .

Details

The formula methods function is based on [formula](#).

Value

The formula methods function should return something similar to [formula](#). The `term.names` methods function should return a character string with the terms in the formula; this includes any intercept (which is denoted by "(Intercept)" as the first element.)

Author(s)

Thomas W. Yee

See Also

[has.interceptvglm](#).

Examples

```
# Example: this is based on a glm example
counts <- c(18,17,15,20,10,20,25,13,12)
outcome <- gl(3, 1, 9); treatment <- gl(3, 3)
vglm.D93 <- vglm(counts ~ outcome + treatment, family = poissonff)
formula(vglm.D93)
pdata <- data.frame(counts, outcome, treatment) # Better style
vglm.D93 <- vglm(counts ~ outcome + treatment, poissonff, data = pdata)
formula(vglm.D93)
term.names(vglm.D93)
responseName(vglm.D93)
has.intercept(vglm.D93)
```

Frank

Frank's Bivariate Distribution

Description

Density, distribution function, and random generation for the (one parameter) bivariate Frank distribution.

Usage

```
dbifrankcop(x1, x2, apar, log = FALSE)
pbifrankcop(q1, q2, apar)
rbifrankcop(n, apar)
```

Arguments

x1, x2, q1, q2	vector of quantiles.
n	number of observations. Same as in runif .
apar	the positive association parameter.
log	Logical. If log = TRUE then the logarithm of the density is returned.

Details

See [bifrankcop](#), the **VGAM** family functions for estimating the association parameter by maximum likelihood estimation, for the formula of the cumulative distribution function and other details.

Value

`dbifrankcop` gives the density, `pbifrankcop` gives the distribution function, and `rbifrankcop` generates random deviates (a two-column matrix).

Author(s)

T. W. Yee

References

Genest, C. (1987). Frank's family of bivariate distributions. *Biometrika*, **74**, 549–555.

See Also

[bifrankcop](#).

Examples

```
## Not run: N <- 100; apar <- exp(2)
xx <- seq(-0.30, 1.30, len = N)
ox <- expand.grid(xx, xx)
zedd <- dbifrankcop(ox[, 1], ox[, 2], apar = apar)
contour(xx, xx, matrix(zedd, N, N))
zedd <- pbifrankcop(ox[, 1], ox[, 2], apar = apar)
contour(xx, xx, matrix(zedd, N, N))

plot(rr <- rbifrankcop(n = 3000, apar = exp(4)))
par(mfrow = c(1, 2))
hist(rr[, 1]); hist(rr[, 2]) # Should be uniform

## End(Not run)
```

Description

Density, distribution function, quantile function and random generation for the three parameter Frechet distribution.

Usage

```
dfrechet(x, location = 0, scale = 1, shape, log = FALSE)
pfrechet(q, location = 0, scale = 1, shape,
         lower.tail = TRUE, log.p = FALSE)
qfrechet(p, location = 0, scale = 1, shape,
         lower.tail = TRUE, log.p = FALSE)
rfrechet(n, location = 0, scale = 1, shape)
```

Arguments

x, q	vector of quantiles.
p	vector of probabilities.
n	number of observations. Passed into runif .
location, scale, shape	the location parameter <i>a</i> , scale parameter <i>b</i> , and shape parameter <i>s</i> .
log	Logical. If log = TRUE then the logarithm of the density is returned.
lower.tail, log.p	Same meaning as in punif or qunif .

Details

See [frechet](#), the **VGAM** family function for estimating the 2 parameters (without location parameter) by maximum likelihood estimation, for the formula of the probability density function and range restrictions on the parameters.

Value

`dfrechet` gives the density, `pfrechet` gives the distribution function, `qfrechet` gives the quantile function, and `rfrechet` generates random deviates.

Author(s)

T. W. Yee and Kai Huang

References

Castillo, E., Hadi, A. S., Balakrishnan, N. and Sarabia, J. S. (2005). *Extreme Value and Related Models with Applications in Engineering and Science*, Hoboken, NJ, USA: Wiley-Interscience.

See Also

[frechet](#).

Examples

```
## Not run: shape <- 5
x <- seq(-0.1, 3.5, length = 401)
plot(x, dfrechet(x, shape = shape), type = "l", ylab = "",
      main = "Frechet density divided into 10 equal areas",
      sub = "Orange = CDF", las = 1)
abline(h = 0, col = "blue", lty = 2)
qq <- qfrechet(seq(0.1, 0.9, by = 0.1), shape = shape)
lines(qq, dfrechet(qq, shape = shape), col = 2, lty = 2, type = "h")
lines(x, pfrechet(q = x, shape = shape), col = "orange")

## End(Not run)
```

 frechet

Frechet Distribution Family Function

Description

Maximum likelihood estimation of the 2-parameter Frechet distribution.

Usage

```
frechet(location = 0, lscale = "loglink",
        lshape = logofflink(offset = -2),
        iscale = NULL, ishape = NULL, nsimEIM = 250, zero = NULL)
```

Arguments

`location` Numeric. Location parameter. It is called a below.
`lscale, lshape` Link functions for the parameters; see [Links](#) for more choices.
`iscale, ishape, zero, nsimEIM`
 See [CommonVGAMffArguments](#) for information.

Details

The (3-parameter) Frechet distribution has a density function that can be written

$$f(y) = \frac{sb}{(y-a)^2} [b/(y-a)]^{s-1} \exp[-(b/(y-a))^s]$$

for $y > a$ and scale parameter $b > 0$. The positive shape parameter is s . The cumulative distribution function is

$$F(y) = \exp[-(b/(y-a))^s].$$

The mean of Y is $a + b\Gamma(1 - 1/s)$ for $s > 1$ (these are returned as the fitted values). The variance of Y is $b^2[\Gamma(1 - 2/s) - \Gamma^2(1 - 1/s)]$ for $s > 2$.

Family `frechet` has a known, and $\log(b)$ and $\log(s - 2)$ are the default linear/additive predictors. The working weights are estimated by simulated Fisher scoring.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Warning

Family function `frechet` may fail for low values of the shape parameter, e.g., near 2 or lower.

Author(s)

T. W. Yee

References

Castillo, E., Hadi, A. S., Balakrishnan, N. and Sarabia, J. S. (2005). *Extreme Value and Related Models with Applications in Engineering and Science*, Hoboken, NJ, USA: Wiley-Interscience.

See Also

[rfrechet](#), [gev](#).

Examples

```
## Not run:
set.seed(123)
fdata <- data.frame(y1 = rfrechet(1000, shape = 2 + exp(1)))
with(fdata, hist(y1))
fit2 <- vglm(y1 ~ 1, frechet, data = fdata, trace = TRUE)
coef(fit2, matrix = TRUE)
Coef(fit2)
head(fitted(fit2))
with(fdata, mean(y1))
head(weights(fit2, type = "working"))
vcov(fit2)

## End(Not run)
```

freund61

*Freund's (1961) Bivariate Extension of the Exponential Distribution***Description**

Estimate the four parameters of the Freund (1961) bivariate extension of the exponential distribution by maximum likelihood estimation.

Usage

```
freund61(la = "loglink", lap = "loglink", lb = "loglink",
        lbp = "loglink", ia = NULL, iap = NULL, ib = NULL,
        ibp = NULL, independent = FALSE, zero = NULL)
```

Arguments

- la, lap, lb, lbp Link functions applied to the (positive) parameters α , α' , β and β' , respectively (the “p” stands for “prime”). See [Links](#) for more choices.
- ia, iap, ib, ibp Initial value for the four parameters respectively. The default is to estimate them all internally.
- independent Logical. If TRUE then the parameters are constrained to satisfy $\alpha = \alpha'$ and $\beta = \beta'$, which implies that y_1 and y_2 are independent and each have an ordinary exponential distribution.
- zero A vector specifying which linear/additive predictors are modelled as intercepts only. The values can be from the set {1,2,3,4}. The default is none of them. See [CommonVGAMffArguments](#) for more information.

Details

This model represents one type of bivariate extension of the exponential distribution that is applicable to certain problems, in particular, to two-component systems which can function if one of the components has failed. For example, engine failures in two-engine planes, paired organs such as peoples' eyes, ears and kidneys. Suppose y_1 and y_2 are random variables representing the lifetimes of two components A and B in a two component system. The dependence between y_1 and y_2 is essentially such that the failure of the B component changes the parameter of the exponential life distribution of the A component from α to α' , while the failure of the A component changes the parameter of the exponential life distribution of the B component from β to β' .

The joint probability density function is given by

$$f(y_1, y_2) = \alpha\beta' \exp(-\beta'y_2 - (\alpha + \beta - \beta')y_1)$$

for $0 < y_1 < y_2$, and

$$f(y_1, y_2) = \beta\alpha' \exp(-\alpha'y_1 - (\alpha + \beta - \alpha')y_2)$$

for $0 < y_2 < y_1$. Here, all four parameters are positive, as well as the responses y_1 and y_2 . Under this model, the probability that component A is the first to fail is $\alpha/(\alpha + \beta)$. The time to the first

failure is distributed as an exponential distribution with rate $\alpha + \beta$. Furthermore, the distribution of the time from first failure to failure of the other component is a mixture of $\text{Exponential}(\alpha')$ and $\text{Exponential}(\beta')$ with proportions $\beta/(\alpha + \beta)$ and $\alpha/(\alpha + \beta)$ respectively.

The marginal distributions are, in general, not exponential. By default, the linear/additive predictors are $\eta_1 = \log(\alpha)$, $\eta_2 = \log(\alpha')$, $\eta_3 = \log(\beta)$, $\eta_4 = \log(\beta')$.

A special case is when $\alpha = \alpha'$ and $\beta = \beta'$, which means that y_1 and y_2 are independent, and both have an ordinary exponential distribution with means $1/\alpha$ and $1/\beta$ respectively.

Fisher scoring is used, and the initial values correspond to the MLEs of an intercept model. Consequently, convergence may take only one iteration.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Note

To estimate all four parameters, it is necessary to have some data where $y_1 < y_2$ and $y_2 < y_1$.

The response must be a two-column matrix, with columns y_1 and y_2 . Currently, the fitted value is a matrix with two columns; the first column has values $(\alpha' + \beta)/(\alpha'(\alpha + \beta))$ for the mean of y_1 , while the second column has values $(\beta' + \alpha)/(\beta'(\alpha + \beta))$ for the mean of y_2 . The variance of y_1 is

$$\frac{(\alpha')^2 + 2\alpha\beta + \beta^2}{(\alpha')^2(\alpha + \beta)^2},$$

the variance of y_2 is

$$\frac{(\beta')^2 + 2\alpha\beta + \alpha^2}{(\beta')^2(\alpha + \beta)^2},$$

the covariance of y_1 and y_2 is

$$\frac{\alpha'\beta' - \alpha\beta}{\alpha'\beta'(\alpha + \beta)^2}.$$

Author(s)

T. W. Yee

References

Freund, J. E. (1961). A bivariate extension of the exponential distribution. *Journal of the American Statistical Association*, **56**, 971–977.

See Also

[exponential](#).

Examples

```

fdata <- data.frame(y1 = rexp(nn <- 1000, rate = exp(1)))
fdata <- transform(fdata, y2 = rexp(nn, rate = exp(2)))
fit1 <- vglm(cbind(y1, y2) ~ 1, freund61, fdata, trace = TRUE)
coef(fit1, matrix = TRUE)
Coef(fit1)
vcov(fit1)
head(fitted(fit1))
summary(fit1)

# y1 and y2 are independent, so fit an independence model
fit2 <- vglm(cbind(y1, y2) ~ 1, freund61(indep = TRUE),
            data = fdata, trace = TRUE)
coef(fit2, matrix = TRUE)
constraints(fit2)
pchisq(2 * (logLik(fit1) - logLik(fit2)), # p-value
      df = df.residual(fit2) - df.residual(fit1),
      lower.tail = FALSE)
lrtest(fit1, fit2) # Better alternative

```

Gaitdbinom

Generally Altered, Inflated, Truncated and Deflated Binomial Distribution

Description

Density, distribution function, quantile function and random generation for the generally altered, inflated, truncated and deflated binomial distribution. Both parametric and nonparametric variants are supported; these are based on finite mixtures of the parent with itself and the multinomial logit model (MLM) respectively.

Usage

```

dgaitdbinom(x, size.p, prob.p,
            a.mix = NULL, a.mlm = NULL,
            i.mix = NULL, i.mlm = NULL,
            d.mix = NULL, d.mlm = NULL, truncate = NULL,
            pobs.mix = 0, pobs.mlm = 0,
            pstr.mix = 0, pstr.mlm = 0,
            pdip.mix = 0, pdip.mlm = 0, byrow.aid = FALSE,
            size.a = size.p, size.i = size.p, size.d = size.p,
            prob.a = prob.p, prob.i = prob.p, prob.d = prob.p,
            log = FALSE, ...)
pgaitdbinom(q, size.p, prob.p,
            a.mix = NULL, a.mlm = NULL,
            i.mix = NULL, i.mlm = NULL,
            d.mix = NULL, d.mlm = NULL, truncate = NULL,
            pobs.mix = 0, pobs.mlm = 0,

```

```

    pstr.mix = 0, pstr.mlm = 0,
    pdip.mix = 0, pdip.mlm = 0, byrow.aid = FALSE,
    size.a = size.p, size.i = size.p, size.d = size.p,
    prob.a = prob.p, prob.i = prob.p, prob.d = prob.p,
    lower.tail = TRUE, ...)
qgaitdbinom(p, size.p, prob.p,
    a.mix = NULL, a.mlm = NULL,
    i.mix = NULL, i.mlm = NULL,
    d.mix = NULL, d.mlm = NULL, truncate = NULL,
    pobs.mix = 0, pobs.mlm = 0,
    pstr.mix = 0, pstr.mlm = 0,
    pdip.mix = 0, pdip.mlm = 0, byrow.aid = FALSE,
    size.a = size.p, size.i = size.p, size.d = size.p,
    prob.a = prob.p, prob.i = prob.p, prob.d = prob.p, ...)
rgaitdbinom(n, size.p, prob.p,
    a.mix = NULL, a.mlm = NULL,
    i.mix = NULL, i.mlm = NULL,
    d.mix = NULL, d.mlm = NULL, truncate = NULL,
    pobs.mix = 0, pobs.mlm = 0,
    pstr.mix = 0, pstr.mlm = 0,
    pdip.mix = 0, pdip.mlm = 0, byrow.aid = FALSE,
    size.a = size.p, size.i = size.p, size.d = size.p,
    prob.a = prob.p, prob.i = prob.p, prob.d = prob.p, ...)

```

Arguments

`x, q, p, n, log, lower.tail`
 Same meaning as in [Binomial](#).

`size.p, prob.p` Same meaning as in [Binomial](#). See [Gaitdpois](#) for generic information.

`size.a, prob.a` See [Gaitdpois](#) for generic information.

`size.i, prob.i` See [Gaitdpois](#) for generic information.

`size.d, prob.d` See [Gaitdpois](#) for generic information.

`truncate` See [Gaitdpois](#) for generic information.

`a.mix, i.mix, d.mix`
 See [Gaitdpois](#) for generic information.

`a.mlm, i.mlm, d.mlm`
 See [Gaitdpois](#) for generic information.

`pstr.mix, pstr.mlm, byrow.aid`
 See [Gaitdpois](#) for generic information.

`pobs.mix, pobs.mlm`
 See [Gaitdpois](#) for generic information.

`pdip.mix, pdip.mlm`
 See [Gaitdpois](#) for generic information.

`...` Arguments such as `max.support` that are ignored. This will occur internally within [dgaitdplot](#).

Details

These functions for the GAITD binomial distribution are analogous to the GAITD Poisson, hence most details have been put in [Gaitdpois](#).

Value

`dgaitdbinom` gives the density, `pgaitdbinom` gives the distribution function, `qgaitdbinom` gives the quantile function, and `rgaitdbinom` generates random deviates. The default values of the arguments correspond to ordinary [dbinom](#), [pbinom](#), [qbinom](#), [rbinom](#) respectively.

Warning

See [Gaitdpois](#) about the dangers of too much inflation and/or deflation on GAITD PMFs, and the difficulties detecting such.

Note

Functions [Posbinom](#) have been moved to **VGAMdata**. It is better to use `dgaitdbinom(x, size, prob, truncate = 0)` instead of `dposbinom(x, size, prob)`, etc.

Author(s)

T. W. Yee.

See Also

[Gaitdpois](#), [Gaitdntbinom](#), [multinomial](#), [Gaitdlog](#), [Gaitdzeta](#).

Examples

```
size <- 20
ivec <- c(6, 10); avec <- c(8, 11); prob <- 0.25; xgrid <- 0:25
tvec <- 14; pobs.a <- 0.05; pstr.i <- 0.15
dvec <- 5; pdip.mlm <- 0.05
(ddd <- dgaitdbinom(xgrid, size, prob.p = prob,
  prob.a = prob + 0.05, truncate = tvec, pobs.mix = pobs.a,
  pdip.mlm = pdip.mlm, d.mlm = dvec,
  pobs.mlm = pobs.a, a.mlm = avec,
  pstr.mix = pstr.i, i.mix = ivec))
## Not run: dgaitdplot(c(size, prob), ylab = "Probability",
  xlab = "x", pobs.mix = pobs.mix,
  pobs.mlm = pobs.a, a.mlm = avec, all.lwd = 3,
  pdip.mlm = pdip.mlm, d.mlm = dvec, fam = "binom",
  pstr.mix = pstr.i, i.mix = ivec, deflation = TRUE,
  main = "GAITD Combo PMF---Binomial Parent")
## End(Not run)
```

Gaitdlog

*Generally Altered, Inflated, Truncated and Deflated Logarithmic Distribution***Description**

Density, distribution function, quantile function and random generation for the generally altered, inflated, truncated and deflated logarithmic distribution. Both parametric and nonparametric variants are supported; these are based on finite mixtures of the parent with itself and the multinomial logit model (MLM) respectively.

Usage

```

dgaitdlog(x, shape.p, a.mix = NULL, a.mlm = NULL,
          i.mix = NULL, i.mlm = NULL,
          d.mix = NULL, d.mlm = NULL, truncate = NULL,
          max.support = Inf, pobs.mix = 0, pobs.mlm = 0,
          pstr.mix = 0, pstr.mlm = 0,
          pdip.mix = 0, pdip.mlm = 0, byrow.aid = FALSE,
          shape.a = shape.p, shape.i = shape.p, shape.d = shape.p,
          log = FALSE)
pgaitdlog(q, shape.p, a.mix = NULL, a.mlm = NULL,
          i.mix = NULL, i.mlm = NULL,
          d.mix = NULL, d.mlm = NULL, truncate = NULL,
          max.support = Inf, pobs.mix = 0, pobs.mlm = 0,
          pstr.mix = 0, pstr.mlm = 0,
          pdip.mix = 0, pdip.mlm = 0, byrow.aid = FALSE,
          shape.a = shape.p, shape.i = shape.p, shape.d = shape.p,
          lower.tail = TRUE)
qgaitdlog(p, shape.p, a.mix = NULL, a.mlm = NULL,
          i.mix = NULL, i.mlm = NULL,
          d.mix = NULL, d.mlm = NULL, truncate = NULL,
          max.support = Inf, pobs.mix = 0, pobs.mlm = 0,
          pstr.mix = 0, pstr.mlm = 0,
          pdip.mix = 0, pdip.mlm = 0, byrow.aid = FALSE,
          shape.a = shape.p, shape.i = shape.p, shape.d = shape.p)
rgaitdlog(n, shape.p, a.mix = NULL, a.mlm = NULL,
          i.mix = NULL, i.mlm = NULL,
          d.mix = NULL, d.mlm = NULL, truncate = NULL,
          max.support = Inf, pobs.mix = 0, pobs.mlm = 0,
          pstr.mix = 0, pstr.mlm = 0,
          pdip.mix = 0, pdip.mlm = 0, byrow.aid = FALSE,
          shape.a = shape.p, shape.i = shape.p, shape.d = shape.p)

```

Arguments

x, q, p, n, log, lower.tail

Same meaning as in [dlog](#).

shape.p, shape.a, shape.i, shape.d
 Same meaning as shape for [dlog](#), i.e., for an ordinary logarithmic distribution.
 See [Gaitdpois](#) for generic information.

truncate, max.support
 See [Gaitdpois](#) for generic information.

a.mix, i.mix, d.mix
 See [Gaitdpois](#) for generic information.

a.mlm, i.mlm, d.mlm
 See [Gaitdpois](#) for generic information.

pobs.mlm, pstr.mlm, pdip.mlm, byrow.aid
 See [Gaitdpois](#) for generic information.

pobs.mix, pstr.mix, pdip.mix
 See [Gaitdpois](#) for generic information.

Details

These functions for the logarithmic distribution are analogous to the Poisson, hence most details have been put in [Gaitdpois](#). These functions do what [Oalog](#), [Oilog](#), [Otlog](#) collectively did plus much more.

Value

[dgaitdlog](#) gives the density, [pgaitdlog](#) gives the distribution function, [qgaitdlog](#) gives the quantile function, and [rgaitdlog](#) generates random deviates. The default values of the arguments correspond to ordinary [dlog](#), [plog](#), [qlog](#), [rlog](#) respectively.

Warning

See [Gaitdpois](#) about the dangers of too much inflation and/or deflation on GAITD PMFs, and the difficulties detecting such.

Note

See [Gaitdpois](#) for general information also relevant to this parent distribution.

Author(s)

T. W. Yee.

See Also

[gaitdlog](#), [Gaitdpois](#), [dgaitdplot](#), [Gaitdzeta](#), [multinomial](#), [Oalog](#), [Oilog](#), [Otlog](#).

Examples

```
ivec <- c(2, 10); avec <- ivec + 1; shape <- 0.995; xgrid <- 0:15
max.support <- 15; pobs.a <- 0.10; pstr.i <- 0.15
dvec <- 1; pdip.mlm <- 0.05
(ddd <- dgaitdlog(xgrid, shape,
  max.support = max.support, pobs.mix = pobs.a,
```

```

pdip.mlm = pdip.mlm, d.mlm = dvec,
a.mix = avec, pstr.mix = pstr.i, i.mix = ivec))
## Not run: dgaitdplot(shape, ylab = "Probability", xlab = "x",
max.support = max.support, pobs.mix = 0,
pobs.mlm = 0, a.mlm = avec, all.lwd = 3,
pdip.mlm = pdip.mlm, d.mlm = dvec, fam = "log",
pstr.mix = pstr.i, i.mix = ivec, deflation = TRUE,
main = "GAITD Combo PMF---Logarithmic Parent")
## End(Not run)

```

gaitdlog

Generally Altered, Inflated, Truncated and Deflated Logarithmic Regression

Description

Fits a generally altered, inflated, truncated and deflated logarithmic regression by MLE. The GAITD combo model having 7 types of special values is implemented. This allows logarithmic mixtures on nested and/or partitioned support as well as a multinomial logit model for altered, inflated and deflated values. Truncation may include the upper tail.

Usage

```

gaitdlog(a.mix = NULL, i.mix = NULL, d.mix = NULL,
a.mlm = NULL, i.mlm = NULL, d.mlm = NULL,
truncate = NULL, max.support = Inf,
zero = c("pobs", "pstr", "pdip"), eq.ap = TRUE, eq.ip = TRUE,
eq.dp = TRUE, parallel.a = FALSE,
parallel.i = FALSE, parallel.d = FALSE,
lshape.p = "logitlink", lshape.a = lshape.p,
lshape.i = lshape.p, lshape.d = lshape.p,
type.fitted = c("mean", "shapes", "pobs.mlm", "pstr.mlm",
"pdip.mlm", "pobs.mix", "pstr.mix", "pdip.mix", "Pobs.mix",
"Pstr.mix", "Pdip.mix", "nonspecial",
"Numer", "Denom.p", "sum.mlm.i", "sum.mix.i", "sum.mlm.d",
"sum.mix.d", "ptrunc.p", "cdf.max.s"),
gshape.p = -expm1(-7 * ppoints(12)), gpstr.mix = ppoints(7) / 3,
gpstr.mlm = ppoints(7) / (3 + length(i.mlm)),
imethod = 1, mux.init = c(0.75, 0.5, 0.75),
ishape.p = NULL, ishape.a = ishape.p,
ishape.i = ishape.p, ishape.d = ishape.p,
ipobs.mix = NULL, ipstr.mix = NULL, ipdip.mix = NULL,
ipobs.mlm = NULL, ipstr.mlm = NULL, ipdip.mlm = NULL,
byrow.aid = FALSE, ishrinkage = 0.95, probs.y = 0.35)

```

Arguments

- truncate, max.support
See [gaitdpoisson](#).
- a.mix, i.mix, d.mix
See [gaitdpoisson](#).
- a.mlm, i.mlm, d.mlm
See [gaitdpoisson](#).
- lshape.p, lshape.a, lshape.i, lshape.d
Link functions. See [gaitdpoisson](#) and [Links](#) for more choices and information. Actually, it is usually a good idea to set these arguments equal to [logffmlink](#) because the log-mean is the first linear/additive predictor so it is like a Poisson regression.
- eq.ap, eq.ip, eq.dp
Single logical each. See [gaitdpoisson](#).
- parallel.a, parallel.i, parallel.d
Single logical each. See [gaitdpoisson](#).
- type.fitted, mux.init
See [gaitdpoisson](#).
- imethod, ipobs.mix, ipstr.mix, ipdip.mix
See [CommonVGAMffArguments](#) and [gaitdpoisson](#) for information.
- ipobs.mlm, ipstr.mlm, ipdip.mlm, byrow.aid
See [CommonVGAMffArguments](#) and [gaitdpoisson](#) for information.
- gpstr.mix, gpstr.mlm
See [CommonVGAMffArguments](#) and [gaitdpoisson](#) for information.
- gshape.p, ishape.p
See [CommonVGAMffArguments](#) and [gaitdpoisson](#) for information. The former argument is used only if the latter is not given. Practical experience has shown that good initial values are needed, so if convergence is not obtained then try a finer grid.
- ishape.a, ishape.i, ishape.d
See [CommonVGAMffArguments](#) and [gaitdpoisson](#) for information.
- probs.y, ishrinkage
See [CommonVGAMffArguments](#) and [gaitdpoisson](#) for information.
- zero
See [gaitdpoisson](#) and [CommonVGAMffArguments](#) for information.

Details

Many details to this family function can be found in [gaitdpoisson](#) because it is also a 1-parameter discrete distribution. This function currently does not handle multiple responses. Further details are at [Gaitdlog](#).

As alluded to above, when there are covariates it is much more interpretable to model the mean rather than the shape parameter. Hence [logffmlink](#) is recommended. (This might become the default in the future.) So installing **VGAMextra** is a good idea.

Apart from the order of the linear/additive predictors, the following are (or should be) equivalent: [gaitdlog\(\)](#) and [logff\(\)](#), [gaitdlog\(a.mix = 1\)](#) and [oalog\(zero = "pobs1"\)](#), [gaitdlog\(i.mix = 1\)](#) and [oilog\(zero = "pstr1"\)](#), [gaitdlog\(truncate = 1\)](#) and [otlog\(\)](#). The functions [oalog](#), [oilog](#) and [otlog](#) have been placed in **VGAMdata**.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), [rrvglm](#) and [vgam](#).

Warning

See [gaitdpoisson](#).

Note

See [gaitdpoisson](#).

Author(s)

T. W. Yee

See Also

[Gaitdlog](#), [logff](#), [logffMlink](#), [Gaitdpois](#), [gaitdpoisson](#), [gaitdzeta](#), [spikeplot](#), [goffset](#), [Trunc](#), [oalog](#), [oilog](#), [otlog](#), [CommonVGAMffArguments](#), [rootogram4](#), [simulate.vlm](#).

Examples

```
## Not run: avec <- c(5, 10) # Alter these values parametrically
ivec <- c(3, 15) # Inflate these values
tvec <- c(6, 7) # Truncate these values
max.support <- 20; set.seed(1)
pobs.a <- pstr.i <- 0.1
gdata <- data.frame(x2 = runif(nn <- 1000))
gdata <- transform(gdata, shape.p = logitlink(2+0.5*x2, inverse = TRUE))
gdata <- transform(gdata,
  y1 = rgaitdlog(nn, shape.p, a.mix = avec, pobs.mix = pobs.a,
    i.mix = ivec, pstr.mix = pstr.i, truncate = tvec,
    max.support = max.support))
gaitdlog(a.mix = avec, i.mix = ivec, max.support = max.support)
with(gdata, table(y1))
spikeplot(with(gdata, y1), las = 1)
fit7 <- vglm(y1 ~ x2, trace = TRUE, data = gdata,
  gaitdlog(i.mix = ivec, truncate = tvec,
    max.support = max.support, a.mix = avec,
    eq.ap = TRUE, eq.ip = TRUE))
head(fitted(fit7, type.fitted = "Pstr.mix"))
head(predict(fit7))
t(coef(fit7, matrix = TRUE)) # Easier to see with t()
summary(fit7)
spikeplot(with(gdata, y1), lwd = 2, ylim = c(0, 0.4))
plotdgaitd(fit7, new.plot = FALSE, offset.x = 0.2, all.lwd = 2)
## End(Not run)
```

Gaitdnbinom	<i>Generally Altered, Inflated, Truncated and Deflated Negative Binomial Distribution</i>
-------------	---

Description

Density, distribution function, quantile function and random generation for the generally altered, inflated, truncated and deflated negative binomial (GAITD-NB) distribution. Both parametric and nonparametric variants are supported; these are based on finite mixtures of the parent with itself and the multinomial logit model (MLM) respectively.

Usage

```
dgaitdnbinom(x, size.p, munb.p,
             a.mix = NULL, a.mlm = NULL,
             i.mix = NULL, i.mlm = NULL,
             d.mix = NULL, d.mlm = NULL, truncate = NULL,
             max.support = Inf, pobs.mix = 0, pobs.mlm = 0,
             pstr.mix = 0, pstr.mlm = 0,
             pdip.mix = 0, pdip.mlm = 0, byrow.aid = FALSE,
             size.a = size.p, size.i = size.p, size.d = size.p,
             munb.a = munb.p, munb.i = munb.p, munb.d = munb.p,
             log = FALSE)
pgaitdnbinom(q, size.p, munb.p,
             a.mix = NULL, a.mlm = NULL,
             i.mix = NULL, i.mlm = NULL,
             d.mix = NULL, d.mlm = NULL, truncate = NULL,
             max.support = Inf, pobs.mix = 0, pobs.mlm = 0,
             pstr.mix = 0, pstr.mlm = 0,
             pdip.mix = 0, pdip.mlm = 0, byrow.aid = FALSE,
             size.a = size.p, size.i = size.p, size.d = size.p,
             munb.a = munb.p, munb.i = munb.p, munb.d = munb.p,
             lower.tail = TRUE)
qgaitdnbinom(p, size.p, munb.p,
             a.mix = NULL, a.mlm = NULL,
             i.mix = NULL, i.mlm = NULL,
             d.mix = NULL, d.mlm = NULL, truncate = NULL,
             max.support = Inf, pobs.mix = 0, pobs.mlm = 0,
             pstr.mix = 0, pstr.mlm = 0,
             pdip.mix = 0, pdip.mlm = 0, byrow.aid = FALSE,
             size.a = size.p, size.i = size.p, size.d = size.p,
             munb.a = munb.p, munb.i = munb.p, munb.d = munb.p)
rgaitdnbinom(n, size.p, munb.p,
             a.mix = NULL, a.mlm = NULL,
             i.mix = NULL, i.mlm = NULL,
             d.mix = NULL, d.mlm = NULL, truncate = NULL,
             max.support = Inf, pobs.mix = 0, pobs.mlm = 0,
```

```
pstr.mix = 0, pstr.mlm = 0,
pdip.mix = 0, pdip.mlm = 0, byrow.aid = FALSE,
size.a = size.p, size.i = size.p, size.d = size.p,
munb.a = munb.p, munb.i = munb.p, munb.d = munb.p)
```

Arguments

`x`, `q`, `p`, `n`, `log`, `lower.tail`
 Same meaning as in [rnbinom](#).

`size.p`, `munb.p` Same meaning as in [rnbinom](#). See [Gaitdpois](#) for generic information.

`size.a`, `munb.a` See [Gaitdpois](#) for generic information.

`size.i`, `munb.i` See [Gaitdpois](#) for generic information.

`size.d`, `munb.d` See [Gaitdpois](#) for generic information.

`truncate`, `max.support`
 See [Gaitdpois](#) for generic information.

`a.mix`, `i.mix`, `d.mix`
 See [Gaitdpois](#) for generic information.

`a.mlm`, `i.mlm`, `d.mlm`
 See [Gaitdpois](#) for generic information.

`pobs.mlm`, `pstr.mlm`, `byrow.aid`
 See [Gaitdpois](#) for generic information.

`pobs.mix`, `pstr.mix`
 See [Gaitdpois](#) for generic information.

`pdip.mix`, `pdip.mlm`
 See [Gaitdpois](#) for generic information.

Details

These functions for the NBD are analogous to the Poisson, hence most details have been put in [Gaitdpois](#). The NBD has two possible parameterizations: one involving a probability (argument begins with `prob`) and the other the mean (beginning with `mu`). Only the latter is supported here.

For now, arguments such as `prob.p` and `prob.a` are no longer supported. That's because `mu` is more likely to be used by most statisticians than `prob`; see [dnbinom](#).

Value

`dgaitdnbinom` gives the density, `pgaitdnbinom` gives the distribution function, `qgaitdnbinom` gives the quantile function, and `rgaitdnbinom` generates random deviates. The default values of the arguments correspond to ordinary [dnbinom](#), [pnbinom](#), [qnbinom](#), [rnbinom](#) respectively.

Warning

See [Gaitdpois](#) about the dangers of too much inflation and/or deflation on GAITD PMFs, and the difficulties detecting such.

Note

Four functions were moved from **VGAM** to **VGAMdata**; they can be seen at [Posnegbin](#). It is preferable to use `dgaitdnbinom(x, size, munb.p = munb, truncate = 0)` instead of `dposnbinom(x, size, munb = munb)`, etc.

Author(s)

T. W. Yee.

See Also

[gaitdnbinomial](#), [Gaitdpois](#), [multinomial](#), [Gaitdbinom](#), [Gaitdlog](#), [Gaitdzeta](#).

Examples

```
size <- 10; xgrid <- 0:25
ivec <- c(5, 6, 10, 14); avec <- c(8, 11); munb <- 10
tvec <- 15; pobs.a <- 0.05; pstr.i <- 0.25
dvec <- 13; pdip.mlm <- 0.03; pobs.mlm <- 0.05
(ddd <- dgaitdnbinom(xgrid, size, munb.p = munb, munb.a = munb + 5,
  truncate = tvec, pobs.mix = pobs.a,
  pdip.mlm = pdip.mlm, d.mlm = dvec,
  pobs.mlm = pobs.a, a.mlm = avec,
  pstr.mix = pstr.i, i.mix = ivec))
## Not run: dgaitdplot(c(size, munb), fam = "nbinom",
  ylab = "Probability", xlab = "x", xlim = c(0, 25),
  truncate = tvec, pobs.mix = pobs.a,
  pobs.mlm = pobs.mlm, a.mlm = avec, all.lwd = 3,
  pdip.mlm = pdip.mlm, d.mlm = dvec,
  pstr.mix = pstr.i, i.mix = ivec, deflation = TRUE,
  main = "GAITD Combo PMF---NB Parent")
## End(Not run)
```

gaitdnbinomial

Generally Altered, Inflated, Truncated and Deflated Negative Binomial Regression

Description

Fits a generally altered, inflated truncated and deflated negative binomial regression by MLE. The GAITD combo model having 7 types of special values is implemented. This allows mixtures of negative binomial distributions on nested and/or partitioned support as well as a multinomial logit model for (nonparametric) altered, inflated and deflated values.

Usage

```

gaitdnbinomial(a.mix = NULL, i.mix = NULL, d.mix = NULL,
  a.mlm = NULL, i.mlm = NULL, d.mlm = NULL,
  truncate = NULL, zero = c("size", "pobs", "pstr", "pdip"),
  eq.ap = TRUE, eq.ip = TRUE, eq.dp = TRUE,
  parallel.a = FALSE, parallel.i = FALSE, parallel.d = FALSE,
  lmunb.p = "loglink",
  lmunb.a = lmunb.p, lmunb.i = lmunb.p, lmunb.d = lmunb.p,
  lsize.p = "loglink",
  lsize.a = lsize.p, lsize.i = lsize.p, lsize.d = lsize.p,
  type.fitted = c("mean", "munbs", "sizes", "pobs.mlm",
    "pstr.mlm", "pdip.mlm", "pobs.mix", "pstr.mix", "pdip.mix",
    "Pobs.mix", "Pstr.mix", "Pdip.mix", "nonspecial", "Numer",
    "Denom.p", "sum.mlm.i", "sum.mix.i",
    "sum.mlm.d", "sum.mix.d", "ptrunc.p", "cdf.max.s"),
  gpstr.mix = ppoints(7) / 3,
  gpstr.mlm = ppoints(7) / (3 + length(i.mlm)),
  imethod = 1, mux.init = c(0.75, 0.5, 0.75, 0.5),
  imunb.p = NULL, imunb.a = imunb.p,
  imunb.i = imunb.p, imunb.d = imunb.p,
  isize.p = NULL, isize.a = isize.p,
  isize.i = isize.p, isize.d = isize.p,
  ipobs.mix = NULL, ipstr.mix = NULL,
  ipdip.mix = NULL, ipobs.mlm = NULL,
  ipstr.mlm = NULL, ipdip.mlm = NULL,
  byrow.aid = FALSE, ishrinkage = 0.95, probs.y = 0.35,
  nsimEIM = 500, cutoff.prob = 0.999, eps.trig = 1e-7,
  nbd.max.support = 4000, max.chunk.MB = 30)

```

Arguments

truncate See [gaitdpoisson](#).

a.mix, i.mix, d.mix See [gaitdpoisson](#).

a.mlm, i.mlm, d.mlm See [gaitdpoisson](#).

lmunb.p, lmunb.a, lmunb.i, lmunb.d
 Link functions pertaining to the mean parameters. See [gaitdpoisson](#) where llambda.p etc. are the equivalent.

lsize.p, lsize.a, lsize.i, lsize.d
 Link functions pertaining to the size parameters. See [NegBinomial](#).

eq.ap, eq.ip, eq.dp
 See [gaitdpoisson](#). These apply to both munb and size parameters simultaneously. See [NegBinomial](#) also.

parallel.a, parallel.i, parallel.d
 See [gaitdpoisson](#).

type.fitted See [gaitdpoisson](#).

gpstr.mix, gpstr.mlm	See gaitdpoisson .
imethod, ipobs.mix, ipstr.mix, ipdip.mix	See gaitdpoisson and CommonVGAMffArguments .
ipobs.mlm, ipstr.mlm, ipdip.mlm	See gaitdpoisson .
mux.init	Numeric, of length 4. General downward multiplier for initial values for the sample proportions (MLEs actually). See gaitdpoisson . The fourth value corresponds to size.
imunb.p, imunb.a, imunb.i, imunb.d	See gaitdpoisson ; imunb.p is similar to ilambda.p, etc.
isize.p, isize.a, isize.i, isize.d	See gaitdpoisson ; isize.p is similar to ilambda.p, etc.
probs.y, ishrinkage	See CommonVGAMffArguments for information.
byrow.aid	Details are at Gaitdpois .
zero	See gaitdpoisson and CommonVGAMffArguments .
nsimEIM, cutoff.prob, eps.trig	See negbinomial .
nbd.max.support, max.chunk.MB	See negbinomial .

Details

The GAITD–NB combo model is the pinnacle of GAITD regression for counts because it potentially handles underdispersion, equidispersion and overdispersion relative to the Poisson, as well as alteration, inflation, deflation and truncation at arbitrary support points. In contrast, [gaitdpoisson](#) cannot handle overdispersion so well. The GAITD–NB is so flexible that it can accommodate up to seven modes.

The full GAITD–NB–NB–MLM–NB–MLM–NB–MLM combo model may be fitted with this family function. There are seven types of special values and all arguments for these may be used in a single model. Here, the MLM represents the nonparametric while the NB refers to the negative binomial mixtures. The defaults for this function correspond to an ordinary negative binomial regression so that [negbinomial](#) is called instead.

While much of the documentation here draws upon [gaitdpoisson](#), there are additional details here because the NBD is a *two* parameter distribution that handles *overdispersion* relative to the Poisson. Consequently, this family function is exceeding flexible and there are many more pitfalls to avoid.

The order of the linear/additive predictors is best explained by an example. Suppose a combo model has $\text{length}(a.\text{mix}) > 3$ and $\text{length}(i.\text{mix}) > 3$, $\text{length}(d.\text{mix}) > 3$, $a.\text{mlm} = 3:5$, $i.\text{mlm} = 6:9$ and $d.\text{mlm} = 10:12$, say. Then $\text{loglink}(\text{munb.p})$ and $\text{loglink}(\text{size.p})$ are the first two. The third is $\text{multilogitlink}(\text{pobs.mix})$ followed by $\text{loglink}(\text{munb.a})$ and $\text{loglink}(\text{size.a})$ because $a.\text{mix}$ is long enough. The sixth is $\text{multilogitlink}(\text{pstr.mix})$ followed by $\text{loglink}(\text{munb.i})$ and $\text{loglink}(\text{size.i})$ because $i.\text{mix}$ is long enough. The ninth is $\text{multilogitlink}(\text{pdip.mix})$ followed by $\text{loglink}(\text{munb.d})$ and $\text{loglink}(\text{size.d})$ because $d.\text{mix}$ is long enough. Next are the probabilities for the $a.\text{mlm}$ values. Then are the probabilities for the $i.\text{mlm}$ values. Lastly are the probabilities for the $d.\text{mlm}$ values. All the probabilities are estimated by one big MLM

and effectively the "(Others)" column of left over probabilities is associated with the nonspecial values. These might be called the *nonspecial baseline probabilities* (NBP) or reserve probabilities. The dimension of the vector of linear/additive predictors here is $M = 21$.

Apart from the order of the linear/additive predictors, the following are (or should be) equivalent: `gaitdnbinomial()` and `negbinomial()`, `gaitdnbinomial(a.mix = 0)` and `zanegbinomial(zero = "pobs0")`, `gaitdnbinomial(i.mix = 0)` and `zinegbinomial(zero = "pstr0")`, `gaitdnbinomial(truncate = 0)` and `posnegbinomial()`. Likewise, if `a.mix` and `i.mix` are assigned a scalar then it effectively moves that scalar to `a.mlm` and `i.mlm` because there is no parameters such as `munb.i` being estimated. Thus `gaitdnbinomial(a.mix = 0)` and `gaitdnbinomial(a.mlm = 0)` are the effectively same, and ditto for `gaitdnbinomial(i.mix = 0)` and `gaitdnbinomial(i.mlm = 0)`.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as `vglm`, `rrvglm` and `vgam`.

The `fitted.values` slot of the fitted object, which should be extracted by the generic function `fitted`, returns the mean μ by default. See the information above on `type.fitted`.

Warning

See [gaitdpoisson](#). Also, having `eq.ap = TRUE`, `eq.ip = TRUE` and `eq.dp = TRUE` is often needed to obtain initial values that are good enough because they borrow strength across the different operators. It is usually easy to relax these assumptions later.

This family function is under constant development and future changes will occur.

Note

If `length(a.mix)` is 1 then effectively this becomes a value of `a.mlm`. If `length(a.mix)` is 2 then an error message will be issued (overfitting really). If `length(a.mix)` is 3 then this is almost overfitting too. Hence `length(a.mix)` should be 4 or more. Ditto for `length(i.mix)` and `length(d.mix)`.

See [gaitdpoisson](#) for notes about numerical problems that can easily arise. With the NBD there is even more potential trouble that can occur. In particular, good initial values are more necessary so it pays to experiment with arguments such as `imunb.p` and `isize.p`, as well as fitting an intercept-only model first before adding covariates and using `etastart`.

Currently `max.support` is missing because only `Inf` is handled. This might change later.

Author(s)

T. W. Yee

References

Yee, T. W. and Ma, C. (2024). Generally altered, inflated, truncated and deflated regression. *Statistical Science*, **39** (in press).

See Also

[Gaitdnbinom](#), [dgaitdplot](#), [multinomial](#), [rootogram4](#), [specials](#), [plotdgaitd](#), [spikeplot](#), [meangaitd](#), [KLD](#), [gaitdpoisson](#), [gaitdlog](#), [gaitdzeta](#), [multilogitlink](#), [multinomial](#), [goffset](#), [Trunc](#), [negbinomial](#), [CommonVGAMffArguments](#), [simulate.vlm](#).

Examples

```
## Not run:
i.mix <- c(5, 10, 12, 16) # Inflate these values parametrically
i.mlm <- c(14, 15) # Inflate these values
a.mix <- c(1, 6, 13, 20) # Alter these values
tvec <- c(3, 11) # Truncate these values
pstr.mlm <- 0.1 # So parallel.i = TRUE
pobs.mix <- pstr.mix <- 0.1; set.seed(1)
gdata <- data.frame(x2 = runif(nn <- 1000))
gdata <- transform(gdata, munb.p = exp(2 + 0.0 * x2),
                  size.p = exp(1))
gdata <- transform(gdata,
                  y1 = rgaitdnbinom(nn, size.p, munb.p, a.mix = a.mix,
                                   i.mix = i.mix,
                                   pobs.mix = pobs.mix, pstr.mix = pstr.mix,
                                   i.mlm = i.mlm, pstr.mlm = pstr.mlm,
                                   truncate = tvec))
gaitdnbinomial(a.mix = a.mix, i.mix = i.mix, i.mlm = i.mlm)
with(gdata, table(y1))
fit1 <- vglm(y1 ~ 1, crit = "coef", trace = TRUE, data = gdata,
            gaitdnbinomial(a.mix = a.mix, i.mix = i.mix,
                           i.mlm = i.mlm,
                           parallel.i = TRUE, eq.ap = TRUE,
                           eq.ip = TRUE, truncate = tvec))
head(fitted(fit1, type.fitted = "Pstr.mix"))
head(predict(fit1))
t(coef(fit1, matrix = TRUE)) # Easier to see with t()
summary(fit1)
spikeplot(with(gdata, y1), lwd = 2)
plotdgaitd(fit1, new.plot = FALSE, offset.x = 0.2, all.lwd = 2)
## End(Not run)
```

Gaitdpois

Generally Altered, Inflated, Truncated and Deflated Poisson Distribution

Description

Density, distribution function, quantile function and random generation for the generally altered, inflated, truncated and deflated Poisson distribution. Both parametric and nonparametric variants are supported; these are based on finite mixtures of the parent with itself and the multinomial logit model (MLM) respectively.

Usage

```

dgaitdpois(x, lambda.p, a.mix = NULL, a.mlm = NULL, i.mix = NULL,
           i.mlm = NULL, d.mix = NULL, d.mlm = NULL, truncate = NULL,
           max.support = Inf, pobs.mix = 0, pobs.mlm = 0, pstr.mix = 0,
           pstr.mlm = 0, pdip.mix = 0, pdip.mlm = 0, byrow.aid = FALSE,
           lambda.a = lambda.p, lambda.i = lambda.p,
           lambda.d = lambda.p, log = FALSE)
pgaitdpois(q, lambda.p, a.mix = NULL, a.mlm = NULL, i.mix = NULL,
           i.mlm = NULL, d.mix = NULL, d.mlm = NULL, truncate = NULL,
           max.support = Inf, pobs.mix = 0, pobs.mlm = 0, pstr.mix = 0,
           pstr.mlm = 0, pdip.mix = 0, pdip.mlm = 0, byrow.aid = FALSE,
           lambda.a = lambda.p, lambda.i = lambda.p,
           lambda.d = lambda.p, lower.tail = TRUE, checkd = FALSE)
qgaitdpois(p, lambda.p, a.mix = NULL, a.mlm = NULL, i.mix = NULL,
           i.mlm = NULL, d.mix = NULL, d.mlm = NULL, truncate = NULL,
           max.support = Inf, pobs.mix = 0, pobs.mlm = 0, pstr.mix = 0,
           pstr.mlm = 0, pdip.mix = 0, pdip.mlm = 0, byrow.aid = FALSE,
           lambda.a = lambda.p, lambda.i = lambda.p, lambda.d = lambda.p)
rgaitdpois(n, lambda.p, a.mix = NULL, a.mlm = NULL, i.mix = NULL,
           i.mlm = NULL, d.mix = NULL, d.mlm = NULL, truncate = NULL,
           max.support = Inf, pobs.mix = 0, pobs.mlm = 0, pstr.mix = 0,
           pstr.mlm = 0, pdip.mix = 0, pdip.mlm = 0, byrow.aid = FALSE,
           lambda.a = lambda.p, lambda.i = lambda.p, lambda.d = lambda.p)

```

Arguments

`x`, `q`, `p`, `n` Same meaning as in [Poisson](#).

`log`, `lower.tail` Same meaning as in [Poisson](#).

`lambda.p`, `lambda.a`, `lambda.i`, `lambda.d`
 Same meaning as in [Poisson](#), i.e., for an ordinary Poisson distribution. The first is for the main *p*arent (or base) distribution. The next two concern the parametric variant and these distributions (usually spikes) may be *altered* and/or *inflated*. The last one concerns the *deflated* variant. Short vectors are recycled.

`truncate`, `max.support`
 numeric; these specify the set of truncated values. The default value of `NULL` means an empty set for the former. The latter is the maximum support value so that any value larger has been truncated (necessary because `truncate = (max.support + 1) : Inf` is not allowed), hence is needed for truncating the upper tail of the distribution. Note that `max(truncate) < max.support` must be satisfied otherwise an error message will be issued.

`a.mix`, `i.mix`, `d.mix`
 Vectors of nonnegative integers; the altered, inflated and deflated values for the parametric variant. Each argument must have unique values only. Assigning argument `a.mix` means that `pobs.mix` will be used. Assigning `i.mix` means that `pstr.mix` will be used. Assigning `d.mix` means that `pdip.mix` will be used. If `a.mix` is of unit length then the default probability mass function (PMF) evaluated at `a.mix` will be `pobs.mix`. So having `a.mix = 0` corresponds to the zero-inflated Poisson distribution (see [Zipois](#)).

<code>a.mlm, i.mlm, d.mlm</code>	Similar to the above, but for the nonparametric (MLM) variant. For example, assigning <code>a.mlm</code> means that <code>pobs.mlm</code> will be used. Collectively, the above 7 arguments represent 7 disjoint sets of special values and they are a proper subset of the support of the distribution.
<code>pobs.mlm, pstr.mlm, pdip.mlm, byrow.aid</code>	The first three arguments are coerced into a matrix of probabilities using <code>byrow.aid</code> to determine the order of the elements (similar to <code>byrow</code> in <code>matrix</code> , and the <code>.aid</code> reinforces the behaviour that it applies to both altered, inflated and deflated cases). The first argument is recycled if necessary to become $n \times \text{length}(a.mlm)$. The second argument becomes $n \times \text{length}(i.mlm)$. The third argument becomes $n \times \text{length}(d.mlm)$. Thus these arguments are not used unless <code>a.mlm</code> , <code>i.mlm</code> and <code>d.mlm</code> are assigned. For deflated models, <code>pdip.mix</code> and <code>pdip.mlm</code> are positive-valued and VGAM will subtract these quantities; the argument deflation has been deprecated.
<code>pobs.mix, pstr.mix, pdip.mix</code>	Vectors of probabilities that are recycled if necessary to length n . The first argument is used when <code>a.mix</code> is not NULL. The second argument is used when <code>i.mix</code> is not NULL. The third argument is used when <code>d.mix</code> is not NULL.
<code>checkd</code>	Logical. If TRUE then the density is computed at <code>floor(q)</code> with the same parameters. This can help detect whether the PMF is invalid. If so, then NaNs are returned. See Example 2 below.

Details

These functions allow any combination of 4 operator types: truncation, alteration, inflation and deflation. The precedence is truncation, then alteration and lastly inflation and deflation. Informally, deflation can be thought of as the opposite of inflation. This order minimizes the potential interference among the operators. Loosely, a set of probabilities is set to 0 by truncation and the remaining probabilities are scaled up. Then a different set of probabilities are set to some values `pobs.mix` and/or `pobs.mlm` and the remaining probabilities are rescaled up. Then another different set of probabilities is inflated by an amount `pstr.mlm` and/or proportional to `pstr.mix` so that individual elements in this set have two sources. Then another different set of probabilities is deflated by an amount `pdip.mlm` and/or proportional to `pdip.mix`. Then all the probabilities are rescaled so that they sum to unity.

Both parametric and nonparametric variants are implemented. They usually have arguments with suffix `.mix` and `.mlm` respectively. The MLM is a loose coupling that effectively separates the *parent* (or *base*) distribution from the altered values. Values inflated nonparametrically effectively have their spikes shaved off. The `.mix` variant has associated with it `lambda.a` and `lambda.i` and `lambda.d` because it is mixture of 4 Poisson distributions with partitioned or nested support.

Any value of the support of the distribution that is altered, inflated, truncated or deflated is called a *special* value. A special value that is altered may mean that its probability increases or decreases relative to the parent distribution. An inflated special value means that its probability has increased, provided alteration elsewhere has not made it decrease in the first case. There are seven types of special values and they are represented by `a.mix`, `a.mlm`, `i.mix`, `i.mlm`, `d.mix`, `d.mlm`, `truncate`.

Terminology-wise, *special* values are altered or inflated or truncated or deflated, and the remaining support points that correspond directly to the parent distribution are *nonspecial* or ordinary. These functions do what `Zapois`, `Zipois`, `Pospois` collectively did plus much more.

In the notation of Yee and Ma (2023) these functions allow for the special cases: (i) GAIT-Pois($\lambda.p$)–Pois($\lambda.a, a.mix, pobs.mix$)–Pois($\lambda.i, i.mix, pstr.mix$); (ii) GAIT-Pois($\lambda.p$)–MLM($a.mlm, pobs.mlm$)–MLM($i.mlm, pstr.mlm$). Model (i) is totally parametric while model (ii) is the most nonparametric possible.

Value

`dgaitdpois` gives the density, `pgaitdpois` gives the distribution function, `qgaitdpois` gives the quantile function, and `rgaitdpois` generates random deviates. The default values of the arguments correspond to ordinary `dpois`, `ppois`, `qpois`, `rpois` respectively.

Warning

It is possible that the GAITD PMF is invalid because of too much inflation and/or deflation. This would result in some probabilities exceeding unity or being negative. Hence `x` should ideally contain these types of special values so that this can be detected. If so then a NaN is returned and a warning is issued, e.g., same as `dnorm(0, 0, sd = -1)`. To help checking, `pgaitdpois(q, ...)` calls `dgaitdpois(floor(q), ...)` if `checkd` is TRUE.

That is, given the parameters, it is impractical to determine whether the PMF is valid. To do this, one would have to compute the PMF at all values of its support and check that they are nonnegative and sum to unity. Hence one must be careful to input values from the parameter space, especially for inflation and deflation. See Example 2 below.

Note

Functions `Pospois` and those similar have been moved to **VGAMdata**. It is better to use `dgaitdpois(x, lambda, truncate = 0)` instead of `dposbinom(x, lambda)`, etc.

Author(s)

T. W. Yee.

References

Yee, T. W. and Ma, C. (2024). Generally altered, inflated, truncated and deflated regression. *Statistical Science*, **39** (in press).

See Also

`gaitdpoisson`, `multinomial`, `specials`, `spikeplot`, `dgaitdplot`, `Zapois`, `Zipois`, `Pospois`, `Poisson`; `Gaitdbinom`, `Gaitdnbinom`, `Gaitdlog`, `Gaitdzeta`.

Examples

```
# Example 1
ivec <- c(6, 14); avec <- c(8, 11); lambda <- 10; xgrid <- 0:25
tvec <- 15; max.support <- 20; pobs.mix <- 0.05; pstr.i <- 0.25
dvec <- 13; pdip.mlm <- 0.05; pobs.mlm <- 0.05
(ddd <- dgaitdpois(xgrid, lambda, lambda.a = lambda + 5,
  truncate = tvec, max.support = max.support, pobs.mix = pobs.mix,
```

```

pobs.mlm = pobs.mlm, a.mlm = avec,
pdip.mlm = pdip.mlm, d.mlm = dvec,
pstr.mix = pstr.i, i.mix = ivec))
## Not run: dgaitdplot(lambda, ylab = "Probability", xlab = "x",
truncate = tvec, max.support = max.support, pobs.mix = pobs.mix,
pobs.mlm = pobs.mlm, a.mlm = avec, all.lwd = 3,
pdip.mlm = pdip.mlm, d.mlm = dvec,
pstr.mix = pstr.i, i.mix = ivec, deflation = TRUE,
main = "GAITD Combo PMF---Poisson Parent")
## End(Not run)

# Example 2: detection of an invalid PMF
xgrid <- 1:3 # Does not cover the special values purposely
(ddd <- dgaitdpois(xgrid, 1, pdip.mlm = 0.1, d.mlm = 5,
pstr.mix = 0.95, i.mix = 0)) # Undetected
xgrid <- 0:13 # Wider range so this detects the problem
(ddd <- dgaitdpois(xgrid, 1, pdip.mlm = 0.1, d.mlm = 5,
pstr.mix = 0.95, i.mix = 0)) # Detected
sum(ddd, na.rm = TRUE) # Something gone awry

```

gaitdpoisson

Generally Altered, Inflated, Truncated and Deflated Poisson Regression

Description

Fits a generally altered, inflated, truncated and deflated Poisson regression by MLE. The GAITD combo model having 7 types of special values is implemented. This allows mixtures of Poissons on nested and/or partitioned support as well as a multinomial logit model for (nonparametric) altered, inflated and deflated values. Truncation may include the upper tail.

Usage

```

gaitdpoisson(a.mix = NULL, i.mix = NULL, d.mix = NULL,
a.mlm = NULL, i.mlm = NULL, d.mlm = NULL,
truncate = NULL, max.support = Inf,
zero = c("pobs", "pstr", "pdip"),
eq.ap = TRUE, eq.ip = TRUE, eq.dp = TRUE,
parallel.a = FALSE, parallel.i = FALSE, parallel.d = FALSE,
llambda.p = "loglink", llambda.a = llambda.p,
llambda.i = llambda.p, llambda.d = llambda.p,
type.fitted = c("mean", "lambdas", "pobs.mlm", "pstr.mlm",
"pdip.mlm", "pobs.mix", "pstr.mix", "pdip.mix",
"Pobs.mix", "Pstr.mix", "Pdip.mix", "nonspecial",
"Numer", "Denom.p", "sum.mlm.i", "sum.mix.i",
"sum.mlm.d", "sum.mix.d", "ptrunc.p",
"cdf.max.s"), gpstr.mix = ppoints(7) / 3,
gpstr.mlm = ppoints(7) / (3 + length(i.mlm)),
imethod = 1, mux.init = c(0.75, 0.5, 0.75),

```

```

ilambda.p = NULL, ilambda.a = ilambda.p,
ilambda.i = ilambda.p, ilambda.d = ilambda.p,
ipobs.mix = NULL, ipstr.mix = NULL, ipdip.mix = NULL,
ipobs.mlm = NULL, ipstr.mlm = NULL, ipdip.mlm = NULL,
byrow.aid = FALSE, ishrinkage = 0.95, probs.y = 0.35)

```

Arguments

`truncate, max.support`

Vector of truncated values, i.e., nonnegative integers. For the first seven arguments (for the *special* values) a NULL stands for an empty set, and the seven sets must be mutually disjoint. Argument `max.support` enables RHS-truncation, i.e., something equivalent to `truncate = (U+1):Inf` for some upper support point `U` specified by `max.support`.

`a.mix, i.mix, d.mix`

Vector of altered and inflated values corresponding to finite mixture models. These are described as *parametric* or structured.

The parameter `lambda.p` is always estimated. If `length(a.mix)` is 1 or more then the parameter `pobs.mix` is estimated. If `length(i.mix)` is 1 or more then the parameter `pstr.mix` is estimated. If `length(d.mix)` is 1 or more then the parameter `pdip.mix` is estimated.

If `length(a.mix)` is 2 or more then the parameter `lambda.a` is estimated. If `length(i.mix)` is 2 or more then the parameter `lambda.i` is estimated. If `length(d.mix)` is 2 or more then the parameter `lambda.d` is estimated.

If `length(a.mix) == 1, length(i.mix) == 1 or length(d.mix) == 1` then `lambda.a, lambda.i` and `lambda.d` are unidentifiable and therefore ignored. In such cases it would be equivalent to moving `a.mix` into `a.mlm`, etc.

Due to its great flexibility, it is easy to misuse this function and ideally the values of the above arguments should be well justified by the application on hand. Adding inappropriate or unnecessary values to these arguments willy-nilly is a recipe for disaster, especially for `i.mix` and `d.mix`. Using `a.mlm` effectively removes a subset of the data from the main analysis, therefore may result in a substantial loss of efficiency. For seeped values, `a.mix, a.mlm, d.mix` and `d.mlm` can be used only. Heaped values can be handled by `i.mlm` and `i.mix`, as well as `a.mix` and `a.mlm`. Because of the NBP reason below, it sometimes may be necessary to specify deflated values to altered values.

`a.mlm, i.mlm, d.mlm`

Vector of altered, inflated and deflated values corresponding to the multinomial logit model (MLM) probabilities of observing those values—see [multinomial](#). These are described as *nonparametric* or unstructured.

`llambda.p, llambda.a, llambda.i, llambda.d`

Link functions for the parent, altered, inflated and deflated distributions respectively. See [Links](#) for more choices and information.

`eq.ap, eq.ip, eq.dp`

Single logical each. Constrain the rate parameters to be equal? See [CommonVGAMffArguments](#) for information. Having all three arguments TRUE gives greater stability in the estimation because of fewer parameters and therefore fewer initial values needed, however if so then one should try relax some of the arguments later.

For the GIT-Pois submodel, after plotting the responses, if the distribution of the spikes above the nominal probabilities has roughly the same shape as the ordinary values then setting `eq.ip = TRUE` would be a good idea so that `lambda.i == lambda.p`. And if `i.mix` is of length 2 or a bit more, then `TRUE` should definitely be entertained. Likewise, for heaped or seeped data, setting `eq.ap = TRUE` (so that `lambda.p == lambda.p`) would be a good idea for the GAT-Pois if the shape of the altered probabilities is roughly the same as the parent distribution.

`parallel.a`, `parallel.i`, `parallel.d`

Single logical each. Constrain the MLM probabilities to be equal? If so then this applies to all `length(a.mlm)` `pobs.mlm` probabilities or all `length(i.mlm)` `pstr.mlm` probabilities or all `length(d.mlm)` `pdip.mlm` probabilities. See [CommonVGAMffArguments](#) for information. The default means that the probabilities are generally unconstrained and unstructured and will follow the shape of the data. See [constraints](#).

`type.fitted`

See [CommonVGAMffArguments](#) and below for information. The first value is the default, and this is usually the unconditional mean. Choosing an irrelevant value may result in an NA being returned and a warning, e.g., "pstr.mlm" for a nonparametric GAT model.

The choice "lambdas" returns a matrix with at least one column and up to three others, corresponding to all those estimated. In order, their `colnames` are "lambda.p", "lambda.a", "lambda.i" and "lambda.d". For other distributions such as [gaitdlog](#) `type.fitted = "shapes"` is permitted and the `colnames` are "shape.p", "shape.a", "shape.i" and "shape.d", etc.

Option "Pobs.mix" provides more detail about "pobs.mix" by returning a matrix whose columns correspond to each altered value; the row sums (`rowSums`) of this matrix is "pobs.mix". Likewise "Pstr.mix" about "pstr.mix" and "Pdip.mix" about "pdip.mix".

The choice "cdf.max.s" is the CDF evaluated at `max.support` using the parent distribution, e.g., `ppois(max.support, lambda.p)` for [gaitdpoisson](#). The value should be 1 if `max.support = Inf` (the default). The choice "nonspecial" is the probability of a nonspecial value. The choices "Denom.p" and "Numer" are quantities found in the GAITD combo PMF and are for convenience only.

The choice `type.fitted = "pobs.mlm"` returns a matrix whose columns are the altered probabilities (Greek symbol ω_s). The choice "pstr.mlm" returns a matrix whose columns are the inflated probabilities (Greek symbol ϕ_s). The choice "pdip.mlm" returns a matrix whose columns are the deflated probabilities (Greek symbol ψ_s).

The choice "ptrunc.p" returns the probability of having a truncated value with respect to the parent distribution. It includes any truncated values in the upper tail beyond `max.support`. The probability of a value less than or equal to `max.support` with respect to the parent distribution is "cdf.max.s".

The choice "sum.mlm.i" adds two terms. This gives the probability of an inflated value, and the formula can be loosely written down as something like "pstr.mlm" + "Numer" * `dpois(i.mlm, lambda.p) / "Denom.p"`. The other three "sum.m*" arguments are similar.

`gpstr.mix`, `gpstr.mlm`

See [CommonVGAMffArguments](#) for information. Gridsearch values for the two parameters. If failure occurs try a finer grid, especially closer to 0, and/or experiment with `mux.init`.

<code>imethod</code> , <code>ipobs.mix</code> , <code>ipstr.mix</code> , <code>ipdip.mix</code>	See CommonVGAMffArguments for information. Good initial values are difficult to compute because of the great flexibility of GAITD regression, therefore it is often necessary to use these arguments. A careful examination of a spikeplot of the data should lead to good choices.
<code>ipobs.mlm</code> , <code>ipstr.mlm</code> , <code>ipdip.mlm</code>	See CommonVGAMffArguments for information.
<code>mux.init</code>	Numeric, of length 3. General downward multiplier for initial values for the sample proportions (MLEs actually). This is under development and more details are forthcoming. In general, 1 means unchanged and values should lie in (0, 1], and values about 0.5 are recommended. The elements apply in order to altered, inflated and deflated (no distinction between mix and MLM).
<code>ilambda.p</code> , <code>ilambda.a</code> , <code>ilambda.i</code> , <code>ilambda.d</code>	Initial values for the rate parameters; see CommonVGAMffArguments for information.
<code>probs.y</code> , <code>ishrinkage</code>	See CommonVGAMffArguments for information.
<code>byrow.aid</code>	Details are at Gaitdpois .
<code>zero</code>	See CommonVGAMffArguments for information. By default, all the MLM probabilities are modelled as simple as possible (intercept-only) to help avoid numerical problems, especially when there are many covariates. The Poisson means are modelled by the covariates, and the default zero vector is pruned of any irrelevant values. To model all the MLM probabilities with covariates set <code>zero = NULL</code> , however, the number of regression coefficients could be excessive. For the MLM probabilities, to model <code>pobs.mix</code> only with covariates set <code>zero = c('pstr', 'pobs.mlm', 'pdip')</code> . Likewise, to model <code>pstr.mix</code> only with covariates set <code>zero = c('pobs', 'pstr.mlm', 'pdip')</code> . It is noted that, amongst other things, zipoisson and zipoissonff differ with respect to <code>zero</code> , and ditto for zapoisson and zapoissonff .

Details

The full GAITD–Pois combo model may be fitted with this family function. There are seven types of special values and all arguments for these may be used in a single model. Here, the MLM represents the nonparametric while the Pois refers to the Poisson mixtures. The defaults for this function correspond to an ordinary Poisson regression so that [poissonff](#) is called instead. A MLM with only one probability to model is equivalent to logistic regression ([binomialff](#) and [logitlink](#)).

The order of the linear/additive predictors is best explained by an example. Suppose a combo model has `length(a.mix) > 2` and `length(i.mix) > 2`, `length(d.mix) > 2`, `a.mlm = 3:5`, `i.mlm = 6:9` and `d.mlm = 10:12`, say. Then `loglink(lambda.p)` is the first. The second is `multilogitlink(pobs.mix)` followed by `loglink(lambda.a)` because `a.mix` is long enough. The fourth is `multilogitlink(pstr.mix)` followed by `loglink(lambda.i)` because `i.mix` is long enough. The sixth is `multilogitlink(pdip.mix)` followed by `loglink(lambda.d)` because `d.mix` is long enough. Next are the probabilities for the `a.mlm` values. Then are the probabilities for the `i.mlm` values. Lastly are the probabilities for the `d.mlm` values. All the probabilities are estimated by one big MLM and effectively the "(Others)" column of left over probabilities is associated with the nonspecial values. These might be called the

nonspecial baseline probabilities (NBP). The dimension of the vector of linear/additive predictors here is $M = 17$.

Two mixture submodels that may be fitted can be abbreviated GAT–Pois or GIT–Pois. For the GAT model the distribution being fitted is a (spliced) mixture of two Poissons with differing (partitioned) support. Likewise, for the GIT model the distribution being fitted is a mixture of two Poissons with nested support. The two rate parameters may be constrained to be equal using eq. ap and eq. ip.

A good first step is to apply [spikeplot](#) for selecting candidate values for altering, inflating and deflating. Deciding between parametrically or nonparametrically can also be determined from examining the spike plot. Misspecified a.mix/a.mlm/i.mix/i.mlm/d.mix/d.mlm will result in convergence problems (setting trace = TRUE is a *very good idea*.) This function currently does not handle multiple responses. Further details are at [Gaitdpois](#).

A well-conditioned data–model combination should pose no difficulties for the automatic starting value selection being successful. Failure to obtain initial values from this self-starting family function indicates the degree of inflation/deflation may be marginal and/or a misspecified model. If this problem is worth surmounting the arguments to focus on especially are mux.init, gpstr.mix, gpstr.mlm, ipdip.mix and ipdip.mlm. See below for the stepping-stone trick.

Apart from the order of the linear/additive predictors, the following are (or should be) equivalent: gaitdpoisson() and poissonff(), gaitdpoisson(a.mix = 0) and zapoisson(zero = "pobs0"), gaitdpoisson(i.mix = 0) and zipoisson(zero = "pstr0"), gaitdpoisson(truncate = 0) and pospoisson(). Likewise, if a.mix and i.mix are assigned a scalar then it effectively moves that scalar to a.mlm and i.mlm because there is no lambda.a or lambda.i being estimated. Thus gaitdpoisson(a.mix = 0) and gaitdpoisson(a.mlm = 0) are the effectively same, and ditto for gaitdpoisson(i.mix = 0) and gaitdpoisson(i.mlm = 0).

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), [rrvglm](#) and [vgam](#).

The fitted.values slot of the fitted object, which should be extracted by the generic function fitted, returns the mean μ by default. See the information above on type.fitted.

Warning

Amateurs tend to be overzealous fitting zero-inflated models when the fitted mean is low—the warning of [ziP](#) should be heeded. For GAITD regression the warning applies more strongly and generally; here to *all* i.mix, i.mlm, d.mix and d.mlm values, not just 0. Even one misspecified special value usually will cause convergence problems.

Default values for this and similar family functions may change in the future, e.g., eq. ap and eq. ip. Important internal changes might occur too, such as the ordering of the linear/additive predictors and the quantities returned as the fitted values.

Using i.mlm requires more caution than a.mlm because gross inflation is ideally needed for it to work safely. Ditto for i.mix versus a.mix. Data exhibiting deflation or little to no inflation will produce numerical problems, hence set trace = TRUE to monitor convergence. More than c.10 IRLS iterations should raise suspicion.

Ranking the four operators by difficulty, the easiest is truncation followed by alteration, then inflation and the most difficult is deflation. The latter needs good initial values and the current default

will probably not work on some data sets. Studying the spikeplot is time very well spent. In general it is very easy to specify an *overfitting* model so it is a good idea to split the data into training and test sets.

This function is quite memory-hungry with respect to `length(c(a.mix, i.mix, d.mix, a.mlm, i.mlm, d.mlm))`. On consuming something different, because all values of the NBP vector need to be positive it pays to be economical with respect to `d.mlm` especially so that one does not consume up probabilities unnecessarily so to speak.

It is often a good idea to set `eq.ip = TRUE`, especially when `length(i.mix)` is not much more than 2 or the values of `i.mix` are not spread over the range of the response. This way the estimation can borrow strength from both the inflated and non-inflated values. If the `i.mix` values form a single small cluster then this can easily create estimation difficulties—the idea is somewhat similar to multicollinearity. The same holds for `d.mix`.

Note

Numerical problems can easily arise because of the exceeding flexibility of this distribution and/or the lack of sizeable inflation/deflation; it is a good idea to gain experience with simulated data first before applying it to real data. Numerical problems may arise if any of the special values are in remote places of the support, e.g., a value y such that `dpois(y, lambda.p)` is very close to 0. This is because the ratio of two tiny values can be unstable.

Good initial values may be difficult to obtain using self-starting procedures, especially when there are covariates. If so, then it is advisable to use a trick: fit an intercept-only model first and then use `etastart = predict(int.only.model)` to fit the model with covariates. This uses the simpler model as a stepping-stone.

The labelling of the linear/additive predictors has been abbreviated to reduce space. For example, `multilogitlink(pobs.mix)` and `multilogitlink(pstr.mix)` would be more accurately `multilogitlink(cbind(pobs.mix, pstr.mix))` because one grand MLM is fitted. This shortening may result in modifications needed in other parts of **VGAM** to compensate.

Because estimation involves a MLM, the restricted parameter space means that if the dip probabilities are large then the NBP may become too close to 0. If this is so then there are tricks to avoid a negative NBP. One of them is to model as many values of `d.mlm` as `d.mix`, hence the dip probabilities become modelled via the deflation distribution instead. Another trick to alter those special values rather than deflating them if the dip probabilities are large.

Due to its complexity, the HDE test `hdeff` is currently unavailable for GAITD regressions.

Randomized quantile residuals (RQRs) are available; see [residualsvglm](#).

Author(s)

T. W. Yee

References

Yee, T. W. and Ma, C. (2024). Generally altered, inflated, truncated and deflated regression. *Statistical Science*, **39** (in press).

See Also

[Gaitdpois](#), [multinomial](#), [rootogram4](#), [specials](#), [plotdgaitd](#), [spikeplot](#), [meangaitd](#), [KLD](#), [goffset](#), [Trunc](#), [gaitdnbinomial](#), [gaitdlog](#), [gaitdzeta](#), [multilogitlink](#), [multinomial](#), [residualsvglm](#), [poissonff](#), [zapoisson](#), [zipoisson](#), [pospoisson](#), [CommonVGAMffArguments](#), [simulate.vlm](#).

Examples

```
## Not run:
i.mix <- c(5, 10) # Inflate these values parametrically
i.mlm <- c(14, 15) # Inflate these values
a.mix <- c(1, 13) # Alter these values
tvec <- c(3, 11) # Truncate these values
pstr.mlm <- 0.1 # So parallel.i = TRUE
pobs.mix <- pstr.mix <- 0.1
max.support <- 20; set.seed(1)
gdata <- data.frame(x2 = runif(nn <- 1000))
gdata <- transform(gdata, lambda.p = exp(2 + 0.0 * x2))
gdata <- transform(gdata,
  y1 = rgaitdpois(nn, lambda.p, a.mix = a.mix, i.mix = i.mix,
    pobs.mix = pobs.mix, pstr.mix = pstr.mix,
    i.mlm = i.mlm, pstr.mlm = pstr.mlm,
    truncate = tvec, max.support = max.support))
gaitdpoisson(a.mix = a.mix, i.mix = i.mix, i.mlm = i.mlm)
with(gdata, table(y1))
fit1 <- vglm(y1 ~ 1, crit = "coef", trace = TRUE, data = gdata,
  gaitdpoisson(a.mix = a.mix, i.mix = i.mix,
    i.mlm = i.mlm, parallel.i = TRUE,
    eq.ap = TRUE, eq.ip = TRUE, truncate =
    tvec, max.support = max.support))
head(fitted(fit1, type.fitted = "Pstr.mix"))
head(predict(fit1))
t(coef(fit1, matrix = TRUE)) # Easier to see with t()
summary(fit1) # No HDE test by default but HDEtest = TRUE is ideal
spikeplot(with(gdata, y1), lwd = 2)
plotdgaitd(fit1, new.plot = FALSE, offset.x = 0.2, all.lwd = 2)

## End(Not run)
```

Gaitdzeta

Generally Altered, Inflated and Truncated and Deflated Zeta Distribution

Description

Density, distribution function, quantile function and random generation for the generally altered, inflated, truncated and deflated zeta distribution. Both parametric and nonparametric variants are supported; these are based on finite mixtures of the parent with itself and the multinomial logit model (MLM) respectively.

Usage

```

dgaitdzeta(x, shape.p, a.mix = NULL, a.mlm = NULL,
           i.mix = NULL, i.mlm = NULL,
           d.mix = NULL, d.mlm = NULL, truncate = NULL,
           max.support = Inf, pobs.mix = 0, pobs.mlm = 0,
           pstr.mix = 0, pstr.mlm = 0,
           pdip.mix = 0, pdip.mlm = 0,
           byrow.aid = FALSE,
           shape.a = shape.p, shape.i = shape.p, shape.d = shape.p,
           log = FALSE)
pgaitdzeta(q, shape.p, a.mix = NULL, a.mlm = NULL,
           i.mix = NULL, i.mlm = NULL,
           d.mix = NULL, d.mlm = NULL, truncate = NULL,
           max.support = Inf, pobs.mix = 0, pobs.mlm = 0,
           pstr.mix = 0, pstr.mlm = 0,
           pdip.mix = 0, pdip.mlm = 0,
           byrow.aid = FALSE,
           shape.a = shape.p, shape.i = shape.p, shape.d = shape.p,
           lower.tail = TRUE)
qgaitdzeta(p, shape.p, a.mix = NULL, a.mlm = NULL,
           i.mix = NULL, i.mlm = NULL,
           d.mix = NULL, d.mlm = NULL, truncate = NULL,
           max.support = Inf, pobs.mix = 0, pobs.mlm = 0,
           pstr.mix = 0, pstr.mlm = 0,
           pdip.mix = 0, pdip.mlm = 0,
           byrow.aid = FALSE,
           shape.a = shape.p, shape.i = shape.p, shape.d = shape.p)
rgaitdzeta(n, shape.p, a.mix = NULL, a.mlm = NULL,
           i.mix = NULL, i.mlm = NULL,
           d.mix = NULL, d.mlm = NULL, truncate = NULL,
           max.support = Inf, pobs.mix = 0, pobs.mlm = 0,
           pstr.mix = 0, pstr.mlm = 0,
           pdip.mix = 0, pdip.mlm = 0,
           byrow.aid = FALSE,
           shape.a = shape.p, shape.i = shape.p, shape.d = shape.p)

```

Arguments

`x`, `q`, `p`, `n`, `log`, `lower.tail`

Same meaning as in [dzeta](#).

`shape.p`, `shape.a`, `shape.i`, `shape.d`

Same meaning as `shape` for [dzeta](#), i.e., for an ordinary zeta distribution. See [Gaitdpois](#) for generic information.

`truncate`, `max.support`

See [Gaitdpois](#) for generic information.

`a.mix`, `i.mix`, `d.mix`

See [Gaitdpois](#) for generic information.

a.mlm, i.mlm, d.mlm

See [Gaitdpois](#) for generic information.

pobs.mlm, pstr.mlm, pdip.mlm, byrow.aid

See [Gaitdpois](#) for generic information.

pobs.mix, pstr.mix, pdip.mix

See [Gaitdpois](#) for generic information.

Details

These functions for the zeta distribution are analogous to the Poisson, hence most details have been put in [Gaitdpois](#). These functions do what [Oazeta](#), [Oizeta](#), [Otzeta](#) collectively did plus much more.

Value

dgaitdzeta gives the density, pgaitdzeta gives the distribution function, qgaitdzeta gives the quantile function, and rgaitdzeta generates random deviates. The default values of the arguments correspond to ordinary [dzeta](#), [pzeta](#), [qzeta](#), [rzeta](#) respectively.

Warning

See [Gaitdpois](#) about the dangers of too much inflation and/or deflation on GAITD PMFs, and the difficulties detecting such.

Note

See [Gaitdpois](#) for general information also relevant to this parent distribution.

Author(s)

T. W. Yee.

See Also

[gaitdzeta](#), [Gaitdpois](#), [dgaitdplot](#), [multinomial](#), [Oazeta](#), [Oizeta](#), [Otzeta](#).

Examples

```
ivec <- c(2, 10); avec <- ivec + 4; shape <- 0.95; xgrid <- 0:29
tvec <- 15; max.support <- 25; pobs.a <- 0.10; pstr.i <- 0.15
(ddd <- dgaitdzeta(xgrid, shape, truncate = tvec,
  max.support = max.support, pobs.mix = pobs.a,
  a.mix = avec, pstr.mix = pstr.i, i.mix = ivec))
## Not run: plot(xgrid, ddd, type = "n", ylab = "Probability",
  xlab = "x", main = "GAIT PMF---Zeta Parent")
mylwd <- 0.5
abline(v = avec, col = 'blue', lwd = mylwd)
abline(v = ivec, col = 'purple', lwd = mylwd)
abline(v = tvec, col = 'tan', lwd = mylwd)
abline(v = max.support, col = 'magenta', lwd = mylwd)
abline(h = c(pobs.a, pstr.i, 0:1), col = 'gray', lty = "dashed")
```

```

lines(xgrid, dzeta(xgrid, shape), col='gray', lty="dashed") # f_{\pi}
lines(xgrid, ddd, type = "h", col = "pink", lwd = 3) # GAIT PMF
points(xgrid[ddd == 0], ddd[ddd == 0], pch = 16, col = 'tan', cex = 2)

## End(Not run)

```

gaitdzeta

Generally Altered, Inflated, Truncated and Deflated Zeta Regression

Description

Fits a generally altered, inflated, truncated and deflated zeta regression by MLE. The GAITD combo model having 7 types of special values is implemented. This allows mixtures of zetas on nested and/or partitioned support as well as a multinomial logit model for altered, inflated and deflated values.

Usage

```

gaitdzeta(a.mix = NULL, i.mix = NULL, d.mix = NULL,
          a.mlm = NULL, i.mlm = NULL, d.mlm = NULL,
          truncate = NULL, max.support = Inf,
          zero = c("pobs", "pstr", "pdip"), eq.ap = TRUE, eq.ip = TRUE,
          eq.dp = TRUE, parallel.a = FALSE,
          parallel.i = FALSE, parallel.d = FALSE,
          lshape.p = "loglink", lshape.a = lshape.p,
          lshape.i = lshape.p, lshape.d = lshape.p,
          type.fitted = c("mean", "shapes", "pobs.mlm", "pstr.mlm",
                          "pdip.mlm", "pobs.mix", "pstr.mix", "pdip.mix", "Pobs.mix",
                          "Pstr.mix", "Pdip.mix", "nonspecial",
                          "Numer", "Denom.p", "sum.mlm.i", "sum.mix.i", "sum.mlm.d",
                          "sum.mix.d", "ptrunc.p", "cdf.max.s"),
          gshape.p = -expm1(-ppoints(7)), gpstr.mix = ppoints(7) / 3,
          gpstr.mlm = ppoints(7) / (3 + length(i.mlm)),
          imethod = 1, mux.init = c(0.75, 0.5, 0.75),
          ishape.p = NULL, ishape.a = ishape.p,
          ishape.i = ishape.p, ishape.d = ishape.p,
          ipobs.mix = NULL, ipstr.mix = NULL, ipdip.mix = NULL,
          ipobs.mlm = NULL, ipstr.mlm = NULL, ipdip.mlm = NULL,
          byrow.aid = FALSE, ishrinkage = 0.95, probs.y = 0.35)

```

Arguments

truncate, max.support

See [gaitdpoisson](#). Only max.support = Inf is allowed because some equations are intractable.

a.mix, i.mix, d.mix

See [gaitdpoisson](#).

a.mlm, i.mlm, d.mlm	See gaitdpoisson .
lshape.p, lshape.a, lshape.i, lshape.d	Link functions. See gaitdpoisson and Links for more choices and information. Actually, it is usually a good idea to set these arguments equal to zetaffMlink because the log-mean is the first linear/additive predictor so it is like a Poisson regression.
eq.ap, eq.ip, eq.dp	Single logical each. See gaitdpoisson
parallel.a, parallel.i, parallel.d	Single logical each. See gaitdpoisson .
type.fitted, mux.init	See gaitdpoisson .
imethod, ipobs.mix, ipstr.mix, ipdip.mix	See CommonVGAMffArguments and gaitdpoisson for information.
ipobs.mlm, ipstr.mlm, ipdip.mlm, byrow.aid	See CommonVGAMffArguments and gaitdpoisson for information.
gpstr.mix, gpstr.mlm	See CommonVGAMffArguments and gaitdpoisson for information.
gshape.p, ishape.p	See CommonVGAMffArguments and gaitdpoisson for information. The former is used only if the latter is not given. Practical experience has shown that good initial values are needed, so if convergence is not obtained then try a finer grid.
ishape.a, ishape.i, ishape.d	See CommonVGAMffArguments and gaitdpoisson for information.
probs.y, ishrinkage	See CommonVGAMffArguments for information.
zero	See gaitdpoisson and CommonVGAMffArguments for information.

Details

Many details to this family function can be found in [gaitdpoisson](#) because it is also a 1-parameter discrete distribution. This function currently does not handle multiple responses. Further details are at [Gaitdzeta](#).

As alluded to above, when there are covariates it is much more interpretable to model the mean rather than the shape parameter. Hence [zetaffMlink](#) is recommended. (This might become the default in the future.) So installing **VGAMextra** is a good idea.

Apart from the order of the linear/additive predictors, the following are (or should be) equivalent: `gaitdzeta()` and `zetaff()`, `gaitdzeta(a.mix = 1)` and `oazeta(zero = "pobs1")`, `gaitdzeta(i.mix = 1)` and `oizeta(zero = "pstr1")`, `gaitdzeta(truncate = 1)` and `otzeta()`. The functions `oazeta`, `oizeta` and `otzeta` have been placed in **VGAMdata**.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as `vglm`, `rrvglm` and `vgam`.

Warning

See [gaitdpoisson](#).

Note

See [gaitdpoisson](#).

Author(s)

T. W. Yee

See Also

[Gaitdzeta](#), [zetaff](#), [zetaffMlink](#), [Gaitdpois](#), [gaitdpoisson](#), [gaitdlog](#), [spikeplot](#), [goffset](#), [Trunc](#), [oazeta](#), [oizeta](#), [otzeta](#), [CommonVGAMffArguments](#), [rootogram4](#), [simulate.vlm](#).

Examples

```
## Not run:
avec <- c(5, 10) # Alter these values parametrically
ivec <- c(3, 15) # Inflate these values
tvec <- c(6, 7) # Truncate these values
set.seed(1); pobs.a <- pstr.i <- 0.1
gdata <- data.frame(x2 = runif(nn <- 1000))
gdata <- transform(gdata, shape.p = logitlink(2, inverse = TRUE))
gdata <- transform(gdata,
  y1 = rgaitdzeta(nn, shape.p, a.mix = avec, pobs.mix = pobs.a,
    i.mix = ivec, pstr.mix = pstr.i, truncate = tvec))
gaitdzeta(a.mix = avec, i.mix = ivec)
with(gdata, table(y1))
spikeplot(with(gdata, y1), las = 1)
fit7 <- vglm(y1 ~ 1, trace = TRUE, data = gdata, crit = "coef",
  gaitdzeta(i.mix = ivec, truncate = tvec,
    a.mix = avec, eq.ap = TRUE, eq.ip = TRUE))
head(fitted(fit7, type.fitted = "Pstr.mix"))
head(predict(fit7))
t(coef(fit7, matrix = TRUE)) # Easier to see with t()
summary(fit7)
spikeplot(with(gdata, y1), lwd = 2, ylim = c(0, 0.6), xlim = c(0, 20))
plotdgaitd(fit7, new.plot = FALSE, offset.x = 0.2, all.lwd = 2)

## End(Not run)
```

gamma1

1-parameter Gamma Regression Family Function

Description

Estimates the 1-parameter gamma distribution by maximum likelihood estimation.

Usage

```
gamma1(link = "loglink", zero = NULL, parallel = FALSE,
       type.fitted = c("mean", "percentiles", "Qlink"),
       percentiles = 50)
```

Arguments

`link` Link function applied to the (positive) *shape* parameter. See [Links](#) for more choices and general information.

`zero, parallel` Details at [CommonVGAMffArguments](#).

`type.fitted, percentiles` See [CommonVGAMffArguments](#) for information. Using "Qlink" is for quantile-links in **VGAMextra**.

Details

The density function is given by

$$f(y) = \exp(-y) \times y^{\text{shape}-1} / \Gamma(\text{shape})$$

for $\text{shape} > 0$ and $y > 0$. Here, $\Gamma(\text{shape})$ is the gamma function, as in [gamma](#). The mean of Y (returned as the default fitted values) is $\mu = \text{shape}$, and the variance is $\sigma^2 = \text{shape}$.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Note

This **VGAM** family function can handle a multiple responses, which is inputted as a matrix.

The parameter *shape* matches with *shape* in [rgamma](#). The argument *rate* in [rgamma](#) is assumed 1 for this family function, so that `scale = 1` is used for calls to [dgamma](#), [qgamma](#), etc.

If *rate* is unknown use the family function [gammaR](#) to estimate it too.

Author(s)

T. W. Yee

References

Most standard texts on statistical distributions describe the 1-parameter gamma distribution, e.g., Forbes, C., Evans, M., Hastings, N. and Peacock, B. (2011). *Statistical Distributions*, Hoboken, NJ, USA: John Wiley and Sons, Fourth edition.

See Also

[gammaR](#) for the 2-parameter gamma distribution, [lgamma1](#), [lindley](#), [simulate.vlm](#), [gammaff.mm](#).

Examples

```
gdata <- data.frame(y = rgamma(n = 100, shape = exp(3)))
fit <- vglm(y ~ 1, gamma1, data = gdata, trace = TRUE, crit = "coef")
coef(fit, matrix = TRUE)
Coef(fit)
summary(fit)
```

gamma2

*2-parameter Gamma Regression Family Function***Description**

Estimates the 2-parameter gamma distribution by maximum likelihood estimation.

Usage

```
gamma2(lmu = "loglink", lshape = "loglink", imethod = 1, ishape = NULL,
       parallel = FALSE, deviance.arg = FALSE, zero = "shape")
```

Arguments

lmu, lshape	Link functions applied to the (positive) <i>mu</i> and <i>shape</i> parameters (called μ and a respectively). See Links for more choices.
ishape	Optional initial value for <i>shape</i> . A NULL means a value is computed internally. If a failure to converge occurs, try using this argument. This argument is ignored if used within cqo ; see the <code>iShape</code> argument of qrrvglm.control instead.
imethod	An integer with value 1 or 2 which specifies the initialization method for the μ parameter. If failure to converge occurs try another value (and/or specify a value for <i>ishape</i>).
deviance.arg	Logical. If TRUE, the deviance function is attached to the object. Under ordinary circumstances, it should be left alone because it really assumes the shape parameter is at the maximum likelihood estimate. Consequently, one cannot use that criterion to minimize within the IRLS algorithm. It should be set TRUE only when used with cqo under the fast algorithm.
zero	See CommonVGAMffArguments for information.
parallel	Details at CommonVGAMffArguments . If <code>parallel = TRUE</code> then the constraint is not applied to the intercept.

Details

This distribution can model continuous skewed responses. The density function is given by

$$f(y; \mu, a) = \frac{\exp(-ay/\mu) \times (ay/\mu)^{a-1} \times a}{\mu \times \Gamma(a)}$$

for $\mu > 0$, $a > 0$ and $y > 0$. Here, $\Gamma(\cdot)$ is the gamma function, as in [gamma](#). The mean of Y is $\mu = \mu$ (returned as the fitted values) with variance $\sigma^2 = \mu^2/a$. If $0 < a < 1$ then the density has a

pole at the origin and decreases monotonically as y increases. If $a = 1$ then this corresponds to the exponential distribution. If $a > 1$ then the density is zero at the origin and is unimodal with mode at $y = \mu - \mu/a$; this can be achieved with `lshape="logloglink"`.

By default, the two linear/additive predictors are $\eta_1 = \log(\mu)$ and $\eta_2 = \log(a)$. This family function implements Fisher scoring and the working weight matrices are diagonal.

This **VGAM** family function handles *multivariate* responses, so that a matrix can be used as the response. The number of columns is the number of species, say, and `zero=-2` means that *all* species have a shape parameter equalling a (different) intercept only.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Note

The response must be strictly positive. A moment estimator for the shape parameter may be implemented in the future.

If `mu` and `shape` are vectors, then `rgamma(n = n, shape = shape, scale = mu/shape)` will generate random gamma variates of this parameterization, etc.; see [GammaDist](#).

Author(s)

T. W. Yee

References

The parameterization of this **VGAM** family function is the 2-parameter gamma distribution described in the monograph

McCullagh, P. and Nelder, J. A. (1989). *Generalized Linear Models*, 2nd ed. London: Chapman & Hall.

See Also

[gamma1](#) for the 1-parameter gamma distribution, [gammaR](#) for another parameterization of the 2-parameter gamma distribution that is directly matched with [rgamma](#), [bigamma.mckay](#) for a bivariate gamma distribution, [gammaff.mm](#) for another, [expexpff](#), [GammaDist](#), [CommonVGAMffArguments](#), [simulate.vlm](#), [negloglink](#).

Examples

```
# Essentially a 1-parameter gamma
gdata <- data.frame(y = rgamma(n = 100, shape = exp(1)))
fit1 <- vglm(y ~ 1, gamma1, data = gdata)
fit2 <- vglm(y ~ 1, gamma2, data = gdata, trace = TRUE, crit = "coef")
coef(fit2, matrix = TRUE)
c(Coef(fit2), colMeans(gdata))

# Essentially a 2-parameter gamma
```

```

gdata <- data.frame(y = rgamma(n = 500, rate = exp(-1), shape = exp(2)))
fit2 <- vglm(y ~ 1, gamma2, data = gdata, trace = TRUE, crit = "coef")
coef(fit2, matrix = TRUE)
c(Coef(fit2), colMeans(gdata))
summary(fit2)

```

gammaff.mm

Multivariate Gamma Family Function: Mathai and Moschopoulos (1992)

Description

Estimate the scale parameter and shape parameters of the Mathai and Moschopoulos (1992) multivariate gamma distribution by maximum likelihood estimation.

Usage

```

gammaff.mm(lscale = "loglink", lshape = "loglink",
           iscale = NULL, ishape = NULL, imethod = 1,
           eq.shapes = FALSE, sh.byrow = TRUE, zero = "shape")

```

Arguments

- `lscale, lshape` Link functions applied to the (positive) parameters b , and s_1, \dots, s_Q respectively. See [Links](#) for more choices. In the future, `lshapes` might be used instead; of course, this link applies to all the shape parameters.
- `iscale, ishape, sh.byrow` Optional initial values. The default is to compute them internally. Argument `sh.byrow` is fed into `byrow` in `matrix` and concerns the ordering of the initial shape parameters; a matrix of dimension n by Q is ultimately constructed. See also [CommonVGAMffArguments](#).
- `eq.shapes` Logical. Constrain the shape parameters to be equal? See also [CommonVGAMffArguments](#).
- `imethod, zero` See [CommonVGAMffArguments](#).

Details

This distribution has the bivariate gamma distribution `bigamma.mckay` as a special case. Let $Q > 1$ be the number of columns of the response matrix y . Then the joint probability density function is given by

$$f(y_1, \dots, y_Q; b, s_1, \dots, s_Q) = y_1^{s_1} (y_2 - y_1)^{s_2} \cdots (y_Q - y_{Q-1})^{s_Q} \exp(-y_Q/b) / [b^{s_Q} \Gamma(s_1) \cdots \Gamma(s_Q)]$$

for $b > 0$, $s_1 > 0$, \dots , $s_Q > 0$ and $0 < y_1 < y_2 < \cdots < y_Q < \infty$. Also, $s_Q^* = s_1 + \cdots + s_Q$. Here, Γ is the [gamma](#) function. By default, the linear/additive predictors are $\eta_1 = \log(b)$, $\eta_2 = \log(s_1)$, \dots , $\eta_M = \log(s_Q)$. Hence $Q = M - 1$. The marginal distributions are gamma, with shape parameters s_1 up to s_Q , but they have a common scale parameter b .

The fitted value returned is a matrix with columns equalling their respective means; for column j it is $\text{sum}(\text{shape}[1:j]) * \text{scale}$.

The correlations are always positive; for columns j and k with $j < k$, the correlation is $\sqrt{\text{sum}(\text{shape}[1:j]) / \text{sum}(\text{shape}[1:k])}$). Hence the variance of column j is $\text{sum}(\text{shape}[1:j]) * \text{scale}^2$.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Note

The response must be a matrix with at least two columns. Apart from the first column, the differences between a column and its LHS adjacent column must all be positive. That is, each row must be strictly increasing.

Author(s)

T. W. Yee

References

Mathai, A. M. and Moschopoulos, P. G. (1992). A form of multivariate gamma distribution. *Ann. Inst. Statist. Math.*, **44**, 97–106.

See Also

[bigamma.mckay](#), [gammaff](#).

Examples

```
## Not run:
data("mbflood", package = "VGAMdata")
mbflood <- transform(mbflood, VdivD = V / D)
fit <- vglm(cbind(Q, y2 = Q + VdivD) ~ 1,
            gammaff.mm, trace = TRUE, data = mbflood)
coef(fit, matrix = TRUE)
Coef(fit)
vcov(fit)
colMeans(depvar(fit)) # Check moments
head(fitted(fit), 1)

## End(Not run)
```

gammahyperbola

*Gamma Hyperbola Bivariate Distribution***Description**

Estimate the parameter of a gamma hyperbola bivariate distribution by maximum likelihood estimation.

Usage

```
gammahyperbola(ltheta = "loglink", itheta = NULL, expected = FALSE)
```

Arguments

<code>ltheta</code>	Link function applied to the (positive) parameter θ . See Links for more choices.
<code>itheta</code>	Initial value for the parameter. The default is to estimate it internally.
<code>expected</code>	Logical. FALSE means the Newton-Raphson (using the observed information matrix) algorithm, otherwise the expected information matrix is used (Fisher scoring algorithm).

Details

The joint probability density function is given by

$$f(y_1, y_2) = \exp(-e^{-\theta} y_1 / \theta - \theta y_2)$$

for $\theta > 0$, $y_1 > 0$, $y_2 > 1$. The random variables Y_1 and Y_2 are independent. The marginal distribution of Y_1 is an exponential distribution with rate parameter $\exp(-\theta)/\theta$. The marginal distribution of Y_2 is an exponential distribution that has been shifted to the right by 1 and with rate parameter θ . The fitted values are stored in a two-column matrix with the marginal means, which are $\theta \exp(\theta)$ and $1 + 1/\theta$.

The default algorithm is Newton-Raphson because Fisher scoring tends to be much slower for this distribution.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Note

The response must be a two-column matrix.

Author(s)

T. W. Yee

References

Reid, N. (2003). Asymptotics and the theory of inference. *Annals of Statistics*, **31**, 1695–1731.

See Also

[exponential](#).

Examples

```
gdata <- data.frame(x2 = runif(nn <- 1000))
gdata <- transform(gdata, theta = exp(-2 + x2))
gdata <- transform(gdata, y1 = rexp(nn, rate = exp(-theta)/theta),
                  y2 = rexp(nn, rate = theta) + 1)
fit <- vglm(cbind(y1, y2) ~ x2, gammahyperbola(expected = TRUE), data = gdata)
coef(fit, matrix = TRUE)
Coef(fit)
head(fitted(fit))
summary(fit)
```

gammaR

2-parameter Gamma Regression Family Function

Description

Estimates the 2-parameter gamma distribution by maximum likelihood estimation.

Usage

```
gammaR(lrate = "loglink", lshape = "loglink", irate = NULL,
       ishape = NULL, lss = TRUE, zero = "shape")
```

Arguments

`lrate`, `lshape` Link functions applied to the (positive) *rate* and *shape* parameters. See [Links](#) for more choices.

`irate`, `ishape` Optional initial values for *rate* and *shape*. A NULL means a value is computed internally. If a failure to converge occurs, try using these arguments.

`zero`, `lss` Details at [CommonVGAMffArguments](#).

Details

The density function is given by

$$f(y; rate, shape) = \exp(-rate \times y) \times y^{shape-1} \times rate^{shape} / \Gamma(shape)$$

for $shape > 0$, $rate > 0$ and $y > 0$. Here, $\Gamma(shape)$ is the gamma function, as in [gamma](#). The mean of Y is $\mu = shape/rate$ (returned as the fitted values) with variance $\sigma^2 = \mu^2/shape = shape/rate^2$. By default, the two linear/additive predictors are $\eta_1 = \log(rate)$ and $\eta_2 = \log(shape)$.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Note

The parameters *rate* and *shape* match with the arguments *rate* and *shape* of [rgamma](#). The order of the arguments agree too. Here, $scale = 1/rate$ is used, so one can use [negloglink](#). Multiple responses are handled.

If $rate = 1$ use the family function [gamma1](#) to estimate *shape*.

The reciprocal of a 2-parameter gamma random variate has an *inverse gamma* distribution. One might write a **VGAM** family function called `invgammaR()` to estimate this, but for now, just feed in the reciprocal of the response.

Author(s)

T. W. Yee

References

Most standard texts on statistical distributions describe the 2-parameter gamma distribution, e.g., Forbes, C., Evans, M., Hastings, N. and Peacock, B. (2011). *Statistical Distributions*, Hoboken, NJ, USA: John Wiley and Sons, Fourth edition.

See Also

[gamma1](#) for the 1-parameter gamma distribution, [gamma2](#) for another parameterization of the 2-parameter gamma distribution, [bigamma.mckay](#) for a bivariate gamma distribution, [gammaff.mm](#) for another, [expexpff](#), [simulate.vlm](#), [rgamma](#), [negloglink](#).

Examples

```
# Essentially a 1-parameter gamma
gdata <- data.frame(y1 = rgamma(n <- 100, shape = exp(1)))
fit1 <- vglm(y1 ~ 1, gamma1, data = gdata, trace = TRUE)
fit2 <- vglm(y1 ~ 1, gammaR, data = gdata, trace = TRUE, crit = "coef")
coef(fit2, matrix = TRUE)
Coef(fit2)

# Essentially a 2-parameter gamma
gdata <- data.frame(y2 = rgamma(n = 500, rate = exp(1), shape = exp(2)))
fit2 <- vglm(y2 ~ 1, gammaR, data = gdata, trace = TRUE, crit = "coef")
coef(fit2, matrix = TRUE)
Coef(fit2)
summary(fit2)
```

 garma

GARMA (Generalized Autoregressive Moving-Average) Models

Description

Fits GARMA models to time series data.

Usage

```
garma(link = "identitylink", p.ar.lag = 1, q.ma.lag = 0,
      coefstart = NULL, step = 1)
```

Arguments

link	Link function applied to the mean response. The default is suitable for continuous responses. The link loglink should be chosen if the data are counts. The link reciprocal can be chosen if the data are counts and the variance assumed for this is μ^2 . The links logitlink , probitlink , clogloglink , and cauchitlink are supported and suitable for binary responses. Note that when the log or logit link is chosen: for log and logit, zero values can be replaced by bvalue. See loglink and logitlink etc. for specific information about each link function.
p.ar.lag	A positive integer, the lag for the autoregressive component. Called p below.
q.ma.lag	A non-negative integer, the lag for the moving-average component. Called q below.
coefstart	Starting values for the coefficients. Assigning this argument is highly recommended. For technical reasons, the argument <code>coefstart</code> in vglm cannot be used.
step	Numeric. Step length, e.g., 0.5 means half-stepsizing.

Details

This function draws heavily on Benjamin *et al.* (1998). See also Benjamin *et al.* (2003). GARMA models extend the ARMA time series model to generalized responses in the exponential family, e.g., Poisson counts, binary responses. Currently, this function is rudimentary and can handle only certain continuous, count and binary responses only. The user must choose an appropriate link for the link argument.

The GARMA(p, q) model is defined by firstly having a response belonging to the exponential family

$$f(y_t|D_t) = \exp \left\{ \frac{y_t \theta_t - b(\theta_t)}{\phi/A_t} + c(y_t, \phi/A_t) \right\}$$

where θ_t and ϕ are the canonical and scale parameters respectively, and A_t are known prior weights. The mean $\mu_t = E(Y_t|D_t) = b'(\theta_t)$ is related to the linear predictor η_t by the link function g .

Here, $D_t = \{x_t, \dots, x_1, y_{t-1}, \dots, y_1, \mu_{t-1}, \dots, \mu_1\}$ is the previous information set. Secondly, the GARMA(p, q) model is defined by

$$g(\mu_t) = \eta_t = x_t^T \beta + \sum_{k=1}^p \phi_k (g(y_{t-k}) - x_{t-k}^T \beta) + \sum_{k=1}^q \theta_k (g(y_{t-k}) - \eta_{t-k}).$$

Parameter vectors β , ϕ and θ are estimated by maximum likelihood.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#).

Warning

This **VGAM** family function is 'non-standard' in that the model does need some coercing to get it into the VGLM framework. Special code is required to get it running. A consequence is that some methods functions may give wrong results when applied to the fitted object.

Note

This function is unpolished and is requires *lots* of improvements. In particular, initialization is *very poor*. Results appear *very* sensitive to quality of initial values. A limited amount of experience has shown that half-stepsizing is often needed for convergence, therefore choosing `crit = "coef"` is not recommended.

Overdispersion is not handled. For binomial responses it is currently best to input a vector of 1s and 0s rather than the `cbind(successes, failures)` because the initialize slot is rudimentary.

Author(s)

T. W. Yee

References

- Benjamin, M. A., Rigby, R. A. and Stasinopoulos, M. D. (1998). Fitting Non-Gaussian Time Series Models. Pages 191–196 in: *Proceedings in Computational Statistics COMPSTAT 1998* by Payne, R. and P. J. Green. Physica-Verlag.
- Benjamin, M. A., Rigby, R. A. and Stasinopoulos, M. D. (2003). Generalized Autoregressive Moving Average Models. *Journal of the American Statistical Association*, **98**: 214–223.
- Zeger, S. L. and Qaqish, B. (1988). Markov regression models for time series: a quasi-likelihood approach. *Biometrics*, **44**: 1019–1031.

Examples

```
gdata <- data.frame(interspike = c(68, 41, 82, 66, 101, 66, 57, 41, 27, 78,
59, 73, 6, 44, 72, 66, 59, 60, 39, 52,
50, 29, 30, 56, 76, 55, 73, 104, 104, 52,
25, 33, 20, 60, 47, 6, 47, 22, 35, 30,
29, 58, 24, 34, 36, 34, 6, 19, 28, 16,
```



```

36, 33, 12, 26, 36, 39, 24, 14, 28, 13,
 2, 30, 18, 17, 28, 9, 28, 20, 17, 12,
19, 18, 14, 23, 18, 22, 18, 19, 26, 27,
23, 24, 35, 22, 29, 28, 17, 30, 34, 17,
20, 49, 29, 35, 49, 25, 55, 42, 29, 16)) # See Zeger and Qaqish (1988)
gdata <- transform(gdata, spikenum = seq(interspike))
bvalue <- 0.1 # .Machine$double.xmin # Boundary value
fit <- vglm(interspike ~ 1, trace = TRUE, data = gdata,
            gamma(loglink(bvalue = bvalue),
                  p = 2, coefstart = c(4, 0.3, 0.4)))
summary(fit)
coef(fit, matrix = TRUE)
Coef(fit) # A bug here
## Not run: with(gdata, plot(interspike, ylim = c(0, 120), las = 1,
  xlab = "Spike Number", ylab = "Inter-Spike Time (ms)", col = "blue"))
with(gdata, lines(spikenum[-(1:fit@misc$plag)], fitted(fit), col = "orange"))
abline(h = mean(with(gdata, interspike)), lty = "dashed", col = "gray")
## End(Not run)

```

GenbetaII

The Generalized Beta II Distribution

Description

Density for the generalized beta II distribution with shape parameters a and p and q , and scale parameter $scale$.

Usage

```
dgenbetaII(x, scale = 1, shape1.a, shape2.p, shape3.q, log = FALSE)
```

Arguments

x vector of quantiles.
 $shape1.a, shape2.p, shape3.q$ positive shape parameters.
 $scale$ positive scale parameter.
 log Logical. If $log = TRUE$ then the logarithm of the density is returned.

Details

See [genbetaII](#), which is the **VGAM** family function for estimating the parameters by maximum likelihood estimation. Several distributions, such as the Singh-Maddala, are special case of this flexible 4-parameter distribution. The product of $shape1.a$ and $shape2.p$ determines the behaviour of the density at the origin.

Value

`dgenbetaII` gives the density.

Author(s)

T. W. Yee

References

Kleiber, C. and Kotz, S. (2003). *Statistical Size Distributions in Economics and Actuarial Sciences*, Hoboken, NJ, USA: Wiley-Interscience.

See Also

[genbetaII](#).

Examples

```
dgenbetaII(0, shape1.a = 1/4, shape2.p = 4, shape3.q = 3)
dgenbetaII(0, shape1.a = 1/4, shape2.p = 2, shape3.q = 3)
dgenbetaII(0, shape1.a = 1/4, shape2.p = 8, shape3.q = 3)
```

genbetaII

Generalized Beta Distribution of the Second Kind

Description

Maximum likelihood estimation of the 4-parameter generalized beta II distribution.

Usage

```
genbetaII(lscale = "loglink", lshape1.a = "loglink",
          lshape2.p = "loglink", lshape3.q = "loglink",
          iscale = NULL, ishape1.a = NULL,
          ishape2.p = NULL, ishape3.q = NULL, lss = TRUE,
          gscale = exp(-5:5), gshape1.a = exp(-5:5),
          gshape2.p = exp(-5:5), gshape3.q = exp(-5:5), zero = "shape")
```

Arguments

`lss` See [CommonVGAMffArguments](#) for important information.

`lshape1.a, lscale, lshape2.p, lshape3.q` Parameter link functions applied to the shape parameter a, scale parameter scale, shape parameter p, and shape parameter q. All four parameters are positive. See [Links](#) for more choices.

`iscale, ishape1.a, ishape2.p, ishape3.q` Optional initial values for the parameters. A NULL means a value is computed internally using the arguments gscale, gshape1.a, etc.

`gscale, gshape1.a, gshape2.p, gshape3.q` See [CommonVGAMffArguments](#) for information. Replaced by iscale, ishape1.a etc. if given.

`zero` The default is to set all the shape parameters to be intercept-only. See [CommonVGAMffArguments](#) for information.

Details

This distribution is most useful for unifying a substantial number of size distributions. For example, the Singh-Maddala, Dagum, Fisk (log-logistic), Lomax (Pareto type II), inverse Lomax, beta distribution of the second kind distributions are all special cases. Full details can be found in Kleiber and Kotz (2003), and Brazauskas (2002). The argument names given here are used by other families that are special cases of this family. Fisher scoring is used here and for the special cases too.

The 4-parameter generalized beta II distribution has density

$$f(y) = ay^{ap-1}/[b^{ap}B(p, q)\{1 + (y/b)^a\}^{p+q}]$$

for $a > 0$, $b > 0$, $p > 0$, $q > 0$, $y \geq 0$. Here B is the beta function, and b is the scale parameter scale, while the others are shape parameters. The mean is

$$E(Y) = b\Gamma(p + 1/a)\Gamma(q - 1/a)/(\Gamma(p)\Gamma(q))$$

provided $-ap < 1 < aq$; these are returned as the fitted values.

This family function handles multiple responses.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Warning

This distribution is very flexible and it is not generally recommended to use this family function when the sample size is small—numerical problems easily occur with small samples. Probably several hundred observations at least are needed in order to estimate the parameters with any level of confidence. Neither is the inclusion of covariates recommended at all—not unless there are several thousand observations. The mean is finite only when $-ap < 1 < aq$, and this can be easily violated by the parameter estimates for small sample sizes. Try fitting some of the special cases of this distribution (e.g., [sinmad](#), [fisk](#), etc.) first, and then possibly use those models for initial values for this distribution.

Note

The default is to use a grid search with respect to all four parameters; this is quite costly and is time consuming. If the self-starting initial values fail, try experimenting with the initial value arguments. Also, the constraint $-ap < 1 < aq$ may be violated as the iterations progress so it pays to monitor convergence, e.g., set `trace = TRUE`. Successful convergence depends on having very good initial values. This is rather difficult for this distribution so that a grid search is conducted by default. One suggestion for increasing the estimation reliability is to set `stepsize = 0.5` and `maxit = 100`; see [vglm.control](#).

Author(s)

T. W. Yee, with help from Victor Miranda.

References

- Kleiber, C. and Kotz, S. (2003). *Statistical Size Distributions in Economics and Actuarial Sciences*, Hoboken, NJ, USA: Wiley-Interscience.
- Brazauskas, V. (2002). Fisher information matrix for the Feller-Pareto distribution. *Statistics & Probability Letters*, **59**, 159–167.

See Also

[dgenbetaII](#), [betaff](#), [betaII](#), [dagum](#), [sinmad](#), [fisk](#), [lomax](#), [inv.lomax](#), [paralogistic](#), [inv.paralogistic](#), [lino](#), [CommonVGAMffArguments](#), [vglm.control](#).

Examples

```
## Not run:
gdata <- data.frame(y = rsinmad(3000, shape1 = exp(1), scale = exp(2),
                          shape3 = exp(1))) # A special case!
fit <- vglm(y ~ 1, genbetaII(lss = FALSE), data = gdata, trace = TRUE)
fit <- vglm(y ~ 1, data = gdata, trace = TRUE,
            genbetaII(ishape1.a = 3, iscale = 7, ishape3.q = 2.3))
coef(fit, matrix = TRUE)
Coef(fit)
summary(fit)

## End(Not run)
```

gengamma.stacy

Generalized Gamma distribution family function

Description

Estimation of the 3-parameter generalized gamma distribution proposed by Stacy (1962).

Usage

```
gengamma.stacy(lscale = "loglink", ld = "loglink", lk = "loglink",
              iscale = NULL, id = NULL, ik = NULL, imethod = 1,
              gscale.mux = exp((-4:4)/2), gshape1.d = exp((-5:5)/2),
              gshape2.k = exp((-5:5)/2), probs.y = 0.3, zero = c("d", "k"))
```

Arguments

- `lscale`, `ld`, `lk` Parameter link function applied to each of the positive parameters b , d and k , respectively. See [Links](#) for more choices.
- `iscale`, `id`, `ik` Initial value for b , d and k , respectively. The defaults mean an initial value is determined internally for each.
- `gscale.mux`, `gshape1.d`, `gshape2.k`
 See [CommonVGAMffArguments](#) for information. Replaced by `iscale`, `id` etc. if given.

imethod, probs.y, zero

See [CommonVGAMffArguments](#) for information.

Details

The probability density function can be written

$$f(y; b, d, k) = db^{-dk}y^{dk-1} \exp[-(y/b)^d]/\Gamma(k)$$

for scale parameter $b > 0$, and Weibull-type shape parameter $d > 0$, gamma-type shape parameter $k > 0$, and $y > 0$. The mean of Y is $b \times \Gamma(k + 1/d)/\Gamma(k)$ (returned as the fitted values), which equals bk if $d = 1$.

There are many special cases, as given in Table 1 of Stacey and Mihram (1965). In the following, the parameters are in the order b, d, k . The special cases are: Exponential $f(y; b, 1, 1)$, Gamma $f(y; b, 1, k)$, Weibull $f(y; b, d, 1)$, Chi Squared $f(y; 2, 1, a/2)$ with a degrees of freedom, Chi $f(y; \sqrt{2}, 2, a/2)$ with a degrees of freedom, Half-normal $f(y; \sqrt{2}, 2, 1/2)$, Circular normal $f(y; \sqrt{2}, 2, 1)$, Spherical normal $f(y; \sqrt{2}, 2, 3/2)$, Rayleigh $f(y; c\sqrt{2}, 2, 1)$ where $c > 0$. Also the log-normal distribution corresponds to when $k = \text{Inf}$.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Warning

Several authors have considered maximum likelihood estimation for the generalized gamma distribution and have found that the Newton-Raphson algorithm does not work very well and that the existence of solutions to the log-likelihood equations is sometimes in doubt. Although Fisher scoring is used here, it is likely that the same problems will be encountered. It appears that large samples are required, for example, the estimator of k became asymptotically normal only with 400 or more observations. It is not uncommon for maximum likelihood estimates to fail to converge even with two or three hundred observations. With covariates, even more observations are needed to increase the chances of convergence. Using covariates is not advised unless the sample size is at least a few thousand, and even if so, modelling 1 or 2 parameters as intercept-only is a very good idea (e.g., `zero = 2:3`). Monitoring convergence is also a very good idea (e.g., `set trace = TRUE`). Half-stepping is not uncommon, and if this occurs, then the results should be viewed with more suspicion.

Note

The notation used here differs from Stacy (1962) and Prentice (1974). Poor initial values may result in failure to converge so if there are covariates and there are convergence problems, try using or checking the `zero` argument (e.g., `zero = 2:3`) or the `ik` argument or the `imethod` argument, etc.

Author(s)

T. W. Yee

References

- Stacy, E. W. (1962). A generalization of the gamma distribution. *Annals of Mathematical Statistics*, **33**(3), 1187–1192.
- Stacy, E. W. and Mihram, G. A. (1965). Parameter estimation for a generalized gamma distribution. *Technometrics*, **7**, 349–358.
- Prentice, R. L. (1974). A log gamma model and its maximum likelihood estimation. *Biometrika*, **61**, 539–544.

See Also

[rgengamma.stacy](#), [gamma1](#), [gamma2](#), [prentice74](#), [simulate.vlm](#), [chisq](#), [lognormal](#), [rayleigh](#), [weibullR](#).

Examples

```
## Not run:
k <- exp(-1); Scale <- exp(1); dd <- exp(0.5); set.seed(1)
gdata <- data.frame(y = rgamma(2000, shape = k, scale = Scale))
gfit <- vglm(y ~ 1, gengamma.stacy, data = gdata, trace = TRUE)
coef(gfit, matrix = TRUE)

## End(Not run)
```

gengammaUC

Generalized Gamma Distribution

Description

Density, distribution function, quantile function and random generation for the generalized gamma distribution with scale parameter *scale*, and parameters *d* and *k*.

Usage

```
dgengamma.stacy(x, scale = 1, d, k, log = FALSE)
pgengamma.stacy(q, scale = 1, d, k,
                 lower.tail = TRUE, log.p = FALSE)
qgengamma.stacy(p, scale = 1, d, k,
                 lower.tail = TRUE, log.p = FALSE)
rgengamma.stacy(n, scale = 1, d, k)
```

Arguments

x, *q* vector of quantiles.

p vector of probabilities.

n number of observations. Same as in [runif](#).

scale the (positive) scale parameter *b*.

`d, k` the (positive) parameters d and k . Both can be thought of as shape parameters, where d is of the Weibull-type and k is of the gamma-type.

`log` Logical. If `log = TRUE` then the logarithm of the density is returned.

`lower.tail, log.p` Same meaning as in `pnorm` or `qnorm`.

Details

See `gengamma.stacy`, the **VGAM** family function for estimating the generalized gamma distribution by maximum likelihood estimation, for formulae and other details. Apart from `n`, all the above arguments may be vectors and are recycled to the appropriate length if necessary.

Value

`dgengamma.stacy` gives the density, `pgengamma.stacy` gives the distribution function, `qgengamma.stacy` gives the quantile function, and `rgengamma.stacy` generates random deviates.

Author(s)

T. W. Yee and Kai Huang

References

Stacy, E. W. and Mihram, G. A. (1965). Parameter estimation for a generalized gamma distribution. *Technometrics*, **7**, 349–358.

See Also

`gengamma.stacy`.

Examples

```
## Not run: x <- seq(0, 14, by = 0.01); d <- 1.5; Scale <- 2; k <- 6
plot(x, dgengamma.stacy(x, Scale, d = d, k = k), type = "l",
     col = "blue", ylim = 0:1,
     main = "Blue is density, orange is the CDF",
     sub = "Purple are 5,10,...,95 percentiles", las = 1, ylab = "")
abline(h = 0, col = "blue", lty = 2)
lines(qgengamma.stacy(seq(0.05, 0.95, by = 0.05), Scale, d = d, k = k),
      dgengamma.stacy(qgengamma.stacy(seq(0.05, 0.95, by = 0.05),
                                     Scale, d = d, k = k),
                    Scale, d = d, k = k), col = "purple", lty = 3, type = "h")
lines(x, pgengamma.stacy(x, Scale, d = d, k = k), col = "orange")
abline(h = 0, lty = 2)
## End(Not run)
```

Genpois0

*Generalized Poisson Distribution (Original Parameterization)***Description**

Density, distribution function, quantile function and random generation for the original parameterization of the generalized Poisson distribution.

Usage

```
dgenpois0(x, theta, lambda = 0, log = FALSE)
pgenpois0(q, theta, lambda = 0, lower.tail = TRUE)
qgenpois0(p, theta, lambda = 0)
rgenpois0(n, theta, lambda = 0, algorithm = c("qgenpois0",
      "inv", "bup", "chdn", "napp", "bran"))
```

Arguments

<code>x, q</code>	Vector of quantiles.
<code>p</code>	Vector of probabilities.
<code>n</code>	Similar to <code>runif</code> .
<code>theta, lambda</code>	See <code>genpoisson0</code> . The default value of <code>lambda</code> corresponds to an ordinary Poisson distribution. <i>Nonnegative</i> values of <code>lambda</code> are currently required.
<code>lower.tail, log</code>	Similar to <code>Poisson</code> .
<code>algorithm</code>	Character. Six choices are available, standing for the <i>qgenpois0</i> , <i>inversion</i> , <i>build-up</i> , <i>chop-down</i> , <i>normal approximation</i> and <i>branching</i> methods. The first one is the default and calls <code>qgenpois0</code> with <code>runif</code> as its first argument. The value inputted may be abbreviated, e.g., <code>alg = "n"</code> . The last 5 algorithms are a direct implementation of Demirtas (2017) and the relative performance of the algorithms are described there—however, the vectorization here may render the comments on relative speed as no longer holding.

Details

Most of the background to these functions are given in `genpoisson0`. Some warnings relevant to this distribution are given there. The complicated range of the parameter `lambda` when negative is no longer supported because the distribution is not normalized. For other GPD variants see `Genpois1`.

Value

`dgenpois0` gives the density, `pgenpois0` gives the distribution function, `qgenpois0` gives the quantile function, and `rgenpois` generates random deviates. For some of these functions such as `dgenpois0` and `pgenpois0` the value `NaN` is returned for elements not satisfying the parameter restrictions, e.g., if $\lambda > 1$. For some of these functions such as `rgenpois0` the input must not contain `NA`s or `NaN`s, etc. since the implemented algorithms are fragile.

Warning

These have not been tested thoroughly.

For `pgentpois0()` `mapply` is called with $\theta:q$ as input, hence will be very slow and memory-hungry for large values of q . Likewise `qgentpois0()` and `rgentpois0()` may suffer from the same limitations.

Note

For `rgentpois0()`: (1). "inv", "bup" and "chdn" appear similar and seem to work okay. (2). "napp" works only when θ is large, away from 0. It suffers from 0-inflation. (3). "bran" has a relatively heavy RHS tail and requires positive λ . More details can be found in Famoye (1997) and Demirtas (2017).

The function `dgenpois0` uses `lfactorial`, which equals `Inf` when x is approximately $1e306$ on many machines. So the density is returned as \emptyset in very extreme cases; see `.Machine`.

Author(s)

T. W. Yee. For `rgentpois0()` the last 5 algorithms are based on code written in H. Demirtas (2017) and vectorized by T. W. Yee; but the "bran" algorithm was rewritten from Famoye (1997).

References

- Demirtas, H. (2017). On accurate and precise generation of generalized Poisson variates. *Communications in Statistics—Simulation and Computation*, **46**, 489–499.
- Famoye, F. (1997). Generalized Poisson random variate generation. *Amer. J. Mathematical and Management Sciences*, **17**, 219–237.

See Also

`genpoisson0`, `Genpois1`, `dpois`.

Examples

```
sum(dgenpois0(0:1000, theta = 2, lambda = 0.5))
## Not run: theta <- 2; lambda <- 0.2; y <- 0:10
proby <- dgenpois0(y, theta = theta, lambda = lambda, log = FALSE)
plot(y, proby, type = "h", col = "blue", lwd = 2, ylab = "Pr(Y=y)",
     main = paste0("Y ~ GP-0(theta=", theta, ", lambda=",
                   lambda, ")"), las = 1, ylim = c(0, 0.3),
     sub = "Orange is the Poisson probability function")
lines(y + 0.1, dpois(y, theta), type = "h", lwd = 2, col = "orange")
## End(Not run)
```

Genpois1	<i>Generalized Poisson Distribution (GP-1 and GP-2 Parameterizations of the Mean)</i>
----------	---

Description

Density, distribution function, quantile function and random generation for two parameterizations (GP-1 and GP-2) of the generalized Poisson distribution of the mean.

Usage

```
dgenpois1(x, meanpar, dispind = 1, log = FALSE)
pgenpois1(q, meanpar, dispind = 1, lower.tail = TRUE)
qgenpois1(p, meanpar, dispind = 1)
rgenpois1(n, meanpar, dispind = 1)
dgenpois2(x, meanpar, disppar = 0, log = FALSE)
pgenpois2(q, meanpar, disppar = 0, lower.tail = TRUE)
qgenpois2(p, meanpar, disppar = 0)
rgenpois2(n, meanpar, disppar = 0)
```

Arguments

x, q	Vector of quantiles.
p	Vector of probabilities.
n	Similar to runif .
meanpar, dispind	The mean and dispersion index (index of dispersion), which are the two parameters for the GP-1. The mean is positive while the dispind is ≥ 1 . The default value of dispind corresponds to an ordinary Poisson distribution.
disppar	The dispersion parameter for the GP-2: $\text{disppar} \geq 0$. The default value of disppar corresponds to an ordinary Poisson distribution.
lower.tail, log	See Genpois0 .

Details

These are wrapper functions for those in [Genpois0](#). The first parameter is the mean, therefore both the GP-1 and GP-2 are recommended for regression and can be compared somewhat to [poissonff](#) and [negbinomial](#). The variance of a GP-1 is $\mu\varphi$ where $\varphi = 1/(1 - \lambda)^2$ is dispind.

The variance of a GP-2 is $\mu(1 + \alpha\mu)^2$ where $\theta = \mu/(1 + \alpha\mu)$, $\lambda = \alpha\mu/(1 + \alpha\mu)$, and α is the dispersion parameter disppar. Thus the variance is linear with respect to the mean for GP-1 while the variance is cubic with respect to the mean for GP-2.

Recall that the *index of dispersion* (also known as the *dispersion index*) is the ratio of the variance and the mean. Also, $\mu = \theta/(1 - \lambda)$ in the original formulation with variance $\theta/(1 - \lambda)^3$. The GP-1 is due to Consul and Famoye (1992). The GP-2 is due to Wang and Famoye (1997).

Value

dgenpois1 and dgenpois2 give the density, pgenpois1 and dgenpois2 give the distribution function, qgenpois1 and dgenpois2 give the quantile function, and rgenpois1 and dgenpois2 generate random deviates. See [Genpois0](#) for more information.

Warning

[Genpois0](#) has warnings that should be heeded.

Author(s)

T. W. Yee.

References

Consul, P. C. and Famoye, F. (1992). Generalized Poisson regression model. *Comm. Statist.—Theory and Meth.*, **2**, 89–109.

Wang, W. and Famoye, F. (1997). Modeling household fertility decisions with generalized Poisson regression. *J. Population Econom.*, **10**, 273–283.

See Also

[Genpois0](#).

Examples

```
sum(dgenpois1(0:1000, meanpar = 5, dispind = 2))
## Not run: dispind <- 5; meanpar <- 5; y <- 0:15
proby <- dgenpois1(y, meanpar = meanpar, dispind)
plot(y, proby, type = "h", col = "blue", lwd = 2, ylab = "P[Y=y]",
      main = paste0("Y ~ GP-1(meanpar=", meanpar, ", dispind=",
                    dispind, ")"), las = 1, ylim = c(0, 0.3),
      sub = "Orange is the Poisson probability function")
lines(y + 0.1, dpois(y, meanpar), type = "h", lwd = 2, col = "orange")
## End(Not run)
```

genpoisson0

Generalized Poisson Regression (Original Parameterization)

Description

Estimation of the two-parameter generalized Poisson distribution (original parameterization).

Usage

```
genpoisson0(ltheta = "loglink", llambda = "logitlink",
            itheta = NULL, ilambda = NULL, imethod = c(1, 1),
            ishrinkage = 0.95, glambda = ppoints(5),
            parallel = FALSE, zero = "lambda")
```

Arguments

- `ltheta, llambda` Parameter link functions for θ and λ . See [Links](#) for more choices. In theory the λ parameter is allowed to be negative to handle underdispersion, however this is no longer supported, hence $0 < \lambda < 1$. The θ parameter is positive, therefore the default is the log link.
- `itheta, ilambda` Optional initial values for λ and θ . The default is to choose values internally.
- `imethod` See [CommonVGAMffArguments](#) for information. Each value is an integer 1 or 2 or 3 which specifies the initialization method for each of the parameters. If failure to converge occurs try another value and/or else specify a value for `ilambda` and/or `itheta`. The argument is recycled to length 2, and the first value corresponds to `theta`, etc.
- `ishrinkage, zero` See [CommonVGAMffArguments](#) for information.
- `glambda, parallel` See [CommonVGAMffArguments](#) for information. Argument `glambda` is similar to `gsigma` there and is currently used only if `imethod[2] = 1`.

Details

The generalized Poisson distribution (GPD) was proposed by Consul and Jain (1973), and it has PMF

$$f(y) = \theta(\theta + \lambda y)^{y-1} \exp(-\theta - \lambda y)/y!$$

for $0 < \theta$ and $y = 0, 1, 2, \dots$. Theoretically, $\max(-1, -\theta/m) \leq \lambda \leq 1$ where $m (\geq 4)$ is the greatest positive integer satisfying $\theta + m\lambda > 0$ when $\lambda < 0$ [and then $Pr(Y = y) = 0$ for $y > m$]. However, there are problems with a negative λ such as it not being normalized, so this family function restricts λ to $(0, 1)$.

This original parameterization is called the GP-0 by **VGAM**, partly because there are two other common parameterizations called the GP-1 and GP-2 (see Yang et al. (2009)), [genpoisson1](#) and [genpoisson2](#)) that are more suitable for regression. However, `genpoisson()` has been simplified to [genpoisson0](#) by only handling positive parameters, hence only overdispersion relative to the Poisson is accommodated. Some of the reasons for this are described in Scollnik (1998), e.g., the probabilities do not sum to unity when `lambda` is negative. To simply things, **VGAM** 1.1-4 and later will only handle positive `lambda`.

An ordinary Poisson distribution corresponds to $\lambda = 0$. The mean (returned as the fitted values) is $E(Y) = \theta/(1 - \lambda)$ and the variance is $\theta/(1 - \lambda)^3$ so that the variance is proportional to the mean, just like the NB-1 and quasi-Poisson.

For more information see Consul and Famoye (2006) for a summary and Consul (1989) for more details.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Warning

Although this family function is far less fragile compared to what used to be called `genpoisson()` it is still a good idea to monitor convergence because equidispersion may result in numerical problems; try `poissonff` instead. And underdispersed data will definitely result in numerical problems and warnings; try `quasipoisson` instead.

Note

This family function replaces `genpoisson()`, and some of the major changes are: (i) the swapping of the linear predictors; (ii) the change from `rhobitlink` to `logitlink` in `llambda` to reflect the no longer handling of underdispersion; (iii) proper Fisher scoring is implemented to give improved convergence.

Notationally, and in the literature too, don't get confused because `theta` (and not `lambda`) here really matches more closely with `lambda` of `dpois`.

This family function handles multiple responses. This distribution is potentially useful for dispersion modelling. Convergence and numerical problems may occur when `lambda` becomes very close to 0 or 1.

Author(s)

T. W. Yee. Easton Huch derived the EIM and it has been implemented in the `weights` slot.

References

- Consul, P. C. and Jain, G. C. (1973). A generalization of the Poisson distribution. *Technometrics*, **15**, 791–799.
- Consul, P. C. and Famoye, F. (2006). *Lagrangian Probability Distributions*, Boston, USA: Birkhauser.
- Jorgensen, B. (1997). *The Theory of Dispersion Models*. London: Chapman & Hall.
- Consul, P. C. (1989). *Generalized Poisson Distributions: Properties and Applications*. New York, USA: Marcel Dekker.
- Yang, Z., Hardin, J. W., Addy, C. L. (2009). A score test for overdispersion in Poisson regression based on the generalized Poisson-2 model. *J. Statist. Plann. Infer.*, **139**, 1514–1521.
- Yee, T. W. (2020). On generalized Poisson regression. *In preparation*.

See Also

`Genpois0`, `genpoisson1`, `genpoisson2`, `poissonff`, `negbinomial`, `Poisson`, `quasipoisson`.

Examples

```
gdata <- data.frame(x2 = runif(nn <- 500))
gdata <- transform(gdata, y1 = rgenpois0(nn, theta = exp(2 + x2),
                                     logitlink(1, inverse = TRUE)))
gfit0 <- vglm(y1 ~ x2, genpoisson0, data = gdata, trace = TRUE)
coef(gfit0, matrix = TRUE)
summary(gfit0)
```

genpoisson1

*Generalized Poisson Regression (GP-1 Parameterization)***Description**

Estimation of the two-parameter generalized Poisson distribution (GP-1 parameterization) which has the variance as a linear function of the mean.

Usage

```
genpoisson1(lmeanpar = "loglink", ldispind = "logloglink",
            parallel = FALSE, zero = "dispind",
            vf1 = FALSE, Form2 = NULL,
            imeanpar = NULL, idispind = NULL, imethod = c(1, 1),
            ishrinkage = 0.95, gdispind = exp(1:5))
```

Arguments

lmeanpar, ldispind Parameter link functions for μ and φ . They are called the *mean parameter* and *dispersion index* respectively. See [Links](#) for more choices. In theory the φ parameter might be allowed to be less than unity to handle underdispersion but this is not supported. The mean is positive so its default is the log link. The dispersion index is > 1 so its default is the log-log link.

vf1, Form2 If `vf1 = TRUE` then `Form2` should be assigned a formula having terms comprising $\eta_2 = \log \log \varphi$. This is similar to [uninormal](#). See [CommonVGAMffArguments](#) for information.

imeanpar, idispind Optional initial values for μ and φ . The default is to choose values internally.

imethod See [CommonVGAMffArguments](#) for information. The argument is recycled to length 2, and the first value corresponds to μ , etc.

ishrinkage, zero See [CommonVGAMffArguments](#) for information.

gdispind, parallel See [CommonVGAMffArguments](#) for information. Argument `gdispind` is similar to `gsigma` there and is currently used only if `imethod[2] = 2`.

Details

This is a variant of the generalized Poisson distribution (GPD) and is similar to the GP-1 referred to by some writers such as Yang, et al. (2009). Compared to the original GP-0 (see [genpoisson0](#)) the GP-1 has $\theta = \mu/\sqrt{\varphi}$ and $\lambda = 1 - 1/\sqrt{\varphi}$ so that the variance is $\mu\varphi$. The first linear predictor by default is $\eta_1 = \log \mu$ so that the GP-1 is more suitable for regression than the GP-1.

This family function can handle only overdispersion relative to the Poisson. An ordinary Poisson distribution corresponds to $\varphi = 1$. The mean (returned as the fitted values) is $E(Y) = \mu$. For overdispersed data, this GP parameterization is a direct competitor of the NB-1 and quasi-Poisson.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Warning

See [genpoisson0](#) for warnings relevant here, e.g., it is a good idea to monitor convergence because of equidispersion and underdispersion.

Author(s)

T. W. Yee.

See Also

[Genpois1](#), [genpoisson0](#), [genpoisson2](#), [poissonff](#), [negbinomial](#), [Poisson](#), [quasipoisson](#).

Examples

```
gdata <- data.frame(x2 = runif(nn <- 500))
gdata <- transform(gdata, y1 = rgenpois1(nn, exp(2 + x2),
                                       logloglink(-1, inverse = TRUE)))
gfit1 <- vglm(y1 ~ x2, genpoisson1, gdata, trace = TRUE)
coef(gfit1, matrix = TRUE)
summary(gfit1)
```

genpoisson2

Generalized Poisson Regression (GP-2 Parameterization)

Description

Estimation of the two-parameter generalized Poisson distribution (GP-2 parameterization) which has the variance as a cubic function of the mean.

Usage

```
genpoisson2(lmeanpar = "loglink", ldisppar = "loglink",
            parallel = FALSE, zero = "disppar",
            vfl = FALSE, oparallel = FALSE,
            imeanpar = NULL, idisppar = NULL, imethod = c(1, 1),
            ishrinkage = 0.95, gdisppar = exp(1:5))
```

Arguments

- `lmeanpar`, `ldisppar`
 Parameter link functions for μ and α . They are called the *mean* and *dispersion parameters* respectively. See [Links](#) for more choices. In theory the α parameter might be allowed to be negative to handle underdispersion but this is not supported. All parameters are positive, therefore the defaults are the log link.
- `imeanpar`, `idisppar`
 Optional initial values for μ and α . The default is to choose values internally.
- `vf1`, `oparallel`
 Argument `oparallel` is similar to `parallel` but uses `rbind(1, -1)` instead. If `vf1 = TRUE` then `oparallel` should be assigned a formula having terms comprising $\eta_1 = \log \mu$, and then the other terms in the main formula are for $\eta_2 = \log \alpha$. See [CommonVGAMffArguments](#) for information.
- `imethod`
 See [CommonVGAMffArguments](#) for information. The argument is recycled to length 2, and the first value corresponds to μ , etc.
- `ishrinkage`, `zero`
 See [CommonVGAMffArguments](#) for information.
- `gdisppar`, `parallel`
 See [CommonVGAMffArguments](#) for information. Argument `gdisppar` is similar to `gsigma` there and is currently used only if `imethod[2] = 2`.

Details

This is a variant of the generalized Poisson distribution (GPD) and called GP-2 by some writers such as Yang, et al. (2009). Compared to the original GP-0 (see [genpoisson0](#)) the GP-2 has $\theta = \mu/(1 + \alpha\mu)$ and $\lambda = \alpha\mu/(1 + \alpha\mu)$ so that the variance is $\mu(1 + \alpha\mu)^2$. The first linear predictor by default is $\eta_1 = \log \mu$ so that the GP-2 is more suitable for regression than the GP-0.

This family function can handle only overdispersion relative to the Poisson. An ordinary Poisson distribution corresponds to $\alpha = 0$. The mean (returned as the fitted values) is $E(Y) = \mu$.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Warning

See [genpoisson0](#) for warnings relevant here, e.g., it is a good idea to monitor convergence because of equidispersion and underdispersion.

Author(s)

T. W. Yee.

References

Letac, G. and Mora, M. (1990). Natural real exponential families with cubic variance functions. *Annals of Statistics* **18**, 1–37.

See Also

[Genpois2](#), [genpoisson0](#), [genpoisson1](#), [poissonff](#), [negbinomial](#), [Poisson](#), [quasipoisson](#).

Examples

```
gdata <- data.frame(x2 = runif(nn <- 500))
gdata <- transform(gdata, y1 = rgenpois2(nn, exp(2 + x2),
                                       loglink(-1, inverse = TRUE)))
gfit2 <- vglm(y1 ~ x2, genpoisson2, gdata, trace = TRUE)
coef(gfit2, matrix = TRUE)
summary(gfit2)
```

genray

The Generalized Rayleigh Distribution

Description

Density, distribution function, quantile function and random generation for the generalized Rayleigh distribution.

Usage

```
dgenray(x, scale = 1, shape, log = FALSE)
pgenray(q, scale = 1, shape, lower.tail = TRUE, log.p = FALSE)
qgenray(p, scale = 1, shape, lower.tail = TRUE, log.p = FALSE)
rgenray(n, scale = 1, shape)
```

Arguments

x, q	vector of quantiles.
p	vector of probabilities.
n	number of observations. If <code>length(n) > 1</code> then the length is taken to be the number required.
scale, shape	positive scale and shape parameters.
log	Logical. If <code>log = TRUE</code> then the logarithm of the density is returned.
lower.tail, log.p	Same meaning as in pnorm or qnorm .

Details

See [genrayleigh](#), the **VGAM** family function for estimating the parameters, for the formula of the probability density function and other details.

Value

`dgenray` gives the density, `pgenray` gives the distribution function, `qgenray` gives the quantile function, and `rgenray` generates random deviates.

Note

We define scale as the reciprocal of the scale parameter used by Kundu and Raqab (2005).

Author(s)

Kai Huang and J. G. Lauder and T. W. Yee

See Also

[genrayleigh](#), [rayleigh](#).

Examples

```
## Not run:
shape <- 0.5; Scale <- 1; nn <- 501
x <- seq(-0.10, 3.0, len = nn)
plot(x, dgenray(x, shape, scale = Scale), type = "l", las = 1, ylim = c(0, 1.2),
      ylab = paste("[dp]genray(shape = ", shape, ", scale = ", Scale, ")"),
      col = "blue", cex.main = 0.8,
      main = "Blue is density, orange is cumulative distribution function",
      sub = "Purple lines are the 10,20,...,90 percentiles")
lines(x, pgenray(x, shape, scale = Scale), col = "orange")
probs <- seq(0.1, 0.9, by = 0.1)
Q <- qgenray(probs, shape, scale = Scale)
lines(Q, dgenray(Q, shape, scale = Scale), col = "purple", lty = 3, type = "h")
lines(Q, pgenray(Q, shape, scale = Scale), col = "purple", lty = 3, type = "h")
abline(h = probs, col = "purple", lty = 3)
max(abs(pgenray(Q, shape, scale = Scale) - probs)) # Should be 0

## End(Not run)
```

genrayleigh

Generalized Rayleigh Distribution Family Function

Description

Estimates the two parameters of the generalized Rayleigh distribution by maximum likelihood estimation.

Usage

```
genrayleigh(lscale = "loglink", lshape = "loglink",
            iscale = NULL,  ishape = NULL,
            tol12 = 1e-05, nsimEIM = 300, zero = 2)
```

Arguments

lscale, lshape	Link function for the two positive parameters, scale and shape. See Links for more choices.
iscale, ishape	Numeric. Optional initial values for the scale and shape parameters.
nsimEIM, zero	See CommonVGAMffArguments .
tol12	Numeric and positive. Tolerance for testing whether the second shape parameter is either 1 or 2. If so then the working weights need to handle these singularities.

Details

The generalized Rayleigh distribution has density function

$$f(y; b = \text{scale}, s = \text{shape}) = (2sy/b^2)e^{-(y/b)^2}(1 - e^{-(y/b)^2})^{s-1}$$

where $y > 0$ and the two parameters, b and s , are positive. The mean cannot be expressed nicely so the median is returned as the fitted values. Applications of the generalized Rayleigh distribution include modeling strength data and general lifetime data. Simulated Fisher scoring is implemented.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Note

We define scale as the reciprocal of the scale parameter used by Kundu and Raqab (2005).

Author(s)

J. G. Lauder and T. W. Yee

References

Kundu, D., Raqab, M. C. (2005). Generalized Rayleigh distribution: different methods of estimations. *Computational Statistics and Data Analysis*, **49**, 187–200.

See Also

[dgenray](#), [rayleigh](#).

Examples

```
## Not run: Scale <- exp(1); shape <- exp(1)
rdata <- data.frame(y = rgenray(n = 1000, scale = Scale, shape = shape))
fit <- vglm(y ~ 1, genrayleigh, data = rdata, trace = TRUE)
c(with(rdata, mean(y)), head(fitted(fit), 1))
coef(fit, matrix = TRUE)
Coef(fit)
summary(fit)
## End(Not run)
```

Description

Density, distribution function, quantile function and random generation for the generalized secant hyperbolic distribution.

Usage

```

dgensh(x, shape, location = 0, scale = 1, tol0 = 1e-4,
       log = FALSE)
pgensh(q, shape, location = 0, scale = 1, tol0 = 1e-4,
       lower.tail = TRUE)
qgensh(p, shape, location = 0, scale = 1, tol0 = 1e-4)
rgensh(n, shape, location = 0, scale = 1, tol0 = 1e-4)

```

Arguments

`x`, `q`, `p`, `n`, `log`, `lower.tail` Similar meaning as in [Normal](#).

`shape` Numeric. Shape parameter, called t in Vaughan (2002). Valid values are $-\pi/2 < t$.

`location`, `scale` Numeric. The location and (positive) scale parameters.

`tol0` Numeric. Used to test whether the shape parameter is close enough to be treated as 0.

Details

This is an implementation of the family of symmetric densities described by Vaughan (2002). By default, the mean and variance are 0 and 1, for all t . Some special (default) cases are: $t = 0$: logistic (which is similar to `stats:dt` with 9 degrees of freedom); $t = -\pi/2$: the standard secant hyperbolic (whence the name); $t = \infty$: `uniform(-sqrt(3), sqrt(3))`.

Value

`dgensh` gives the density, `pgensh` gives the distribution function, `qgensh` gives the quantile function, and `rgensh` generates random deviates.

Warning

Numerical problems may occur when some argument values are extreme.

Author(s)

T. W. Yee.

See Also

[gensh](#), [logistic](#), [hypersecant](#), [Logistic](#).

Examples

```
x <- seq(-2, 4, by = 0.01)
loc <- 1; shape <- -pi / 2
## Not run: plot(x, dgensh(x, shape, loc), type = "l",
  main = "Blue is density, orange is the CDF",
  ylim = 0:1, las = 1, ylab = "",
  sub = "Purple are 5, 10, ..., 95 percentiles",
  col = "blue")
abline(h = 0, col = "blue", lty = 2)
lines(qgensh((1:19) / 20, shape, loc), type = "h",
  dgensh(qgensh((1:19) / 20, shape, loc),
  shape, loc), col = "purple", lty = 3)
lines(x, pgensh(x, shape, loc), col = "orange")
abline(h = 0, lty = 2)
## End(Not run)

pp <- (1:19) / 20 # Test two functions
max(abs(pgensh(qgensh(pp, shape, loc),
  shape, loc) - pp)) # Should be 0
```

gensh

Generalized Secant Hyperbolic Regression Family Function

Description

Estimation of the parameters of the generalized secant hyperbolic distribution.

Usage

```
gensh(shape, llocation = "identitylink",
  lscale = "loglink", zero = "scale",
  ilocation = NULL, iscale = NULL, imethod = 1,
  glocation.mux = exp((-4:4)/2),
  gscale.mux = exp((-4:4)/2),
  probs.y = 0.3, tol0 = 1e-4)
```

Arguments

shape Numeric of length 1. Shape parameter, called t in Vaughan (2002). Valid values are $-\pi/2 < t$.

llocation, lscale Parameter link functions applied to the two parameters. See [Links](#) for more choices. See [CommonVGAMffArguments](#) for more information.

zero, imethod See [CommonVGAMffArguments](#) for information.

`ilocation, iscale` See [CommonVGAMffArguments](#) for information.
`glocation.mux, gscale.mux` See [CommonVGAMffArguments](#) for information.
`probs.y, tol0` See [CommonVGAMffArguments](#) for information.

Details

The probability density function of the hyperbolic secant distribution is given by

$$f(y; a, b, s) = [(c_1/b) \exp(c_2 z)] / [\exp(2c_2 z) + 2C_3 \exp(c_2 z) + 1]$$

for shape parameter $-\pi < s$ and all real y . The scalars c_1, c_2, C_3 are functions of s . The mean of Y is the location parameter a (returned as the fitted values). All moments of the distribution are finite.

Further details about the parameterization can be found in Vaughan (2002). Fisher scoring is implemented and it has a diagonal EIM. More details are at [Gensh](#).

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Author(s)

T. W. Yee

References

Vaughan, D. C. (2002). The generalized secant hyperbolic distribution and its properties. *Communications in Statistics—Theory and Methods*, **31**(2): 219–238.

See Also

[hypersecant](#), [logistic](#).

Examples

```

sh <- -pi / 2; loc <- 2
hdata <- data.frame(x2 = rnorm(nn <- 200))
hdata <- transform(hdata, y = rgensh(nn, sh, loc))
fit <- vglm(y ~ x2, gensh(sh), hdata, trace = TRUE)
coef(fit, matrix = TRUE)

```

geometric

Geometric (Truncated and Untruncated) Distributions

Description

Maximum likelihood estimation for the geometric and truncated geometric distributions.

Usage

```
geometric(link = "logitlink", expected = TRUE, imethod = 1,
          iprob = NULL, zero = NULL)
truncgeometric(upper.limit = Inf,
               link = "logitlink", expected = TRUE, imethod = 1,
               iprob = NULL, zero = NULL)
```

Arguments

link	Parameter link function applied to the probability parameter p , which lies in the unit interval. See Links for more choices.
expected	Logical. Fisher scoring is used if expected = TRUE, else Newton-Raphson.
iprob, imethod, zero	See CommonVGAMffArguments for details.
upper.limit	Numeric. Upper values. As a vector, it is recycled across responses first. The default value means both family functions should give the same result.

Details

A random variable Y has a 1-parameter geometric distribution if $P(Y = y) = p(1 - p)^y$ for $y = 0, 1, 2, \dots$. Here, p is the probability of success, and Y is the number of (independent) trials that are fails until a success occurs. Thus the response Y should be a non-negative integer. The mean of Y is $E(Y) = (1 - p)/p$ and its variance is $Var(Y) = (1 - p)/p^2$. The geometric distribution is a special case of the negative binomial distribution (see [negbinomial](#)). The geometric distribution is also a special case of the Borel distribution, which is a Lagrangian distribution. If Y has a geometric distribution with parameter p then $Y + 1$ has a positive-geometric distribution with the same parameter. Multiple responses are permitted.

For `truncgeometric()`, the (upper) truncated geometric distribution can have response integer values from 0 to `upper.limit`. It has density $\text{prob} * (1 - \text{prob})^y / [1 - (1 - \text{prob})^{(1 + \text{upper.limit})}]$.

For a generalized truncated geometric distribution with integer values L to U , say, subtract L from the response and feed in $U - L$ as the upper limit.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as `vglm`, and `vgam`.

Author(s)

T. W. Yee. Help from Viet Hoang Quoc is gratefully acknowledged.

References

Forbes, C., Evans, M., Hastings, N. and Peacock, B. (2011). *Statistical Distributions*, Hoboken, NJ, USA: John Wiley and Sons, Fourth edition.

See Also

[negbinomial](#), [Geometric](#), [betageometric](#), [expgeometric](#), [zageometric](#), [zigeometric](#), [rbetageom](#), [simulate.vlm](#).

Examples

```
gdata <- data.frame(x2 = runif(nn <- 1000) - 0.5)
gdata <- transform(gdata, x3 = runif(nn) - 0.5,
                  x4 = runif(nn) - 0.5)
gdata <- transform(gdata, eta = -1.0 - 1.0 * x2 + 2.0 * x3)
gdata <- transform(gdata, prob = logitlink(eta, inverse = TRUE))
gdata <- transform(gdata, y1 = rgeom(nn, prob))
with(gdata, table(y1))
fit1 <- vglm(y1 ~ x2 + x3 + x4, geometric, data = gdata, trace = TRUE)
coef(fit1, matrix = TRUE)
summary(fit1)

# Truncated geometric (between 0 and upper.limit)
upper.limit <- 5
tdata <- subset(gdata, y1 <= upper.limit)
nrow(tdata) # Less than nn
fit2 <- vglm(y1 ~ x2 + x3 + x4, truncgeometric(upper.limit),
            data = tdata, trace = TRUE)
coef(fit2, matrix = TRUE)

# Generalized truncated geometric (between lower.limit and upper.limit)
lower.limit <- 1
upper.limit <- 8
gtdata <- subset(gdata, lower.limit <= y1 & y1 <= upper.limit)
with(gtdata, table(y1))
nrow(gtdata) # Less than nn
fit3 <- vglm(y1 - lower.limit ~ x2 + x3 + x4,
            truncgeometric(upper.limit - lower.limit),
            data = gtdata, trace = TRUE)
coef(fit3, matrix = TRUE)
```


Description

Retrieve one component of the list `.smart.prediction` from `smartpredenv`.

Usage

```
get.smart()
```

Details

`get.smart` is used in "read" mode within a smart function: it retrieves parameters saved at the time of fitting, and is used for prediction. `get.smart` is only used in smart functions such as [sm.poly](#); `get.smart.prediction` is only used in modelling functions such as [lm](#) and [glm](#). The function `get.smart` gets only a part of `.smart.prediction` whereas `get.smart.prediction` gets the entire `.smart.prediction`.

Value

Returns with one list component of `.smart.prediction` from `smartpredenv`, in fact, `.smart.prediction[[]smart.prediction`. The whole procedure mimics a first-in first-out stack (better known as a *queue*).

Side Effects

The variable `.smart.prediction.counter` in `smartpredenv` is incremented beforehand, and then written back to `smartpredenv`.

See Also

[get.smart.prediction](#).

Examples

```
print(sm.min1)
```

`get.smart.prediction` *Retrieves ".smart.prediction"*

Description

Retrieves `.smart.prediction` from `smartpredenv`.

Usage

```
get.smart.prediction()
```

Details

A smart modelling function such as [lm](#) allows smart functions such as [sm.bs](#) to write to a data structure called `.smart.prediction` in `smartpredenv`. At the end of fitting, `get.smart.prediction` retrieves this data structure. It is then attached to the object, and used for prediction later.

Value

Returns with the list `.smart.prediction` from `smartpredenv`.

See Also

[get.smart](#), [lm](#).

Examples

```
## Not run:
fit$smart <- get.smart.prediction() # Put at the end of lm()

## End(Not run)
```

 gev

Generalized Extreme Value Regression Family Function

Description

Maximum likelihood estimation of the 3-parameter generalized extreme value (GEV) distribution.

Usage

```
gev(llocation = "identitylink", lscale = "loglink",
    lshape = logofflink(offset = 0.5), percentiles = c(95, 99),
    ilocation = NULL, iscale = NULL, ishape = NULL, imethod = 1,
    gprobs.y = (1:9)/10, gscale.mux = exp((-5:5)/6),
    gshape = (-5:5) / 11 + 0.01,
    iprobs.y = NULL, tolshape0 = 0.001,
    type.fitted = c("percentiles", "mean"),
    zero = c("scale", "shape"))
gevff(llocation = "identitylink", lscale = "loglink",
      lshape = logofflink(offset = 0.5), percentiles = c(95, 99),
      ilocation = NULL, iscale = NULL, ishape = NULL, imethod = 1,
      gprobs.y = (1:9)/10, gscale.mux = exp((-5:5)/6),
      gshape = (-5:5) / 11 + 0.01,
      iprobs.y = NULL, tolshape0 = 0.001,
      type.fitted = c("percentiles", "mean"), zero = c("scale", "shape"))
```

Arguments

`llocation`, `lscale`, `lshape`

Parameter link functions for μ , σ and ξ respectively. See [Links](#) for more choices.

For the shape parameter, the default [logofflink](#) link has an offset called A below; and then the linear/additive predictor is $\log(\xi + A)$ which means that $\xi > -A$. For technical reasons (see **Details**) it is a good idea for $A = 0.5$.

percentiles	Numeric vector of percentiles used for the fitted values. Values should be between 0 and 100. This argument is ignored if <code>type.fitted = "mean"</code> .
type.fitted	See CommonVGAMffArguments for information. The default is to use the percentiles argument. If "mean" is chosen, then the mean $\mu + \sigma(\Gamma(1 - \xi) - 1)/\xi$ is returned as the fitted values, and these are only defined for $\xi < 1$.
ilocation, iscale, ishape	Numeric. Initial value for the location parameter, σ and ξ . A NULL means a value is computed internally. The argument <code>ishape</code> is more important than the other two. If a failure to converge occurs, or even to obtain initial values occurs, try assigning <code>ishape</code> some value (positive or negative; the sign can be very important). Also, in general, a larger value of <code>iscale</code> tends to be better than a smaller value.
imethod	Initialization method. Either the value 1 or 2. If both methods fail then try using <code>ishape</code> . See CommonVGAMffArguments for information.
gshape	Numeric vector. The values are used for a grid search for an initial value for ξ . See CommonVGAMffArguments for information.
gprobs.y, gscale.mux, iprobs.y	Numeric vectors, used for the initial values. See CommonVGAMffArguments for information.
tolshape0	Passed into dgev when computing the log-likelihood.
zero	A specifying which linear/additive predictors are modelled as intercepts only. The values can be from the set {1,2,3} corresponding respectively to μ , σ , ξ . If <code>zero = NULL</code> then all linear/additive predictors are modelled as a linear combination of the explanatory variables. For many data sets having <code>zero = 3</code> is a good idea. See CommonVGAMffArguments for information.

Details

The GEV distribution function can be written

$$G(y) = \exp(-[(y - \mu)/\sigma]_+^{-1/\xi})$$

where $\sigma > 0$, $-\infty < \mu < \infty$, and $1 + \xi(y - \mu)/\sigma > 0$. Here, $x_+ = \max(x, 0)$. The μ , σ , ξ are known as the *location*, *scale* and *shape* parameters respectively. The cases $\xi > 0$, $\xi < 0$, $\xi = 0$ correspond to the Frechet, reverse Weibull, and Gumbel types respectively. It can be noted that the Gumbel (or Type I) distribution accommodates many commonly-used distributions such as the normal, lognormal, logistic, gamma, exponential and Weibull.

For the GEV distribution, the k th moment about the mean exists if $\xi < 1/k$. Provided they exist, the mean and variance are given by $\mu + \sigma\{\Gamma(1 - \xi) - 1\}/\xi$ and $\sigma^2\{\Gamma(1 - 2\xi) - \Gamma^2(1 - \xi)\}/\xi^2$ respectively, where Γ is the gamma function.

Smith (1985) established that when $\xi > -0.5$, the maximum likelihood estimators are completely regular. To have some control over the estimated ξ try using `lshape = logofflink(offset = 0.5)`, say, or `lshape = extlogitlink(min = -0.5, max = 0.5)`, say.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Warning

Currently, if an estimate of ξ is too close to 0 then an error may occur for `gev()` with multivariate responses. In general, `gevff()` is more reliable than `gev()`.

Fitting the GEV by maximum likelihood estimation can be numerically fraught. If $1 + \xi(y - \mu)/\sigma \leq 0$ then some crude evasive action is taken but the estimation process can still fail. This is particularly the case if `vgam` with `s` is used; then smoothing is best done with `vglm` with regression splines (`bs` or `ns`) because `vglm` implements half-stepsizing whereas `vgam` doesn't (half-stepsizing helps handle the problem of straying outside the parameter space.)

Note

The **VGAM** family function `gev` can handle a multivariate (matrix) response, cf. multiple responses. If so, each row of the matrix is sorted into descending order and NAs are put last. With a vector or one-column matrix response using `gevff` will give the same result but be faster and it handles the $\xi = 0$ case. The function `gev` implements Tawn (1988) while `gevff` implements Prescott and Walden (1980).

Function `egev()` has been replaced by the new family function `gevff()`. It now conforms to the usual **VGAM** philosophy of having M1 linear predictors per (independent) response. This is the usual way multiple responses are handled. Hence `vglm(cbind(y1, y2)... , gevff, ...)` will have 6 linear predictors and it is possible to constrain the linear predictors so that the answer is similar to `gev()`. Missing values in the response of `gevff()` will be deleted; this behaviour is the same as with almost every other **VGAM** family function.

The shape parameter ξ is difficult to estimate accurately unless there is a lot of data. Convergence is slow when ξ is near -0.5 . Given many explanatory variables, it is often a good idea to make sure `zero = 3`. The range restrictions of the parameter ξ are not enforced; thus it is possible for a violation to occur.

Successful convergence often depends on having a reasonably good initial value for ξ . If failure occurs try various values for the argument `ishape`, and if there are covariates, having `zero = 3` is advised.

Author(s)

T. W. Yee

References

- Yee, T. W. and Stephenson, A. G. (2007). Vector generalized linear and additive extreme value models. *Extremes*, **10**, 1–19.
- Tawn, J. A. (1988). An extreme-value theory model for dependent observations. *Journal of Hydrology*, **101**, 227–250.
- Prescott, P. and Walden, A. T. (1980). Maximum likelihood estimation of the parameters of the generalized extreme-value distribution. *Biometrika*, **67**, 723–724.
- Smith, R. L. (1985). Maximum likelihood estimation in a class of nonregular cases. *Biometrika*, **72**, 67–90.

See Also

[rgev](#), [gumbel](#), [gumbelfff](#), [guplot](#), [rlplot.gevfff](#), [gpd](#), [weibullR](#), [frechet](#), [extlogitlink](#), [oxtemp](#), [venice](#), [CommonVGAMffArguments](#).

Examples

```
## Not run:
# Multivariate example
fit1 <- vgam(cbind(r1, r2) ~ s(year, df = 3), gev(zero = 2:3),
            data = venice, trace = TRUE)
coef(fit1, matrix = TRUE)
head(fitted(fit1))
par(mfrow = c(1, 2), las = 1)
plot(fit1, se = TRUE, lcol = "blue", scol = "forestgreen",
     main = "Fitted mu(year) function (centered)", cex.main = 0.8)
with(venice, matplot(year, depvar(fit1)[, 1:2], ylab = "Sea level (cm)",
                    col = 1:2, main = "Highest 2 annual sea levels", cex.main = 0.8))
with(venice, lines(year, fitted(fit1)[,1], lty = "dashed", col = "blue"))
legend("topleft", lty = "dashed", col = "blue", "Fitted 95 percentile")

# Univariate example
(fit <- vglm(maxtemp ~ 1, gevfff, data = oxtemp, trace = TRUE))
head(fitted(fit))
coef(fit, matrix = TRUE)
Coef(fit)
vcov(fit)
vcov(fit, untransform = TRUE)
sqrt(diag(vcov(fit))) # Approximate standard errors
rlplot(fit)

## End(Not run)
```

 gevUC

The Generalized Extreme Value Distribution

Description

Density, distribution function, quantile function and random generation for the generalized extreme value distribution (GEV) with location parameter location, scale parameter scale and shape parameter shape.

Usage

```
dgev(x, location = 0, scale = 1, shape = 0, log = FALSE,
     tolshape0 = sqrt(.Machine$double.eps))
pgev(q, location = 0, scale = 1, shape = 0, lower.tail = TRUE, log.p = FALSE)
qgev(p, location = 0, scale = 1, shape = 0, lower.tail = TRUE, log.p = FALSE)
rgev(n, location = 0, scale = 1, shape = 0)
```

Arguments

x, q	vector of quantiles.
p	vector of probabilities.
n	number of observations. If <code>length(n) > 1</code> then the length is taken to be the number required.
location	the location parameter μ .
scale	the (positive) scale parameter σ . Must consist of positive values.
shape	the shape parameter ξ .
log	Logical. If <code>log = TRUE</code> then the logarithm of the density is returned.
lower.tail, log.p	Same meaning as in <code>punif</code> or <code>qunif</code> .
tolshape0	Positive numeric. Threshold/tolerance value for resting whether ξ is zero. If the absolute value of the estimate of ξ is less than this value then it will be assumed zero and a Gumbel distribution will be used.

Details

See `gev`, the **VGAM** family function for estimating the 3 parameters by maximum likelihood estimation, for formulae and other details. Apart from `n`, all the above arguments may be vectors and are recycled to the appropriate length if necessary.

Value

`dgev` gives the density, `pgev` gives the distribution function, `qgev` gives the quantile function, and `rgev` generates random deviates.

Note

The default value of $\xi = 0$ means the default distribution is the Gumbel.

Currently, these functions have different argument names compared with those in the **evd** package.

Author(s)

T. W. Yee

References

Coles, S. (2001). *An Introduction to Statistical Modeling of Extreme Values*. London: Springer-Verlag.

See Also

`gev`, `gevff`, `vglm.control`.

Examples

```

loc <- 2; sigma <- 1; xi <- -0.4
pgev(qgev(seq(0.05, 0.95, by = 0.05), loc, sigma, xi), loc, sigma, xi)
## Not run: x <- seq(loc - 3, loc + 3, by = 0.01)
plot(x, dgev(x, loc, sigma, xi), type = "l", col = "blue", ylim = c(0, 1),
     main = "Blue is density, orange is the CDF",
     sub = "Purple are 10,...,90 percentiles", ylab = "", las = 1)
abline(h = 0, col = "blue", lty = 2)
lines(qgev(seq(0.1, 0.9, by = 0.1), loc, sigma, xi),
     dgev(qgev(seq(0.1, 0.9, by = 0.1), loc, sigma, xi), loc, sigma, xi),
     col = "purple", lty = 3, type = "h")
lines(x, pgev(x, loc, sigma, xi), type = "l", col = "orange")
abline(h = (0:10)/10, lty = 2, col = "gray50")

## End(Not run)

```

gew

General Electric and Westinghouse Data

Description

General Electric and Westinghouse capital data.

Usage

```
data(gew)
```

Format

A data frame with 20 observations on the following 7 variables. All variables are numeric vectors. Variables ending in `.g` correspond to General Electric and those ending in `.w` are Westinghouse.

year The observations are the years from 1934 to 1953

invest.g, invest.w investment figures. These are I = Gross investment = additions to plant and equipment plus maintenance and repairs in millions of dollars deflated by P_1 .

capital.g, capital.w capital stocks. These are C = The stock of plant and equipment = accumulated sum of net additions to plant and equipment deflated by P_1 minus depreciation allowance deflated by P_3 .

value.g, value.w market values. These are F = Value of the firm = price of common and preferred shares at December 31 (or average price of December 31 and January 31 of the following year) times number of common and preferred shares outstanding plus total book value of debt at December 31 in millions of dollars deflated by P_2 .

Details

These data are a subset of a table in Boot and de Wit (1960), also known as the Grunfeld data. It is used a lot in econometrics, e.g., for seemingly unrelated regressions (see [SURff](#)).

Here, P_1 = Implicit price deflator of producers durable equipment (base 1947), P_2 = Implicit price deflator of G.N.P. (base 1947), P_3 = Depreciation expense deflator = ten years moving average of wholesale price index of metals and metal products (base 1947).

Source

Table 10 of: Boot, J. C. G. and de Wit, G. M. (1960) Investment Demand: An Empirical Contribution to the Aggregation Problem. *International Economic Review*, **1**, 3–30.

Grunfeld, Y. (1958) The Determinants of Corporate Investment. Unpublished PhD Thesis (Chicago).

References

Zellner, A. (1962). An efficient method of estimating seemingly unrelated regressions and tests for aggregation bias. *Journal of the American Statistical Association*, **57**, 348–368.

See Also

[SURff](#), <http://statmath.wu.ac.at/~zeileis/grunfeld> (the link might now be stale).

Examples

```
str(gew)
```

goffset

GAITD Offset for the GTE Method

Description

Utility function to create a matrix of log-offset values, to help facilitate the Generally-Truncated-Expansion method

Usage

```
goffset(mux, n,
        a.mix = NULL, i.mix = NULL, d.mix = NULL,
        a.mlm = NULL, i.mlm = NULL, d.mlm = NULL, par1or2 = 1)
```


Arguments

<code>mux</code>	Multiplier. Usually a small positive integer. Must be positive. The value 1 means no change.
<code>n</code>	Number of rows. A positive integer, it should be the number of rows of the data frame containing the data.
<code>a.mix, i.mix, d.mix</code>	See, e.g., gaitdpoisson .
<code>a.mlm, i.mlm, d.mlm</code>	See, e.g., gaitdpoisson .
<code>par1or2</code>	Number of parameters of the parent distribution. Set <code>par1or2 = 2</code> for gaitdnbinomial , else the default value should be used.

Details

This function is intended to make the Generally-Truncated-Expansion (GTE) method easier for the user. It only makes sense if the linear predictors(s) are log of the mean of the parent distribution, which is the usual case for [gaitdpoisson](#) and [gaitdnbinomial](#). However, for [gaitdlog](#) and [gaitdzeta](#) one should be using [logffMlink](#) and [zetaffMlink](#).

Without this function, the user must do quite a lot of book-keeping to know which columns of the offset matrix is to be assigned $\log(\text{mux})$. This can be rather laborious.

In the fictional example below the response is underdispersed with respect to a Poisson distribution and doubling the response achieves approximate equidispersion.

Value

A matrix with `n` rows and the same number of columns that a GAITD regression would produce for its matrix of linear predictors. The matrix can be inputted into [vglm](#) by assigning the `offset` argument.

Note

This function is still in a developmental stage. The order of the arguments might change, hence it's safest to invoke it with full specification.

See Also

[gaitdpoisson](#), [gaitdlog](#), [gaitdzeta](#), [gaitdnbinomial](#), [Trunc](#), [offset](#).

Examples

```
i.mix <- c(5, 10, 15, 20); a.mlm <- 13; mymux <- 2
goffset(mymux, 10, i.mix = i.mix, a.mlm = a.mlm)
## Not run: org1 <- with(gdata, range(y)) # Original range of the data
vglm(mymux * y ~ 1,
     offset = goffset(mymux, nrow(gdata), i.mix = i.mix, a.mlm = a.mlm),
     gaitdpoisson(a.mlm = mymux * a.mlm, i.mix = mymux * i.mix,
                 truncate = Trunc(org1, mymux)),
     data = gdata)
```

```
## End(Not run)
```

Gompertz

Gompertz Distribution

Description

Density, cumulative distribution function, quantile function and random generation for the Gompertz distribution.

Usage

```
dgompertz(x, scale = 1, shape, log = FALSE)
pgompertz(q, scale = 1, shape, lower.tail = TRUE, log.p = FALSE)
qgompertz(p, scale = 1, shape, lower.tail = TRUE, log.p = FALSE)
rgompertz(n, scale = 1, shape)
```

Arguments

x, q	vector of quantiles.
p	vector of probabilities.
n	number of observations. Same as in runif .
log	Logical. If log = TRUE then the logarithm of the density is returned.
lower.tail, log.p	Same meaning as in pnorm or qnorm .
scale, shape	positive scale and shape parameters.

Details

See [gompertz](#) for details.

Value

dgompertz gives the density, pgompertz gives the cumulative distribution function, qgompertz gives the quantile function, and rgompertz generates random deviates.

Author(s)

T. W. Yee and Kai Huang

See Also

[gompertz](#), [dgumbel](#), [dmakeham](#).

Examples

```

probs <- seq(0.01, 0.99, by = 0.01)
Shape <- exp(1); Scale <- exp(1)
max(abs(pgompertz(qgompertz(p = probs, Scale, shape = Shape),
                  Scale, shape = Shape) - probs)) # Should be 0

## Not run: x <- seq(-0.1, 1.0, by = 0.001)
plot(x, dgompertz(x, Scale, shape = Shape), type = "l", las = 1,
     main = "Blue is density, orange is the CDF", col = "blue",
     sub = "Purple lines are the 10,20,...,90 percentiles",
     ylab = "")
abline(h = 0, col = "blue", lty = 2)
lines(x, pgompertz(x, Scale, shape = Shape), col = "orange")
probs <- seq(0.1, 0.9, by = 0.1)
Q <- qgompertz(probs, Scale, shape = Shape)
lines(Q, dgompertz(Q, Scale, shape = Shape), col = "purple",
      lty = 3, type = "h")
pgompertz(Q, Scale, shape = Shape) - probs # Should be all zero
abline(h = probs, col = "purple", lty = 3)
## End(Not run)

```

gompertz

*Gompertz Regression Family Function***Description**

Maximum likelihood estimation of the 2-parameter Gompertz distribution.

Usage

```

gompertz(lscale = "loglink", lshape = "loglink",
         iscale = NULL,  ishape = NULL,
         nsimEIM = 500, zero = NULL, nowarning = FALSE)

```

Arguments

nowarning	Logical. Suppress a warning? Ignored for VGAM 0.9-7 and higher.
lshape, lscale	Parameter link functions applied to the shape parameter a, scale parameter scale. All parameters are positive. See Links for more choices.
ishape, iscale	Optional initial values. A NULL means a value is computed internally.
nsimEIM, zero	See CommonVGAMffArguments .

Details

The Gompertz distribution has a cumulative distribution function

$$F(x; \alpha, \beta) = 1 - \exp[-(\alpha/\beta) \times (\exp(\beta x) - 1)]$$

which leads to a probability density function

$$f(x; \alpha, \beta) = \alpha \exp(\beta x) \exp[-(\alpha/\beta) \times (\exp(\beta x) - 1)]$$

for $\alpha > 0$, $\beta > 0$, $x > 0$. Here, β is called the scale parameter *scale*, and α is called the shape parameter (one could refer to α as a location parameter and β as a shape parameter—see Lenart (2014)). The mean involves an exponential integral function. Simulated Fisher scoring is used and multiple responses are handled.

The Makeham distribution has an additional parameter compared to the Gompertz distribution. If X is defined to be the result of sampling from a Gumbel distribution until a negative value Z is produced, then $X = -Z$ has a Gompertz distribution.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Warning

The same warnings in [makeham](#) apply here too.

Author(s)

T. W. Yee

References

Lenart, A. (2014). The moments of the Gompertz distribution and maximum likelihood estimation of its parameters. *Scandinavian Actuarial Journal*, **2014**, 255–277.

See Also

[dgompertz](#), [makeham](#), [simulate.vlm](#).

Examples

```
## Not run:
gdata <- data.frame(x2 = runif(nn <- 1000))
gdata <- transform(gdata, eta1 = -1,
                  eta2 = -1 + 0.2 * x2,
                  ceta1 = 1,
                  ceta2 = -1 + 0.2 * x2)
gdata <- transform(gdata, shape1 = exp(eta1),
                  shape2 = exp(eta2),
                  scale1 = exp(ceta1),
                  scale2 = exp(ceta2))
gdata <- transform(gdata, y1 = rgompertz(nn, scale = scale1, shape = shape1),
                  y2 = rgompertz(nn, scale = scale2, shape = shape2))

fit1 <- vglm(y1 ~ 1, gompertz, data = gdata, trace = TRUE)
fit2 <- vglm(y2 ~ x2, gompertz, data = gdata, trace = TRUE)
```

```

coef(fit1, matrix = TRUE)
Coef(fit1)
summary(fit1)
coef(fit2, matrix = TRUE)
summary(fit2)

## End(Not run)

```

gpd

*Generalized Pareto Distribution Regression Family Function***Description**

Maximum likelihood estimation of the 2-parameter generalized Pareto distribution (GPD).

Usage

```

gpd(threshold = 0, lscale = "loglink", lshape = logofflink(offset = 0.5),
    percentiles = c(90, 95), iscale = NULL, ishape = NULL,
    tolshape0 = 0.001, type.fitted = c("percentiles", "mean"),
    imethod = 1, zero = "shape")

```

Arguments

threshold	Numeric, values are recycled if necessary. The threshold value(s), called μ below.
lscale	Parameter link function for the scale parameter σ . See Links for more choices.
lshape	Parameter link function for the shape parameter ξ . See Links for more choices. The default constrains the parameter to be greater than -0.5 because if $\xi \leq -0.5$ then Fisher scoring does not work. See the Details section below for more information. For the shape parameter, the default logofflink link has an offset called A below; and then the second linear/additive predictor is $\log(\xi + A)$ which means that $\xi > -A$. The working weight matrices are positive definite if $A = 0.5$.
percentiles	Numeric vector of percentiles used for the fitted values. Values should be between 0 and 100. See the example below for illustration. This argument is ignored if <code>type.fitted = "mean"</code> .
type.fitted	See CommonVGAMffArguments for information. The default is to use the percentiles argument. If "mean" is chosen, then the mean $\mu + \sigma/(1 - \xi)$ is returned as the fitted values, and these are only defined for $\xi < 1$.
iscale, ishape	Numeric. Optional initial values for σ and ξ . The default is to use <code>imethod</code> and compute a value internally for each parameter. Values of <code>ishape</code> should be between -0.5 and 1. Values of <code>iscale</code> should be positive.
tolshape0	Passed into dgpd when computing the log-likelihood.

<code>imethod</code>	Method of initialization, either 1 or 2. The first is the method of moments, and the second is a variant of this. If neither work, try assigning values to arguments <code>ishape</code> and/or <code>iscale</code> .
<code>zero</code>	Can be an integer-valued vector specifying which linear/additive predictors are modelled as intercepts only. For one response, the value should be from the set {1,2} corresponding respectively to σ and ξ . It is often a good idea for the σ parameter only to be modelled through a linear combination of the explanatory variables because the shape parameter is probably best left as an intercept only: <code>zero = 2</code> . Setting <code>zero = NULL</code> means both parameters are modelled with explanatory variables. See CommonVGAMffArguments for more details.

Details

The distribution function of the GPD can be written

$$G(y) = 1 - [1 + \xi(y - \mu)/\sigma]_+^{-1/\xi}$$

where μ is the location parameter (known, with value threshold), $\sigma > 0$ is the scale parameter, ξ is the shape parameter, and $h_+ = \max(h, 0)$. The function $1 - G$ is known as the *survivor function*. The limit $\xi \rightarrow 0$ gives the *shifted exponential* as a special case:

$$G(y) = 1 - \exp[-(y - \mu)/\sigma].$$

The support is $y > \mu$ for $\xi > 0$, and $\mu < y < \mu - \sigma/\xi$ for $\xi < 0$.

Smith (1985) showed that if $\xi \leq -0.5$ then this is known as the nonregular case and problems/difficulties can arise both theoretically and numerically. For the (regular) case $\xi > -0.5$ the classical asymptotic theory of maximum likelihood estimators is applicable; this is the default.

Although for $\xi < -0.5$ the usual asymptotic properties do not apply, the maximum likelihood estimator generally exists and is superefficient for $-1 < \xi < -0.5$, so it is “better” than normal. When $\xi < -1$ the maximum likelihood estimator generally does not exist as it effectively becomes a two parameter problem.

The mean of Y does not exist unless $\xi < 1$, and the variance does not exist unless $\xi < 0.5$. So if you want to fit a model with finite variance use `lshape = "extlogitlink"`.

Value

An object of class “`vglmff`” (see [vglmff-class](#)). The object is used by modelling functions such as `vglm` and `vgam`. However, for this **VGAM** family function, `vglm` is probably preferred over `vgam` when there is smoothing.

Warning

Fitting the GPD by maximum likelihood estimation can be numerically fraught. If $1 + \xi(y - \mu)/\sigma \leq 0$ then some crude evasive action is taken but the estimation process can still fail. This is particularly the case if `vgam` with `s` is used. Then smoothing is best done with `vglm` with regression splines (`bs` or `ns`) because `vglm` implements half-stepsizing whereas `vgam` doesn't. Half-stepsizing helps handle the problem of straying outside the parameter space.

Note

The response in the formula of `vglm` and `vgam` is y . Internally, $y - \mu$ is computed. This **VGAM** family function can handle a multiple responses, which is inputted as a matrix. The response stored on the object is the original uncentred data.

With functions `rgpd`, `dgpd`, etc., the argument `location` matches with the argument `threshold` here.

Author(s)

T. W. Yee

References

- Yee, T. W. and Stephenson, A. G. (2007). Vector generalized linear and additive extreme value models. *Extremes*, **10**, 1–19.
- Coles, S. (2001). *An Introduction to Statistical Modeling of Extreme Values*. London: Springer-Verlag.
- Smith, R. L. (1985). Maximum likelihood estimation in a class of nonregular cases. *Biometrika*, **72**, 67–90.

See Also

[rgpd](#), [meplot](#), [gev](#), [paretoff](#), [vglm](#), [vgam](#), [s](#).

Examples

```
# Simulated data from an exponential distribution (xi = 0)
Threshold <- 0.5
gdata <- data.frame(y1 = Threshold + rexp(n = 3000, rate = 2))
fit <- vglm(y1 ~ 1, gpd(threshold = Threshold), data = gdata, trace = TRUE)
head(fitted(fit))
summary(depvar(fit)) # The original uncentred data
coef(fit, matrix = TRUE) # xi should be close to 0
Coef(fit)
summary(fit)

head(fit@extra$threshold) # Note the threshold is stored here

# Check the 90 percentile
ii <- depvar(fit) < fitted(fit)[1, "90%"]
100 * table(ii) / sum(table(ii)) # Should be 90%

# Check the 95 percentile
ii <- depvar(fit) < fitted(fit)[1, "95%"]
100 * table(ii) / sum(table(ii)) # Should be 95%

## Not run: plot(depvar(fit), col = "blue", las = 1,
                main = "Fitted 90% and 95% quantiles")
matlines(1:length(depvar(fit)), fitted(fit), lty = 2:3, lwd = 2)
## End(Not run)
```

```

# Another example
gdata <- data.frame(x2 = runif(nn <- 2000))
Threshold <- 0; xi <- exp(-0.8) - 0.5
gdata <- transform(gdata, y2 = rgpd(nn, scale = exp(1 + 0.1*x2), shape = xi))
fit <- vglm(y2 ~ x2, gpd(Threshold), data = gdata, trace = TRUE)
coef(fit, matrix = TRUE)

## Not run: # Nonparametric fits
# Not so recommended:
fit1 <- vgam(y2 ~ s(x2), gpd(Threshold), data = gdata, trace = TRUE)
par(mfrow = c(2, 1))
plot(fit1, se = TRUE, scol = "blue")
# More recommended:
fit2 <- vglm(y2 ~ sm.bs(x2), gpd(Threshold), data = gdata, trace = TRUE)
plot(as(fit2, "vgam"), se = TRUE, scol = "blue")
## End(Not run)

```

gpdUC

*The Generalized Pareto Distribution***Description**

Density, distribution function, quantile function and random generation for the generalized Pareto distribution (GPD) with location parameter location, scale parameter scale and shape parameter shape.

Usage

```

dgpdc(x, location = 0, scale = 1, shape = 0, log = FALSE,
      tolshape0 = sqrt(.Machine$double.eps))
pgpdc(q, location = 0, scale = 1, shape = 0,
      lower.tail = TRUE, log.p = FALSE)
qgpdc(p, location = 0, scale = 1, shape = 0,
      lower.tail = TRUE, log.p = FALSE)
rgpdc(n, location = 0, scale = 1, shape = 0)

```

Arguments

x, q	vector of quantiles.
p	vector of probabilities.
n	number of observations. If <code>length(n) > 1</code> then the length is taken to be the number required.
location	the location parameter μ .
scale	the (positive) scale parameter σ .

shape	the shape parameter ξ .
log	Logical. If log = TRUE then the logarithm of the density is returned.
lower.tail, log.p	Same meaning as in punif or qunif .
tolshape0	Positive numeric. Threshold/tolerance value for resting whether ξ is zero. If the absolute value of the estimate of ξ is less than this value then it will be assumed zero and an exponential distribution will be used.

Details

See [gpd](#), the **VGAM** family function for estimating the two parameters by maximum likelihood estimation, for formulae and other details. Apart from n, all the above arguments may be vectors and are recycled to the appropriate length if necessary.

Value

dgpd gives the density, pgpd gives the distribution function, qgpd gives the quantile function, and rgpd generates random deviates.

Note

The default values of all three parameters, especially $\xi = 0$, means the default distribution is the exponential.

Currently, these functions have different argument names compared with those in the **evd** package.

Author(s)

T. W. Yee and Kai Huang

References

Coles, S. (2001). *An Introduction to Statistical Modeling of Extreme Values*. London: Springer-Verlag.

See Also

[gpd](#), [Exponential](#).

Examples

```
## Not run: loc <- 2; sigma <- 1; xi <- -0.4
x <- seq(loc - 0.2, loc + 3, by = 0.01)
plot(x, dgpd(x, loc, sigma, xi), type = "l", col = "blue",
      main = "Blue is density, red is the CDF", ylim = c(0, 1),
      sub = "Purple are 5,10,...,95 percentiles", ylab = "", las = 1)
abline(h = 0, col = "blue", lty = 2)
lines(qgpd(seq(0.05, 0.95, by = 0.05), loc, sigma, xi),
      dgpd(qgpd(seq(0.05, 0.95, by = 0.05), loc, sigma, xi), loc, sigma, xi),
      col = "purple", lty = 3, type = "h")
lines(x, pgpd(x, loc, sigma, xi), type = "l", col = "red")
```

```
abline(h = 0, lty = 2)
pgpd(qgpd(seq(0.05, 0.95, by = 0.05), loc, sigma, xi), loc, sigma, xi)
## End(Not run)
```

grain.us

Grain Prices Data in USA

Description

A 4-column matrix.

Usage

```
data(grain.us)
```

Format

The columns are:

wheat.flour numeric

corn numeric

wheat numeric

rye numeric

Details

Monthly averages of grain prices in the United States for wheat flour, corn, wheat, and rye for the period January 1961 through October 1972. The units are US dollars per 100 pound sack for wheat flour, and per bushel for corn, wheat and rye.

Source

Ahn and Reinsel (1988).

References

Ahn, S. K and Reinsel, G. C. (1988). Nested reduced-rank autoregressive models for multiple time series. *Journal of the American Statistical Association*, **83**, 849–856.

Examples

```
## Not run:
cgrain <- scale(grain.us, scale = FALSE) # Center the time series only
fit <- vglm(cgrain ~ 1, rrar(Rank = c(4, 1)),
           epsilon = 1e-3, stepsize = 0.5, trace = TRUE, maxit = 50)
summary(fit)

## End(Not run)
```

grc

Row-Column Interaction Models including Goodman's RC Association Model and Unconstrained Quadratic Ordination

Description

Fits a Goodman's RC association model (GRC) to a matrix of counts, and more generally, row-column interaction models (RCIMs). RCIMs allow for unconstrained quadratic ordination (UQO).

Usage

```
grc(y, Rank = 1, Index.corner = 2:(1 + Rank),
    str0 = 1, summary.arg = FALSE, h.step = 1e-04, ...)
rcim(y, family = poissonff, Rank = 0, M1 = NULL,
     weights = NULL, which.linpred = 1,
     Index.corner = ifelse(is.null(str0), 0, max(str0)) + 1:Rank,
     rprefix = "Row.", cprefix = "Col.", iprefix = "X2.",
     offset = 0, str0 = if (Rank) 1 else NULL,
     summary.arg = FALSE, h.step = 0.0001,
     rbaseline = 1, cbaseline = 1,
     has.intercept = TRUE, M = NULL, rindex = 2:nrow(y),
     cindex = 2:ncol(y), iindex = 2:nrow(y), ...)
```

Arguments

y	For <code>grc()</code> : a matrix of counts. For <code>rcim()</code> : a general matrix response depending on family. Output from <code>table()</code> is acceptable; it is converted into a matrix. Note that y should be at least 3 by 3 in dimension.
family	A VGAM family function. By default, the first linear/additive predictor is fitted using main effects plus an optional rank-Rank interaction term. Not all family functions are suitable or make sense. All other linear/additive predictors are fitted using an intercept-only, so it has a common value over all rows and columns. For example, <code>zipoissonff</code> may be suitable for counts but not <code>zipoisson</code> because of the ordering of the linear/additive predictors. If the VGAM family function does not have an <code>infos</code> slot then <code>M1</code> needs to be inputted (the number of linear predictors for an ordinary (usually univariate) response, aka <i>M</i>). The VGAM family function also needs to be able to handle multiple responses (currently not all of them can do this).
Rank	An integer from the set $\{0, \dots, \min(\text{nrow}(y), \text{ncol}(y))\}$. This is the dimension of the fit in terms of the interaction. For <code>grc()</code> this argument must be positive. A value of 0 means no interactions (i.e., main effects only); each row and column is represented by an indicator variable.
weights	Prior weights. Fed into <code>rrvglm</code> or <code>vglm</code> .
which.linpred	Single integer. Specifies which linear predictor is modelled as the sum of an intercept, row effect, column effect plus an optional interaction term. It should be one value from the set $1:M1$.

<code>Index.corner</code>	A vector of Rank integers. These are used to store the Rank by Rank identity matrix in the A matrix; corner constraints are used.
<code>rprefix, cprefix, iprefix</code>	Character, for rows and columns and interactions respectively. For labelling the indicator variables.
<code>offset</code>	Numeric. Either a matrix of the right dimension, else a single numeric expanded into such a matrix.
<code>str0</code>	Ignored if Rank = 0, else an integer from the set $\{1, \dots, \min(\text{nrow}(y), \text{ncol}(y))\}$, specifying the row that is used as the structural zero. Passed into <code>rrvglm.control</code> if Rank > 0. Set <code>str0 = NULL</code> for none.
<code>summary.arg</code>	Logical. If TRUE then a summary is returned. If TRUE then y may be the output (fitted object) of <code>grc()</code> .
<code>h.step</code>	A small positive value that is passed into <code>summary.rrvglm()</code> . Only used when <code>summary.arg = TRUE</code> .
<code>...</code>	Arguments that are passed into <code>rrvglm.control()</code> .
<code>M1</code>	The number of linear predictors of the VGAM family function for an ordinary (univariate) response. Then the number of linear predictors of the <code>rcim()</code> fit is usually the number of columns of y multiplied by M1. The default is to evaluate the <code>infos</code> slot of the VGAM family function to try to evaluate it; see <code>vglmff-class</code> . If this information is not yet supplied by the family function then the value needs to be inputted manually using this argument.
<code>rbaseline, cbaseline</code>	Baseline reference levels for the rows and columns. Currently stored on the object but not used.
<code>has.intercept</code>	Logical. Include an intercept?
<code>M, cindex</code>	<i>M</i> is the usual VGAM <i>M</i> , viz. the number of linear/additive predictors in total. Also, <code>cindex</code> means column index, and these point to the columns of y which are part of the vector of linear/additive predictor <i>main effects</i> . For <code>family = multinomial</code> it is necessary to input these arguments as <code>M = ncol(y)-1</code> and <code>cindex = 2:(ncol(y)-1)</code> .
<code>rindex, iindex</code>	<code>rindex</code> means row index, and these are similar to <code>cindex</code> . <code>iindex</code> means interaction index, and these are similar to <code>cindex</code> .

Details

Goodman's RC association model fits a reduced-rank approximation to a table of counts. A Poisson model is assumed. The log of each cell mean is decomposed as an intercept plus a row effect plus a column effect plus a reduced-rank component. The latter can be collectively written $A \%*\% t(C)$, the product of two 'thin' matrices. Indeed, A and C have Rank columns. By default, the first column and row of the interaction matrix $A \%*\% t(C)$ is chosen to be structural zeros, because `str0 = 1`. This means the first row of A are all zeros.

This function uses `options()`\$`contrasts` to set up the row and column indicator variables. In particular, Equation (4.5) of Yee and Hastie (2003) is used. These are called `Row.` and `Col.` (by default) followed by the row or column number.

The function `rcim()` is more general than `grc()`. Its default is a no-interaction model of `grc()`, i.e., rank-0 and a Poisson distribution. This means that each row and column has a dummy variable associated with it. The first row and first column are baseline. The power of `rcim()` is that many **VGAM** family functions can be assigned to its `family` argument. For example, `uninormal` fits something in between a 2-way ANOVA with and without interactions, `alaplac2` with `Rank = 0` is something like `medpolish`. Others include `zipoissonff` and `negbinomial`. Hopefully one day *all* **VGAM** family functions will work when assigned to the `family` argument, although the result may not have meaning.

Unconstrained quadratic ordination (UQO) can be performed using `rcim()` and `grc()`. This has been called *unconstrained Gaussian ordination* in the literature, however the word *Gaussian* has two meanings which is confusing; it is better to use *quadratic* because the bell-shape response surface is meant. UQO is similar to CQO (`cqo`) except there are no environmental/explanatory variables. Here, a GLM is fitted to each column (species) that is a quadratic function of hypothetical latent variables or gradients. Thus each row of the response has an associated site score, and each column of the response has an associated optimum and tolerance matrix. UQO can be performed here under the assumption that all species have the same tolerance matrices. See Yee and Hadi (2014) for details. It is not recommended that presence/absence data be inputted because the information content is so low for each site-species cell. The example below uses Poisson counts.

Value

An object of class "grc", which currently is the same as an "rrvglm" object. Currently, a rank-0 `rcim()` object is of class `rcim0-class`, else of class "rcim" (this may change in the future).

Warning

The function `rcim()` is experimental at this stage and may have bugs. Quite a lot of expertise is needed when fitting and in its interpretation thereof. For example, the constraint matrices applies the reduced-rank regression to the first (see `which.linpred`) linear predictor and the other linear predictors are intercept-only and have a common value throughout the entire data set. This means that, by default, `family = zipoissonff` is appropriate but not `family = zipoisson`. Else set `family = zipoisson` and `which.linpred = 2`. To understand what is going on, do examine the constraint matrices of the fitted object, and reconcile this with Equations (4.3) to (4.5) of Yee and Hastie (2003).

The functions temporarily create a permanent data frame called `.grc.df` or `.rcim.df`, which used to be needed by `summary.rrvglm()`. Then these data frames are deleted before exiting the function. If an error occurs then the data frames may be present in the workspace.

Note

These functions set up the indicator variables etc. before calling `rrvglm` or `vglm`. The `...` is passed into `rrvglm.control` or `vglm.control`, This means, e.g., `Rank = 1` is default for `grc()`.

The data should be labelled with `rownames` and `colnames`. Setting `trace = TRUE` is recommended to monitor convergence. Using `criterion = "coefficients"` can result in slow convergence.

If `summary = TRUE` then `y` can be a "grc" object, in which case a summary can be returned. That is, `grc(y, summary = TRUE)` is equivalent to `summary(grc(y))`. It is not possible to plot a `grc(y, summary = TRUE)` or `rcim(y, summary = TRUE)` object.

Author(s)

Thomas W. Yee, with assistance from Alfian F. Hadi.

References

Yee, T. W. and Hastie, T. J. (2003). Reduced-rank vector generalized linear models. *Statistical Modelling*, **3**, 15–41.

Yee, T. W. and Hadi, A. F. (2014). Row-column interaction models, with an R implementation. *Computational Statistics*, **29**, 1427–1445.

Goodman, L. A. (1981). Association models and canonical correlation in the analysis of cross-classifications having ordered categories. *Journal of the American Statistical Association*, **76**, 320–334.

See Also

[rrvglm](#), [rrvglm.control](#), [rrvglm-class](#), [summary.grc](#), [moffset](#), [Rcim](#), [Select](#), [Qvar](#), [plotrcim0](#), [cqo](#), [multinomial](#), [alcoff](#), [crashi](#), [auc](#), [olym08](#), [olym12](#), [poissonff](#), [medpolish](#).

Examples

```
# Example 1: Undergraduate enrolments at Auckland University in 1990
fitted(grc1 <- grc(auc))
summary(grc1)

grc2 <- grc(auc, Rank = 2, Index.corner = c(2, 5))
fitted(grc2)
summary(grc2)

model3 <- rcim(auc, Rank = 1, fam = multinomial,
              M = ncol(auc)-1, cindex = 2:(ncol(auc)-1), trace = TRUE)
fitted(model3)
summary(model3)

# Median polish but not 100 percent reliable. Maybe call alaplace2()...
## Not run:
rcim0 <- rcim(auc, fam = alaplace1(tau = 0.5), trace=FALSE, maxit = 500)
round(fitted(rcim0), digits = 0)
round(100 * (fitted(rcim0) - auc) / auc, digits = 0) # Discrepancy
devar(rcim0)
round(coef(rcim0, matrix = TRUE), digits = 2)
Coef(rcim0, matrix = TRUE)
# constraints(rcim0)
names(constraints(rcim0))

# Compare with medpolish():
(med.a <- medpolish(auc))
fv <- med.a$overall + outer(med.a$row, med.a$col, "+")
round(100 * (fitted(rcim0) - fv) / fv) # Hopefully should be all 0s

## End(Not run)
```

```

# Example 2: 2012 Summer Olympic Games in London
## Not run: top10 <- head(olymp12, 10)
grc1.oly12 <- with(top10, grc(cbind(gold, silver, bronze)))
round(fitted(grc1.oly12))
round(resid(grc1.oly12, type = "response"), digits = 1) # Resp. resid
summary(grc1.oly12)
Coef(grc1.oly12)

## End(Not run)

# Example 3: UQO; see Yee and Hadi (2014)
## Not run:
n <- 100; p <- 5; S <- 10
pdata <- rcqo(n, p, S, es.opt = FALSE, eq.max = FALSE,
             eq.tol = TRUE, sd.latvar = 0.75) # Poisson counts
true.nu <- attr(pdata, "latvar") # The 'truth'; site scores
attr(pdata, "tolerances") # The 'truth'; tolerances

Y <- Select(pdata, "y", sort = FALSE) # Y matrix (n x S); the "y" vars
uqo.rcim1 <- rcim(Y, Rank = 1,
                str0 = NULL, # Delta covers entire n x M matrix
                iindex = 1:nrow(Y), # RRR covers the entire Y
                has.intercept = FALSE) # Suppress the intercept

# Plot 1
par(mfrow = c(2, 2))
plot(attr(pdata, "optimums"), Coef(uqo.rcim1)@A,
     col = "blue", type = "p", main = "(a) UQO optimums",
     xlab = "True optimums", ylab = "Estimated (UQO) optimums")
mylm <- lm(Coef(uqo.rcim1)@A ~ attr(pdata, "optimums"))
abline(coef = coef(mylm), col = "orange", lty = "dashed")

# Plot 2
fill.val <- NULL # Choose this for the new parameterization
plot(attr(pdata, "latvar"), c(fill.val, concoef(uqo.rcim1)),
     las = 1, col = "blue", type = "p", main = "(b) UQO site scores",
     xlab = "True site scores", ylab = "Estimated (UQO) site scores" )
mylm <- lm(c(fill.val, concoef(uqo.rcim1)) ~ attr(pdata, "latvar"))
abline(coef = coef(mylm), col = "orange", lty = "dashed")

# Plots 3 and 4
myform <- attr(pdata, "formula")
plut <- cqo(myform, family = poissonff,
           eq.tol = FALSE, trace = FALSE, data = pdata)
clut <- cqo(Select(pdata, "y", sort = FALSE) ~ scale(latvar(uqo.rcim1)),
           family = poissonff, eq.tol = FALSE, trace = FALSE, data = pdata)
lvplot(plut, lcol = 1:S, y = TRUE, pcol = 1:S, pch = 1:S, pcex = 0.5,
       main = "(c) CQO fitted to the original data",
       xlab = "Estimated (CQO) site scores")
lvplot(clut, lcol = 1:S, y = TRUE, pcol = 1:S, pch = 1:S, pcex = 0.5,

```

```

main = "(d) CQ0 fitted to the scaled UQ0 site scores",
xlab = "Estimated (UQ0) site scores")

## End(Not run)

```

gumbel

Gumbel Regression Family Function

Description

Maximum likelihood estimation of the 2-parameter Gumbel distribution.

Usage

```

gumbel(llocation = "identitylink", lscale = "loglink",
       iscale = NULL, R = NA, percentiles = c(95, 99),
       mpv = FALSE, zero = NULL)
gumbelff(llocation = "identitylink", lscale = "loglink",
         iscale = NULL, R = NA, percentiles = c(95, 99),
         zero = "scale", mpv = FALSE)

```

Arguments

llocation, lscale	Parameter link functions for μ and σ . See Links for more choices.
iscale	Numeric and positive. Optional initial value for σ . Recycled to the appropriate length. In general, a larger value is better than a smaller value. A NULL means an initial value is computed internally.
R	Numeric. Maximum number of values possible. See Details for more details.
percentiles	Numeric vector of percentiles used for the fitted values. Values should be between 0 and 100. This argument uses the argument R if assigned. If percentiles = NULL then the mean will be returned as the fitted values.
mpv	Logical. If mpv = TRUE then the <i>median predicted value</i> (MPV) is computed and returned as the (last) column of the fitted values. This argument is ignored if percentiles = NULL. See Details for more details.
zero	A vector specifying which linear/additive predictors are modelled as intercepts only. The value (possibly values) can be from the set {1, 2} corresponding respectively to μ and σ . By default all linear/additive predictors are modelled as a linear combination of the explanatory variables. See CommonVGAMffArguments for more details.

Details

The Gumbel distribution is a generalized extreme value (GEV) distribution with *shape* parameter $\xi = 0$. Consequently it is more easily estimated than the GEV. See [gev](#) for more details.

The quantity R is the maximum number of observations possible, for example, in the Venice data below, the top 10 daily values are recorded for each year, therefore $R = 365$ because there are about 365 days per year. The MPV is the value of the response such that the probability of obtaining a value greater than the MPV is 0.5 out of R observations. For the Venice data, the MPV is the sea level such that there is an even chance that the highest level for a particular year exceeds the MPV. If `mpv = TRUE` then the column labelled "MPV" contains the MPVs when `fitted()` is applied to the fitted object.

The formula for the mean of a response Y is $\mu + \sigma \times Euler$ where *Euler* is a constant that has value approximately equal to 0.5772. The formula for the percentiles are (if R is not given) $\mu - \sigma \times \log[-\log(P/100)]$ where P is the percentile argument value(s). If R is given then the percentiles are $\mu - \sigma \times \log[R(1 - P/100)]$.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Warning

When R is not given (the default) the fitted percentiles are that of the data, and not of the overall population. For example, in the example below, the 50 percentile is approximately the running median through the data, however, the data are the highest sea level measurements recorded each year (it therefore equates to the median predicted value or MPV).

Note

Like many other usual **VGAM** family functions, `gumbelff()` handles (independent) multiple responses.

`gumbel()` can handle more of a multivariate response, i.e., a matrix with more than one column. Each row of the matrix is sorted into descending order. Missing values in the response are allowed but require `na.action = na.pass`. The response matrix needs to be padded with any missing values. With a multivariate response one has a matrix y , say, where $y[, 2]$ contains the second order statistics, etc.

Author(s)

T. W. Yee

References

- Yee, T. W. and Stephenson, A. G. (2007). Vector generalized linear and additive extreme value models. *Extremes*, **10**, 1–19.
- Smith, R. L. (1986). Extreme value theory based on the r largest annual events. *Journal of Hydrology*, **86**, 27–43.

Rosen, O. and Cohen, A. (1996). Extreme percentile regression. In: Haerdle, W. and Schimek, M. G. (eds.), *Statistical Theory and Computational Aspects of Smoothing: Proceedings of the COMPSTAT '94 Satellite Meeting held in Semmering, Austria, 27–28 August 1994*, pp.200–214, Heidelberg: Physica-Verlag.

Coles, S. (2001). *An Introduction to Statistical Modeling of Extreme Values*. London: Springer-Verlag.

See Also

[rgumbel](#), [dgumbelII](#), [cens.gumbel](#), [guplot](#), [gev](#), [gevff](#), [venice](#).

Examples

```
# Example 1: Simulated data
gdata <- data.frame(y1 = rgumbel(n = 1000, loc = 100, scale = exp(1)))
fit1 <- vglm(y1 ~ 1, gumbelff(perc = NULL), data = gdata, trace = TRUE)
coef(fit1, matrix = TRUE)
Coef(fit1)
head(fitted(fit1))
with(gdata, mean(y1))

# Example 2: Venice data
(fit2 <- vglm(cbind(r1, r2, r3, r4, r5) ~ year, data = venice,
              gumbel(R = 365, mpv = TRUE), trace = TRUE))
head(fitted(fit2))
coef(fit2, matrix = TRUE)
sqrt(diag(vcov(summary(fit2)))) # Standard errors

# Example 3: Try a nonparametric fit -----
# Use the entire data set, including missing values
# Same as as.matrix(venice[, paste0("r", 1:10)]):
Y <- Select(venice, "r", sort = FALSE)
fit3 <- vgam(Y ~ s(year, df = 3), gumbel(R = 365, mpv = TRUE),
             data = venice, trace = TRUE, na.action = na.pass)
depar(fit3)[4:5, ] # NAs used to pad the matrix

## Not run: # Plot the component functions
par(mfrow = c(2, 3), mar = c(6, 4, 1, 2) + 0.3, xpd = TRUE)
plot(fit3, se = TRUE, lcol = "blue", scol = "limegreen", lty = 1,
     lwd = 2, slwd = 2, slty = "dashed")

# Quantile plot --- plots all the fitted values
qtplot(fit3, mpv = TRUE, lcol = c(1, 2, 5), tcol = c(1, 2, 5), lwd = 2,
       pcol = "blue", tadj = 0.1, ylab = "Sea level (cm)")

# Plot the 99 percentile only
year <- venice[["year"]]
matplot(year, Y, ylab = "Sea level (cm)", type = "n")
matpoints(year, Y, pch = "*", col = "blue")
lines(year, fitted(fit3)[, "99%"], lwd = 2, col = "orange")

# Check the 99 percentiles with a smoothing spline.
```

```
# Nb. (1-0.99) * 365 = 3.65 is approx. 4, meaning the 4th order
# statistic is approximately the 99 percentile.
plot(year, Y[, 4], ylab = "Sea level (cm)", type = "n",
      main = "Orange is 99 percentile, Green is a smoothing spline")
points(year, Y[, 4], pch = "4", col = "blue")
lines(year, fitted(fit3)[, "99%"], lty = 1, col = "orange")
lines(smooth.spline(year, Y[, 4], df = 4), col = "limegreen", lty = 2)

## End(Not run)
```

Gumbel-II

The Gumbel-II Distribution

Description

Density, cumulative distribution function, quantile function and random generation for the Gumbel-II distribution.

Usage

```
dgumbelII(x, scale = 1, shape, log = FALSE)
pgumbelII(q, scale = 1, shape, lower.tail = TRUE, log.p = FALSE)
qgumbelII(p, scale = 1, shape, lower.tail = TRUE, log.p = FALSE)
rgumbelII(n, scale = 1, shape)
```

Arguments

x, q	vector of quantiles.
p	vector of probabilities.
n	number of observations. Same as in runif .
log	Logical. If log = TRUE then the logarithm of the density is returned.
lower.tail, log.p	Same meaning as in pnorm or qnorm .
shape, scale	positive shape and scale parameters.

Details

See [gumbelII](#) for details.

Value

dgumbelII gives the density, pgumbelIII gives the cumulative distribution function, qgumbelIII gives the quantile function, and rgumbelIII generates random deviates.

Author(s)

T. W. Yee and Kai Huang

See Also

[gumbelII](#), [dgumbel](#).

Examples

```
probs <- seq(0.01, 0.99, by = 0.01)
Scale <- exp(1); Shape <- exp( 0.5);
max(abs(pgumbelII(qgumbelII(p = probs, shape = Shape, Scale),
                shape = Shape, Scale) - probs)) # Should be 0

## Not run: x <- seq(-0.1, 10, by = 0.01);
plot(x, dgumbelII(x, shape = Shape, Scale), type = "l", col = "blue",
     main = "Blue is density, orange is the CDF", las = 1,
     sub = "Red lines are the 10,20,...,90 percentiles",
     ylab = "", ylim = 0:1)
abline(h = 0, col = "blue", lty = 2)
lines(x, pgumbelII(x, shape = Shape, Scale), col = "orange")
probs <- seq(0.1, 0.9, by = 0.1)
Q <- qgumbelII(probs, shape = Shape, Scale)
lines(Q, dgumbelII(Q, Scale, Shape), col = "red", lty = 3, type = "h")
pgumbelII(Q, shape = Shape, Scale) - probs # Should be all zero
abline(h = probs, col = "red", lty = 3)
## End(Not run)
```

gumbelII

Gumbel-II Regression Family Function

Description

Maximum likelihood estimation of the 2-parameter Gumbel-II distribution.

Usage

```
gumbelII(lscale = "loglink", lshape = "loglink", iscale = NULL, ishape = NULL,
        probs.y = c(0.2, 0.5, 0.8), perc.out = NULL, imethod = 1,
        zero = "shape", nowarning = FALSE)
```

Arguments

nowarning	Logical. Suppress a warning?
lshape, lscale	Parameter link functions applied to the (positive) shape parameter (called <i>s</i> below) and (positive) scale parameter (called <i>b</i> below). See Links for more choices.
Parameter link functions applied to the	
ishape, iscale	Optional initial values for the shape and scale parameters.
imethod	See weibullR .

zero, probs.y	Details at CommonVGAMffArguments .
perc.out	If the fitted values are to be quantiles then set this argument to be the percentiles of these, e.g., 50 for median.

Details

The Gumbel-II density for a response Y is

$$f(y; b, s) = sy^{s-1} \exp[-(y/b)^s] / (b^s)$$

for $b > 0$, $s > 0$, $y > 0$. The cumulative distribution function is

$$F(y; b, s) = \exp[-(y/b)^s].$$

The mean of Y is $b\Gamma(1 - 1/s)$ (returned as the fitted values) when $s > 1$, and the variance is $b^2\Gamma(1 - 2/s)$ when $s > 2$. This distribution looks similar to [weibullR](#), and is due to Gumbel (1954).

This **VGAM** family function currently does not handle censored data. Fisher scoring is used to estimate the two parameters. Probably similar regularity conditions hold for this distribution compared to the Weibull distribution.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Note

See [weibullR](#). This **VGAM** family function handles multiple responses.

Author(s)

T. W. Yee

References

Gumbel, E. J. (1954). Statistical theory of extreme values and some practical applications. *Applied Mathematics Series*, volume 33, U.S. Department of Commerce, National Bureau of Standards, USA.

See Also

[dgumbelII](#), [gumbel](#), [gev](#).

Examples

```
gdata <- data.frame(x2 = runif(nn <- 1000))
gdata <- transform(gdata, heta1 = +1,
                  heta2 = -1 + 0.1 * x2,
                  ceta1 = 0,
                  ceta2 = 1)
```

```

gdata <- transform(gdata, shape1 = exp(heta1),
                  shape2 = exp(heta2),
                  scale1 = exp(ceta1),
                  scale2 = exp(ceta2))
gdata <- transform(gdata,
                  y1 = rgumbelIII(nn, scale = scale1, shape = shape1),
                  y2 = rgumbelIII(nn, scale = scale2, shape = shape2))

fit <- vglm(cbind(y1, y2) ~ x2,
           gumbelIII(zero = c(1, 2, 3)), data = gdata, trace = TRUE)
coef(fit, matrix = TRUE)
vcov(fit)
summary(fit)

```

gumbelUC

The Gumbel Distribution

Description

Density, distribution function, quantile function and random generation for the Gumbel distribution with location parameter location and scale parameter scale.

Usage

```

dgumbel(x, location = 0, scale = 1, log = FALSE)
pgumbel(q, location = 0, scale = 1, lower.tail = TRUE, log.p = FALSE)
qgumbel(p, location = 0, scale = 1, lower.tail = TRUE, log.p = FALSE)
rgumbel(n, location = 0, scale = 1)

```

Arguments

x, q	vector of quantiles.
p	vector of probabilities.
n	number of observations. If <code>length(n) > 1</code> then the length is taken to be the number required.
location	the location parameter μ . This is not the mean of the Gumbel distribution (see Details below).
scale	the scale parameter σ . This is not the standard deviation of the Gumbel distribution (see Details below).
log	Logical. If <code>log = TRUE</code> then the logarithm of the density is returned.
lower.tail, log.p	Same meaning as in punif or qunif .

Details

The Gumbel distribution is a special case of the *generalized extreme value* (GEV) distribution where the shape parameter $\xi = 0$. The latter has 3 parameters, so the Gumbel distribution has two. The Gumbel distribution function is

$$G(y) = \exp\left(-\exp\left[-\frac{y-\mu}{\sigma}\right]\right)$$

where $-\infty < y < \infty$, $-\infty < \mu < \infty$ and $\sigma > 0$. Its mean is

$$\mu - \sigma * \gamma$$

and its variance is

$$\sigma^2 * \pi^2 / 6$$

where γ is Euler's constant (which can be obtained as `-digamma(1)`).

See [gumbel](#), the **VGAM** family function for estimating the two parameters by maximum likelihood estimation, for formulae and other details. Apart from `n`, all the above arguments may be vectors and are recycled to the appropriate length if necessary.

Value

`dgumbel` gives the density, `pgumbel` gives the distribution function, `qgumbel` gives the quantile function, and `rgumbel` generates random deviates.

Note

The **VGAM** family function [gumbel](#) can estimate the parameters of a Gumbel distribution using maximum likelihood estimation.

Author(s)

T. W. Yee

References

Coles, S. (2001). *An Introduction to Statistical Modeling of Extreme Values*. London: Springer-Verlag.

See Also

[gumbel](#), [gumbelff](#), [gev](#), [dgomperz](#).

Examples

```
mu <- 1; sigma <- 2;
y <- rgumbel(n = 100, loc = mu, scale = sigma)
c(mean(y), mu - sigma * digamma(1)) # Sample and population means
c(var(y), sigma^2 * pi^2 / 6) # Sample and population variances

## Not run: x <- seq(-2.5, 3.5, by = 0.01)
```

```

loc <- 0; sigma <- 1
plot(x, dgumbel(x, loc, sigma), type = "l", col = "blue",
     main = "Blue is density, red is the CDF", ylim = c(0, 1),
     sub = "Purple are 5,10,...,95 percentiles", ylab = "", las = 1)
abline(h = 0, col = "blue", lty = 2)
lines(qgumbel(seq(0.05, 0.95, by = 0.05), loc, sigma),
      dgumbel(qgumbel(seq(0.05, 0.95, by = 0.05), loc, sigma), loc, sigma),
      col = "purple", lty = 3, type = "h")
lines(x, pgumbel(x, loc, sigma), type = "l", col = "red")
abline(h = 0, lty = 2)
## End(Not run)

```

guplot

Gumbel Plot

Description

Produces a Gumbel plot, a diagnostic plot for checking whether the data appears to be from a Gumbel distribution.

Usage

```

guplot(object, ...)
guplot.default(y, main = "Gumbel Plot",
              xlab = "Reduced data", ylab = "Observed data", type = "p", ...)
guplot.vlm(object, ...)

```

Arguments

<code>y</code>	A numerical vector. NAs etc. are not allowed.
<code>main</code>	Character. Overall title for the plot.
<code>xlab</code>	Character. Title for the x axis.
<code>ylab</code>	Character. Title for the y axis.
<code>type</code>	Type of plot. The default means points are plotted.
<code>object</code>	An object that inherits class "vlm", usually of class <code>vglm-class</code> or <code>vgam-class</code> .
<code>...</code>	Graphical argument passed into <code>plot</code> . See <code>par</code> for an exhaustive list. The arguments <code>xlim</code> and <code>ylim</code> are particularly useful.

Details

If Y has a Gumbel distribution then plotting the sorted values y_i versus the *reduced values* r_i should appear linear. The reduced values are given by

$$r_i = -\log(-\log(p_i))$$

where p_i is the i th plotting position, taken here to be $(i - 0.5)/n$. Here, n is the number of observations. Curvature upwards/downwards may indicate a Frchet/Weibull distribution, respectively. Outliers may also be detected using this plot.

The function `guplot` is generic, and `guplot.default` and `guplot.vlm` are some methods functions for Gumbel plots.

Value

A list is returned invisibly with the following components.

x	The reduced data.
y	The sorted y data.

Note

The Gumbel distribution is a special case of the GEV distribution with shape parameter equal to zero.

Author(s)

T. W. Yee

References

Coles, S. (2001). *An Introduction to Statistical Modeling of Extreme Values*. London: Springer-Verlag.

Gumbel, E. J. (1958). *Statistics of Extremes*. New York, USA: Columbia University Press.

See Also

[gumbel](#), [gumbelff](#), [gev](#), [venice](#).

Examples

```
## Not run: guplot(rnorm(500), las = 1) -> ii
names(ii)

guplot(with(venice, r1), col = "blue") # Venice sea levels data

## End(Not run)
```

has.interceptvlm	<i>Has a Fitted VGLM Got an Intercept Term?</i>
------------------	---

Description

Looks at the formula to see if it has an intercept term.

Usage

```
has.intercept(object, ...)
has.interceptvlm(object, form.number = 1, ...)
```

Arguments

object	A fitted model object.
form.number	Formula number, is 1 or 2. which correspond to the arguments formula and form2 respectively.
...	Arguments that are might be passed from one function to another.

Details

This methods function is a simple way to determine whether a fitted `vglm` object etc. has an intercept term or not. It is not entirely foolproof because one might suppress the intercept from the formula and then add in a variable in the formula that has a constant value.

Value

Returns a single logical.

Author(s)

Thomas W. Yee

See Also

[formulavlm](#), [termsvlm](#).

Examples

```
# Example: this is based on a glm example
counts <- c(18,17,15,20,10,20,25,13,12)
outcome <- gl(3, 1, 9); treatment <- gl(3, 3)
pdata <- data.frame(counts, outcome, treatment) # Better style
vglm.D93 <- vglm(counts ~ outcome + treatment, poissonff, data = pdata)
formula(vglm.D93)
term.names(vglm.D93)
responseName(vglm.D93)
has.intercept(vglm.D93)
```

Description

When complete, a suite of functions that can be used to compute some of the regression (leave-one-out deletion) diagnostics, for the VGLM class.

Usage

```

hatvalues(model, ...)
hatvaluesvml(model, type = c("diagonal", "matrix", "centralBlocks"), ...)
hatplot(model, ...)
hatplot.vlm(model, multiplier = c(2, 3), lty = "dashed",
             xlab = "Observation", ylab = "Hat values", ylim = NULL, ...)
dfbetavlm(model, maxit.new = 1,
           trace.new = FALSE,
           smallno = 1.0e-8, ...)

```

Arguments

model	an R object, typically returned by <code>vglm</code> .
type	Character. The default is the first choice, which is a $nM \times nM$ matrix. If type = "matrix" then the <i>entire</i> hat matrix is returned. If type = "centralBlocks" then n central $M \times M$ block matrices, in matrix-band format.
multiplier	Numeric, the multiplier. The usual rule-of-thumb is that values greater than two or three times the average leverage (at least for the linear model) should be checked.
lty, xlab, ylab, ylim	Graphical parameters, see <code>par</code> etc. The default of <code>ylim</code> is <code>c(0, max(hatvalues(model)))</code> which means that if the horizontal dashed lines cannot be seen then there are no particularly influential observations.
maxit.new, trace.new, smallno	Having <code>maxit.new = 1</code> will give a one IRLS step approximation from the ordinary solution (and no warnings!). Else having <code>maxit.new = 10</code> , say, should usually mean convergence will occur for all observations when they are removed one-at-a-time. Else having <code>maxit.new = 2</code> , say, should usually mean some lack of convergence will occur when observations are removed one-at-a-time. Setting <code>trace.new = TRUE</code> will produce some running output at each IRLS iteration and for each individual row of the model matrix. The argument <code>smallno</code> multiplies each value of the original prior weight (often unity); setting it identically to zero will result in an error, but setting a very small value effectively removes that observation.
...	further arguments, for example, graphical parameters for <code>hatplot.vlm()</code> .

Details

The invocation `hatvalues(vglmObject)` should return a $n \times M$ matrix of the diagonal elements of the hat (projection) matrix of a `vglm` object. To do this, the QR decomposition of the object is retrieved or reconstructed, and then straightforward calculations are performed.

The invocation `hatplot(vglmObject)` should plot the diagonal of the hat matrix for each of the M linear/additive predictors. By default, two horizontal dashed lines are added; hat values higher than these ought to be checked.

Note

It is hoped, soon, that the full suite of functions described at [influence.measures](#) will be written for VGLMs. This will enable general regression deletion diagnostics to be available for the entire VGLM class.

Author(s)

T. W. Yee.

See Also

[vglm](#), [cumulative](#), [influence.measures](#).

Examples

```
# Proportional odds model, p.179, in McCullagh and Nelder (1989)
pneumo <- transform(pneumo, let = log(exposure.time))
fit <- vglm(cbind(normal, mild, severe) ~ let, cumulative, data = pneumo)
hatvalues(fit) # n x M matrix, with positive values
all.equal(sum(hatvalues(fit)), fit@rank) # Should be TRUE
## Not run: par(mfrow = c(1, 2))
hatplot(fit, ylim = c(0, 1), las = 1, col = "blue")
## End(Not run)
```

hdeff

Hauck-Donner Effects: A Detection Test for Wald Tests

Description

A detection test for the Hauck-Donner effect on each regression coefficient of a VGLM regression or 2 x 2 table.

Usage

```
hdeff(object, ...)
hdeff.vglm(object, derivative = NULL, se.arg = FALSE,
           subset = NULL, theta0 = 0, hstep = 0.005,
           fd.only = FALSE, ...)
hdeff.numeric(object, byrow = FALSE, ...)
hdeff.matrix(object, ...)
```

Arguments

object Usually a [vglm](#) object. Although only a limited number of family functions have an analytical solution to the HDE detection test ([binomialff](#), [borel.tanner](#), [cumulative](#), [erlang](#), [felix](#), [lindley](#), [poissonff](#), [topple](#), [uninormal](#), [zipoissonff](#), and [zipoisson](#); hopefully some more will be implemented in the short future!)

the finite-differences (FDs) method can be applied to almost all **VGAM** family functions to get a numerical solution.

Alternatively `object` may represent a 2×2 table of *positive* counts. If so, then the first row corresponds to $x_2 = 0$ (baseline group) and the second row $x_2 = 1$. The first column corresponds to $y = 0$ (failure) and the second column $y = 1$ (success).

Another alternative is that `object` is a numerical vector of length 4, representing a 2×2 table of *positive* counts. If so then it is fed into `hdeff.matrix` using the argument `byrow`, which matches `matrix`. See the examples below.

<code>derivative</code>	Numeric. Either 1 or 2. Currently only a few models having one linear predictor are handled analytically for <code>derivative = 2</code> , e.g., <code>binomialff</code> , <code>poissonff</code> . However, the numerical method can return the first two derivatives for almost all models.
<code>se.arg</code>	Logical. If TRUE then the derivatives of the standard errors are returned as well, because usually the derivatives of the Wald statistics are of central interest. Requires <code>derivative</code> to be assigned the value 1 or 2 for this argument to operate.
<code>subset</code>	Logical or vector of indices, to select the regression coefficients of interest. The default is to select all coefficients. Recycled if necessary if logical. If numeric then they should comprise elements from $1 : \text{length}(\text{coef}(\text{object}))$. This argument can be useful for computing the derivatives of a Cox regression (<code>coxph</code>) fitted using artificially created Poisson data; then there are many coefficients that are effectively nuisance parameters.
<code>theta0</code>	Numeric. Vector recycled to the necessary length which is the number of regression coefficients. The null hypotheses for the regression coefficients are that they equal those respective values, and the alternative hypotheses are all two-sided. It is not recommended that argument <code>subset</code> be used if a vector of values is assigned here because <code>theta0[subset]</code> is implied and might not work.
<code>hstep</code>	Positive numeric and recycled to length 2; it is the so-called <i>step size</i> when using finite-differences and is often called h in the calculus literature, e.g., $f'(x)$ is approximately $(f(x+h) - f(x))/h$. For the 2nd-order partial derivatives, there are two step sizes and hence this argument is recycled to length 2. The default is to have the same values. The 1st-order derivatives use the first value only. It is recommended that a few values of this argument be tried because values of the first and second derivatives can vary accordingly. If any values are too large then the derivatives may be inaccurate; and if too small then the derivatives may be unstable and subject to too much round-off/cancellation error (in fact it may create an error or a NA).
<code>fd.only</code>	Logical; if TRUE then finite-differences are used to estimate the derivatives even if an analytical solution has been coded. By default, finite-differences will be used when an analytical solution has not been implemented. It is possible that NAs are returned. If so, and if <code>fd.only = FALSE</code> , then a warning is issued and a recursive call is made with <code>fd.only = TRUE</code> —this is more likely to return an answer without any NAs.
<code>byrow</code>	Logical; fed into <code>matrix</code> if <code>object</code> is a vector of length 4 so that there are two choices in the order of the elements.
<code>...</code>	currently unused but may be used in the future for further arguments passed into the other methods functions.

Details

Almost all of statistical inference based on the likelihood assumes that the parameter estimates are located in the interior of the parameter space. The nonregular case of being located on the boundary is not considered very much and leads to very different results from the regular case. Practically, an important question is: how close is close to the boundary? One might answer this as: the parameter estimates are too close to the boundary when the Hauck-Donner effect (HDE) is present, whereby the Wald statistic becomes aberrant.

Hauck and Donner (1977) first observed an aberration of the Wald test statistic not monotonically increasing as a function of increasing distance between the parameter estimate and the null value. This "disturbing" and "undesirable" underappreciated effect has since been observed in other regression models by various authors. This function computes the first, and possibly second, derivative of the Wald statistic for each regression coefficient. A negative value of the first derivative is indicative of the HDE being present. More information can be obtained from [hdeffsev](#) regarding HDE severity: there may be none, faint, weak, moderate, strong and extreme amounts of HDE present.

In general, most models have derivatives that are computed numerically using finite-difference approximations. The reason is that it takes a lot of work to program in the analytical solution (this includes a few very common models, such as [poissonff](#) and [binomialff](#), where the first two derivatives have been implemented).

Value

By default this function returns a labelled logical vector; a TRUE means the HDE is affirmative for that coefficient (negative slope). Hence ideally all values are FALSE. Any TRUE values suggests that the MLE is too near the boundary of the parameter space, and that the p-value for that regression coefficient is biased upwards. When present a highly significant variable might be deemed nonsignificant, and thus the HDE can create havoc for variable selection. If the HDE is present then more accurate p-values can generally be obtained by conducting a likelihood ratio test (see [lrt.stat.vlm](#)) or Rao's score test (see [score.stat.vlm](#)); indeed the default of [wald.stat.vlm](#) does not suffer from the HDE.

Setting `deriv = 1` returns a numerical vector of first derivatives of the Wald statistics. Setting `deriv = 2` returns a 2-column matrix of first and second derivatives of the Wald statistics. Then setting `se.arg = TRUE` returns an additional 1 or 2 columns.

Some 2nd derivatives are NA if only a partial analytic solution has been programmed in.

For those **VGAM** family functions whose HDE test has not yet been implemented explicitly (the vast majority of them), finite-difference approximations to the derivatives will be used—see the arguments `hstep` and `fd.only` for getting some control on them.

Note

The function [summaryvglm](#) conducts the HDE detection test if possible and prints out a line at the bottom if the HDE is detected for some regression coefficients. By "if possible", only a few family functions are exempt and they have an `infos` slot with component `hadof = FALSE`; such as [normal.vcm](#), [rec.normal](#) because it uses the BFGS-IRLS method for computing the working weights. For these few a NULL is returned by [hdeff](#).

If the second derivatives are of interest then it is recommended that `crit = "c"` be added to the fitting so that a slightly more accurate model results (usually one more IRLS iteration). This is

because the FD approximation is very sensitive to values of the working weights, so they need to be computed accurately. Occasionally, if the coefficient is close to 0, then its Wald statistic's second derivative may be unusually large in magnitude (this could be due to something such as roundoff error).

This function is currently under development and may change a little in the short future. For HDE severity measures see [hdeffsev](#).

Author(s)

Thomas W. Yee.

References

Hauck, J. W. W. and A. Donner (1977). Wald's test as applied to hypotheses in logit analysis. *Journal of the American Statistical Association*, **72**, 851–853.

Yee, T. W. (2022). On the Hauck-Donner effect in Wald tests: Detection, tipping points and parameter space characterization, *Journal of the American Statistical Association*, **117**, 1763–1774. [doi:10.1080/01621459.2021.1886936](https://doi.org/10.1080/01621459.2021.1886936).

Yee, T. W. (2021). Some new results concerning the Hauck-Donner effect. *Manuscript in preparation*.

See Also

[summaryvglm](#), [hdeffsev](#), [alogitlink](#), [asinlink](#), [vglm](#), [lrt.stat](#), [score.stat](#), [wald.stat](#), [confintvglm](#), [profilevglm](#).

Examples

```
pneumo <- transform(pneumo, let = log(exposure.time))
fit <- vglm(cbind(normal, mild, severe) ~ let, data = pneumo,
           trace = TRUE, crit = "c", # Get some more accuracy
           cumulative(reverse = TRUE, parallel = TRUE))
cumulative()@infos()$hadof # Analytical solution implemented
hdeff(fit)
hdeff(fit, deriv = 1) # Analytical solution
hdeff(fit, deriv = 2) # It is a partial analytical solution
hdeff(fit, deriv = 2, se.arg = TRUE,
      fd.only = TRUE) # All derivatives solved numerically by FDs

# 2 x 2 table of counts
R0 <- 25; N0 <- 100 # Hauck Donner (1977) data set
mymat <- c(N0-R0, R0, 8, 92) # HDE present
(mymat <- matrix(mymat, 2, 2, byrow = TRUE))
hdeff(mymat)
hdeff(c(mymat)) # Input is a vector
hdeff(c(t(mymat)), byrow = TRUE) # Reordering of the data
```

hdeffsev

*Hauck-Donner Effect: Severity Measures***Description**

Computes the severity of the Hauck-Donner effect for each regression coefficient of a VGML regression.

Usage

```
hdeffsev(x, y, dy, ddy, allofit = FALSE, eta0 = 0, COPS0 = eta0,
         severity.table = c("None", "Faint", "Weak",
                             "Moderate", "Strong", "Extreme", "Undetermined"))
hdeffsev2(x, y, dy, ddy, allofit = FALSE, ndepends = FALSE,
          eta0 = 0,
          severity.table = c("None", "Faint", "Weak",
                              "Moderate", "Strong", "Extreme",
                              "Undetermined")[if (ndepends) TRUE else
                                                c(1, 4, 6, 7)], tol0 = 0.1)
```

Arguments

<code>x, y</code>	Numeric vectors; <code>x</code> are the estimates (sorted), and <code>y</code> are the signed Wald statistics.
<code>dy, ddy</code>	Numeric vectors; the first and second derivatives of the Wald statistics. They can be computed by hdeff .
<code>allofit</code>	Logical. If TRUE then other quantities are returned in a list. The default is a vector with elements selected from the argument <code>severity.table</code> .
<code>severity.table</code>	Character vector with 6 values plus the last value for initialization. Usually users should not assign anything to this argument.
<code>eta0</code>	Numeric. The hypothesized value. The default is appropriate for most symmetric binomialff links, and also for poissonff regression with the natural parameter.
<code>ndepends</code>	Logical. Use boundaries that depend on the sample size n ? For <code>hdeffsev2</code> the default is to use boundaries that do not depend on n , and consequently there are fewer severity measures. These do not use the normal or tangent lines; instead they are based only on the signs of <code>dy</code> and <code>ddy</code> .
<code>COPS0</code>	Numeric. See Yee (2023).
<code>tol0</code>	Numeric. Any estimate whose absolute value is less than <code>tol0</code> is assigned the first value of the argument <code>severity.table</code> , i.e., none. This is to handle a singularity at the origin: the estimates might be extremely close to 0.

Details

Note: The function hdeffsev has a bug or two in it but they should be fixed later this year (2024).

Function hdeffsev is currently rough-and-ready. It is possible to use the first two derivatives obtained from hdeff to categorize the severity of the the Hauck-Donner effect (HDE). It is effectively assumed that, starting at the origin and going right, the curve is made up of a convex segment followed by a concave segment and then the convex segment. Midway in the concave segment the first derivative is 0, and beyond that the HDE is really manifest because the derivative remains negative.

For "None" the estimate lies on the convex part of the curve near the origin, hence there is very little HDE at all.

For "Weak" the estimate lies on the concave part of the curve but the Wald statistic is still increasing as estimate gets away from 0, hence it is only a mild form of the HDE.

For "Moderate", "Strong" and "Extreme" the Wald statistic is decreasing as the estimate gets away from $\eta\theta$, hence it really does exhibit the HDE. It is recommended that `lrt.stat` be used to compute LRT p-values, as they do not suffer from the HDE.

Value

By default this function (hdeffsev) returns a labelled vector with elements selected from `severity.table`. If `allofit = TRUE` then Yee (2022) gives details about some of the other list components, e.g., a quantity called zeta is the normal line projected onto the x-axis, and its first derivative gives additional information about the position of the estimate along the curve.

Note

These functions are likely to change in the short future because it is experimental and far from complete. Improvements are intended.

The severity measures ideally should be based on tangent lines rather than normal lines so that the boundaries are independent of the sample size n . Hence such boundaries differ a little from Yee (2022) which had a mixture of such.

The functions were written specifically for `binomialff`, but they should work for some other family functions.

Currently, in order for "Strong" to be assigned correctly, at least one such value is needed on the LHS and/or RHS each. From those, two other boundary points are obtained so that it creates two intervals.

Author(s)

Thomas W. Yee.

References

Yee, T. W. (2022). On the Hauck-Donner effect in Wald tests: Detection, tipping points and parameter space characterization, *Journal of the American Statistical Association*, **117**, 1763–1774. doi:10.1080/01621459.2021.1886936.

Yee, T. W. (2023). Some new results concerning the Wald tests and the parameter space. *In review*.

See Also

[seglines](#), [hdeff](#).

Examples

```
deg <- 4 # myfun is a function that approximates the HDE
myfun <- function(x, deriv = 0) switch(as.character(deriv),
  '0' = x^deg * exp(-x),
  '1' = (deg * x^(deg-1) - x^deg) * exp(-x),
  '2' = (deg*(deg-1)*x^(deg-2) - 2*deg*x^(deg-1) + x^deg)*exp(-x))

xgrid <- seq(0, 10, length = 101)
ansm <- hdeffsev(xgrid, myfun(xgrid), myfun(xgrid, deriv = 1),
  myfun(xgrid, deriv = 2), allofit = TRUE)

digg <- 4
cbind(severity = ansm$sev,
  fun = round(myfun(xgrid), digg),
  deriv1 = round(myfun(xgrid, deriv = 1), digg),
  deriv2 = round(myfun(xgrid, deriv = 2), digg),
  zderiv1 = round(1 + (myfun(xgrid, deriv = 1))^2 +
    myfun(xgrid, deriv = 2) * myfun(xgrid), digg))
```

hormone

Hormone Assay Data

Description

A hormone assay data set from Carroll and Ruppert (1988).

Usage

```
data(hormone)
```

Format

A data frame with 85 observations on the following 2 variables.

X a numeric vector, suitable as the x-axis in a scatter plot. The reference method.

Y a numeric vector, suitable as the y-axis in a scatter plot. The test method.

Details

The data is given in Table 2.4 of Carroll and Ruppert (1988), and was downloaded from <http://www.stat.tamu.edu/~carroll> prior to 2019. The book describes the data as follows. The data are the results of two assay methods for hormone data; the scale of the data as presented is not particularly meaningful, and the original source of the data refused permission to divulge further information. As in a similar example of Leurgans (1980), the old or reference method is being used to predict the new or test method. The overall goal is to see whether we can reproduce the test-method measurements with the reference-method measurements. Thus calibration might be of interest for the data.

References

- Carroll, R. J. and Ruppert, D. (1988). *Transformation and Weighting in Regression*. New York, USA: Chapman & Hall.
- Leurgans, S. (1980). Evaluating laboratory measurement techniques. *Biostatistics Casebook*. Eds.: Miller, R. G. Jr., and Efron, B. and Brown, B. W. Jr., and Moses, L. New York, USA: Wiley.
- Yee, T. W. (2014). Reduced-rank vector generalized linear models with two linear predictors. *Computational Statistics and Data Analysis*, **71**, 889–902.

See Also

[uninormal](#), [rrvglm](#).

Examples

```
## Not run:
data(hormone)
summary(hormone)

modelI <- rrvglm(Y ~ 1 + X, data = hormone, trace = TRUE,
                uninormal(zero = NULL, lsd = "identitylink", imethod = 2))

# Alternative way to fit modelI
modelI.other <- vglm(Y ~ 1 + X, data = hormone, trace = TRUE,
                   uninormal(zero = NULL, lsd = "identitylink"))

# Inferior to modelI
modelII <- vglm(Y ~ 1 + X, data = hormone, trace = TRUE,
              family = uninormal(zero = NULL))

logLik(modelI)
logLik(modelII) # Less than logLik(modelI)

# Reproduce the top 3 equations on p.65 of Carroll and Ruppert (1988).
# They are called Equations (1)--(3) here.

# Equation (1)
hormone <- transform(hormone, rX = 1 / X)
clist <- list("(Intercept)" = diag(2), X = diag(2), rX = rbind(0, 1))
fit1 <- vglm(Y ~ 1 + X + rX, family = uninormal(zero = NULL),
            constraints = clist, data = hormone, trace = TRUE)
coef(fit1, matrix = TRUE)
summary(fit1) # Actually, the intercepts do not seem significant
plot(Y ~ X, hormone, col = "blue")
lines(fitted(fit1) ~ X, hormone, col = "orange")

# Equation (2)
fit2 <- rrvglm(Y ~ 1 + X, uninormal(zero = NULL), hormone, trace = TRUE)
coef(fit2, matrix = TRUE)
plot(Y ~ X, hormone, col = "blue")
lines(fitted(fit2) ~ X, hormone, col = "red")
```

```

# Add +- 2 SEs
lines(fitted(fit2) + 2 * exp(predict(fit2)[, "loglink(sd)"]) ~ X,
      hormone, col = "orange")
lines(fitted(fit2) - 2 * exp(predict(fit2)[, "loglink(sd)"]) ~ X,
      hormone, col = "orange")

# Equation (3)
# Does not fit well because the loglink link for the mean is not good.
fit3 <- rrvglm(Y ~ 1 + X, maxit = 300, data = hormone, trace = TRUE,
              uninormal(lmean = "loglink", zero = NULL))
coef(fit3, matrix = TRUE)
plot(Y ~ X, hormone, col = "blue") # Does not look okay.
lines(exp(predict(fit3)[, 1]) ~ X, hormone, col = "red")
# Add +- 2 SEs
lines(fitted(fit3) + 2 * exp(predict(fit3)[, "loglink(sd)"]) ~ X,
      hormone, col = "orange")
lines(fitted(fit3) - 2 * exp(predict(fit3)[, "loglink(sd)"]) ~ X,
      hormone, col = "orange")

## End(Not run)

```

hspider

Hunting Spider Data

Description

Abundance of hunting spiders in a Dutch dune area.

Usage

```
data(hspider)
```

Format

A data frame with 28 observations (sites) on the following 18 variables.

WaterCon Log percentage of soil dry mass.

BareSand Log percentage cover of bare sand.

FallTwig Log percentage cover of fallen leaves and twigs.

CoveMoss Log percentage cover of the moss layer.

CoveHerb Log percentage cover of the herb layer.

ReflLux Reflection of the soil surface with cloudless sky.

Alopacce Abundance of *Alopecosa accentuata*.

Alopcune Abundance of *Alopecosa cuneata*.

Alopfabr Abundance of *Alopecosa fabrilis*.

Arctlute Abundance of *Arctosa lutetiana*.

Arctperi Abundance of *Arctosa perita*.
Auloalbi Abundance of *Aulonia albimana*.
Pardlugu Abundance of *Pardosa lugubris*.
Pardmont Abundance of *Pardosa monticola*.
Pardnigr Abundance of *Pardosa nigriceps*.
Pardpull Abundance of *Pardosa pullata*.
Trocterr Abundance of *Trochosa terricola*.
Zoraspin Abundance of *Zora spinimana*.

Details

The data, which originally came from Van der Aart and Smeek-Enserink (1975) consists of abundances (numbers trapped over a 60 week period) and 6 environmental variables. There were 28 sites.

This data set has been often used to illustrate ordination, e.g., using canonical correspondence analysis (CCA). In the example below, the data is used for constrained quadratic ordination (CQO; formerly called canonical Gaussian ordination or CGO), a numerically intensive method that has many superior qualities. See [cqo](#) for details.

References

Van der Aart, P. J. M. and Smeek-Enserink, N. (1975). Correlations between distributions of hunting spiders (Lycosidae, Ctenidae) and environmental characteristics in a dune area. *Netherlands Journal of Zoology*, **25**, 1–45.

Examples

```
summary(hspider)

## Not run:
# Standardize the environmental variables:
hspider[, 1:6] <- scale(subset(hspider, select = WaterCon:ReflLux))

# Fit a rank-1 binomial CAO
hsbin <- hspider # Binary species data
hsbin[, -(1:6)] <- as.numeric(hsbin[, -(1:6)] > 0)
set.seed(123)
ahsb1 <- cao(cbind(Alopcune, Arctlute, Auloalbi, Zoraspin) ~
             WaterCon + ReflLux,
             family = binomialff(multiple.responses = TRUE),
             df1.nl = 2.2, Bestof = 3, data = hsbin)
par(mfrow = 2:1, las = 1)
lvplot(ahsb1, type = "predictors", llwd = 2,
       ylab = "logitlink(p)", lcol = 1:9)
persp(ahsb1, rug = TRUE, col = 1:10, lwd = 2)
coef(ahsb1)

## End(Not run)
```

huber2

*Huber's Least Favourable Distribution Family Function***Description**

M-estimation of the two parameters of Huber's least favourable distribution. The one parameter case is also implemented.

Usage

```
huber1(llocation = "identitylink", k = 0.862, imethod = 1)
huber2(llocation = "identitylink", lscale = "loglink",
       k = 0.862, imethod = 1, zero = "scale")
```

Arguments

llocation, lscale	Link functions applied to the location and scale parameters. See Links for more choices.
k	Tuning constant. See rhuber for more information.
imethod, zero	See CommonVGAMffArguments for information. The default value of zero means the scale parameter is modelled as intercept-only.

Details

Huber's least favourable distribution family function is popular for resistant/robust regression. The center of the distribution is normal and its tails are double exponential.

By default, the mean is the first linear/additive predictor (returned as the fitted values; this is the location parameter), and the log of the scale parameter is the second linear/additive predictor. The Fisher information matrix is diagonal; Fisher scoring is implemented.

The **VGAM** family function `huber1()` estimates only the location parameter. It assumes a scale parameter of unit value.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Note

Warning: actually, `huber2()` may be erroneous since the first derivative is not continuous when there are two parameters to estimate. `huber1()` is fine in this respect.

The response should be univariate.

Author(s)

T. W. Yee. Help was given by Arash Ardalan.

References

Huber, P. J. and Ronchetti, E. (2009). *Robust Statistics*, 2nd ed. New York: Wiley.

See Also

[rhuber](#), [uninormal](#), [laplace](#), [CommonVGAMffArguments](#).

Examples

```
set.seed(1231); NN <- 30; coef1 <- 1; coef2 <- 10
hdata <- data.frame(x2 = sort(runif(NN)))
hdata <- transform(hdata, y = rhuber(NN, mu = coef1 + coef2 * x2))

hdata$x2[1] <- 0.0 # Add an outlier
hdata$y[1] <- 10

fit.huber2 <- vglm(y ~ x2, huber2(imethod = 3), hdata, trace = TRUE)
fit.huber1 <- vglm(y ~ x2, huber1(imethod = 3), hdata, trace = TRUE)

coef(fit.huber2, matrix = TRUE)
summary(fit.huber2)

## Not run: # Plot the results
plot(y ~ x2, data = hdata, col = "blue", las = 1)
lines(fitted(fit.huber2) ~ x2, data = hdata, col = "darkgreen", lwd = 2)

fit.lm <- lm(y ~ x2, hdata) # Compare to a LM:
lines(fitted(fit.lm) ~ x2, data = hdata, col = "lavender", lwd = 3)

# Compare to truth:
lines(coef1 + coef2 * x2 ~ x2, data = hdata, col = "orange",
      lwd = 2, lty = "dashed")

legend("bottomright", legend = c("truth", "huber", "lm"),
      col = c("orange", "darkgreen", "lavender"),
      lty = c("dashed", "solid", "solid"), lwd = c(2, 2, 3))
## End(Not run)
```

Huggins89.t1

Table 1 of Huggins (1989)

Description

Simulated capture data set for the linear logistic model depending on an occasion covariate and an individual covariate for 10 trapping occasions and 20 individuals.

Usage

```
data(Huggins89table1)
data(Huggins89.t1)
```

Format

The format is a data frame.

Details

Table 1 of Huggins (1989) gives this toy data set. Note that variables t_1, \dots, t_{10} are occasion-specific variables. They correspond to the response variables y_1, \dots, y_{10} which have values 1 for capture and 0 for not captured.

Both Huggins89table1 and Huggins89.t1 are identical. The latter used variables beginning with z, not t, and may be withdrawn very soon.

References

Huggins, R. M. (1989). On the statistical analysis of capture experiments. *Biometrika*, **76**, 133–140.

Examples

```
## Not run:
Huggins89table1 <-
  transform(Huggins89table1, x3.tij = t01,
            T02 = t02, T03 = t03, T04 = t04, T05 = t05, T06 = t06,
            T07 = t07, T08 = t08, T09 = t09, T10 = t10)
small.table1 <- subset(Huggins89table1,
  y01 + y02 + y03 + y04 + y05 + y06 + y07 + y08 + y09 + y10 > 0)
# fit.tbh is the bottom equation on p.133.
# It is a M_tbh model.
fit.tbh <-
  vglm(cbind(y01, y02, y03, y04, y05, y06, y07, y08, y09, y10) ~
    x2 + x3.tij,
    xij = list(x3.tij ~ t01 + t02 + t03 + t04 + t05 +
              t06 + t07 + t08 + t09 + t10 +
              T02 + T03 + T04 + T05 + T06 +
              T07 + T08 + T09 + T10 - 1),
    posbernoulli.tb(parallel.t = TRUE ~ x2 + x3.tij),
    data = small.table1, trace = TRUE,
    form2 = ~ x2 + x3.tij +
            t01 + t02 + t03 + t04 + t05 + t06 +
            t07 + t08 + t09 + t10 +
            T02 + T03 + T04 + T05 + T06 +
            T07 + T08 + T09 + T10)

# These results differ a bit from Huggins (1989), probably because
# two animals had to be removed here (they were never caught):
coef(fit.tbh) # First element is the behavioural effect
sqrt(diag(vcov(fit.tbh))) # SEs
constraints(fit.tbh, matrix = TRUE)
```



```

summary(fit.tbh, presid = FALSE)
fit.tbh@extra$N.hat      # Estimate of the population site N; cf. 20.86
fit.tbh@extra$SE.N.hat  # Its standard error; cf. 1.87 or 4.51

fit.th <-
  vglm(cbind(y01, y02, y03, y04, y05, y06, y07, y08, y09, y10) ~ x2,
        posbernoulli.t, data = small.table1, trace = TRUE)
coef(fit.th)
constraints(fit.th)
coef(fit.th, matrix = TRUE) # M_th model
summary(fit.th, presid = FALSE)
fit.th@extra$N.hat      # Estimate of the population size N
fit.th@extra$SE.N.hat  # Its standard error

fit.bh <-
  vglm(cbind(y01, y02, y03, y04, y05, y06, y07, y08, y09, y10) ~ x2,
        posbernoulli.b(I2 = FALSE), data = small.table1, trace = TRUE)
coef(fit.bh)
constraints(fit.bh)
coef(fit.bh, matrix = TRUE) # M_bh model
summary(fit.bh, presid = FALSE)
fit.bh@extra$N.hat
fit.bh@extra$SE.N.hat

fit.h <-
  vglm(cbind(y01, y02, y03, y04, y05, y06, y07, y08, y09, y10) ~ x2,
        posbernoulli.b, data = small.table1, trace = TRUE)
coef(fit.h, matrix = TRUE) # M_h model (version 1)
coef(fit.h)
summary(fit.h, presid = FALSE)
fit.h@extra$N.hat
fit.h@extra$SE.N.hat

Fit.h <-
  vglm(cbind(y01, y02, y03, y04, y05, y06, y07, y08, y09, y10) ~ x2,
        posbernoulli.t(parallel.t = TRUE ~ x2),
        data = small.table1, trace = TRUE)
coef(Fit.h)
coef(Fit.h, matrix = TRUE) # M_h model (version 2)
summary(Fit.h, presid = FALSE)
Fit.h@extra$N.hat
Fit.h@extra$SE.N.hat

## End(Not run)

```

Description

The hunua data frame has 392 rows and 18 columns. Altitude is explanatory, and there are binary responses (presence/absence = 1/0 respectively) for 17 plant species.

Usage

```
data(hunua)
```

Format

This data frame contains the following columns:

agaaus Agathis australis, or Kauri
beitaw Beilschmiedia tawa, or Tawa
corlae Corynocarpus laevigatus
cyadea Cyathea dealbata
cyamed Cyathea medullaris
daccup Dacrydium cupressinum
dacdac Dacrycarpus dacrydioides
eladen Elaeocarpus dentatus
hedarb Hedycarya arborea
hohpop Species name unknown
kniexc Knightia excelsa, or Rewarewa
kuneri Kunzea ericoides
lepsco Leptospermum scoparium
metrob Metrosideros robusta
neslan Nestegis lanceolata
rhosap Rhopalostylis sapida
vitluc Vitex lucens, or Puriri
altitude meters above sea level

Details

These were collected from the Hunua Ranges, a small forest in southern Auckland, New Zealand. At 392 sites in the forest, the presence/absence of 17 plant species was recorded, as well as the altitude. Each site was of area size $200m^2$.

Source

Dr Neil Mitchell, University of Auckland.

See Also

[waitakere](#).

Examples

```
# Fit a GAM using vgam() and compare it with the Waitakere Ranges one
fit.h <- vgam(agaaus ~ s(altitude, df = 2), binomialff, data = hunua)
## Not run:
plot(fit.h, se = TRUE, lcol = "orange", scol = "orange",
      llwd = 2, slwd = 2, main = "Orange is Hunua, Blue is Waitakere")
## End(Not run)
head(predict(fit.h, hunua, type = "response"))

fit.w <- vgam(agaaus ~ s(altitude, df = 2), binomialff, data = waitakere)
## Not run:
plot(fit.w, se = TRUE, lcol = "blue", scol = "blue", add = TRUE)
## End(Not run)
head(predict(fit.w, hunua, type = "response")) # Same as above?
```

Hurea

The Husler-Reiss Angular Surface Distribution

Description

Density for the Husler-Reiss angular surface distribution.

Usage

```
dhurea(x, shape, log = FALSE)
```

Arguments

x	Same as Uniform .
shape	the positive (shape) parameter. It is often called <i>lambda</i> and it might not be a shape parameter at all.
log	Logical. If log = TRUE then the logarithm of the density is returned.

Details

See [hurea](#), the **VGAM** family function for estimating the (shape) parameter *s* by maximum likelihood estimation, for the formula of the probability density function.

Value

dhurea gives the density.

Warning

The cases $x == 0$, $x == 1$, $shape == 0$ and $shape == Inf$ may not be handled correctly.

Note

Difficulties are encountered as the shape parameter approaches 0 with respect to [integrate](#) because the density converges to a degenerate distribution with probability mass at 0 and 1. That is, when s is around 0.5 the density is “u” shaped and the area around the endpoints becomes concentrated at the two points. See the examples below. Approximately, the density is “u” shaped for $s < 1$ and unimodal for $s > 2$.

Author(s)

T. W. Yee

See Also

[hurea](#).

Examples

```
integrate(dhurea, 0, 1, shape = 0.20) # Incorrect
integrate(dhurea, 0, 1, shape = 0.35) # struggling but okay
## Not run: x <- seq(0, 1, length = 501)
par(mfrow = c(2, 2))
plot(x, dhurea(x, 0.7), col = "blue", type = "l")
plot(x, dhurea(x, 1.1), col = "blue", type = "l")
plot(x, dhurea(x, 1.4), col = "blue", type = "l")
plot(x, dhurea(x, 3.0), col = "blue", type = "l")

## End(Not run)
```

hurea

Husler-Reiss Angular Surface Distribution Family Function

Description

Estimating the parameter of the Husler-Reiss angular surface distribution by maximum likelihood estimation.

Usage

```
hurea(lshape = "loglink", zero = NULL, nrfs = 1,
      gshape = exp(3 * ppoints(5) - 1), parallel = FALSE)
```

Arguments

`lshape`, `gshape` Details at [CommonVGAMffArguments](#).
`nrfs`, `zero`, `parallel`
 Details at [CommonVGAMffArguments](#).

Details

The Husler-Reiss angular surface distribution has a probability density function that can be written

$$f(y; s) = (s / (4 * \text{sqrt}(2 * \text{pi}) * y(1 - y)^2)) \exp(-(2 + s^2 * \text{logity})^2 / [8s^2])$$

for $0 < y < 1$ and positive shape parameter s . The mean of Y is currently unknown to me, as well as its quantiles. Hence s is currently returned as the fitted values. Fisher-scoring is implemented.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Note

This **VGAM** family function handles multiple responses. It may struggle and/or fail when s is close to 0. Some comments about "u"-shaped versus unimodal densities accommodated by this distribution are at [dhurea](#).

Author(s)

T. W. Yee

References

Mhalla, L. and de Carvalho, M. and Chavez-Demoulin, V. (2019). Regression-type models for extremal dependence. *Scandinavian Journal of Statistics*, **46**, 1141–1167.

See Also

[hurea](#).

Examples

```
nn <- 100; set.seed(1)
hdata <- data.frame(x2 = runif(nn))
hdata <-
  transform(hdata, # Cannot generate proper random variates!
    y1 = rbeta(nn, shape1 = 0.5, shape2 = 0.5), # "U" shaped
    y2 = rnorm(nn, 0.65, sd = exp(-3 - 4 * x2)))
# Multiple responses:
hfit <- vglm(cbind(y1, y2) ~ x2, hurea, hdata, trace = TRUE)
coef(hfit, matrix = TRUE)
summary(hfit)
```

hyperg *Hypergeometric Family Function*

Description

Family function for a hypergeometric distribution where either the number of white balls or the total number of white and black balls are unknown.

Usage

```
hyperg(N = NULL, D = NULL, lprob = "logitlink", iprob = NULL)
```

Arguments

N	Total number of white and black balls in the urn. Must be a vector with positive values, and is recycled, if necessary, to the same length as the response. One of N and D must be specified.
D	Number of white balls in the urn. Must be a vector with positive values, and is recycled, if necessary, to the same length as the response. One of N and D must be specified.
lprob	Link function for the probabilities. See Links for more choices.
iprob	Optional initial value for the probabilities. The default is to choose initial values internally.

Details

Consider the scenario from [dhyper](#) where there are $N = m + n$ balls in an urn, where m are white and n are black. A simple random sample (i.e., *without* replacement) of k balls is taken. The response here is the sample *proportion* of white balls. In this document, N is $N = m + n$, D is m (for the number of “defectives”, in quality control terminology, or equivalently, the number of marked individuals). The parameter to be estimated is the population proportion of white balls, viz. $prob = m/(m + n)$.

Depending on which one of N and D is inputted, the estimate of the other parameter can be obtained from the equation $prob = m/(m + n)$, or equivalently, $prob = D/N$. However, the log-factorials are computed using [lgamma](#) and both m and n are not restricted to being integer. Thus if an integer N is to be estimated, it will be necessary to evaluate the likelihood function at integer values about the estimate, i.e., at `trunc(Nhat)` and `ceiling(Nhat)` where Nhat is the (real) estimate of N .

Value

An object of class “`vglmff`” (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), [vgam](#), [rrvglm](#), [cqo](#), and [cao](#).

Warning

No checking is done to ensure that certain values are within range, e.g., $k \leq N$.

Note

The response can be of one of three formats: a factor (first level taken as success), a vector of proportions of success, or a 2-column matrix (first column = successes) of counts. The argument weights in the modelling function can also be specified. In particular, for a general vector of proportions, you will need to specify weights because the number of trials is needed.

Author(s)

Thomas W. Yee

References

Forbes, C., Evans, M., Hastings, N. and Peacock, B. (2011). *Statistical Distributions*, Hoboken, NJ, USA: John Wiley and Sons, Fourth edition.

See Also

[dhyper](#), [binomialff](#).

Examples

```
nn <- 100
m <- 5 # Number of white balls in the population
k <- rep(4, len = nn) # Sample sizes
n <- 4 # Number of black balls in the population
y <- rhyper(nn = nn, m = m, n = n, k = k)
yprop <- y / k # Sample proportions

# N is unknown, D is known. Both models are equivalent:
fit <- vglm(cbind(y,k-y) ~ 1, hyperg(D = m), trace = TRUE, crit = "c")
fit <- vglm(yprop ~ 1, hyperg(D = m), weight = k, trace = TRUE, crit = "c")

# N is known, D is unknown. Both models are equivalent:
fit <- vglm(cbind(y, k-y) ~ 1, hyperg(N = m+n), trace = TRUE, crit = "l")
fit <- vglm(yprop ~ 1, hyperg(N = m+n), weight = k, trace = TRUE, crit = "l")

coef(fit, matrix = TRUE)
Coef(fit) # Should be equal to the true population proportion
unique(m / (m+n)) # The true population proportion
fit@extra
head(fitted(fit))
summary(fit)
```

hypersecant

Hyperbolic Secant Regression Family Function

Description

Estimation of the parameter of the hyperbolic secant distribution.

Usage

```
hypersecant(link.theta = extlogitlink(min = -pi/2, max = pi/2),
            init.theta = NULL)
hypersecant01(link.theta = extlogitlink(min = -pi/2, max = pi/2),
              init.theta = NULL)
```

Arguments

`link.theta` Parameter link function applied to the parameter θ . See [Links](#) for more choices.

`init.theta` Optional initial value for θ . If failure to converge occurs, try some other value. The default means an initial value is determined internally.

Details

The probability density function of the hyperbolic secant distribution is given by

$$f(y; \theta) = \exp(\theta y + \log(\cos(\theta))) / (2 \cosh(\pi y/2)),$$

for parameter $-\pi/2 < \theta < \pi/2$ and all real y . The mean of Y is $\tan(\theta)$ (returned as the fitted values). Morris (1982) calls this model NEF-HS (Natural Exponential Family-Hyperbolic Secant). It is used to generate NEFs, giving rise to the class of NEF-GHS (G for Generalized).

Another parameterization is used for `hypersecant01()`: let $Y = (\text{logit}U)/\pi$. Then this uses

$$f(u; \theta) = (\cos(\theta)/\pi) \times u^{-0.5+\theta/\pi} \times (1-u)^{-0.5-\theta/\pi},$$

for parameter $-\pi/2 < \theta < \pi/2$ and $0 < u < 1$. Then the mean of U is $0.5 + \theta/\pi$ (returned as the fitted values) and the variance is $(\pi^2 - 4\theta^2)/(8\pi^2)$.

For both parameterizations Newton-Raphson is same as Fisher scoring.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Author(s)

T. W. Yee

References

Jorgensen, B. (1997). *The Theory of Dispersion Models*. London: Chapman & Hall.

Morris, C. N. (1982). Natural exponential families with quadratic variance functions. *The Annals of Statistics*, **10**(1), 65–80.

See Also

[gensh](#), [extlogitlink](#).

Examples

```

hdata <- data.frame(x2 = rnorm(nn <- 200))
hdata <- transform(hdata, y = rnorm(nn)) # Not very good data!
fit1 <- vglm(y ~ x2, hypersecant, hdata, trace = TRUE, crit = "c")
coef(fit1, matrix = TRUE)
fit1@misc$earg

# Not recommended:
fit2 <- vglm(y ~ x2, hypersecant(link = "identitylink"), hdata)
coef(fit2, matrix = TRUE)
fit2@misc$earg

```

Hzeta

*Haight's Zeta Distribution***Description**

Density, distribution function, quantile function and random generation for Haight's zeta distribution with parameter shape.

Usage

```

dhzeta(x, shape, log = FALSE)
phzeta(q, shape, log.p = FALSE)
qhzeta(p, shape)
rhzeta(n, shape)

```

Arguments

x, q, p, n	Same meaning as runif .
shape	The positive shape parameter. Called α below.
log, log.p	Same meaning as in pnorm or qnorm .

Details

The probability function is

$$f(x) = (2x - 1)^{-\alpha} - (2x + 1)^{-\alpha},$$

where $\alpha > 0$ and $x = 1, 2, \dots$

Value

dhzeta gives the density, phzeta gives the distribution function, qhzeta gives the quantile function, and rhzeta generates random deviates.

Note

Given some response data, the **VGAM** family function [hzeta](#) estimates the parameter shape.

Author(s)

T. W. Yee and Kai Huang

See Also

[hzeta](#), [zeta](#), [zetaff](#), [simulate.vlm](#).

Examples

```
dhzeta(1:20, 2.1)
rhzeta(20, 2.1)

round(1000 * dhzeta(1:8, 2))
table(rhzeta(1000, 2))

## Not run:  shape <- 1.1; x <- 1:10
plot(x, dhzeta(x, shape = shape), type = "h", ylim = 0:1,
     sub = paste("shape =", shape), las = 1, col = "blue",
     ylab = "Probability", lwd = 2,
     main = "Haight's zeta: blue = density; orange = CDF")
lines(x+0.1, phzeta(x, shape = shape), col = "orange", lty = 3, lwd = 2,
     type = "h")

## End(Not run)
```

hzeta

Haight's Zeta Family Function

Description

Estimating the parameter of Haight's zeta distribution

Usage

```
hzeta(lshape = "logloglink", ishape = NULL, nsimEIM = 100)
```

Arguments

`lshape` Parameter link function for the parameter, called α below. See [Links](#) for more choices. Here, a log-log link keeps the parameter greater than one, meaning the mean is finite.

`ishape, nsimEIM` See [CommonVGAMffArguments](#) for more information.

Details

The probability function is

$$f(y) = (2y - 1)^{(-\alpha)} - (2y + 1)^{(-\alpha)},$$

where the parameter $\alpha > 0$ and $y = 1, 2, \dots$. The function `dhzeta` computes this probability function. The mean of Y , which is returned as fitted values, is $(1 - 2^{-\alpha})\zeta(\alpha)$ provided $\alpha > 1$, where ζ is Riemann's zeta function. The mean is a decreasing function of α . The mean is infinite if $\alpha \leq 1$, and the variance is infinite if $\alpha \leq 2$.

Value

An object of class "vglmff" (see `vglmff-class`). The object is used by modelling functions such as `vglm` and `vgam`.

Author(s)

T. W. Yee

References

Johnson N. L., Kemp, A. W. and Kotz S. (2005). *Univariate Discrete Distributions*, 3rd edition, pp.533–4. Hoboken, New Jersey: Wiley.

See Also

[Hzeta](#), [zeta](#), [zetaff](#), [loglog](#), [simulate.vlm](#).

Examples

```
shape <- exp(exp(-0.1)) # The parameter
hdata <- data.frame(y = rhzeta(n = 1000, shape))
fit <- vglm(y ~ 1, hzeta, data = hdata, trace = TRUE, crit = "coef")
coef(fit, matrix = TRUE)
Coef(fit) # Useful for intercept-only models; should be same as shape
c(with(hdata, mean(y)), head(fitted(fit), 1))
summary(fit)
```

iam

Index from Array to Matrix

Description

Maps the elements of an array containing symmetric positive-definite matrices to a matrix with sufficient columns to hold them (called matrix-band format.)

Usage

```
iam(j, k, M, both = FALSE, diag = TRUE)
```

Arguments

j	Usually an integer from the set $\{1:M\}$ giving the row number of an element. However, the argument can also be a vector of length M , for selecting an entire row or column, e.g., <code>iam(1:M, 1, M)</code> or <code>iam(1, 1:M, M)</code> .
k	An integer from the set $\{1:M\}$ giving the column number of an element.
M	The number of linear/additive predictors. This is the dimension of each positive-definite symmetric matrix.
both	Logical. Return both the row and column indices? See below for more details.
diag	Logical. Return the indices for the diagonal elements? If FALSE then only the strictly upper triangular part of the matrix elements are used.

Details

Suppose we have n symmetric positive-definite square matrices, each M by M , and these are stored in an array of dimension $c(n,M,M)$. Then these can be more compactly represented by a matrix of dimension $c(n,K)$ where K is an integer between M and $M*(M+1)/2$ inclusive. The mapping between these two representations is given by this function. It firstly enumerates by the diagonal elements, followed by the band immediately above the diagonal, then the band above that one, etc. The last element is $(1, M)$. This function performs the mapping from elements (j, k) of symmetric positive-definite square matrices to the columns of another matrix representing such. This is called the *matrix-band* format and is used by the **VGAM** package.

Value

This function has a dual purpose depending on the value of `both`. If `both = FALSE` then the column number corresponding to the j - k element of the matrix is returned. If `both = TRUE` then j and k are ignored and a list with the following components are returned.

<code>row.index</code>	The row indices of the upper triangular part of the matrix (This may or may not include the diagonal elements, depending on the argument <code>diagonal</code>).
<code>col.index</code>	The column indices of the upper triangular part of the matrix (This may or may not include the diagonal elements, depending on the argument <code>diagonal</code>).

Note

This function is used in the `weight` slot of many **VGAM** family functions (see [vglmff-class](#)), especially those whose M is determined by the data, e.g., [dirichlet](#), [multinomial](#).

Author(s)

T. W. Yee

See Also

[vglmff-class](#).

Examples

```

iam(1, 2, M = 3) # The 4th coln represents elt (1,2) of a 3x3 matrix
iam(NULL, NULL, M = 3, both = TRUE) # Return the row & column indices

dirichlet()@weight

M <- 4
temp1 <- iam(NA, NA, M = M, both = TRUE)
mat1 <- matrix(NA, M, M)
mat1[cbind(temp1$row, temp1$col)] = 1:length(temp1$row)
mat1 # More commonly used

temp2 <- iam(NA, NA, M = M, both = TRUE, diag = FALSE)
mat2 <- matrix(NA, M, M)
mat2[cbind(temp2$row, temp2$col)] = 1:length(temp2$row)
mat2 # Rarely used

```

identitylink

*Identity Link Function***Description**

Computes the identity transformation, including its inverse and the first two derivatives.

Usage

```

identitylink(theta, inverse = FALSE, deriv = 0, short = TRUE,
             tag = FALSE)
negidentitylink(theta, inverse = FALSE, deriv = 0, short = TRUE,
                tag = FALSE)

```

Arguments

theta Numeric or character. See below for further details.
inverse, deriv, short, tag
 Details at [Links](#).

Details

The identity link function $g(\theta) = \theta$ should be available to every parameter estimated by the **VGAM** library. However, it usually results in numerical problems because the estimates lie outside the permitted range. Consequently, the result may contain Inf, -Inf, NA or NaN.

The function `negidentitylink` is the negative-identity link function and corresponds to $g(\theta) = -\theta$. This is useful for some models, e.g., in the literature supporting the `gevff` function it seems that half of the authors use $\xi = -k$ for the shape parameter and the other half use k instead of ξ .

Value

For `identitylink()`: for `deriv = 0`, the identity of θ , i.e., θ when `inverse = FALSE`, and if `inverse = TRUE` then $1/\theta$. For `deriv = 1`, then the function returns $d\eta/d\theta$ as a function of θ if `inverse = FALSE`, else if `inverse = TRUE` then it returns the reciprocal.

For `negidentitylink()`: the results are similar to `identitylink()` except for a sign change in most cases.

Author(s)

Thomas W. Yee

References

McCullagh, P. and Nelder, J. A. (1989). *Generalized Linear Models*, 2nd ed. London: Chapman & Hall.

See Also

[Links](#), [loglink](#), [logitlink](#), [probitlink](#), [powerlink](#).

Examples

```
identitylink((-5):5)
identitylink((-5):5, deriv = 1)
identitylink((-5):5, deriv = 2)
negidentitylink((-5):5)
negidentitylink((-5):5, deriv = 1)
negidentitylink((-5):5, deriv = 2)
```

Influence

Influence Function (S4 generic) of a Fitted Model

Description

Returns a matrix containing the influence function of a fitted model, e.g., a "vglm" object.

Usage

```
Influence(object, ...)
Influence.vglm(object, weighted = TRUE, ...)
```

Arguments

<code>object</code>	an object, especially that of class "vglm"—see vglm-class . Currently other classes such as "vgam" are not yet implemented.
<code>weighted</code>	Logical. Include the prior weights? Currently only TRUE is accepted. This might change in the future and/or the default value might change.
<code>...</code>	any additional arguments such as to allow or disallow the prior weights.

Details

Influence functions are useful in fields such as sample survey theory, e.g., **survey**, **svyVGAM**. For each $i = 1, \dots, n$, the formula is approximately $-IU$ where I is the weighted Fisher information matrix and U is the i th score vector.

Value

An n by p vlm matrix.

Warning

This function is currently experimental and defaults may change. Use with caution! The functions here should not be confused with [lm.influence](#).

See Also

[vglm](#), [vglm-class](#), [survey](#).

Examples

```
pneumo <- transform(pneumo, let = log(exposure.time))
fit <- vglm(cbind(normal, mild, severe) ~ let, acat, data = pneumo)
coef(fit) # 8-vector
Influence(fit) # 8 x 4
all(abs(colSums(Influence(fit))) < 1e-6) # TRUE
```

 inv.binomial

Inverse Binomial Distribution Family Function

Description

Estimates the two parameters of an inverse binomial distribution by maximum likelihood estimation.

Usage

```
inv.binomial(lrho = extlogitlink(min = 0.5, max = 1),
            llambda = "loglink", irho = NULL, ilambda = NULL, zero = NULL)
```

Arguments

`lrho`, `llambda` Link function for the ρ and λ parameters. See [Links](#) for more choices.

`irho`, `ilambda` Numeric. Optional initial values for ρ and λ .

`zero` See [CommonVGAMffArguments](#).

Details

The inverse binomial distribution of Yanagimoto (1989) has density function

$$f(y; \rho, \lambda) = \frac{\lambda \Gamma(2y + \lambda)}{\Gamma(y + 1) \Gamma(y + \lambda + 1)} \{\rho(1 - \rho)\}^y \rho^\lambda$$

where $y = 0, 1, 2, \dots$ and $\frac{1}{2} < \rho < 1$, and $\lambda > 0$. The first two moments exist for $\rho > \frac{1}{2}$; then the mean is $\lambda(1 - \rho)/(2\rho - 1)$ (returned as the fitted values) and the variance is $\lambda\rho(1 - \rho)/(2\rho - 1)^3$. The inverse binomial distribution is a special case of the generalized negative binomial distribution of Jain and Consul (1971). It holds that $Var(Y) > E(Y)$ so that the inverse binomial distribution is overdispersed compared with the Poisson distribution.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Note

This **VGAM** family function only works reasonably well with intercept-only models. Good initial values are needed; if convergence failure occurs use `irho` and/or `ilambda`.

Some elements of the working weight matrices use the expected information matrix while other elements use the observed information matrix. Yet to do: using the mean and the reciprocal of λ results in an EIM that is diagonal.

Author(s)

T. W. Yee

References

Yanagimoto, T. (1989). The inverse binomial distribution as a statistical model. *Communications in Statistics: Theory and Methods*, **18**, 3625–3633.

Jain, G. C. and Consul, P. C. (1971). A generalized negative binomial distribution. *SIAM Journal on Applied Mathematics*, **21**, 501–513.

Jorgensen, B. (1997). *The Theory of Dispersion Models*. London: Chapman & Hall

See Also

[negbinomial](#), [poissonff](#).

Examples

```
idata <- data.frame(y = rnbinom(n <- 1000, mu = exp(3), size = exp(1)))
fit <- vglm(y ~ 1, inv.binomial, data = idata, trace = TRUE)
with(idata, c(mean(y), head(fitted(fit), 1)))
summary(fit)
coef(fit, matrix = TRUE)
Coef(fit)
```



```
sum(weights(fit)) # Sum of the prior weights
sum(weights(fit, type = "work")) # Sum of the working weights
```

Inv.gaussian

The Inverse Gaussian Distribution

Description

Density, distribution function and random generation for the inverse Gaussian distribution.

Usage

```
dinv.gaussian(x, mu, lambda, log = FALSE)
pinv.gaussian(q, mu, lambda)
rinv.gaussian(n, mu, lambda)
```

Arguments

x, q	vector of quantiles.
n	number of observations. If <code>length(n) > 1</code> then the length is taken to be the number required.
mu	the mean parameter.
lambda	the λ parameter.
log	Logical. If <code>log = TRUE</code> then the logarithm of the density is returned.

Details

See [inv.gaussianff](#), the **VGAM** family function for estimating both parameters by maximum likelihood estimation, for the formula of the probability density function.

Value

`dinv.gaussian` gives the density, `pinv.gaussian` gives the distribution function, and `rinv.gaussian` generates random deviates.

Note

Currently `qinv.gaussian` is unavailable.

Author(s)

T. W. Yee

References

Johnson, N. L. and Kotz, S. and Balakrishnan, N. (1994). *Continuous Univariate Distributions*, 2nd edition, Volume 1, New York: Wiley.

Taraldsen, G. and Lindqvist, B. H. (2005). The multiple roots simulation algorithm, the inverse Gaussian distribution, and the sufficient conditional Monte Carlo method. *Preprint Statistics No. 4/2005*, Norwegian University of Science and Technology, Trondheim, Norway.

See Also

[inv.gaussianff](#), [waldff](#).

Examples

```
## Not run: x <- seq(-0.05, 4, len = 300)
plot(x, dinv.gaussian(x, mu = 1, lambda = 1), type = "l",
      col = "blue", las = 1, main =
        "blue is density, orange is cumulative distribution function")
abline(h = 0, col = "gray", lty = 2)
lines(x, pinv.gaussian(x, mu = 1, lambda = 1), type = "l", col = "orange")
## End(Not run)
```

inv.gaussianff

Inverse Gaussian Distribution Family Function

Description

Estimates the two parameters of the inverse Gaussian distribution by maximum likelihood estimation.

Usage

```
inv.gaussianff(lmu = "loglink", llambda = "loglink",
              imethod = 1, ilambda = NULL,
              parallel = FALSE, ishrinkage = 0.99, zero = NULL)
```

Arguments

`lmu, llambda` Parameter link functions for the μ and λ parameters. See [Links](#) for more choices.

`ilambda, parallel`

See [CommonVGAMffArguments](#) for more information. If `parallel = TRUE` then the constraint is not applied to the intercept.

`imethod, ishrinkage, zero`

See [CommonVGAMffArguments](#) for information.

Details

The standard (“canonical”) form of the inverse Gaussian distribution has a density that can be written as

$$f(y; \mu, \lambda) = \sqrt{\lambda/(2\pi y^3)} \exp(-\lambda(y - \mu)^2/(2y\mu^2))$$

where $y > 0$, $\mu > 0$, and $\lambda > 0$. The mean of Y is μ and its variance is μ^3/λ . By default, $\eta_1 = \log(\mu)$ and $\eta_2 = \log(\lambda)$. The mean is returned as the fitted values. This **VGAM** family function can handle multiple responses (inputted as a matrix).

Value

An object of class “vglmff” (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), [rrvglm](#) and [vgam](#).

Note

The inverse Gaussian distribution can be fitted (to a certain extent) using the usual GLM framework involving a scale parameter. This family function is different from that approach in that it estimates both parameters by full maximum likelihood estimation.

Author(s)

T. W. Yee

References

- Johnson, N. L. and Kotz, S. and Balakrishnan, N. (1994). *Continuous Univariate Distributions*, 2nd edition, Volume 1, New York: Wiley.
- Forbes, C., Evans, M., Hastings, N. and Peacock, B. (2011). *Statistical Distributions*, Hoboken, NJ, USA: John Wiley and Sons, Fourth edition.

See Also

[Inv.gaussian](#), [waldff](#), [bisa](#).

The R package **SuppDists** has several functions for evaluating the density, distribution function, quantile function and generating random numbers from the inverse Gaussian distribution.

Examples

```
idata <- data.frame(x2 = runif(nn <- 1000))
idata <- transform(idata, mymu = exp(2 + 1 * x2),
                  Lambda = exp(2 + 1 * x2))
idata <- transform(idata, y = rinv.gaussian(nn, mu = mymu, Lambda))
fit1 <- vglm(y ~ x2, inv.gaussianff, data = idata, trace = TRUE)
rrig <- rrvglm(y ~ x2, inv.gaussianff, data = idata, trace = TRUE)
coef(fit1, matrix = TRUE)
coef(rrig, matrix = TRUE)
Coef(rrig)
summary(fit1)
```

 Inv.lomax

The Inverse Lomax Distribution

Description

Density, distribution function, quantile function and random generation for the inverse Lomax distribution with shape parameter p and scale parameter $scale$.

Usage

```
dinv.lomax(x, scale = 1, shape2.p, log = FALSE)
pinv.lomax(q, scale = 1, shape2.p, lower.tail = TRUE, log.p = FALSE)
qinv.lomax(p, scale = 1, shape2.p, lower.tail = TRUE, log.p = FALSE)
rinv.lomax(n, scale = 1, shape2.p)
```

Arguments

<code>x, q</code>	vector of quantiles.
<code>p</code>	vector of probabilities.
<code>n</code>	number of observations. If <code>length(n) > 1</code> , the length is taken to be the number required.
<code>shape2.p</code>	shape parameter.
<code>scale</code>	scale parameter.
<code>log</code>	Logical. If <code>log = TRUE</code> then the logarithm of the density is returned.
<code>lower.tail, log.p</code>	Same meaning as in pnorm or qnorm .

Details

See [inv.lomax](#), which is the **VGAM** family function for estimating the parameters by maximum likelihood estimation.

Value

`dinv.lomax` gives the density, `pinv.lomax` gives the distribution function, `qinv.lomax` gives the quantile function, and `rinv.lomax` generates random deviates.

Note

The inverse Lomax distribution is a special case of the 4-parameter generalized beta II distribution.

Author(s)

T. W. Yee

References

Kleiber, C. and Kotz, S. (2003). *Statistical Size Distributions in Economics and Actuarial Sciences*, Hoboken, NJ, USA: Wiley-Interscience.

See Also

[inv.lomax](#), [genbetaII](#).

Examples

```
idata <- data.frame(y = rinv.lomax(n = 1000, exp(2), exp(1)))
fit <- vglm(y ~ 1, inv.lomax, idata, trace = TRUE, crit = "coef")
coef(fit, matrix = TRUE)
Coef(fit)
```

inv.lomax

Inverse Lomax Distribution Family Function

Description

Maximum likelihood estimation of the 2-parameter inverse Lomax distribution.

Usage

```
inv.lomax(lscale = "loglink", lshape2.p = "loglink", iscale = NULL,
  ishape2.p = NULL, imethod = 1, gscale = exp(-5:5),
  gshape2.p = exp(-5:5), probs.y = c(0.25, 0.5, 0.75),
  zero = "shape2.p")
```

Arguments

lscale, lshape2.p

Parameter link functions applied to the (positive) parameters b , and p . See [Links](#) for more choices.

iscale, ishape2.p, imethod, zero

See [CommonVGAMffArguments](#) for information. For `imethod = 2` a good initial value for `ishape2.p` is needed to obtain a good estimate for the other parameter.

gscale, gshape2.p

See [CommonVGAMffArguments](#) for information.

probs.y

See [CommonVGAMffArguments](#) for information.

Details

The 2-parameter inverse Lomax distribution is the 4-parameter generalized beta II distribution with shape parameters $a = q = 1$. It is also the 3-parameter Dagum distribution with shape parameter $a = 1$, as well as the beta distribution of the second kind with $q = 1$. More details can be found in Kleiber and Kotz (2003).

The inverse Lomax distribution has density

$$f(y) = py^{p-1}/[b^p\{1 + y/b\}^{p+1}]$$

for $b > 0$, $p > 0$, $y \geq 0$. Here, b is the scale parameter scale, and p is a shape parameter. The mean does not seem to exist; the *median* is returned as the fitted values. This family function handles multiple responses.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Note

See the notes in [genbetaII](#).

Author(s)

T. W. Yee

References

Kleiber, C. and Kotz, S. (2003). *Statistical Size Distributions in Economics and Actuarial Sciences*, Hoboken, NJ, USA: Wiley-Interscience.

See Also

[inv.lomax](#), [genbetaII](#), [betaII](#), [dagum](#), [sinmad](#), [fisk](#), [lomax](#), [paralogistic](#), [inv.paralogistic](#), [simulate.vlm](#).

Examples

```
idata <- data.frame(y = rinv.lomax(2000, sc = exp(2), exp(1)))
fit <- vglm(y ~ 1, inv.lomax, data = idata, trace = TRUE)
fit <- vglm(y ~ 1, inv.lomax(iscale = exp(3)), data = idata,
           trace = TRUE, epsilon = 1e-8, crit = "coef")
coef(fit, matrix = TRUE)
Coef(fit)
summary(fit)
```

 Inv.paralogistic *The Inverse Paralogistic Distribution*

Description

Density, distribution function, quantile function and random generation for the inverse paralogistic distribution with shape parameters a and p , and scale parameter $scale$.

Usage

```
dinv.paralogistic(x, scale = 1, shape1.a, log = FALSE)
pinv.paralogistic(q, scale = 1, shape1.a, lower.tail = TRUE,
                  log.p = FALSE)
qinv.paralogistic(p, scale = 1, shape1.a, lower.tail = TRUE,
                  log.p = FALSE)
rinv.paralogistic(n, scale = 1, shape1.a)
```

Arguments

<code>x, q</code>	vector of quantiles.
<code>p</code>	vector of probabilities.
<code>n</code>	number of observations. If <code>length(n) > 1</code> , the length is taken to be the number required.
<code>shape1.a</code>	shape parameter.
<code>scale</code>	scale parameter.
<code>log</code>	Logical. If <code>log = TRUE</code> then the logarithm of the density is returned.
<code>lower.tail, log.p</code>	Same meaning as in pnorm or qnorm .

Details

See [inv.paralogistic](#), which is the **VGAM** family function for estimating the parameters by maximum likelihood estimation.

Value

`dinv.paralogistic` gives the density, `pinv.paralogistic` gives the distribution function, `qinv.paralogistic` gives the quantile function, and `rinv.paralogistic` generates random deviates.

Note

The inverse paralogistic distribution is a special case of the 4-parameter generalized beta II distribution.

Author(s)

T. W. Yee

References

Kleiber, C. and Kotz, S. (2003). *Statistical Size Distributions in Economics and Actuarial Sciences*, Hoboken, NJ, USA: Wiley-Interscience.

See Also

[inv.paralogistic](#), [genbetaII](#).

Examples

```
idata <- data.frame(y = rinv.paralogistic(3000, exp(1), sc = exp(2)))
fit <- vglm(y ~ 1, inv.paralogistic(lss = FALSE, ishape1.a = 2.1),
           data = idata, trace = TRUE, crit = "coef")
coef(fit, matrix = TRUE)
Coef(fit)
```

inv.paralogistic	<i>Inverse Paralogistic Distribution Family Function</i>
------------------	--

Description

Maximum likelihood estimation of the 2-parameter inverse paralogistic distribution.

Usage

```
inv.paralogistic(lscale = "loglink", lshape1.a = "loglink",
                iscale = NULL, ishape1.a = NULL, imethod = 1,
                lss = TRUE, gscale = exp(-5:5),
                gshape1.a = seq(0.75, 4, by = 0.25), probs.y = c(0.25, 0.5,
                0.75), zero = "shape")
```

Arguments

`lss` See [CommonVGAMffArguments](#) for important information.

`lshape1.a`, `lscale` Parameter link functions applied to the (positive) parameters `a` and `scale`. See [Links](#) for more choices.

`iscale`, `ishape1.a`, `imethod`, `zero` See [CommonVGAMffArguments](#) for information. For `imethod = 2` a good initial value for `ishape1.a` is needed to obtain a good estimate for the other parameter.

`gscale`, `gshape1.a` See [CommonVGAMffArguments](#) for information.

`probs.y` See [CommonVGAMffArguments](#) for information.

Details

The 2-parameter inverse paralogistic distribution is the 4-parameter generalized beta II distribution with shape parameter $q = 1$ and $a = p$. It is the 3-parameter Dagum distribution with $a = p$. More details can be found in Kleiber and Kotz (2003).

The inverse paralogistic distribution has density

$$f(y) = a^2 y^{a^2-1} / [b^{a^2} \{1 + (y/b)^a\}^{a+1}]$$

for $a > 0$, $b > 0$, $y \geq 0$. Here, b is the scale parameter scale, and a is the shape parameter. The mean is

$$E(Y) = b \Gamma(a + 1/a) \Gamma(1 - 1/a) / \Gamma(a)$$

provided $a > 1$; these are returned as the fitted values. This family function handles multiple responses.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Note

See the notes in [genbetaII](#).

Author(s)

T. W. Yee

References

Kleiber, C. and Kotz, S. (2003). *Statistical Size Distributions in Economics and Actuarial Sciences*, Hoboken, NJ, USA: Wiley-Interscience.

See Also

[Inv.paralogistic](#), [genbetaII](#), [betaII](#), [dagum](#), [sinmad](#), [fisk](#), [inv.lomax](#), [lomax](#), [paralogistic](#), [simulate.vlm](#).

Examples

```
## Not run:
idata <- data.frame(y = rinv.paralogistic(3000, exp(1), sc = exp(2)))
fit <- vglm(y ~ 1, inv.paralogistic(lss = FALSE), idata, trace = TRUE)
fit <- vglm(y ~ 1, inv.paralogistic(imethod = 2, ishape1.a = 4),
           data = idata, trace = TRUE, crit = "coef")
coef(fit, matrix = TRUE)
Coef(fit)
summary(fit)
## End(Not run)
```

`is.buggy`*Does the Fitted Object Suffer from a Known Bug?*

Description

Checks to see if a fitted object suffers from some known bug.

Usage

```
is.buggy(object, ...)  
is.buggy.vlm(object, each.term = FALSE, ...)
```

Arguments

<code>object</code>	A fitted VGAM object, e.g., from vgam .
<code>each.term</code>	Logical. If TRUE then a logical is returned for each term.
<code>...</code>	Unused for now.

Details

It is known that [vgam](#) with `s` terms do not correctly handle constraint matrices (`cmat`, say) when `crossprod(cmat)` is not diagonal. This function detects whether this is so or not. Note that probably all **VGAM** family functions have defaults where all `crossprod(cmat)`s are diagonal, therefore do not suffer from this bug. It is more likely to occur if the user inputs constraint matrices using the `constraints` argument (and setting `zero = NULL` if necessary).

Second-generation VGAMs based on [sm.ps](#) are a modern alternative to using `s`. It does not suffer from this bug. However, G2-VGAMs require a reasonably large sample size in order to work more reliably.

Value

The default is a single logical (TRUE if any term is TRUE), otherwise a vector of such with each element corresponding to a term. If the value is TRUE then I suggest replacing the VGAM by a similar model fitted by [vglm](#) and using regression splines, e.g., [bs](#), [ns](#).

Note

When the bug is fixed this function may be withdrawn, otherwise always return FALSEs!

Author(s)

T. W. Yee

See Also

[vgam](#), [vglm](#), [s](#), [sm.ps](#), [bs](#), [ns](#).

Examples

```

fit1 <- vgam(cbind(agaaus, kniexc) ~ s(altitude, df = c(3, 4)),
            binomialff(multiple.responses = TRUE), data = hunua)
is.buggy(fit1) # Okay
is.buggy(fit1, each.term = TRUE) # No terms are buggy
fit2 <-
  vgam(cbind(agaaus, kniexc) ~ s(altitude, df = c(3, 4)),
        binomialff(multiple.responses = TRUE), data = hunua,
        constraints =
          list("(Intercept)" = diag(2),
              "s(altitude, df = c(3, 4))" = matrix(c(1, 1, 0, 1), 2, 2)))
is.buggy(fit2) # TRUE
is.buggy(fit2, each.term = TRUE)
constraints(fit2)

# fit2b is an approximate alternative to fit2:
fit2b <-
  vglm(cbind(agaaus, kniexc) ~ bs(altitude, df=3) + bs(altitude, df=4),
        binomialff(multiple.responses = TRUE), data = hunua,
        constraints =
          list("(Intercept)" = diag(2),
              "bs(altitude, df = 3)" = rbind(1, 1),
              "bs(altitude, df = 4)" = rbind(0, 1)))
is.buggy(fit2b) # Okay
is.buggy(fit2b, each.term = TRUE)
constraints(fit2b)

```

is.crossing

Quantile Crossing Detection

Description

Returns a logical from testing whether an object such as an `extlogF1()` VGLM object has crossing quantiles.

Usage

```
is.crossing.vglm(object, ...)
```

Arguments

<code>object</code>	an object such as a <code>vglm</code> object with family function <code>extlogF1</code> .
<code>...</code>	additional optional arguments. Currently unused.

Details

This function was specifically written for a `vglm` with family function `extlogF1`. It examines the fitted quantiles to see if any cross. Note that if one uses regression splines such as `bs` and `ns` then it is possible that they cross at values of the covariate space that are not represented by actual data. One could use linear interpolation between fitted values to get around this problem.

Value

A logical. If TRUE then one can try fit a similar model by combining columns of the constraint matrices so that crossing no longer holds; see [fix.crossing](#). For LMS-Box-Cox type quantile regression models it is impossible for the quantiles to cross, by definition, hence FALSE is returned; see [lms.bcn](#).

See Also

[extlogF1](#), [fix.crossing](#), [lms.bcn](#), [vglm](#).

Examples

```
## Not run: ooo <- with(bmi.nz, order(age))
bmi.nz <- bmi.nz[ooo, ] # Sort by age
with(bmi.nz, plot(age, BMI, col = "blue"))
mytau <- c(50, 93, 95, 97) / 100 # Some quantiles are quite close
fit1 <- vglm(BMI ~ ns(age, 7), extlogF1(mytau), bmi.nz, trace = TRUE)
plot(BMI ~ age, bmi.nz, col = "blue", las = 1,
     main = "Partially parallel (darkgreen) & nonparallel quantiles",
     sub = "Crossing quantiles are orange")
is.crossing(fit1)
matlines(with(bmi.nz, age), fitted(fit1), lty = 1, col = "orange")
## End(Not run)
```

is.parallel

Parallelism Constraint Matrices

Description

Returns a logical vector from a test of whether an object such as a matrix or VGLM object corresponds to a parallelism assumption.

Usage

```
is.parallel.matrix(object, ...)
is.parallel.vglm(object, type = c("term", "lm"), ...)
```

Arguments

object	an object such as a constraint matrix or a vglm object.
type	passed into constraints .
...	additional optional arguments. Currently unused.

Details

These functions may be useful for categorical models such as [propodds](#), [cumulative](#), [acat](#), [cratio](#), [sratio](#), [multinomial](#).

Value

A vector of logicals, testing whether each constraint matrix is a one-column matrix of ones. Note that parallelism can still be thought of as holding if the constraint matrix has a non-zero but constant values, however, this is currently not implemented. No checking is done that the constraint matrices have the same number of rows.

See Also

[constraints](#), [vglm](#).

Examples

```
## Not run: require("VGAMdata")
fit <- vglm(educ ~ sm.bs(age) * sex + ethnicity,
           cumulative(parallel = TRUE), head(xs.nz, 200))
is.parallel(fit)
is.parallel(fit, type = "lm") # For each column of the LM matrix

## End(Not run)
```

is.smart

Test For a Smart Object

Description

Tests an object to see if it is smart.

Usage

```
is.smart(object)
```

Arguments

object a function or a fitted model.

Details

If object is a function then this function looks to see whether object has the logical attribute "smart". If so then this is returned, else FALSE.

If object is a fitted model then this function looks to see whether object@smart.prediction or object\$smart.prediction exists. If it does and it is not equal to list(smarts.arg=FALSE) then a TRUE is returned, else FALSE. The reason for this is because, e.g., lm(..., smart=FALSE) and vglm(..., smart=FALSE), will return such a specific list.

Writers of smart functions manually have to assign this attribute to their smart function after it has been written.

Value

Returns TRUE or FALSE, according to whether the object is smart or not.

Examples

```
is.smart(sm.min1) # TRUE
is.smart(sm.poly) # TRUE
library(splines)
is.smart(sm.bs) # TRUE
is.smart(sm.ns) # TRUE
is.smart(tan) # FALSE
## Not run:
udata <- data.frame(x2 = rnorm(9))
fit1 <- vglm(rnorm(9) ~ x2, uninormal, data = udata)
is.smart(fit1) # TRUE
fit2 <- vglm(rnorm(9) ~ x2, uninormal, data = udata, smart = FALSE)
is.smart(fit2) # FALSE
fit2@smart.prediction

## End(Not run)
```

is.zero

Zero Constraint Matrices

Description

Returns a logical vector from a test of whether an object such as a matrix or VGLM object corresponds to a 'zero' assumption.

Usage

```
is.zero.matrix(object, ...)
is.zero.vglm(object, ...)
```

Arguments

object an object such as a coefficient matrix of a `vglm` object, or a `vglm` object.
... additional optional arguments. Currently unused.

Details

These functions test the effect of the zero argument on a `vglm` object or the coefficient matrix of a `vglm` object. The latter is obtained by `coef(vglmObject, matrix = TRUE)`.

Value

A vector of logicals, testing whether each linear/additive predictor has the zero argument applied to it. It is TRUE if that linear/additive predictor is intercept-only, i.e., all other regression coefficients are set to zero.

No checking is done for the intercept term at all, i.e., that it was estimated in the first place.

See Also

[constraints](#), [vglm](#), [CommonVGAMffArguments](#).

Examples

```
coalminers <- transform(coalminers, Age = (age - 42) / 5)
fit <- vglm(cbind(nBnW,nBW,BnW,BW) ~ Age, binom2.or(zero = NULL),
           data = coalminers)
is.zero(fit)
is.zero(coef(fit, matrix = TRUE))
```

kendall.tau

Kendall's Tau Statistic

Description

Computes Kendall's Tau, which is a rank-based correlation measure, between two vectors.

Usage

```
kendall.tau(x, y, exact = FALSE, max.n = 3000)
```

Arguments

<code>x, y</code>	Numeric vectors. Must be of equal length. Ideally their values are continuous and not too discrete. Let $\text{length}(x)$ be N , say.
<code>exact</code>	Logical. If TRUE then the exact value is computed.
<code>max.n</code>	Numeric. If <code>exact = FALSE</code> and $\text{length}(x)$ is more than <code>max.n</code> then a random sample of <code>max.n</code> pairs are chosen.

Details

Kendall's tau is a measure of dependency in a bivariate distribution. Loosely, two random variables are *concordant* if large values of one random variable are associated with large values of the other random variable. Similarly, two random variables are *disconcordant* if large values of one random variable are associated with small values of the other random variable. More formally, if $(x[i] - x[j]) \cdot (y[i] - y[j]) > 0$ then that comparison is concordant ($i \neq j$). And if $(x[i] - x[j]) \cdot (y[i] - y[j]) < 0$ then that comparison is disconcordant ($i \neq j$). Out of $\text{choose}(N, 2)$ comparisons, let c and d be the number of concordant and disconcordant pairs. Then Kendall's tau can be estimated by $(c - d)/(c + d)$. If there are ties then half the ties are deemed concordant and half disconcordant so that $(c - d)/(c + d + t)$ is used.

Value

Kendall's tau, which lies between -1 and 1 .

Warning

If `length(x)` is large then the cost is $O(N^2)$, which is expensive! Under these circumstances it is not advisable to set `exact = TRUE` or `max.n` to a very large number.

See Also

[binormalcop](#), [cor](#).

Examples

```
N <- 5000; x <- 1:N; y <- runif(N)
true.rho <- -0.8
ymat <- rbinorm(N, cov12 = true.rho) # Bivariate normal, aka N_2
x <- ymat[, 1]
y <- ymat[, 2]

## Not run: plot(x, y, col = "blue")

kendall.tau(x, y) # A random sample is taken here
kendall.tau(x, y) # A random sample is taken here

kendall.tau(x, y, exact = TRUE) # Costly if length(x) is large
kendall.tau(x, y, max.n = N)   # Same as exact = TRUE

(rhohat <- sin(kendall.tau(x, y) * pi / 2)) # Holds for N_2 actually
true.rho # rhohat should be near this value
```

KLD

Kullback-Leibler Divergence

Description

Calculates the Kullback-Leibler divergence for certain fitted model objects

Usage

```
KLD(object, ...)
KLDvglm(object, ...)
```

Arguments

<code>object</code>	Some VGAM object, for example, having class <code>vglm-class</code> . Currently object must be intercept-only.
<code>...</code>	Other possible arguments fed into <code>KLDvglm</code> in order to compute the KLD.

Details

The *Kullback-Leibler divergence* (KLD), or *relative entropy*, is a measure of how one probability distribution differs from a second reference probability distribution. Currently the **VGAM** package computes the KLD for GAITD regression models (e.g., see [gaitdpoisson](#) and [gaitdnbinomial](#)) where the reference distribution is the (unscaled) parent or base distribution. For such, the formula for the KLD simplifies somewhat. Hence one can obtain a quantitative measure for the overall effect of altering, inflating, truncating and deflating certain (special) values.

Value

Returns a numeric nonnegative value with the corresponding KLD. A 0 value means no difference between an ordinary parent or base distribution.

Warning

Numerical problems might occur if any of the evaluated probabilities of the unscaled parent distribution are very close to 0.

Author(s)

T. W. Yee.

References

Kullback, S. and Leibler, R. A. (1951). On information and sufficiency. *Annals of Mathematical Statistics*, **22**, 79–86.

McKendrick, A. G. (1925). Applications of mathematics to medical problems. *Proc. Edinb. Math. Soc.*, **44**, 98–130.

See Also

[gaitdpoisson](#), [gaitdnbinomial](#).

Examples

```
# McKendrick (1925): Data from 223 Indian village households
cholera <- data.frame(ncases = 0:4, # Number of cholera cases,
                    wfreq = c(168, 32, 16, 6, 1)) # Frequencies
fit7 <- vglm(ncases ~ 1, gaitdpoisson(i.nlm = 0, ilambda.p = 1),
            weight = wfreq, data = cholera, trace = TRUE)
coef(fit7, matrix = TRUE)
KLD(fit7)
```

Description

Density, distribution function, quantile function and random generation for the Kumaraswamy distribution.

Usage

```
dkumar(x, shape1, shape2, log = FALSE)
pkumar(q, shape1, shape2, lower.tail = TRUE, log.p = FALSE)
qkumar(p, shape1, shape2, lower.tail = TRUE, log.p = FALSE)
rkumar(n, shape1, shape2)
```

Arguments

x, q	vector of quantiles.
p	vector of probabilities.
n	number of observations. If <code>length(n) > 1</code> then the length is taken to be the number required.
shape1, shape2	positive shape parameters.
log	Logical. If <code>log = TRUE</code> then the logarithm of the density is returned.
lower.tail, log.p	Same meaning as in pnorm or qnorm .

Details

See [kumar](#), the **VGAM** family function for estimating the parameters, for the formula of the probability density function and other details.

Value

`dkumar` gives the density, `pkumar` gives the distribution function, `qkumar` gives the quantile function, and `rkumar` generates random deviates.

Author(s)

T. W. Yee and Kai Huang

See Also

[kumar](#).

Examples

```
## Not run:
shape1 <- 2; shape2 <- 2; nn <- 201; # shape1 <- shape2 <- 0.5;
x <- seq(-0.05, 1.05, len = nn)
plot(x, dkumar(x, shape1, shape2), type = "l", las = 1,
      ylab = paste("dkumar(shape1 = ", shape1,
                    ", shape2 = ", shape2, ")"),
      col = "blue", cex.main = 0.8, ylim = c(0,1.5),
      main = "Blue is density, orange is the CDF",
      sub = "Red lines are the 10,20,...,90 percentiles")
lines(x, pkumar(x, shape1, shape2), col = "orange")
probs <- seq(0.1, 0.9, by = 0.1)
Q <- qkumar(probs, shape1, shape2)
lines(Q, dkumar(Q, shape1, shape2), col = "red", lty = 3, type = "h")
lines(Q, pkumar(Q, shape1, shape2), col = "red", lty = 3, type = "h")
abline(h = probs, col = "red", lty = 3)
max(abs(pkumar(Q, shape1, shape2) - probs)) # Should be 0

## End(Not run)
```

 kumar

Kumaraswamy Regression Family Function

Description

Estimates the two parameters of the Kumaraswamy distribution by maximum likelihood estimation.

Usage

```
kumar(lshape1 = "loglink", lshape2 = "loglink",
      ishape1 = NULL, ishape2 = NULL,
      gshape1 = exp(2*ppoints(5) - 1), tol12 = 1.0e-4, zero = NULL)
```

Arguments

lshape1, lshape2	Link function for the two positive shape parameters, respectively, called a and b below. See Links for more choices.
ishape1, ishape2	Numeric. Optional initial values for the two positive shape parameters.
tol12	Numeric and positive. Tolerance for testing whether the second shape parameter is either 1 or 2. If so then the working weights need to handle these singularities.
gshape1	Values for a grid search for the first shape parameter. See CommonVGAMffArguments for more information.
zero	See CommonVGAMffArguments .

Details

The Kumaraswamy distribution has density function

$$f(y; a = \text{shape1}, b = \text{shape2}) = aby^{a-1}(1 - y^a)^{b-1}$$

where $0 < y < 1$ and the two shape parameters, a and b , are positive. The mean is $b \times \text{Beta}(1 + 1/a, b)$ (returned as the fitted values) and the variance is $b \times \text{Beta}(1 + 2/a, b) - (b \times \text{Beta}(1 + 1/a, b))^2$. Applications of the Kumaraswamy distribution include the storage volume of a water reservoir. Fisher scoring is implemented. Handles multiple responses (matrix input).

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Author(s)

T. W. Yee

References

- Kumaraswamy, P. (1980). A generalized probability density function for double-bounded random processes. *Journal of Hydrology*, **46**, 79–88.
- Jones, M. C. (2009). Kumaraswamy's distribution: A beta-type distribution with some tractability advantages. *Statistical Methodology*, **6**, 70–81.

See Also

[dkumar](#), [betaff](#), [simulate.vlm](#).

Examples

```
shape1 <- exp(1); shape2 <- exp(2)
kdata <- data.frame(y = rkumar(n = 1000, shape1, shape2))
fit <- vglm(y ~ 1, kumar, data = kdata, trace = TRUE)
c(with(kdata, mean(y)), head(fitted(fit), 1))
coef(fit, matrix = TRUE)
Coef(fit)
summary(fit)
```

lakeO

Annual catches on Lake Otamangakau from October 1974 to October 1989

Description

Rainbow and brown trout catches by a Mr Swainson at Lake Otamangakau in the central North Island of New Zealand during the 1970s and 1980s.

Usage

```
data(lake0)
```

Format

A data frame with 15 observations on the following 5 variables.

`year` a numeric vector, the season began on 1 October of the year and ended 12 months later.

`total.fish` a numeric vector, the total number of fish caught during the season. Simply the sum of brown and rainbow trout.

`brown` a numeric vector, the number of brown trout (*Salmo trutta*) caught.

`rainbow` a numeric vector, the number of rainbow trout (*Oncorhynchus mykiss*) caught.

`visits` a numeric vector, the number of visits during the season that the angler made to the lake. It is necessary to assume that the visits were of an equal time length in order to interpret the usual Poisson regressions.

Details

The data was extracted from the season summaries at Lake Otamangakau by Anthony Swainson for the seasons 1974–75 to 1988–89. Mr Swainson was one of a small group of regular fly fishing anglers and kept a diary of his catches. Lake Otamangakau is a lake of area 1.8 squared km and has a maximum depth of about 12m, and is located in the central North Island of New Zealand. It is trout-infested and known for its trophy-sized fish.

See also [trap0](#).

Source

Table 7.2 of the reference below. Thanks to Dr Michel Dedual for a copy of the report and for help reading the final year's data. The report is available from TWY on request.

References

Dedual, M. and MacLean, G. and Rowe, D. and Cudby, E., *The Trout Population and Fishery of Lake Otamangakau—Interim Report*. National Institute of Water and Atmospheric Research, Hamilton, New Zealand. Consultancy Report Project No. ELE70207, (Dec 1996).

Examples

```
data(lake0)
lake0
summary(lake0)
```

lambertW

*The Lambert W Function***Description**

Computes the Lambert W function for real values.

Usage

```
lambertW(x, tolerance = 1e-10, maxit = 50)
```

Arguments

<code>x</code>	A vector of reals.
<code>tolerance</code>	Accuracy desired.
<code>maxit</code>	Maximum number of iterations of third-order Halley's method.

Details

The Lambert W function is the root of the equation $W(z) \exp(W(z)) = z$ for complex z . If z is real and $-1/e < z < 0$ then it has two possible real values, and currently only the upper branch (often called W_0) is computed so that a value that is ≥ -1 is returned.

Value

This function returns the principal branch of the W function for *real* z . It returns $W(z) \geq -1$, and NA for $z < -1/e$.

Note

If convergence does not occur then increase the value of `maxit` and/or `tolerance`.

Yet to do: add an argument `lbranch = TRUE` to return the lower branch (often called W_{-1}) for real $-1/e \leq z < 0$; this would give $W(z) \leq -1$.

Author(s)

T. W. Yee

References

Corless, R. M. and Gonnet, G. H. and Hare, D. E. G. and Jeffrey, D. J. and Knuth, D. E. (1996). On the Lambert W function. *Advances in Computational Mathematics*, **5**(4), 329–359.

See Also

[log](#), [exp](#), [bell](#). There is also a package called **LambertW**.

Examples

```
## Not run:
curve(lambertW, -exp(-1), 3, xlim = c(-1, 3), ylim = c(-2, 1),
      las = 1, col = "orange", n = 1001)
abline(v = -exp(-1), h = -1, lwd = 2, lty = "dotted", col = "gray")
abline(h = 0, v = 0, lty = "dashed", col = "blue")
## End(Not run)
```

laplace

*Laplace Regression Family Function***Description**

Maximum likelihood estimation of the 2-parameter classical Laplace distribution.

Usage

```
laplace(llocation = "identitylink", lscale = "loglink",
        ilocation = NULL, iscale = NULL, imethod = 1, zero = "scale")
```

Arguments

`llocation`, `lscale` Character. Parameter link functions for location parameter a and scale parameter b . See [Links](#) for more choices.

`ilocation`, `iscale` Optional initial values. If given, it must be numeric and values are recycled to the appropriate length. The default is to choose the value internally.

`imethod` Initialization method. Either the value 1 or 2.

`zero` See [CommonVGAMffArguments](#) for information.

Details

The Laplace distribution is often known as the *double-exponential* distribution and, for modelling, has heavier tail than the normal distribution. The Laplace density function is

$$f(y) = \frac{1}{2b} \exp\left(-\frac{|y-a|}{b}\right)$$

where $-\infty < y < \infty$, $-\infty < a < \infty$ and $b > 0$. Its mean is a and its variance is $2b^2$. This parameterization is called the *classical Laplace distribution* by Kotz et al. (2001), and the density is symmetric about a .

For $y \sim 1$ (where y is the response) the maximum likelihood estimate (MLE) for the location parameter is the sample median, and the MLE for b is `mean(abs(y-location))` (replace location by its MLE if unknown).

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Warning

This family function has not been fully tested. The MLE regularity conditions do *not* hold for this distribution, therefore misleading inferences may result, e.g., in the summary and vcov of the object. Hence this family function might be withdrawn from **VGAM** in the future.

Note

This family function uses Fisher scoring. Convergence may be slow for non-intercept-only models; half-stepping is frequently required.

Author(s)

T. W. Yee

References

Kotz, S., Kozubowski, T. J. and Podgorski, K. (2001). *The Laplace distribution and generalizations: a revisit with applications to communications, economics, engineering, and finance*, Boston: Birkhauser.

See Also

[rlaplace](#), [alaplace2](#) (which differs slightly from this parameterization), [exponential](#), [median](#).

Examples

```
ldata <- data.frame(y = rlaplace(nn <- 100, 2, scale = exp(1)))
fit <- vglm(y ~ 1, laplace, ldata, trace = TRUE)
coef(fit, matrix = TRUE)
Coef(fit)
with(ldata, median(y))
```

```
ldata <- data.frame(x = runif(nn <- 1001))
ldata <- transform(ldata, y = rlaplace(nn, 2, scale = exp(-1 + 1*x)))
coef(vglm(y ~ x, laplace(iloc = 0.2, imethod = 2, zero = 1), ldata,
  trace = TRUE), matrix = TRUE)
```


Description

Density, distribution function, quantile function and random generation for the Laplace distribution with location parameter `location` and scale parameter `scale`.

Usage

```
dlaplace(x, location = 0, scale = 1, log = FALSE)
plaplace(q, location = 0, scale = 1, lower.tail = TRUE, log.p = FALSE)
qlaplace(p, location = 0, scale = 1, lower.tail = TRUE, log.p = FALSE)
rlaplace(n, location = 0, scale = 1)
```

Arguments

<code>x, q</code>	vector of quantiles.
<code>p</code>	vector of probabilities.
<code>n</code>	number of observations. Same as in runif .
<code>location</code>	the location parameter a , which is the mean.
<code>scale</code>	the scale parameter b . Must consist of positive values.
<code>log</code>	Logical. If <code>log = TRUE</code> then the logarithm of the density is returned.
<code>lower.tail, log.p</code>	Same meaning as in pnorm or qnorm .

Details

The Laplace distribution is often known as the double-exponential distribution and, for modelling, has heavier tail than the normal distribution. The Laplace density function is

$$f(y) = \frac{1}{2b} \exp\left(-\frac{|y-a|}{b}\right)$$

where $-\infty < y < \infty$, $-\infty < a < \infty$ and $b > 0$. The mean is a and the variance is $2b^2$.

See [laplace](#), the **VGAM** family function for estimating the two parameters by maximum likelihood estimation, for formulae and details. Apart from `n`, all the above arguments may be vectors and are recycled to the appropriate length if necessary.

Value

`dlaplace` gives the density, `plaplace` gives the distribution function, `qlaplace` gives the quantile function, and `rlaplace` generates random deviates.

Author(s)

T. W. Yee and Kai Huang

References

Forbes, C., Evans, M., Hastings, N. and Peacock, B. (2011). *Statistical Distributions*, Hoboken, NJ, USA: John Wiley and Sons, Fourth edition.

See Also

[laplace](#).

Examples

```
loc <- 1; b <- 2
y <- rlaplace(n = 100, loc = loc, scale = b)
mean(y) # sample mean
loc     # population mean
var(y)  # sample variance
2 * b^2 # population variance

## Not run: loc <- 0; b <- 1.5; x <- seq(-5, 5, by = 0.01)
plot(x, dlaplace(x, loc, b), type = "l", col = "blue",
     main = "Blue is density, orange is the CDF", ylim = c(0,1),
     sub = "Purple are 5,10,...,95 percentiles", las = 1, ylab = "")
abline(h = 0, col = "blue", lty = 2)
lines(qlaplace(seq(0.05,0.95,by = 0.05), loc, b),
      dlaplace(qlaplace(seq(0.05, 0.95, by = 0.05), loc, b), loc, b),
      col = "purple", lty = 3, type = "h")
lines(x, plaplace(x, loc, b), type = "l", col = "orange")
abline(h = 0, lty = 2)
## End(Not run)

plaplace(qlaplace(seq(0.05, 0.95, by = 0.05), loc, b), loc, b)
```

latvar

Latent Variables

Description

Generic function for the *latent variables* of a model.

Usage

```
latvar(object, ...)
lv(object, ...)
```

Arguments

object An object for which the extraction of latent variables is meaningful.

... Other arguments fed into the specific methods function of the model. Sometimes they are fed into the methods function for [Coef](#).

Details

Latent variables occur in reduced-rank regression models, as well as in quadratic and additive ordination models. For the latter two, latent variable values are often called *site scores* by ecologists. Latent variables are linear combinations of the explanatory variables.

Value

The value returned depends specifically on the methods function invoked.

Warning

`latvar` and `lv` are identical, but the latter will be deprecated soon.

Latent variables are not really applicable to `vglm/vgam` models.

Author(s)

Thomas W. Yee

References

Yee, T. W. and Hastie, T. J. (2003). Reduced-rank vector generalized linear models. *Statistical Modelling*, **3**, 15–41.

Yee, T. W. (2004). A new technique for maximum-likelihood canonical Gaussian ordination. *Ecological Monographs*, **74**, 685–701.

Yee, T. W. (2006). Constrained additive ordination. *Ecology*, **87**, 203–213.

See Also

`latvar.qrrvglm`, `latvar.rrvglm`, `latvar.cao`, `lvplot`.

Examples

```
## Not run:
hspider[, 1:6] <- scale(hspider[, 1:6]) # Standardized environmental vars
set.seed(123)
p1 <- cao(cbind(Pardlugu, Pardmont, Pardnigr, Pardpull, Zoraspin) ~
  WaterCon + BareSand + FallTwig + CoveMoss + CoveHerb + ReflLux,
  family = poissonff, data = hspider, Rank = 1, df1.nl =
  c(Zoraspin = 2.5, 3), Bestof = 3, Crowlpositive = TRUE)

var(latvar(p1)) # Scaled to unit variance # Scaled to unit variance
c(latvar(p1))  # Estimated site scores

## End(Not run)
```

leipnik

*Leipnik Regression Family Function***Description**

Estimates the two parameters of a (transformed) Leipnik distribution by maximum likelihood estimation.

Usage

```
leipnik(lmu = "logitlink", llambda = logofflink(offset = 1),
       imu = NULL, ilambda = NULL)
```

Arguments

`lmu`, `llambda` Link function for the μ and λ parameters. See [Links](#) for more choices.
`imu`, `ilambda` Numeric. Optional initial values for μ and λ .

Details

The (transformed) Leipnik distribution has density function

$$f(y; \mu, \lambda) = \frac{\{y(1-y)\}^{-\frac{1}{2}}}{\text{Beta}(\frac{\lambda+1}{2}, \frac{1}{2})} \left[1 + \frac{(y-\mu)^2}{y(1-y)} \right]^{-\frac{\lambda}{2}}$$

where $0 < y < 1$ and $\lambda > -1$. The mean is μ (returned as the fitted values) and the variance is $1/\lambda$.

Jorgensen (1997) calls the above the **transformed** Leipnik distribution, and if $y = (x + 1)/2$ and $\mu = (\theta + 1)/2$, then the distribution of X as a function of x and θ is known as the the (untransformed) Leipnik distribution. Here, both x and θ are in $(-1, 1)$.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), [rrvglm](#) and [vgam](#).

Note

Convergence may be slow or fail. Until better initial value estimates are forthcoming try assigning the argument `ilambda` some numerical value if it fails to converge. Currently, Newton-Raphson is implemented, not Fisher scoring. Currently, this family function probably only really works for intercept-only models, i.e., $y \sim 1$ in the formula.

Author(s)

T. W. Yee

References

- Jorgensen, B. (1997). *The Theory of Dispersion Models*. London: Chapman & Hall
- Johnson, N. L. and Kotz, S. and Balakrishnan, N. (1995). *Continuous Univariate Distributions*, 2nd edition, Volume 2, New York: Wiley. (pages 612–617).

See Also

[mccullagh89](#).

Examples

```
ldata <- data.frame(y = rnorm(2000, 0.5, 0.1)) # Improper data
fit <- vglm(y ~ 1, leipnik(ilambda = 1), ldata, trace = TRUE)
head(fitted(fit))
with(ldata, mean(y))
summary(fit)
coef(fit, matrix = TRUE)
Coef(fit)

sum(weights(fit)) # Sum of the prior weights
sum(weights(fit, type = "work")) # Sum of the working weights
```

lerch

Lerch Phi Function

Description

Computes the Lerch Phi function.

Usage

```
lerch(x, s, v, tolerance = 1.0e-10, iter = 100)
```

Arguments

x, s, v	Numeric. This function recycles values of x, s, and v if necessary.
tolerance	Numeric. Accuracy required, must be positive and less than 0.01.
iter	Maximum number of iterations allowed to obtain convergence. If iter is too small then a result of NA may occur; if so, try increasing its value.

Details

Also known as the Lerch transcendent, it can be defined by an integral involving analytical continuation. An alternative definition is the series

$$\Phi(x, s, v) = \sum_{n=0}^{\infty} \frac{x^n}{(n+v)^s}$$

which converges for $|x| < 1$ as well as for $|x| = 1$ with $s > 1$. The series is undefined for integers $v \leq 0$. Actually, x may be complex but this function only works for real x . The algorithm used is based on the relation

$$\Phi(x, s, v) = x^m \Phi(x, s, v + m) + \sum_{n=0}^{m-1} \frac{x^n}{(n+v)^s}.$$

See the URL below for more information. This function is a wrapper function for the C code described below.

Value

Returns the value of the function evaluated at the values of x , s , v . If the above ranges of x and v are not satisfied, or some numeric problems occur, then this function will return an NA for those values. (The C code returns 6 possible return codes, but this is not passed back up to the R level.)

Warning

This function has not been thoroughly tested and contains limitations, for example, the zeta function cannot be computed with this function even though $\zeta(s) = \Phi(x = 1, s, v = 1)$. Several numerical problems can arise, such as lack of convergence, overflow and underflow, especially near singularities. If any problems occur then an NA will be returned. For example, if $|x| = 1$ and $s > 1$ then convergence may be so slow that changing tolerance and/or `iter` may be needed to get an answer (that is treated cautiously).

Note

There are a number of special cases, e.g., the Riemann zeta-function is $\zeta(s) = \Phi(x = 1, s, v = 1)$. Another example is the Hurwitz zeta function $\zeta(s, v) = \Phi(x = 1, s, v = v)$. The special case of $s = 1$ corresponds to the hypergeometric ${}_2F_1$, and this is implemented in the **gsl** package. The Lerch Phi function should not be confused with the Lerch zeta function though they are quite similar.

Author(s)

S. V. Aksenov and U. D. Jentschura wrote the C code (called Version 1.00). The R wrapper function was written by T. Yee.

References

Originally the code was found at <http://aksenov.freeshell.org/lerchphi/source/lerchphi.c>.
Bateman, H. (1953). *Higher Transcendental Functions*. Volume 1. McGraw-Hill, NY, USA.

See Also

[zeta](#).

Examples

```
## Not run:
s <- 2; v <- 1; x <- seq(-1.1, 1.1, length = 201)
plot(x, lerch(x, s = s, v = v), type = "l", col = "blue",
     las = 1, main = paste0("lerch(x, s = ", s, ", v = ", v, ")"))
abline(v = 0, h = 1, lty = "dashed", col = "gray")

## End(Not run)
```

leukemia

Acute Myelogenous Leukemia Survival Data

Description

Survival in patients with Acute Myelogenous Leukemia

Usage

```
data(leukemia)
```

Format

time: survival or censoring time
status: censoring status
x: maintenance chemotherapy given? (factor)

Note

This data set has been transferred from **survival** and renamed from aml to leukemia.

Source

Rupert G. Miller (1997). *Survival Analysis*. John Wiley & Sons.

levy

*Levy Distribution Family Function***Description**

Estimates the scale parameter of the Levy distribution by maximum likelihood estimation.

Usage

```
levy(location = 0, lscale = "loglink", iscale = NULL)
```

Arguments

location	Location parameter. Must have a known value. Called a below.
lscale	Parameter link function for the (positive) scale parameter b . See Links for more choices.
iscale	Initial value for the b parameter. By default, an initial value is chosen internally.

Details

The Levy distribution is one of three stable distributions whose density function has a tractable form. The formula for the density is

$$f(y; b) = \sqrt{\frac{b}{2\pi}} \exp\left(\frac{-b}{2(y-a)}\right) / (y-a)^{3/2}$$

where $a < y < \infty$ and $b > 0$. Note that if a is very close to $\min(y)$ (where y is the response), then numerical problem will occur. The mean does not exist. The median is returned as the fitted values.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Author(s)

T. W. Yee

References

Nolan, J. P. (2005). *Stable Distributions: Models for Heavy Tailed Data*.

See Also

The Nolan article was at <http://academic2.american.edu/~jpnolan/stable/chap1.pdf>.

Examples

```

nn <- 1000; loc1 <- 0; loc2 <- 10
myscale <- 1 # log link ==> 0 is the answer
ldata <-
  data.frame(y1 = loc1 + myscale/rnorm(nn)^2, # Levy(myscale, a)
            y2 = rlevy(nn, loc = loc2, scale = exp(+2)))
# Cf. Table 1.1 of Nolan for Levy(1,0)
with(ldata, sum(y1 > 1) / length(y1)) # Should be 0.6827
with(ldata, sum(y1 > 2) / length(y1)) # Should be 0.5205

fit1 <- vglm(y1 ~ 1, levy(location = loc1), ldata, trace = TRUE)
coef(fit1, matrix = TRUE)
Coef(fit1)
summary(fit1)
head(weights(fit1, type = "work"))

fit2 <- vglm(y2 ~ 1, levy(location = loc2), ldata, trace = TRUE)
coef(fit2, matrix = TRUE)
Coef(fit2)
c(median = with(ldata, median(y2)),
  fitted.median = head(fitted(fit2), 1))

```

lgamma1

*Log-gamma Distribution Family Function***Description**

Estimation of the parameter of the standard and nonstandard log-gamma distribution.

Usage

```

lgamma1(lshape = "loglink", ishape = NULL)
lgamma3(llocation = "identitylink", lscale = "loglink",
        lshape = "loglink", ilocation = NULL, iscale = NULL, ishape = 1,
        zero = c("scale", "shape"))

```

Arguments

llocation, lscale	Parameter link function applied to the location parameter a and the positive scale parameter b . See Links for more choices.
lshape	Parameter link function applied to the positive shape parameter k . See Links for more choices.
ishape	Initial value for k . If given, it must be positive. If failure to converge occurs, try some other value. The default means an initial value is determined internally.
ilocation, iscale	Initial value for a and b . The defaults mean an initial value is determined internally for each.

zero An integer-valued vector specifying which linear/additive predictors are modelled as intercepts only. The values must be from the set {1,2,3}. The default value means none are modelled as intercept-only terms. See [CommonVGAMffArguments](#) for more information.

Details

The probability density function of the standard log-gamma distribution is given by

$$f(y; k) = \exp[ky - \exp(y)]/\Gamma(k),$$

for parameter $k > 0$ and all real y . The mean of Y is `digamma(k)` (returned as the fitted values) and its variance is `trigamma(k)`.

For the non-standard log-gamma distribution, one replaces y by $(y - a)/b$, where a is the location parameter and b is the positive scale parameter. Then the density function is

$$f(y) = \exp[k(y - a)/b - \exp((y - a)/b)]/(b\Gamma(k)).$$

The mean and variance of Y are $a + b*\text{digamma}(k)$ (returned as the fitted values) and $b^2 * \text{trigamma}(k)$, respectively.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as `vglm`, and `vgam`.

Note

The standard log-gamma distribution can be viewed as a generalization of the standard type 1 extreme value density: when $k = 1$ the distribution of $-Y$ is the standard type 1 extreme value distribution.

The standard log-gamma distribution is fitted with `lgamma1` and the non-standard (3-parameter) log-gamma distribution is fitted with `lgamma3`.

Author(s)

T. W. Yee

References

Kotz, S. and Nadarajah, S. (2000). *Extreme Value Distributions: Theory and Applications*, pages 48–49, London: Imperial College Press.

Johnson, N. L. and Kotz, S. and Balakrishnan, N. (1995). *Continuous Univariate Distributions*, 2nd edition, Volume 2, p.89, New York: Wiley.

See Also

[rlgamma](#), [gengamma.stacy](#), [prentice74](#), [gamma1](#), [lgamma](#).

Examples

```

ldata <- data.frame(y = rlgamma(100, shape = exp(1)))
fit <- vglm(y ~ 1, lgamma1, ldata, trace = TRUE, crit = "coef")
summary(fit)
coef(fit, matrix = TRUE)
Coef(fit)

ldata <- data.frame(x2 = runif(nn <- 5000)) # Another example
ldata <- transform(ldata, loc = -1 + 2 * x2, Scale = exp(1))
ldata <- transform(ldata, y = rlgamma(nn, loc, sc = Scale, sh = exp(0)))
fit2 <- vglm(y ~ x2, lgamma3, data = ldata, trace = TRUE, crit = "c")
coef(fit2, matrix = TRUE)

```

lgammaUC

*The Log-Gamma Distribution***Description**

Density, distribution function, quantile function and random generation for the log-gamma distribution with location parameter location, scale parameter scale and shape parameter k.

Usage

```

dlgamma(x, location = 0, scale = 1, shape = 1, log = FALSE)
plgamma(q, location = 0, scale = 1, shape = 1,
        lower.tail = TRUE, log.p = FALSE)
qlgamma(p, location = 0, scale = 1, shape = 1,
        lower.tail = TRUE, log.p = FALSE)
rlgamma(n, location = 0, scale = 1, shape = 1)

```

Arguments

x, q	vector of quantiles.
p	vector of probabilities.
n	number of observations. Same as runif .
location	the location parameter a .
scale	the (positive) scale parameter b .
shape	the (positive) shape parameter k .
log	Logical. If log = TRUE then the logarithm of the density is returned.
lower.tail, log.p	Same meaning as in pnorm or qnorm .

Details

See [lgamma1](#), the **VGAM** family function for estimating the one parameter standard log-gamma distribution by maximum likelihood estimation, for formulae and other details. Apart from n, all the above arguments may be vectors and are recycled to the appropriate length if necessary.

Value

dlgamma gives the density, plgamma gives the distribution function, qlgamma gives the quantile function, and rlgamma generates random deviates.

Note

The **VGAM** family function [lgamma3](#) is for the three parameter (nonstandard) log-gamma distribution.

Author(s)

T. W. Yee and Kai Huang

References

Kotz, S. and Nadarajah, S. (2000). *Extreme Value Distributions: Theory and Applications*, pages 48–49, London: Imperial College Press.

See Also

[lgamma1](#), [prentice74](#).

Examples

```
## Not run: loc <- 1; Scale <- 1.5; shape <- 1.4
x <- seq(-3.2, 5, by = 0.01)
plot(x, dlgamma(x, loc = loc, Scale = Scale, shape = shape), type = "l",
      col = "blue", ylim = 0:1,
      main = "Blue is density, orange is the CDF",
      sub = "Red are 5,10,...,95 percentiles", las = 1, ylab = "")
abline(h = 0, col = "blue", lty = 2)
lines(qlgamma(seq(0.05, 0.95, by = 0.05), loc = loc, Scale, sh = shape),
      dlgamma(qlgamma(seq(0.05, 0.95, by = 0.05), loc = loc, sc = Scale,
                     shape = shape),
            loc = loc, Scale, shape = shape), col = "red", lty = 3, type = "h")
lines(x, plgamma(x, loc = loc, Scale, shape = shape), col = "orange")
abline(h = 0, lty = 2)
## End(Not run)
```

Description

Density, cumulative distribution function, and random generation for the Lindley distribution.

Usage

```
dlind(x, theta, log = FALSE)
plind(q, theta, lower.tail = TRUE, log.p = FALSE)
rlind(n, theta)
```

Arguments

x, q	vector of quantiles.
n	number of observations. Same as in runif .
log	Logical. If log = TRUE then the logarithm of the density is returned.
theta	positive parameter.
lower.tail, log.p	Same meaning as in pnorm or qnorm .

Details

See [lindley](#) for details.

Value

dlind gives the density, plind gives the cumulative distribution function, and rlind generates random deviates.

Author(s)

T. W. Yee and Kai Huang

See Also

[lindley](#).

Examples

```
theta <- exp(-1); x <- seq(0.0, 17, length = 700)
dlind(0:10, theta)
## Not run:
plot(x, dlind(x, theta), type = "l", las = 1, col = "blue",
      main = "dlind(x, theta = exp(-1))")
abline(h = 1, col = "grey", lty = "dashed")
## End(Not run)
```

lindley

*1-parameter Lindley Distribution***Description**

Estimates the (1-parameter) Lindley distribution by maximum likelihood estimation.

Usage

```
lindley(link = "loglink", itheta = NULL, zero = NULL)
```

Arguments

`link` Link function applied to the (positive) parameter. See [Links](#) for more choices.
`itheta, zero` See [CommonVGAMffArguments](#) for information.

Details

The density function is given by

$$f(y; \theta) = \theta^2(1 + y) \exp(-\theta y)/(1 + \theta)$$

for $\theta > 0$ and $y > 0$. The mean of Y (returned as the fitted values) is $\mu = (\theta + 2)/(\theta(\theta + 1))$. The variance is $(\theta^2 + 4\theta + 2)/(\theta(\theta + 1))^2$.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Note

This **VGAM** family function can handle multiple responses (inputted as a matrix). Fisher scoring is implemented.

Author(s)

T. W. Yee

References

- Lindley, D. V. (1958). Fiducial distributions and Bayes' theorem. *Journal of the Royal Statistical Society, Series B, Methodological*, **20**, 102–107.
- Ghitany, M. E. and Atieh, B. and Nadarajah, S. (2008). Lindley distribution and its application. *Math. Comput. Simul.*, **78**, 493–506.

See Also

[dlind](#), [gammaR](#), [simulate.vlm](#).

Examples

```
ldata <- data.frame(y = rbind(n = 1000, theta = exp(3)))
fit <- vglm(y ~ 1, lindley, data = ldata, trace = TRUE, crit = "coef")
coef(fit, matrix = TRUE)
Coef(fit)
summary(fit)
```

linkfun	<i>Link Functions for VGLMs</i>
---------	---------------------------------

Description

Returns the link functions, and parameter names, for *vector generalized linear models* (VGLMs).

Usage

```
linkfun(object, ...)
linkfunvlm(object, earg = FALSE, ...)
```

Arguments

object	An object which has parameter link functions, e.g., has class "vglm".
earg	Logical. Return the extra arguments associated with each link function? If TRUE then a list is returned.
...	Arguments that might be used in the future.

Details

All fitted VGLMs have a link function applied to each parameter. This function returns these, and optionally, the extra arguments associated with them.

Value

Usually just a (named) character string, with the link functions in order. It is named with the parameter names. If earg = TRUE then a list with the following components.

link	The default output.
earg	The extra arguments, in order.

Note

Presently, the multinomial logit model has only one link function, [multilogitlink](#), so a warning is not issued for that link. For other models, if the number of link functions does not equal M then a warning may be issued.

Author(s)

Thomas W. Yee

See Also[linkfun](#), [multilogitlink](#), [vglm](#).**Examples**

```
pneumo <- transform(pneumo, let = log(exposure.time))
fit1 <- vglm(cbind(normal, mild, severe) ~ let, propodds, data = pneumo)
coef(fit1, matrix = TRUE)
linkfun(fit1)
linkfun(fit1, earg = TRUE)

fit2 <- vglm(cbind(normal, mild, severe) ~ let, multinomial, data = pneumo)
coef(fit2, matrix = TRUE)
linkfun(fit2)
linkfun(fit2, earg = TRUE)
```

Links

*Link functions for VGLM/VGAM/etc. families***Description**

The **VGAM** package provides a number of (parameter) link functions which are described in general here. Collectively, they offer the user considerable choice and flexibility for modelling data.

Usage

```
TypicalVGAMlink(theta, someParameter = 0, bvalue = NULL, inverse = FALSE,
                 deriv = 0, short = TRUE, tag = FALSE)
```

Arguments

theta	Numeric or character. This is usually θ (default) but can sometimes be η , depending on the other arguments. If theta is character then inverse and deriv are ignored. The name theta should always be the name of the first argument.
someParameter	Some parameter, e.g., an offset.
bvalue	Boundary value, positive if given. If $0 < \theta$ then values of theta which are less than or equal to 0 can be replaced by bvalue before computing the link function value. Values of theta which are greater than or equal to 1 can be replaced by 1 minus bvalue before computing the link function value. The value <code>bvalue = .Machine\$double.eps</code> is sometimes a reasonable value, or something slightly higher.
inverse	Logical. If TRUE and <code>deriv = 0</code> then the inverse link value θ is returned, hence the argument theta is really η . In all other cases, the argument theta is really θ .

deriv	Integer. Either 0, 1, or 2, specifying the order of the derivative. Most link functions handle values up to 3 or 4. Some links can even handle values up to 9 but may suffer from catastrophic cancellation near the boundaries, be inefficient and slow because they are based on <code>D</code> .
short, tag	Logical. These are used for labelling the blurb slot of a <code>vglmff-class</code> object. These arguments are used only if <code>theta</code> is character, and gives the formula for the link in character form. If <code>tag = TRUE</code> then the result is preceded by a little more information.

Details

Almost all **VGAM** link functions have something similar to the argument list as given above. In this help file we have $\eta = g(\theta)$ where g is the link function, θ is the parameter and η is the linear/additive predictor. The link g must be strictly monotonic and twice-differentiable in its range.

The following is a brief enumeration of all **VGAM** link functions.

For parameters lying between 0 and 1 (e.g., probabilities): [logitlink](#), [probitlink](#), [clogloglink](#), [cauchitlink](#), [foldsqrtlink](#), [logclink](#).

For positive parameters (i.e., greater than 0): [loglink](#), [negloglink](#), [sqrtlink](#), [powerlink](#).

For parameters greater than 1: [logloglink](#), [loglogloglink](#) (greater than e).

For parameters between -1 and 1 : [fisherzlink](#), [rhobitlink](#).

For parameters between finite A and B : [extlogitlink](#), [logofflink](#) ($B = \infty$).

For unrestricted parameters (i.e., any value): [identitylink](#), [negidentitylink](#), [reciprocallink](#), [negreciprocallink](#).

Value

Returns one of: the link function value or its first or second derivative, the inverse link or its first or second derivative, or a character description of the link.

Here are the general details. If `inverse = FALSE` and `deriv = 0` (default) then the ordinary link function $\eta = g(\theta)$ is returned.

If `inverse = TRUE` and `deriv = 0` then the inverse link function value is returned, hence `theta` is really η (the only occasion this happens).

If `inverse = FALSE` and `deriv = 1` then it is $d\eta/d\theta$ as a function of θ . If `inverse = FALSE` and `deriv = 2` then it is $d^2\eta/d\theta^2$ as a function of θ .

If `inverse = TRUE` and `deriv = 1` then it is $d\theta/d\eta$ as a function of θ . If `inverse = TRUE` and `deriv = 2` then it is $d^2\theta/d\eta^2$ as a function of θ .

It is only when `deriv = 1` that `linkfun(theta, deriv = 1, inverse = TRUE)` and `linkfun(theta, deriv = 1, inverse = FALSE)` are *reciprocals* of each other. In particular, `linkfun(theta, deriv = 2, inverse = TRUE)` and `linkfun(theta, deriv = 2, inverse = FALSE)` are *not* reciprocals of each other in general.

Warning

The output of link functions changed at **VGAM** 0.9-9 (date was around 2015-07). Formerly, `linkfun(theta, deriv = 1)` is now `linkfun(theta, deriv = 1, inverse = TRUE)`, or equivalently,

$1 / \text{linkfun}(\text{theta}, \text{deriv} = 1, \text{inverse} = \text{TRUE})$. Also, formerly, $\text{linkfun}(\text{theta}, \text{deriv} = 2)$ was $1 / \text{linkfun}(\text{theta}, \text{deriv} = 2, \text{inverse} = \text{TRUE})$. This was a bug. Altogether, these are big changes and the user should beware!

In **VGAM** 1.0-7 (January 2019) all link function names were made to end in the characters "link", e.g., `loglink` replaces `loge`, `logitlink` replaces `logit`. For this most of them were renamed. Upward compatibility holds for older link function names, however, users should adopt the new names immediately.

Note

VGAM link functions are generally not compatible with other functions outside the package. In particular, they won't work with `glm` or any other package for fitting GAMs.

From October 2006 onwards, all **VGAM** family functions will only contain one default value for each link argument rather than giving a vector of choices. For example, rather than `binomialff(link = c("logitlink", "probitlink", "clogloglink", "cauchitlink", "identitylink"), ...)` it is now `binomialff(link = "logitlink", ...)`. No checking will be done to see if the user's choice is reasonable. This means that the user can write his/her own **VGAM** link function and use it within any **VGAM** family function. Altogether this provides greater flexibility. The downside is that the user must specify the *full* name of the link function, by either assigning the link argument the full name as a character string, or just the name itself. See the examples below.

From August 2012 onwards, a major change in link functions occurred. Argument `esigma` (and the like such as `earg`) used to be in **VGAM** prior to version 0.9-0 (released during the 2nd half of 2012). The major change is that arguments such as `offset` that used to be passed in via those arguments can now be done directly through the link function. For example, `gev(lshape = "logofflink", eshape = list(offset = 0.5))` is replaced by `gev(lshape = logofflink(offset = 0.5))`. The `@misc` slot no longer has `link` and `earg` components, but two other components replace these. Functions such as `dtheta.deta()`, `d2theta.deta2()`, `d3theta.deta3()`, `eta2theta()`, `theta2eta()` are modified.

From January 2019 onwards, all link function names ended in "link". See above for details.

Author(s)

T. W. Yee

References

McCullagh, P. and Nelder, J. A. (1989). *Generalized Linear Models*, 2nd ed. London: Chapman & Hall.

See Also

[TypicalVGAMfamilyFunction](#), [linkfun](#), [vglm](#), [vgam](#), [rrvglm](#), [cgo](#), [cao](#).

Examples

```
logitlink("a")
logitlink("a", short = FALSE)
logitlink("a", short = FALSE, tag = TRUE)
```

```

logofflink(1:5, offset = 1) # Same as log(1:5 + 1)
powerlink(1:5, power = 2) # Same as (1:5)^2

## Not run: # This is old and no longer works:
logofflink(1:5, earg = list(offset = 1))
powerlink(1:5, earg = list(power = 2))

## End(Not run)

fit1 <- vgam(agaaus ~ altitude,
            binomialff(link = "clogloglink"), hunua) # best
fit2 <- vgam(agaaus ~ altitude,
            binomialff(link = clogloglink ), hunua) # okay

## Not run:
# This no longer works since "clog" is not a valid VGAM link function:
fit3 <- vgam(agaaus ~ altitude,
            binomialff(link = "clog"), hunua) # not okay

# No matter what the link, the estimated var-cov matrix is the same
y <- rbeta(n = 1000, shape1 = exp(0), shape2 = exp(1))
fit1 <- vglm(y ~ 1, betaR(lshape1 = "identitylink",
                        lshape2 = "identitylink"),
            trace = TRUE, crit = "coef")
fit2 <- vglm(y ~ 1, betaR(lshape1 = logofflink(offset = 1.1),
                        lshape2 = logofflink(offset = 1.1)), trace=TRUE)
vcov(fit1, untransform = TRUE)
vcov(fit1, untransform = TRUE) -
vcov(fit2, untransform = TRUE) # Should be all 0s
\dontrun{ # This is old:
fit1@misc$earg # Some 'special' parameters
fit2@misc$earg # Some 'special' parameters are here
}

par(mfrow = c(2, 2))
p <- seq(0.05, 0.95, len = 200) # A rather restricted range
x <- seq(-4, 4, len = 200)
plot(p, logitlink(p), type = "l", col = "blue")
plot(x, logitlink(x, inverse = TRUE), type = "l", col = "blue")
plot(p, logitlink(p, deriv=1), type="l", col="blue") # 1 / (p*(1-p))
plot(p, logitlink(p, deriv=2), type="l", col="blue") # (2*p-1)/(p*(1-p))^2

## End(Not run)

```

Description

Density, distribution function, quantile function and random generation for the generalized beta distribution, as proposed by Libby and Novick (1982).

Usage

```
dlino(x, shape1, shape2, lambda = 1, log = FALSE)
plino(q, shape1, shape2, lambda = 1, lower.tail = TRUE, log.p = FALSE)
qlino(p, shape1, shape2, lambda = 1, lower.tail = TRUE, log.p = FALSE)
rlino(n, shape1, shape2, lambda = 1)
```

Arguments

x, q	vector of quantiles.
p	vector of probabilities.
n	number of observations. Same as in runif .
shape1, shape2, lambda	see lino .
log	Logical. If log = TRUE then the logarithm of the density is returned.
lower.tail, log.p	Same meaning as in pnorm or qnorm .

Details

See [lino](#), the **VGAM** family function for estimating the parameters, for the formula of the probability density function and other details.

Value

dlino gives the density, plino gives the distribution function, qlino gives the quantile function, and rlino generates random deviates.

Author(s)

T. W. Yee and Kai Huang

See Also

[lino](#).

Examples

```
## Not run: lambda <- 0.4; shape1 <- exp(1.3); shape2 <- exp(1.3)
x <- seq(0.0, 1.0, len = 101)
plot(x, dlino(x, shape1 = shape1, shape2 = shape2, lambda = lambda),
     type = "l", col = "blue", las = 1, ylab = "",
     main = "Blue is PDF, orange is the CDF",
     sub = "Purple lines are the 10,20,...,90 percentiles")
abline(h = 0, col = "blue", lty = 2)
```

```

lines(x, plino(x, shape1, shape2, lambda = lambda), col = "orange")
probs <- seq(0.1, 0.9, by = 0.1)
Q <- qlino(probs, shape1 = shape1, shape2 = shape2, lambda = lambda)
lines(Q, dlino(Q, shape1 = shape1, shape2 = shape2, lambda = lambda),
      col = "purple", lty = 3, type = "h")
plino(Q, shape1, shape2, lambda = lambda) - probs # Should be all 0

## End(Not run)

```

lino

*Generalized Beta Distribution Family Function***Description**

Maximum likelihood estimation of the 3-parameter generalized beta distribution as proposed by Libby and Novick (1982).

Usage

```

lino(lshape1 = "loglink", lshape2 = "loglink", llambda = "loglink",
     ishape1 = NULL, ishape2 = NULL, ilambda = 1, zero = NULL)

```

Arguments

lshape1, lshape2	Parameter link functions applied to the two (positive) shape parameters a and b . See Links for more choices.
llambda	Parameter link function applied to the parameter λ . See Links for more choices.
ishape1, ishape2, ilambda	Initial values for the parameters. A NULL value means one is computed internally. The argument <code>ilambda</code> must be numeric, and the default corresponds to a standard beta distribution.
zero	Can be an integer-valued vector specifying which linear/additive predictors are modelled as intercepts only. Here, the values must be from the set $\{1,2,3\}$ which correspond to a, b, λ , respectively. See CommonVGAMffArguments for more information.

Details

Proposed by Libby and Novick (1982), this distribution has density

$$f(y; a, b, \lambda) = \frac{\lambda^a y^{a-1} (1-y)^{b-1}}{B(a, b) \{1 - (1-\lambda)y\}^{a+b}}$$

for $a > 0$, $b > 0$, $\lambda > 0$, $0 < y < 1$. Here B is the beta function (see [beta](#)). The mean is a complicated function involving the Gauss hypergeometric function. If X has a lino distribution with parameters `shape1`, `shape2`, `lambda`, then $Y = \lambda X / (1 - (1 - \lambda)X)$ has a standard beta distribution with parameters `shape1`, `shape2`.

Since $\log(\lambda) = 0$ corresponds to the standard beta distribution, a summary of the fitted model performs a t-test for whether the data belongs to a standard beta distribution (provided the [loglink](#) link for λ is used; this is the default).

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Note

The fitted values, which is usually the mean, have not been implemented yet. Currently the median is returned as the fitted values.

Although Fisher scoring is used, the working weight matrices are positive-definite only in a certain region of the parameter space. Problems with this indicate poor initial values or an ill-conditioned model or insufficient data etc.

This model is can be difficult to fit. A reasonably good value of `ilambda` seems to be needed so if the self-starting initial values fail, try experimenting with the initial value arguments. Experience suggests `ilambda` is better a little larger, rather than smaller, compared to the true value.

Author(s)

T. W. Yee

References

Libby, D. L. and Novick, M. R. (1982). Multivariate generalized beta distributions with applications to utility assessment. *Journal of Educational Statistics*, **7**, 271–294.

Gupta, A. K. and Nadarajah, S. (2004). *Handbook of Beta Distribution and Its Applications*, NY: Marcel Dekker, Inc.

See Also

[Lino](#), [genbetaII](#).

Examples

```
ldata <- data.frame(y1 = rbeta(n = 1000, exp(0.5), exp(1))) # Std beta
fit <- vglm(y1 ~ 1, lino, data = ldata, trace = TRUE)
coef(fit, matrix = TRUE)
Coef(fit)
head(fitted(fit))
summary(fit)

# Nonstandard beta distribution
ldata <- transform(ldata, y2 = rlino(1000, shape1 = exp(1),
                                   shape2 = exp(2), lambda = exp(1)))
fit2 <- vglm(y2 ~ 1,
            lino(lshape1 = "identitylink", lshape2 = "identitylink",
                ilamb = 10), data = ldata, trace = TRUE)
```

```
coef(fit2, matrix = TRUE)
```

lirat	<i>Low-iron Rat Teratology Data</i>
-------	-------------------------------------

Description

Low-iron rat teratology data.

Usage

```
data(lirat)
```

Format

A data frame with 58 observations on the following 4 variables.

N Litter size.

R Number of dead fetuses.

hb Hemoglobin level.

grp Group number. Group 1 is the untreated (low-iron) group, group 2 received injections on day 7 or day 10 only, group 3 received injections on days 0 and 7, and group 4 received injections weekly.

Details

The following description comes from Moore and Tsiatis (1991). The data comes from the experimental setup from Shepard et al. (1980), which is typical of studies of the effects of chemical agents or dietary regimens on fetal development in laboratory rats.

Female rats were put in iron-deficient diets and divided into 4 groups. One group of controls was given weekly injections of iron supplement to bring their iron intake to normal levels, while another group was given only placebo injections. Two other groups were given fewer iron-supplement injections than the controls. The rats were made pregnant, sacrificed 3 weeks later, and the total number of fetuses and the number of dead fetuses in each litter were counted.

For each litter the number of dead fetuses may be considered to be $\text{Binomial}(N, p)$ where N is the litter size and p is the probability of a fetus dying. The parameter p is expected to vary from litter to litter, therefore the total variance of the proportions will be greater than that predicted by a binomial model, even when the covariates for hemoglobin level and experimental group are accounted for.

Source

Moore, D. F. and Tsiatis, A. (1991) Robust Estimation of the Variance in Moment Methods for Extra-binomial and Extra-Poisson Variation. *Biometrics*, **47**, 383–401.

References

Shepard, T. H., Mackler, B. and Finch, C. A. (1980). Reproductive studies in the iron-deficient rat. *Teratology*, **22**, 329–334.

Examples

```
## Not run:
# cf. Figure 3 of Moore and Tsiatis (1991)
plot(R / N ~ hb, data = lirat, pch = as.character(grp), col = grp,
     las = 1, xlab = "Hemoglobin level", ylab = "Proportion Dead")
## End(Not run)
```

lms.bcg

LMS Quantile Regression with a Box-Cox transformation to a Gamma Distribution

Description

LMS quantile regression with the Box-Cox transformation to the gamma distribution.

Usage

```
lms.bcg(percentiles = c(25, 50, 75), zero = c("lambda", "sigma"),
        llambda = "identitylink", lmu = "identitylink", lsigma = "loglink",
        idf.mu = 4, idf.sigma = 2, ilambda = 1, isigma = NULL)
```

Arguments

percentiles A numerical vector containing values between 0 and 100, which are the quantiles. They will be returned as ‘fitted values’.

zero See [lms.bcn](#). See [CommonVGAMffArguments](#) for more information.

llambda, lmu, lsigma See [lms.bcn](#).

idf.mu, idf.sigma See [lms.bcn](#).

ilambda, isigma See [lms.bcn](#).

Details

Given a value of the covariate, this function applies a Box-Cox transformation to the response to best obtain a gamma distribution. The parameters chosen to do this are estimated by maximum likelihood or penalized maximum likelihood. Similar details can be found at [lms.bcn](#).

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), [rrvglm](#) and [vgam](#).

Warning

This **VGAM** family function comes with the same warnings as [lms.bcn](#). Also, the expected value of the second derivative with respect to lambda may be incorrect (my calculations do not agree with the Lopatzidis and Green manuscript.)

Note

Similar notes can be found at [lms.bcn](#).

Author(s)

Thomas W. Yee

References

Lopatatzidis A. and Green, P. J. (unpublished manuscript). Semiparametric quantile regression using the gamma distribution.

Yee, T. W. (2004). Quantile regression via vector generalized additive models. *Statistics in Medicine*, **23**, 2295–2315.

See Also

[lms.bcn](#), [lms.yjn](#), [qtplot.lmscreg](#), [deplot.lmscreg](#), [cdf.lmscreg](#), [bmi.nz](#), [amlexponential](#).

Examples

```
# This converges, but deplot(fit) and qtplot(fit) do not work
fit0 <- vglm(BMI ~ sm.bs(age, df = 4), lms.bcg, bmi.nz, trace = TRUE)
coef(fit0, matrix = TRUE)
## Not run:
par(mfrow = c(1, 1))
plotvgam(fit0, se = TRUE) # Plot mu function (only)

## End(Not run)

# Use a trick: fit0 is used for initial values for fit1.
fit1 <- vgam(BMI ~ s(age, df = c(4, 2)), etastart = predict(fit0),
            lms.bcg(zero = 1), bmi.nz, trace = TRUE)

# Difficult to get a model that converges. Here, we prematurely
# stop iterations because it fails near the solution.
fit2 <- vgam(BMI ~ s(age, df = c(4, 2)), maxit = 4,
            lms.bcg(zero = 1, ilam = 3), bmi.nz, trace = TRUE)
summary(fit1)
head(predict(fit1))
head(fitted(fit1))
head(bmi.nz)
# Person 1 is near the lower quartile of BMI amongst people his age
head(cdf(fit1))

## Not run:
# Quantile plot
par(bty = "l", mar=c(5, 4, 4, 3) + 0.1, xpd = TRUE)
qtplot(fit1, percentiles=c(5, 50, 90, 99), main = "Quantiles",
       xlim = c(15, 90), las = 1, ylab = "BMI", lwd = 2, lcol = 4)

# Density plot
```

```

ygrid <- seq(15, 43, len = 100) # BMI ranges
par(mfrow = c(1, 1), lwd = 2)
(aa <- deplot(fit1, x0 = 20, y = ygrid, xlab = "BMI", col = "black",
  main = "PDFs at Age = 20 (black), 42 (red) and 55 (blue)"))
aa <- deplot(fit1, x0 = 42, y = ygrid, add=TRUE, llty=2, col="red")
aa <- deplot(fit1, x0 = 55, y = ygrid, add=TRUE, llty=4, col="blue",
  Attach = TRUE)
aa@post$deplot # Contains density function values

## End(Not run)

```

lms.bcn

LMS Quantile Regression with a Box-Cox Transformation to Normality

Description

LMS quantile regression with the Box-Cox transformation to normality.

Usage

```

lms.bcn(percentiles = c(25, 50, 75), zero = c("lambda", "sigma"),
  llambda = "identitylink", lmu = "identitylink",
  lsigma = "loglink", idf.mu = 4, idf.sigma = 2, ilambda = 1,
  isigma = NULL, tol0 = 0.001)

```

Arguments

percentiles	A numerical vector containing values between 0 and 100, which are the quantiles. They will be returned as ‘fitted values’.
zero	Can be an integer-valued vector specifying which linear/additive predictors are modelled as intercepts only. The values must be from the set {1,2,3}. The default value usually increases the chance of successful convergence. Setting zero = NULL means they all are functions of the covariates. For more information see CommonVGAMffArguments .
llambda, lmu, lsigma	Parameter link functions applied to the first, second and third linear/additive predictors. See Links for more choices, and CommonVGAMffArguments .
idf.mu	Degrees of freedom for the cubic smoothing spline fit applied to get an initial estimate of mu. See vsMOOTH.spline .
idf.sigma	Degrees of freedom for the cubic smoothing spline fit applied to get an initial estimate of sigma. See vsMOOTH.spline . This argument may be assigned NULL to get an initial value using some other algorithm.
ilambda	Initial value for lambda. If necessary, it is recycled to be a vector of length n where n is the number of (independent) observations.

isigma	Optional initial value for sigma. If necessary, it is recycled to be a vector of length n . The default value, NULL, means an initial value is computed in the @initialize slot of the family function.
tol0	Small positive number, the tolerance for testing if lambda is equal to zero.

Details

Given a value of the covariate, this function applies a Box-Cox transformation to the response to best obtain normality. The parameters chosen to do this are estimated by maximum likelihood or penalized maximum likelihood.

In more detail, the basic idea behind this method is that, for a fixed value of x , a Box-Cox transformation of the response Y is applied to obtain standard normality. The 3 parameters (λ , μ , σ , which start with the letters “L-M-S” respectively, hence its name) are chosen to maximize a penalized log-likelihood (with [vgam](#)). Then the appropriate quantiles of the standard normal distribution are back-transformed onto the original scale to get the desired quantiles. The three parameters may vary as a smooth function of x .

The Box-Cox power transformation here of the Y , given x , is

$$Z = [(Y/\mu(x))^{\lambda(x)} - 1]/(\sigma(x) \lambda(x))$$

for $\lambda(x) \neq 0$. (The singularity at $\lambda(x) = 0$ is handled by a simple function involving a logarithm.) Then Z is assumed to have a standard normal distribution. The parameter $\sigma(x)$ must be positive, therefore **VGAM** chooses $\eta(x)^T = (\lambda(x), \mu(x), \log(\sigma(x)))$ by default. The parameter μ is also positive, but while $\log(\mu)$ is available, it is not the default because μ is more directly interpretable. Given the estimated linear/additive predictors, the 100α percentile can be estimated by inverting the Box-Cox power transformation at the 100α percentile of the standard normal distribution.

Of the three functions, it is often a good idea to allow $\mu(x)$ to be more flexible because the functions $\lambda(x)$ and $\sigma(x)$ usually vary more smoothly with x . This is somewhat reflected in the default value for the argument zero, viz. zero = c(1, 3).

Value

An object of class “vglmff” (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), [rrvglm](#) and [vgam](#).

Warning

The computations are not simple, therefore convergence may fail. Set trace = TRUE to monitor convergence if it isn’t set already. Convergence failure will occur if, e.g., the response is bimodal at any particular value of x . In case of convergence failure, try different starting values. Also, the estimate may diverge quickly near the solution, in which case try prematurely stopping the iterations by assigning maxits to be the iteration number corresponding to the highest likelihood value.

One trick is to fit a simple model and use it to provide initial values for a more complex model; see in the examples below.

Note

The response must be positive because the Box-Cox transformation cannot handle negative values. In theory, the LMS-Yeo-Johnson-normal method can handle both positive and negative values.

In general, the lambda and sigma functions should be more smoother than the mean function. Having $\text{zero} = 1$, $\text{zero} = 3$ or $\text{zero} = c(1, 3)$ is often a good idea. See the example below.

Author(s)

Thomas W. Yee

References

Cole, T. J. and Green, P. J. (1992). Smoothing Reference Centile Curves: The LMS Method and Penalized Likelihood. *Statistics in Medicine*, **11**, 1305–1319.

Green, P. J. and Silverman, B. W. (1994). *Nonparametric Regression and Generalized Linear Models: A Roughness Penalty Approach*, London: Chapman & Hall.

Yee, T. W. (2004). Quantile regression via vector generalized additive models. *Statistics in Medicine*, **23**, 2295–2315.

See Also

[lms.bcg](#), [lms.yjn](#), [qtplot.lmscreg](#), [deplot.lmscreg](#), [cdf.lmscreg](#), [eCDF](#), [extlogF1](#), [alaplace1](#), [amlnormal](#), [denorm](#), [CommonVGAMffArguments](#).

Examples

```
## Not run: require("VGAMdata")
mysub <- subset(xs.nz, sex == "M" & ethnicity == "Maori" & study1)
mysub <- transform(mysub, BMI = weight / height^2)
BMIData <- na.omit(mysub)
BMIData <- subset(BMIData, BMI < 80 & age < 65,
                 select = c(age, BMI)) # Delete an outlier
summary(BMIData)

fit <- vgam(BMI ~ s(age, df = c(4, 2)), lms.bcn(zero = 1), BMIData)

par(mfrow = c(1, 2))
plot(fit, scol = "blue", se = TRUE) # The two centered smooths

head(predict(fit))
head(fitted(fit))
head(BMIData)
head(cdf(fit)) # Person 46 is probably overweight, given his age
100 * colMeans(c(deivar(fit)) < fitted(fit)) # Empirical proportions

# Correct for "vgam" objects but not very elegant:
fit@family@linkinv(eta = predict(fit, data.frame(age = 60)),
                  extra = list(percentiles = c(10, 50)))

if (FALSE) {
```

```

# These work for "vglm" objects:
fit2 <- vglm(BMI ~ bs(age, df = 4), lms.bcn(zero = 3), BMIdata)
predict(fit2, percentiles = c(10, 50),
        newdata = data.frame(age = 60), type = "response")
head(fitted(fit2, percentiles = c(10, 50))) # Different percentiles
}

# Convergence problems? Use fit0 for initial values for fit1
fit0 <- vgam(BMI ~ s(age, df = 4), lms.bcn(zero = c(1, 3)), BMIdata)
fit1 <- vgam(BMI ~ s(age, df = c(4, 2)), lms.bcn(zero = 1), BMIdata,
            etastart = predict(fit0))

## End(Not run)

## Not run: # Quantile plot
par(bty = "l", mar = c(5, 4, 4, 3) + 0.1, xpd = TRUE)
qtplot(fit, percentiles = c(5, 50, 90, 99), main = "Quantiles",
       xlim = c(15, 66), las = 1, ylab = "BMI", lwd = 2, lcol = 4)

# Density plot
ygrid <- seq(15, 43, len = 100) # BMI ranges
par(mfrow = c(1, 1), lwd = 2)
(aa <- deplot(fit, x0 = 20, y = ygrid, xlab = "BMI", col = "black",
             main = "PDFs at Age = 20 (black), 42 (red) and 55 (blue)"))
aa <- deplot(fit, x0 = 42, y = ygrid, add = TRUE, llty = 2, col = "red")
aa <- deplot(fit, x0 = 55, y = ygrid, add = TRUE, llty = 4, col = "blue",
            Attach = TRUE)
aa@post$deplot # Contains density function values

## End(Not run)

```

lms.yjn

LMS Quantile Regression with a Yeo-Johnson Transformation to Normality

Description

LMS quantile regression with the Yeo-Johnson transformation to normality. This family function is experimental and the LMS-BCN family function is recommended instead.

Usage

```

lms.yjn(percentiles = c(25, 50, 75), zero = c("lambda", "sigma"),
        llambda = "identitylink", lsigma = "loglink",
        idf.mu = 4, idf.sigma = 2,
        ilambda = 1, isigma = NULL, rule = c(10, 5),
        yoffset = NULL, diagW = FALSE, iters.diagW = 6)
lms.yjn2(percentiles = c(25, 50, 75), zero = c("lambda", "sigma"),
         llambda = "identitylink", lmu = "identitylink", lsigma = "loglink",
         idf.mu = 4, idf.sigma = 2, ilambda = 1.0,
         isigma = NULL, yoffset = NULL, nsimEIM = 250)

```

Arguments

percentiles	A numerical vector containing values between 0 and 100, which are the quantiles. They will be returned as ‘fitted values’.
zero	See lms.bcn . See CommonVGAMffArguments for more information.
llambda, lmu, lsigma	See lms.bcn .
idf.mu, idf.sigma	See lms.bcn .
ilambda, isigma	See lms.bcn .
rule	Number of abscissae used in the Gaussian integration scheme to work out elements of the weight matrices. The values given are the possible choices, with the first value being the default. The larger the value, the more accurate the approximation is likely to be but involving more computational expense.
yoffset	A value to be added to the response y , for the purpose of centering the response before fitting the model to the data. The default value, NULL, means $-\text{median}(y)$ is used, so that the response actually used has median zero. The <code>yoffset</code> is saved on the object and used during prediction.
diagW	Logical. This argument is offered because the expected information matrix may not be positive-definite. Using the diagonal elements of this matrix results in a higher chance of it being positive-definite, however convergence will be very slow. If TRUE, then the first <code>iters.diagW</code> iterations will use the diagonal of the expected information matrix. The default is FALSE, meaning faster convergence.
iters.diagW	Integer. Number of iterations in which the diagonal elements of the expected information matrix are used. Only used if <code>diagW = TRUE</code> .
nsimEIM	See CommonVGAMffArguments for more information.

Details

Given a value of the covariate, this function applies a Yeo-Johnson transformation to the response to best obtain normality. The parameters chosen to do this are estimated by maximum likelihood or penalized maximum likelihood. The function `lms.yjn2()` estimates the expected information matrices using simulation (and is consequently slower) while `lms.yjn()` uses numerical integration. Try the other if one function fails.

Value

An object of class “`vglmff`” (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Warning

The computations are not simple, therefore convergence may fail. In that case, try different starting values.

The generic function `predict`, when applied to a `lms.yjn` fit, does not add back the `yoffset` value. As described above, this family function is experimental and the LMS-BCN family function is recommended instead.

Note

The response may contain both positive and negative values. In contrast, the LMS-Box-Cox-normal and LMS-Box-Cox-gamma methods only handle a positive response because the Box-Cox transformation cannot handle negative values.

Some other notes can be found at [lms.bcn](#).

Author(s)

Thomas W. Yee

References

Yeo, I.-K. and Johnson, R. A. (2000). A new family of power transformations to improve normality or symmetry. *Biometrika*, **87**, 954–959.

Yee, T. W. (2004). Quantile regression via vector generalized additive models. *Statistics in Medicine*, **23**, 2295–2315.

Yee, T. W. (2002). An Implementation for Regression Quantile Estimation. Pages 3–14. In: Haerle, W. and Ronz, B., *Proceedings in Computational Statistics COMPSTAT 2002*. Heidelberg: Physica-Verlag.

See Also

[lms.bcn](#), [lms.bcg](#), [qtplot.lmscreg](#), [deplot.lmscreg](#), [cdf.lmscreg](#), [bmi.nz](#), [amlnormal](#).

Examples

```
fit <- vgam(BMI ~ s(age, df = 4), lms.yjn, bmi.nz, trace = TRUE)
head(predict(fit))
head(fitted(fit))
head(bmi.nz)
# Person 1 is near the lower quartile of BMI amongst people his age
head(cdf(fit))

## Not run:
# Quantile plot
par(bty = "l", mar = c(5, 4, 4, 3) + 0.1, xpd = TRUE)
qtplot(fit, percentiles = c(5, 50, 90, 99), main = "Quantiles",
       xlim = c(15, 90), las = 1, ylab = "BMI", lwd = 2, lcol = 4)

# Density plot
ygrid <- seq(15, 43, len = 100) # BMI ranges
par(mfrow = c(1, 1), lwd = 2)
(Z <- deplot(fit, x0 = 20, y = ygrid, xlab = "BMI", col = "black",
            main = "PDFs at Age = 20 (black), 42 (red) and 55 (blue)"))
Z <- deplot(fit, x0 = 42, y = ygrid, add = TRUE, llty = 2, col = "red")
Z <- deplot(fit, x0 = 55, y = ygrid, add = TRUE, llty = 4, col = "blue",
            Attach = TRUE)
with(Z@post, deplot) # Contains PDF values; == a@post$deplot

## End(Not run)
```

Log

Logarithmic Distribution

Description

Density, distribution function, quantile function, and random generation for the logarithmic distribution.

Usage

```
dlog(x, shape, log = FALSE)
plog(q, shape, lower.tail = TRUE, log.p = FALSE)
qlog(p, shape)
rlog(n, shape)
```

Arguments

<code>x, q, p, n, lower.tail</code>	Same interpretation as in runif .
<code>shape</code>	The shape parameter value c described in in logff .
<code>log, log.p</code>	Logical. If <code>log.p = TRUE</code> then all probabilities p are given as $\log(p)$.

Details

The details are given in [logff](#).

Value

`dlog` gives the density, `plog` gives the distribution function, `qlog` gives the quantile function, and `rlog` generates random deviates.

Note

Given some response data, the **VGAM** family function [logff](#) estimates the parameter shape. For `plog()`, if argument `q` contains large values and/or `q` is long in length then the memory requirements may be very high. Very large values in `q` are handled by an approximation by Owen (1965).

Author(s)

T. W. Yee

References

Forbes, C., Evans, M., Hastings, N. and Peacock, B. (2011). *Statistical Distributions*, Hoboken, NJ, USA: John Wiley and Sons, Fourth edition.

See Also

[logff](#), [Gaitdlog](#), [Oilog](#), [Otlog](#).

Examples

```
dlog(1:20, 0.5)
rlog(20, 0.5)

## Not run: shape <- 0.8; x <- 1:10
plot(x, dlog(x, shape = shape), type = "h", ylim = 0:1,
      sub = "shape=0.8", las = 1, col = "blue", ylab = "shape",
      main = "Logarithmic distribution: blue=PDF; orange=CDF")
lines(x + 0.1, plog(x, shape), col = "orange", lty = 3, type = "h")
## End(Not run)
```

log1mexp

Logarithms with an Unit Offset and Exponential Term

Description

Computes $\log(1 + \exp(x))$ and $\log(1 - \exp(-x))$ accurately.

Usage

```
log1mexp(x)
log1pexp(x)
```

Arguments

`x` A vector of reals (numeric). Complex numbers not allowed since [expm1](#) and [log1p](#) do not handle these.

Details

Computes $\log(1 + \exp(x))$ and $\log(1 - \exp(-x))$ accurately. An adjustment is made when x is away from 0 in value.

Value

`log1mexp(x)` gives the value of $\log(1 - \exp(-x))$.

`log1pexp(x)` gives the value of $\log(1 + \exp(x))$.

Note

If NA or NaN is present in the input, the corresponding output will be NA.

Author(s)

This is a direct translation of the function in Martin Maechler's (2012) paper by Xiangjie Xue and T. W. Yee.

References

Maechler, Martin (2012). Accurately Computing $\log(1-\exp(-|a|))$. Assessed from the **Rmpfr** package.

See Also

[log1p](#), [expm1](#), [exp](#), [log](#)

Examples

```
x <- c(10, 50, 100, 200, 400, 500, 800, 1000, 1e4, 1e5, 1e20, Inf, NA)
log1pexp(x)
log(1 + exp(x)) # Naive; suffers from overflow
log1mexp(x)
log(1 - exp(-x))
y <- -x
log1pexp(y)
log(1 + exp(y)) # Naive; suffers from inaccuracy
```

logclink

Complementary-log Link Function

Description

Computes the Complementary-log Transformation, Including its Inverse and the First Two Derivatives.

Usage

```
logclink(theta, bvalue = NULL, inverse = FALSE, deriv = 0,
         short = TRUE, tag = FALSE)
```

Arguments

theta Numeric or character. See below for further details.
bvalue See [Links](#).
inverse, deriv, short, tag
 Details at [Links](#).

Details

The complementary-log link function is suitable for parameters that are less than unity. Numerical values of theta close to 1 or out of range result in Inf, -Inf, NA or NaN.

Value

For `deriv = 0`, the log of theta, i.e., $\log(1-\text{theta})$ when `inverse = FALSE`, and if `inverse = TRUE` then $1-\exp(\text{theta})$.

For `deriv = 1`, then the function returns $d \text{ eta} / d \text{ theta}$ as a function of theta if `inverse = FALSE`, else if `inverse = TRUE` then it returns the reciprocal.

Here, all logarithms are natural logarithms, i.e., to base e .

Note

Numerical instability may occur when theta is close to 1. One way of overcoming this is to use `bvalue`.

Author(s)

Thomas W. Yee

References

McCullagh, P. and Nelder, J. A. (1989). *Generalized Linear Models*, 2nd ed. London: Chapman & Hall.

See Also

[Links](#), [loglink](#), [clogloglink](#), [logloglink](#), [logofflink](#).

Examples

```
## Not run:
logclink(seq(-0.2, 1.1, by = 0.1)) # Has NAs

## End(Not run)
logclink(seq(-0.2, 1.1, by=0.1), bvalue=1-.Machine$double.eps) # Has no NAs
```

logF

Natural Exponential Family Generalized Hyperbolic Secant Distribution Family Function

Description

Maximum likelihood estimation of the 2-parameter log F distribution.

Usage

```
logF(lshape1 = "loglink", lshape2 = "loglink",
     ishape1 = NULL, ishape2 = 1, imethod = 1)
```

Arguments

lshape1, lshape2	Parameter link functions for the shape parameters. Called α and β respectively. See Links for more choices.
ishape1, ishape2	Optional initial values for the shape parameters. If given, it must be numeric and values are recycled to the appropriate length. The default is to choose the value internally. See CommonVGAMffArguments for more information.
imethod	Initialization method. Either the value 1, 2, or . . . See CommonVGAMffArguments for more information.

Details

The density for this distribution is

$$f(y; \alpha, \beta) = \exp(\alpha y) / [B(\alpha, \beta)(1 + e^y)^{\alpha + \beta}]$$

where y is real, $\alpha > 0$, $\beta > 0$, $B(., .)$ is the beta function [beta](#).

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Author(s)

Thomas W. Yee

References

Jones, M. C. (2008). On a class of distributions with simple exponential tails. *Statistica Sinica*, **18**(3), 1101–1110.

See Also

[dlogF](#), [extlogF1](#), [logff](#).

Examples

```
nn <- 1000
ldata <- data.frame(y1 = rnorm(nn, +1, sd = exp(2)), # Not proper data
                   x2 = rnorm(nn, -1, sd = exp(2)),
                   y2 = rnorm(nn, -1, sd = exp(2))) # Not proper data
fit1 <- vglm(y1 ~ 1, logF, ldata, trace = TRUE)
fit2 <- vglm(y2 ~ x2, logF, ldata, trace = TRUE)
coef(fit2, matrix = TRUE)
summary(fit2)
vcov(fit2)

head(fitted(fit1))
with(ldata, mean(y1))
max(abs(head(fitted(fit1)) - with(ldata, mean(y1))))
```

logff	<i>Logarithmic Distribution</i>
-------	---------------------------------

Description

Estimating the (single) parameter of the logarithmic distribution.

Usage

```
logff(lshape = "logitlink", gshape = -expm1(-7 * ppoints(4)),
      zero = NULL)
```

Arguments

lshape	Parameter link function for the parameter c , which lies between 0 and 1. See Links for more choices and information. Soon <code>logfflink()</code> will hopefully be available for event-rate data.
gshape, zero	Details at CommonVGAMffArguments . Practical experience shows that having the initial value for c being close to the solution is quite important.

Details

The logarithmic distribution is a generalized power series distribution that is based specifically on the logarithmic series (scaled to a probability function). Its probability function is $f(y) = ac^y/y$, for $y = 1, 2, 3, \dots$, where $0 < c < 1$ (called shape), and $a = -1/\log(1 - c)$. The mean is $ac/(1 - c)$ (returned as the fitted values) and variance is $ac(1 - ac)/(1 - c)^2$. When the sample mean is large, the value of c tends to be very close to 1, hence it could be argued that `logitlink` is not the best choice.

Value

An object of class "vg1mff" (see [vg1mff-class](#)). The object is used by modelling functions such as `vglm`, and `vgam`.

Note

The function `log` computes the natural logarithm. In the **VGAM** library, a link function with option `loglink` corresponds to this.

Multiple responses are permitted.

The “logarithmic distribution” has various meanings in the literature. Sometimes it is also called the *log-series distribution*. Some others call some continuous distribution on $[a, b]$ by the name “logarithmic distribution”.

Author(s)

T. W. Yee

References

- Johnson N. L., Kemp, A. W. and Kotz S. (2005). *Univariate Discrete Distributions*, 3rd edition, ch.7. Hoboken, New Jersey: Wiley.
- Forbes, C., Evans, M., Hastings, N. and Peacock, B. (2011) *Statistical Distributions*, Hoboken, NJ, USA: John Wiley and Sons, Fourth edition.

See Also

[Log](#), [gaitdlog](#), [oalog](#), [oilog](#), [otlog](#), [log](#), [loglink](#), [logofflink](#), [explogff](#), [simulate.vlm](#).

Examples

```
nn <- 1000
ldata <- data.frame(y = rlog(nn, shape = logitlink(0.2, inv = TRUE)))
fit <- vglm(y ~ 1, logff, data = ldata, trace = TRUE, crit = "c")
coef(fit, matrix = TRUE)
Coef(fit)
## Not run: with(ldata, spikeplot(y, col = "blue", capped = TRUE))
x <- seq(1, with(ldata, max(y)), by = 1)
with(ldata, lines(x + 0.1, dlog(x, Coef(fit)[1]), col = "orange",
                 type = "h", lwd = 2))
## End(Not run)

# Example: Corbet (1943) butterfly Malaya data
corbet <- data.frame(nindiv = 1:24,
                    ofreq = c(118, 74, 44, 24, 29, 22, 20, 19, 20, 15, 12,
                              14, 6, 12, 6, 9, 9, 6, 10, 10, 11, 5, 3, 3))
fit <- vglm(nindiv ~ 1, logff, data = corbet, weights = ofreq)
coef(fit, matrix = TRUE)
shapehat <- Coef(fit)["shape"]
pdf2 <- dlog(x = with(corbet, nindiv), shape = shapehat)
print(with(corbet, cbind(nindiv, ofreq, fitted = pdf2 * sum(ofreq))),
      digits = 1)
```

logistic

Logistic Distribution Family Function

Description

Estimates the location and scale parameters of the logistic distribution by maximum likelihood estimation.

Usage

```
logistic1(llocation = "identitylink", scale.arg = 1, imethod = 1)
logistic(llocation = "identitylink", lscale = "loglink",
         ilocation = NULL, iscale = NULL, imethod = 1, zero = "scale")
```

Arguments

llocation, lscale	Parameter link functions applied to the location parameter l and scale parameter s . See Links for more choices, and CommonVGAMffArguments for more information.
scale.arg	Known positive scale parameter (called s below).
ilocation, iscale	See CommonVGAMffArguments for information.
imethod, zero	See CommonVGAMffArguments for information.

Details

The two-parameter logistic distribution has a density that can be written as

$$f(y; l, s) = \frac{\exp[-(y-l)/s]}{s(1 + \exp[-(y-l)/s])^2}$$

where $s > 0$ is the scale parameter, and l is the location parameter. The response $-\infty < y < \infty$. The mean of Y (which is the fitted value) is l and its variance is $\pi^2 s^2 / 3$.

A logistic distribution with scale = 0.65 (see [dlogis](#)) resembles [dt](#) with $df = 7$; see [logistic1](#) and [studentt](#).

[logistic1](#) estimates the location parameter only while [logistic](#) estimates both parameters. By default, $\eta_1 = l$ and $\eta_2 = \log(s)$ for [logistic](#).

[logistic](#) can handle multiple responses.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), [rrvglm](#) and [vgam](#).

Note

Fisher scoring is used, and the Fisher information matrix is diagonal.

Author(s)

T. W. Yee

References

- Johnson, N. L. and Kotz, S. and Balakrishnan, N. (1994). *Continuous Univariate Distributions*, 2nd edition, Volume 1, New York: Wiley. Chapter 15.
- Forbes, C., Evans, M., Hastings, N. and Peacock, B. (2011). *Statistical Distributions*, Hoboken, NJ, USA: John Wiley and Sons, Fourth edition.
- Castillo, E., Hadi, A. S., Balakrishnan, N. and Sarabia, J. S. (2005). *Extreme Value and Related Models with Applications in Engineering and Science*, Hoboken, NJ, USA: Wiley-Interscience, p.130.
- deCani, J. S. and Stine, R. A. (1986). A Note on Deriving the Information Matrix for a Logistic Distribution, *The American Statistician*, **40**, 220–222.

See Also

[rlogis](#), [CommonVGAMffArguments](#), [logitlink](#), [gensh](#), [cumulative](#), [bilogistic](#), [simulate.vlm](#).

Examples

```
# Location unknown, scale known
ldata <- data.frame(x2 = runif(nn <- 500))
ldata <- transform(ldata, y1 = rlogis(nn, loc = 1+5*x2, sc = exp(2)))
fit1 <- vglm(y1 ~ x2, logistic1(scale = exp(2)), ldata, trace = TRUE)
coef(fit1, matrix = TRUE)

# Both location and scale unknown
ldata <- transform(ldata, y2 = rlogis(nn, loc = 1 + 5*x2, exp(x2)))
fit2 <- vglm(cbind(y1, y2) ~ x2, logistic, data = ldata, trace = TRUE)
coef(fit2, matrix = TRUE)
vcov(fit2)
summary(fit2)
```

logitlink

Logit Link Function

Description

Computes the logit transformation, including its inverse and the first nine derivatives.

Usage

```
logitlink(theta, bvalue = NULL, inverse = FALSE, deriv = 0,
  short = TRUE, tag = FALSE)
extlogitlink(theta, min = 0, max = 1, bminvalue = NULL,
  bmaxvalue = NULL, inverse = FALSE, deriv = 0,
  short = TRUE, tag = FALSE)
```

Arguments

theta Numeric or character. See below for further details.

bvalue, bminvalue, bmaxvalue See [Links](#). These are boundary values. For `extlogitlink`, values of `theta` less than or equal to A or greater than or equal to B can be replaced by `bminvalue` and `bmaxvalue`.

min, max For `extlogitlink`, `min` gives A , `max` gives B , and for out of range values, `bminvalue` and `bmaxvalue`.

inverse, deriv, short, tag Details at [Links](#).

Details

The logit link function is very commonly used for parameters that lie in the unit interval. It is the inverse CDF of the logistic distribution. Numerical values of θ close to 0 or 1 or out of range result in Inf, -Inf, NA or NaN.

The *extended* logit link function `extlogitlink` should be used more generally for parameters that lie in the interval (A, B) , say. The formula is

$$\log((\theta - A)/(B - \theta))$$

and the default values for A and B correspond to the ordinary logit function. Numerical values of θ close to A or B or out of range result in Inf, -Inf, NA or NaN. However these can be replaced by values `bminvalue` and `bmaxvalue` first before computing the link function.

Value

For `logitlink` with `deriv = 0`, the logit of θ , i.e., $\log(\theta/(1-\theta))$ when `inverse = FALSE`, and if `inverse = TRUE` then $\exp(\theta)/(1+\exp(\theta))$.

For `deriv = 1`, then the function returns $d\theta/d\theta$ as a function of θ if `inverse = FALSE`, else if `inverse = TRUE` then it returns the reciprocal.

Here, all logarithms are natural logarithms, i.e., to base e .

Note

Numerical instability may occur when θ is close to 1 or 0 (for `logitlink`), or close to A or B for `extlogitlink`. One way of overcoming this is to use, e.g., `bvalue`.

In terms of the threshold approach with cumulative probabilities for an ordinal response this link function corresponds to the univariate logistic distribution (see [logistic](#)).

Author(s)

Thomas W. Yee

References

McCullagh, P. and Nelder, J. A. (1989). *Generalized Linear Models*, 2nd ed. London: Chapman & Hall.

See Also

[Links](#), [alogitlink](#), [asinlink](#), [logitoffsetlink](#), [probitlink](#), [clogloglink](#), [cauchitlink](#), [logistic1](#), [loglink](#), [Logistic](#), [multilogitlink](#).

Examples

```
p <- seq(0.01, 0.99, by = 0.01)
logitlink(p)
max(abs(logitlink(logitlink(p), inverse = TRUE) - p)) # 0?

p <- c(seq(-0.02, 0.02, by = 0.01), seq(0.97, 1.02, by = 0.01))
```

```

logitlink(p) # Has NAs
logitlink(p, bvalue = .Machine$double.eps) # Has no NAs

p <- seq(0.9, 2.2, by = 0.1)
extlogitlink(p, min = 1, max = 2,
             bminvalue = 1 + .Machine$double.eps,
             bmaxvalue = 2 - .Machine$double.eps) # Has no NAs

## Not run: par(mfrow = c(2,2), lwd = (mylwd <- 2))
y <- seq(-4, 4, length = 100)
p <- seq(0.01, 0.99, by = 0.01)
for (d in 0:1) {
  myinv <- (d > 0)
  matplot(p, cbind( logitlink(p, deriv = d, inv = myinv),
                  probitlink(p, deriv = d, inv = myinv)), las = 1,
          type = "n", col = "purple", ylab = "transformation",
          main = if (d == 0) "Some probability link functions"
                else "1 / first derivative")
  lines(p, logitlink(p, deriv = d, inverse = myinv), col = "limegreen")
  lines(p, probitlink(p, deriv = d, inverse = myinv), col = "purple")
  lines(p, clogloglink(p, deriv = d, inverse = myinv), col = "chocolate")
  lines(p, cauchitlink(p, deriv = d, inverse = myinv), col = "tan")
  if (d == 0) {
    abline(v = 0.5, h = 0, lty = "dashed")
    legend(0, 4.5, c("logitlink", "probitlink",
                  "clogloglink", "cauchitlink"), col = c("limegreen", "purple",
                  "chocolate", "tan"), lwd = mylwd)
  } else
    abline(v = 0.5, lty = "dashed")
}

for (d in 0) {
  matplot(y, cbind(logitlink(y, deriv = d, inverse = TRUE),
                  probitlink(y, deriv = d, inverse = TRUE)), las = 1,
          type = "n", col = "purple", xlab = "transformation", ylab = "p",
          main = if (d == 0) "Some inverse probability link functions"
                else "First derivative")
  lines(y, logitlink(y, deriv = d, inv = TRUE), col = "limegreen")
  lines(y, probitlink(y, deriv = d, inv = TRUE), col = "purple")
  lines(y, clogloglink(y, deriv = d, inv = TRUE), col = "chocolate")
  lines(y, cauchitlink(y, deriv = d, inv = TRUE), col = "tan")
  if (d == 0) {
    abline(h = 0.5, v = 0, lty = "dashed")
    legend(-4, 1, c("logitlink", "probitlink", "clogloglink",
                  "cauchitlink"), col = c("limegreen", "purple",
                  "chocolate", "tan"), lwd = mylwd)
  }
}

p <- seq(0.21, 0.59, by = 0.01)
plot(p, extlogitlink(p, min = 0.2, max = 0.6), xlim = c(0, 1),
     type = "l", col = "black", ylab = "transformation",
     las = 1, main = "extlogitlink(p, min = 0.2, max = 0.6)")

```

```
par(lwd = 1)
## End(Not run)
```

logitoffsetlink	<i>Logit-with-an-Offset Link Function</i>
-----------------	---

Description

Computes the logitoffsetlink transformation, including its inverse and the first two derivatives.

Usage

```
logitoffsetlink(theta, offset = 0, inverse = FALSE, deriv = 0,
                short = TRUE, tag = FALSE)
```

Arguments

theta Numeric or character. See below for further details.
offset The offset value(s), which must be non-negative. It is called K below.
inverse, deriv, short, tag
 Details at [Links](#).

Details

This link function allows for some asymmetry compared to the ordinary [logitlink](#) link. The formula is

$$\log(\theta/(1 - \theta) - K)$$

and the default value for the offset K is corresponds to the ordinary [logitlink](#) link. When inverse = TRUE will mean that the value will lie in the interval $(K/(1 + K), 1)$.

Value

For logitoffsetlink with deriv = 0, the logitoffsetlink of theta, i.e., $\log(\theta/(1-\theta) - K)$ when inverse = FALSE, and if inverse = TRUE then $(K + \exp(\theta))/(1 + \exp(\theta) + K)$.

For deriv = 1, then the function returns $d \text{ eta} / d \text{ theta}$ as a function of theta if inverse = FALSE, else if inverse = TRUE then it returns the reciprocal.

Here, all logarithms are natural logarithms, i.e., to base e .

Note

This function is numerical less stability than [logitlink](#).

Author(s)

Thomas W. Yee

References

Komori, O. and Eguchi, S. et al., 2016. An asymmetric logistic model for ecological data. *Methods in Ecology and Evolution*, 7.

See Also

[Links](#), [logitlink](#).

Examples

```
p <- seq(0.05, 0.99, by = 0.01); myoff <- 0.05
logitoffsetlink(p, myoff)
max(abs(logitoffsetlink(logitoffsetlink(p, myoff),
  myoff, inverse = TRUE) - p)) # Should be 0
```

logLik.vlm

Extract Log-likelihood for VGMLs/VGAMs/etc.

Description

Calculates the log-likelihood value or the element-by-element contributions of the log-likelihood.

Usage

```
## S3 method for class 'vlm'
logLik(object, summation = TRUE, ...)
```

Arguments

object	Some VGAM object, for example, having class vglmff-class .
summation	Logical, apply sum ? If FALSE then a n -vector or n -row matrix (with the number of responses as the number of columns) is returned. Each element is the contribution to the log-likelihood.
...	Currently unused. In the future: other possible arguments fed into <code>logLik</code> in order to compute the log-likelihood.

Details

By default, this function returns the log-likelihood of the object. Thus this code relies on the log-likelihood being defined, and computed, for the object.

Value

Returns the log-likelihood of the object. If `summation = FALSE` then a n -vector or n -row matrix (with the number of responses as the number of columns) is returned. Each element is the contribution to the log-likelihood. The prior weights are assimilated within the answer.

Warning

Not all **VGAM** family functions have had the summation checked.

Note

Not all **VGAM** family functions currently have the summation argument implemented.

Author(s)

T. W. Yee.

See Also

VGLMs are described in [vglm-class](#); VGAMs are described in [vgam-class](#); RR-VGLMs are described in [rrvglm-class](#); [AIC](#); [anova.vglm](#).

Examples

```

zdata <- data.frame(x2 = runif(nn <- 50))
zdata <- transform(zdata, Ps01 = logitlink(-0.5      , inverse = TRUE),
                  Ps02 = logitlink( 0.5      , inverse = TRUE),
                  lambda1 = loglink(-0.5 + 2*x2, inverse = TRUE),
                  lambda2 = loglink( 0.5 + 2*x2, inverse = TRUE))
zdata <- transform(zdata, y1 = rzipois(nn, lambda = lambda1, pstr0 = Ps01),
                  y2 = rzipois(nn, lambda = lambda2, pstr0 = Ps02))

with(zdata, table(y1)) # Eyeball the data
with(zdata, table(y2))
fit2 <- vglm(cbind(y1, y2) ~ x2, zipoisson(zero = NULL), data = zdata)

logLik(fit2) # Summed over the two responses
sum(logLik(fit2, sum = FALSE)) # For checking purposes
(ll.matrix <- logLik(fit2, sum = FALSE)) # nn x 2 matrix
colSums(ll.matrix) # log-likelihood for each response

```

loglinb2

Loglinear Model for Two Binary Responses

Description

Fits a loglinear model to two binary responses.

Usage

```
loglinb2(exchangeable = FALSE, zero = "u12")
```

Arguments

- exchangeable Logical. If TRUE, the two marginal probabilities are constrained to be equal. Should be set TRUE for ears, eyes, etc. data.
- zero Which linear/additive predictors are modelled as intercept-only? A NULL means none of them. See [CommonVGAMffArguments](#) for more information.

Details

The model is

$$P(Y_1 = y_1, Y_2 = y_2) = \exp(u_0 + u_1 y_1 + u_2 y_2 + u_{12} y_1 y_2)$$

where y_1 and y_2 are 0 or 1, and the parameters are u_1, u_2, u_{12} . The normalizing parameter u_0 can be expressed as a function of the other parameters, viz.,

$$u_0 = -\log[1 + \exp(u_1) + \exp(u_2) + \exp(u_1 + u_2 + u_{12})].$$

The linear/additive predictors are $(\eta_1, \eta_2, \eta_3)^T = (u_1, u_2, u_{12})^T$.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), [rrvglm](#) and [vgam](#).

When fitted, the `fitted.values` slot of the object contains the four joint probabilities, labelled as $(Y_1, Y_2) = (0,0), (0,1), (1,0), (1,1)$, respectively.

Note

The response must be a two-column matrix of ones and zeros only. This is more restrictive than [binom2.or](#), which can handle more types of input formats. Note that each of the 4 combinations of the multivariate response need to appear in the data set. After estimation, the response attached to the object is also a two-column matrix; possibly in the future it might change into a four-column matrix.

Author(s)

Thomas W. Yee

References

Yee, T. W. and Wild, C. J. (2001). Discussion to: "Smoothing spline ANOVA for multivariate Bernoulli observations, with application to ophthalmology data (with discussion)" by Gao, F., Wahba, G., Klein, R., Klein, B. *Journal of the American Statistical Association*, **96**, 127–160.

McCullagh, P. and Nelder, J. A. (1989). *Generalized Linear Models*, 2nd ed. London: Chapman & Hall.

See Also

[binom2.or](#), [binom2.rho](#), [loglinb3](#).

Examples

```

coalminers <- transform(coalminers, Age = (age - 42) / 5)
# Get the n x 4 matrix of counts
fit0 <- vglm(cbind(nBnW,nBW,BnW,BW) ~ Age, binom2.or, coalminers)
counts <- round(c(weights(fit0, type = "prior") * depvar(fit0))
# Create a n x 2 matrix response for loglinb2()
# bwmat <- matrix(c(0,0, 0,1, 1,0, 1,1), 4, 2, byrow = TRUE)
bwmat <- cbind(bln = c(0,0,1,1), wheeze = c(0,1,0,1))
matof1 <- matrix(1, nrow(counts), 1)
newminers <-
  data.frame(bln      = kronecker(matof1, bwmat[, 1]),
             wheeze   = kronecker(matof1, bwmat[, 2]),
             wt       = c(t(counts)),
             Age      = with(coalminers, rep(age, rep(4, length(age))))))
newminers <- newminers[with(newminers, wt) > 0,]

fit <- vglm(cbind(bln,wheeze) ~ Age, loglinb2(zero = NULL),
           weight = wt, data = newminers)
coef(fit, matrix = TRUE) # Same! (at least for the log odds-ratio)
summary(fit)

# Try reconcile this with McCullagh and Nelder (1989), p.234
(0.166-0.131) / 0.027458 # 1.275 is approximately 1.25

```

loglinb3

*Loglinear Model for Three Binary Responses***Description**

Fits a loglinear model to three binary responses.

Usage

```
loglinb3(exchangeable = FALSE, zero = c("u12", "u13", "u23"))
```

Arguments

exchangeable Logical. If TRUE, the three marginal probabilities are constrained to be equal.

zero Which linear/additive predictors are modelled as intercept-only? A NULL means none. See [CommonVGAMffArguments](#) for further information.

Details

The model is $P(Y_1 = y_1, Y_2 = y_2, Y_3 = y_3) =$

$$\exp(u_0 + u_1 y_1 + u_2 y_2 + u_3 y_3 + u_{12} y_1 y_2 + u_{13} y_1 y_3 + u_{23} y_2 y_3)$$

where y_1 , y_2 and y_3 are 0 or 1, and the parameters are u_1 , u_2 , u_3 , u_{12} , u_{13} , u_{23} . The normalizing parameter u_0 can be expressed as a function of the other parameters. Note that a third-order association parameter, u_{123} for the product $y_1 y_2 y_3$, is assumed to be zero for this family function.

The linear/additive predictors are $(\eta_1, \eta_2, \dots, \eta_6)^T = (u_1, u_2, u_3, u_{12}, u_{13}, u_{23})^T$.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), [rrvglm](#) and [vgam](#).

When fitted, the `fitted.values` slot of the object contains the eight joint probabilities, labelled as $(Y_1, Y_2, Y_3) = (0,0,0), (0,0,1), (0,1,0), (0,1,1), (1,0,0), (1,0,1), (1,1,0), (1,1,1)$, respectively.

Note

The response must be a 3-column matrix of ones and zeros only. Note that each of the 8 combinations of the multivariate response need to appear in the data set, therefore data sets will need to be large in order for this family function to work. After estimation, the response attached to the object is also a 3-column matrix; possibly in the future it might change into a 8-column matrix.

Author(s)

Thomas W. Yee

References

Yee, T. W. and Wild, C. J. (2001). Discussion to: "Smoothing spline ANOVA for multivariate Bernoulli observations, with application to ophthalmology data (with discussion)" by Gao, F., Wahba, G., Klein, R., Klein, B. *Journal of the American Statistical Association*, **96**, 127–160.

McCullagh, P. and Nelder, J. A. (1989). *Generalized Linear Models*, 2nd ed. London: Chapman & Hall.

See Also

[loglinb2](#), [binom2.or](#), [hunua](#).

Examples

```
lfit <- vglm(cbind(cyadea, beitaw, kniexc) ~ altitude, loglinb3,
            data = hunua, trace = TRUE)
coef(lfit, matrix = TRUE)
head(fitted(lfit))
summary(lfit)
```

loglink

Log Link Function, and Variants

Description

Computes the log transformation, including its inverse and the first two derivatives.

Usage

```
loglink(theta, bvalue = NULL, inverse = FALSE, deriv = 0,
        short = TRUE, tag = FALSE)
negloglink(theta, bvalue = NULL, inverse = FALSE, deriv = 0,
           short = TRUE, tag = FALSE)
logneglink(theta, bvalue = NULL, inverse = FALSE, deriv = 0,
           short = TRUE, tag = FALSE)
```

Arguments

theta Numeric or character. See below for further details.
bvalue See [Links](#).
inverse, deriv, short, tag
 Details at [Links](#).

Details

The log link function is very commonly used for parameters that are positive. Here, all logarithms are natural logarithms, i.e., to base e . Numerical values of theta close to 0 or out of range result in Inf, -Inf, NA or NaN.

The function loglink computes $\log(\theta)$ whereas negloglink computes $-\log(\theta) = \log(1/\theta)$.

The function logneglink computes $\log(-\theta)$, hence is suitable for parameters that are negative, e.g., a trap-shy effect in [posbernoulli.b](#).

Value

The following concerns loglink. For deriv = 0, the log of theta, i.e., $\log(\text{theta})$ when inverse = FALSE, and if inverse = TRUE then $\exp(\text{theta})$. For deriv = 1, then the function returns $d \text{ eta} / d \text{ theta}$ as a function of theta if inverse = FALSE, else if inverse = TRUE then it returns the reciprocal.

Note

This function was called loge to avoid conflict with the [log](#) function. Numerical instability may occur when theta is close to 0 unless bvalue is used.

Author(s)

Thomas W. Yee

References

McCullagh, P. and Nelder, J. A. (1989). *Generalized Linear Models*, 2nd ed. London: Chapman & Hall.

See Also

[Links](#), [explink](#), [logitlink](#), [logclink](#), [logloglink](#), [log](#), [logofflink](#), [lambertW](#), [posbernoulli.b](#).

Examples

```
## Not run: loglink(seq(-0.2, 0.5, by = 0.1))
  loglink(seq(-0.2, 0.5, by = 0.1), bvalue = .Machine$double.xmin)
negloglink(seq(-0.2, 0.5, by = 0.1))
negloglink(seq(-0.2, 0.5, by = 0.1), bvalue = .Machine$double.xmin)
## End(Not run)
logneglink(seq(-0.5, -0.2, by = 0.1))
```

logloglink

Log-log and Log-log-log Link Functions

Description

Computes the two transformations, including their inverse and the first two derivatives.

Usage

```
logloglink(theta, bvalue = NULL, inverse = FALSE, deriv = 0,
           short = TRUE, tag = FALSE)
loglogloglink(theta, bvalue = NULL, inverse = FALSE, deriv = 0,
              short = TRUE, tag = FALSE)
```

Arguments

theta	Numeric or character. See below for further details.
bvalue	Values of theta which are less than or equal to 1 or e can be replaced by bvalue before computing the link function value. The component name bvalue stands for “boundary value”. See Links for more information.
inverse, deriv, short, tag	Details at Links .

Details

The log-log link function is commonly used for parameters that are greater than unity. Similarly, the log-log-log link function is applicable for parameters that are greater than e . Numerical values of theta close to 1 or e or out of range result in Inf, -Inf, NA or NaN. One possible application of loglogloglink() is to the k parameter (also called size) of [negbinomial](#) to Poisson-like data but with only a small amount of overdispersion; then k is a large number relative to munb. In such situations a [loglink](#) or [loglog](#) link may not be sufficient to draw the estimate toward the interior of the parameter space. Using a more stronger link function can help mitigate the Hauck-Donner effect [hdeff](#).

Value

For `logloglink()`: for `deriv = 0`, the log of $\log(\theta)$, i.e., $\log(\log(\theta))$ when `inverse = FALSE`, and if `inverse = TRUE` then $\exp(\exp(\theta))$.

For `loglogloglink()`: for `deriv = 0`, the log of $\log(\log(\theta))$, i.e., $\log(\log(\log(\theta)))$ when `inverse = FALSE`, and if `inverse = TRUE` then $\exp(\exp(\exp(\theta)))$.

For `deriv = 1`, then the function returns $d\theta/d\eta$ as a function of θ if `inverse = FALSE`, else if `inverse = TRUE` then it returns the reciprocal.

Here, all logarithms are natural logarithms, i.e., to base e .

Note

Numerical instability may occur when θ is close to 1 or e unless `bvalue` is used.

References

McCullagh, P. and Nelder, J. A. (1989). *Generalized Linear Models*, 2nd ed. London: Chapman & Hall.

See Also

[Links](#), [loglink](#), [logofflink](#).

Examples

```
x <- seq(0.8, 1.5, by = 0.1)
logloglink(x) # Has NAs
logloglink(x, bvalue = 1.0 + .Machine$double.eps) # Has no NAs

x <- seq(1.01, 10, len = 100)
logloglink(x)
max(abs(logloglink(logloglink(x), inverse = TRUE) - x)) # 0?
```

lognormal

Lognormal Distribution

Description

Maximum likelihood estimation of the (univariate) lognormal distribution.

Usage

```
lognormal(lmeanlog = "identitylink", lsdlog = "loglink", zero = "sdlog")
```

Arguments

<code>lmeanlog, lsdlog</code>	Parameter link functions applied to the mean and (positive) σ (standard deviation) parameter. Both of these are on the log scale. See Links for more choices.
<code>zero</code>	Specifies which linear/additive predictor is modelled as intercept-only. For <code>lognormal()</code> , the values can be from the set $\{1,2\}$ which correspond to μ , σ , respectively. See CommonVGAMffArguments for more information.

Details

A random variable Y has a 2-parameter lognormal distribution if $\log(Y)$ is distributed $N(\mu, \sigma^2)$. The expected value of Y , which is

$$E(Y) = \exp(\mu + 0.5\sigma^2)$$

and not μ , make up the fitted values. The variance of Y is

$$Var(Y) = [\exp(\sigma^2) - 1] \exp(2\mu + \sigma^2).$$

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as `vglm`, and `vgam`.

Author(s)

T. W. Yee

References

Kleiber, C. and Kotz, S. (2003). *Statistical Size Distributions in Economics and Actuarial Sciences*, Hoboken, NJ, USA: Wiley-Interscience.

See Also

[Lognormal](#), [uninormal](#), [CommonVGAMffArguments](#), [simulate.vlm](#).

Examples

```
ldata2 <- data.frame(x2 = runif(nn <- 1000))
ldata2 <- transform(ldata2, y1 = rlnorm(nn, 1 + 2 * x2, sd = exp(-1)),
                    y2 = rlnorm(nn, 1, sd = exp(-1 + x2)))
fit1 <- vglm(y1 ~ x2, lognormal(zero = 2), data = ldata2, trace = TRUE)
fit2 <- vglm(y2 ~ x2, lognormal(zero = 1), data = ldata2, trace = TRUE)
coef(fit1, matrix = TRUE)
coef(fit2, matrix = TRUE)
```

logofflink	<i>Log Link Function with an Offset</i>
------------	---

Description

Computes the log transformation with an offset, including its inverse and the first two derivatives.

Usage

```
logofflink(theta, offset = 0, inverse = FALSE, deriv = 0,
           short = TRUE, tag = FALSE)
log1plink(theta, offset = 0, inverse = FALSE, deriv = 0,
          short = TRUE, tag = FALSE)
```

Arguments

theta	Numeric or character. See below for further details.
offset	Offset value. See Links . For log1plink this argument should not be used because the offset is implicitly unity .
inverse, deriv, short, tag	Details at Links .

Details

The log-offset link function is very commonly used for parameters that are greater than a certain value. In particular, it is defined by $\log(\text{theta} + \text{offset})$ where `offset` is the offset value. For example, if `offset = 0.5` then the value of `theta` is restricted to be greater than -0.5 .

Numerical values of `theta` close to $-\text{offset}$ or out of range result in `Inf`, `-Inf`, `NA` or `NaN`.

The offset is implicitly 1 in [log1plink](#). It is equivalent to `logofflink(offset = 1)` but is more accurate if $\text{abs}(\text{theta})$ is tiny. It may be used for `lrho` in [extbetabinomial](#) provided an offset $\log(\text{size} - 1)$ for η_2 is included.

Value

For `deriv = 0`, the log of `theta+offset`, i.e., $\log(\text{theta} + \text{offset})$ when `inverse = FALSE`, and if `inverse = TRUE` then $\exp(\text{theta}) - \text{offset}$.

For `deriv = 1`, then the function returns $d \text{ theta} / d \text{ eta}$ as a function of `theta` if `inverse = FALSE`, else if `inverse = TRUE` then it returns the reciprocal.

Here, all logarithms are natural logarithms, i.e., to base e .

Note

The default means this function is identical to [loglink](#).

Numerical instability may occur when `theta` is close to $-\text{offset}$.

Author(s)

Thomas W. Yee

References

McCullagh, P. and Nelder, J. A. (1989). *Generalized Linear Models*, 2nd ed. London: Chapman & Hall.

See Also

[Links](#), [loglink](#), [extbetabinomial](#).

Examples

```
## Not run:
logofflink(seq(-0.2, 0.5, by = 0.1))
logofflink(seq(-0.2, 0.5, by = 0.1), offset = 0.5)
  log(seq(-0.2, 0.5, by = 0.1) + 0.5)
## End(Not run)
```

Lomax

The Lomax Distribution

Description

Density, distribution function, quantile function and random generation for the Lomax distribution with scale parameter *scale* and shape parameter *q*.

Usage

```
dlomax(x, scale = 1, shape3.q, log = FALSE)
plomax(q, scale = 1, shape3.q, lower.tail = TRUE, log.p = FALSE)
qlomax(p, scale = 1, shape3.q, lower.tail = TRUE, log.p = FALSE)
rlomax(n, scale = 1, shape3.q)
```

Arguments

<i>x</i> , <i>q</i>	vector of quantiles.
<i>p</i>	vector of probabilities.
<i>n</i>	number of observations. If <code>length(n) > 1</code> , the length is taken to be the number required.
<i>scale</i>	scale parameter.
<i>shape3.q</i>	shape parameter.
<i>log</i>	Logical. If <code>log = TRUE</code> then the logarithm of the density is returned.
<i>lower.tail</i> , <i>log.p</i>	Same meaning as in pnorm or qnorm .

Details

See [lomax](#), which is the **VGAM** family function for estimating the parameters by maximum likelihood estimation.

Value

`dlomax` gives the density, `plomax` gives the distribution function, `qlomax` gives the quantile function, and `rlomax` generates random deviates.

Note

The Lomax distribution is a special case of the 4-parameter generalized beta II distribution.

Author(s)

T. W. Yee and Kai Huang

References

Kleiber, C. and Kotz, S. (2003). *Statistical Size Distributions in Economics and Actuarial Sciences*, Hoboken, NJ, USA: Wiley-Interscience.

See Also

[lomax](#), [genbetaII](#).

Examples

```
probs <- seq(0.1, 0.9, by = 0.1)
max(abs(plomax(qlomax(p = probs, shape3.q = 1),
              shape3.q = 1) - probs)) # Should be 0

## Not run: par(mfrow = c(1, 2))
x <- seq(-0.01, 5, len = 401)
plot(x, dexp(x), type = "l", col = "black", ylab = "", ylim = c(0, 3),
      main = "Black is std exponential, others are dlomax(x, shape3.q)")
lines(x, dlomax(x, shape3.q = 1), col = "orange")
lines(x, dlomax(x, shape3.q = 2), col = "blue")
lines(x, dlomax(x, shape3.q = 5), col = "green")
legend("topright", col = c("orange", "blue", "green"), lty = rep(1, 3),
      legend = paste("shape3.q =", c(1, 2, 5)))

plot(x, pexp(x), type = "l", col = "black", ylab = "", las = 1,
      main = "Black is std exponential, others are plomax(x, shape3.q)")
lines(x, plomax(x, shape3.q = 1), col = "orange")
lines(x, plomax(x, shape3.q = 2), col = "blue")
lines(x, plomax(x, shape3.q = 5), col = "green")
legend("bottomright", col = c("orange", "blue", "green"), lty = rep(1, 3),
      legend = paste("shape3.q =", c(1, 2, 5)))

## End(Not run)
```

lomax

*Lomax Distribution Family Function***Description**

Maximum likelihood estimation of the 2-parameter Lomax distribution.

Usage

```
lomax(lscale = "loglink", lshape3.q = "loglink", iscale = NULL,
      ishape3.q = NULL, imethod = 1, gscale = exp(-5:5),
      gshape3.q = seq(0.75, 4, by = 0.25),
      probs.y = c(0.25, 0.5, 0.75), zero = "shape")
```

Arguments

lscale, lshape3.q

Parameter link function applied to the (positive) parameters scale and q. See [Links](#) for more choices.

iscale, ishape3.q, imethod

See [CommonVGAMffArguments](#) for information. For imethod = 2 a good initial value for iscale is needed to obtain a good estimate for the other parameter.

gscale, gshape3.q, zero, probs.y

See [CommonVGAMffArguments](#).

Details

The 2-parameter Lomax distribution is the 4-parameter generalized beta II distribution with shape parameters $a = p = 1$. It is probably more widely known as the Pareto (II) distribution. It is also the 3-parameter Singh-Maddala distribution with shape parameter $a = 1$, as well as the beta distribution of the second kind with $p = 1$. More details can be found in Kleiber and Kotz (2003).

The Lomax distribution has density

$$f(y) = q/[b\{1 + y/b\}^{1+q}]$$

for $b > 0$, $q > 0$, $y \geq 0$. Here, b is the scale parameter scale, and q is a shape parameter. The cumulative distribution function is

$$F(y) = 1 - [1 + (y/b)]^{-q}.$$

The mean is

$$E(Y) = b/(q - 1)$$

provided $q > 1$; these are returned as the fitted values. This family function handles multiple responses.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Note

See the notes in [genbetaII](#).

Author(s)

T. W. Yee

References

Kleiber, C. and Kotz, S. (2003). *Statistical Size Distributions in Economics and Actuarial Sciences*, Hoboken, NJ, USA: Wiley-Interscience.

See Also

[Lomax](#), [genbetaII](#), [betaII](#), [dagum](#), [sinmad](#), [fisk](#), [inv.lomax](#), [paralogistic](#), [inv.paralogistic](#), [simulate.vlm](#).

Examples

```
ldata <- data.frame(y = rlomax(n = 1000, scale = exp(1), exp(2)))
fit <- vglm(y ~ 1, lomax, data = ldata, trace = TRUE)
coef(fit, matrix = TRUE)
Coef(fit)
summary(fit)
```

lpossums

Leadbeater's Possums

Description

Abundance of Leadbeater's Possums observed in the field.

Usage

```
data(lpossums)
```

Format

A data frame with the following variables.

number Values between 0 and 10 excluding 6.

ofreq Observed frequency, i.e., the number of sites.

Details

A small data set recording the abundance of Leadbeater's Possums *Gymnobelideus leadbeateri* observed in the montane ash forests of the Central Highlands of Victoria, in south-eastern Australia. There are 151 3-hectare sites. The data has more 0s than usual relative to the Poisson, as well as exhibiting overdispersion too.

Source

Welsh, A. H., Cunningham, R. B., Donnelly, C. F. and Lindenmayer, D. B. (1996). Modelling the abundances of rare species: statistical models for counts with extra zeros. *Ecological Modelling*, **88**, 297–308.

See Also

[zipoissonff](#).

Examples

```
lpossums
(samplemean <- with(lpossums, weighted.mean(number, ofreq)))
with(lpossums, var(rep(number, times = ofreq)) / samplemean)
sum(with(lpossums, ofreq))
## Not run: spikeplot(with(lpossums, rep(number, times = ofreq)),
  main = "Leadbeater's possums", col = "blue", xlab = "Number")
## End(Not run)
```

lqnorm

Minimizing the L-q norm Family Function

Description

Minimizes the L-q norm of residuals in a linear model.

Usage

```
lqnorm(qpower = 2, link = "identitylink",
  imethod = 1, imu = NULL, ishrinkage = 0.95)
```

Arguments

qpower	A single numeric, must be greater than one, called q below. The absolute value of residuals are raised to the power of this argument, and then summed. This quantity is minimized with respect to the regression coefficients.
link	Link function applied to the ‘mean’ μ . See Links for more details.
imethod	Must be 1, 2 or 3. See CommonVGAMffArguments for more information. Ignored if imu is specified.
imu	Numeric, optional initial values used for the fitted values. The default is to use imethod = 1.
ishrinkage	How much shrinkage is used when initializing the fitted values. The value must be between 0 and 1 inclusive, and a value of 0 means the individual response values are used, and a value of 1 means the median or mean is used. This argument is used in conjunction with imethod = 3.

Details

This function minimizes the objective function

$$\sum_{i=1}^n w_i (|y_i - \mu_i|)^q$$

where q is the argument `power`, $\eta_i = g(\mu_i)$ where g is the link function, and η_i is the vector of linear/additive predictors. The prior weights w_i can be inputted using the `weights` argument of `vlm/vglm/vgam` etc.; it should be just a vector here since this function handles only a single vector or one-column response.

Numerical problem will occur when q is too close to one. Probably reasonable values range from 1.5 and up, say. The value $q = 2$ corresponds to ordinary least squares while $q = 1$ corresponds to the MLE of a double exponential (Laplace) distribution. The procedure becomes more sensitive to outliers the larger the value of q .

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as `vglm`, and `vgam`.

Warning

Convergence failure is common, therefore the user is advised to be cautious and monitor convergence!

Note

This **VGAM** family function is an initial attempt to provide a more robust alternative for regression and/or offer a little more flexibility than least squares. The `@misc` slot of the fitted object contains a list component called `objectiveFunction` which is the value of the objective function at the final iteration.

Author(s)

Thomas W. Yee

References

Yee, T. W. and Wild, C. J. (1996). Vector generalized additive models. *Journal of the Royal Statistical Society, Series B, Methodological*, **58**, 481–493.

See Also

[uninormal](#).

Examples

```

set.seed(123)
ldata <- data.frame(x = sort(runif(nn <- 10)))
realfun <- function(x) 4 + 5*x
ldata <- transform(ldata, y = realfun(x) + rnorm(nn, sd = exp(-1)))
# Make the first observation an outlier
ldata <- transform(ldata, y = c(4*y[1], y[-1]), x = c(-1, x[-1]))
fit <- vglm(y ~ x, lqnorm(qpower = 1.2), data = ldata)
coef(fit, matrix = TRUE)
head(fitted(fit))
fit@misc$qpower
fit@misc$objectiveFunction

## Not run:
# Graphical check
with(ldata, plot(x, y,
  main = paste0("LS = red, lqnorm = blue (qpower = ",
    fit@misc$qpower, "), truth = black"), col = "blue"))
lmfit <- lm(y ~ x, data = ldata)
with(ldata, lines(x, fitted(fit), col = "blue"))
with(ldata, lines(x, lmfit$fitted, col = "red"))
with(ldata, lines(x, realfun(x), col = "black"))
## End(Not run)

```

lrt.stat

Likelihood Ratio Test Statistics Evaluated at the Null Values

Description

Generic function that computes likelihood ratio test (LRT) statistics evaluated at the null values (consequently they do not suffer from the Hauck-Donner effect).

Usage

```

lrt.stat(object, ...)
lrt.stat.vlm(object, values0 = 0, subset = NULL, omit1s = TRUE,
  all.out = FALSE, trace = FALSE, ...)

```

Arguments

object, values0, subset
 Same as in [wald.stat.vlm](#).

omit1s, all.out, trace
 Same as in [wald.stat.vlm](#).

...
 Ignored for now.

Details

When `summary()` is applied to a `vglm` object a 4-column Wald table is produced. The corresponding p-values are generally viewed as inferior to those from a likelihood ratio test (LRT). For example, the Hauck and Donner (1977) effect (HDE) produces p-values that are biased upwards (see [hdeff](#)). Other reasons are that the Wald test is often less accurate (especially in small samples) and is not invariant to parameterization. By default, this function returns p-values based on the LRT by deleting one column at a time from the big VLM matrix and then restarting IRLS to obtain convergence (hopefully). Twice the difference between the log-likelihoods (or equivalently, the difference in the deviances if they are defined) is asymptotically chi-squared with 1 degree of freedom. One might expect the p-values from this function therefore to be more accurate and not suffer from the HDE. Thus this function is a recommended alternative (if it works) to `summaryvglm` for testing for the significance of a regression coefficient.

Value

By default, a vector of signed square root of the LRT statistics; these are asymptotically standard normal under the null hypotheses. If `all.out = TRUE` then a list is returned with the following components: `lrt.stat` the signed LRT statistics, `pvalues` the 2-sided p-values, `lrt.stat2` the usual LRT statistic, `values0` the null values.

Warning

See [wald.stat.vlm](#).

Author(s)

T. W. Yee.

See Also

[score.stat](#), [wald.stat](#), [summaryvglm](#), [anova.vglm](#), [vglm](#), [lrtest](#), [confintvglm](#), [pchisq](#), [profilevglm](#), [hdeff](#).

Examples

```
set.seed(1)
pneumo <- transform(pneumo, let = log(exposure.time),
                   x3 = rnorm(nrow(pneumo)))
fit <- vglm(cbind(normal, mild, severe) ~ let, propodds, pneumo)
cbind(coef(summary(fit)),
      "signed LRT stat" = lrt.stat(fit, omit1s = FALSE))
summary(fit, lrt0 = TRUE) # Easy way to get it
```

 lrtest *Likelihood Ratio Test of Nested Models*

Description

lrtest is a generic function for carrying out likelihood ratio tests. The default method can be employed for comparing nested VGLMs (see details below).

Usage

```
lrtest(object, ...)
```

```
lrtest_vglm(object, ..., no.warning = FALSE, name = NULL)
```

Arguments

object	a <code>vglm</code> object. See below for details.
...	further object specifications passed to methods. See below for details.
no.warning	logical; if TRUE then no warning is issued. For example, setting TRUE might be a good idea when testing for linearity of a variable for a "pvgam" object.
name	a function for extracting a suitable name/description from a fitted model object. By default the name is queried by calling <code>formula</code> .

Details

lrtest is intended to be a generic function for comparisons of models via asymptotic likelihood ratio tests. The default method consecutively compares the fitted model object `object` with the models passed in `...`. Instead of passing the fitted model objects in `...`, several other specifications are possible. The updating mechanism is the same as for `waldtest()` in **lmtree**: the models in `...` can be specified as integers, characters (both for terms that should be eliminated from the previous model), update formulas or fitted model objects. Except for the last case, the existence of an `update` method is assumed. See `waldtest()` in **lmtree** for details.

Subsequently, an asymptotic likelihood ratio test for each two consecutive models is carried out: Twice the difference in log-likelihoods (as derived by the `logLik` methods) is compared with a Chi-squared distribution.

Value

An object of class "VGAManova" which contains a slot with the log-likelihood, degrees of freedom, the difference in degrees of freedom, likelihood ratio Chi-squared statistic and corresponding p value. These are printed by `stats:::print.anova()`; see `anova`.

Warning

Several **VGAM** family functions implement distributions which do not satisfying the usual regularity conditions needed for the LRT to work. No checking or warning is given for these.

Note

The code was adapted directly from **lmtest** (written by T. Hothorn, A. Zeileis, G. Millo, D. Mitchell) and made to work for VGLMs and S4. This help file also was adapted from **lmtest**.

Approximate LRTs might be applied to VGAMs, as produced by **vgam**, but it is probably better in inference to use **vglm** with regression splines (**bs** and **ns**). This methods function should not be applied to other models such as those produced by **rrvglm**, by **cqo**, by **cao**.

See Also

lmtest, **vglm**, **lrt.stat.vlm**, **score.stat.vlm**, **wald.stat.vlm**, **anova.vglm**.

Examples

```
set.seed(1)
pneumo <- transform(pneumo, let = log(exposure.time),
                    x3 = runif(nrow(pneumo)))
fit1 <- vglm(cbind(normal, mild, severe) ~ let, propodds, pneumo)
fit2 <- vglm(cbind(normal, mild, severe) ~ let + x3, propodds, pneumo)
fit3 <- vglm(cbind(normal, mild, severe) ~ let, cumulative, pneumo)
# Various equivalent specifications of the LR test for testing x3
(ans1 <- lrtest(fit2, fit1))
ans2 <- lrtest(fit2, 2)
ans3 <- lrtest(fit2, "x3")
ans4 <- lrtest(fit2, . ~ . - x3)
c(all.equal(ans1, ans2), all.equal(ans1, ans3), all.equal(ans1, ans4))

# Doing it manually
(testStatistic <- 2 * (logLik(fit2) - logLik(fit1)))
(pval <- pchisq(testStatistic, df = df.residual(fit1) - df.residual(fit2),
               lower.tail = FALSE))

(ans4 <- lrtest(fit3, fit1)) # Test P0 (parallelism) assumption
```

lvplot

*Latent Variable Plot***Description**

Generic function for a *latent variable plot* (also known as an *ordination diagram* by ecologists).

Usage

```
lvplot(object, ...)
```

Arguments

object An object for a latent variable plot is meaningful.

... Other arguments fed into the specific methods function of the model. They usually are graphical parameters, and sometimes they are fed into the methods function for **Coef**.

Details

Latent variables occur in reduced-rank regression models, as well as in quadratic and additive ordination. For the latter, latent variables are often called the *site scores*. Latent variable plots were coined by Yee (2004), and have the latent variable as at least one of its axes.

Value

The value returned depends specifically on the methods function invoked.

Note

Latent variables are not really applicable to [vglm/vgam](#) models.

Author(s)

Thomas W. Yee

References

- Yee, T. W. (2004). A new technique for maximum-likelihood canonical Gaussian ordination. *Ecological Monographs*, **74**, 685–701.
- Yee, T. W. (2006). Constrained additive ordination. *Ecology*, **87**, 203–213.

See Also

[lvplot.qrrvglm](#), [lvplot.cao](#), [latvar](#), [trplot](#).

Examples

```
## Not run:
hspider[,1:6] <- scale(hspider[,1:6]) # Stdz environmental vars
set.seed(123)
p1 <- cao(cbind(Pardlugu, Pardmont, Pardnigr, Pardpull, Zoraspin) ~
  WaterCon + BareSand + FallTwig +
  CoveMoss + CoveHerb + ReflLux,
  family = poissonff, data = hspider, Bestof = 3,
  df1.nl = c(Zoraspin = 2.5, 3), Crowlpositive = TRUE)
index <- 1:ncol(depvar(p1))
lvplot(p1, lcol = index, pcol = index, y = TRUE, las = 1)

## End(Not run)
```

 lvplot.qrrvglm *Latent Variable Plot for QO models*

Description

Produces an ordination diagram (latent variable plot) for quadratic ordination (QO) models. For rank-1 models, the x-axis is the first ordination/constrained/canonical axis. For rank-2 models, the x- and y-axis are the first and second ordination axes respectively.

Usage

```
lvplot.qrrvglm(object, varI.latvar = FALSE, refResponse = NULL,
  add = FALSE, show.plot = TRUE,
  rug = TRUE, y = FALSE, type = c("fitted.values", "predictors"),
  xlab = paste0("Latent Variable", if (Rank == 1) "" else " 1"),
  ylab = if (Rank == 1) switch(type, predictors = "Predictors",
    fitted.values = "Fitted values") else "Latent Variable 2",
  pcex = par()$cex, pcol = par()$col, pch = par()$pch,
  lty = par()$lty, lcol = par()$col, llwd = par()$lwd,
  label.arg = FALSE, adj.arg = -0.1,
  ellipse = 0.95, Absolute = FALSE, elty = par()$lty,
  ecol = par()$col, elwd = par()$lwd, egrid = 200,
  chull.arg = FALSE, cpty = 2, ccol = par()$col, clwd = par()$lwd,
  cpch = " ",
  C = FALSE, OriginC = c("origin", "mean"),
  Clty = par()$lty, Ccol = par()$col, Clwd = par()$lwd,
  Ccex = par()$cex, Cadj.arg = -0.1, stretchC = 1,
  sites = FALSE, spch = NULL, scol = par()$col, scex = par()$cex,
  sfont = par()$font, check.ok = TRUE, jitter.y = FALSE, ...)
```

Arguments

object	A CQO object.
varI.latvar	Logical that is fed into Coef.qrrvglm .
refResponse	Integer or character that is fed into Coef.qrrvglm .
add	Logical. Add to an existing plot? If FALSE, a new plot is made.
show.plot	Logical. Plot it?
rug	Logical. If TRUE, a rug plot is plotted at the foot of the plot (applies to rank-1 models only). These values are jittered to expose ties.
y	Logical. If TRUE, the responses will be plotted (applies only to rank-1 models and if type = "fitted.values").
type	Either "fitted.values" or "predictors", specifies whether the y-axis is on the response or eta-scales respectively.
xlab	Caption for the x-axis. See par .

y1ab	Caption for the y-axis. See par .
pcex	Character expansion of the points. Here, for rank-1 models, points are the response y data. For rank-2 models, points are the optimums. See the cex argument in par .
pcol	Color of the points. See the col argument in par .
pch	Either an integer specifying a symbol or a single character to be used as the default in plotting points. See par . The pch argument can be of length M , the number of species.
l1ty	Line type. Rank-1 models only. See the lty argument of par .
lcol	Line color. Rank-1 models only. See the col argument of par .
llwd	Line width. Rank-1 models only. See the lwd argument of par .
label.arg	Logical. Label the optimums and C? (applies only to rank-2 models only).
adj.arg	Justification of text strings for labelling the optimums (applies only to rank-2 models only). See the adj argument of par .
ellipse	Numerical, of length 0 or 1 (applies only to rank-2 models only). If Absolute is TRUE then ellipse should be assigned a value that is used for the elliptical contouring. If Absolute is FALSE then ellipse should be assigned a value between 0 and 1, for example, setting ellipse = 0.9 means an ellipse with contour = 90% of the maximum will be plotted about each optimum. If ellipse is a negative value, then the function checks that the model is an equal-tolerances model and varI.latvar = FALSE, and if so, plots circles with radius -ellipse. For example, setting ellipse = -1 will result in circular contours that have unit radius (in latent variable units). If ellipse is NULL or FALSE then no ellipse is drawn around the optimums.
Absolute	Logical. If TRUE, the contours corresponding to ellipse are on an absolute scale. If FALSE, the contours corresponding to ellipse are on a relative scale.
elty	Line type of the ellipses. See the lty argument of par .
ecol	Line color of the ellipses. See the col argument of par .
elwd	Line width of the ellipses. See the lwd argument of par .
egrid	Numerical. Line resolution of the ellipses. Choosing a larger value will result in smoother ellipses. Useful when ellipses are large.
chull.arg	Logical. Add a convex hull around the site scores?
clty	Line type of the convex hull. See the lty argument of par .
ccol	Line color of the convex hull. See the col argument of par .
clwd	Line width of the convex hull. See the lwd argument of par .
cpch	Character to be plotted at the intersection points of the convex hull. Having white spaces means that site labels are not obscured there. See the pch argument of par .
C	Logical. Add C (represented by arrows emanating from OriginC) to the plot?
OriginC	Character or numeric. Where the arrows representing C emanate from. If character, it must be one of the choices given. By default the first is chosen. The value "origin" means $c(0, 0)$. The value "mean" means the sample mean of the latent variables (centroid). Alternatively, the user may specify a numerical vector of length 2.

Clty	Line type of the arrows representing C . See the lty argument of par .
Ccol	Line color of the arrows representing C . See the col argument of par .
Clwd	Line width of the arrows representing C . See the lwd argument of par .
Ccex	Numeric. Character expansion of the labelling of C . See the cex argument of par .
Cadj.arg	Justification of text strings when labelling C . See the adj argument of par .
stretchC	Numerical. Stretching factor for C . Instead of using C , stretchC * C is used.
sites	Logical. Add the site scores (aka latent variable values, nu's) to the plot? (applies only to rank-2 models only).
spch	Plotting character of the site scores. The default value of NULL means the row labels of the data frame are used. They often are the site numbers. See the pch argument of par .
scol	Color of the site scores. See the col argument of par .
scex	Character expansion of the site scores. See the cex argument of par .
sfont	Font used for the site scores. See the font argument of par .
check.ok	Logical. Whether a check is performed to see that noRRR = ~ 1 was used. It doesn't make sense to have a latent variable plot unless this is so.
jitter.y	Logical. If y is plotted, jitter it first? This may be useful for counts and proportions.
...	Arguments passed into the plot function when setting up the entire plot. Useful arguments here include xlim and ylim.

Details

This function only works for rank-1 and rank-2 QRR-VGLMs with argument noRRR = ~ 1.

For unequal-tolerances models, the latent variable axes can be rotated so that at least one of the tolerance matrices is diagonal; see [Coef.qrrvglm](#) for details.

Arguments beginning with “p” correspond to the points e.g., pcex and pcol correspond to the size and color of the points. Such “p” arguments should be vectors of length 1, or *n*, the number of sites. For the rank-2 model, arguments beginning with “p” correspond to the optimums.

Value

Returns a matrix of latent variables (site scores) regardless of whether a plot was produced or not.

Warning

Interpretation of a latent variable plot (CQO diagram) is potentially very misleading in terms of distances if (i) the tolerance matrices of the species are unequal and (ii) the contours of these tolerance matrices are not included in the ordination diagram.

Note

A species which does not have an optimum will not have an ellipse drawn even if requested, i.e., if its tolerance matrix is not positive-definite.

Plotting **C** gives a visual display of the weights (loadings) of each of the variables used in the linear combination defining each latent variable.

The arguments `elty`, `ecol` and `elwd`, may be replaced in the future by `llty`, `lcol` and `llwd`, respectively.

For rank-1 models, a similar function to this one is [perspqrvglm](#). It plots the fitted values on a more fine grid rather than at the actual site scores here. The result is a collection of smooth bell-shaped curves. However, it has the weakness that the plot is more divorced from the data; the user thinks it is the truth without an appreciation of the statistical variability in the estimates.

In the example below, the data comes from an equal-tolerances model. The species' tolerance matrices are all the identity matrix, and the optimums are at (0,0), (1,1) and (-2,0) for species 1, 2, 3 respectively.

Author(s)

Thomas W. Yee

References

Yee, T. W. (2004). A new technique for maximum-likelihood canonical Gaussian ordination. *Ecological Monographs*, **74**, 685–701.

See Also

[lvplot](#), [perspqrvglm](#), [Coef.qrrvglm](#), [par](#), [cqo](#).

Examples

```
set.seed(123); nn <- 200
cdata <- data.frame(x2 = rnorm(nn), # Mean 0 (needed when I.tol=TRUE)
                  x3 = rnorm(nn), # Mean 0 (needed when I.tol=TRUE)
                  x4 = rnorm(nn)) # Mean 0 (needed when I.tol=TRUE)
cdata <- transform(cdata, latvar1 = x2 + x3 - 2*x4,
                  latvar2 = -x2 + x3 + 0*x4)
# Nb. latvar2 is weakly correlated with latvar1
cdata <- transform(cdata,
                  lambda1 = exp(6 - 0.5 * (latvar1-0)^2 - 0.5 * (latvar2-0)^2),
                  lambda2 = exp(5 - 0.5 * (latvar1-1)^2 - 0.5 * (latvar2-1)^2),
                  lambda3 = exp(5 - 0.5 * (latvar1+2)^2 - 0.5 * (latvar2-0)^2))
cdata <- transform(cdata,
                  spp1 = rpois(nn, lambda1),
                  spp2 = rpois(nn, lambda2),
                  spp3 = rpois(nn, lambda3))
set.seed(111)
## Not run:
p2 <- cqo(cbind(spp1, spp2, spp3) ~ x2 + x3 + x4, poissonff,
          data = cdata, Rank = 2, I.tolerances = TRUE,
```

```

      Crowlpositive = c(TRUE, FALSE)) # deviance = 505.81
if (deviance(p2) > 506) stop("suboptimal fit obtained")
sort(deviance(p2, history = TRUE)) # A history of the iterations
Coef(p2)

## End(Not run)

## Not run:
lvplot(p2, sites = TRUE, spch = "*", scol = "darkgreen", scex = 1.5,
  chull = TRUE, label = TRUE, Absolute = TRUE, ellipse = 140,
  adj = -0.5, pcol = "blue", pcex = 1.3, las = 1, Ccol = "orange",
  C = TRUE, Cadj = c(-0.3, -0.3, 1), Clwd = 2, Ccex = 1.4,
  main = paste("Contours at Abundance = 140 with",
    "convex hull of the site scores"))
## End(Not run)
## Not run:
var(latvar(p2)) # A diagonal matrix, i.e., uncorrelated latent vars
var(latvar(p2, varI.latvar = TRUE)) # Identity matrix
Tol(p2)[, , 1:2] # Identity matrix
Tol(p2, varI.latvar = TRUE)[, , 1:2] # A diagonal matrix

## End(Not run)

```

lvplot.rrvglm

Latent Variable Plot for RR-VGLMs

Description

Produces an *ordination diagram* (also known as a *biplot* or *latent variable plot*) for *reduced-rank vector generalized linear models* (RR-VGLMs). For rank-2 models only, the x- and y-axis are the first and second canonical axes respectively.

Usage

```

lvplot.rrvglm(object,
  A = TRUE, C = TRUE, scores = FALSE, show.plot = TRUE,
  groups = rep(1, n), gapC = sqrt(sum(par()$cxy^2)), scaleA = 1,
  xlab = "Latent Variable 1", ylab = "Latent Variable 2",
  Alabels = if (length(object@misc$predictors.names))
  object@misc$predictors.names else param.names("LP", M),
  Aadj = par()$adj, Acex = par()$cex, Acol = par()$col,
  Apch = NULL,
  Clabels = rownames(Cmat), Cadj = par()$adj,
  Ccex = par()$cex, Ccol = par()$col, Clty = par()$lty,
  Clwd = par()$lwd,
  chull.arg = FALSE, ccex = par()$cex, ccol = par()$col,
  clty = par()$lty, clwd = par()$lwd,
  spch = NULL, scex = par()$cex, scol = par()$col,
  slabels = rownames(x2mat), ...)

```

Arguments

object	Object of class "rrvglm".
A	Logical. Allow the plotting of A ?
C	Logical. Allow the plotting of C ? If TRUE then C is represented by arrows emanating from the origin.
scores	Logical. Allow the plotting of the n scores? The scores are the values of the latent variables for each observation.
show.plot	Logical. Plot it? If FALSE, no plot is produced and the matrix of scores (n latent variable values) is returned. If TRUE, the rank of object need not be 2.
groups	A vector whose distinct values indicate which group the observation belongs to. By default, all the observations belong to a single group. Useful for the multinomial logit model (see multinomial).
gapC	The gap between the end of the arrow and the text labelling of C , in latent variable units.
scaleA	Numerical value that is multiplied by A , so that C is divided by this value.
xlab	Caption for the x-axis. See par .
ylab	Caption for the y-axis. See par .
Alabels	Character vector to label A . Must be of length M .
Aadj	Justification of text strings for labelling A . See the adj argument of par .
Acex	Numeric. Character expansion of the labelling of A . See the cex argument of par .
Acol	Line color of the arrows representing C . See the col argument of par .
Apch	Either an integer specifying a symbol or a single character to be used as the default in plotting points. See par . The pch argument can be of length M , the number of species.
Clabels	Character vector to label C . Must be of length p^2 .
Cadj	Justification of text strings for labelling C . See the adj argument of par .
Ccex	Numeric. Character expansion of the labelling of C . See the cex argument of par .
Ccol	Line color of the arrows representing C . See the col argument of par .
Clty	Line type of the arrows representing C . See the lty argument of par .
Clwd	Line width of the arrows representing C . See the lwd argument of par .
chull.arg	Logical. Plot the convex hull of the scores? This is done for each group (see the group argument).
ccex	Numeric. Character expansion of the labelling of the convex hull. See the cex argument of par .
ccol	Line color of the convex hull. See the col argument of par .
clty	Line type of the convex hull. See the lty argument of par .
clwd	Line width of the convex hull. See the lwd argument of par .

spch	Either an integer specifying a symbol or a single character to be used as the default in plotting points. See par . The spch argument can be of length M , number of species.
scex	Numeric. Character expansion of the labelling of the scores. See the cex argument of par .
scol	Line color of the arrows representing \mathbf{C} . See the col argument of par .
slabels	Character vector to label the scores. Must be of length n .
...	Arguments passed into the plot function when setting up the entire plot. Useful arguments here include <code>xlim</code> and <code>ylim</code> .

Details

For RR-VGLMs, a *biplot* and a *latent variable* plot coincide. In general, many of the arguments starting with “A” refer to \mathbf{A} (of length M), “C” to \mathbf{C} (of length p^2), “c” to the convex hull (of length `length(unique(groups))`), and “s” to scores (of length n).

As the result is a biplot, its interpretation is based on the inner product.

Value

The matrix of scores (n latent variable values) is returned regardless of whether a plot was produced or not.

Note

The functions `lvplot.rrvglm` and `biplot.rrvglm` are equivalent.

In the example below the predictor variables are centered, which is a good idea.

Author(s)

Thomas W. Yee

References

Yee, T. W. and Hastie, T. J. (2003). Reduced-rank vector generalized linear models. *Statistical Modelling*, **3**, 15–41.

See Also

[lvplot](#), [par](#), [rrvglm](#), [Coef.rrvglm](#), [rrvglm.control](#).

Examples

```
set.seed(1)
nn <- nrow(pneumo) # x1--x3 are some unrelated covariates
pneumo <-
  transform(pneumo, slet = scale(log(exposure.time)),
            imag = severe + 3, # Fictional!
            x1 = rnorm(nn), x2 = rnorm(nn), x3 = rnorm(nn))
fit <-
```

```

rrvglm(cbind(normal, mild, severe, imag) ~ slet + x1 + x2 + x3,
#       Corner = FALSE, Uncorrel = TRUE, # orig.
       multinomial, data = pneumo, Rank = 2)
## Not run:
lvplot(fit, chull = TRUE, scores = TRUE, clty = 2, ccol = 4,
       scol = "red", Ccol = "green3", Clwd = 2, Ccex = 2,
       main = "Biplot of some fictitious data")
## End(Not run)

```

machinists

Machinists Accidents

Description

A small count data set involving 414 machinists from a three months study, of accidents around the end of WWI.

Usage

```
data(machinists)
```

Format

A data frame with the following variables.

accidents The number of accidents

ofreq Observed frequency, i.e., the number of machinists with that many accidents

Details

The data was collected over a period of three months. There were 414 machinists in total. Also, there were data collected over six months, but it is not given here.

Source

Incidence of Industrial Accidents. Report No. 4 (Industrial Fatigue Research Board), Stationery Office, London, 1919.

References

Greenwood, M. and Yule, G. U. (1920). An Inquiry into the Nature of Frequency Distributions Representative of Multiple Happenings with Particular Reference to the Occurrence of Multiple Attacks of Disease or of Repeated Accidents. *Journal of the Royal Statistical Society*, **83**, 255–279.

See Also

[negbinomial](#), [poissonff](#).

Examples

```

machinists
mean(with(machinists, rep(accidents, times = ofreq)))
var(with(machinists, rep(accidents, times = ofreq)))
## Not run: barplot(with(machinists, ofreq),
  names.arg = as.character(with(machinists, accidents)),
  main = "Machinists accidents",
  col = "lightblue", las = 1,
  ylab = "Frequency", xlab = "accidents")
## End(Not run)

```

 Makeham

The Makeham Distribution

Description

Density, cumulative distribution function, quantile function and random generation for the Makeham distribution.

Usage

```

dmakeham(x, scale = 1, shape, epsilon = 0, log = FALSE)
pmakeham(q, scale = 1, shape, epsilon = 0, lower.tail = TRUE,
  log.p = FALSE)
qmakeham(p, scale = 1, shape, epsilon = 0, lower.tail = TRUE,
  log.p = FALSE)
rmakeham(n, scale = 1, shape, epsilon = 0)

```

Arguments

x, q	vector of quantiles.
p	vector of probabilities.
n	number of observations. Same as in runif .
log	Logical. If log = TRUE then the logarithm of the density is returned.
lower.tail, log.p	Same meaning as in pnorm or qnorm .
scale, shape	positive scale and shape parameters.
epsilon	another parameter. Must be non-negative. See below.

Details

See [makeham](#) for details. The default value of `epsilon = 0` corresponds to the Gompertz distribution. The function [pmakeham](#) uses [lambertW](#).

Value

dmakeham gives the density, pmakeham gives the cumulative distribution function, qmakeham gives the quantile function, and rmakeham generates random deviates.

Author(s)

T. W. Yee and Kai Huang

References

Jodra, P. (2009). A closed-form expression for the quantile function of the Gompertz-Makeham distribution. *Mathematics and Computers in Simulation*, **79**, 3069–3075.

See Also

[makeham](#), [lambertW](#).

Examples

```
probs <- seq(0.01, 0.99, by = 0.01)
Shape <- exp(-1); Scale <- exp(1); eps = Epsilon <- exp(-1)
max(abs(pmakeham(qmakeham(probs, sca = Scale, Shape, eps = Epsilon),
  sca = Scale, Shape, eps = Epsilon) - probs)) # Should be 0

## Not run: x <- seq(-0.1, 2.0, by = 0.01);
plot(x, dmakeham(x, sca = Scale, Shape, eps = Epsilon), type = "l",
  main = "Blue is density, orange is the CDF",
  sub = "Purple lines are the 10,20,...,90 percentiles",
  col = "blue", las = 1, ylab = "")
abline(h = 0, col = "blue", lty = 2)
lines(x, pmakeham(x, sca = Scale, Shape, eps = Epsilon), col = "orange")
probs <- seq(0.1, 0.9, by = 0.1)
Q <- qmakeham(probs, sca = Scale, Shape, eps = Epsilon)
lines(Q, dmakeham(Q, sca = Scale, Shape, eps = Epsilon),
  col = "purple", lty = 3, type = "h")
pmakeham(Q, sca = Scale, Shape, eps = Epsilon) - probs # Should be all 0
abline(h = probs, col = "purple", lty = 3)
## End(Not run)
```

makeham

Makeham Regression Family Function

Description

Maximum likelihood estimation of the 3-parameter Makeham distribution.

Usage

```
makeham(lscale = "loglink", lshape = "loglink", lepsilon = "loglink",
        iscale = NULL,  lshape = NULL,  ieppsilon = NULL,
        gscale = exp(-5:5), gshape = exp(-5:5), geppsilon = exp(-4:1),
        nsimEIM = 500, oim.mean = TRUE, zero = NULL, nowarning = FALSE)
```

Arguments

`nowarning` Logical. Suppress a warning? Ignored for **VGAM** 0.9-7 and higher.

`lshape, lscale, lepsilon` Parameter link functions applied to the shape parameter `shape`, scale parameter `scale`, and other parameter `epsilon`. All parameters are treated as positive here (cf. `dmakeham` allows `epsilon = 0`, etc.). See [Links](#) for more choices.

`ishape, iscale, ieppsilon` Optional initial values. A `NULL` means a value is computed internally. A value must be given for `iepsilon` currently, and this is a sensitive parameter!

`gshape, gscale, geppsilon` See [CommonVGAMffArguments](#).

`nsimEIM, zero` See [CommonVGAMffArguments](#). Argument `probs.y` is used only when `imethod = 2`.

`oim.mean` To be currently ignored.

Details

The Makeham distribution, which adds another parameter to the Gompertz distribution, has cumulative distribution function

$$F(y; \alpha, \beta, \varepsilon) = 1 - \exp \left\{ -y\varepsilon + \frac{\alpha}{\beta} [1 - e^{\beta y}] \right\}$$

which leads to a probability density function

$$f(y; \alpha, \beta, \varepsilon) = [\varepsilon + \alpha e^{\beta y}] \exp \left\{ -y\varepsilon + \frac{\alpha}{\beta} [1 - e^{\beta y}] \right\},$$

for $\alpha > 0$, $\beta > 0$, $\varepsilon \geq 0$, $y > 0$. Here, β is called the scale parameter `scale`, and α is called a shape parameter. The moments for this distribution do not appear to be available in closed form.

Simulated Fisher scoring is used and multiple responses are handled.

Value

An object of class `"vglmff"` (see [vglmff-class](#)). The object is used by modelling functions such as `vglm`, and `vgam`.

Warning

A lot of care is needed because this is a rather difficult distribution for parameter estimation, especially when the shape parameter is large relative to the scale parameter. If the self-starting initial values fail then try experimenting with the initial value arguments, especially `iepsilon`. Successful convergence depends on having very good initial values. More improvements could be made here. Also, monitor convergence by setting `trace = TRUE`.

A trick is to fit a [gompertz](#) distribution and use it for initial values; see below. However, this family function is currently numerically fraught.

Author(s)

T. W. Yee

See Also

[dmakeham](#), [gompertz](#), [simulate.vlm](#).

Examples

```
## Not run: set.seed(123)
mdata <- data.frame(x2 = runif(nn <- 1000))
mdata <- transform(mdata, eta1 = -1,
                  ceta1 = 1,
                  eeta1 = -2)
mdata <- transform(mdata, shape1 = exp(eta1),
                  scale1 = exp(ceta1),
                  epsil1 = exp(eeta1))
mdata <- transform(mdata,
                  y1 = rmakeham(nn, shape = shape1, scale = scale1, eps = epsil1))

# A trick is to fit a Gompertz distribution first
fit0 <- vglm(y1 ~ 1, gompertz, data = mdata, trace = TRUE)
fit1 <- vglm(y1 ~ 1, makeham, data = mdata,
            etastart = cbind(predict(fit0), log(0.1)), trace = TRUE)

coef(fit1, matrix = TRUE)
summary(fit1)

## End(Not run)
```

Description

Marginal effects for the multinomial logit model and cumulative logit/probit/... models and continuation ratio models and stopping ratio models and adjacent categories models: the derivative of the fitted probabilities with respect to each explanatory variable.

Usage

```
margeff(object, subset = NULL, ...)
```

Arguments

object	A <code>vglm</code> object, with one of the following family functions: <code>multinomial</code> , <code>cumulative</code> , <code>cratio</code> , <code>sratio</code> , <code>acat</code> , <code>poissonff</code> , <code>negbinomial</code> or <code>posnegbinomial</code> .
subset	Numerical or logical vector, denoting the required observation(s). Recycling is used if possible. The default means all observations.
...	further arguments passed into the other methods functions.

Details

Computes the derivative of the fitted probabilities of the categorical response model with respect to each explanatory variable. Formerly one big function, this function now uses S4 dispatch to break up the computations.

The function `margeff()` is *not* generic. However, it calls the function `margeffS4VGAM()` which *is*. This is based on the class of the `VGAMff` argument, and it uses the S4 function `setMethod` to correctly dispatch to the required methods function. The inheritance is given by the `vfamil` slot of the **VGAM** family function.

Value

A p by $M + 1$ by n array, where p is the number of explanatory variables and the (hopefully) nominal response has $M + 1$ levels, and there are n observations.

In general, if `is.numeric(subset)` and `length(subset) == 1` then a p by $M + 1$ matrix is returned.

Warning

Care is needed in interpretation, e.g., the change is not universally accurate for a unit change in each explanatory variable because eventually the ‘new’ probabilities may become negative or greater than unity. Also, the ‘new’ probabilities will not sum to one.

This function is not applicable for models with data-dependent terms such as `bs` and `poly`. Also the function should not be applied to models with any terms that have generated more than one column of the LM model matrix, such as `bs` and `poly`. For such try using numerical methods such as finite-differences. The formula in `object` should comprise of simple terms of the form $\sim x_2 + x_3 + x_4$, etc.

Some numerical problems may occur if the fitted values are close to 0 or 1 for the `cratio` and `sratio` models. Models with offsets may result in an incorrect answer.

Note

For `multinomial` this function should handle any value of `refLevel` and also any constraint matrices. However, it does not currently handle the `xij` or `form2` arguments, nor `vgam` objects.

If marginal effects are to be computed for some values not equal to those used in the training set, then the `@x` and the `@predictors` slots both need to be assigned. See Example 3 below.

Some other limitations are imposed, e.g., for `acat` models only a `loglink` link is allowed.

Author(s)

T. W. Yee, with some help and motivation from Stasha Rmandic.

See Also

[multinomial](#), [cumulative](#), [propodds](#), [acat](#), [cratio](#), [sratio](#), [poissonff](#), [negbinomial](#), [vglm](#).

Examples

```
# Not a good example for multinomial() since the response is ordinal!!
ii <- 3; hh <- 1/100
pneumo <- transform(pneumo, let = log(exposure.time))
fit <- vglm(cbind(normal, mild, severe) ~ let, multinomial, pneumo)
fit <- vglm(cbind(normal, mild, severe) ~ let,
            cumulative(reverse = TRUE, parallel = TRUE),
            data = pneumo)
fitted(fit)[ii, ]

mynewdata <- with(pneumo, data.frame(let = let[ii] + hh))
(newp <- predict(fit, newdata = mynewdata, type = "response"))

# Compare the difference. Should be the same as hh --> 0.
round((newp-fitted(fit)[ii, ]) / hh, 3) # Finite-diff approxn
round(margeff(fit, subset = ii)["let",], 3)

# Other examples
round(margeff(fit), 3)
round(margeff(fit, subset = 2)["let",], 3)
round(margeff(fit, subset = c(FALSE, TRUE))["let",,], 3) # Recycling
round(margeff(fit, subset = c(2, 4, 6, 8))["let",,], 3)

# Example 3; margeffs at a new value
mynewdata2a <- data.frame(let = 2) # New value
mynewdata2b <- data.frame(let = 2 + hh) # For finite-diff approxn
(neweta2 <- predict(fit, newdata = mynewdata2a))
fit@x[1, ] <- c(1, unlist(mynewdata2a))
fit@predictors[1, ] <- neweta2 # Needed
max(abs(margeff(fit, subset = 1)["let", ] - (
  predict(fit, newdata = mynewdata2b, type = "response") -
  predict(fit, newdata = mynewdata2a, type = "response"))) / hh
)) # Should be 0
```

Description

Some marital data mainly from a large NZ company collected in the early 1990s.

Usage

```
data(marital.nz)
```

Format

A data frame with 6053 observations on the following 3 variables.

`age` a numeric vector, age in years

`ethnicity` a factor with levels European Maori Other Polynesian. Only Europeans are included in the data set.

`mstatus` a factor with levels Divorced/Separated, Married/Partnered, Single, Widowed.

Details

This is a subset of a data set collected from a self-administered questionnaire administered in a large New Zealand workforce observational study conducted during 1992–3. The data were augmented by a second study consisting of retirees. The data can be considered a reasonable representation of the white male New Zealand population in the early 1990s.

Source

Clinical Trials Research Unit, University of Auckland, New Zealand.

References

See [bmi.nz](#) and [chest.nz](#).

Examples

```
summary(marital.nz)
```

Max

Maximums

Description

Generic function for the *maximums* (maxima) of a model.

Usage

```
Max(object, ...)
```

Arguments

`object` An object for which the computation or extraction of a maximum (or maximums) is meaningful.

`...` Other arguments fed into the specific methods function of the model. Sometimes they are fed into the methods function for [Coef](#).

Details

Different models can define a maximum in different ways. Many models have no such notion or definition.

Maximums occur in quadratic and additive ordination, e.g., CQO or CAO. For these models the maximum is the fitted value at the optimum. For quadratic ordination models there is a formula for the optimum but for additive ordination models the optimum must be searched for numerically. If it occurs on the boundary, then the optimum is undefined. For a valid optimum, the fitted value at the optimum is the maximum.

Value

The value returned depends specifically on the methods function invoked.

Author(s)

Thomas W. Yee

References

Yee, T. W. (2004). A new technique for maximum-likelihood canonical Gaussian ordination. *Ecological Monographs*, **74**, 685–701.

Yee, T. W. (2006). Constrained additive ordination. *Ecology*, **87**, 203–213.

See Also

Max.qrrvglm, [Tol](#), [Opt](#).

Examples

```
## Not run:
set.seed(111) # This leads to the global solution
hspider[,1:6] <- scale(hspider[,1:6]) # Standardized environmental vars
p1 <- cqo(cbind(Alopacce, Alopcone, Alopfabr, Arctlute, Arctperi,
              Auloalbi, Pardlugu, Pardmont, Pardnigr, Pardpull,
              Trocterr, Zoraspin) ~
          WaterCon + BareSand + FallTwig + CoveMoss + CoveHerb + ReflLux,
          poissonff, Bestof = 2, data = hspider, Crowlpositive = FALSE)
Max(p1)

index <- 1:ncol(depvar(p1))
persp(p1, col = index, las = 1, llwd = 2)
abline(h = Max(p1), lty = 2, col = index)

## End(Not run)
```

Maxwell

The Maxwell Distribution

Description

Density, distribution function, quantile function and random generation for the Maxwell distribution.

Usage

```
dmaxwell(x, rate, log = FALSE)
pmaxwell(q, rate, lower.tail = TRUE, log.p = FALSE)
qmaxwell(p, rate, lower.tail = TRUE, log.p = FALSE)
rmaxwell(n, rate)
```

Arguments

x, q, p, n	Same as Uniform .
rate	the (rate) parameter.
log	Logical. If log = TRUE then the logarithm of the density is returned.
lower.tail, log.p	Same meaning as in pnorm or qnorm .

Details

See [maxwell](#), the **VGAM** family function for estimating the (rate) parameter a by maximum likelihood estimation, for the formula of the probability density function.

Value

`dmaxwell` gives the density, `pmaxwell` gives the distribution function, `qmaxwell` gives the quantile function, and `rmaxwell` generates random deviates.

Note

The Maxwell distribution is related to the Rayleigh distribution.

Author(s)

T. W. Yee and Kai Huang

References

Balakrishnan, N. and Nevzorov, V. B. (2003). *A Primer on Statistical Distributions*. Hoboken, New Jersey: Wiley.

See Also

[maxwell](#), [Rayleigh](#), [rayleigh](#).

Examples

```
## Not run: rate <- 3; x <- seq(-0.5, 3, length = 100)
plot(x, dmaxwell(x, rate = rate), type = "l", col = "blue",
     main = "Blue is density, orange is CDF", ylab = "", las = 1,
     sub = "Purple lines are the 10,20,...,90 percentiles")
abline(h = 0, col = "blue", lty = 2)
lines(x, pmaxwell(x, rate = rate), type = "l", col = "orange")
probs <- seq(0.1, 0.9, by = 0.1)
Q <- qmaxwell(probs, rate = rate)
lines(Q, dmaxwell(Q, rate), col = "purple", lty = 3, type = "h")
lines(Q, pmaxwell(Q, rate), col = "purple", lty = 3, type = "h")
abline(h = probs, col = "purple", lty = 3)
max(abs(pmaxwell(Q, rate) - probs)) # Should be zero

## End(Not run)
```

maxwell

Maxwell Regression Family Function

Description

Estimating the parameter of the Maxwell distribution by maximum likelihood estimation.

Usage

```
maxwell(link = "loglink", zero = NULL, parallel = FALSE,
        type.fitted = c("mean", "percentiles", "Qlink"),
        percentiles = 50)
```

Arguments

link Parameter link function applied to a , which is called the parameter rate. See [Links](#) for more choices and information; a log link is the default because the parameter is positive. More information is at [CommonVGAMffArguments](#).

zero, parallel See [CommonVGAMffArguments](#).

type.fitted, percentiles See [CommonVGAMffArguments](#) for information. Using "Qlink" is for quantile-links in **VGAMextra**.

Details

The Maxwell distribution, which is used in the area of thermodynamics, has a probability density function that can be written

$$f(y; a) = \sqrt{2/\pi} a^{3/2} y^2 \exp(-0.5ay^2)$$

for $y > 0$ and $a > 0$. The mean of Y is $\sqrt{8/(a\pi)}$ (returned as the fitted values), and its variance is $(3\pi - 8)/(\pi a)$.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), [rrvglm](#) and [vgam](#).

Note

Fisher-scoring and Newton-Raphson are the same here. A related distribution is the Rayleigh distribution. This **VGAM** family function handles multiple responses. This **VGAM** family function can be mimicked by `poisson.points(ostatistic = 1.5, dimension = 2)`.

Author(s)

T. W. Yee

References

von Seggern, D. H. (1993). *CRC Standard Curves and Surfaces*, Boca Raton, FL, USA: CRC Press.

See Also

[Maxwell](#), [rayleigh](#), [poisson.points](#).

Examples

```
mdata <- data.frame(y = rmaxwell(1000, rate = exp(2)))
fit <- vglm(y ~ 1, maxwell, mdata, trace = TRUE, crit = "coef")
coef(fit, matrix = TRUE)
Coef(fit)
```

Description

Estimates the two parameters of the McCullagh (1989) distribution by maximum likelihood estimation.

Usage

```
mccullagh89(ltheta = "rhitlink", lnu = logofflink(offset = 0.5),
            itheta = NULL, inu = NULL, zero = NULL)
```

Arguments

`ltheta`, `lnu` Link functions for the θ and ν parameters. See [Links](#) for general information.

`itheta`, `inu` Numeric. Optional initial values for θ and ν . The default is to internally compute them.

`zero` See [CommonVGAMffArguments](#) for information.

Details

The McCullagh (1989) distribution has density function

$$f(y; \theta, \nu) = \frac{\{1 - y^2\}^{\nu - \frac{1}{2}}}{(1 - 2\theta y + \theta^2)^\nu \text{Beta}(\nu + \frac{1}{2}, \frac{1}{2})}$$

where $-1 < y < 1$ and $-1 < \theta < 1$. This distribution is equation (1) in that paper. The parameter ν satisfies $\nu > -1/2$, therefore the default is to use an log-offset link with offset equal to 0.5, i.e., $\eta_2 = \log(\nu + 0.5)$. The mean is of Y is $\nu\theta/(1 + \nu)$, and these are returned as the fitted values.

This distribution is related to the Leipnik distribution (see Johnson et al. (1995)), is related to ultraspherical functions, and under certain conditions, arises as exit distributions for Brownian motion. Fisher scoring is implemented here and it uses a diagonal matrix so the parameters are globally orthogonal in the Fisher information sense. McCullagh (1989) also states that, to some extent, θ and ν have the properties of a location parameter and a precision parameter, respectively.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), [rrvglm](#) and [vgam](#).

Note

Convergence may be slow or fail unless the initial values are reasonably close. If a failure occurs, try assigning the argument `inu` and/or `itheta`. Figure 1 of McCullagh (1989) gives a broad range of densities for different values of θ and ν , and this could be consulted for obtaining reasonable initial values if all else fails.

Author(s)

T. W. Yee

References

McCullagh, P. (1989). Some statistical properties of a family of continuous univariate distributions. *Journal of the American Statistical Association*, **84**, 125–129.

Johnson, N. L. and Kotz, S. and Balakrishnan, N. (1995). *Continuous Univariate Distributions*, 2nd edition, Volume 2, New York: Wiley. (pages 612–617).

See Also

[leipnik](#), [rhobitlink](#), [logofflink](#).

Examples

```
# Limit as theta = 0, nu = Inf:
mdata <- data.frame(y = rnorm(1000, sd = 0.2))
fit <- vglm(y ~ 1, mccullagh89, data = mdata, trace = TRUE)
head(fitted(fit))
with(mdata, mean(y))
summary(fit)
coef(fit, matrix = TRUE)
Coef(fit)
```

meangaitd

Mean of the GAITD Combo Density

Description

Returns the mean of a 1- or 2-parameter GAITD combo probability mass function.

Usage

```
meangaitd(theta.p, fam = c("pois", "log", "zeta"),
  a.mix = NULL, i.mix = NULL, d.mix = NULL,
  a.mlm = NULL, i.mlm = NULL, d.mlm = NULL,
  truncate = NULL, max.support = Inf,
  pobs.mix = 0, pobs.mlm = 0,
  pstr.mix = 0, pstr.mlm = 0,
  pdip.mix = 0, pdip.mlm = 0, byrow.aid = FALSE,
  theta.a = theta.p, theta.i = theta.p, theta.d = theta.p, ...)
```

Arguments

theta.p	Same as dgaitdplot ; usually of length 1 but may be of length 2.
fam	Same as dgaitdplot . The default is the first one. All other choices are listed in that vector.
a.mix, i.mix, a.mlm, i.mlm	Same as dgaitdplot .
d.mix, d.mlm	Same as dgaitdplot .
truncate, max.support	Same as dgaitdplot .
pobs.mix, pobs.mlm, byrow.aid	Same as dgaitdplot .
pstr.mix, pstr.mlm, pdip.mix, pdip.mlm	Same as dgaitdplot .

theta.a, theta.i, theta.d
Same as [dgaitdplot](#).
... Currently unused.

Details

This function returns the mean of the PMF of the GAITD combo model. Many of its arguments are the same as [dgaitdplot](#). More functionality may be added in the future, such as returning the variance.

Value

The mean.

Note

This utility function may change a lot in the future.

Author(s)

T. W. Yee.

See Also

[dgaitdplot](#), [Gaitdpois](#), [gaitdpoisson](#).

Examples

```
i.mix <- seq(0, 15, by = 5)
lambda.p <- 10
meangaitd(lambda.p, a.mix = i.mix + 1, i.mix = i.mix,
           max.support = 17, pobs.mix = 0.1, pstr.mix = 0.1)
```

melbmaxtemp

Melbourne Daily Maximum Temperatures

Description

Melbourne daily maximum temperatures in degrees Celsius over the ten-year period 1981–1990.

Usage

```
data(melbmaxtemp)
```

Format

A vector with 3650 observations.

Details

This is a time series data from Melbourne, Australia. It is commonly used to give a difficult quantile regression problem since the data is bimodal. That is, a hot day is likely to be followed by either an equally hot day or one much cooler. However, an independence assumption is typically made.

References

Hyndman, R. J. and Bashtannyk, D. M. and Grunwald, G. K. (1996). Estimating and visualizing conditional densities. *J. Comput. Graph. Statist.*, **5**(4), 315–336.

See Also

[lms.bcn.](#)

Examples

```
summary(melbmaxtemp)
## Not run:
melb <- data.frame(today      = melbmaxtemp[-1],
                   yesterday = melbmaxtemp[-length(melbmaxtemp)])
plot(today ~ yesterday, data = melb,
     xlab = "Yesterday's Max Temperature",
     ylab = "Today's Max Temperature", cex = 1.4, type = "n")
points(today ~ yesterday, melb, pch = 0, cex = 0.50, col = "blue")
abline(a = 0, b = 1, lty = 3)

## End(Not run)
```

meplot

Mean Excess Plot

Description

Mean excess plot (also known as a mean residual life plot), a diagnostic plot for the generalized Pareto distribution (GPD).

Usage

```
meplot(object, ...)
meplot.default(y, main = "Mean Excess Plot",
               xlab = "Threshold", ylab = "Mean Excess", lty = c(2, 1:2),
               conf = 0.95, col = c("blue", "black", "blue"), type = "l", ...)
meplot.vlm(object, ...)
```

Arguments

<code>y</code>	A numerical vector. NAs etc. are not allowed.
<code>main, xlab, ylab</code>	Character. Overall title for the plot, and titles for the x- and y-axes.
<code>lty</code>	Line type. The second value is for the mean excess value, the first and third values are for the envelope surrounding the confidence interval.
<code>conf</code>	Confidence level. The default results in approximate 95 percent confidence intervals for each mean excess value.
<code>col</code>	Colour of the three lines.
<code>type</code>	Type of plot. The default means lines are joined between the mean excesses and also the upper and lower limits of the confidence intervals.
<code>object</code>	An object that inherits class "v1m", usually of class <code>vglm-class</code> or <code>vgam-class</code> .
<code>...</code>	Graphical argument passed into <code>plot</code> . See <code>par</code> for an exhaustive list. The arguments <code>xlim</code> and <code>ylim</code> are particularly useful.

Details

If Y has a GPD with scale parameter σ and shape parameter $\xi < 1$, and if $y > 0$, then

$$E(Y - u | Y > u) = \frac{\sigma + \xi u}{1 - \xi}.$$

It is a linear function in u , the threshold. Note that $Y - u$ is called the *excess* and values of Y greater than u are called *exceedances*. The empirical versions used by these functions is to use sample means to estimate the left hand side of the equation. Values of u in the plot are the values of y itself. If the plot is roughly a straight line then the GPD is a good fit; this plot can be used to select an appropriate threshold value. See `gpd` for more details. If the plot is flat then the data may be exponential, and if it is curved then it may be Weibull or gamma. There is often a lot of variance/fluctuation at the RHS of the plot due to fewer observations.

The function `meplot` is generic, and `meplot.default` and `meplot.v1m` are some methods functions for mean excess plots.

Value

A list is returned invisibly with the following components.

<code>threshold</code>	The x axis values.
<code>meanExcess</code>	The y axis values. Each value is a sample mean minus a value u .
<code>plusminus</code>	The amount which is added or subtracted from the mean excess to give the confidence interval. The last value is a NA because it is based on one observation.

Note

The function is designed for speed and not accuracy, therefore huge data sets with extremely large values may cause failure (the function `cumsum` is used.) Ties may not be well handled.

Author(s)

T. W. Yee

References

Davison, A. C. and Smith, R. L. (1990). Models for exceedances over high thresholds (with discussion). *Journal of the Royal Statistical Society, Series B, Methodological*, **52**, 393–442.

Coles, S. (2001). *An Introduction to Statistical Modeling of Extreme Values*. London: Springer-Verlag.

See Also

[gpd](#).

Examples

```
## Not run: meplot(with(venice90, sealevel), las = 1) -> ii
names(ii)
abline(h = ii$meanExcess[1], col = "orange", lty = "dashed")

par(mfrow = c(2, 2))
for (ii in 1:4)
  meplot(rgpd(1000), col = c("orange", "blue", "orange"))

## End(Not run)
```

micmen

Michaelis-Menten Model

Description

Fits a Michaelis-Menten nonlinear regression model.

Usage

```
micmen(rpar = 0.001, divisor = 10, init1 = NULL, init2 = NULL,
       imethod = 1, oim = TRUE, link1 = "identitylink",
       link2 = "identitylink", firstDeriv = c("nsimEIM", "rpar"),
       probs.x = c(0.15, 0.85), nsimEIM = 500, dispersion = 0,
       zero = NULL)
```

Arguments

rpar	Numeric. Initial positive ridge parameter. This is used to create positive-definite weight matrices.
divisor	Numerical. The divisor used to divide the ridge parameter at each iteration until it is very small but still positive. The value of divisor should be greater than one.
init1, init2	Numerical. Optional initial value for the first and second parameters, respectively. The default is to use a self-starting value.

link1, link2	Parameter link function applied to the first and second parameters, respectively. See Links for more choices.
dispersion	Numerical. Dispersion parameter.
firstDeriv	Character. Algorithm for computing the first derivatives and working weights. The first is the default.
imethod, probs.x	See CommonVGAMffArguments for information.
nsimEIM, zero	See CommonVGAMffArguments for information.
oim	Use the OIM? See CommonVGAMffArguments for information.

Details

The Michaelis-Menten model is given by

$$E(Y_i) = (\theta_1 u_i) / (\theta_2 + u_i)$$

where θ_1 and θ_2 are the two parameters.

The relationship between iteratively reweighted least squares and the Gauss-Newton algorithm is given in Wedderburn (1974). However, the algorithm used by this family function is different. Details are given at the Author's web site.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Warning

This function is not (nor could ever be) entirely reliable. Plotting the fitted function and monitoring convergence is recommended.

Note

The regressor values u_i are inputted as the RHS of the form2 argument. It should just be a simple term; no smart prediction is used. It should just a single vector, therefore omit the intercept term. The LHS of the formula form2 is ignored.

To predict the response at new values of u_i one must assign the @extra slot in the fitted object these values, e.g., see the example below.

Numerical problems may occur. If so, try setting some initial values for the parameters. In the future, several self-starting initial values will be implemented.

Author(s)

T. W. Yee

References

- Seber, G. A. F. and Wild, C. J. (1989). *Nonlinear Regression*, New York: Wiley.
- Wedderburn, R. W. M. (1974). Quasi-likelihood functions, generalized linear models, and the Gauss-Newton method. *Biometrika*, **61**, 439–447.
- Bates, D. M. and Watts, D. G. (1988). *Nonlinear Regression Analysis and Its Applications*, New York: Wiley.

See Also

[enzyme](#).

Examples

```
mfit <- vglm(velocity ~ 1, micmen, data = enzyme, trace = TRUE,
            crit = "coef", form2 = ~ conc - 1)
summary(mfit)

## Not run:
plot(velocity ~ conc, enzyme, xlab = "concentration", las = 1,
     col = "blue",
     main = "Michaelis-Menten equation for the enzyme data",
     ylim = c(0, max(velocity)), xlim = c(0, max(conc)))
points(fitted(mfit) ~ conc, enzyme, col = 2, pch = "+", cex = 2)

# This predicts the response at a finer grid:
newenzyme <- data.frame(conc = seq(0, max(with(enzyme, conc)),
                             len = 200))
mfit@extra$Xm2 <- newenzyme$conc # This is needed for prediction
lines(predict(mfit, newenzyme, "response") ~ conc, newenzyme,
      col = "red")
## End(Not run)
```

mills.ratio

Mills Ratio

Description

Computes the Mills ratio.

Usage

```
mills.ratio(x)
mills.ratio2(x)
```

Arguments

x Numeric (real).

Details

The Mills ratio here is $\text{dnorm}(x) / \text{pnorm}(x)$ (some use $(1 - \text{pnorm}(x)) / \text{dnorm}(x)$). Some care is needed as x approaches $-\text{Inf}$; when x is very negative then its value approaches $-x$.

Value

`mills.ratio` returns the Mills ratio, and `mills.ratio2` returns $\text{dnorm}(x) * \text{dnorm}(x) / \text{pnorm}(x)$.

Author(s)

T. W. Yee

References

Mills, J. P. (1926). Table of the ratio: area to bounding ordinate, for any portion of normal curve. *Biometrika*. **18**(3/4), 395–400.

See Also

[Normal](#), [tobit](#), [cens.poisson](#).

Examples

```
## Not run:
curve(mills.ratio, -5, 5, col = "orange", las = 1)
curve(mills.ratio, -5, 5, col = "orange", las = 1, log = "y")

## End(Not run)
```

mix2exp

Mixture of Two Exponential Distributions

Description

Estimates the three parameters of a mixture of two exponential distributions by maximum likelihood estimation.

Usage

```
mix2exp(lphi = "logitlink", llambda = "loglink", iphi = 0.5,
        il1 = NULL, il2 = NULL, qmu = c(0.8, 0.2), nsimEIM = 100,
        zero = "phi")
```

Arguments

lphi, llambda	Link functions for the parameters ϕ and λ . The latter is the rate parameter and note that the mean of an ordinary exponential distribution is $1/\lambda$. See Links for more choices.
iphi, il1, il2	Initial value for ϕ , and optional initial value for λ_1 and λ_2 . The last two have values that must be positive. The default is to compute initial values internally using the argument qmu.
qmu	Vector with two values giving the probabilities relating to the sample quantiles for obtaining initial values for λ_1 and λ_2 . The two values are fed in as the probs argument into quantile .
nsimEIM, zero	See CommonVGAMffArguments .

Details

The probability density function can be loosely written as

$$f(y) = \phi \text{Exponential}(\lambda_1) + (1 - \phi) \text{Exponential}(\lambda_2)$$

where ϕ is the probability an observation belongs to the first group, and $y > 0$. The parameter ϕ satisfies $0 < \phi < 1$. The mean of Y is $\phi/\lambda_1 + (1 - \phi)/\lambda_2$ and this is returned as the fitted values. By default, the three linear/additive predictors are $(\text{logit}(\phi), \log(\lambda_1), \log(\lambda_2))^T$.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Warning

This **VGAM** family function requires care for a successful application. In particular, good initial values are required because of the presence of local solutions. Therefore running this function with several different combinations of arguments such as iphi, il1, il2, qmu is highly recommended. Graphical methods such as [hist](#) can be used as an aid.

This **VGAM** family function is experimental and should be used with care.

Note

Fitting this model successfully to data can be difficult due to local solutions, uniqueness problems and ill-conditioned data. It pays to fit the model several times with different initial values and check that the best fit looks reasonable. Plotting the results is recommended. This function works better as λ_1 and λ_2 become more different. The default control argument trace = TRUE is to encourage monitoring convergence.

Author(s)

T. W. Yee

See Also

[rexp](#), [exponential](#), [mix2poisson](#).

Examples

```
## Not run: lambda1 <- exp(1); lambda2 <- exp(3)
(phi <- logitlink(-1, inverse = TRUE))
mdata <- data.frame(y1 = rexp(nn <- 1000, lambda1))
mdata <- transform(mdata, y2 = rexp(nn, lambda2))
mdata <- transform(mdata, Y = ifelse(runif(nn) < phi, y1, y2))
fit <- vglm(Y ~ 1, mix2exp, data = mdata, trace = TRUE)
coef(fit, matrix = TRUE)

# Compare the results with the truth
round(rbind('Estimated' = Coef(fit),
           'Truth' = c(phi, lambda1, lambda2)), digits = 2)

with(mdata, hist(Y, prob = TRUE, main = "Orange=estimate, blue=truth"))
abline(v = 1 / Coef(fit)[c(2, 3)], lty = 2, col = "orange", lwd = 2)
abline(v = 1 / c(lambda1, lambda2), lty = 2, col = "blue", lwd = 2)

## End(Not run)
```

mix2normal

Mixture of Two Univariate Normal Distributions

Description

Estimates the five parameters of a mixture of two univariate normal distributions by maximum likelihood estimation.

Usage

```
mix2normal(lphi = "logitlink", lmu = "identitylink", lsd =
  "loglink", iphi = 0.5, imu1 = NULL, imu2 = NULL, isd1 =
  NULL, isd2 = NULL, qmu = c(0.2, 0.8), eq.sd = TRUE,
  nsimEIM = 100, zero = "phi")
```

Arguments

lphi, lmu, lsd	Link functions for the parameters ϕ , μ , and σ . See Links for more choices.
iphi	Initial value for ϕ , whose value must lie between 0 and 1.
imu1, imu2	Optional initial value for μ_1 and μ_2 . The default is to compute initial values internally using the argument qmu.
isd1, isd2	Optional initial value for σ_1 and σ_2 . The default is to compute initial values internally based on the argument qmu. Currently these are not great, therefore using these arguments where practical is a good idea.

qmu	Vector with two values giving the probabilities relating to the sample quantiles for obtaining initial values for μ_1 and μ_2 . The two values are fed in as the probs argument into quantile .
eq.sd	Logical indicating whether the two standard deviations should be constrained to be equal. If TRUE then the appropriate constraint matrices will be used.
nsimEIM	See CommonVGAMffArguments .
zero	May be an integer vector specifying which linear/additive predictors are modelled as intercept-only. If given, the value or values can be from the set $\{1, 2, \dots, 5\}$. The default is the first one only, meaning ϕ is a single parameter even when there are explanatory variables. Set zero = NULL to model all linear/additive predictors as functions of the explanatory variables. See CommonVGAMffArguments for more information.

Details

The probability density function can be loosely written as

$$f(y) = \phi N(\mu_1, \sigma_1) + (1 - \phi) N(\mu_2, \sigma_2)$$

where ϕ is the probability an observation belongs to the first group. The parameters μ_1 and μ_2 are the means, and σ_1 and σ_2 are the standard deviations. The parameter ϕ satisfies $0 < \phi < 1$. The mean of Y is $\phi\mu_1 + (1 - \phi)\mu_2$ and this is returned as the fitted values. By default, the five linear/additive predictors are $(\text{logit}(\phi), \mu_1, \log(\sigma_1), \mu_2, \log(\sigma_2))^T$. If eq.sd = TRUE then $\sigma_1 = \sigma_2$ is enforced.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Warning

Numerical problems can occur and half-stepping is not uncommon. If failure to converge occurs, try inputting better initial values, e.g., by using iphi, qmu, imu1, imu2, isd1, isd2, etc.

This **VGAM** family function is experimental and should be used with care.

Note

Fitting this model successfully to data can be difficult due to numerical problems and ill-conditioned data. It pays to fit the model several times with different initial values and check that the best fit looks reasonable. Plotting the results is recommended. This function works better as μ_1 and μ_2 become more different.

Convergence can be slow, especially when the two component distributions are not well separated. The default control argument trace = TRUE is to encourage monitoring convergence. Having eq.sd = TRUE often makes the overall optimization problem easier.

Author(s)

T. W. Yee

References

- McLachlan, G. J. and Peel, D. (2000). *Finite Mixture Models*. New York: Wiley.
- Everitt, B. S. and Hand, D. J. (1981). *Finite Mixture Distributions*. London: Chapman & Hall.

See Also

[uninormal](#), [Normal](#), [mix2poisson](#).

Examples

```
## Not run: mu1 <- 99; mu2 <- 150; nn <- 1000
sd1 <- sd2 <- exp(3)
(phi <- logitlink(-1, inverse = TRUE))
rrn <- runif(nn)
mdata <- data.frame(y = ifelse(rrn < phi, rnorm(nn, mu1, sd1),
                              rnorm(nn, mu2, sd2)))
fit <- vglm(y ~ 1, mix2normal(eq.sd = TRUE), data = mdata)

# Compare the results
cfit <- coef(fit)
round(rbind('Estimated' = c(logitlink(cfit[1], inverse = TRUE),
                             cfit[2], exp(cfit[3]), cfit[4]),
           'Truth' = c(phi, mu1, sd1, mu2)), digits = 2)

# Plot the results
xx <- with(mdata, seq(min(y), max(y), len = 200))
plot(xx, (1-phi) * dnorm(xx, mu2, sd2), type = "l", xlab = "y",
      main = "red = estimate, blue = truth",
      col = "blue", ylab = "Density")
phi.est <- logitlink(coef(fit)[1], inverse = TRUE)
sd.est <- exp(coef(fit)[3])
lines(xx, phi*dnorm(xx, mu1, sd1), col = "blue")
lines(xx, phi.est * dnorm(xx, Coef(fit)[2], sd.est), col = "red")
lines(xx, (1-phi.est)*dnorm(xx, Coef(fit)[4], sd.est), col="red")
abline(v = Coef(fit)[c(2,4)], lty = 2, col = "red")
abline(v = c(mu1, mu2), lty = 2, col = "blue")

## End(Not run)
```

mix2poisson

Mixture of Two Poisson Distributions

Description

Estimates the three parameters of a mixture of two Poisson distributions by maximum likelihood estimation.

Usage

```
mix2poisson(lphi = "logitlink", llambda = "loglink",
            iphi = 0.5, il1 = NULL, il2 = NULL,
            qmu = c(0.2, 0.8), nsimEIM = 100, zero = "phi")
```

Arguments

lphi, llambda	Link functions for the parameter ϕ and λ . See Links for more choices.
iphi	Initial value for ϕ , whose value must lie between 0 and 1.
il1, il2	Optional initial value for λ_1 and λ_2 . These values must be positive. The default is to compute initial values internally using the argument qmu.
qmu	Vector with two values giving the probabilities relating to the sample quantiles for obtaining initial values for λ_1 and λ_2 . The two values are fed in as the probs argument into quantile .
nsimEIM, zero	See CommonVGAMffArguments .

Details

The probability function can be loosely written as

$$P(Y = y) = \phi \text{Poisson}(\lambda_1) + (1 - \phi) \text{Poisson}(\lambda_2)$$

where ϕ is the probability an observation belongs to the first group, and $y = 0, 1, 2, \dots$. The parameter ϕ satisfies $0 < \phi < 1$. The mean of Y is $\phi\lambda_1 + (1 - \phi)\lambda_2$ and this is returned as the fitted values. By default, the three linear/additive predictors are $(\text{logit}(\phi), \log(\lambda_1), \log(\lambda_2))^T$.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Warning

This **VGAM** family function requires care for a successful application. In particular, good initial values are required because of the presence of local solutions. Therefore running this function with several different combinations of arguments such as iphi, il1, il2, qmu is highly recommended. Graphical methods such as [hist](#) can be used as an aid.

With grouped data (i.e., using the weights argument) one has to use a large value of nsimEIM; see the example below.

This **VGAM** family function is experimental and should be used with care.

Note

The response must be integer-valued since [dpois](#) is invoked.

Fitting this model successfully to data can be difficult due to local solutions and ill-conditioned data. It pays to fit the model several times with different initial values, and check that the best fit looks reasonable. Plotting the results is recommended. This function works better as λ_1 and λ_2 become more different. The default control argument trace = TRUE is to encourage monitoring convergence.

Author(s)

T. W. Yee

See Also[rpois](#), [poissonff](#), [mix2normal](#).**Examples**

```
## Not run: # Example 1: simulated data
nn <- 1000
mu1 <- exp(2.5) # Also known as lambda1
mu2 <- exp(3)
(phi <- logitlink(-0.5, inverse = TRUE))
mdata <- data.frame(y = rpois(nn, ifelse(runif(nn) < phi, mu1, mu2)))
mfit <- vglm(y ~ 1, mix2poisson, data = mdata)
coef(mfit, matrix = TRUE)

# Compare the results with the truth
round(rbind('Estimated' = Coef(mfit), 'Truth' = c(phi, mu1, mu2)), 2)

ty <- with(mdata, table(y))
plot(names(ty), ty, type = "h", main = "Orange=estimate, blue=truth",
      ylab = "Frequency", xlab = "y")
abline(v = Coef(mfit)[-1], lty = 2, col = "orange", lwd = 2)
abline(v = c(mu1, mu2), lty = 2, col = "blue", lwd = 2)

# Example 2: London Times data (Lange, 1997, p.31)
ltdata1 <- data.frame(deaths = 0:9,
                     freq = c(162,267,271, 185,111,61,27,8,3,1))
ltdata2 <- data.frame(y = with(ltdata1, rep(deaths, freq)))

# Usually this does not work well unless nsimEIM is large
Mfit <- vglm(deaths ~ 1, weight = freq, data = ltdata1,
            mix2poisson(iphi=0.3, il1=1, il2=2.5, nsimEIM=5000))

# This works better in general
Mfit = vglm(y ~ 1, mix2poisson(iphi=0.3, il1=1, il2=2.5), ltdata2)
coef(Mfit, matrix = TRUE)
Coef(Mfit)

## End(Not run)
```

Description

Estimates the three independent parameters of the the MNSs blood group system.

Usage

```
MNSs(link = "logitlink", imS = NULL, ims = NULL, inS = NULL)
```

Arguments

`link` Link function applied to the three parameters. See [Links](#) for more choices.

`imS, ims, inS` Optional initial value for `mS`, `ms` and `nS` respectively. A NULL means they are computed internally.

Details

There are three independent parameters: `m_S`, `m_s`, `n_S`, say, so that $n_s = 1 - m_S - m_s - n_S$. We let the eta vector (transposed) be $(g(m_S), g(m_s), g(n_S))$ where g is the link function.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Note

The input can be a 6-column matrix of counts, where the columns are MS, Ms, MNS, MNs, NS, Ns (in order). Alternatively, the input can be a 6-column matrix of proportions (so each row adds to 1) and the `weights` argument is used to specify the total number of counts for each row.

Author(s)

T. W. Yee

References

Elandt-Johnson, R. C. (1971). *Probability Models and Statistical Methods in Genetics*, New York: Wiley.

See Also

[AA.Aa.aa](#), [AB.Ab.aB.ab](#), [ABO](#), [A1A2A3](#).

Examples

```
# Order matters only:
y <- cbind(MS = 295, Ms = 107, MNS = 379, MNs = 322, NS = 102, Ns = 214)
fit <- vglm(y ~ 1, MNSs("logitlink", .25, .28, .08), trace = TRUE)
fit <- vglm(y ~ 1, MNSs(link = logitlink), trace = TRUE, crit = "coef")
Coef(fit)
rbind(y, sum(y)*fitted(fit))
sqrt(diag(vcov(fit)))
```

model.framevlm	<i>Construct the Model Frame of a VLM Object</i>
----------------	--

Description

This function returns a [data.frame](#) with the variables. It is applied to an object which inherits from class "vlm" (e.g., a fitted model of class "vg1m").

Usage

```
model.framevlm(object, setupsmart = TRUE, wrapupsmart = TRUE, ...)
```

Arguments

object	a model object from the VGAM R package that inherits from a <i>vector linear model</i> (VLM), e.g., a model of class "vg1m".
...	further arguments such as data, na.action, subset. See model.frame for more information on these.
setupsmart, wrapupsmart	Logical. Arguments to determine whether to use smart prediction.

Details

Since object is an object which inherits from class "vlm" (e.g., a fitted model of class "vg1m"), the method will either returned the saved model frame used when fitting the model (if any, selected by argument model = TRUE) or pass the call used when fitting on to the default method.

This code implements *smart prediction* (see [smartpred](#)).

Value

A [data.frame](#) containing the variables used in the object plus those specified in ...

References

Chambers, J. M. (1992). *Data for models*. Chapter 3 of *Statistical Models in S* eds J. M. Chambers and T. J. Hastie, Wadsworth & Brooks/Cole.

See Also

[model.frame](#), [model.matrixvlm](#), [predictvg1m](#), [smartpred](#).

Examples

```
# Illustrates smart prediction
pneumo <- transform(pneumo, let = log(exposure.time))
fit <- vglm(cbind(normal,mild, severe) ~ poly(c(scale(let)), 2),
           multinomial, pneumo, trace = TRUE, x = FALSE)
class(fit)

check1 <- head(model.frame(fit))
check1
check2 <- model.frame(fit, data = head(pneumo))
check2
all.equal(unlist(check1), unlist(check2)) # Should be TRUE

q0 <- head(predict(fit))
q1 <- head(predict(fit, newdata = pneumo))
q2 <- predict(fit, newdata = head(pneumo))
all.equal(q0, q1) # Should be TRUE
all.equal(q1, q2) # Should be TRUE
```

model.matrixqrrvglm *Construct the Model Matrix of a QRR-VGLM Object*

Description

Creates a model matrix. Two types can be returned: a large one (class "v1m" or one that inherits from this such as "vglm") or a small one (such as returned if it were of class "1m").

Usage

```
model.matrixqrrvglm(object, type = c("latvar", "1m", "v1m"), ...)
```

Arguments

object	an object of a class "qrrvglm", i.e., a <code>cqo</code> object.
type	Type of model (or design) matrix returned. The first is the default. The value "latvar" is model matrix mainly comprising of the latent variable values (sometimes called the <i>site scores</i>). The value "1m" is the LM matrix directly corresponding to the formula argument. The value "v1m" is the big VLM model matrix <i>given C</i> .
...	further arguments passed to or from other methods.

Details

This function creates one of several design matrices from object. For example, this can be a small LM object or a big VLM object.

When type = "v1m" this function calls `fnumat2R()` to construct the big model matrix *given C*. That is, the constrained coefficients are assumed known, so that something like a large Poisson or logistic

regression is set up. This is because all responses are fitted simultaneously here. The columns are labelled in the following order and with the following prefixes: "A" for the A matrix (linear in the latent variables), "D" for the D matrix (quadratic in the latent variables), "x1." for the $B1$ matrix (usually contains the intercept; see the argument `noRRR` in `qrrvglm.control`).

Value

The design matrix *after scaling* for a regression model with the specified formula and data. By *after scaling*, it is meant that it matches the output of `coef(qrrvglmObject)` rather than the original scaling of the fitted object.

See Also

[model.matrixvlm](#), [cqo](#), [vcovqrrvglm](#).

Examples

```
## Not run:
set.seed(1); n <- 40; p <- 3; S <- 4; myrank <- 1
mydata <- rcqo(n, p, S, Rank = myrank, es.opt = TRUE, eq.max = TRUE)
(myform <- attr(mydata, "formula"))
mycqo <- cqo(myform, poissonff, data = mydata,
             I.tol = TRUE, Rank = myrank, Bestof = 5)
model.matrix(mycqo, type = "latvar")
model.matrix(mycqo, type = "lm")
model.matrix(mycqo, type = "vlm")

## End(Not run)
```

model.matrixvlm

Construct the Design Matrix of a VLM Object

Description

Creates a design matrix. Two types can be returned: a large one (class "vlm" or one that inherits from this such as "vglm") or a small one (such as returned if it were of class "lm").

Usage

```
model.matrixvlm(object, type = c("vlm", "lm", "lm2", "bothlm2"),
               linpred.index = NULL, label.it = TRUE, ...)
```

Arguments

object	an object of a class that inherits from the <i>vector linear model</i> (VLM).
type	Type of design matrix returned. The first is the default. The value "vlm" is the VLM model matrix corresponding to the formula argument. The value "lm" is the LM model matrix corresponding to the formula argument. The value "lm2" is the second (LM) model matrix corresponding to the form2 argument. The value "bothlm2" means both LM and VLM model matrices.

linpred.index	Vector of integers. The index for a linear/additive predictor, it must have values from the set 1:M. Also, if <code>length(linpred.index) == 1</code> then <code>type = "lm"</code> must be assigned, whereas if <code>length(linpred.index) > 1</code> then <code>type = "vlm"</code> must be assigned. Then it returns a subset of the VLM matrix corresponding to the <code>linpred.index</code> th linear/additive predictor(s); this is a LM-type matrix when it is of unit length. Currently some attributes are returned, but these may change in value in the future because of ongoing development work.
label.it	Logical. Label the row and columns with character names? If FALSE, time and memory might be saved if the big model matrix is very large. The argument is only used when <code>type = "vlm"</code> .
...	further arguments passed to or from other methods. These include <code>data</code> (which is a data frame created with <code>model.framevlm</code>), <code>contrasts.arg</code> , and <code>xlev</code> . See <code>model.matrix</code> for more information.

Details

This function creates a design matrix from object. This can be a small LM object or a big VLM object (default). The latter is constructed from the former and the constraint matrices.

This code implements *smart prediction* (see `smartpred`).

Value

The design matrix for a regression model with the specified formula and data. If `type = "bothlm2"` then a list is returned with components "X" and "Xm2".

Sometimes (especially if `x = TRUE` when calling `vglm`) the model matrix has attributes: "assign" ("lm"-type) and "vassign" ("vlm"-type) and "orig.assign.lm" ("lm"-type). These are used internally a lot for bookkeeping, especially regarding the columns of both types of model matrices. In particular, constraint matrices and variable selection relies on this information a lot. The "orig.assign.lm" is the ordinary "assign" attribute for `lm` and `glm` objects.

References

Chambers, J. M. (1992). *Data for models*. Chapter 3 of *Statistical Models in S* eds J. M. Chambers and T. J. Hastie, Wadsworth & Brooks/Cole.

See Also

`model.matrix`, `model.framevlm`, `predictvglm`, `smartpred`, `constraints.vlm`, `trim.constraints`, `add1.vglm`, `drop1.vglm`, `step4vglm`.

Examples

```
# (I) Illustrates smart prediction ,,,,,,,,,,,,,,,,,,,,,,
pneumo <- transform(pneumo, let = log(exposure.time))
fit <- vglm(cbind(normal, mild, severe) ~
            sm.poly(c(sm.scale(let)), 2),
            multinomial, data = pneumo, trace = TRUE, x = FALSE)
class(fit)
fit@smart.prediction # Data-dependent parameters
```

```

fit@x # Not saved on the object
model.matrix(fit)
model.matrix(fit, linpred.index = 1, type = "lm")
model.matrix(fit, linpred.index = 2, type = "lm")

(Check1 <- head(model.matrix(fit, type = "lm")))
(Check2 <- model.matrix(fit, data = head(pneumo), type = "lm"))
all.equal(c(Check1), c(Check2)) # Should be TRUE

q0 <- head(predict(fit))
q1 <- head(predict(fit, newdata = pneumo))
q2 <- predict(fit, newdata = head(pneumo))
all.equal(q0, q1) # Should be TRUE
all.equal(q1, q2) # Should be TRUE

# (II) Attributes ,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,
fit2 <- vglm(cbind(normal, mild, severe) ~ let, # x = TRUE
            multinomial, data = pneumo, trace = TRUE)
fit2@x # "lm"-type; saved on the object; note the attributes
model.matrix(fit2, type = "lm") # Note the attributes
model.matrix(fit2, type = "vlm") # Note the attributes

```

moffset

Matrix Offset

Description

Modify a matrix by shifting successive elements.

Usage

```

moffset(mat, roffset = 0, coffset = 0, postfix = "",
        rprefix = "Row.", cprefix = "Col.")

```

Arguments

mat	Data frame or matrix. This ought to have at least three rows and three columns. The elements are shifted in the order of <code>c(mat)</code> , i.e., going down successive columns, as the columns go from left to right. Wrapping of values is done.
roffset, coffset	Numeric or character. If numeric, the amount of shift (offset) for each row and column. The default is no change to <code>mat</code> . If character, the offset is computed by matching with the row or column names. For example, for the <code>alcoff</code> , put <code>roffset = "6"</code> means that we make an effective day's dataset start from 6:00 am, and this wraps around to include midnight to 05.59 am on the next day.
postfix	Character. Modified rows and columns are renamed by pasting this argument to the end of each name. The default is no change.
rprefix, cprefix	Same as <code>rcim</code> .

Details

This function allows a matrix to be rearranged so that element (roffset + 1, coffset + 1) becomes the (1, 1) element. The elements are assumed to be ordered in the same way as the elements of `c(mat)`,

This function is applicable to, e.g., `alcoff`, where it is useful to define the *effective day* as starting at some other hour than midnight, e.g., 6.00am. This is because partying on Friday night continues on into Saturday morning, therefore it is more interpretable to use the effective day when considering a daily effect.

This is a data preprocessing function for `rcim` and `plotrcim0`. The differences between `Rcim` and `moffset` is that `Rcim` only reorders the level of the rows and columns so that the data is shifted but not moved. That is, a value in one row stays in that row, and ditto for column. But in `moffset` values in one column can be moved to a previous column. See the examples below.

Value

A matrix of the same dimensional as its input.

Note

The input `mat` should have row names and column names.

Author(s)

T. W. Yee, Alfian F. Hadi.

See Also

`Rcim`, `rcim`, `plotrcim0`, `alcoff`, `crashi`.

Examples

```
# Some day's data is moved to previous day:
moffset(alcoff, 3, 2, "*")
Rcim(alcoff, 3 + 1, 2 + 1) # Data does not move as much.
alcoff # Original data
moffset(alcoff, 3, 2, "*") -
Rcim(alcoff, 3+1, 2+1) # Note the differences

# An 'effective day' data set:
alcoff.e <- moffset(alcoff, roffset = "6", postfix = "*")
fit.o <- rcim(alcoff) # default baselines are 1st row and col
fit.e <- rcim(alcoff.e) # default baselines are 1st row and col

## Not run: par(mfrow = c(2, 2), mar = c(9, 4, 2, 1))
plot(fit.o, rsub = "Not very interpretable",
     csub = "Not very interpretable")
plot(fit.e, rsub = "More interpretable",
     csub = "More interpretable")

## End(Not run)
```

```
# Some checking
all.equal(moffset(alcoff), alcoff) # Should be no change
moffset(alcoff, 1, 1, "*")
moffset(alcoff, 2, 3, "*")
moffset(alcoff, 1, 0, "*")
moffset(alcoff, 0, 1, "*")
moffset(alcoff, "6", "Mon", "*") # This one is good

# Customise row and column baselines
fit2 <- rcim(Rcim(alcoff.e, rbaseline = "11", cbaseline = "Mon*"))
```

multilogitlink

Multi-logit Link Function

Description

Computes the multilogit transformation, including its inverse and the first two derivatives.

Usage

```
multilogitlink(theta, refLevel = "(Last)", M = NULL, whitespace = FALSE,
               bvalue = NULL, inverse = FALSE, deriv = 0, all.derivs = FALSE,
               short = TRUE, tag = FALSE)
```

Arguments

theta Numeric or character. See below for further details.
refLevel, M, whitespace See [multinomial](#).
bvalue See [Links](#).
all.derivs Logical. This is currently experimental only.
inverse, deriv, short, tag Details at [Links](#).

Details

The `multilogitlink()` link function is a generalization of the [logitlink](#) link to M levels/classes. It forms the basis of the [multinomial](#) logit model. It is sometimes called the *multi-logit* link or the *multinomial logit* link; some people use *softmax* too. When its inverse function is computed it returns values which are positive and add to unity.

Value

For `multilogitlink` with `deriv = 0`, the multilogit of θ , i.e., $\log(\theta[, j]/\theta[, M+1])$ when `inverse = FALSE`, and if `inverse = TRUE` then $\exp(\theta[, j])/(1+\text{rowSums}(\exp(\theta)))$.

For `deriv = 1`, then the function returns $d\eta/d\theta$ as a function of θ if `inverse = FALSE`, else if `inverse = TRUE` then it returns the reciprocal.

Here, all logarithms are natural logarithms, i.e., to base e .

Note

Numerical instability may occur when theta is close to 1 or 0 (for `multilogitlink`). One way of overcoming this is to use, e.g., `bvalue`. Currently `care.exp()` is used to avoid NAs being returned if the probability is too close to 1.

Author(s)

Thomas W. Yee

References

McCullagh, P. and Nelder, J. A. (1989). *Generalized Linear Models*, 2nd ed. London: Chapman & Hall.

See Also

[Links](#), [multinomial](#), [logitlink](#), [gaitdpoisson](#), [normal.vcm](#), [CommonVGAMffArguments](#).

Examples

```
pneumo <- transform(pneumo, let = log(exposure.time))
fit <- vglm(cbind(normal, mild, severe) ~ let, # For illustration only!
           multinomial, trace = TRUE, data = pneumo)
fitted(fit)
predict(fit)

multilogitlink(fitted(fit))
multilogitlink(fitted(fit)) - predict(fit) # Should be all 0s

multilogitlink(predict(fit), inverse = TRUE) # rowSums() add to unity
multilogitlink(predict(fit), inverse = TRUE, refLevel = 1)
multilogitlink(predict(fit), inverse = TRUE) -
fitted(fit) # Should be all 0s

multilogitlink(fitted(fit), deriv = 1)
multilogitlink(fitted(fit), deriv = 2)
```

multinomial

Multinomial Logit Model

Description

Fits a multinomial logit model (MLM) to a (preferably unordered) factor response.

Usage

```
multinomial(zero = NULL, parallel = FALSE, nointercept = NULL,
            refLevel = "(Last)", ynames = FALSE,
            imethod = 1, imu = NULL, byrow.arg = FALSE,
            Thresh = NULL, Trev = FALSE,
            Tref = if (Trev) "M" else 1, whitespace = FALSE)
```

Arguments

- zero** Can be an integer-valued vector specifying which linear/additive predictors are modelled as intercepts only. Any values must be from the set $\{1, 2, \dots, M\}$. The default value means none are modelled as intercept-only terms. See [CommonVGAMffArguments](#) for more information.
- parallel** A logical, or formula specifying which terms have equal/unequal coefficients.
- ynames** Logical. If TRUE then "mu[, 1]" is replaced by the probability of the first named response category, etc. (e.g., "P[normal]"), so that the output is more readable, albeit less compact. This is seen in output such as `predict(fit)` and `coef(fit, matrix = TRUE)`. Of course, "mu" stands for the fitted probabilities, and it remains the default for upward compatibility and predictability.
- nointercept, whitespace** See [CommonVGAMffArguments](#) for details.
- imu, byrow.arg** See [CommonVGAMffArguments](#) for details.
- refLevel** Either a (1) single positive integer or (2) a value of the factor or (3) a character string. If inputted as an integer then it specifies which column of the response matrix is the reference or baseline level. The default is the *last* one (the $(M + 1)$ th one). If used, this argument will be usually assigned the value 1. If inputted as a value of a factor then beware of missing values of certain levels of the factor (`drop.unused.levels = TRUE` or `drop.unused.levels = FALSE`). See the example below. If inputted as a character string then this should be equal to (A) one of the levels of the factor response, else (B) one of the column names of the matrix response of counts; e.g., `vglm(cbind(normal, mild, severe) ~ let, multinomial(refLevel = "severe"), data = pneumo)` if it was (incorrectly because the response is ordinal) applied to the [pneumo](#) data set. Another example is `vglm(ethnicity ~ age, multinomial(refLevel = "European"), data = xs.nz)` if it was applied to the [xs.nz](#) data set.
- imethod** Choosing 2 will use the mean sample proportions of each column of the response matrix, which corresponds to the MLEs for intercept-only models. See [CommonVGAMffArguments](#) for more details.
- Thresh, Trev, Tref** Same as [cumulative](#). Because these arguments concern the intercepts, they should not be confused with the *stereotype* model where they would be applied to the **A** matrix instead.

Details

In this help file the response Y is assumed to be a factor with unordered values $1, 2, \dots, M + 1$, so that M is the number of linear/additive predictors η_j .

The default model can be written

$$\eta_j = \log(P[Y = j]/P[Y = M + 1])$$

where η_j is the j th linear/additive predictor. Here, $j = 1, \dots, M$, and η_{M+1} is 0 by definition. That is, the last level of the factor, or last column of the response matrix, is taken as the reference level or baseline—this is for identifiability of the parameters. The reference or baseline level can be changed with the `refLevel` argument.

In almost all the literature, the constraint matrices associated with this family of models are known. For example, setting `parallel = TRUE` will make all constraint matrices (including the intercept) equal to a vector of M 1's; to suppress the intercepts from being parallel then set `parallel = FALSE ~ 1`. If the constraint matrices are unknown and to be estimated, then this can be achieved by fitting the model as a reduced-rank vector generalized linear model (RR-VGLM; see `rrvglm`). In particular, a multinomial logit model with unknown constraint matrices is known as a *stereotype* model (Anderson, 1984), and can be fitted with `rrvglm`.

The above details correspond to the ordinary MLM where all the levels are *altered* (in the terminology of GAITD regression).

Value

An object of class "vglmff" (see `vglmff-class`). The object is used by modelling functions such as `vglm`, `rrvglm` and `vgam`.

Warning

No check is made to verify that the response is nominal.

See `CommonVGAMffArguments` for more warnings.

Note

The response should be either a matrix of counts (with row sums that are all positive), or a factor. In both cases, the `y` slot returned by `vglm/vgam/rrvglm` is the matrix of sample proportions.

The multinomial logit model is more appropriate for a nominal (unordered) factor response than for an ordinal (ordered) factor response. Models more suited for the latter include those based on cumulative probabilities, e.g., `cumulative`.

`multinomial` is prone to numerical difficulties if the groups are separable and/or the fitted probabilities are close to 0 or 1. The fitted values returned are estimates of the probabilities $P[Y = j]$ for $j = 1, \dots, M + 1$. See `safeBinaryRegression` for the logistic regression case.

Here is an example of the usage of the `parallel` argument. If there are covariates `x2`, `x3` and `x4`, then `parallel = TRUE ~ x2 + x3 - 1` and `parallel = FALSE ~ x4` are equivalent. This would constrain the regression coefficients for `x2` and `x3` to be equal; those of the intercepts and `x4` would be different.

In Example 4 below, a conditional logit model is fitted to an artificial data set that explores how cost and travel time affect people's decision about how to travel to work. Walking is the baseline group. The variable `Cost.car` is the difference between the cost of travel to work by car and walking, etc. The variable `Time.car` is the difference between the travel duration/time to work by car and walking, etc. For other details about the `xij` argument see `vglm.control` and `fill1`.

The `multinom` function in the `nnet` package uses the first level of the factor as baseline, whereas the last level of the factor is used here. Consequently the estimated regression coefficients differ.

Author(s)

Thomas W. Yee

References

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- Yee, T. W. and Hastie, T. J. (2003). Reduced-rank vector generalized linear models. *Statistical Modelling*, **3**, 15–41.
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See Also

[multilogitlink](#), [margeff](#), [cumulative](#), [acat](#), [cratio](#), [sratio](#), [CM.equid](#), [CommonVGAMffArguments](#), [dirichlet](#), [dirmultinomial](#), [rrvglm](#), [fill1](#), [Multinomial](#), [gaitdpoisson](#), [Gaitdpois](#), [iris](#).

Examples

```
# Example 1: Regn spline VGAM: marital status versus age
data(marital.nz)
ooo <- with(marital.nz, order(age))
om.nz <- marital.nz[ooo, ]
fit1 <- vglm(mstatus ~ sm.bs(age), multinomial, om.nz)
coef(fit1, matrix = TRUE) # Mostly meaningless
## Not run: with(om.nz,
matplot(age, fitted(fit1), type = "l", las = 1, lwd = 2))
legend("topright", leg = colnames(fitted(fit1)),
      lty = 1:4, col = 1:4, lwd = 2)
## End(Not run)

# Example 2a: a simple example
ycounts <- t(rmultinom(10, size = 20, prob = c(0.1, 0.2, 0.8)))
fit <- vglm(ycounts ~ 1, multinomial)
head(fitted(fit)) # Proportions
fit@prior.weights # NOT recommended for the prior weights
weights(fit, type = "prior", matrix = FALSE) # The better method
```

```

deivar(fit)          # Sample proportions; same as fit@y
constraints(fit)     # Constraint matrices

# Example 2b: Different reference level used as the baseline
fit2 <- vglm(ycounts ~ 1, multinomial(refLevel = 2))
coef(fit2, matrix = TRUE)
coef(fit , matrix = TRUE) # Easy to reconcile this output with fit2

# Example 3: The response is a factor.
nn <- 10
dframe3 <- data.frame(yfac = gl(3, nn, labels = c("Ctrl",
        "Trt1", "Trt2")),
        x2 = runif(3 * nn))
myrefLevel <- with(dframe3, yfac[12])
fit3a <- vglm(yfac ~ x2, multinomial(refLevel = myrefLevel), dframe3)
fit3b <- vglm(yfac ~ x2, multinomial(refLevel = 2), dframe3)
coef(fit3a, matrix = TRUE) # "Trt1" is the reference level
coef(fit3b, matrix = TRUE) # "Trt1" is the reference level
margeff(fit3b)

# Example 4: Fit a rank-1 stereotype model
fit4 <- rrvglm(Country ~ Width + Height + HP, multinomial, car.all)
coef(fit4) # Contains the C matrix
constraints(fit4)$HP # The A matrix
coef(fit4, matrix = TRUE) # The B matrix
Coef(fit4)@C # The C matrix
concoef(fit4) # Better to get the C matrix this way
Coef(fit4)@A # The A matrix
svd(coef(fit4, matrix = TRUE)[-1, ])$d # Has rank 1; = C %*% t(A)
# Classification (but watch out for NAs in some of the variables):
apply(fitted(fit4), 1, which.max) # Classification
# Classification:
colnames(fitted(fit4))[apply(fitted(fit4), 1, which.max)]
apply(predict(fit4, car.all, type = "response"),
        1, which.max) # Ditto

# Example 5: Using the xij argument (aka conditional logit model)
set.seed(111)
nn <- 100 # Number of people who travel to work
M <- 3 # There are M+1 models of transport to go to work
ycounts <- matrix(0, nn, M+1)
ycounts[cbind(1:nn, sample(x = M+1, size = nn, replace = TRUE))] = 1
dimnames(ycounts) <- list(NULL, c("bus", "train", "car", "walk"))
gotowork <- data.frame(cost.bus = runif(nn), time.bus = runif(nn),
        cost.train= runif(nn), time.train= runif(nn),
        cost.car = runif(nn), time.car = runif(nn),
        cost.walk = runif(nn), time.walk = runif(nn))
gotowork <- round(gotowork, digits = 2) # For convenience
gotowork <- transform(gotowork,
        Cost.bus = cost.bus - cost.walk,
        Cost.car = cost.car - cost.walk,
        Cost.train = cost.train - cost.walk,

```

```

      Cost      = cost.train - cost.walk, # for labelling
      Time.bus  = time.bus   - time.walk,
      Time.car  = time.car   - time.walk,
      Time.train = time.train - time.walk,
      Time      = time.train - time.walk) # for labelling
fit <- vglm(ycounts ~ Cost + Time,
  multinomial(parall = TRUE ~ Cost + Time - 1),
  xij = list(Cost ~ Cost.bus + Cost.train + Cost.car,
            Time ~ Time.bus + Time.train + Time.car),
  form2 = ~ Cost + Cost.bus + Cost.train + Cost.car +
          Time + Time.bus + Time.train + Time.car,
  data = gotowork, trace = TRUE)
head(model.matrix(fit, type = "lm")) # LM model matrix
head(model.matrix(fit, type = "vlm")) # Big VLM model matrix
coef(fit)
coef(fit, matrix = TRUE)
constraints(fit)
summary(fit)
max(abs(predict(fit) - predict(fit, new = gotowork))) # Should be 0

```

N1binom

Linear Model and Binomial Mixed Data Type Distribution

Description

Density, and random generation for the (four parameter bivariate) Linear Model–Bernoulli copula distribution.

Usage

```

dN1binom(x1, x2, mean = 0, sd = 1, prob, apar = 0,
  copula = "gaussian", log = FALSE)
rN1binom(n, mean = 0, sd = 1, prob,
  apar = 0, copula = "gaussian")

```

Arguments

`x1, x2` vector of quantiles. The valid values of `x2` are 0 and 1.

`n` number of observations. Same as [rnorm](#).

`copula` See [N1binomial](#).

`mean, sd, prob, apar` See [N1binomial](#).

`log` Logical. If TRUE then the logarithm is returned.

Details

See [N1binomial](#), the **VGAM** family functions for estimating the parameter by maximum likelihood estimation, for details.

Value

dN1binom gives the probability density/mass function, rN1binom generates random deviate and returns a two-column matrix.

Author(s)

T. W. Yee

See Also

[N1binomial](#), [rnorm](#), [rbinom](#).

Examples

```
## Not run:
nn <- 1000; apar <- rhobitlink(1.5, inverse = TRUE)
prob <- logitlink(0.5, inverse = TRUE)
mymu <- 1; sdev <- exp(1)
mat <- rN1binom(nn, mymu, sdev, prob, apar)
bndata <- data.frame(y1 = mat[, 1], y2 = mat[, 2])
with(bndata, plot(jitter(y1), jitter(y2), col = "blue"))

## End(Not run)
```

N1binomial

Linear Model and Binomial Mixed Data Type Family Function

Description

Estimate the four parameters of the (bivariate) N_1 -binomial copula mixed data type model by maximum likelihood estimation.

Usage

```
N1binomial(lmean = "identitylink", lsd = "loglink",
           lvar = "loglink", lprob = "logitlink", lapar = "rhobitlink",
           zero = c(if (var.arg) "var" else "sd", "apar"),
           nnodes = 20, copula = "gaussian", var.arg = FALSE,
           imethod = 1, isd = NULL, iprob = NULL, iapar = NULL)
```

Arguments

lmean, lsd, lvar, lprob, lapar

Details at [CommonVGAMffArguments](#). See [Links](#) for more link function choices.

imethod, isd, iprob, iapar

Initial values. Details at [CommonVGAMffArguments](#).

zero

Details at [CommonVGAMffArguments](#).

nnodes	Number of nodes and weights for the Gauss–Hermite (GH) quadrature. While a higher value should be more accurate, setting an excessive value runs the risk of evaluating some special functions near the boundary of the parameter space and producing numerical problems.
copula	Type of copula used. Currently only the bivariate normal is used but more might be implemented in the future.
var.arg	See uninormal .

Details

The bivariate response comprises Y_1 from the linear model having parameters mean and sd for μ_1 and σ_1 , and the binary Y_2 having parameter prob for its mean μ_2 . The joint probability density/mass function is $P(y_1, Y_2 = 0) = \phi_1(y_1; \mu_1, \sigma_1)(1 - \Delta)$ and $P(y_1, Y_2 = 1) = \phi_1(y_1; \mu_1, \sigma_1)\Delta$ where Δ adjusts μ_2 according to the *association parameter* α . The quantity Δ is $\Phi((\Phi^{-1}(\mu_2) - \alpha Z_1)/\sqrt{1 - \alpha^2})$. The quantity Z_1 is $(Y_1 - \mu_1)/\sigma_1$. Thus there is an underlying bivariate normal distribution, and a copula is used to bring the two marginal distributions together. Here, $-1 < \alpha < 1$, and Φ is the cumulative distribution function [pnorm](#) of a standard univariate normal.

The first marginal distribution is a normal distribution for the linear model. The second column of the response must have values 0 or 1, e.g., Bernoulli random variables. When $\alpha = 0$ then $\Delta = \mu_2$. Together, this family function combines [uninormal](#) and [binomialff](#). If the response are correlated then a more efficient joint analysis should result.

This **VGAM** family function cannot handle multiple responses. Only a two-column matrix is allowed. The two-column fitted value matrix has columns μ_1 and μ_2 .

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Note

This **VGAM** family function is fragile. Because the EIMs are approximated by GH quadrature it is found that convergence may be a little slower than for other models whose EIM is tractable. Also, the log-likelihood may be flat at the MLE with respect to α especially because the correlation between the two mixed data types may be weak.

It pays to set `trace = TRUE` to monitor convergence, especially when `abs(apar)` is high.

Author(s)

T. W. Yee

References

Song, P. X.-K. (2007). *Correlated Data Analysis: Modeling, Analytics, and Applications*. Springer.

See Also

[rN1binom](#), [N1poisson](#), [binormalcop](#), [uninormal](#), [binomialff](#), [pnorm](#).

Examples

```

nn <- 1000; mymu <- 1; sdev <- exp(1)
apar <- rhobitlink(0.5, inverse = TRUE)
prob <- logitlink(0.5, inverse = TRUE)
mat <- rN1binom(nn, mymu, sdev, prob, apar)
nbdata <- data.frame(y1 = mat[, 1], y2 = mat[, 2])
fit1 <- vglm(cbind(y1, y2) ~ 1, N1binomial,
             nbdata, trace = TRUE)
coef(fit1, matrix = TRUE)
Coef(fit1)
head(fitted(fit1))
summary(fit1)
confint(fit1)

```

N1pois

*Linear Model and Poisson Mixed Data Type Distribution***Description**

Density, and random generation for the (four parameter bivariate) Linear Model–Poisson copula distribution.

Usage

```

dN1pois(x1, x2, mean = 0, sd = 1, lambda, apar = 0,
        doff = 5, copula = "gaussian", log = FALSE)
rN1pois(n, mean = 0, sd = 1, lambda, apar = 0,
        doff = 5, copula = "gaussian")

```

Arguments

x1, x2	vector of quantiles. The valid values of x2 are nonnegative integers.
n	number of observations. Same as rnorm .
copula	See N1poisson .
mean, sd, lambda, apar	See N1poisson .
doff	See N1poisson .
log	Logical. If TRUE then the logarithm is returned.

Details

See [N1poisson](#), the **VGAM** family functions for estimating the parameter by maximum likelihood estimation, for details.

Value

`dN1pois` gives the probability density/mass function, `rN1pois` generates random deviate and returns a two-column matrix.

Author(s)

T. W. Yee

See Also[N1poisson](#), [rnorm](#), [rpois](#).**Examples**

```
## Not run:
nn <- 1000; mymu <- 1; sdev <- exp(1)
apar <- rhobitlink(0.4, inverse = TRUE)
lambda <- loglink(1, inverse = TRUE)
mat <- rN1pois(nn, mymu, sdev, lambda, apar)
pndata <- data.frame(y1 = mat[, 1], y2 = mat[, 2])
with(pndata, plot(jitter(y1), jitter(y2), col = 4))

## End(Not run)
```

N1poisson

*Linear Model and Poisson Mixed Data Type Family Function***Description**

Estimate the four parameters of the (bivariate) N_1 -Poisson copula mixed data type model by maximum likelihood estimation.

Usage

```
N1poisson(lmean = "identitylink", lsd = "loglink",
  lvar = "loglink", llambda = "loglink", lapar = "rhobitlink",
  zero = c(if (var.arg) "var" else "sd", "apar"),
  doff = 5, nnodes = 20, copula = "gaussian",
  var.arg = FALSE, imethod = 1, isd = NULL,
  ilambda = NULL, iapar = NULL)
```

Arguments

lmean, lsd, lvar, llambda, lapar

Details at [CommonVGAMffArguments](#). See [Links](#) for more link function choices.
The second response is primarily controlled by the parameter λ_2 .

imethod, isd, ilambda, iapar

Initial values. Details at [CommonVGAMffArguments](#).

zero

Details at [CommonVGAMffArguments](#).

doff	Numeric of unit length, the denominator offset $\delta > 0$. A monotonic transformation $\Delta^* = \lambda_2^{2/3}/(\delta + \lambda_2^{2/3})$ is taken to map the Poisson mean onto the unit interval. This argument is δ . The default reflects the property that the normal approximation to the Poisson work wells for $\lambda_2 \geq 10$ or thereabouts, hence that value is mapped to the origin by <code>qnorm</code> . That's because $10^{2/3}$ is approximately 5. It is known that the λ_2 rate parameter raised to the power of $2/3$ is a transformation that approximates the normal density more closely. Alternatively, <code>delta</code> may be assigned a single negative value. If so, then $\Delta^* = \log(1+\lambda_2)/[\delta + \log(1+\lambda_2)]$ is used. For this, <code>doff = -log1p(10)</code> is suggested.
nnodes, copula	Details at N1binomial .
var.arg	See uninormal .

Details

The bivariate response comprises Y_1 from a linear model having parameters mean and sd for μ_1 and σ_1 , and the Poisson count Y_2 having parameter `lambda` for its mean λ_2 . The joint probability density/mass function is $P(y_1, Y_2 = y_2) = \phi_1(y_1; \mu_1, \sigma_1) \exp(-h^{-1}(\Delta)) [h^{-1}(\Delta)]^{y_2} / y_2!$ where Δ adjusts λ_2 according to the *association parameter* α . The quantity Δ is $\Phi((\Phi^{-1}(h(\lambda_2)) - \alpha Z_1) / \sqrt{1 - \alpha^2})$ where h maps λ_2 onto the unit interval. The quantity Z_1 is $(Y_1 - \mu_1) / \sigma_1$. Thus there is an underlying bivariate normal distribution, and a copula is used to bring the two marginal distributions together. Here, $-1 < \alpha < 1$, and Φ is the cumulative distribution function `pnorm` of a standard univariate normal.

The first marginal distribution is a normal distribution for the linear model. The second column of the response must have nonnegative integer values. When $\alpha = 0$ then $\Delta = \Delta^*$. Together, this family function combines [uninormal](#) and [poissonff](#). If the response are correlated then a more efficient joint analysis should result.

The second marginal distribution allows for overdispersion relative to an ordinary Poisson distribution—a property due to α .

This **VGAM** family function cannot handle multiple responses. Only a two-column matrix is allowed. The two-column fitted value matrix has columns μ_1 and λ_2 .

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Note

This **VGAM** family function is based on [N1binomial](#) and shares many properties with it. It pays to set `trace = TRUE` to monitor convergence, especially when `abs(apar)` is high.

Author(s)

T. W. Yee

See Also

[rN1pois](#), [N1binomial](#), [binormalcop](#), [uninormal](#), [poissonff](#), [dpois](#).

Examples

```

apar <- rhobitlink(0.3, inverse = TRUE)
nn <- 1000; mymu <- 1; sdev <- exp(1)
lambda <- loglink(1, inverse = TRUE)
mat <- rN1pois(nn, mymu, sdev, lambda, apar)
npdata <- data.frame(y1 = mat[, 1], y2 = mat[, 2])
with(npdata, var(y2) / mean(y2)) # Overdispersion
fit1 <- vglm(cbind(y1, y2) ~ 1, N1poisson,
            npdata, trace = TRUE)
coef(fit1, matrix = TRUE)
Coef(fit1)
head(fitted(fit1))
summary(fit1)
confint(fit1)

```

Nakagami

Nakagami Distribution

Description

Density, cumulative distribution function, quantile function and random generation for the Nakagami distribution.

Usage

```

dnaka(x, scale = 1, shape, log = FALSE)
pnaka(q, scale = 1, shape, lower.tail = TRUE, log.p = FALSE)
qnaka(p, scale = 1, shape, ...)
rnaka(n, scale = 1, shape, Smallno = 1.0e-6)

```

Arguments

x, q	vector of quantiles.
p	vector of probabilities.
n	number of observations. Same as in runif .
scale, shape	arguments for the parameters of the distribution. See nakagami for more details. For rnaka , arguments shape and scale must be of length 1.
Smallno	Numeric, a small value used by the rejection method for determining the upper limit of the distribution. That is, $pnaka(U) > 1 - \text{Smallno}$ where U is the upper limit.
...	Arguments that can be passed into uniroot .
log	Logical. If log = TRUE then the logarithm of the density is returned.
lower.tail, log.p	Same meaning as in pnorm or qnorm .

Details

See [nakagami](#) for more details.

Value

dnaka gives the density, pnaka gives the cumulative distribution function, qnaka gives the quantile function, and rnaka generates random deviates.

Author(s)

T. W. Yee and Kai Huang

See Also

[nakagami](#).

Examples

```
## Not run: x <- seq(0, 3.2, len = 200)
plot(x, dgamma(x, shape = 1), type = "n", col = "black", ylab = "",
      ylim = c(0,1.5), main = "dnaka(x, shape = shape)")
lines(x, dnaka(x, shape = 1), col = "orange")
lines(x, dnaka(x, shape = 2), col = "blue")
lines(x, dnaka(x, shape = 3), col = "green")
legend(2, 1.0, col = c("orange","blue","green"), lty = rep(1, len = 3),
      legend = paste("shape =", c(1, 2, 3)))

plot(x, pnorm(x), type = "n", col = "black", ylab = "",
      ylim = 0:1, main = "pnaka(x, shape = shape)")
lines(x, pnaka(x, shape = 1), col = "orange")
lines(x, pnaka(x, shape = 2), col = "blue")
lines(x, pnaka(x, shape = 3), col = "green")
legend(2, 0.6, col = c("orange","blue","green"), lty = rep(1, len = 3),
      legend = paste("shape =", c(1, 2, 3)))
## End(Not run)

probs <- seq(0.1, 0.9, by = 0.1)
pnaka(qnaka(p = probs, shape = 2), shape = 2) - probs # Should be all 0
```

Description

Estimation of the two parameters of the Nakagami distribution by maximum likelihood estimation.

Usage

```
nakagami(lscale = "loglink", lshape = "loglink", iscale = 1,
         ishape = NULL, nowarning = FALSE, zero = "shape")
```

Arguments

nowarning	Logical. Suppress a warning?
lscale, lshape	Parameter link functions applied to the <i>scale</i> and <i>shape</i> parameters. Log links ensure they are positive. See Links for more choices and information.
iscale, ishape	Optional initial values for the shape and scale parameters. For ishape, a NULL value means it is obtained in the initialize slot based on the value of iscale. For iscale, assigning a NULL means a value is obtained in the initialize slot, however, setting another numerical value is recommended if convergence fails or is too slow.
zero	See CommonVGAMffArguments .

Details

The Nakagami distribution, which is useful for modelling wireless systems such as radio links, can be written

$$f(y) = 2(shape/scale)^{shape} y^{2 \times shape - 1} \exp(-shape \times y^2 / scale) / \Gamma(shape)$$

for $y > 0$, $shape > 0$, $scale > 0$. The mean of Y is $\sqrt{scale/shape} \times \Gamma(shape + 0.5) / \Gamma(shape)$ and these are returned as the fitted values. By default, the linear/additive predictors are $\eta_1 = \log(scale)$ and $\eta_2 = \log(shape)$. Fisher scoring is implemented.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Note

The Nakagami distribution is also known as the Nakagami- m distribution, where $m = shape$ here. Special cases: $m = 0.5$ is a one-sided Gaussian distribution and $m = 1$ is a Rayleigh distribution. The second moment is $E(Y^2) = m$.

If Y has a Nakagami distribution with parameters $shape$ and $scale$ then Y^2 has a gamma distribution with shape parameter $shape$ and scale parameter $scale/shape$.

Author(s)

T. W. Yee

References

Nakagami, M. (1960). The m -distribution: a general formula of intensity distribution of rapid fading, pp.3–36 in: *Statistical Methods in Radio Wave Propagation*. W. C. Hoffman, Ed., New York: Pergamon.

See Also

[rnaka](#), [gamma2](#), [rayleigh](#).

Examples

```
nn <- 1000; shape <- exp(0); Scale <- exp(1)
ndata <- data.frame(y1 = sqrt(rgamma(nn, shape = shape, scale = Scale/shape)))
nfit <- vglm(y1 ~ 1, nakagami, data = ndata, trace = TRUE, crit = "coef")
ndata <- transform(ndata, y2 = rnaka(nn, scale = Scale, shape = shape))
nfit <- vglm(y2 ~ 1, nakagami(iscale = 3), data = ndata, trace = TRUE)
head(fitted(nfit))
with(ndata, mean(y2))
coef(nfit, matrix = TRUE)
(Cfit <- Coef(nfit))
## Not run: sy <- with(ndata, sort(y2))
hist(with(ndata, y2), prob = TRUE, main = "", xlab = "y", ylim = c(0, 0.6),
      col = "lightblue")
lines(dnaka(sy, scale = Cfit["scale"], shape = Cfit["shape"]) ~ sy,
      data = ndata, col = "orange")
## End(Not run)
```

nbcancelink

Negative Binomial Canonical Link Function

Description

Computes the negative binomial canonical link transformation, including its inverse and the first two derivatives.

Usage

```
nbcancelink(theta, size = NULL, wrt.param = NULL, bvalue = NULL,
            inverse = FALSE, deriv = 0, short = TRUE, tag = FALSE)
```

Arguments

theta	Numeric or character. Typically the mean of a negative binomial distribution (NBD). See below for further details.
size, wrt.param	size contains the k matrix which must be of a conformable dimension as theta. Also, if <code>deriv > 0</code> then <code>wrt.param</code> is either 1 or 2 (1 for with respect to the first parameter, and 2 for with respect to the second parameter (size)).
bvalue	Details at Links .
inverse, deriv, short, tag	Details at Links .

Details

The NBD canonical link is $\log(\theta/(\theta + k))$ where θ is the NBD mean. The canonical link is used for theoretically relating the NBD to GLM class.

This link function was specifically written for `negbinomial` and `negbinomial.size`, and should not be used elsewhere (these **VGAM** family functions have code that specifically handles `nbcanlink(.)`.)

Estimation with the NB canonical link has a somewhat interesting history. If we take the problem as beginning with the admission of McCullagh and Nelder (1983; first edition, p.195) [see also McCullagh and Nelder (1989, p.374)] that the NB is little used in applications and has a “problematical” canonical link then it appears only one other publicized attempt was made to solve the problem seriously. This was Hilbe, who produced a defective solution. However, Miranda and Yee (2023) solve this four-decade old problem using total derivatives and it is implemented by using `nbcanlink` with `negbinomial`. Note that early versions of **VGAM** had a defective solution.

Value

For `deriv = 0`, the above equation when `inverse = FALSE`, and if `inverse = TRUE` then `kmatrix / expm1(-theta)` where `theta` is really `eta`. For `deriv = 1`, then the function returns $d\eta/d\theta$ as a function of `theta` if `inverse = FALSE`, else if `inverse = TRUE` then it returns the reciprocal.

Note

While theoretically nice, this function is not recommended in general since its value is always negative (linear predictors ought to be unbounded in general). A `loglink` link for argument `lm` is recommended instead.

Numerical instability may occur when `theta` is close to 0 or 1. Values of `theta` which are less than or equal to 0 can be replaced by `bvalue` before computing the link function value. See [Links](#).

Author(s)

Victor Miranda and Thomas W. Yee.

References

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- McCullagh, P. and Nelder, J. A. (1989). *Generalized Linear Models*, 2nd ed. London: Chapman & Hall.
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- Yee, T. W. (2014). Reduced-rank vector generalized linear models with two linear predictors. *Computational Statistics and Data Analysis*, **71**, 889–902.

See Also

`negbinomial`, `negbinomial.size`.

Examples

```

nbcancelink("mu", short = FALSE)

mymu <- 1:10 # Test some basic operations:
kmatrix <- cbind(runif(length(mymu)))
eta1 <- nbcancelink(mymu, size = kmatrix)
ans2 <- nbcancelink(eta1, size = kmatrix, inverse = TRUE)
max(abs(ans2 - mymu)) # Should be 0

## Not run: mymu <- seq(0.5, 10, length = 101)
kmatrix <- matrix(10, length(mymu), 1)
plot(nbcancelink(mymu, size = kmatrix) ~ mymu, las = 1,
      type = "l", col = "blue", xlab = expression({mu}))

## End(Not run)

# Estimate the parameters from some simulated data
ndata <- data.frame(x2 = runif(nn <- 500))
ndata <- transform(ndata, eta1 = -1 - 1 * x2, # eta1 < 0
                  size1 = exp(1),
                  size2 = exp(2))
ndata <- transform(ndata,
                  mu1 = nbcancelink(eta1, size = size1, inverse = TRUE),
                  mu2 = nbcancelink(eta1, size = size2, inverse = TRUE))
ndata <- transform(ndata, y1 = rnbinom(nn, mu = mu1, size1),
                  y2 = rnbinom(nn, mu = mu2, size2))
summary(ndata)

nbcfit <-
  vglm(cbind(y1, y2) ~ x2, # crit = "c",
       negbinomial(lmu = "nbcancelink"),
       data = ndata, trace = TRUE)
coef(nbcfit, matrix = TRUE)
summary(nbcfit)

```

negbinomial

*Negative Binomial Distribution Family Function***Description**

Maximum likelihood estimation of the two parameters of a negative binomial distribution.

Usage

```

negbinomial(zero = "size", parallel = FALSE, deviance.arg = FALSE,
            type.fitted = c("mean", "quantiles"),
            percentiles = c(25, 50, 75), vfl = FALSE,
            mds.min = 1e-3, nsimEIM = 500, cutoff.prob = 0.999,
            eps.trig = 1e-7, max.support = 4000, max.chunk.MB = 30,

```

```

      lmu = "loglink", lsize = "loglink",
      imethod = 1, imu = NULL, iprobs.y = NULL,
      gprobs.y = ppoints(6), isize = NULL,
      gsize.mux = exp(c(-30, -20, -15, -10, -6:3))
polya(zero = "size", type.fitted = c("mean", "prob"),
      mds.min = 1e-3, nsimEIM = 500, cutoff.prob = 0.999,
      eps.trig = 1e-7, max.support = 4000, max.chunk.MB = 30,
      lprob = "logitlink", lsize = "loglink", imethod = 1, iprob = NULL,
      iprobs.y = NULL, gprobs.y = ppoints(6), isize = NULL,
      gsize.mux = exp(c(-30, -20, -15, -10, -6:3)), imunb = NULL)
polyaR(zero = "size", type.fitted = c("mean", "prob"),
      mds.min = 1e-3, nsimEIM = 500, cutoff.prob = 0.999,
      eps.trig = 1e-7, max.support = 4000, max.chunk.MB = 30,
      lsize = "loglink", lprob = "logitlink", imethod = 1, iprob = NULL,
      iprobs.y = NULL, gprobs.y = ppoints(6), isize = NULL,
      gsize.mux = exp(c(-30, -20, -15, -10, -6:3)), imunb = NULL)

```

Arguments

- zero** Can be an integer-valued vector, and if so, then it is usually assigned -2 or 2 . Specifies which of the two linear/additive predictors are modelled as an intercept only. By default, the k parameter (after `lsize` is applied) is modelled as a single unknown number that is estimated. It can be modelled as a function of the explanatory variables by setting `zero = NULL`; this has been called a NB-H model by Hilbe (2011). A negative value means that the value is recycled, so setting -2 means all k are intercept-only. See [CommonVGAMffArguments](#) for more information.
- lmu, lsize, lprob** Link functions applied to the μ , k and p parameters. See [Links](#) for more choices. Note that the μ , k and p parameters are the `mu`, `size` and `prob` arguments of [rnbinom](#) respectively. Common alternatives for `lsize` are [negloglink](#) and [reciprocallink](#), and [logloglink](#) (if $k > 1$).
- imu, imunb, isize, iprob** Optional initial values for the mean and k and p . For k , if failure to converge occurs then try different values (and/or use `imethod`). For a S -column response, `isize` can be of length S . A value `NULL` means an initial value for each response is computed internally using a gridsearch based on `gsize.mux`. The last argument is ignored if used within [cqo](#); see the `iKvector` argument of [qrrvglm.control](#) instead. In the future `isize` and `iprob` might be depreciated.
- nsimEIM** This argument is used for computing the diagonal element of the *expected information matrix* (EIM) corresponding to k based on the *simulated Fisher scoring* (SFS) algorithm. See [CommonVGAMffArguments](#) for more information and the notes below. SFS is one of two algorithms for computing the EIM elements (so that both algorithms may be used on a given data set). SFS is faster than the exact method when Q_{\max} is large.
- cutoff.prob** Fed into the `p` argument of [qnbinom](#) in order to obtain an upper limit for the approximate support of the distribution, called Q_{\max} , say. Similarly, the value $1-p$ is fed into the `p` argument of [qnbinom](#) in order to obtain a lower limit for

the approximate support of the distribution, called Q_{\min} , say. Hence the approximate support is $Q_{\min}:Q_{\max}$. This argument should be a numeric and close to 1 but never exactly 1. Used to specify how many terms of the infinite series for computing the second diagonal element of the EIM are actually used. The closer this argument is to 1, the more accurate the standard errors of the regression coefficients will be. If this argument is too small, convergence will take longer.

`max.chunk.MB`, `max.support`

`max.support` is used to describe the eligibility of individual observations to have their EIM computed by the *exact method*. Here, we are concerned about computing the EIM wrt k . The exact method algorithm operates separately on each response variable, and it constructs a large matrix provided that the number of columns is less than `max.support`. If so, then the computations are done in chunks, so that no more than about `max.chunk.MB` megabytes of memory is used at a time (actually, it is proportional to this amount). Regarding eligibility of this algorithm, each observation must have the length of the vector, starting from the `1-cutoff.prob` quantile and finishing up at the `cutoff.prob` quantile, less than `max.support` (as its approximate support). If you have abundant memory then you might try setting `max.chunk.MB = Inf`, but then the computations might take a very long time. Setting `max.chunk.MB = 0` or `max.support = 0` will force the EIM to be computed using the SFS algorithm only (this *used to be* the default method for *all* the observations). When the fitted values of the model are large and k is small, the computation of the EIM will be costly with respect to time and memory if the exact method is used. Hence the argument `max.support` limits the cost in terms of time. For intercept-only models `max.support` is multiplied by a number (such as 10) because only one inner product needs be computed. Note: `max.support` is an upper bound and limits the number of terms dictated by the `eps.trig` argument.

`mds.min`

Numeric. Minimum value of the NBD mean divided by size parameter. The closer this ratio is to 0, the closer the distribution is to a Poisson. Iterations will stop when an estimate of k is so large, relative to the mean, than it is below this threshold (this is treated as a boundary of the parameter space).

`vfl`

Logical. Fit the Variance–variance Factorized Loglinear (VFL) model? If TRUE then the constraint matrix `rbind(0, -1)` is assigned to all covariates which are not parallel. Hence `parallel` must be used in conjunction with this argument to specify the set of covariates used for modelling the mean. Note that the constraint matrix for the intercept should be parallel too. The overall resulting parameterization is the same as Evans (1953). Some general information is at [CommonVGAMffArguments](#).

`eps.trig`

Numeric. A small positive value used in the computation of the EIMs. It focusses on the denominator of the terms of a series. Each term in the series (that is used to approximate an infinite series) has a value greater than `size / sqrt(eps.trig)`, thus very small terms are ignored. It's a good idea to set a smaller value that will result in more accuracy, but it will require a greater computing time (when k is close to 0). And adjustment to `max.support` may be needed. In particular, the quantity computed by special means is $\psi'(k) - E[\psi'(Y + k)]$, which is the difference between two [trigamma](#) functions. It is part of the calculation of the EIM with respect to the size parameter.

<code>gsize.mux</code>	Similar to <code>gsigma</code> in CommonVGAMffArguments . However, this grid is multiplied by the initial estimates of the NBD mean parameter. That is, it is on a relative scale rather than on an absolute scale. If the counts are very large in value then convergence fail might occur; if so, then try a smaller value such as <code>gsize.mux = exp(-40)</code> .
<code>type.fitted, percentiles</code>	See CommonVGAMffArguments for more information.
<code>deviance.arg</code>	Logical. If TRUE, the deviance is computed <i>after</i> convergence. It only works in the NB-2 model. It is also necessary to set <code>criterion = "coefficients"</code> or <code>half.step = FALSE</code> since one cannot use that criterion properly for the minimization within the IRLS algorithm. It should be set TRUE when used with <code>cqo</code> under the fast algorithm.
<code>imethod</code>	An integer with value 1 or 2 etc. which specifies the initialization method for the μ parameter. If failure to converge occurs try another value and/or else specify a value for <code>iprobs.y</code> and/or else specify a value for <code>isize</code> .
<code>parallel</code>	Setting <code>parallel = TRUE</code> is useful in order to get something similar to quasipoisson or what is known as NB-1. If <code>parallel = TRUE</code> then the parallelism constraint does not apply to any intercept term. You should set <code>zero = NULL</code> too if <code>parallel = TRUE</code> to avoid a conflict. See CommonVGAMffArguments for more information. Argument <code>vf1</code> requires the use of <code>parallel</code> to fit the VFL model.
<code>gprobs.y</code>	A vector representing a grid; passed into the <code>probs</code> argument of quantile when <code>imethod = 1</code> to obtain an initial value for the mean of each response. Is overwritten by any value of <code>iprobs.y</code> .
<code>iprobs.y</code>	Passed into the <code>probs</code> argument of quantile when <code>imethod = 1</code> to obtain an initial value for the mean of each response. Overwrites any value of <code>gprobs.y</code> . This argument might be deleted in the future.

Details

The negative binomial distribution (NBD) can be motivated in several ways, e.g., as a Poisson distribution with a mean that is gamma distributed. There are several common parametrizations of the NBD. The one used by `negbinomial()` uses the mean μ and an *index* parameter k , both which are positive. Specifically, the density of a random variable Y is

$$f(y; \mu, k) = \binom{y+k-1}{y} \left(\frac{\mu}{\mu+k} \right)^y \left(\frac{k}{k+\mu} \right)^k$$

where $y = 0, 1, 2, \dots$, and $\mu > 0$ and $k > 0$. Note that the *dispersion* parameter is $1/k$, so that as k approaches infinity the NBD approaches a Poisson distribution. The response has variance $Var(Y) = \mu + \mu^2/k$. When fitted, the `fitted.values` slot of the object contains the estimated value of the μ parameter, i.e., of the mean $E(Y)$. It is common for some to use $\alpha = 1/k$ as the ancillary or heterogeneity parameter; so common alternatives for `lsize` are [negloglink](#) and [reciprocallink](#).

For *polya* the density is

$$f(y; p, k) = \binom{y+k-1}{y} (1-p)^y p^k$$

where $y = 0, 1, 2, \dots$, and $k > 0$ and $0 < p < 1$.

Family function `polyaR()` is the same as `polya()` except the order of the two parameters are switched. The reason is that `polyaR()` tries to match with `rnbinom` closely in terms of the argument order, etc. Should the probability parameter be of primary interest, probably, users will prefer using `polya()` rather than `polyaR()`. Possibly `polyaR()` will be decommissioned one day.

The NBD can be coerced into the classical GLM framework with one of the parameters being of interest and the other treated as a nuisance/scale parameter (this is implemented in the **MASS** library). The **VGAM** family function `negbinomial()` treats both parameters on the same footing, and estimates them both by full maximum likelihood estimation.

The parameters μ and k are independent (diagonal EIM), and the confidence region for k is extremely skewed so that its standard error is often of no practical use. The parameter $1/k$ has been used as a measure of aggregation. For the NB-C the EIM is not diagonal.

These **VGAM** family functions handle *multiple* responses, so that a response matrix can be inputted. The number of columns is the number of species, say, and setting `zero = -2` means that *all* species have a k equalling a (different) intercept only.

Conlisk, et al. (2007) show that fitting the NBD to presence-absence data will result in identifiability problems. However, the model is identifiable if the response values include 0, 1 and 2.

For the NB canonical link (NB-C), its estimation has a somewhat interesting history. Some details are at [nbcancelink](#).

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), [rrvglm](#) and [vgam](#).

Warning

Poisson regression corresponds to k equalling infinity. If the data is Poisson or close to Poisson, numerical problems may occur. Some corrective measures are taken, e.g., k is effectively capped (relative to the mean) during estimation to some large value and a warning is issued. And setting `stepsize = 0.5` for half stepping is probably a good idea too when the data is extreme.

The NBD is a strictly unimodal distribution. Any data set that does not exhibit a mode (somewhere in the middle) makes the estimation problem difficult. Set `trace = TRUE` to monitor convergence.

These functions are fragile; the maximum likelihood estimate of the index parameter is fraught (see Lawless, 1987). Other alternatives to `negbinomial` are to fit a NB-1 or RR-NB (aka NB-P) model; see Yee (2014). Also available are the NB-C, NB-H and NB-G. Assigning values to the `isize` argument may lead to a local solution, and smaller values are preferred over large values when using this argument.

If one wants to force SFS to be used on all observations, then set `max.support = 0` or `max.chunk.MB = 0`. If one wants to force the exact method to be used for all observations, then set `max.support = Inf`. If the computer has *much* memory, then trying `max.chunk.MB = Inf` and `max.support = Inf` may provide a small speed increase. If SFS is used at all, then the working weights (`@weights`) slot of the fitted object will be a matrix; otherwise that slot will be a 0×0 matrix.

An alternative to the NBD is the generalized Poisson distribution, [genpoisson1](#), [genpoisson2](#) and [genpoisson0](#), since that also handles overdispersion wrt Poisson. It has one advantage in that its EIM can be computed straightforwardly.

Yet to do: write a family function which uses the methods of moments estimator for k .

Note

These 3 functions implement 2 common parameterizations of the negative binomial (NB). Some people called the NB with integer k the *Pascal* distribution, whereas if k is real then this is the *Polya* distribution. I don't. The one matching the details of `rnbinom` in terms of p and k is `polya()`.

For `polya()` the code may fail when p is close to 0 or 1. It is not yet compatible with `cqo` or `cao`.

Suppose the response is called `yamat`. For `negbinomial()` the diagonal element of the *expected information matrix* (EIM) for parameter k involves an infinite series; consequently SFS (see `nsimEIM`) is used as the backup algorithm only. SFS should be better if $\max(\text{yamat})$ is large, e.g., $\max(\text{yamat}) > 1000$, or if there are any outliers in `yamat`. The default algorithm involves a finite series approximation to the support $\emptyset: \text{Inf}$; the arguments `max.memory`, `min.size` and `cutoff.prob` are pertinent.

Regardless of the algorithm used, convergence problems may occur, especially when the response has large outliers or is large in magnitude. If convergence failure occurs, try using arguments (in recommended decreasing order) `max.support`, `nsimEIM`, `cutoff.prob`, `iprob.y`, `imethod`, `isize`, `zero`, `max.chunk.MB`.

The function `negbinomial` can be used by the fast algorithm in `cqo`, however, setting `eq.tolerances = TRUE` and `I.tolerances = FALSE` is recommended.

In the first example below (Bliss and Fisher, 1953), from each of 6 McIntosh apple trees in an orchard that had been sprayed, 25 leaves were randomly selected. On each of the leaves, the number of adult female European red mites were counted.

There are two special uses of `negbinomial` for handling count data. Firstly, when used by `rrvglm` this results in a continuum of models in between and inclusive of quasi-Poisson and negative binomial regression. This is known as a reduced-rank negative binomial model (*RR-NB*). It fits a negative binomial log-linear regression with variance function $Var(Y) = \mu + \delta_1 \mu^{\delta_2}$ where δ_1 and δ_2 are parameters to be estimated by MLE. Confidence intervals are available for δ_2 , therefore it can be decided upon whether the data are quasi-Poisson or negative binomial, if any.

Secondly, the use of `negbinomial` with `parallel = TRUE` inside `vglm` can result in a model similar to `quasipoisson`. This is named the *NB-I* model. The dispersion parameter is estimated by MLE whereas `glm` uses the method of moments. In particular, it fits a negative binomial log-linear regression with variance function $Var(Y) = \phi_0 \mu$ where ϕ_0 is a parameter to be estimated by MLE. Confidence intervals are available for ϕ_0 .

Author(s)

Thomas W. Yee, and with a lot of help by Victor Miranda to get it going with `nbcancelink`.

References

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- Conlisk, E. and Conlisk, J. and Harte, J. (2007). The impossibility of estimating a negative binomial clustering parameter from presence-absence data: A comment on He and Gaston. *The American Naturalist* **170**, 651–654.
- Evans, D. A. (1953). Experimental evidence concerning contagious distributions in ecology. *Biometrika*, **40**(1–2), 186–211.

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See Also

[quasipoisson](#), [gaitdnbinomial](#), [poissonff](#), [zinegbinomial](#), [negbinomial.size](#) (e.g., NB-G), [nbcnlink](#) (NB-C), [posnegbinomial](#), [genpoisson1](#), [genpoisson2](#), [genpoisson0](#), [inv.binomial](#), [NegBinomial](#), [rrvglm](#), [cao](#), [cqo](#), [CommonVGAMffArguments](#), [simulate.vlm](#), [ppoints](#), [margeff](#).

Examples

```
## Not run:
# Example 1: apple tree data (Bliss and Fisher, 1953)
appletree <- data.frame(y = 0:7, w = c(70, 38, 17, 10, 9, 3, 2, 1))
fit <- vglm(y ~ 1, negbinomial(deviance = TRUE), data = appletree,
           weights = w, crit = "coef") # Obtain the deviance
fit <- vglm(y ~ 1, negbinomial(deviance = TRUE), data = appletree,
           weights = w, half.step = FALSE) # Alternative method
summary(fit)
coef(fit, matrix = TRUE)
Coef(fit) # For intercept-only models
deviance(fit) # NB2 only; needs 'crit="coef"' & 'deviance=T' above

# Example 2: simulated data with multiple responses
ndata <- data.frame(x2 = runif(nn <- 200))
ndata <- transform(ndata, y1 = rnbinom(nn, exp(1), mu = exp(3+x2)),
                  y2 = rnbinom(nn, exp(0), mu = exp(2-x2)))
fit1 <- vglm(cbind(y1, y2) ~ x2, negbinomial, ndata, trace = TRUE)
coef(fit1, matrix = TRUE)

# Example 3: large counts implies SFS is used
ndata <- transform(ndata, y3 = rnbinom(nn, exp(1), mu = exp(10+x2)))
with(ndata, range(y3)) # Large counts
fit2 <- vglm(y3 ~ x2, negbinomial, data = ndata, trace = TRUE)
coef(fit2, matrix = TRUE)
head(weights(fit2, type = "working")) # Non-empty; SFS was used

# Example 4: a NB-1 to estimate a NB with Var(Y)=phi0*mu
nn <- 200 # Number of observations
phi0 <- 10 # Specify this; should be greater than unity
delta0 <- 1 / (phi0 - 1)
mydata <- data.frame(x2 = runif(nn), x3 = runif(nn))
```

```

mydata <- transform(mydata, mu = exp(2 + 3 * x2 + 0 * x3))
mydata <- transform(mydata, y3 = rbinom(nn, delta0 * mu, mu = mu))
plot(y3 ~ x2, data = mydata, pch = "+", col = "blue",
     main = paste("Var(Y) = ", phi0, " * mu", sep = ""), las = 1)
nb1 <- vglm(y3 ~ x2 + x3, negbinomial(parallel = TRUE, zero = NULL),
           data = mydata, trace = TRUE)
# Extracting out some quantities:
cnb1 <- coef(nb1, matrix = TRUE)
mydiff <- (cnb1["(Intercept)", "loglink(size)"] -
          cnb1["(Intercept)", "loglink(mu)"])
delta0.hat <- exp(mydiff)
(phi.hat <- 1 + 1 / delta0.hat) # MLE of phi
summary(nb1)
# Obtain a 95 percent confidence interval for phi0:
myvec <- rbind(-1, 1, 0, 0)
(se.mydiff <- sqrt(t(myvec) %*% vcov(nb1) %*% myvec))
ci.mydiff <- mydiff + c(-1.96, 1.96) * c(se.mydiff)
ci.delta0 <- ci.exp.mydiff <- exp(ci.mydiff)
(ci.phi0 <- 1 + 1 / rev(ci.delta0)) # The 95
Confint.nb1(nb1) # Quick way to get it
# cf. moment estimator:
summary(glm(y3 ~ x2 + x3, quasipoisson, mydata))$disper

## End(Not run)

```

negbinomial.size

Negative Binomial Distribution Family Function With Known Size

Description

Maximum likelihood estimation of the mean parameter of a negative binomial distribution with known size parameter.

Usage

```

negbinomial.size(size = Inf, lmu = "loglink", imu = NULL,
                iprobs.y = 0.35, imethod = 1,
                ishrinkage = 0.95, zero = NULL)

```

Arguments

size	Numeric, positive. Same as argument size of <code>rbinom</code> . If the response is a matrix then this is recycled to a matrix of the same dimension, by row (<code>matrix</code> with <code>byrow = TRUE</code>).
lmu, imu	Same as <code>negbinomial</code> .
iprobs.y, imethod	Same as <code>negbinomial</code> .
zero, ishrinkage	Same as <code>negbinomial</code> . See CommonVGAMffArguments for information.

Details

This **VGAM** family function estimates only the mean parameter of the negative binomial distribution. See [negbinomial](#) for general information. Setting `size = 1` gives what might be called the NB-G (geometric model; see Hilbe (2011)). The default, `size = Inf`, corresponds to the Poisson distribution.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), [rrvglm](#) and [vgam](#).

Note

If `lmu = "nbcnlink"` in `negbinomial.size()` then the `size` argument here should be assigned and these values are recycled.

Author(s)

Thomas W. Yee

References

Hilbe, J. M. (2011). *Negative Binomial Regression*, 2nd Edition. Cambridge: Cambridge University Press.

Yee, T. W. (2014). Reduced-rank vector generalized linear models with two linear predictors. *Computational Statistics and Data Analysis*, **71**, 889–902.

See Also

[negbinomial](#), [nbcnlink](#) (NB-C model), [poissonff](#), [rnbinom](#), [simulate.vlm](#).

Examples

```
# Simulated data with various multiple responses
size1 <- exp(1); size2 <- exp(2); size3 <- exp(0); size4 <- Inf
ndata <- data.frame(x2 = runif(nn <- 1000))
ndata <- transform(ndata, eta1 = -1 - 2 * x2, # eta1 must be negative
                  size1 = size1)
ndata <- transform(ndata,
                  mu1 = nbcnlink(eta1, size = size1, inv = TRUE))
ndata <- transform(ndata,
                  y1 = rnbinom(nn, mu = mu1, size = size1), # NB-C
                  y2 = rnbinom(nn, mu = exp(2 - x2), size = size2),
                  y3 = rnbinom(nn, mu = exp(3 + x2), size = size3), # NB-G
                  y4 = rpois(nn, lambda = exp(1 + x2)))

# Also known as NB-C with size known (Hilbe, 2011)
fit1 <- vglm(y1 ~ x2, negbinomial.size(size = size1, lmu = "nbcnlink"),
            data = ndata, trace = TRUE)
coef(fit1, matrix = TRUE)
```

```

head(fit1@misc$size) # size saved here

fit2 <- vglm(cbind(y2, y3, y4) ~ x2, data = ndata, trace = TRUE,
            negbinomial.size(size = c(size2, size3, size4)))
coef(fit2, matrix = TRUE)
head(fit2@misc$size) # size saved here

```

normal.vcm

Univariate Normal Distribution as a Varying-Coefficient Model

Description

Maximum likelihood estimation of all the coefficients of a LM where each of the usual regression coefficients is modelled with other explanatory variables via parameter link functions. Thus this is a basic varying-coefficient model.

Usage

```

normal.vcm(link.list = list("(Default)" = "identitylink"),
           earg.list = list("(Default)" = list()),
           lsd = "loglink", lvar = "loglink",
           esd = list(), evar = list(),
           var.arg = FALSE, imethod = 1,
           icoefficients = NULL, isd = NULL, zero = "sd",
           sd.inflation.factor = 2.5)

```

Arguments

link.list, earg.list

Link functions and extra arguments applied to the coefficients of the LM, excluding the standard deviation/variance. See [CommonVGAMffArguments](#) for more information. The default is for an identity link to be applied to each of the regression coefficients.

lsd, esd, lvar, evar

Link function and extra argument applied to the standard deviation/variance. See [CommonVGAMffArguments](#) for more information. Same as [uninormal](#).

icoefficients Optional initial values for the coefficients. Recycled to length $M - 1$ (does not include the standard deviation/variance). Try using this argument if there is a link function that is not programmed explicitly to handle range restrictions in the initialize slot.

var.arg, imethod, isd

Same as, or similar to, [uninormal](#).

zero

See [CommonVGAMffArguments](#) for more information. The default applies to the last one, viz. the standard deviation/variance parameter.

sd.inflation.factor

Numeric, should be greater than 1. The initial value of the standard deviation is multiplied by this, unless isd is inputted. Experience has shown that it is safer to start off with a larger value rather than a smaller one.

Details

This function allows all the usual LM regression coefficients to be modelled as functions of other explanatory variables via parameter link functions. For example, we may want some of them to be positive. Or we may want a subset of them to be positive and add to unity. So a class of such models have been named *varying-coefficient models* (VCMs).

The usual linear model is specified through argument `form2`. As with all other **VGAM** family functions, the linear/additive predictors are specified through argument `formula`.

The `multilogitlink` link allows a subset of the coefficients to be positive and add to unity. Either none or more than one call to `multilogitlink` is allowed. The last variable will be used as the baseline/reference group, and therefore excluded from the estimation.

By default, the log of the standard deviation is the last linear/additive predictor. It is recommended that this parameter be estimated as intercept-only, for numerical stability.

Technically, the Fisher information matrix is of unit-rank for all but the last parameter (the standard deviation/variance). Hence an approximation is used that pools over all the observations.

This **VGAM** family function cannot handle multiple responses. Also, this function will probably not have the full capabilities of the class of varying-coefficient models as described by Hastie and Tibshirani (1993). However, it should be able to manage some simple models, especially involving the following links: `identitylink`, `loglink`, `logofflink`, `logloglink`, `logitlink`, `probitlink`, `cauchitlink`, `clogloglink`, `rhobitlink`, `fisherzlink`.

Value

An object of class "vglmff" (see `vglmff-class`). The object is used by modelling functions such as `vglm`, and `vgam`.

Warning

This **VGAM** family function is fragile. One should monitor convergence, and possibly enter initial values especially when there are non-`identity`-link functions. If the initial value of the standard deviation/variance is too small then numerical problems may occur. One trick is to fit an intercept-only model and feed its `predict()` output into argument `etastart` of a more complicated model. The use of the zero argument is recommended in order to keep models as simple as possible.

Note

The standard deviation/variance parameter is best modelled as intercept-only.

Yet to do: allow an argument such as `parallel` that enables many of the coefficients to be equal.

Fix a bug: `Coef()` does not work for intercept-only models.

Author(s)

T. W. Yee

References

Hastie, T. and Tibshirani, R. (1993). Varying-coefficient models. *J. Roy. Statist. Soc. Ser. B*, **55**, 757–796.

See Also

[uninormal](#), [lm](#).

Examples

```

ndata <- data.frame(x2 = runif(nn <- 2000))
# Note that coeff1 + coeff2 + coeff5 == 1. So try "multilogitlink".
myoffset <- 10
ndata <- transform(ndata,
  coeff1 = 0.25, # "multilogitlink"
  coeff2 = 0.25, # "multilogitlink"
  coeff3 = exp(-0.5), # "loglink"
# "logofflink" link:
  coeff4 = logofflink(+0.5, offset = myoffset, inverse = TRUE),
  coeff5 = 0.50, # "multilogitlink"
  coeff6 = 1.00, # "identitylink"
  v2 = runif(nn),
  v3 = runif(nn),
  v4 = runif(nn),
  v5 = rnorm(nn),
  v6 = rnorm(nn))
ndata <- transform(ndata,
  Coeff1 = 0.25 - 0 * x2,
  Coeff2 = 0.25 - 0 * x2,
  Coeff3 = logitlink(-0.5 - 1 * x2, inverse = TRUE),
  Coeff4 = logloglink( 0.5 - 1 * x2, inverse = TRUE),
  Coeff5 = 0.50 - 0 * x2,
  Coeff6 = 1.00 + 1 * x2)
ndata <- transform(ndata,
  y1 = coeff1 * 1 +
    coeff2 * v2 +
    coeff3 * v3 +
    coeff4 * v4 +
    coeff5 * v5 +
    coeff6 * v6 + rnorm(nn, sd = exp(0)),
  y2 = Coeff1 * 1 +
    Coeff2 * v2 +
    Coeff3 * v3 +
    Coeff4 * v4 +
    Coeff5 * v5 +
    Coeff6 * v6 + rnorm(nn, sd = exp(0)))

# An intercept-only model
fit1 <- vglm(y1 ~ 1,
  form2 = ~ 1 + v2 + v3 + v4 + v5 + v6,
  normal.vcm(link.list = list("(Intercept)" = "multilogitlink",
    "v2" = "multilogitlink",
    "v3" = "loglink",
    "v4" = "logofflink",
    "(Default)" = "identitylink",
    "v5" = "multilogitlink"),
  earg.list = list("(Intercept)" = list(),

```

```

                                "v2"      = list(),
                                "v4"      = list(offset = myoffset),
                                "v3"      = list(),
                                "(Default)" = list(),
                                "v5"      = list()),
                                zero = c(1:2, 6)),
                                data = ndata, trace = TRUE)
coef(fit1, matrix = TRUE)
summary(fit1)
# This works only for intercept-only models:
multilogitlink(rbind(coef(fit1, matrix = TRUE)[1, c(1, 2)]), inverse = TRUE)

# A model with covariate x2 for the regression coefficients
fit2 <- vglm(y2 ~ 1 + x2,
             form2 = ~ 1 + v2 + v3 + v4 + v5 + v6,
             normal.vcm(link.list = list("(Intercept)" = "multilogitlink",
                                         "v2"          = "multilogitlink",
                                         "v3"          = "logitlink",
                                         "v4"          = "logloglink",
                                         "(Default)"   = "identitylink",
                                         "v5"          = "multilogitlink"),
             earg.list = list("(Intercept)" = list(),
                              "v2"         = list(),
                              "v3"         = list(),
                              "v4"         = list(),
                              "(Default)"  = list(),
                              "v5"         = list()),
             zero = c(1:2, 6)),
             data = ndata, trace = TRUE)

coef(fit2, matrix = TRUE)
summary(fit2)

```

nparam.vlm

Number of Parameters

Description

Returns the number of parameters in a fitted model object.

Usage

```

nparam(object, ...)
nparam.vlm(object, dpar = TRUE, ...)
nparam.vgam(object, dpar = TRUE, linear.only = FALSE, ...)
nparam.rrvglm(object, dpar = TRUE, ...)
nparam.drrvglm(object, dpar = TRUE, ...)
nparam.qrrvglm(object, dpar = TRUE, ...)
nparam.rrvgam(object, dpar = TRUE, ...)

```

Arguments

object	Some VGAM object, for example, having class <code>vglmff-class</code> .
...	Other possible arguments fed into the function.
dpar	Logical, include any (estimated) dispersion parameters as a parameter?
linear.only	Logical, include only the number of linear (parametric) parameters?

Details

The code was copied from the `AIC()` methods functions.

Value

Returns a numeric value with the corresponding number of parameters. For `vgam` objects, this may be real rather than integer, because the nonlinear degrees of freedom is real-valued.

Warning

This code has not been double-checked.

Author(s)

T. W. Yee.

See Also

VGLMs are described in `vglm-class`; VGAMs are described in `vgam-class`; RR-VGLMs are described in `rrvglm-class`; `AICvlm`.

Examples

```
pneumo <- transform(pneumo, let = log(exposure.time))
(fit1 <- vglm(cbind(normal, mild, severe) ~ let, propodds, data = pneumo))
coef(fit1)
coef(fit1, matrix = TRUE)
nparam(fit1)
(fit2 <- vglm(hits ~ 1, poissonff, weights = ofreq, data = V1))
coef(fit2)
coef(fit2, matrix = TRUE)
nparam(fit2)
nparam(fit2, dpar = FALSE)
```

olympics

2008 and 2012 Summer Olympic Final Medal Count Data

Description

Final medal count, by country, for the Summer 2008 and 2012 Olympic Games.

Usage

```
data(olymp08)
data(olymp12)
```

Format

A data frame with 87 or 85 observations on the following 6 variables.

rank a numeric vector, overall ranking of the countries.

country a factor.

gold a numeric vector, number of gold medals.

silver a numeric vector, number of silver medals.

bronze a numeric vector, number of bronze medals.

totalmedal a numeric vector, total number of medals.

Details

The events were held during (i) August 8–24, 2008, in Beijing; and (ii) 27 July–12 August, 2012, in London.

References

The official English website was/is <http://en.beijing2008.cn> and <http://www.london2012.com>. Help from Viet Hoang Quoc is gratefully acknowledged.

See Also

[grc](#).

Examples

```
summary(olymp08)
summary(olymp12)
## maybe str(olymp08) ; plot(olymp08) ...
## Not run: par(mfrow = c(1, 2))
myylim <- c(0, 55)
with(head(olymp08, n = 8),
barplot(rbind(gold, silver, bronze),
col = c("gold", "grey", "brown"), # No "silver" or "bronze"!
```

```

#       "gold", "grey71", "chocolate4",
names.arg = country, cex.names = 0.5, ylim = myylim,
beside = TRUE, main = "2008 Summer Olympic Final Medal Count",
ylab = "Medal count", las = 1,
sub = "Top 8 countries; 'gold'=gold, 'grey'=silver, 'brown'=bronze"))
with(head(olymp12, n = 8),
barplot(rbind(gold, silver, bronze),
col = c("gold", "grey", "brown"), # No "silver" or "bronze"!
names.arg = country, cex.names = 0.5, ylim = myylim,
beside = TRUE, main = "2012 Summer Olympic Final Medal Count",
ylab = "Medal count", las = 1,
sub = "Top 8 countries; 'gold'=gold, 'grey'=silver, 'brown'=bronze"))
## End(Not run)

```

Opt

Optimums

Description

Generic function for the *optimums* (or optima) of a model.

Usage

```
Opt(object, ...)
```

Arguments

object	An object for which the computation or extraction of an optimum (or optimums) is meaningful.
...	Other arguments fed into the specific methods function of the model. Sometimes they are fed into the methods function for Coef .

Details

Different models can define an optimum in different ways. Many models have no such notion or definition.

Optimums occur in quadratic and additive ordination, e.g., CQO or CAO. For these models the optimum is the value of the latent variable where the maximum occurs, i.e., where the fitted value achieves its highest value. For quadratic ordination models there is a formula for the optimum but for additive ordination models the optimum must be searched for numerically. If it occurs on the boundary, then the optimum is undefined. At an optimum, the fitted value of the response is called the *maximum*.

Value

The value returned depends specifically on the methods function invoked.

Note

In ordination, the optimum of a species is sometimes called the *species score*.

Author(s)

Thomas W. Yee

References

Yee, T. W. (2004). A new technique for maximum-likelihood canonical Gaussian ordination. *Ecological Monographs*, **74**, 685–701.

Yee, T. W. (2006). Constrained additive ordination. *Ecology*, **87**, 203–213.

See Also

Opt.qrrvglm, Max, Tol.

Examples

```
## Not run:
set.seed(111) # This leads to the global solution
hspider[,1:6] <- scale(hspider[,1:6]) # Standardized environmental vars
p1 <- cqa(cbind(Alopacce, Alopcone, Alopfabr, Arctlute, Arctperi,
              Auloalbi, Pardlugu, Pardmont, Pardnigr, Pardpull,
              Trocterr, Zoraspin) ~
          WaterCon + BareSand + FallTwig + CoveMoss + CoveHerb + ReflLux,
          family = poissonff, data = hspider, Crowlpositive = FALSE)
Opt(p1)

clr <- (1:(ncol(depvar(p1))+1))[-7] # Omits yellow
persp(p1, col = clr, las = 1, main = "Vertical lines at the optimums")
abline(v = Opt(p1), lty = 2, col = clr)

## End(Not run)
```

ordpoisson

Ordinal Poisson Family Function

Description

Fits a Poisson regression where the response is ordinal (the Poisson counts are grouped between known cutpoints).

Usage

```
ordpoisson(cutpoints, countdata = FALSE, NOS = NULL,
           Levels = NULL, init.mu = NULL, parallel = FALSE,
           zero = NULL, link = "loglink")
```

Arguments

cutpoints	Numeric. The cutpoints, K_l . These must be non-negative integers. Inf values may be included. See below for further details.
countdata	Logical. Is the response (LHS of formula) in count-data format? If not then the response is a matrix or vector with values $1, 2, \dots, L$, say, where L is the number of levels. Such input can be generated with <code>cut</code> with argument <code>labels = FALSE</code> . If <code>countdata = TRUE</code> then the response is expected to be in the same format as <code>devar(fit)</code> where <code>fit</code> is a fitted model with <code>ordpoisson</code> as the VGAM family function. That is, the response is matrix of counts with L columns (if <code>NOS = 1</code>).
NOS	Integer. The number of species, or more generally, the number of response random variates. This argument must be specified when <code>countdata = TRUE</code> . Usually <code>NOS = 1</code> .
Levels	Integer vector, recycled to length <code>NOS</code> if necessary. The number of levels for each response random variate. This argument should agree with <code>cutpoints</code> . This argument must be specified when <code>countdata = TRUE</code> .
init.mu	Numeric. Initial values for the means of the Poisson regressions. Recycled to length <code>NOS</code> if necessary. Use this argument if the default initial values fail (the default is to compute an initial value internally).
parallel, zero, link	See <code>poissonff</code> . See CommonVGAMffArguments for information.

Details

This **VGAM** family function uses maximum likelihood estimation (Fisher scoring) to fit a Poisson regression to each column of a matrix response. The data, however, is ordinal, and is obtained from known integer cutpoints. Here, $l = 1, \dots, L$ where L ($L \geq 2$) is the number of levels. In more detail, let $Y^* = l$ if $K_{l-1} < Y \leq K_l$ where the K_l are the cutpoints. We have $K_0 = -\infty$ and $K_L = \infty$. The response for this family function corresponds to Y^* but we are really interested in the Poisson regression of Y .

If `NOS=1` then the argument `cutpoints` is a vector (K_1, K_2, \dots, K_L) where the last value (`Inf`) is optional. If `NOS>1` then the vector should have `NOS-1` `Inf` values separating the cutpoints. For example, if there are `NOS=3` responses, then something like `ordpoisson(cut = c(0, 5, 10, Inf, 20, 30, Inf, 0, 10, 40, Inf))` is valid.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as `vglm` and `vgam`.

Warning

The input requires care as little to no checking is done. If `fit` is the fitted object, have a look at `fit@extra` and `devar(fit)` to check.

Note

Sometimes there are no observations between two cutpoints. If so, the arguments `Levels` and `NOS` need to be specified too. See below for an example.

Author(s)

Thomas W. Yee

References

Yee, T. W. (2020). *Ordinal ordination with normalizing link functions for count data*, (in preparation).

See Also

[poissonff](#), [polf](#), [ordered](#).

Examples

```
set.seed(123) # Example 1
x2 <- runif(n <- 1000); x3 <- runif(n)
mymu <- exp(3 - 1 * x2 + 2 * x3)
y1 <- rpois(n, lambda = mymu)
cutpts <- c(-Inf, 20, 30, Inf)
fcutpts <- cutpts[is.finite(cutpts)] # finite cutpoints
ystar <- cut(y1, breaks = cutpts, labels = FALSE)
## Not run:
plot(x2, x3, col = ystar, pch = as.character(ystar))

## End(Not run)
table(ystar) / sum(table(ystar))
fit <- vglm(ystar ~ x2 + x3, fam = ordpoisson(cutpoi = fcutpts))
head(depvar(fit)) # This can be input if countdata = TRUE
head(fitted(fit))
head(predict(fit))
coef(fit, matrix = TRUE)
fit@extra

# Example 2: multivariate and there are no obsns between some cutpoints
cutpts2 <- c(-Inf, 0, 9, 10, 20, 70, 200, 201, Inf)
fcutpts2 <- cutpts2[is.finite(cutpts2)] # finite cutpoints
y2 <- rpois(n, lambda = mymu) # Same model as y1
ystar2 <- cut(y2, breaks = cutpts2, labels = FALSE)
table(ystar2) / sum(table(ystar2))
fit <- vglm(cbind(ystar, ystar2) ~ x2 + x3, fam =
  ordpoisson(cutpoi = c(fcutpts, Inf, fcutpts2, Inf),
    Levels = c(length(fcutpts)+1, length(fcutpts2)+1),
    parallel = TRUE), trace = TRUE)
coef(fit, matrix = TRUE)
fit@extra
constraints(fit)
summary(depvar(fit)) # Some columns have all zeros
```

ordsup

*Ordinal Superiority Measures***Description**

Ordinal superiority measures for the linear model and cumulative link models: the probability that an observation from one distribution falls above an independent observation from the other distribution, adjusted for explanatory variables in a model.

Usage

```
ordsup(object, ...)
ordsup.vglm(object, all.vars = FALSE, confint = FALSE, ...)
```

Arguments

object	A <code>vglm</code> fit. Currently it must be one of: <code>cumulative</code> , <code>uninormal</code> . The links for <code>cumulative</code> must be <code>logitlink</code> or <code>probitlink</code> , and <code>parallel = TRUE</code> is also needed. For <code>uninormal</code> the mean must use <code>identitylink</code> and model the sd as intercept-only.
all.vars	Logical. The default is to use explanatory variables which are binary, but all variables are used (except the intercept) if set to <code>TRUE</code> .
confint	Logical. If <code>TRUE</code> then <code>confintvglm</code> is called to return confidence intervals for γ and Δ . By default, Wald intervals are produced, but they can be replaced by profile intervals by setting <code>method = "profile"</code> .
...	Parameters that can be fed into <code>confintvglm</code> , e.g., <code>level = 0.95</code> and <code>method = c("wald", "profile")</code> .

Details

Details are given in Agresti and Kateri (2017) and this help file draws directly from this. This function returns two quantities for comparing two groups on an ordinal categorical response variable, while adjusting for other explanatory variables. They are called “ordinal superiority” measures, and the two groups can be compared without supplementary explanatory variables. Let Y_1 and Y_2 be independent random variables from groups A and B, say, for a quantitative ordinal categorical scale. Then $\Delta = P(Y_1 > Y_2) - P(Y_2 > Y_1)$ summarizes their relative size. A second quantity is $\gamma = P(Y_1 > Y_2) - 0.5 \times P(Y_2 = Y_1)$. Then $\Delta = 2 \times \gamma - 1$, whereas $\gamma = (\Delta + 1)/2$. The range of γ is $[0, 1]$, while the range of Δ is $[-1, 1]$. The examples below are based on that paper. This function is currently implemented for a very limited number of specific models.

Value

By default, a list with components `gamma` and `Delta`, where each is a vector with elements corresponding to binary explanatory variables (i.e., 0 or 1), and if no explanatory variables are binary then a `NULL` is returned. If `confint = TRUE` then the list contains 4 more components: `lower.gamma`, `upper.gamma`, `Lower.Delta`, `Upper.Delta`.

Author(s)

Thomas W. Yee

References

Agresti, A. and Kateri, M. (2017). Ordinal probability effect measures for group comparisons in multinomial cumulative link models. *Biometrics*, **73**, 214–219.

See Also

[cumulative](#), [propodds](#), [uninormal](#).

Examples

```
## Not run:
Mental <- read.table("http://www.stat.ufl.edu/~aa/glm/data/Mental.dat",
                    header = TRUE) # Make take a while to load in
Mental$impair <- ordered(Mental$impair)
pfit3 <- vglm(impair ~ ses + life, data = Mental,
             cumulative(link = "probitlink", reverse = FALSE, parallel = TRUE))
coef(pfit3, matrix = TRUE)
ordsup(pfit3) # The 'ses' variable is binary

# Fit a crude LM
fit7 <- vglm(as.numeric(impair) ~ ses + life, uninormal, data = Mental)
coef(fit7, matrix = TRUE) # 'sd' is estimated by MLE
ordsup(fit7)
ordsup(fit7, all.vars = TRUE) # Some output may not be meaningful
ordsup(fit7, confint = TRUE, method = "profile")

## End(Not run)
```

 oxtemp

Oxford Temperature Data

Description

Annual maximum temperatures collected at Oxford, UK.

Usage

```
data(oxtemp)
```

Format

A data frame with 80 observations on the following 2 variables.

maxtemp Annual maximum temperatures (in degrees Fahrenheit).

year The values 1901 to 1980.

Details

The data were collected from 1901 to 1980.

Source

Unknown.

Examples

```
## Not run: fit <- vglm(maxtemp ~ 1, gevff, data = oxtemp, trace = TRUE)
```

Paralogistic

The Paralogistic Distribution

Description

Density, distribution function, quantile function and random generation for the paralogistic distribution with shape parameter a and scale parameter $scale$.

Usage

```
dparalogistic(x, scale = 1, shape1.a, log = FALSE)
pparalogistic(q, scale = 1, shape1.a, lower.tail = TRUE, log.p = FALSE)
qparalogistic(p, scale = 1, shape1.a, lower.tail = TRUE, log.p = FALSE)
rparalogistic(n, scale = 1, shape1.a)
```

Arguments

<code>x, q</code>	vector of quantiles.
<code>p</code>	vector of probabilities.
<code>n</code>	number of observations. If <code>length(n) > 1</code> , the length is taken to be the number required.
<code>shape1.a</code>	shape parameter.
<code>scale</code>	scale parameter.
<code>log</code>	Logical. If <code>log=TRUE</code> then the logarithm of the density is returned.
<code>lower.tail, log.p</code>	Same meaning as in pnorm or qnorm .

Details

See [paralogistic](#), which is the **VGAM** family function for estimating the parameters by maximum likelihood estimation.

Value

`dparalogistic` gives the density, `pparalogistic` gives the distribution function, `qparalogistic` gives the quantile function, and `rparalogistic` generates random deviates.

Note

The paralogistic distribution is a special case of the 4-parameter generalized beta II distribution.

Author(s)

T. W. Yee and Kai Huang

References

Kleiber, C. and Kotz, S. (2003). *Statistical Size Distributions in Economics and Actuarial Sciences*, Hoboken, NJ, USA: Wiley-Interscience.

See Also

[paralogistic](#), [genbetaII](#).

Examples

```
pdata <- data.frame(y = rparalogistic(n = 3000, scale = exp(1), exp(2)))
fit <- vglm(y ~ 1, paralogistic(lss = FALSE, ishape1.a = 4.1),
           data = pdata, trace = TRUE)
coef(fit, matrix = TRUE)
Coef(fit)
```

paralogistic

Paralogistic Distribution Family Function

Description

Maximum likelihood estimation of the 2-parameter paralogistic distribution.

Usage

```
paralogistic(lscale = "loglink", lshape1.a = "loglink", iscale = NULL,
            ishape1.a = NULL, imethod = 1, lss = TRUE, gscale = exp(-5:5),
            gshape1.a = seq(0.75, 4, by = 0.25), probs.y = c(0.25, 0.5, 0.75),
            zero = "shape")
```

Arguments

`lss` See [CommonVGAMffArguments](#) for important information.

`lshape1.a, lscale` Parameter link functions applied to the (positive) parameters a and scale. See [Links](#) for more choices.

`iscale, ishape1.a, imethod, zero` See [CommonVGAMffArguments](#) for information. For `imethod = 2` a good initial value for `ishape1.a` is needed to obtain good estimates for the other parameter.

gscale, gshape1.a

See [CommonVGAMffArguments](#) for information.

probs.y

See [CommonVGAMffArguments](#) for information.

Details

The 2-parameter paralogistic distribution is the 4-parameter generalized beta II distribution with shape parameter $p = 1$ and $a = q$. It is the 3-parameter Singh-Maddala distribution with $a = q$. More details can be found in Kleiber and Kotz (2003).

The 2-parameter paralogistic has density

$$f(y) = a^2 y^{a-1} / [b^a \{1 + (y/b)^a\}^{1+a}]$$

for $a > 0$, $b > 0$, $y \geq 0$. Here, b is the scale parameter scale, and a is the shape parameter. The mean is

$$E(Y) = b \Gamma(1 + 1/a) \Gamma(a - 1/a) / \Gamma(a)$$

provided $a > 1$; these are returned as the fitted values. This family function handles multiple responses.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Note

See the notes in [genbetaII](#).

Author(s)

T. W. Yee

References

Kleiber, C. and Kotz, S. (2003). *Statistical Size Distributions in Economics and Actuarial Sciences*, Hoboken, NJ, USA: Wiley-Interscience.

See Also

[Paralogistic](#), [sinmad](#), [genbetaII](#), [betaII](#), [dagum](#), [fisk](#), [inv.lomax](#), [lomax](#), [inv.paralogistic](#).

Examples

```
## Not run:
pdata <- data.frame(y = rparalogistic(n = 3000, exp(1), scale = exp(1)))
fit <- vglm(y ~ 1, paralogistic(lss = FALSE), data = pdata, trace = TRUE)
fit <- vglm(y ~ 1, paralogistic(ishape1.a = 2.3, iscale = 5),
           data = pdata, trace = TRUE)
coef(fit, matrix = TRUE)
Coef(fit)
```

```
summary(fit)
## End(Not run)
```

Pareto

The Pareto Distribution

Description

Density, distribution function, quantile function and random generation for the Pareto(I) distribution with parameters scale and shape.

Usage

```
dpareto(x, scale = 1, shape, log = FALSE)
ppareto(q, scale = 1, shape, lower.tail = TRUE, log.p = FALSE)
qpareto(p, scale = 1, shape, lower.tail = TRUE, log.p = FALSE)
rpareto(n, scale = 1, shape)
```

Arguments

x, q	vector of quantiles.
p	vector of probabilities.
n	number of observations. Same as in runif .
scale, shape	the α and k parameters.
log	Logical. If log = TRUE then the logarithm of the density is returned.
lower.tail, log.p	Same meaning as in pnorm or qnorm .

Details

See [paretoff](#), the **VGAM** family function for estimating the parameter k by maximum likelihood estimation, for the formula of the probability density function and the range restrictions imposed on the parameters.

Value

dpareto gives the density, ppareto gives the distribution function, qpareto gives the quantile function, and rpareto generates random deviates.

Author(s)

T. W. Yee and Kai Huang

References

Forbes, C., Evans, M., Hastings, N. and Peacock, B. (2011). *Statistical Distributions*, Hoboken, NJ, USA: John Wiley and Sons, Fourth edition.

See Also

[paretoff](#), [ParetoIV](#).

Examples

```
alpha <- 3; k <- exp(1); x <- seq(2.8, 8, len = 300)
## Not run:
plot(x, dpareto(x, scale = alpha, shape = k), type = "l",
     main = "Pareto density split into 10 equal areas")
abline(h = 0, col = "blue", lty = 2)
qvec <- qpareto(seq(0.1, 0.9, by = 0.1), scale = alpha, shape = k)
lines(qvec, dpareto(qvec, scale = alpha, shape = k),
      col = "purple", lty = 3, type = "h")

## End(Not run)
pvec <- seq(0.1, 0.9, by = 0.1)
qvec <- qpareto(pvec, scale = alpha, shape = k)
ppareto(qvec, scale = alpha, shape = k)
qpareto(ppareto(qvec, scale = alpha, shape = k),
        scale = alpha, shape = k) - qvec # Should be 0
```

 paretoff

Pareto and Truncated Pareto Distribution Family Functions

Description

Estimates one of the parameters of the Pareto(I) distribution by maximum likelihood estimation. Also includes the upper truncated Pareto(I) distribution.

Usage

```
paretoff(scale = NULL, lshape = "loglink")
truncpareto(lower, upper, lshape = "loglink", ishape = NULL, imethod = 1)
```

Arguments

lshape	Parameter link function applied to the parameter k . See Links for more choices. A log link is the default because k is positive.
scale	Numeric. The parameter α below. If the user inputs a number then it is assumed known with this value. The default means it is estimated by maximum likelihood estimation, which means $\min(y)$ is used, where y is the response vector.
lower, upper	Numeric. Lower and upper limits for the truncated Pareto distribution. Each must be positive and of length 1. They are called α and U below.
ishape	Numeric. Optional initial value for the shape parameter. A NULL means a value is obtained internally. If failure to converge occurs try specifying a value, e.g., 1 or 2.
imethod	See CommonVGAMffArguments for information. If failure to converge occurs then try specifying a value for ishape.

Details

A random variable Y has a Pareto distribution if

$$P[Y > y] = C/y^k$$

for some positive k and C . This model is important in many applications due to the power law probability tail, especially for large values of y .

The Pareto distribution, which is used a lot in economics, has a probability density function that can be written

$$f(y; \alpha, k) = k\alpha^k/y^{k+1}$$

for $0 < \alpha < y$ and $0 < k$. The α is called the *scale* parameter, and it is either assumed *known* or else $\min(y)$ is used. The parameter k is called the *shape* parameter. The mean of Y is $\alpha k/(k - 1)$ provided $k > 1$. Its variance is $\alpha^2 k/((k - 1)^2(k - 2))$ provided $k > 2$.

The upper truncated Pareto distribution has a probability density function that can be written

$$f(y) = k\alpha^k/[y^{k+1}(1 - (\alpha/U)^k)]$$

for $0 < \alpha < y < U < \infty$ and $k > 0$. Possibly, better names for k are the *index* and *tail* parameters. Here, α and U are known. The mean of Y is $k\alpha^k(U^{1-k} - \alpha^{1-k})/[(1 - k)(1 - (\alpha/U)^k)]$.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Warning

The usual or unbounded Pareto distribution has two parameters (called α and k here) but the family function `paretoff` estimates only k using iteratively reweighted least squares. The MLE of the α parameter lies on the boundary and is $\min(y)$ where y is the response. Consequently, using the default argument values, the standard errors are incorrect when one does a summary on the fitted object. If the user inputs a value for `alpha` then it is assumed known with this value and then summary on the fitted object should be correct. Numerical problems may occur for small k , e.g., $k < 1$.

Note

Outside of economics, the Pareto distribution is known as the Bradford distribution.

For `paretoff`, if the estimate of k is less than or equal to unity then the fitted values will be NAs. Also, `paretoff` fits the Pareto(I) distribution. See [paretoIV](#) for the more general Pareto(IV/III/II) distributions, but there is a slight change in notation: $s = k$ and $b = \alpha$.

In some applications the Pareto law is truncated by a natural upper bound on the probability tail. The upper truncated Pareto distribution has three parameters (called α , U and k here) but the family function `truncpareto()` estimates only k . With known lower and upper limits, the ML estimator of k has the usual properties of MLEs. Aban (2006) discusses other inferential details.

Author(s)

T. W. Yee

References

Forbes, C., Evans, M., Hastings, N. and Peacock, B. (2011). *Statistical Distributions*, Hoboken, NJ, USA: John Wiley and Sons, Fourth edition.

Aban, I. B., Meerschaert, M. M. and Panorska, A. K. (2006). Parameter estimation for the truncated Pareto distribution, *Journal of the American Statistical Association*, **101**(473), 270–277.

See Also

[Pareto](#), [Truncpareto](#), [paretoIV](#), [gpd](#), [benini1](#).

Examples

```
alpha <- 2; kay <- exp(3)
pdata <- data.frame(y = rpareto(n = 1000, scale = alpha, shape = kay))
fit <- vglm(y ~ 1, paretoff, data = pdata, trace = TRUE)
fit@extra # The estimate of alpha is here
head(fitted(fit))
with(pdata, mean(y))
coef(fit, matrix = TRUE)
summary(fit) # Standard errors are incorrect!!

# Here, alpha is assumed known
fit2 <- vglm(y ~ 1, paretoff(scale = alpha), data = pdata, trace = TRUE)
fit2@extra # alpha stored here
head(fitted(fit2))
coef(fit2, matrix = TRUE)
summary(fit2) # Standard errors are okay

# Upper truncated Pareto distribution
lower <- 2; upper <- 8; kay <- exp(2)
pdata3 <- data.frame(y = rtruncpareto(n = 100, lower = lower,
                                     upper = upper, shape = kay))
fit3 <- vglm(y ~ 1, truncpareto(lower, upper), data = pdata3, trace = TRUE)
coef(fit3, matrix = TRUE)
c(fit3@misc$lower, fit3@misc$upper)
```

ParetoIV

The Pareto(IV/III/II) Distributions

Description

Density, distribution function, quantile function and random generation for the Pareto(IV/III/II) distributions.

Usage

```

dparetoIV(x, location = 0, scale = 1, inequality = 1, shape = 1,
          log = FALSE)
pparetoIV(q, location = 0, scale = 1, inequality = 1, shape = 1,
          lower.tail = TRUE, log.p = FALSE)
qparetoIV(p, location = 0, scale = 1, inequality = 1, shape = 1,
          lower.tail = TRUE, log.p = FALSE)
rparetoIV(n, location = 0, scale = 1, inequality = 1, shape = 1)
dparetoIII(x, location = 0, scale = 1, inequality = 1, log = FALSE)
pparetoIII(q, location = 0, scale = 1, inequality = 1,
          lower.tail = TRUE, log.p = FALSE)
qparetoIII(p, location = 0, scale = 1, inequality = 1,
          lower.tail = TRUE, log.p = FALSE)
rparetoIII(n, location = 0, scale = 1, inequality = 1)
dparetoII(x, location = 0, scale = 1, shape = 1, log = FALSE)
pparetoII(q, location = 0, scale = 1, shape = 1,
          lower.tail = TRUE, log.p = FALSE)
qparetoII(p, location = 0, scale = 1, shape = 1,
          lower.tail = TRUE, log.p = FALSE)
rparetoII(n, location = 0, scale = 1, shape = 1)
dparetoI(x, scale = 1, shape = 1, log = FALSE)
pparetoI(q, scale = 1, shape = 1,
          lower.tail = TRUE, log.p = FALSE)
qparetoI(p, scale = 1, shape = 1,
          lower.tail = TRUE, log.p = FALSE)
rparetoI(n, scale = 1, shape = 1)

```

Arguments

x, q	vector of quantiles.
p	vector of probabilities.
n	number of observations. Same as in runif .
location	the location parameter.
scale, shape, inequality	the (positive) scale, inequality and shape parameters.
log	Logical. If log = TRUE then the logarithm of the density is returned.
lower.tail, log.p	Same meaning as in pnorm or qnorm .

Details

For the formulas and other details see [paretoIV](#).

Value

Functions beginning with the letters d give the density, p give the distribution function, q give the quantile function, and r generates random deviates.

Note

The functions [dpqr]paretoI are the same as [dpqr]pareto except for a slight change in notation: $s = k$ and $b = \alpha$; see [Pareto](#).

Author(s)

T. W. Yee and Kai Huang

References

Brazauskas, V. (2003). Information matrix for Pareto(IV), Burr, and related distributions. *Comm. Statist. Theory and Methods* **32**, 315–325.

Arnold, B. C. (1983). *Pareto Distributions*. Fairland, Maryland: International Cooperative Publishing House.

See Also

[paretoIV](#), [Pareto](#).

Examples

```
## Not run:
x <- seq(-0.2, 4, by = 0.01)
loc <- 0; Scale <- 1; ineq <- 1; shape <- 1.0
plot(x, dparetoIV(x, loc, Scale, ineq, shape), type = "l",
     main = "Blue is density, orange is the CDF", col = "blue",
     sub = "Purple are 5,10,...,95 percentiles", ylim = 0:1,
     las = 1, ylab = "")
abline(h = 0, col = "blue", lty = 2)
Q <- qparetoIV(seq(0.05, 0.95, by = 0.05), loc, Scale, ineq, shape)
lines(Q, dparetoIV(Q, loc, Scale, ineq, shape), col = "purple",
     lty = 3, type = "h")
lines(x, pparetoIV(x, loc, Scale, ineq, shape), col = "orange")
abline(h = 0, lty = 2)

## End(Not run)
```

paretoIV

Pareto(IV/III/II) Distribution Family Functions

Description

Estimates three of the parameters of the Pareto(IV) distribution by maximum likelihood estimation. Some special cases of this distribution are also handled.

Usage

```

paretoIV(location = 0, lscale = "loglink", linequality = "loglink",
         lshape = "loglink", iscale = 1, iinequality = 1, ishape = NULL,
         imethod = 1)
paretoIII(location = 0, lscale = "loglink", linequality = "loglink",
          iscale = NULL, iinequality = NULL)
paretoII(location = 0, lscale = "loglink", lshape = "loglink",
          iscale = NULL, ishape = NULL)

```

Arguments

location	Location parameter, called a below. It is assumed known.
lscale, linequality, lshape	Parameter link functions for the scale parameter (called b below), inequality parameter (called g below), and shape parameter (called s below). See Links for more choices. A log link is the default for all because all these parameters are positive.
iscale, iinequality, ishape	Initial values for the parameters. A NULL value means that it is obtained internally. If convergence failure occurs, use these arguments to input some alternative initial values.
imethod	Method of initialization for the shape parameter. Currently only values 1 and 2 are available. Try the other value if convergence failure occurs.

Details

The Pareto(IV) distribution, which is used in actuarial science, economics, finance and telecommunications, has a cumulative distribution function that can be written

$$F(y) = 1 - [1 + ((y - a)/b)^{1/g}]^{-s}$$

for $y > a$, $b > 0$, $g > 0$ and $s > 0$. The a is called the *location* parameter, b the *scale* parameter, g the *inequality* parameter, and s the *shape* parameter.

The location parameter is assumed known otherwise the Pareto(IV) distribution will not be a regular family. This assumption is not too restrictive in modelling because in typical applications this parameter is known, e.g., in insurance and reinsurance it is pre-defined by a contract and can be represented as a deductible or a retention level.

The inequality parameter is so-called because of its interpretation in the economics context. If we choose a unit shape parameter value and a zero location parameter value then the inequality parameter is the Gini index of inequality, provided $g \leq 1$.

The fitted values are currently the median, e.g., `qparetoIV` is used for `paretoIV()`.

There are a number of special cases of the Pareto(IV) distribution. These include the Pareto(I), Pareto(II), Pareto(III), and Burr family of distributions. Denoting $PIV(a, b, g, s)$ as the Pareto(IV) distribution, the Burr distribution $Burr(b, g, s)$ is $PIV(a = 0, b, 1/g, s)$, the Pareto(III) distribution $PIII(a, b, g)$ is $PIV(a, b, g, s = 1)$, the Pareto(II) distribution $PII(a, b, s)$ is $PIV(a, b, g = 1, s)$, and the Pareto(I) distribution $PI(b, s)$ is $PIV(b, b, g = 1, s)$. Thus the Burr distribution can be fitted using the [negloglink](#) link function and using the default `location=0` argument. The

Pareto(I) distribution can be fitted using [paretoff](#) but there is a slight change in notation: $s = k$ and $b = \alpha$.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Warning

The Pareto(IV) distribution is very general, for example, special cases include the Pareto(I), Pareto(II), Pareto(III), and Burr family of distributions. [Johnson et al. (1994) says on p.19 that fitting Type IV by ML is very difficult and rarely attempted]. Consequently, reasonably good initial values are recommended, and convergence to a local solution may occur. For this reason setting `trace=TRUE` is a good idea for monitoring the convergence. Large samples are ideally required to get reasonable results.

Note

The extra slot of the fitted object has a component called "location" which stores the location parameter value(s).

Author(s)

T. W. Yee

References

- Johnson N. L., Kotz S., and Balakrishnan N. (1994). *Continuous Univariate Distributions, Volume 1*, 2nd ed. New York: Wiley.
- Brazauskas, V. (2003). Information matrix for Pareto(IV), Burr, and related distributions. *Comm. Statist. Theory and Methods* **32**, 315–325.
- Arnold, B. C. (1983). *Pareto Distributions*. Fairland, Maryland: International Cooperative Publishing House.

See Also

[ParetoIV](#), [paretoff](#), [gpd](#).

Examples

```
pdata <- data.frame(y = rparetoIV(2000, scale = exp(1),
                              ineq = exp(-0.3), shape = exp(1)))
## Not run: par(mfrow = c(2, 1))
with(pdata, hist(y)); with(pdata, hist(log(y)))
## End(Not run)
fit <- vglm(y ~ 1, paretoIV, data = pdata, trace = TRUE)
head(fitted(fit))
summary(pdata)
coef(fit, matrix = TRUE)
```

```
Coef(fit)
summary(fit)
```

Perks

The Perks Distribution

Description

Density, cumulative distribution function, quantile function and random generation for the Perks distribution.

Usage

```
dperks(x, scale = 1, shape, log = FALSE)
pperks(q, scale = 1, shape, lower.tail = TRUE, log.p = FALSE)
qperks(p, scale = 1, shape, lower.tail = TRUE, log.p = FALSE)
rperks(n, scale = 1, shape)
```

Arguments

<code>x, q</code>	vector of quantiles.
<code>p</code>	vector of probabilities.
<code>n</code>	number of observations. Same as in runif .
<code>log</code>	Logical. If <code>log = TRUE</code> then the logarithm of the density is returned.
<code>lower.tail, log.p</code>	Same meaning as in pnorm or qnorm .
<code>shape, scale</code>	positive shape and scale parameters.

Details

See [perks](#) for details.

Value

`dperks` gives the density, `pperks` gives the cumulative distribution function, `qperks` gives the quantile function, and `rperks` generates random deviates.

Author(s)

T. W. Yee and Kai Huang

See Also

[perks](#).

Examples

```

probs <- seq(0.01, 0.99, by = 0.01)
Shape <- exp(-1.0); Scale <- exp(1);
max(abs(pperks(qperks(p = probs, Shape, Scale),
              Shape, Scale) - probs)) # Should be 0

## Not run: x <- seq(-0.1, 0.7, by = 0.01);
plot(x, dperks(x, Shape, Scale), type = "l", col = "blue", las = 1,
     main = "Blue is density, orange is cumulative distribution function",
     sub = "Purple lines are the 10,20,...,90 percentiles",
     ylab = "", ylim = 0:1)
abline(h = 0, col = "blue", lty = 2)
lines(x, pperks(x, Shape, Scale), col = "orange")
probs <- seq(0.1, 0.9, by = 0.1)
Q <- qperks(probs, Shape, Scale)
lines(Q, dperks(Q, Shape, Scale), col = "purple", lty = 3, type = "h")
pperks(Q, Shape, Scale) - probs # Should be all zero
abline(h = probs, col = "purple", lty = 3)
## End(Not run)

```

perks

*Perks Distribution Family Function***Description**

Maximum likelihood estimation of the 2-parameter Perks distribution.

Usage

```

perks(lscale = "loglink", lshape = "loglink",
     iscale = NULL,  ishape = NULL,
     gscale = exp(-5:5), gshape = exp(-5:5),
     nsimEIM = 500, oim.mean = FALSE, zero = NULL,
     nowarning = FALSE)

```

Arguments

nowarning	Logical. Suppress a warning? Ignored for VGAM 0.9-7 and higher.
lscale, lshape	Parameter link functions applied to the shape parameter shape, scale parameter scale. All parameters are treated as positive here See Links for more choices.
iscale, ishape	Optional initial values. A NULL means a value is computed internally.
gscale, gshape	See CommonVGAMffArguments .
nsimEIM, zero	See CommonVGAMffArguments .
oim.mean	To be currently ignored.

Details

The Perks distribution has cumulative distribution function

$$F(y; \alpha, \beta) = 1 - \left\{ \frac{1 + \alpha}{1 + \alpha e^{\beta y}} \right\}^{1/\beta}$$

which leads to a probability density function

$$f(y; \alpha, \beta) = [1 + \alpha]^{1/\beta} \alpha e^{\beta y} / (1 + \alpha e^{\beta y})^{1+1/\beta}$$

for $\alpha > 0$, $\beta > 0$, $y > 0$. Here, β is called the scale parameter *scale*, and α is called a shape parameter. The moments for this distribution do not appear to be available in closed form.

Simulated Fisher scoring is used and multiple responses are handled.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Warning

A lot of care is needed because this is a rather difficult distribution for parameter estimation. If the self-starting initial values fail then try experimenting with the initial value arguments, especially *iscale*. Successful convergence depends on having very good initial values. Also, monitor convergence by setting *trace* = TRUE.

Author(s)

T. W. Yee

References

- Perks, W. (1932). On some experiments in the graduation of mortality statistics. *Journal of the Institute of Actuaries*, **63**, 12–40.
- Richards, S. J. (2012). A handbook of parametric survival models for actuarial use. *Scandinavian Actuarial Journal*. 1–25.

See Also

[dperks](#), [simulate.vlm](#).

Examples

```
## Not run: set.seed(123)
pdata <- data.frame(x2 = runif(nn <- 1000)) # x2 unused
pdata <- transform(pdata, eta1 = -1,
                  ceta1 = 1)
pdata <- transform(pdata, shape1 = exp(eta1),
                  scale1 = exp(ceta1))
pdata <- transform(pdata, y1 = rperks(nn, sh = shape1, sc = scale1))
```

```
fit1 <- vglm(y1 ~ 1, perks, data = pdata, trace = TRUE)
coef(fit1, matrix = TRUE)
summary(fit1)

## End(Not run)
```

perspqrvglm

Perspective plot for QRR-VGLMs

Description

Produces a perspective plot for a CQO model (QRR-VGLM). It is only applicable for rank-1 or rank-2 models with argument `noRRR = ~ 1`.

Usage

```
perspqrvglm(x, varI.latvar = FALSE, refResponse = NULL, show.plot = TRUE,
            xlim = NULL, ylim = NULL, zlim = NULL,
            gridlength = if (Rank == 1) 301 else c(51,51),
            which.species = NULL,
            xlab = if (Rank == 1) "Latent Variable" else "Latent Variable 1",
            ylab = if (Rank == 1) "Expected Value" else "Latent Variable 2",
            zlab = "Expected value", labelSpecies = FALSE,
            stretch = 1.05, main = "", ticktype = "detailed",
            col = if (Rank == 1) par()$col else "white",
            llty = par()$lty, llwd = par()$lwd,
            add1 = FALSE, ...)
```

Arguments

<code>x</code>	Object of class "qrrvglm", i.e., a constrained quadratic ordination (CQO) object.
<code>varI.latvar</code>	Logical that is fed into <code>Coef.qrrvglm</code> .
<code>refResponse</code>	Integer or character that is fed into <code>Coef.qrrvglm</code> .
<code>show.plot</code>	Logical. Plot it?
<code>xlim, ylim</code>	Limits of the x- and y-axis. Both are numeric of length 2. See <code>par</code> .
<code>zlim</code>	Limits of the z-axis. Numeric of length 2. Ignored if rank is 1. See <code>par</code> .
<code>gridlength</code>	Numeric. The fitted values are evaluated on a grid, and this argument regulates the fineness of the grid. If Rank = 2 then the argument is recycled to length 2, and the two numbers are the number of grid points on the x- and y-axes respectively.
<code>which.species</code>	Numeric or character vector. Indicates which species are to be plotted. The default is to plot all of them. If numeric, it should contain values in the set $\{1, 2, \dots, S\}$ where S is the number of species.
<code>xlab, ylab</code>	Character caption for the x-axis and y-axis. By default, a suitable caption is found. See the <code>xlab</code> argument in <code>plot</code> or <code>title</code> .

zlab	Character caption for the z-axis. Used only if Rank = 2. By default, a suitable caption is found. See the xlab argument in plot or title .
labelSpecies	Logical. Whether the species should be labelled with their names. Used for Rank = 1 only. The position of the label is just above the species' maximum.
stretch	Numeric. A value slightly more than 1, this argument adjusts the height of the y-axis. Used for Rank = 1 only.
main	Character, giving the title of the plot. See the main argument in plot or title .
ticktype	Tick type. Used only if Rank = 2. See persp for more information.
col	Color. See persp for more information.
lty	Line type. Rank-1 models only. See the lty argument of par .
lwd	Line width. Rank-1 models only. See the lwd argument of par .
add1	Logical. Add to an existing plot? Used only for rank-1 models.
...	Arguments passed into persp . Useful arguments here include theta and phi, which control the position of the eye.

Details

For a rank-1 model, a perspective plot is similar to [lvplot.qrrvglm](#) but plots the curves along a fine grid and there is no rugplot to show the site scores.

For a rank-2 model, a perspective plot has the first latent variable as the x-axis, the second latent variable as the y-axis, and the expected value (fitted value) as the z-axis. The result of a CQO is that each species has a response surface with elliptical contours. This function will, at each grid point, work out the maximum fitted value over all the species. The resulting response surface is plotted. Thus rare species will be obscured and abundant species will dominate the plot. To view rare species, use the `which.species` argument to select a subset of the species.

A perspective plot will be performed if `noRRR` = ~ 1, and Rank = 1 or 2. Also, all the tolerance matrices of those species to be plotted must be positive-definite.

Value

For a rank-2 model, a list with the following components.

<code>fitted</code>	A $(G_1 \times G_2)$ by M matrix of fitted values on the grid. Here, G_1 and G_2 are the two values of <code>gridlength</code> .
<code>latvar1grid</code> , <code>latvar2grid</code>	The grid points for the x-axis and y-axis.
<code>max.fitted</code>	A G_1 by G_2 matrix of maximum of the fitted values over all species. These are the values that are plotted on the z-axis.

For a rank-1 model, the components `latvar2grid` and `max.fitted` are NULL.

Note

Yee (2004) does not refer to perspective plots. Instead, contour plots via [lvplot.qrrvglm](#) are used.

For rank-1 models, a similar function to this one is [lvplot.qrrvglm](#). It plots the fitted values at the actual site score values rather than on a fine grid here. The result has the advantage that the user sees the curves as a direct result from a model fitted to data whereas here, it is easy to think that the smooth bell-shaped curves are the truth because the data is more of a distance away.

Author(s)

Thomas W. Yee

References

Yee, T. W. (2004). A new technique for maximum-likelihood canonical Gaussian ordination. *Ecological Monographs*, **74**, 685–701.

See Also

[persp](#), [cqo](#), [Coef.qrrvglm](#), [lvplot.qrrvglm](#), [par](#), [title](#).

Examples

```
## Not run:
hspider[, 1:6] <- scale(hspider[, 1:6]) # Good idea when I.tolerances = TRUE
set.seed(111)
r1 <- cqo(cbind(Alopacce, Alopcone, Alopfabr, Arctlute, Arctperi,
              Auloalbi, Pardmont, Pardnigr, Pardpull, Trocterr) ~
          WaterCon + BareSand + FallTwig + CoveMoss + CoveHerb + ReflLux,
          poissonff, data = hspider, trace = FALSE, I.tolerances = TRUE)
set.seed(111) # r2 below is an ill-conditioned model
r2 <- cqo(cbind(Alopacce, Alopcone, Alopfabr, Arctlute, Arctperi,
              Auloalbi, Pardmont, Pardnigr, Pardpull, Trocterr) ~
          WaterCon + BareSand + FallTwig + CoveMoss + CoveHerb + ReflLux,
          isd.lv = c(2.4, 1.0), Muxfactor = 3.0, trace = FALSE,
          poissonff, data = hspider, Rank = 2, eq.tolerances = TRUE)

sort(deviance(r1, history = TRUE)) # A history of all the fits
sort(deviance(r2, history = TRUE)) # A history of all the fits
if (deviance(r2) > 857) stop("suboptimal fit obtained")

persp(r1, xlim = c(-6, 5), col = 1:4, label = TRUE)

# Involves all species
persp(r2, xlim = c(-6, 5), ylim = c(-4, 5), theta = 10, phi = 20, zlim = c(0, 220))
# Omit the two dominant species to see what is behind them
persp(r2, xlim = c(-6, 5), ylim = c(-4, 5), theta = 10, phi = 20, zlim = c(0, 220),
      which = (1:10)[-c(8, 10)]) # Use zlim to retain the original z-scale

## End(Not run)
```

Description

The first two derivatives of the incomplete gamma integral.

Usage

```
pgamma.deriv(q, shape, tmax = 100)
```

Arguments

q, shape As in [pgamma](#) but these must be vectors of positive values only and finite.
 tmax Maximum number of iterations allowed in the computation (per q value).

Details

Write $x = q$ and $\text{shape} = a$. The first and second derivatives with respect to q and a are returned. This function is similar in spirit to [pgamma](#); define

$$P(a, x) = \frac{1}{\Gamma(a)} \int_0^x t^{a-1} e^{-t} dt$$

so that $P(a, x)$ is `pgamma(x, a)`. Currently a 6-column matrix is returned (in the future this may change and an argument may be supplied so that only what is required by the user is computed.)

The computations use a series expansion for $a \leq x \leq 1$ or $x < a$, else otherwise a continued fraction expansion. Machine overflow can occur for large values of x when x is much greater than a .

Value

The first 5 columns, running from left to right, are the derivatives with respect to: x, x^2, a, a^2, xa . The 6th column is $P(a, x)$ (but it is not as accurate as calling [pgamma](#) directly).

Note

If convergence does not occur then try increasing the value of `tmax`.

Yet to do: add more arguments to give greater flexibility in the accuracy desired and to compute only quantities that are required by the user.

Author(s)

T. W. Yee wrote the wrapper function to the Fortran subroutine written by R. J. Moore. The subroutine was modified to run using double precision. The original code came from <http://lib.stat.cmu.edu/apstat/187>. but this website has since become stale.

References

Moore, R. J. (1982). Algorithm AS 187: Derivatives of the Incomplete Gamma Integral. *Journal of the Royal Statistical Society, Series C (Applied Statistics)*, **31**(3), 330–335.

See Also

[pgamma.deriv.unscaled](#), [pgamma](#).

Examples

```
x <- seq(2, 10, length = 501)
head(ans <- pgamma.deriv(x, 2))
## Not run: par(mfrow = c(2, 3))
for (jay in 1:6)
  plot(x, ans[, jay], type = "l", col = "blue", cex.lab = 1.5,
       cex.axis = 1.5, las = 1, log = "x",
       main = colnames(ans)[jay], xlab = "q", ylab = "")
## End(Not run)
```

pgamma.deriv.unscaled *Derivatives of the Incomplete Gamma Integral (Unscaled Version)*

Description

The first two derivatives of the incomplete gamma integral with scaling.

Usage

```
pgamma.deriv.unscaled(q, shape)
```

Arguments

q, shape As in `pgamma` and `pgamma.deriv` but these must be vectors of positive values only and finite.

Details

Define

$$G(x, a) = \int_0^x t^{a-1} e^{-t} dt$$

so that $G(x, a)$ is `pgamma(x, a) * gamma(a)`. Write $x = q$ and $\text{shape} = a$. The 0th and first and second derivatives with respect to a of G are returned. This function is similar in spirit to `pgamma.deriv` but here there is no gamma function to scale things. Currently a 3-column matrix is returned (in the future this may change and an argument may be supplied so that only what is required by the user is computed.) This function is based on Wingo (1989).

Value

The 3 columns, running from left to right, are the 0:2th derivatives with respect to a .

Warning

These function seems inaccurate for $q = 1$ and $q = 2$; see the plot below.

Author(s)

T. W. Yee.

References

See [truncweibull](#).

See Also

[pgamma.deriv](#), [pgamma](#).

Examples

```
x <- 3; aa <- seq(0.3, 04, by = 0.01)
ans.u <- pgamma.deriv.unscaled(x, aa)
head(ans.u)

## Not run: par(mfrow = c(1, 3))
for (jay in 1:3) {
  plot(aa, ans.u[, jay], type = "l", col = "blue", cex.lab = 1.5,
       cex.axis = 1.5, las = 1, main = colnames(ans.u)[jay],
       log = "", xlab = "shape", ylab = "")
  abline(h = 0, v = 1:2, lty = "dashed", col = "gray") # Inaccurate at 1 and 2
}

## End(Not run)
```

plotdeplot.lmscreg *Density Plot for LMS Quantile Regression*

Description

Plots a probability density function associated with a LMS quantile regression.

Usage

```
plotdeplot.lmscreg(answer, y.arg, add.arg = FALSE,
  xlab = "", ylab = "density", xlim = NULL, ylim = NULL,
  lty.arg = par()$lty, col.arg = par()$col,
  llwd.arg = par()$lwd, ...)
```

Arguments

answer	Output from functions of the form <code>deplot.???</code> where <code>???</code> is the name of the VGAM LMS family function, e.g., <code>lms.yjn</code> . See below for details.
y.arg	Numerical vector. The values of the response variable at which to evaluate the density. This should be a grid that is fine enough to ensure the plotted curves are smooth.
add.arg	Logical. Add the density to an existing plot?
xlab, ylab	Caption for the x- and y-axes. See par .
xlim, ylim	Limits of the x- and y-axes. See par .

lty.arg	Line type. See the lty argument of par .
col.arg	Line color. See the col argument of par .
llwd.arg	Line width. See the lwd argument of par .
...	Arguments passed into the plot function when setting up the entire plot. Useful arguments here include main and las.

Details

The above graphical parameters offer some flexibility when plotting the quantiles.

Value

The list answer, which has components

newdata	The argument newdata above from the argument list of deplot.lmscreg , or a one-row data frame constructed out of the x0 argument.
y	The argument y.arg above.
density	Vector of the density function values evaluated at y.arg.

Note

While the graphical arguments of this function are useful to the user, this function should not be called directly.

Author(s)

Thomas W. Yee

References

Yee, T. W. (2004). Quantile regression via vector generalized additive models. *Statistics in Medicine*, **23**, 2295–2315.

See Also

[deplot.lmscreg](#).

Examples

```
fit <- vgam(BMI ~ s(age, df = c(4,2)), lms.bcn(zero = 1), bmi.nz)
## Not run: y = seq(15, 43, by = 0.25)
deplot(fit, x0 = 20, y = y, xlab = "BMI", col = "green", llwd = 2,
       main = "BMI distribution at ages 20 (green), 40 (blue), 60 (orange)")
deplot(fit, x0 = 40, y = y, add = TRUE, col = "blue", llwd = 2)
deplot(fit, x0 = 60, y = y, add = TRUE, col = "orange", llwd = 2) -> aa

names(aa$post$deplot)
aa$post$deplot$newdata
head(aa$post$deplot$y)
head(aa$post$deplot$density)
## End(Not run)
```

Description

Given a GAITD regression object, plots the probability mass function.

Usage

```
plotdgaitd(object, ...)  
plotdgaitd.vglm(object, ...)
```

Arguments

object	A fitted GAITD combo regression, e.g., gaitdpoisson .
...	Graphical arguments passed into dgaitdplot .

Details

This is meant to be a more convenient function for plotting the PMF of the GAITD combo model from a fitted regression model. The fit should be intercept-only and the distribution should have 1 or 2 parameters. Currently it should work for a [gaitdpoisson](#) fit. As much information as needed such as the special values is extracted from the object and fed into [dgaitdplot](#).

Value

Same as [dgaitdplot](#).

Note

This function is subject to change.

Author(s)

T. W. Yee.

See Also

[dgaitdplot](#), [spikeplot](#), [gaitdpoisson](#).

Examples

```
## Not run:  
example(gaitdpoisson)  
gaitpfit2 <-  
  vglm(y1 ~ 1, crit = "coef", trace = TRUE, data = gdata,  
    gaitdpoisson(a.mix = a.mix, i.mix = i.mix,  
      i.nlm = i.nlm, eq.ap = TRUE, eq.ip = TRUE,
```

```

                                truncate = tvec, max.support = max.support))
plotdgaitd(gaitpfit2)

## End(Not run)

```

plotqrrvglm

Model Diagnostic Plots for QRR-VGLMs

Description

The residuals of a QRR-VGLM are plotted for model diagnostic purposes.

Usage

```

plotqrrvglm(object, rtype = c("response", "pearson", "deviance", "working"),
            ask = FALSE,
            main = paste(Rtype, "residuals vs latent variable(s)",
                        xlab = "Latent Variable",
                        I.tolerances = object@control$eq.tolerances, ...)

```

Arguments

object	An object of class "qrrvglm".
rtype	Character string giving residual type. By default, the first one is chosen.
ask	Logical. If TRUE, the user is asked to hit the return key for the next plot.
main	Character string giving the title of the plot.
xlab	Character string giving the x-axis caption.
I.tolerances	Logical. This argument is fed into <code>Coef(object, I.tolerances = I.tolerances)</code> .
...	Other plotting arguments (see par).

Details

Plotting the residuals can be potentially very useful for checking that the model fit is adequate.

Value

The original object.

Note

An ordination plot of a QRR-VGLM can be obtained by [lvplot.qrrvglm](#).

Author(s)

Thomas W. Yee

Arguments

fitted.values	Matrix of fitted values.
object	A VGAM quantile regression model, i.e., an object produced by modelling functions such as <code>vglm</code> and <code>vgam</code> with a family function beginning with "lms.", e.g., <code>lms.yjn</code> .
newdata	Data frame at which predictions are made. By default, the original data are used.
percentiles	Numerical vector with values between 0 and 100 that specify the percentiles (quantiles). The default is to use the percentiles when fitting the model. For example, the value 50 corresponds to the median.
lp	Length of percentiles.
add.arg	Logical. Add the quantiles to an existing plot?
y	Logical. Add the response as points to the plot?
spline.fit	Logical. Add a spline curve to the plot?
label	Logical. Add the percentiles (as text) to the plot?
size.label	Numeric. How much room to leave at the RHS for the label. It is in percent (of the range of the primary variable).
xlab	Caption for the x-axis. See <code>par</code> .
ylab	Caption for the y-axis. See <code>par</code> .
pch	Plotting character. See <code>par</code> .
pcex	Character expansion of the points. See <code>par</code> .
pcol.arg	Color of the points. See the <code>col</code> argument of <code>par</code> .
xlim	Limits of the x-axis. See <code>par</code> .
ylim	Limits of the y-axis. See <code>par</code> .
lty.arg	Line type. Line type. See the <code>lty</code> argument of <code>par</code> .
lcol.arg	Color of the lines. See the <code>col</code> argument of <code>par</code> .
llwd.arg	Line width. See the <code>lwd</code> argument of <code>par</code> .
tcol.arg	Color of the text (if <code>label</code> is TRUE). See the <code>col</code> argument of <code>par</code> .
tadj	Text justification. See the <code>adj</code> argument of <code>par</code> .
...	Arguments passed into the <code>plot</code> function when setting up the entire plot. Useful arguments here include <code>main</code> and <code>las</code> .

Details

The above graphical parameters offer some flexibility when plotting the quantiles.

Value

The matrix of fitted values.

Note

While the graphical arguments of this function are useful to the user, this function should not be called directly.

Author(s)

Thomas W. Yee

References

Yee, T. W. (2004). Quantile regression via vector generalized additive models. *Statistics in Medicine*, **23**, 2295–2315.

See Also

[qtplot.lmscreg](#).

Examples

```
## Not run:
fit <- vgam(BMI ~ s(age, df = c(4,2)), lms.bcn(zero = 1), data = bmi.nz)
qtplot(fit)
qtplot(fit, perc = c(25,50,75,95), lcol = "blue", tcol = "blue", llwd = 2)

## End(Not run)
```

plotrcim0

Main Effects Plot for a Row-Column Interaction Model (RCIM)

Description

Produces a main effects plot for Row-Column Interaction Models (RCIMs).

Usage

```
plotrcim0(object, centered = TRUE, which.plots = c(1, 2),
          hline0 = TRUE, hlty = "dashed", hcol = par()$col, hlwd = par()$lwd,
          rfirst = 1, cfirst = 1,
          rtype = "h", ctype = "h",
          rcex.lab = 1, rcex.axis = 1, rtick = FALSE,
          ccex.lab = 1, ccex.axis = 1, ctick = FALSE,
          rmain = "Row effects", rsub = "",
          rxlab = "", rylab = "Row effects",
          cmain = "Column effects", csub = "",
          cxlab = "", cylab = "Column effects",
          rcol = par()$col, ccol = par()$col,
          no.warning = FALSE, ...)
```

Arguments

object	An rcim object. This should be of rank-0, i.e., main effects only and no interactions.
which.plots	Numeric, describing which plots are to be plotted. The row effects plot is 1 and the column effects plot is 2. Set the value 0, say, for no plots at all.
centered	Logical. If TRUE then the row and column effects are centered (but not scaled) by scale . If FALSE then the raw effects are used (of which the first are zero by definition).
hline0, hlty, hcol, hlwd	hline0 is logical. If TRUE then a horizontal line is plotted at 0 and the other arguments describe this line. Probably having hline0 = TRUE only makes sense when centered = TRUE.
rfirst, cfirst	rfirst is the level of row that is placed first in the row effects plot, etc.
rmain, cmain	Character. rmain is the main label in the row effects plot, etc.
rtype, ctype, rsub, csub	See the type and sub arguments of plot.default .
rxlab, rylab, cxlab, cylab	Character. For the row effects plot, rxlab is xlab and rylab is ylab; see par . Ditto for cxlab and cylab for the column effects plot.
rcex.lab, ccex.lab	Numeric. rcex.lab is cex for the row effects plot label, etc.
rcex.axis, ccex.axis	Numeric. rcex.axis is the cex argument for the row effects axis label, etc.
rtick, ctick	Logical. If rtick = TRUE then add ticks to the row effects plot, etc.
rcol, ccol	rcol give a colour for the row effects plot, etc.
no.warning	Logical. If TRUE then no warning is issued if the model is not rank-0.
...	Arguments fed into plot.default , etc.

Details

This function plots the row and column effects of a rank-0 RCIM. As the result is a main effects plot of a regression analysis, its interpretation when centered = FALSE is relative to the baseline (reference level) of a row and column, and should also be considered in light of the link function used. Many arguments that start with "r" refer to the row effects plot, and "c" for the column effects plot.

Value

The original object with the post slot assigned additional information from the plot.

Note

This function should be only used to plot the object of rank-0 RCIM. If the rank is positive then it will issue a warning.

Using an argument ylim will mean the row and column effects are plotted on a common scale; see [plot.window](#).

Author(s)

T. W. Yee, A. F. Hadi.

See Also

[moffset](#) [Rcim](#), [rcim](#).

Examples

```
alcoff.e <- moffset(alcoff, "6", "Mon", postfix = "*") # Effective day
fit0 <- rcim(alcoff.e, family = poissonff)
## Not run: par(oma = c(0, 0, 4, 0), mfrow = 1:2) # For all plots below too
ii <- plot(fit0, rcol = "blue", ccol = "orange",
          lwd = 4, ylim = c(-2, 2), # A common ylim
          cylab = "Effective daily effects", rylab = "Hourly effects",
          rxlab = "Hour", cxlab = "Effective day")
ii@post # Endowed with additional information

## End(Not run)

# Negative binomial example
## Not run:
fit1 <- rcim(alcoff.e, negbinomial, trace = TRUE)
plot(fit1, ylim = c(-2, 2))
## End(Not run)

# Univariate normal example
fit2 <- rcim(alcoff.e, uninormal, trace = TRUE)
## Not run: plot(fit2, ylim = c(-200, 400))

# Median-polish example
## Not run:
fit3 <- rcim(alcoff.e, alaplace1(tau = 0.5), maxit = 1000, trace = FALSE)
plot(fit3, ylim = c(-200, 250))
## End(Not run)

# Zero-inflated Poisson example on "crashp" (no 0s in alcoff)
## Not run:
cbind(rowSums(crashp)) # Easy to see the data
cbind(colSums(crashp)) # Easy to see the data
fit4 <- rcim(Rcim(crashp, rbaseline = "5", cbaseline = "Sun"),
            zipoissonff, trace = TRUE)
plot(fit4, ylim = c(-3, 3))
## End(Not run)
```

Description

Component functions of a `vgam-class` object can be plotted with `plotvgam()`. These are on the scale of the linear/additive predictor.

Usage

```
plotvgam(x, newdata = NULL, y = NULL, residuals = NULL,
         rugplot = TRUE, se = FALSE, scale = 0, raw = TRUE,
         offset.arg = 0, deriv.arg = 0, overlay = FALSE,
         type.residuals = c("deviance", "working", "pearson", "response"),
         plot.arg = TRUE, which.term = NULL, which.cf = NULL,
         control = plotvgam.control(...), varxij = 1, ...)
```

Arguments

<code>x</code>	A fitted VGAM object, e.g., produced by <code>vgam</code> , <code>vglm</code> , or <code>rrvglm</code> .
<code>newdata</code>	Data frame. May be used to reconstruct the original data set.
<code>y</code>	Unused.
<code>residuals</code>	Logical. If TRUE then residuals are plotted. See <code>type.residuals</code>
<code>rugplot</code>	Logical. If TRUE then a rug plot is plotted at the foot of each plot. These values are jittered to expose ties.
<code>se</code>	Logical. If TRUE then approximate ± 2 pointwise standard error bands are included in the plot.
<code>scale</code>	Numerical. By default, each plot will have its own y-axis scale. However, by specifying a value, each plot's y-axis scale will be at least <code>scale</code> wide.
<code>raw</code>	Logical. If TRUE then the smooth functions are those obtained directly by the algorithm, and are plotted without having to premultiply with the constraint matrices. If FALSE then the smooth functions have been premultiply by the constraint matrices. The <code>raw</code> argument is directly fed into <code>predict.vgam()</code> .
<code>offset.arg</code>	Numerical vector of length <code>r</code> . These are added to the component functions. Useful for separating out the functions when <code>overlay</code> is TRUE. If <code>overlay</code> is TRUE and there is one covariate then using the intercept values as the offsets can be a good idea.
<code>deriv.arg</code>	Numerical. The order of the derivative. Should be assigned an small integer such as 0, 1, 2. Only applying to <code>s()</code> terms, it plots the derivative.
<code>overlay</code>	Logical. If TRUE then component functions of the same covariate are overlaid on each other. The functions are centered, so <code>offset.arg</code> can be useful when <code>overlay</code> is TRUE.
<code>type.residuals</code>	if <code>residuals</code> is TRUE then the first possible value of this vector, is used to specify the type of residual.
<code>plot.arg</code>	Logical. If FALSE then no plot is produced.
<code>which.term</code>	Character or integer vector containing all terms to be plotted, e.g., <code>which.term = c("s(age)", "s(height"))</code> or <code>which.term = c(2, 5, 9)</code> . By default, all are plotted.

which.cf	An integer-valued vector specifying which linear/additive predictors are to be plotted. The values must be from the set $\{1, 2, \dots, r\}$. By default, all are plotted.
control	Other control parameters. See plotvgam.control .
...	Other arguments that can be fed into plotvgam.control . This includes line colors, line widths, line types, etc.
varxij	Positive integer. Used if <code>xij</code> of vglm.control was used, this chooses which inner argument the component is plotted against. This argument is related to <code>raw = TRUE</code> and terms such as <code>NS(dum1, dum2)</code> and constraint matrices that have more than one column. The default would plot the smooth against <code>dum1</code> but setting <code>varxij = 2</code> could mean plotting the smooth against <code>dum2</code> . See the VGAM website for further information.

Details

In this help file M is the number of linear/additive predictors, and r is the number of columns of the constraint matrix of interest.

Many of `plotvgam()`'s options can be found in [plotvgam.control](#), e.g., line types, line widths, colors.

Value

The original object, but with the `preplot` slot of the object assigned information regarding the plot.

Note

While `plot(fit)` will work if `class(fit)` is "vgam", it is necessary to use `plotvgam(fit)` explicitly otherwise.

`plotvgam()` is quite buggy at the moment.

Author(s)

Thomas W. Yee

See Also

[vgam](#), [plotvgam.control](#), [predict.vgam](#), [plotvglm](#), [vglm](#).

Examples

```
coalminers <- transform(coalminers, Age = (age - 42) / 5)
fit <- vgam(cbind(nBnW, nBW, BnW, BW) ~ s(Age),
            binom2.or(zero = NULL), data = coalminers)
## Not run: par(mfrow = c(1,3))
plot(fit, se = TRUE, ylim = c(-3, 2), las = 1)
plot(fit, se = TRUE, which.cf = 1:2, lcol = "blue", scol = "orange",
     ylim = c(-3, 2))
plot(fit, se = TRUE, which.cf = 1:2, lcol = "blue", scol = "orange",
     overlay = TRUE)
## End(Not run)
```

plotvgam.control *Control Function for plotvgam()*

Description

Provides default values for many arguments available for plotvgam().

Usage

```
plotvgam.control(which.cf = NULL,
  xlim = NULL, ylim = NULL, llty = par()$lty, slty = "dashed",
  pcex = par()$cex, pch = par()$pch, pcol = par()$col,
  lcol = par()$col, rcol = par()$col, scol = par()$col,
  llwd = par()$lwd, slwd = par()$lwd, add.arg = FALSE,
  one.at.a.time = FALSE, .include.dots = TRUE, noxmean = FALSE,
  shade = FALSE, shcol = "gray80", main = "", ...)
```

Arguments

which.cf	Integer vector specifying which component functions are to be plotted (for each covariate). Must have values from the set $\{1, 2, \dots, M\}$.
xlim	Range for the x-axis.
ylim	Range for the y-axis.
llty	Line type for the fitted functions (lines). Fed into par(lty).
slty	Line type for the standard error bands. Fed into par(lty).
pcex	Character expansion for the points (residuals). Fed into par(cex).
pch	Character used for the points (residuals). Same as par(pch).
pcol	Color of the points. Fed into par(col).
lcol	Color of the fitted functions (lines). Fed into par(col).
rcol	Color of the rug plot. Fed into par(col).
scol	Color of the standard error bands. Fed into par(col).
llwd	Line width of the fitted functions (lines). Fed into par(lwd).
slwd	Line width of the standard error bands. Fed into par(lwd).
add.arg	Logical. If TRUE then the plot will be added to an existing plot, otherwise a new plot will be made.
one.at.a.time	Logical. If TRUE then the plots are done one at a time, with the user having to hit the return key between the plots.
.include.dots	Not to be used by the user.
noxmean	Logical. If TRUE then the point at the mean of x , which is added when standard errors are specified and it thinks the function is linear, is not added. One might use this argument if ylab is specified.

shade, shcol shade is logical; if TRUE then the pointwise SE band is shaded gray by default. The colour can be adjusted by setting shcol. These arguments are ignored unless se = TRUE and overlay = FALSE; If shade = TRUE then scol is ignored.

main Character vector, recycled to the number needed.

... Other arguments that may be fed into par().

In the above, M is the number of linear/additive predictors.

Details

The most obvious features of `plotvglm` can be controlled by the above arguments.

Value

A list with values matching the arguments.

Author(s)

Thomas W. Yee

References

Yee, T. W. and Wild, C. J. (1996). Vector generalized additive models. *Journal of the Royal Statistical Society, Series B, Methodological*, **58**, 481–493.

See Also

`plotvgam`.

Examples

```
plotvglm.control(lcol = c("red", "blue"), scol = "darkgreen", se = TRUE)
```

plotvglm

Plots for VGLMs

Description

Currently this function plots the Pearson residuals versus the linear predictors (M plots) and plots the Pearson residuals versus the hat values (M plots).

Usage

```
plotvglm(x, which = "(All)", ...)
```

Arguments

<code>x</code>	An object of class "vglm" (see vglm-class) or inherits from that class.
<code>which</code>	If a subset of the plots is required, specify a subset of the numbers 1:(2*M). The default is to plot them all.
<code>...</code>	Arguments fed into the primitive plot functions.

Details

This function is under development. Currently it plots the Pearson residuals against the predicted values (on the transformed scale) and the hat values. There are $2M$ plots in total, therefore users should call [par](#) to assign, e.g., the `mfrow` argument. Note: Section 3.7 of Yee (2015) describes the Pearson residuals and hat values for VGLMs.

Value

Returns the object invisibly.

Author(s)

T. W. Yee

See Also

[plotvgam](#), [plotvgam.control](#), [vglm](#).

Examples

```
## Not run:
ndata <- data.frame(x2 = runif(nn <- 200))
ndata <- transform(ndata, y1 = rnbino(nn, mu = exp(3+x2), size = exp(1)))
fit1 <- vglm(y1 ~ x2, negbinomial, data = ndata, trace = TRUE)
coef(fit1, matrix = TRUE)
par(mfrow = c(2, 2))
plot(fit1)

# Manually produce the four plots
plot(fit1, which = 1, col = "blue", las = 1, main = "main1")
abline(h = 0, lty = "dashed", col = "gray50")
plot(fit1, which = 2, col = "blue", las = 1, main = "main2")
abline(h = 0, lty = "dashed", col = "gray50")
plot(fit1, which = 3, col = "blue", las = 1, main = "main3")
plot(fit1, which = 4, col = "blue", las = 1, main = "main4")

## End(Not run)
```

pneumo

Pneumoconiosis in Coalminers Data

Description

The pneumo data frame has 8 rows and 4 columns. Exposure time is explanatory, and there are 3 ordinal response variables.

Usage

```
data(pneumo)
```

Format

This data frame contains the following columns:

exposure.time a numeric vector, in years

normal a numeric vector, counts

mild a numeric vector, counts

severe a numeric vector, counts

Details

These were collected from coalface workers. In the original data set, the two most severe categories were combined.

Source

Ashford, J.R., 1959. An approach to the analysis of data for semi-quantal responses in biological assay. *Biometrics*, **15**, 573–581.

References

McCullagh, P. and Nelder, J. A. (1989). *Generalized Linear Models*, 2nd ed. London: Chapman & Hall.

See Also

[cumulative](#).

Examples

```
# Fit the proportional odds model, p.179, in McCullagh and Nelder (1989)
pneumo <- transform(pneumo, let = log(exposure.time))
vglm(cbind(normal, mild, severe) ~ let, propodds, data = pneumo)
```

poisson.points *Poisson-points-on-a-plane/volume Distances Distribution*

Description

Estimating the density parameter of the distances from a fixed point to the u -th nearest point, in a plane or volume.

Usage

```
poisson.points(ostatistic, dimension = 2, link = "loglink",
               idensity = NULL, imethod = 1)
```

Arguments

ostatistic	Order statistic. A single positive value, usually an integer. For example, the value 5 means the response are the distances of the fifth nearest value to that point (usually over many planes or volumes). Non-integers are allowed because the value 1.5 coincides with maxwell when <code>dimension = 2</code> . Note: if <code>ostatistic = 1</code> and <code>dimension = 2</code> then this VGAM family function coincides with rayleigh .
dimension	The value 2 or 3; 2 meaning a plane and 3 meaning a volume.
link	Parameter link function applied to the (positive) density parameter, called λ below. See Links for more choices.
idensity	Optional initial value for the parameter. A NULL value means a value is obtained internally. Use this argument if convergence failure occurs.
imethod	An integer with value 1 or 2 which specifies the initialization method for λ . If failure to converge occurs try another value and/or else specify a value for <code>idensity</code> .

Details

Suppose the number of points in any region of area A of the plane is a Poisson random variable with mean λA (i.e., λ is the *density* of the points). Given a fixed point P , define D_1, D_2, \dots to be the distance to the nearest point to P , second nearest to P , etc. This **VGAM** family function estimates λ since the probability density function for D_u is easily derived, $u = 1, 2, \dots$. Here, u corresponds to the argument `ostatistic`.

Similarly, suppose the number of points in any volume V is a Poisson random variable with mean λV where, once again, λ is the *density* of the points. This **VGAM** family function estimates λ by specifying the argument `ostatistic` and using `dimension = 3`.

The mean of D_u is returned as the fitted values. Newton-Raphson is the same as Fisher-scoring.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), [rrvglm](#) and [vgam](#).

Warning

Convergence may be slow if the initial values are far from the solution. This often corresponds to the situation when the response values are all close to zero, i.e., there is a high density of points.

Formulae such as the means have not been fully checked.

Author(s)

T. W. Yee

See Also

[poissonff](#), [maxwell](#), [rayleigh](#).

Examples

```
pdata <- data.frame(y = rgamma(10, shape = exp(-1))) # Not proper data!
ostat <- 2
fit <- vglm(y ~ 1, poisson.points(ostat, 2), data = pdata,
            trace = TRUE, crit = "coef")
fit <- vglm(y ~ 1, poisson.points(ostat, 3), data = pdata,
            trace = TRUE, crit = "coef") # Slow convergence?
fit <- vglm(y ~ 1, poisson.points(ostat, 3, idensi = 1), data = pdata,
            trace = TRUE, crit = "coef")
head(fitted(fit))
with(pdata, mean(y))
coef(fit, matrix = TRUE)
Coef(fit)
```

poissonff

Poisson Regression

Description

Family function for a generalized linear model fitted to Poisson responses.

Usage

```
poissonff(link = "loglink", imu = NULL, imethod = 1,
           parallel = FALSE, zero = NULL, bred = FALSE,
           earg.link = FALSE, type.fitted = c("mean", "quantiles"),
           percentiles = c(25, 50, 75))
```

Arguments

link	Link function applied to the mean or means. See Links for more choices and information.
parallel	A logical or formula. Used only if the response is a matrix.
imu, imethod	See CommonVGAMffArguments for more information.
zero	Can be an integer-valued vector specifying which linear/additive predictors are modelled as intercepts only. The values must be from the set $\{1, 2, \dots, M\}$, where M is the number of columns of the matrix response. See CommonVGAMffArguments for more information.
bred, earg.link	Details at CommonVGAMffArguments . Setting bred = TRUE should work for multiple responses and all VGAM link functions; it has been tested for loglink , identity but further testing is required.
type.fitted, percentiles	Details at CommonVGAMffArguments .

Details

M defined above is the number of linear/additive predictors. With overdispersed data try [negbinomial](#).

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), [vgam](#), [rrvglm](#), [cqo](#), and [cao](#).

Warning

With multiple responses, assigning a known dispersion parameter for *each* response is not handled well yet. Currently, only a single known dispersion parameter is handled well.

Note

This function will handle a matrix response automatically.

Regardless of whether the dispersion parameter is to be estimated or not, its value can be seen from the output from the `summary()` of the object.

Author(s)

Thomas W. Yee

References

McCullagh, P. and Nelder, J. A. (1989). *Generalized Linear Models*, 2nd ed. London: Chapman & Hall.

See Also

[Links](#), [hdeff.vglm](#), [negbinomial](#), [genpoisson1](#), [genpoisson2](#), [genpoisson0](#), [gaitdpoisson](#), [zipoisson](#), [N1poisson](#), [pospoisson](#), [skellam](#), [mix2poisson](#), [cens.poisson](#), [ordpoisson](#), [amlpoisson](#), [inv.binomial](#), [simulate.vlm](#), [loglink](#), [polf](#), [rrvglm](#), [cqo](#), [cao](#), [binomialff](#), [poisson](#), [Poisson](#), [poisson.points](#), [ruge](#), [V1](#), [V2](#), [residualsvglm](#), [margeff](#).

Examples

```
poissonff()

set.seed(123)
pdata <- data.frame(x2 = rnorm(nn <- 100))
pdata <- transform(pdata, y1 = rpois(nn, exp(1 + x2)),
                  y2 = rpois(nn, exp(1 + x2)))
(fit1 <- vglm(cbind(y1, y2) ~ x2, poissonff, data = pdata))
(fit2 <- vglm(y1 ~ x2, poissonff(bred = TRUE), data = pdata))
coef(fit1, matrix = TRUE)
coef(fit2, matrix = TRUE)

nn <- 200
cdata <- data.frame(x2 = rnorm(nn), x3 = rnorm(nn), x4 = rnorm(nn))
cdata <- transform(cdata, lv1 = 0 + x3 - 2*x4)
cdata <- transform(cdata, lambda1 = exp(3 - 0.5 * (lv1-0)^2),
                  lambda2 = exp(2 - 0.5 * (lv1-1)^2),
                  lambda3 = exp(2 - 0.5 * ((lv1+4)/2)^2))
cdata <- transform(cdata, y1 = rpois(nn, lambda1),
                  y2 = rpois(nn, lambda2),
                  y3 = rpois(nn, lambda3))
## Not run: lvplot(p1, y = TRUE, lcol = 2:4, pch = 2:4, pcol = 2:4, rug = FALSE)
```

PoissonPoints

Poisson Points Distribution

Description

Density for the PoissonPoints distribution.

Usage

```
dpois.points(x, lambda, ostatic, dimension = 2, log = FALSE)
```

Arguments

x	vector of quantiles.
lambda	the mean density of points.
ostatic	positive values, usually integers.
dimension	Either 2 and/or 3.
log	Logical; if TRUE, the logarithm is returned.

Details

See [poisson.points](#), the **VGAM** family function for estimating the parameters, for the formula of the probability density function and other details.

Value

`dpois.points` gives the density.

See Also

[poisson.points](#), [dpois](#), [Maxwell](#).

Examples

```
## Not run: lambda <- 1; xvec <- seq(0, 2, length = 400)
plot(xvec, dpois.points(xvec, lambda, ostat = 1, dimension = 2),
     type = "l", las = 1, col = "blue",
     sub = "First order statistic",
     main = paste("PDF of PoissonPoints distribution with lambda = ",
                  lambda, " and on the plane", sep = ""))
## End(Not run)
```

 Polono

The Poisson Lognormal Distribution

Description

Density, distribution function and random generation for the Poisson lognormal distribution.

Usage

```
dpolono(x, meanlog = 0, sdlog = 1, bigx = 170, ...)
ppolono(q, meanlog = 0, sdlog = 1,
        isOne = 1 - sqrt(.Machine$double.eps), ...)
rpolono(n, meanlog = 0, sdlog = 1)
```

Arguments

<code>x, q</code>	vector of quantiles.
<code>n</code>	number of observations. If <code>length(n) > 1</code> then the length is taken to be the number required.
<code>meanlog, sdlog</code>	the mean and standard deviation of the normal distribution (on the log scale). They match the arguments in Lognormal .

<code>bigx</code>	Numeric. This argument is for handling large values of x and/or when <code>integrate</code> fails. A first order Taylor series approximation [Equation (7) of Bulmer (1974)] is used at values of x that are greater or equal to this argument. For <code>bigx = 10</code> , he showed that the approximation has a relative error less than 0.001 for values of <code>meanlog</code> and <code>sdlog</code> “likely to be encountered in practice”. The argument can be assigned <code>Inf</code> in which case the approximation is not used.
<code>isOne</code>	Used to test whether the cumulative probabilities have effectively reached unity.
<code>...</code>	Arguments passed into <code>integrate</code> .

Details

The Poisson lognormal distribution is similar to the negative binomial in that it can be motivated by a Poisson distribution whose mean parameter comes from a right skewed distribution (gamma for the negative binomial and lognormal for the Poisson lognormal distribution).

Value

`dpolono` gives the density, `ppolono` gives the distribution function, and `rpolono` generates random deviates.

Note

By default, `dpolono` involves numerical integration that is performed using `integrate`. Consequently, computations are very slow and numerical problems may occur (if so then the use of `...` may be needed). Alternatively, for extreme values of x , `meanlog`, `sdlog`, etc., the use of `bigx = Inf` avoids the call to `integrate`, however the answer may be a little inaccurate.

For the maximum likelihood estimation of the 2 parameters a **VGAM** family function called `polono()`, say, has not been written yet.

Author(s)

T. W. Yee. Some anonymous soul kindly wrote `ppolono()` and improved the original `dpolono()`.

References

Bulmer, M. G. (1974). On fitting the Poisson lognormal distribution to species-abundance data. *Biometrics*, **30**, 101–110.

See Also

[lognormal](#), [poissonff](#), [negbinomial](#).

Examples

```
meanlog <- 0.5; sdlog <- 0.5; yy <- 0:19
sum(proby <- dpolono(yy, m = meanlog, sd = sdlog)) # Should be 1
max(abs(cumsum(proby) - ppolono(yy, m = meanlog, sd = sdlog))) # 0?

## Not run: opar = par(no.readonly = TRUE)
par(mfrow = c(2, 2))
```

```

plot(yy, proby, type = "h", col = "blue", ylab = "P[Y=y]", log = "",
     main = paste0("Poisson lognormal(m = ", meanlog,
                  ", sdl = ", sdlog, ")"))

y <- 0:190 # More extreme values; use the approxn & plot on a log scale
(sum(proby <- dpolono(y, m = meanlog, sd = sdlog, bigx = 100))) # 1?
plot(y, proby, type = "h", col = "blue", ylab = "P[Y=y] (log)", log = "y",
     main = paste0("Poisson lognormal(m = ", meanlog,
                  ", sdl = ", sdlog, ")")) # Note the kink at bigx

# Random number generation
table(y <- rpolono(n = 1000, m = meanlog, sd = sdlog))
hist(y, breaks = ((-1):max(y))+0.5, prob = TRUE, border = "blue")
par(opar)
## End(Not run)

```

posbernoulli.b

Positive Bernoulli Family Function with Behavioural Effects

Description

Fits a GLM-/GAM-like model to multiple Bernoulli responses where each row in the capture history matrix response has at least one success (capture). Capture history behavioural effects are accommodated.

Usage

```

posbernoulli.b(link = "logitlink", drop.b = FALSE ~ 1,
              type.fitted = c("likelihood.cond", "mean.uncond"), I2 = FALSE,
              ipcapture = NULL, iprecapture = NULL,
              p.small = 1e-4, no.warning = FALSE)

```

Arguments

link, drop.b, ipcapture, iprecapture

See [CommonVGAMffArguments](#) for information about these arguments. By default the parallelism assumption does not apply to the intercept. With an intercept-only model setting drop.b = TRUE ~ 1 results in the M_0/M_h model.

I2 Logical. This argument is used for terms that are not parallel. If TRUE then the constraint matrix `diag(2)` (the general default constraint matrix in **VGAM**) is used, else `cbind(0:1, 1)`. The latter means the first element/column corresponds to the behavioural effect. Consequently it and its standard error etc. can be accessed directly without subtracting two quantities.

type.fitted Details at [posbernoulli.tb](#).

p.small, no.warning

See [posbernoulli.t](#).


```

M.b <- vglm(cbind(y1, y2, y3, y4, y5, y6) ~ 1,
            posbernoulli.b, data = deermice, trace = TRUE)
coef(M.b)[("(Intercept):1"] # Behavioural effect on logit scale
coef(M.b, matrix = TRUE)
constraints(M.b, matrix = TRUE)
summary(M.b, presid = FALSE)

# Fit a M_bh model
M.bh <- vglm(cbind(y1, y2, y3, y4, y5, y6) ~ sex + weight,
            posbernoulli.b, data = deermice, trace = TRUE)
coef(M.bh, matrix = TRUE)
coef(M.bh)[("(Intercept):1"] # Behavioural effect on logit scale
# (2,1) elt is for the behavioural effect:
constraints(M.bh)[["(Intercept)"]]
summary(M.bh, presid = FALSE) # Significant trap-happy effect
# Approx. 95 percent confidence for the behavioural effect:
SE.M.bh <- coef(summary(M.bh))[("(Intercept):1", "Std. Error")]
coef(M.bh)[("(Intercept):1"] + c(-1, 1) * 1.96 * SE.M.bh

# Fit a M_h model
M.h <- vglm(cbind(y1, y2, y3, y4, y5, y6) ~ sex + weight,
            posbernoulli.b(drop.b = TRUE ~ sex + weight),
            data = deermice, trace = TRUE)
coef(M.h, matrix = TRUE)
constraints(M.h, matrix = TRUE)
summary(M.h, presid = FALSE)

# Fit a M_0 model
M.0 <- vglm(cbind( y1 + y2 + y3 + y4 + y5 + y6,
                  6 - y1 - y2 - y3 - y4 - y5 - y6) ~ 1,
            posbinomial, data = deermice, trace = TRUE)
coef(M.0, matrix = TRUE)
summary(M.0, presid = FALSE)

# Simulated data set ,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,
set.seed(123); nTimePts <- 5; N <- 1000 # N is the popn size
pdata <- rposbern(N, nTimePts=nTimePts, pvars=2, is.popn=TRUE)
nrow(pdata) # < N (because some animals were never captured)
# The truth: xcoeffs are c(-2, 1, 2) and cap.effect = +1

M.bh.2 <- vglm(cbind(y1, y2, y3, y4, y5) ~ x2,
            posbernoulli.b, data = pdata, trace = TRUE)
coef(M.bh.2)
coef(M.bh.2, matrix = TRUE)
constraints(M.bh.2, matrix = TRUE)
summary(M.bh.2, presid = FALSE)
head(depvar(M.bh.2)) # Capture history response matrix
head(M.bh.2@extra$cap.hist1) # Info on its capture history
head(M.bh.2@extra$cap1) # When it was first captured
head(fitted(M.bh.2)) # Depends on capture history
(trap.effect <- coef(M.bh.2)[("(Intercept):1"]) # Should be +1
head(model.matrix(M.bh.2, type = "v1m"), 21)

```

```

head(pdata)
summary(pdata)
dim(depvar(M.bh.2))
vcov(M.bh.2)

M.bh.2@extra$N.hat # Population size estimate; should be about N
M.bh.2@extra$SE.N.hat # SE of the estimate of the population size
# An approximate 95 percent confidence interval:
round(M.bh.2@extra$N.hat + c(-1, 1)*1.96* M.bh.2@extra$SE.N.hat, 1)

## End(Not run)

```

posbernoulli.t

Positive Bernoulli Family Function with Time Effects

Description

Fits a GLM/GAM-like model to multiple Bernoulli responses where each row in the capture history matrix response has at least one success (capture). Sampling occasion effects are accommodated.

Usage

```

posbernoulli.t(link = "logitlink", parallel.t = FALSE ~ 1,
  iprob = NULL, p.small = 1e-4, no.warning = FALSE,
  type.fitted = c("probs", "onempall0"))

```

Arguments

link, iprob, parallel.t

See [CommonVGAMffArguments](#) for information. By default, the parallelism assumption does not apply to the intercept. Setting `parallel.t = FALSE ~ -1`, or equivalently `parallel.t = FALSE ~ 0`, results in the M_0/M_h model.

p.small, no.warning

A small probability value used to give a warning for the Horvitz–Thompson estimator. Any estimated probability value less than `p.small` will result in a warning, however, setting `no.warning = TRUE` will suppress this warning if it occurs. This is because the Horvitz–Thompson estimator is the sum of the reciprocal of such probabilities, therefore any probability that is too close to 0 will result in an unstable estimate.

type.fitted

See [CommonVGAMffArguments](#) for information. The default is to return a matrix of probabilities. If "onempall0" is chosen then the the probability that each animal is captured at least once in the course of the study is returned. The abbreviation stands for one minus the probability of all 0s, and the quantity appears in the denominator of the usual formula.

Details

These models (commonly known as M_t or M_{th} (no prefix h means it is an intercept-only model) in the capture–recapture literature) operate on a capture history matrix response of 0s and 1s ($n \times \tau$). Each column is a sampling occasion where animals are potentially captured (e.g., a field trip), and each row is an individual animal. Capture is a 1, else a 0. No removal of animals from the population is made (closed population), e.g., no immigration or emigration. Each row of the response matrix has at least one capture. Once an animal is captured for the first time, it is marked/tagged so that its future capture history can be recorded. Then it is released immediately back into the population to remix. It is released immediately after each recapture too. It is assumed that the animals are independent and that, for a given animal, each sampling occasion is independent. And animals do not lose their marks/tags, and all marks/tags are correctly recorded.

The number of linear/additive predictors is equal to the number of sampling occasions, i.e., $M = \tau$, say. The default link functions are $(\text{logit } p_1, \dots, \text{logit } p_\tau)^T$ where each p_j denotes the probability of capture at time point j . The fitted value returned is a matrix of probabilities of the same dimension as the response matrix.

A conditional likelihood is maximized here using Fisher scoring. Each sampling occasion has a separate probability that is modelled here. The probabilities can be constrained to be equal by setting `parallel.t = FALSE ~ 0`; then the results are effectively the same as `posbinomial` except the binomial constants are not included in the log-likelihood. If `parallel.t = TRUE ~ 0` then each column should have at least one 1 and at least one 0.

It is well-known that some species of animals are affected by capture, e.g., trap-shy or trap-happy. This **VGAM** family function does *not* allow any behavioral effect to be modelled (`posbernoulli.b` and `posbernoulli.tb` do) because the denominator of the likelihood function must be free of behavioral effects.

Value

An object of class "vglmff" (see `vglmff-class`). The object is used by modelling functions such as `vglm`, and `vgam`.

Upon fitting the extra slot has a (list) component called `N.hat` which is a point estimate of the population size N (it is the Horvitz-Thompson (1952) estimator). And there is a component called `SE.N.hat` containing its standard error.

Note

The weights argument of `vglm` need not be assigned, and the default is just a matrix of ones.

Fewer numerical problems are likely to occur for `parallel.t = TRUE`. Data-wise, each sampling occasion may need at least one success (capture) and one failure. Less stringent conditions in the data are needed when `parallel.t = TRUE`. Ditto when parallelism is applied to the intercept too.

The response matrix is returned unchanged; i.e., not converted into proportions like `posbinomial`. If the response matrix has column names then these are used in the labelling, else `prob1`, `prob2`, etc. are used.

Using `AIC()` or `BIC()` to compare `posbernoulli.t`, `posbernoulli.b`, `posbernoulli.tb` models with a `posbinomial` model requires `posbinomial(omit.constant = TRUE)` because one needs to remove the normalizing constant from the log-likelihood function. See `posbinomial` for an example.

Author(s)

Thomas W. Yee.

References

Huggins, R. M. (1991). Some practical aspects of a conditional likelihood approach to capture experiments. *Biometrics*, **47**, 725–732.

Huggins, R. M. and Hwang, W.-H. (2011). A review of the use of conditional likelihood in capture–recapture experiments. *International Statistical Review*, **79**, 385–400.

Otis, D. L. and Burnham, K. P. and White, G. C. and Anderson, D. R. (1978). Statistical inference from capture data on closed animal populations, *Wildlife Monographs*, **62**, 3–135.

Yee, T. W. and Stoklosa, J. and Huggins, R. M. (2015). The **VGAM** package for capture–recapture data using the conditional likelihood. *Journal of Statistical Software*, **65**, 1–33. doi:10.18637/jss.v065.i05.

See Also

[posbernoulli.b](#), [posbernoulli.tb](#), [Select](#), [deermice](#), [Huggins89table1](#), [Huggins89.t1](#), [dposbern](#), [rposbern](#), [posbinomial](#), [AICv1m](#), [BICv1m](#), [prinia](#).

Examples

```
M.t <- vglm(cbind(y1, y2, y3, y4, y5, y6) ~ 1, posbernoulli.t,
           data = deermice, trace = TRUE)
coef(M.t, matrix = TRUE)
constraints(M.t, matrix = TRUE)
summary(M.t, presid = FALSE)

M.h.1 <- vglm(Select(deermice, "y") ~ sex + weight, trace = TRUE,
            posbernoulli.t(parallel.t = FALSE ~ -1), deermice)
coef(M.h.1, matrix = TRUE)
constraints(M.h.1)
summary(M.h.1, presid = FALSE)
head(depvar(M.h.1)) # Response capture history matrix
dim(depvar(M.h.1))

M.th.2 <- vglm(cbind(y1, y2, y3, y4, y5, y6) ~ sex + weight,
            posbernoulli.t(parallel.t = FALSE), deermice)
# Test the parallelism assumption wrt sex and weight:
lrtest(M.h.1, M.th.2)
coef(M.th.2)
coef(M.th.2, matrix = TRUE)
constraints(M.th.2)
summary(M.th.2, presid = FALSE)
head(model.matrix(M.th.2, type = "v1m"), 21)

M.th.2@extra$N.hat # Population size estimate; should be about N
M.th.2@extra$SE.N.hat # SE of the estimate of the population size
# An approximate 95 percent confidence interval:
round(M.th.2@extra$N.hat + c(-1, 1)*1.96* M.th.2@extra$SE.N.hat, 1)
```

```
# Fit a M_h model, effectively the parallel M_t model:
deermice <- transform(deermice, ysum = y1 + y2 + y3 + y4 + y5 + y6,
                      tau = 6)
M.h.3 <- vglm(cbind(ysum, tau - ysum) ~ sex + weight,
             posbinomial(omit.constant = TRUE), data = deermice)
max(abs(coef(M.h.1) - coef(M.h.3))) # Should be zero
# Difference is due to the binomial constants:
logLik(M.h.3) - logLik(M.h.1)
```

 posbernoulli.tb

Positive Bernoulli Family Function with Time and Behavioural Effects

Description

Fits a GLM/GAM-like model to multiple Bernoulli responses where each row in the capture history matrix response has at least one success (capture). Sampling occasion effects and behavioural effects are accommodated.

Usage

```
posbernoulli.tb(link = "logitlink", parallel.t = FALSE ~ 1,
               parallel.b = FALSE ~ 0, drop.b = FALSE ~ 1,
               type.fitted = c("likelihood.cond", "mean.uncond"),
               imethod = 1, iprob = NULL,
               p.small = 1e-4, no.warning = FALSE,
               ridge.constant = 0.0001, ridge.power = -4)
```

Arguments

link, imethod, iprob

See [CommonVGAMffArguments](#) for information.

parallel.t, parallel.b, drop.b

A logical, or formula with a logical as the response. See [CommonVGAMffArguments](#) for information. The parallel.-type arguments specify whether the constraint matrices have a parallelism assumption for the temporal and behavioural effects. Argument parallel.t means parallel with respect to time, and matches the same argument name in [posbernoulli.t](#).

Suppose the model is intercept-only. Setting parallel.t = FALSE ~ 0 results in the M_b model. Setting drop.b = FALSE ~ 0 results in the M_t model because it drops columns off the constraint matrices corresponding to any behavioural effect. Setting parallel.t = FALSE ~ 0 and setting parallel.b = FALSE ~ 0 results in the M_b model. Setting parallel.t = FALSE ~ 0, parallel.b = FALSE ~ 0 and drop.b = FALSE ~ 0 results in the M_0 model. Note the default for parallel.t and parallel.b may be unsuitable for data sets which have a large τ because of the large number of parameters; it might be too flexible. If it is desired to have the behaviour affect some of the other covariates then set drop.b = TRUE ~ 0.

The default model has a different intercept for each sampling occasion, a time-parallelism assumption for all other covariates, and a dummy variable representing a single behavioural effect (also in the intercept).

The most flexible model is to set `parallel.b = TRUE ~ 0`, `parallel.t = TRUE ~ 0` and `drop.b = TRUE ~ 0`. This means that all possible temporal and behavioural effects are estimated, for the intercepts and other covariates. Such a model is *not* recommended; it will contain a lot of parameters.

- `type.fitted` Character, one of the choices for the type of fitted value returned. The default is the first one. Partial matching is okay. For "likelihood.cond": the probability defined by the conditional likelihood. For "mean.uncond": the unconditional mean, which should agree with `colMeans` applied to the response matrix for intercept-only models.
- `ridge.constant`, `ridge.power` Determines the ridge parameters at each IRLS iteration. They are the constant and power (exponent) for the ridge adjustment for the working weight matrices (the capture probability block matrix, hence the first τ diagonal values). At iteration a of the IRLS algorithm a positive value is added to the first τ diagonal elements of the working weight matrices to make them positive-definite. This adjustment is the mean of the diagonal elements of wz multiplied by $K \times a^p$ where K is `ridge.constant` and p is `ridge.power`. This is always positive but decays to zero as iterations proceed (provided p is negative etc.).
- `p.small`, `no.warning` See [posbernoulli.t](#).

Details

This model (commonly known as M_{tb}/M_{tbb} in the capture–recapture literature) operates on a response matrix of 0s and 1s ($n \times \tau$). See [posbernoulli.t](#) for information that is in common. It allows time and behavioural effects to be modelled.

Evidently, the expected information matrix (EIM) seems *not* of full rank (especially in early iterations), so `ridge.constant` and `ridge.power` are used to *try* fix up the problem. The default link functions are $(\text{logit } p_{c1}, \dots, \text{logit } p_{c\tau}, \text{logit } p_{r2}, \dots, \text{logit } p_{r\tau})^T$ where the subscript c denotes capture, the subscript r denotes recapture, and it is not possible to recapture the animal at sampling occasion 1. Thus $M = 2\tau - 1$. The parameters are currently prefixed by `pcapture` and `precapture` for the capture and recapture probabilities. This **VGAM** family function may be further modified in the future.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as `vglm`, and `vgam`.

Note

It is a good idea to apply the parallelism assumption to each sampling occasion except possibly with respect to the intercepts. Also, a simple behavioural effect such as being modelled using the intercept is recommended; if the behavioural effect is not parallel and/or allowed to apply to other

covariates then there will probably be too many parameters, and hence, numerical problems. See `M_tbh.1` below.

It is a good idea to monitor convergence. Simpler models such as the M_0/M_h models are best fitted with [posbernoulli.t](#) or [posbernoulli.b](#) or [posbinomial](#).

Author(s)

Thomas W. Yee.

References

See [posbernoulli.t](#).

See Also

[posbernoulli.b](#) (including `N.hat`), [posbernoulli.t](#), [posbinomial](#), [Select](#), [fill1](#), [Huggins89table1](#), [Huggins89.t1](#), [deermice](#), [prinia](#).

Examples

```
## Not run:
# Example 1: simulated data
nTimePts <- 5 # (aka tau == # of sampling occasions)
nnn <- 1000 # Number of animals
pdata <- rposbern(n = nnn, nTimePts = nTimePts, pvars = 2)
dim(pdata); head(pdata)

M_tbh.1 <- vglm(cbind(y1, y2, y3, y4, y5) ~ x2,
               posbernoulli.tb, data = pdata, trace = TRUE)
coef(M_tbh.1) # First element is the behavioural effect
coef(M_tbh.1, matrix = TRUE)
constraints(M_tbh.1, matrix = TRUE)
summary(M_tbh.1, resid = FALSE) # Std errors are approximate
head(fitted(M_tbh.1))
head(model.matrix(M_tbh.1, type = "v1m"), 21)
dim(depvar(M_tbh.1))

M_tbh.2 <- vglm(cbind(y1, y2, y3, y4, y5) ~ x2,
               posbernoulli.tb(parallel.t = FALSE ~ 0),
               data = pdata, trace = TRUE)
coef(M_tbh.2) # First element is the behavioural effect
coef(M_tbh.2, matrix = TRUE)
constraints(M_tbh.2, matrix = TRUE)
summary(M_tbh.2, resid = FALSE) # Std errors are approximate
head(fitted(M_tbh.2))
head(model.matrix(M_tbh.2, type = "v1m"), 21)
dim(depvar(M_tbh.2))

# Example 2: deermice subset data
fit1 <- vglm(cbind(y1, y2, y3, y4, y5, y6) ~ sex + weight,
             posbernoulli.t, data = deermice, trace = TRUE)
coef(fit1)
```



```

coef(fit1, matrix = TRUE)
constraints(fit1, matrix = TRUE)
summary(fit1, presid = FALSE) # Standard errors are approximate

# fit1 is the same as Fit1 (a M_{th} model):
Fit1 <- vglm(cbind(y1, y2, y3, y4, y5, y6) ~ sex + weight,
            posbernoulli.tb(drop.b = TRUE ~ sex + weight,
                           parallel.t = TRUE), # But not for the intercept
            data = deermice, trace = TRUE)
constraints(Fit1)

## End(Not run)

```

posbernUC

Positive Bernoulli Sequence Model

Description

Density, and random generation for multiple Bernoulli responses where each row in the response matrix has at least one success.

Usage

```

rposbern(n, nTimePts = 5, pvars = length(xcoeff),
         xcoeff = c(-2, 1, 2), Xmatrix = NULL, cap.effect = 1,
         is.popn = FALSE, link = "logitlink", earg.link = FALSE)
dposbern(x, prob, prob0 = prob, log = FALSE)

```

Arguments

x	response vector or matrix. Should only have 0 and 1 values, at least two columns, and each row should have at least one 1.
nTimePts	Number of sampling occasions. Called τ in posbernoulli.b and posbernoulli.t .
n	number of observations. Usually a single positive integer, else the length of the vector is used. See argument <code>is.popn</code> .
is.popn	Logical. If TRUE then argument <code>n</code> is the population size and what is returned may have substantially less rows than <code>n</code> . That is, if an animal has at least one one in its sequence then it is returned, else that animal is not returned because it never was captured.
Xmatrix	Optional X matrix. If given, the X matrix is not generated internally.
cap.effect	Numeric, the capture effect. Added to the linear predictor if captured previously. A positive or negative value corresponds to a trap-happy and trap-shy effect respectively.
pvars	Number of other numeric covariates that make up the linear predictor. Labelled <code>x1, x2, ...</code> , where the first is an intercept, and the others are independent standard runif random variates. The first <code>pvars</code> elements of <code>xcoeff</code> are used.

xcoeff	The regression coefficients of the linear predictor. These correspond to x_1 , x_2 , ..., and the first is for the intercept. The length of xcoeff must be at least pvars.
link, earg.link	The former is used to generate the probabilities for capture at each occasion. Other details at CommonVGAMffArguments .
prob, prob0	Matrix of probabilities for the numerator and denominators respectively. The default does <i>not</i> correspond to the M_b model since the M_b model has a denominator which involves the capture history.
log	Logical. Return the logarithm of the answer?

Details

The form of the conditional likelihood is described in [posbernoulli.b](#) and/or [posbernoulli.t](#) and/or [posbernoulli.tb](#). The denominator is equally shared among the elements of the matrix x .

Value

rposbern returns a data frame with some attributes. The function generates random deviates (τ columns labelled y_1 , y_2 , ...) for the response. Some indicator columns are also included (those starting with ch are for previous capture history). The default setting corresponds to a M_{bh} model that has a single trap-happy effect. Covariates x_1 , x_2 , ... have the same affect on capture/recapture at every sampling occasion (see the argument parallel.t in, e.g., [posbernoulli.tb](#)).

The function dposbern gives the density,

Note

The r-type function is experimental only and does not follow the usual conventions of r-type R functions. It may change a lot in the future. The d-type function is more conventional and is less likely to change.

Author(s)

Thomas W. Yee.

See Also

[posbernoulli.tb](#), [posbernoulli.b](#), [posbernoulli.t](#).

Examples

```
rposbern(n = 10)
attributes(pdata <- rposbern(n = 100))
M.bh <- vglm(cbind(y1, y2, y3, y4, y5) ~ x2 + x3,
             posbernoulli.b(I2 = FALSE), pdata, trace = TRUE)
constraints(M.bh)
summary(M.bh)
```

posbinomial *Positive Binomial Distribution Family Function*

Description

Fits a positive binomial distribution.

Usage

```
posbinomial(link = "logitlink", multiple.responses = FALSE,
            parallel = FALSE, omit.constant = FALSE, p.small = 1e-4,
            no.warning = FALSE, zero = NULL)
```

Arguments

`link`, `multiple.responses`, `parallel`, `zero`
 Details at [CommonVGAMffArguments](#).

`omit.constant` Logical. If TRUE then the constant (`lchoose(size, size * yprop)`) is omitted from the loglikelihood calculation. If the model is to be compared using `AIC()` or `BIC()` (see [AICv1m](#) or [BICv1m](#)) to the likes of [posbernoulli.tb](#) etc. then it is important to set `omit.constant = TRUE` because all models then will not have any normalizing constants in the likelihood function. Hence they become comparable. This is because the M_0 Otis et al. (1978) model coincides with `posbinomial()`. See below for an example. Also see [posbernoulli.t](#) regarding estimating the population size (`N.hat` and `SE.N.hat`) if the number of trials is the same for all observations.

`p.small`, `no.warning`
 See [posbernoulli.t](#).

Details

The positive binomial distribution is the ordinary binomial distribution but with the probability of zero being zero. Thus the other probabilities are scaled up (i.e., divided by $1 - P(Y = 0)$). The fitted values are the ordinary binomial distribution fitted values, i.e., the usual mean.

In the capture–recapture literature this model is called the M_0 if it is an intercept-only model. Otherwise it is called the M_h when there are covariates. It arises from a sum of a sequence of τ -Bernoulli random variates subject to at least one success (capture). Here, each animal has the same probability of capture or recapture, regardless of the τ sampling occasions. Independence between animals and between sampling occasions etc. is assumed.

Value

An object of class "vg1mff" (see [vg1mff-class](#)). The object is used by modelling functions such as `vg1m`, and `vgam`.

Warning

Under- or over-flow may occur if the data is ill-conditioned.

Note

The input for this family function is the same as [binomialff](#).

If `multiple.responses = TRUE` then each column of the matrix `response` should be a count (the number of successes), and the `weights` argument should be a matrix of the same dimension as the `response` containing the number of trials. If `multiple.responses = FALSE` then the `response` input should be the same as [binomialff](#).

Yet to be done: a `quasi.posbinomial()` which estimates a dispersion parameter.

Author(s)

Thomas W. Yee

References

Otis, D. L. et al. (1978). Statistical inference from capture data on closed animal populations, *Wildlife Monographs*, **62**, 3–135.

Patil, G. P. (1962). Maximum likelihood estimation for generalised power series distributions and its application to a truncated binomial distribution. *Biometrika*, **49**, 227–237.

Pearson, K. (1913). *A Monograph on Albinism in Man*. Drapers Company Research Memoirs.

See Also

[posbernoulli.b](#), [posbernoulli.t](#), [posbernoulli.tb](#), [binomialff](#), [AICvlm](#), [BICvlm](#), [simulate.vlm](#).

Examples

```
# Albinotic children in families with 5 kids (from Patil, 1962) , , , ,
albinos <- data.frame(y = c(rep(1, 25), rep(2, 23), rep(3, 10), 4, 5),
                    n = rep(5, 60))
fit1 <- vglm(cbind(y, n-y) ~ 1, posbinomial, albinos, trace = TRUE)
summary(fit1)
Coef(fit1) # = MLE of p = 0.3088
head(fitted(fit1))
sqrt(vcov(fit1, untransform = TRUE)) # SE = 0.0322

# Fit a M_0 model (Otis et al. 1978) to the deermice data , , , , , , , ,
M.0 <- vglm(cbind( y1 + y2 + y3 + y4 + y5 + y6,
                  6 - y1 - y2 - y3 - y4 - y5 - y6) ~ 1, trace = TRUE,
            posbinomial(omit.constant = TRUE), data = deermice)
coef(M.0, matrix = TRUE)
Coef(M.0)
constraints(M.0, matrix = TRUE)
summary(M.0)
c( N.hat = M.0@extra$N.hat, # As tau = 6, i.e., 6 Bernoulli trials
  SE.N.hat = M.0@extra$SE.N.hat) # per obsn is the same for each obsn

# Compare it to the M_b using AIC and BIC
M.b <- vglm(cbind(y1, y2, y3, y4, y5, y6) ~ 1, trace = TRUE,
            posbernoulli.b, data = deermice)
```

```
sort(c(M.0 = AIC(M.0), M.b = AIC(M.b))) # Ok since omit.constant=TRUE
sort(c(M.0 = BIC(M.0), M.b = BIC(M.b))) # Ok since omit.constant=TRUE
```

Posgeom

Positive-Geometric Distribution

Description

Density, distribution function, quantile function and random generation for the positive-geometric distribution.

Usage

```
dposgeom(x, prob, log = FALSE)
pposgeom(q, prob)
qposgeom(p, prob)
rposgeom(n, prob)
```

Arguments

x, q	vector of quantiles.
p	vector of probabilities.
n	number of observations. Fed into runif .
prob	vector of probabilities of success (of an ordinary geometric distribution). Short vectors are recycled.
log	logical.

Details

The positive-geometric distribution is a geometric distribution but with the probability of a zero being zero. The other probabilities are scaled to add to unity. The mean therefore is $1/prob$.

As *prob* decreases, the positive-geometric and geometric distributions become more similar. Like similar functions for the geometric distribution, a zero value of *prob* is not permitted here.

Value

`dposgeom` gives the density, `pposgeom` gives the distribution function, `qposgeom` gives the quantile function, and `rposgeom` generates random deviates.

Author(s)

T. W. Yee

See Also

[zageometric](#), [zigeometric](#), [rgeom](#).

Examples

```

prob <- 0.75; y <- rposgeom(n = 1000, prob)
table(y)
mean(y) # Sample mean
1 / prob # Population mean

(ii <- dposgeom(0:7, prob))
cumsum(ii) - pposgeom(0:7, prob) # Should be 0s
table(rposgeom(100, prob))

table(qposgeom(runif(1000), prob))
round(dposgeom(1:10, prob) * 1000) # Should be similar

## Not run:
x <- 0:5
barplot(rbind(dposgeom(x, prob), dgeom(x, prob)),
        beside = TRUE, col = c("blue", "orange"),
        main = paste("Positive geometric(", prob, ") (blue) vs",
                    "geometric(", prob, ") (orange)", sep = ""),
        names.arg = as.character(x), las = 1, lwd = 2)
## End(Not run)

```

posnegbinomial

Positive Negative Binomial Distribution Family Function

Description

Maximum likelihood estimation of the two parameters of a positive negative binomial distribution.

Usage

```

posnegbinomial(zero = "size",
               type.fitted = c("mean", "munb", "prob0"),
               mds.min = 0.001, nsimEIM = 500, cutoff.prob = 0.999,
               eps.trig = 1e-07, max.support = 4000, max.chunk.MB = 30,
               lmunb = "loglink", lsize = "loglink", imethod = 1,
               imunb = NULL, iprobs.y = NULL,
               gprobs.y = ppoints(8), isize = NULL,
               gsize.mux = exp(c(-30, -20, -15, -10, -6:3)))

```

Arguments

lmunb	Link function applied to the munb parameter, which is the mean μ_{nb} of an ordinary negative binomial distribution. See Links for more choices.
lsize	Parameter link function applied to the dispersion parameter, called k. See Links for more choices.

isize	Optional initial value for k , an index parameter. The value $1/k$ is known as a dispersion parameter. If failure to converge occurs try different values (and/or use <code>imethod</code>). If necessary this vector is recycled to length equal to the number of responses. A value NULL means an initial value for each response is computed internally using a range of values.
nsimEIM, zero, eps.trig	See CommonVGAMffArguments .
mds.min, iprobs.y, cutoff.prob	Similar to negbinomial .
imunb, max.support	Similar to negbinomial .
max.chunk.MB, gsize.mux	Similar to negbinomial .
imethod, gprobs.y	See negbinomial .
type.fitted	See CommonVGAMffArguments for details.

Details

The positive negative binomial distribution is an ordinary negative binomial distribution but with the probability of a zero response being zero. The other probabilities are scaled to sum to unity.

This family function is based on [negbinomial](#) and most details can be found there. To avoid confusion, the parameter `munb` here corresponds to the mean of an ordinary negative binomial distribution [negbinomial](#). The mean of `posnegbinomial` is

$$\mu_{nb}/(1 - p(0))$$

where $p(0) = (k/(k + \mu_{nb}))^k$ is the probability an ordinary negative binomial distribution has a zero value.

The parameters `munb` and `k` are not independent in the positive negative binomial distribution, whereas they are in the ordinary negative binomial distribution.

This function handles *multiple* responses, so that a matrix can be used as the response. The number of columns is the number of species, say, and setting `zero = -2` means that *all* species have a `k` equalling a (different) intercept only.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), [rrvglm](#) and [vgam](#).

Warning

This family function is fragile; at least two cases will lead to numerical problems. Firstly, the positive-Poisson model corresponds to `k` equalling infinity. If the data is positive-Poisson or close to positive-Poisson, then the estimated `k` will diverge to `Inf` or some very large value. Secondly, if the data is clustered about the value 1 because the `munb` parameter is close to 0 then numerical problems will also occur. Users should set `trace = TRUE` to monitor convergence. In the situation

when both cases hold, the result returned (which will be untrustworthy) will depend on the initial values.

The negative binomial distribution (NBD) is a strictly unimodal distribution. Any data set that does not exhibit a mode (in the middle) makes the estimation problem difficult. The positive NBD inherits this feature. Set `trace = TRUE` to monitor convergence.

See the example below of a data set where `posbinomial()` fails; the so-called solution is *extremely* poor. This is partly due to a lack of a unimodal shape because the number of counts decreases only. This long tail makes it very difficult to estimate the mean parameter with any certainty. The result too is that the size parameter is numerically fraught.

This **VGAM** family function inherits the same warnings as [negbinomial](#). And if k is much less than 1 then the estimation may be slow.

Note

If the estimated k is very large then fitting a [pospoisson](#) model is a good idea.

If both `munb` and k are large then it may be necessary to decrease `eps.trig` and increase `max.support` so that the EIMs are positive-definite, e.g., `eps.trig = 1e-8` and `max.support = Inf`.

Author(s)

Thomas W. Yee

References

- Barry, S. C. and Welsh, A. H. (2002). Generalized additive modelling and zero inflated count data. *Ecological Modelling*, **157**, 179–188.
- Williamson, E. and Bretherton, M. H. (1964). Tables of the logarithmic series distribution. *Annals of Mathematical Statistics*, **35**, 284–297.

See Also

[gaitdnbinomial](#), [pospoisson](#), [negbinomial](#), [zanegbinomial](#), [rnbinom](#), [CommonVGAMffArguments](#), [corbet](#), [logff](#), [simulate.vlm](#), [margeff](#).

Examples

```
## Not run:
pdata <- data.frame(x2 = runif(nn <- 1000))
pdata <- transform(pdata,
  y1 = rgaitdnbinom(nn, exp(1), munb.p = exp(0+2*x2), truncate = 0),
  y2 = rgaitdnbinom(nn, exp(3), munb.p = exp(1+2*x2), truncate = 0))
fit <- vglm(cbind(y1, y2) ~ x2, posnegbinomial, pdata, trace = TRUE)
coef(fit, matrix = TRUE)
dim(depvar(fit)) # Using dim(fit@y) is not recommended

# Another artificial data example
pdata2 <- data.frame(munb = exp(2), size = exp(3)); nn <- 1000
pdata2 <- transform(pdata2,
```



```

y3 = rgaitdnbinom(nn, size, munb.p = munb,
                  truncate = 0))
with(pdata2, table(y3))
fit <- vglm(y3 ~ 1, posnegbinomial, data = pdata2, trace = TRUE)
coef(fit, matrix = TRUE)
with(pdata2, mean(y3)) # Sample mean
head(with(pdata2, munb/(1-(size/(size+munb))^size)), 1) # Popn mean
head(fitted(fit), 3)
head(predict(fit), 3)

# Example: Corbet (1943) butterfly Malaya data
fit <- vglm(ofreq ~ 1, posnegbinomial, weights = species, corbet)
coef(fit, matrix = TRUE)
Coef(fit)
(khat <- Coef(fit)["size"])
pdf2 <- dgaitdnbinom(with(corbet, ofreq), khat,
                    munb.p = fitted(fit), truncate = 0)
print(with(corbet,
           cbind(ofreq, species, fitted = pdf2*sum(species))), dig = 1)
with(corbet,
matplot(ofreq, cbind(species, fitted = pdf2*sum(species)), las = 1,
        xlab = "Observed frequency (of individual butterflies)",
        type = "b", ylab = "Number of species", col = c("blue", "orange"),
        main = "blue 1s = observe; orange 2s = fitted"))

# Data courtesy of Maxim Gerashchenko causes posbinomial() to fail
pnbd.fail <- data.frame(
y1 = c(1:16, 18:21, 23:28, 33:38, 42, 44, 49:51, 55, 56, 58,
59, 61:63, 66, 73, 76, 94, 107, 112, 124, 190, 191, 244),
ofreq = c(130, 80, 38, 23, 22, 11, 21, 14, 6, 7, 9, 9, 9, 4, 4, 5, 1,
4, 6, 1, 3, 2, 4, 3, 4, 5, 3, 1, 2, 1, 1, 4, 1, 2, 2, 1, 3,
1, 1, 2, 2, 2, 1, 3, 2, 1, 1, 1, 1, 1, 1, 1, 1))
fit.fail <- vglm(y1 ~ 1, weights = ofreq, posnegbinomial,
                trace = TRUE, data = pnbd.fail)

## End(Not run)

```

Description

Density, distribution function, quantile function and random generation for the univariate positive-normal distribution.

Usage

```

dposnorm(x, mean = 0, sd = 1, log = FALSE)
pposnorm(q, mean = 0, sd = 1, lower.tail = TRUE, log.p = FALSE)

```

```
qposnorm(p, mean = 0, sd = 1, lower.tail = TRUE, log.p = FALSE)
rposnorm(n, mean = 0, sd = 1)
```

Arguments

`x, q` vector of quantiles.
`p` vector of probabilities.
`n` number of observations. If `length(n) > 1` then the length is taken to be the number required.
`mean, sd, log, lower.tail, log.p`
 see [rnorm](#).

Details

See [posnormal](#), the **VGAM** family function for estimating the parameters, for the formula of the probability density function and other details.

Value

`dposnorm` gives the density, `pposnorm` gives the distribution function, `qposnorm` gives the quantile function, and `rposnorm` generates random deviates.

Author(s)

T. W. Yee

See Also

[posnormal](#).

Examples

```
## Not run: m <- 0.8; x <- seq(-1, 4, len = 501)
plot(x, dposnorm(x, m = m), type = "l", las = 1, ylim = 0:1,
      ylab = paste("posnorm(m = ", m, ", sd = 1)"), col = "blue",
      main = "Blue is density, orange is the CDF",
      sub = "Purple lines are the 10,20,...,90 percentiles")
abline(h = 0, col = "grey")
lines(x, pposnorm(x, m = m), col = "orange", type = "l")
probs <- seq(0.1, 0.9, by = 0.1)
Q <- qposnorm(probs, m = m)
lines(Q, dposnorm(Q, m = m), col = "purple", lty = 3, type = "h")
lines(Q, pposnorm(Q, m = m), col = "purple", lty = 3, type = "h")
abline(h = probs, col = "purple", lty = 3)
max(abs(pposnorm(Q, m = m) - probs)) # Should be 0

## End(Not run)
```

posnormal

Positive Normal Distribution Family Function

Description

Fits a positive (univariate) normal distribution.

Usage

```
posnormal(lmean = "identitylink", lsd = "loglink",
          eq.mean = FALSE, eq.sd = FALSE,
          gmean = exp((-5:5)/2), gsd = exp((-1:5)/2),
          imean = NULL, isd = NULL, probs.y = 0.10, imethod = 1,
          nsimEIM = NULL, zero = "sd")
```

Arguments

`lmean, lsd` Link functions for the mean and standard deviation parameters of the usual univariate normal distribution. They are μ and σ respectively. See [Links](#) for more choices.

`gmean, gsd, imethod` See [CommonVGAMffArguments](#) for more information. `gmean` and `gsd` currently operate on a multiplicative scale, on the sample mean and the sample standard deviation, respectively.

`imean, isd` Optional initial values for μ and σ . A NULL means a value is computed internally. See [CommonVGAMffArguments](#) for more information.

`eq.mean, eq.sd` See [CommonVGAMffArguments](#) for more information. The fact that these arguments are supported results in default constraint matrices being a *permutation* of the identity matrix (effectively *trivial* constraints).

`zero, nsimEIM, probs.y` See [CommonVGAMffArguments](#) for information.

Details

The positive normal distribution is the ordinary normal distribution but with the probability of zero or less being zero. The rest of the probability density function is scaled up. Hence the probability density function can be written

$$f(y) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{1}{2}(y - \mu)^2/\sigma^2\right) / [1 - \Phi(-\mu/\sigma)]$$

where $\Phi()$ is the cumulative distribution function of a standard normal ([pnorm](#)). Equivalently, this is

$$f(y) = \frac{1}{\sigma} \frac{\phi((y - \mu)/\sigma)}{1 - \Phi(-\mu/\sigma)}$$

where $\phi()$ is the probability density function of a standard normal distribution ([dnorm](#)).

The mean of Y is

$$E(Y) = \mu + \sigma \frac{\phi(-\mu/\sigma)}{1 - \Phi(-\mu/\sigma)}.$$

This family function handles multiple responses.

Value

An object of class "vglmfff" (see [vglmfff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Warning

It is recommended that `trace = TRUE` be used to monitor convergence; sometimes the estimated mean is `-Inf` and the estimated mean standard deviation is `Inf`, especially when the sample size is small. Under- or over-flow may occur if the data is ill-conditioned.

Note

The response variable for this family function is the same as [uninormal](#) except positive values are required. Reasonably good initial values are needed.

The distribution of the reciprocal of a positive normal random variable is known as an alpha distribution.

Author(s)

Thomas W. Yee

See Also

[uninormal](#), [tobit](#).

Examples

```
pdata <- data.frame(Mean = 1.0, SD = exp(1.0))
pdata <- transform(pdata, y = rposnorm(n <- 1000, m = Mean, sd = SD))

## Not run: with(pdata, hist(y, prob = TRUE, border = "blue",
  main = paste("posnorm(m =", Mean[1], ", sd =", round(SD[1], 2), ")"))))
## End(Not run)
fit <- vglm(y ~ 1, posnormal, data = pdata, trace = TRUE)
coef(fit, matrix = TRUE)
(Cfit <- Coef(fit))
mygrid <- with(pdata, seq(min(y), max(y), len = 200))
## Not run: lines(mygrid, dposnorm(mygrid, Cfit[1], Cfit[2]), col = "red")
```

 pospoisson

Positive Poisson Distribution Family Function

Description

Fits a positive Poisson distribution.

Usage

```
pospoisson(link = "loglink", type.fitted = c("mean", "lambda",
      "prob0"), expected = TRUE, ilambda = NULL, imethod = 1,
      zero = NULL, gt.1 = FALSE)
```

Arguments

link	Link function for the usual mean (lambda) parameter of an ordinary Poisson distribution. See Links for more choices.
expected	Logical. Fisher scoring is used if expected = TRUE, else Newton-Raphson.
ilambda, imethod, zero	See CommonVGAMffArguments for information.
type.fitted	See CommonVGAMffArguments for details.
gt.1	Logical. Enforce lambda > 1? The default is to enforce lambda > 0.

Details

The positive Poisson distribution is the ordinary Poisson distribution but with the probability of zero being zero. Thus the other probabilities are scaled up (i.e., divided by $1 - P[Y = 0]$). The mean, $\lambda/(1 - \exp(-\lambda))$, can be obtained by the extractor function `fitted` applied to the object.

A related distribution is the zero-inflated Poisson, in which the probability $P[Y = 0]$ involves another parameter ϕ . See [zipoisson](#).

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), [rrvglm](#) and [vgam](#).

Warning

Under- or over-flow may occur if the data is ill-conditioned.

Note

This family function can handle multiple responses.

Yet to be done: a quasi .pospoisson which estimates a dispersion parameter.

Author(s)

Thomas W. Yee

References

Coleman, J. S. and James, J. (1961). The equilibrium size distribution of freely-forming groups. *Sociometry*, **24**, 36–45.

See Also

[Gaitdpois](#), [gaitdpoisson](#), [posnegbinomial](#), [poissonff](#), [zapoisson](#), [zipoisson](#), [simulate.vlm](#), [otpospoisson](#), [Pospois](#).

Examples

```
# Data from Coleman and James (1961)
cjdata <- data.frame(y = 1:6, freq = c(1486, 694, 195, 37, 10, 1))
fit <- vglm(y ~ 1, pospoisson, data = cjdata, weights = freq)
Coef(fit)
summary(fit)
fitted(fit)

pdata <- data.frame(x2 = runif(nn <- 1000)) # Artificial data
pdata <- transform(pdata, lambda = exp(1 - 2 * x2))
pdata <- transform(pdata, y1 = rgaitdpois(nn, lambda, truncate = 0))
with(pdata, table(y1))
fit <- vglm(y1 ~ x2, pospoisson, data = pdata, trace = TRUE, crit = "coef")
coef(fit, matrix = TRUE)
```

powerlink

Power Link Function

Description

Computes the power transformation, including its inverse and the first two derivatives.

Usage

```
powerlink(theta, power = 1, inverse = FALSE, deriv = 0,
          short = TRUE, tag = FALSE)
```

Arguments

theta Numeric or character. See below for further details.
power This denotes the power or exponent.
inverse, deriv, short, tag
 Details at [Links](#).

Details

The power link function raises a parameter by a certain value of power. Care is needed because it is very easy to get numerical problems, e.g., if $\text{power} = 0.5$ and θ is negative.

Value

For `powerlink` with `deriv = 0`, then θ raised to the power of `power`. And if `inverse = TRUE` then θ raised to the power of $1/\text{power}$.

For `deriv = 1`, then the function returns $d\theta/d\eta$ as a function of θ if `inverse = FALSE`, else if `inverse = TRUE` then it returns the reciprocal.

Note

Numerical problems may occur for certain combinations of θ and `power`. Consequently this link function should be used with caution.

Author(s)

Thomas W. Yee

See Also

[Links](#), [loglink](#).

Examples

```
powerlink("a", power = 2, short = FALSE, tag = TRUE)
powerlink(x <- 1:5)
powerlink(x, power = 2)
max(abs(powerlink(powerlink(x, power = 2),
                  power = 2, inverse = TRUE) - x)) # Should be 0
powerlink(x <- (-5):5, power = 0.5) # Has NAs

# 1/2 = 0.5
pdata <- data.frame(y = rbeta(n = 1000, shape1 = 2^2, shape2 = 3^2))
fit <- vglm(y ~ 1, betaR(lshape1 = powerlink(power = 0.5), i1 = 3,
                    lshape2 = powerlink(power = 0.5), i2 = 7), data = pdata)
t(coef(fit, matrix = TRUE))
Coef(fit) # Useful for intercept-only models
vcov(fit, untransform = TRUE)
```

prats

Pregnant Rats Toxicological Experiment Data

Description

A small toxicological experiment data. The subjects are fetuses from two randomized groups of pregnant rats, and they were given a placebo or chemical treatment. The number with birth defects were recorded, as well as each litter size.

Usage

```
data(prats)
```

Format

A data frame with the following variables.

treatment A 0 means control; a 1 means the chemical treatment.

alive, litter.size The number of fetuses alive at 21 days, out of the number of fetuses alive at 4 days (the litter size).

Details

The data concerns a toxicological experiment where the subjects are fetuses from two randomized groups of 16 pregnant rats each, and they were given a placebo or chemical treatment. The number with birth defects and the litter size were recorded. Half the rats were fed a control diet during pregnancy and lactation, and the diet of the other half was treated with a chemical. For each litter the number of pups alive at 4 days and the number of pups that survived the 21 day lactation period, were recorded.

Source

Weil, C. S. (1970) Selection of the valid number of sampling units and a consideration of their combination in toxicological studies involving reproduction, teratogenesis or carcinogenesis. *Food and Cosmetics Toxicology*, **8**(2), 177–182.

References

Williams, D. A. (1975). The Analysis of Binary Responses From Toxicological Experiments Involving Reproduction and Teratogenicity. *Biometrics*, **31**(4), 949–952.

See Also

[betabinomial](#), [betabinomialff](#).

Examples

```
prats
colSums(subset(prats, treatment == 0))
colSums(subset(prats, treatment == 1))
summary(prats)
```

predictqrrvglm *Predict Method for a CQO fit*

Description

Predicted values based on a constrained quadratic ordination (CQO) object.

Usage

```
predictqrrvglm(object, newdata = NULL,
  type = c("link", "response", "latvar", "terms"),
  se.fit = FALSE, deriv = 0, dispersion = NULL,
  extra = object@extra, varI.latvar = FALSE, refResponse = NULL, ...)
```

Arguments

object	Object of class inheriting from "qrrvglm".
newdata	An optional data frame in which to look for variables with which to predict. If omitted, the fitted linear predictors are used.
type, se.fit, dispersion, extra	See predictvglm .
deriv	Derivative. Currently only 0 is handled.
varI.latvar, refResponse	Arguments passed into Coef.qrrvglm .
...	Currently undocumented.

Details

Obtains predictions from a fitted CQO object. Currently there are lots of limitations of this function; it is unfinished.

Value

See [predictvglm](#).

Note

This function is not robust and has not been checked fully.

Author(s)

T. W. Yee

References

Yee, T. W. (2004). A new technique for maximum-likelihood canonical Gaussian ordination. *Ecological Monographs*, **74**, 685–701.

See Also

[cqo](#), [calibrate.qrrvglm](#).

Examples

```
## Not run: set.seed(1234)
hspider[, 1:6] <- scale(hspider[, 1:6]) # Standardize the X vars
p1 <- cqo(cbind(Alopacce, Alopcone, Alopfabr, Arctlute,
               Arctperi, Auloalbi, Pardlugu, Pardmont,
               Pardnigr, Pardpull, Trocterr, Zoraspin) ~
          WaterCon + BareSand + FallTwig + CoveMoss + CoveHerb + ReflLux,
          poissonff, data = hspider, Crowlpositive = FALSE, I.toler = TRUE)
sort(deviance(p1, history = TRUE)) # A history of all the iterations
head(predict(p1))

# The following should be all 0s:
max(abs(predict(p1, newdata = head(hspider)) - head(predict(p1))))
max(abs(predict(p1, newdata = head(hspider), type = "res")-head(fitted(p1))))

## End(Not run)
```

predictvglm

Predict Method for a VGLM fit

Description

Predicted values based on a vector generalized linear model (VGLM) object.

Usage

```
predictvglm(object, newdata = NULL,
            type = c("link", "response", "terms"),
            se.fit = FALSE, deriv = 0, dispersion = NULL,
            untransform = FALSE,
            type.fitted = NULL, percentiles = NULL, ...)
```

Arguments

object	Object of class inheriting from "vglm", e.g., vglm .
newdata	An optional data frame in which to look for variables with which to predict. If omitted, the fitted linear predictors are used.
type	The value of this argument can be abbreviated. The type of prediction required. The default is the first one, meaning on the scale of the linear predictors. This should be a $n \times M$ matrix. The alternative "response" is on the scale of the response variable, and depending on the family function, this may or may not be the mean. Often this is the fitted value, e.g., <code>fitted(vglmObject)</code> (see fittedvglm). Note that the

	response is output from the @linkinv slot, where the eta argument is the $n \times M$ matrix of linear predictors.
	The "terms" option returns a matrix giving the fitted values of each term in the model formula on the linear predictor scale. The terms have been centered.
se.fit	logical: return standard errors?
deriv	Non-negative integer. Currently this must be zero. Later, this may be implemented for general values.
dispersion	Dispersion parameter. This may be inputted at this stage, but the default is to use the dispersion parameter of the fitted model.
type.fitted	Some VGAM family functions have an argument by the same name. If so, then one can obtain fitted values by setting type = "response" and choosing a value of type.fitted from what's available. If type.fitted = "quantiles" is available then the percentiles argument can be used to specify what quantile values are requested.
percentiles	Used only if type.fitted = "quantiles" is available and is selected.
untransform	Logical. Reverses any parameter link function. This argument only works if type = "link", se.fit = FALSE, deriv = 0. Setting untransform = TRUE does not work for all VGAM family functions; only ones where there is a one-to-one correspondence between a simple link function and a simple parameter might work.
...	Arguments passed into predictvglm.

Details

Obtains predictions and optionally estimates standard errors of those predictions from a fitted `vglm` object. By default, each row of the matrix returned can be written as η_i^T , comprising of M components or linear predictors. If there are any offsets, these *are* included.

This code implements *smart prediction* (see `smartpred`).

Value

If `se.fit = FALSE`, a vector or matrix of predictions. If `se.fit = TRUE`, a list with components

fitted.values	Predictions
se.fit	Estimated standard errors
df	Degrees of freedom
sigma	The square root of the dispersion parameter (but these are being phased out in the package)

Warning

This function may change in the future.

Note

Setting `se.fit = TRUE` and `type = "response"` will generate an error.

The arguments `type.fitted` and `percentiles` are provided in this function to give more convenience than modifying the extra slot directly.

Author(s)

Thomas W. Yee

References

Yee, T. W. (2015). *Vector Generalized Linear and Additive Models: With an Implementation in R*. New York, USA: Springer.

Yee, T. W. and Hastie, T. J. (2003). Reduced-rank vector generalized linear models. *Statistical Modelling*, **3**, 15–41.

See Also

[predict](#), [vglm](#), [predictvglm](#), [smartpred](#), [calibrate](#).

Examples

```
# Illustrates smart prediction
pneumo <- transform(pneumo, let = log(exposure.time))
fit <- vglm(cbind(normal, mild, severe) ~ poly(c(scale(let)), 2),
           propodds, pneumo, trace = TRUE, x.arg = FALSE)
class(fit)

(q0 <- head(predict(fit)))
(q1 <- predict(fit, newdata = head(pneumo)))
(q2 <- predict(fit, newdata = head(pneumo)))
all.equal(q0, q1) # Should be TRUE
all.equal(q1, q2) # Should be TRUE

head(predict(fit))
head(predict(fit, untransform = TRUE))

p0 <- head(predict(fit, type = "response"))
p1 <- head(predict(fit, type = "response", newdata = pneumo))
p2 <- head(predict(fit, type = "response", newdata = pneumo))
p3 <- head(fitted(fit))
all.equal(p0, p1) # Should be TRUE
all.equal(p1, p2) # Should be TRUE
all.equal(p2, p3) # Should be TRUE

predict(fit, type = "terms", se = TRUE)
```

Description

Estimation of a 3-parameter log-gamma distribution described by Prentice (1974).

Usage

```
prentice74(llocation = "identitylink", lscale = "loglink",
           lshape = "identitylink", ilocation = NULL, iscale = NULL,
           ishape = NULL, imethod = 1,
           glocation.mux = exp((-4:4)/2), gscale.mux = exp((-4:4)/2),
           gshape = qt(ppoints(6), df = 1), probs.y = 0.3,
           zero = c("scale", "shape"))
```

Arguments

`llocation`, `lscale`, `lshape`
Parameter link function applied to the location parameter a , positive scale parameter b and the shape parameter q , respectively. See [Links](#) for more choices.

`ilocation`, `iscale`
Initial value for a and b , respectively. The defaults mean an initial value is determined internally for each.

`ishape`
Initial value for q . If failure to converge occurs, try some other value. The default means an initial value is determined internally.

`imethod`, `zero` See [CommonVGAMffArguments](#) for information.

`glocation.mux`, `gscale.mux`, `gshape`, `probs.y`
See [CommonVGAMffArguments](#) for information.

Details

The probability density function is given by

$$f(y; a, b, q) = |q| \exp(w/q^2 - e^w) / (b \Gamma(1/q^2)),$$

for shape parameter $q \neq 0$, positive scale parameter $b > 0$, location parameter a , and all real y . Here, $w = (y - a)q/b + \psi(1/q^2)$ where ψ is the digamma function, [digamma](#). The mean of Y is a (returned as the fitted values). This is a different parameterization compared to [lgamma3](#).

Special cases: $q = 0$ is the normal distribution with standard deviation b , $q = -1$ is the extreme value distribution for maximums, $q = 1$ is the extreme value distribution for minima (Weibull). If $q > 0$ then the distribution is left skew, else $q < 0$ is right skew.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Warning

The special case $q = 0$ is not handled, therefore estimates of q too close to zero may cause numerical problems.

Note

The notation used here differs from Prentice (1974): $\alpha = a$, $\sigma = b$. Fisher scoring is used.

Author(s)

T. W. Yee

References

Prentice, R. L. (1974). A log gamma model and its maximum likelihood estimation. *Biometrika*, **61**, 539–544.

See Also

[lgamma3](#), [lgamma](#), [gengamma.stacy](#).

Examples

```
pdata <- data.frame(x2 = runif(nn <- 1000))
pdata <- transform(pdata, loc = -1 + 2*x2, Scale = exp(1))
pdata <- transform(pdata, y = rlgamma(nn, loc = loc, scale = Scale, shape = 1))
fit <- vglm(y ~ x2, prentice74(zero = 2:3), data = pdata, trace = TRUE)
coef(fit, matrix = TRUE) # Note the coefficients for location
```

 prinia

Yellow-bellied Prinia

Description

A data frame with yellow-bellied Prinia.

Usage

```
data(prinia)
```

Format

A data frame with 151 observations on the following 23 variables.

length a numeric vector, the scaled wing length (zero mean and unit variance).

fat a numeric vector, fat index; originally 1 (no fat) to 4 (very fat) but converted to 0 (no fat) versus 1 otherwise.

cap a numeric vector, number of times the bird was captured or recaptured.

noncap a numeric vector, number of times the bird was not captured.

y01, y02, y03, y04, y05, y06 a numeric vector of 0s and 1s; for noncapture and capture resp.

y07, y08, y09, y10, y11, y12 same as above.

y13, y14, y15, y16, y17, y18, y19 same as above.

Details

The yellow-bellied Prinia *Prinia flaviventris* is a common bird species located in Southeast Asia. A capture–recapture experiment was conducted at the Mai Po Nature Reserve in Hong Kong during 1991, where captured individuals had their wing lengths measured and fat index recorded. A total of 19 weekly capture occasions were considered, where 151 distinct birds were captured.

More generally, the prinias are a genus of small insectivorous birds, and are sometimes referred to as *wren-warblers*. They are a little-known group of the tropical and subtropical Old World, the roughly 30 species being divided fairly equally between Africa and Asia.

Source

Thanks to Paul Yip for permission to make this data available.

Hwang, W.-H. and Huggins, R. M. (2007) Application of semiparametric regression models in the analysis of capture–recapture experiments. *Australian and New Zealand Journal of Statistics* **49**, 191–202.

Examples

```
head(prinia)
summary(prinia)
rowSums(prinia[, c("cap", "noncap")]) # 19s

# Fit a positive-binomial distribution (M.h) to the data:
fit1 <- vglm(cbind(cap, noncap) ~ length + fat, posbinomial, prinia)

# Fit another positive-binomial distribution (M.h) to the data:
# The response input is suitable for posbernoulli.*-type functions.
fit2 <- vglm(cbind(y01, y02, y03, y04, y05, y06, y07, y08, y09, y10,
                  y11, y12, y13, y14, y15, y16, y17, y18, y19) ~
            length + fat, posbernoulli.b(drop.b = FALSE ~ 0), prinia)
```

probitlink

Probit Link Function

Description

Computes the probit transformation, including its inverse and the first two derivatives.

Usage

```
probitlink(theta, bvalue = NULL, inverse = FALSE, deriv = 0,
           short = TRUE, tag = FALSE)
```

Arguments

theta Numeric or character. See below for further details.
bvalue See [Links](#).
inverse, deriv, short, tag
 Details at [Links](#).

Details

The probit link function is commonly used for parameters that lie in the unit interval. It is the inverse CDF of the standard normal distribution. Numerical values of theta close to 0 or 1 or out of range result in Inf, -Inf, NA or NaN.

Value

For `deriv = 0`, the probit of theta, i.e., `qnorm(theta)` when `inverse = FALSE`, and if `inverse = TRUE` then `pnorm(theta)`.

For `deriv = 1`, then the function returns $d \text{ eta} / d \text{ theta}$ as a function of theta if `inverse = FALSE`, else if `inverse = TRUE` then it returns the reciprocal.

Note

Numerical instability may occur when theta is close to 1 or 0. One way of overcoming this is to use `bvalue`.

In terms of the threshold approach with cumulative probabilities for an ordinal response this link function corresponds to the univariate normal distribution (see [uninormal](#)).

Author(s)

Thomas W. Yee

References

McCullagh, P. and Nelder, J. A. (1989). *Generalized Linear Models*, 2nd ed. London: Chapman & Hall.

See Also

[Links](#), [logitlink](#), [clogloglink](#), [cauchitlink](#), [Normal](#).

Examples

```
p <- seq(0.01, 0.99, by = 0.01)
probitlink(p)
max(abs(probitlink(probitlink(p), inverse = TRUE) - p)) # Should be 0

p <- c(seq(-0.02, 0.02, by = 0.01), seq(0.97, 1.02, by = 0.01))
probitlink(p) # Has NAs
probitlink(p, bvalue = .Machine$double.eps) # Has no NAs

## Not run: p <- seq(0.01, 0.99, by = 0.01); par(lwd = (mylwd <- 2))
plot(p, logitlink(p), type = "l", col = "limegreen", ylab = "transformation",
     las = 1, main = "Some probability link functions")
lines(p, probitlink(p), col = "purple")
lines(p, clogloglink(p), col = "chocolate")
lines(p, cauchitlink(p), col = "tan")
abline(v = 0.5, h = 0, lty = "dashed")
legend(0.1, 4, c("logitlink", "probitlink", "clogloglink", "cauchitlink"),
```



```
col = c("limegreen", "purple", "chocolate", "tan"), lwd = mylwd)
par(lwd = 1)
## End(Not run)
```

profilevglm

Method for Profiling vglm Objects

Description

Investigates the profile log-likelihood function for a fitted model of class "vglm".

Usage

```
profilevglm(object, which = 1:p.vlm, alpha = 0.01,
            maxsteps = 10, del = zmax/5, trace = NULL, ...)
```

Arguments

object	the original fitted model object.
which	the original model parameters which should be profiled. This can be a numeric or character vector. By default, all parameters are profiled.
alpha	highest significance level allowed for the profiling.
maxsteps	maximum number of points to be used for profiling each parameter.
del	suggested change on the scale of the profile t-statistics. Default value chosen to allow profiling at about 10 parameter values.
trace	logical: should the progress of profiling be reported? The default is to use the trace value from the fitted object; see vglm.control for details.
...	further arguments passed to or from other methods.

Details

This function is called by [confintvglm](#) to do the profiling. See also [profile.glm](#) for details.

Value

A list of classes "profile.glm" and "profile" with an element for each parameter being profiled. The elements are data-frames with two variables

par.vals	a matrix of parameter values for each fitted model.
tau	the profile t-statistics.

Author(s)

T. W. Yee adapted this function from [profile.glm](#), written originally by D. M. Bates and W. N. Venables. (For S in 1996.) The help file was also used as a template.

See Also

[vglm](#), [confintvglm](#), [lrt.stat](#), [profile](#), [profile.glm](#), [plot.profile](#) in **MASS** or **stats**.

Examples

```
pneumo <- transform(pneumo, let = log(exposure.time))
fit1 <- vglm(cbind(normal, mild, severe) ~ let, propodds,
            trace = TRUE, data = pneumo)
pfit1 <- profile(fit1, trace = FALSE)
confint(fit1, method = "profile", trace = FALSE)
```

propodds

Proportional Odds Model for Ordinal Regression

Description

Fits the proportional odds model to a (preferably ordered) factor response.

Usage

```
propodds(reverse = TRUE, whitespace = FALSE, ynames = FALSE,
         Thresh = NULL, Trev = reverse, Tref = if (Trev) "M" else 1)
```

Arguments

`reverse`, `whitespace`
Logical. Fed into arguments of the same name in [cumulative](#).

`ynames`
See [multinomial](#) for information.

`Thresh`, `Trev`, `Tref`
Fed into arguments of the same name in [cumulative](#).

Details

The *proportional odds model* is a special case from the class of *cumulative link models*. It involves a logit link applied to cumulative probabilities and a strong *parallelism* assumption. A parallelism assumption means there is less chance of numerical problems because the fitted probabilities will remain between 0 and 1; however the *parallelism* assumption ought to be checked, e.g., via a likelihood ratio test. This **VGAM** family function is merely a shortcut for `cumulative(reverse = reverse, link = "logit", parallel = TRUE)`. Please see [cumulative](#) for more details on this model.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Warning

No check is made to verify that the response is ordinal if the response is a matrix; see [ordered](#).

Author(s)

Thomas W. Yee

References

See [cumulative](#).

See Also

[cumulative](#), [R2latvar](#).

Examples

```
# Fit the proportional odds model, McCullagh and Nelder (1989,p.179)
pneumo <- transform(pneumo, let = log(exposure.time))
(fit <- vglm(cbind(normal, mild, severe) ~ let, propodds, pneumo))
depvar(fit) # Sample proportions
weights(fit, type = "prior") # Number of observations
coef(fit, matrix = TRUE)
constraints(fit) # Constraint matrices
summary(fit)

# Check that the model is linear in let -----
fit2 <- vgam(cbind(normal, mild, severe) ~ s(let, df = 2), propodds,
            pneumo)
## Not run: plot(fit2, se = TRUE, lcol = 2, scol = 2)

# Check the proportional odds assumption with a LRT -----
(fit3 <- vglm(cbind(normal, mild, severe) ~ let,
             cumulative(parallel = FALSE, reverse = TRUE), pneumo))
pchisq(deviance(fit) - deviance(fit3),
       df = df.residual(fit) - df.residual(fit3), lower.tail = FALSE)
lrtest(fit3, fit) # Easier
```

Description

Plots the fitted probabilities for some very simplified special cases of categorical data analysis models.

Usage

```
prplot(object, control = prplot.control(...), ...)

prplot.control(xlab = NULL, ylab = "Probability", main = NULL, xlim = NULL,
  ylim = NULL, lty = par()$lty, col = par()$col, rcol = par()$col,
  lwd = par()$lwd, rlwd = par()$lwd, las = par()$las, rug.arg = FALSE, ...)
```

Arguments

object	Currently only an cumulative object. This includes a propodds object since that VGAM family function is a special case of cumulative .
control	List containing some basic graphical parameters.
xlab, ylab, main, xlim, ylim, lty	See par and ... below.
col, rcol, lwd, rlwd, las, rug.arg	See par and ... below. Arguments starting with r refer to the rug. Argument rug.arg is logical: add a rug for the distinct values of the explanatory variable?
...	Arguments such as xlab which are fed into <code>prplot.control()</code> . Only a small selection of graphical arguments from par are offered.

Details

For models involving one term in the RHS of the formula this function plots the fitted probabilities against the single explanatory variable.

Value

The object is returned invisibly with the `preplot` slot assigned. This is obtained by a call to `plotvgam()`.

Note

This function is rudimentary.

See Also

[cumulative](#).

Examples

```
pneumo <- transform(pneumo, let = log(exposure.time))
fit <- vglm(cbind(normal, mild, severe) ~ let, propodds, data = pneumo)
M <- npred(fit) # Or fit@misc$M
## Not run: prplot(fit)
prplot(fit, lty = 1:M, col = (1:M)+2, rug = TRUE, las = 1,
  ylim = c(0, 1), rlwd = 2)
## End(Not run)
```

put.smart	<i>Adds a List to the End of the List “.smart.prediction”</i>
-----------	---

Description

Adds a list to the end of the list `.smart.prediction` in `smartpredenv`.

Usage

```
put.smart(smart)
```

Arguments

`smart` a list containing parameters needed later for smart prediction.

Details

`put.smart` is used in "write" mode within a smart function. It saves parameters at the time of model fitting, which are later used for prediction. The function `put.smart` is the opposite of [get.smart](#), and both deal with the same contents.

Value

Nothing is returned.

Side Effects

The variable `.smart.prediction.counter` in `smartpredenv` is incremented beforehand, and `.smart.prediction[.smart]` is assigned the list `smart`. If the list `.smart.prediction` in `smartpredenv` is not long enough to hold `smart`, then it is made larger, and the variable `.max.smart` in `smartpredenv` is adjusted accordingly.

See Also

[get.smart](#).

Examples

```
print(sm.min1)
```

qrrvglm.control

*Control Function for QRR-VGLMs (CQO)***Description**

Algorithmic constants and parameters for a constrained quadratic ordination (CQO), by fitting a *quadratic reduced-rank vector generalized linear model* (QRR-VGLM), are set using this function. It is the control function for `cqo`.

Usage

```
qrrvglm.control(Rank = 1, Bestof = if (length(Cinit)) 1 else 10,
  checkwz = TRUE, Cinit = NULL, Crowlpositive = TRUE,
  epsilon = 1.0e-06, EqualTolerances = NULL, eq.tolerances = TRUE,
  Etamat.colmax = 10, FastAlgorithm = TRUE, GradientFunction = TRUE,
  Hstep = 0.001, isd.latvar = rep_len(c(2, 1, rep_len(0.5, Rank)),
  Rank), ikvector = 0.1, iShape = 0.1, ITolerances = NULL,
  I.tolerances = FALSE, maxitl = 40, imethod = 1,
  Maxit.optim = 250, MUXfactor = rep_len(7, Rank),
  noRRR = ~ 1, Norrr = NA, optim.maxit = 20,
  Parscale = if (I.tolerances) 0.001 else 1.0,
  sd.Cinit = 0.02, SmallNo = 5.0e-13, trace = TRUE,
  Use.Init.Poisson.QO = TRUE,
  wzepsilon = .Machine$double.eps^0.75, ...)
```

Arguments

In the following, R is the Rank, M is the number of linear predictors, and S is the number of responses (species). Thus $M = S$ for binomial and Poisson responses, and $M = 2S$ for the negative binomial and 2-parameter gamma distributions.

Rank	The numerical rank R of the model, i.e., the number of ordination axes. Must be an element from the set $\{1, 2, \dots, \min(M, p_2)\}$ where the vector of explanatory variables x is partitioned into (x_1, x_2) , which is of dimension $p_1 + p_2$. The variables making up x_1 are given by the terms in the <code>noRRR</code> argument, and the rest of the terms comprise x_2 .
Bestof	Integer. The best of <code>Bestof</code> models fitted is returned. This argument helps guard against local solutions by (hopefully) finding the global solution from many fits. The argument has value 1 if an initial value for C is inputted using <code>Cinit</code> .
checkwz	logical indicating whether the diagonal elements of the working weight matrices should be checked whether they are sufficiently positive, i.e., greater than <code>wzepsilon</code> . If not, any values less than <code>wzepsilon</code> are replaced with this value.
Cinit	Optional initial C matrix, which must be a p_2 by R matrix. The default is to apply <code>.Init.Poisson.QO()</code> to obtain initial values.

<code>Crow1positive</code>	Logical vector of length <code>Rank</code> (recycled if necessary): are the elements of the first row of C positive? For example, if <code>Rank</code> is 4, then specifying <code>Crow1positive = c(FALSE, TRUE)</code> will force $C[1, 1]$ and $C[1, 3]$ to be negative, and $C[1, 2]$ and $C[1, 4]$ to be positive. This argument allows for a reflection in the ordination axes because the coefficients of the latent variables are unique up to a sign.
<code>epsilon</code>	Positive numeric. Used to test for convergence for GLMs fitted in C . Larger values mean a loosening of the convergence criterion. If an error code of 3 is reported, try increasing this value.
<code>eq.tolerances</code>	Logical indicating whether each (quadratic) predictor will have equal tolerances. Having <code>eq.tolerances = TRUE</code> can help avoid numerical problems, especially with binary data. Note that the estimated (common) tolerance matrix may or may not be positive-definite. If it is then it can be scaled to the R by R identity matrix, i.e., made equivalent to <code>I.tolerances = TRUE</code> . Setting <code>I.tolerances = TRUE</code> will <i>force</i> a common R by R identity matrix as the tolerance matrix to the data even if it is not appropriate. In general, setting <code>I.tolerances = TRUE</code> is preferred over <code>eq.tolerances = TRUE</code> because, if it works, it is much faster and uses less memory. However, <code>I.tolerances = TRUE</code> requires the environmental variables to be scaled appropriately. See Details for more details.
<code>EqualTolerances</code>	Defunct argument. Use <code>eq.tolerances</code> instead.
<code>Etamat.colmax</code>	Positive integer, no smaller than <code>Rank</code> . Controls the amount of memory used by <code>.Init.Poisson.QO()</code> . It is the maximum number of columns allowed for the pseudo-response and its weights. In general, the larger the value, the better the initial value. Used only if <code>Use.Init.Poisson.QO = TRUE</code> .
<code>FastAlgorithm</code>	Logical. Whether a new fast algorithm is to be used. The fast algorithm results in a large speed increases compared to Yee (2004). Some details of the fast algorithm are found in Appendix A of Yee (2006). Setting <code>FastAlgorithm = FALSE</code> will give an error.
<code>GradientFunction</code>	Logical. Whether <code>optim</code> 's argument <code>gr</code> is used or not, i.e., to compute gradient values. Used only if <code>FastAlgorithm</code> is <code>TRUE</code> . The default value is usually faster on most problems.
<code>Hstep</code>	Positive value. Used as the step size in the finite difference approximation to the derivatives by <code>optim</code> .
<code>isd.latvar</code>	Initial standard deviations for the latent variables (site scores). Numeric, positive and of length R (recycled if necessary). This argument is used only if <code>I.tolerances = TRUE</code> . Used by <code>.Init.Poisson.QO()</code> to obtain initial values for the constrained coefficients C adjusted to a reasonable value. It adjusts the spread of the site scores relative to a common species tolerance of 1 for each ordination axis. A value between 0.5 and 10 is recommended; a value such as 10 means that the range of the environmental space is very large relative to the niche width of the species. The successive values should decrease because the first ordination axis should have the most spread of site scores, followed by the second ordination axis, etc.
<code>iKvector, iShape</code>	Numeric, recycled to length S if necessary. Initial values used for estimating the positive k and λ parameters of the negative binomial and 2-parameter

	gamma distributions respectively. For further information see negbinomial and gamma2 . These arguments override the <code>ik</code> and <code>ishape</code> arguments in negbinomial and gamma2 .
<code>I.tolerances</code>	Logical. If TRUE then the (common) tolerance matrix is the R by R identity matrix by definition. Note that having <code>I.tolerances = TRUE</code> implies <code>eq.tolerances = TRUE</code> , but not vice versa. Internally, the quadratic terms will be treated as offsets (in GLM jargon) and so the models can potentially be fitted very efficiently. <i>However, it is a very good idea to center and scale all numerical variables in the x_2 vector.</i> See Details for more details. The success of <code>I.tolerances = TRUE</code> often depends on suitable values for <code>isd.latvar</code> and/or <code>MUXfactor</code> .
<code>ITolerances</code>	Defunct argument. Use <code>I.tolerances</code> instead.
<code>maxitl</code>	Maximum number of times the optimizer is called or restarted. Most users should ignore this argument.
<code>imethod</code>	Method of initialization. A positive integer 1 or 2 or 3 etc. depending on the VGAM family function. Currently it is used for negbinomial and gamma2 only, and used within the C.
<code>Maxit.optim</code>	Positive integer. Number of iterations given to the function <code>optim</code> at each of the <code>optim.maxit</code> iterations.
<code>MUXfactor</code>	Multiplication factor for detecting large offset values. Numeric, positive and of length R (recycled if necessary). This argument is used only if <code>I.tolerances = TRUE</code> . Offsets are -0.5 multiplied by the sum of the squares of all R latent variable values. If the latent variable values are too large then this will result in numerical problems. By too large, it is meant that the standard deviation of the latent variable values are greater than <code>MUXfactor[r] * isd.latvar[r]</code> for $r=1:Rank$ (this is why centering and scaling all the numerical predictor variables in x_2 is recommended). A value about 3 or 4 is recommended. If failure to converge occurs, try a slightly lower value.
<code>optim.maxit</code>	Positive integer. Number of times <code>optim</code> is invoked. At iteration i , the i th value of <code>Maxit.optim</code> is fed into <code>optim</code> .
<code>noRRR</code>	Formula giving terms that are <i>not</i> to be included in the reduced-rank regression (or formation of the latent variables), i.e., those belong to x_1 . Those variables which do not make up the latent variable (reduced-rank regression) correspond to the B_1 matrix. The default is to omit the intercept term from the latent variables.
<code>Norrr</code>	Defunct. Please use <code>noRRR</code> . Use of <code>Norrr</code> will become an error soon.
<code>Parscale</code>	Numerical and positive-valued vector of length C (recycled if necessary). Passed into <code>optim(..., control = list(parscale = Parscale))</code> ; the elements of C become $C / Parscale$. Setting <code>I.tolerances = TRUE</code> results in line searches that are very large, therefore C has to be scaled accordingly to avoid large step sizes. See Details for more information. It's probably best to leave this argument alone.
<code>sd.Cinit</code>	Standard deviation of the initial values for the elements of C . These are normally distributed with mean zero. This argument is used only if <code>Use.Init.Poisson.Q0 = FALSE</code> and C is not inputted using <code>Cinit</code> .

trace	Logical indicating if output should be produced for each iteration. The default is TRUE because the calculations are numerically intensive, meaning it may take a long time, so that the user might think the computer has locked up if trace = FALSE.
SmallNo	Positive numeric between <code>.Machine\$double.eps</code> and <code>0.0001</code> . Used to avoid under- or over-flow in the IRLS algorithm. Used only if <code>FastAlgorithm</code> is TRUE.
Use.Init.Poisson.QO	Logical. If TRUE then the function <code>.Init.Poisson.QO()</code> is used to obtain initial values for the canonical coefficients C . If FALSE then random numbers are used instead.
wzepsilon	Small positive number used to test whether the diagonals of the working weight matrices are sufficiently positive.
...	Ignored at present.

Details

Recall that the central formula for CQO is

$$\eta = B_1^T x_1 + A\nu + \sum_{m=1}^M (\nu^T D_m \nu) e_m$$

where x_1 is a vector (usually just a 1 for an intercept), x_2 is a vector of environmental variables, $\nu = C^T x_2$ is a R -vector of latent variables, e_m is a vector of 0s but with a 1 in the m th position. QRR-VGLMs are an extension of RR-VGLMs and allow for maximum likelihood solutions to constrained quadratic ordination (CQO) models.

Having `I.tolerances = TRUE` means all the tolerance matrices are the order- R identity matrix, i.e., it *forces* bell-shaped curves/surfaces on all species. This results in a more difficult optimization problem (especially for 2-parameter models such as the negative binomial and gamma) because of overflow errors and it appears there are more local solutions. To help avoid the overflow errors, scaling C by the factor `Parscale` can help enormously. Even better, scaling C by specifying `isd.latvar` is more understandable to humans. If failure to converge occurs, try adjusting `Parscale`, or better, setting `eq.tolerances = TRUE` (and hope that the estimated tolerance matrix is positive-definite). To fit an equal-tolerances model, it is firstly best to try setting `I.tolerances = TRUE` and varying `isd.latvar` and/or `MUXfactor` if it fails to converge. If it still fails to converge after many attempts, try setting `eq.tolerances = TRUE`, however this will usually be a lot slower because it requires a lot more memory.

With a $R > 1$ model, the latent variables are always uncorrelated, i.e., the variance-covariance matrix of the site scores is a diagonal matrix.

If setting `eq.tolerances = TRUE` is used and the common estimated tolerance matrix is positive-definite then that model is effectively the same as the `I.tolerances = TRUE` model (the two are transformations of each other). In general, `I.tolerances = TRUE` is numerically more unstable and presents a more difficult problem to optimize; the arguments `isd.latvar` and/or `MUXfactor` often must be assigned some good value(s) (possibly found by trial and error) in order for convergence to occur. Setting `I.tolerances = TRUE` *forces* a bell-shaped curve or surface onto all the species data, therefore this option should be used with deliberation. If unsuitable, the resulting fit may be very misleading. Usually it is a good idea for the user to set `eq.tolerances = FALSE` to see which

species appear to have a bell-shaped curve or surface. Improvements to the fit can often be achieved using transformations, e.g., nitrogen concentration to log nitrogen concentration.

Fitting a CAO model (see [cao](#)) first is a good idea for pre-examining the data and checking whether it is appropriate to fit a CQO model.

Value

A list with components matching the input names.

Warning

The default value of `Bestof` is a bare minimum for many datasets, therefore it will be necessary to increase its value to increase the chances of obtaining the global solution.

Note

When `I.tolerances = TRUE` it is a good idea to apply [scale](#) to all the numerical variables that make up the latent variable, i.e., those of x_2 . This is to make them have mean 0, and hence avoid large offset values which cause numerical problems.

This function has many arguments that are common with [rrvglm.control](#) and [vglm.control](#).

It is usually a good idea to try fitting a model with `I.tolerances = TRUE` first, and if convergence is unsuccessful, then try `eq.tolerances = TRUE` and `I.tolerances = FALSE`. Ordination diagrams with `eq.tolerances = TRUE` have a natural interpretation, but with `eq.tolerances = FALSE` they are more complicated and requires, e.g., contours to be overlaid on the ordination diagram (see [lvplot.qrrvglm](#)).

In the example below, an equal-tolerances CQO model is fitted to the hunting spiders data. Because `I.tolerances = TRUE`, it is a good idea to center all the x_2 variables first. Upon fitting the model, the actual standard deviation of the site scores are computed. Ideally, the `isd.latvar` argument should have had this value for the best chances of getting good initial values. For comparison, the model is refitted with that value and it should run more faster and reliably.

Author(s)

Thomas W. Yee

References

Yee, T. W. (2004). A new technique for maximum-likelihood canonical Gaussian ordination. *Ecological Monographs*, **74**, 685–701.

Yee, T. W. (2006). Constrained additive ordination. *Ecology*, **87**, 203–213.

See Also

[cqo](#), [rcqo](#), [Coef.qrrvglm](#), [Coef.qrrvglm-class](#), [optim](#), [binomialff](#), [poissonff](#), [negbinomial](#), [gamma2](#).

Examples

```
## Not run: # Poisson CQO with equal tolerances
set.seed(111) # This leads to the global solution
hspider[,1:6] <- scale(hspider[,1:6]) # Good when I.tolerances = TRUE
p1 <- cqo(cbind(Alopacce, Alopcone, Alopfabr,
               Arctlute, Arctperi, Auloalbi,
               Pardlugu, Pardmont, Pardnigr,
               Pardpull, Trocterr, Zoraspin) ~
          WaterCon + BareSand + FallTwig +
          CoveMoss + CoveHerb + ReflLux,
          poissonff, data = hspider, eq.tolerances = TRUE)
sort(deviance(p1, history = TRUE)) # Iteration history

(isd.latvar <- apply(latvar(p1), 2, sd)) # Approx isd.latvar

# Refit the model with better initial values
set.seed(111) # This leads to the global solution
p1 <- cqo(cbind(Alopacce, Alopcone, Alopfabr,
               Arctlute, Arctperi, Auloalbi,
               Pardlugu, Pardmont, Pardnigr,
               Pardpull, Trocterr, Zoraspin) ~
          WaterCon + BareSand + FallTwig +
          CoveMoss + CoveHerb + ReflLux,
          I.tolerances = TRUE, poissonff, data = hspider,
          isd.latvar = isd.latvar) # Note this
sort(deviance(p1, history = TRUE)) # Iteration history

## End(Not run)
```

qtplot.gumbel

Quantile Plot for Gumbel Regression

Description

Plots quantiles associated with a Gumbel model.

Usage

```
qtplot.gumbel(object, show.plot = TRUE,
              y.arg = TRUE, spline.fit = FALSE, label = TRUE,
              R = object@misc$R, percentiles = object@misc$percentiles,
              add.arg = FALSE, mpv = object@misc$mpv,
              xlab = NULL, ylab = "", main = "",
              pch = par()$pch, pcol.arg = par()$col,
              lty.arg = par()$lty, lcol.arg = par()$col, llwd.arg = par()$lwd,
              tcol.arg = par()$col, tadj = 1, ...)
```

Arguments

object	A VGAM extremes model of the Gumbel type, produced by modelling functions such as <code>vglm</code> and <code>vgam</code> , and with a family function that is either <code>gumbel</code> or <code>gumbelff</code> .
show.plot	Logical. Plot it? If FALSE no plot will be done.
y.arg	Logical. Add the raw data on to the plot?
spline.fit	Logical. Use a spline fit through the fitted percentiles? This can be useful if there are large gaps between some values along the covariate.
label	Logical. Label the percentiles?
R	See <code>gumbel</code> .
percentiles	See <code>gumbel</code> .
add.arg	Logical. Add the plot to an existing plot?
mpv	See <code>gumbel</code> .
xlab	Caption for the x-axis. See <code>par</code> .
ylab	Caption for the y-axis. See <code>par</code> .
main	Title of the plot. See <code>title</code> .
pch	Plotting character. See <code>par</code> .
pcol.arg	Color of the points. See the <code>col</code> argument of <code>par</code> .
lty.arg	Line type. Line type. See the <code>lty</code> argument of <code>par</code> .
lcol.arg	Color of the lines. See the <code>col</code> argument of <code>par</code> .
lwd.arg	Line width. See the <code>lwd</code> argument of <code>par</code> .
tcol.arg	Color of the text (if <code>label</code> is TRUE). See the <code>col</code> argument of <code>par</code> .
tadj	Text justification. See the <code>adj</code> argument of <code>par</code> .
...	Arguments passed into the plot function when setting up the entire plot. Useful arguments here include <code>sub</code> and <code>las</code> .

Details

There should be a single covariate such as time. The quantiles specified by percentiles are plotted.

Value

The object with a list called `qtplot` in the post slot of `object`. (If `show.plot = FALSE` then just the list is returned.) The list contains components

<code>fitted.values</code>	The percentiles of the response, possibly including the MPV.
<code>percentiles</code>	The percentiles (small vector of values between 0 and 100).

Note

Unlike `gumbel`, one cannot have `percentiles = NULL`.

Author(s)

Thomas W. Yee

See Also[gumbel](#).**Examples**

```

ymat <- as.matrix(venice[, paste("r", 1:10, sep = "")])
fit1 <- vgam(ymat ~ s(year, df = 3), gumbel(R = 365, mpv = TRUE),
            data = venice, trace = TRUE, na.action = na.pass)
head(fitted(fit1))

## Not run: par(mfrow = c(1, 1), bty = "l", xpd = TRUE, las = 1)
qtplot(fit1, mpv = TRUE, lcol = c(1, 2, 5), tcol = c(1, 2, 5),
       lwd = 2, pcol = "blue", tadj = 0.4, ylab = "Sea level (cm)")

qtplot(fit1, perc = 97, mpv = FALSE, lcol = 3, tcol = 3,
       lwd = 2, tadj = 0.4, add = TRUE) -> saved
head(saved$post$qtplot$fitted)

## End(Not run)

```

qtplot.lmscreg

*Quantile Plot for LMS Quantile Regression***Description**

Plots quantiles associated with a LMS quantile regression.

Usage

```

qtplot.lmscreg(object, newdata = NULL,
               percentiles = object@misc$percentiles,
               show.plot = TRUE, ...)

```

Arguments

object	A VGAM quantile regression model, i.e., an object produced by modelling functions such as vglm and vgam with a family function beginning with "lms.", e.g., lms.yjn .
newdata	Optional data frame for computing the quantiles. If missing, the original data is used.
percentiles	Numerical vector with values between 0 and 100 that specify the percentiles (quantiles). The default are the percentiles used when the model was fitted.
show.plot	Logical. Plot it? If FALSE no plot will be done.
...	Graphical parameter that are passed into plotqtplot.lmscreg .

Details

The ‘primary’ variable is defined as the main covariate upon which the regression or smoothing is performed. For example, in medical studies, it is often the age. In **VGAM**, it is possible to handle more than one covariate, however, the primary variable must be the first term after the intercept.

Value

A list with the following components.

`fitted.values` A vector of fitted percentile values.

`percentiles` The percentiles used.

Note

[plotqtplot.lmscreg](#) does the actual plotting.

Author(s)

Thomas W. Yee

References

Yee, T. W. (2004). Quantile regression via vector generalized additive models. *Statistics in Medicine*, **23**, 2295–2315.

See Also

[plotqtplot.lmscreg](#), [deplot.lmscreg](#), [lms.bcn](#), [lms.bcg](#), [lms.yjn](#).

Examples

```
## Not run:
fit <- vgam(BMI ~ s(age, df = c(4, 2)), lms.bcn(zero=1), bmi.nz)
qtplot(fit)
qtplot(fit, perc = c(25, 50, 75, 95), lcol = 4, tcol = 4, llwd = 2)

## End(Not run)
```

Qvar

Quasi-variances Preprocessing Function

Description

Takes a [vglm](#) fit or a variance-covariance matrix, and preprocesses it for [rcim](#) and [uninormal](#) so that quasi-variances can be computed.

Usage

```
Qvar(object, factorname = NULL, which.linpred = 1,
      coef.indices = NULL, labels = NULL,
      dispersion = NULL, reference.name = "(reference)", estimates = NULL)
```

Arguments

object	A "vglm" object or a variance-covariance matrix, e.g., <code>vcov(vglm.object)</code> . The former is preferred since it contains all the information needed. If a matrix then <code>factorname</code> and/or <code>coef.indices</code> should be specified to identify the factor.
which.linpred	A single integer from the set $1:M$. Specifies which linear predictor to use. Let the value of <code>which.linpred</code> be called j . Then the factor should appear in that linear predictor, hence the j th row of the constraint matrix corresponding to the factor should have at least one nonzero value. Currently the j th row must have exactly one nonzero value because programming it for more than one nonzero value is difficult.
factorname	Character. If the <code>vglm</code> object contains more than one factor as explanatory variable then this argument should be the name of the factor of interest. If <code>object</code> is a variance-covariance matrix then this argument should also be specified.
labels	Character. Optional, for labelling the variance-covariance matrix.
dispersion	Numeric. Optional, passed into <code>vcov()</code> with the same argument name.
reference.name	Character. Label for for the reference level.
coef.indices	Optional numeric vector of length at least 3 specifying the indices of the factor from the variance-covariance matrix.
estimates	an optional vector of estimated coefficients (redundant if <code>object</code> is a model).

Details

Suppose a factor with L levels is an explanatory variable in a regression model. By default, R treats the first level as baseline so that its coefficient is set to zero. It estimates the other $L - 1$ coefficients, and with its associated standard errors, this is the conventional output. From the complete variance-covariance matrix one can compute L quasi-variances based on all pairwise difference of the coefficients. They are based on an approximation, and can be treated as uncorrelated. In minimizing the relative (not absolute) errors it is not hard to see that the estimation involves a RCIM (`rcim`) with an exponential link function (`explink`).

If `object` is a model, then at least one of `factorname` or `coef.indices` must be non-NULL. The value of `coef.indices`, if non-NULL, determines which rows and columns of the model's variance-covariance matrix to use. If `coef.indices` contains a zero, an extra row and column are included at the indicated position, to represent the zero variances and covariances associated with a reference level. If `coef.indices` is NULL, then `factorname` should be the name of a factor effect in the model, and is used in order to extract the necessary variance-covariance estimates.

Quasi-variances were first implemented in R with `qvcalc`. This implementation draws heavily from that.

Value

A L by L matrix whose i - j element is the logarithm of the variance of the i th coefficient minus the j th coefficient, for all values of i and j . The diagonal elements are arbitrary and are set to zero.

The matrix has an attribute that corresponds to the prior weight matrix; it is accessed by `uninormal` and replaces the usual `weights` argument. of `vglm`. This weight matrix has ones on the off-diagonals and some small positive number on the diagonals.

Warning

Negative quasi-variances may occur (one of them and only one), though they are rare in practice. If so then numerical problems may occur. See `qvcalc()` for more information.

Note

This is an adaptation of `qvcalc()` in `qvcalc`. It should work for all `vglm` models with one linear predictor, i.e., $M = 1$. For $M > 1$ the factor should appear only in one of the linear predictors.

It is important to set `maxit` to be larger than usual for `rcim` since convergence is slow. Upon successful convergence the i th row effect and the i th column effect should be equal. A simple computation involving the fitted and predicted values allows the quasi-variances to be extracted (see example below).

A function to plot *comparison intervals* has not been written here.

Author(s)

T. W. Yee, based heavily on `qvcalc()` in `qvcalc` written by David Firth.

References

Firth, D. (2003). Overcoming the reference category problem in the presentation of statistical models. *Sociological Methodology* **33**, 1–18.

Firth, D. and de Menezes, R. X. (2004). Quasi-variances. *Biometrika* **91**, 65–80.

Yee, T. W. and Hadi, A. F. (2014). Row-column interaction models, with an R implementation. *Computational Statistics*, **29**, 1427–1445.

See Also

`rcim`, `vglm`, `qvar`, `uninormal`, `explink`, `qvcalc()` in `qvcalc`, `ships`.

Examples

```
# Example 1
data("ships", package = "MASS")

Shipmodel <- vglm(incidents ~ type + year + period,
                 poissonff, offset = log(service),
                 trace = TRUE, model = TRUE,
                 data = ships, subset = (service > 0))
```



```

# Easiest form of input
fit1 <- rcim(Qvar(Shipmodel, "type"), uninormal("explink"), maxit = 99)
qvar(fit1)           # Easy method to get the quasi-variances
qvar(fit1, se = TRUE) # Easy method to get the quasi-standard errors

(quasiVar <- exp(diag(fitted(fit1))) / 2)           # Version 1
(quasiVar <- diag(predict(fit1)[, c(TRUE, FALSE)]) / 2) # Version 2
(quasiSE <- sqrt(quasiVar))

# Another form of input
fit2 <- rcim(Qvar(Shipmodel, coef.ind = c(0, 2:5), reference.name = "typeA"),
             uninormal("explink"), maxit = 99)
## Not run: qvplot(fit2, col = "green", lwd = 3, scol = "blue", slwd = 2, las = 1)

# The variance-covariance matrix is another form of input (not recommended)
fit3 <- rcim(Qvar(cbind(0, rbind(0, vcov(Shipmodel)[2:5, 2:5])),
                labels = c("typeA", "typeB", "typeC", "typeD", "typeE"),
                estimates = c(typeA = 0, coef(Shipmodel)[2:5])),
            uninormal("explink"), maxit = 99)
(QuasiVar <- exp(diag(fitted(fit3))) / 2)           # Version 1
(QuasiVar <- diag(predict(fit3)[, c(TRUE, FALSE)]) / 2) # Version 2
(QuasiSE <- sqrt(quasiVar))
## Not run: qvplot(fit3)

# Example 2: a model with M > 1 linear predictors
## Not run: require("VGAMdata")
xs.nz.f <- subset(xs.nz, sex == "F")
xs.nz.f <- subset(xs.nz.f, !is.na(babies) & !is.na(age) & !is.na(ethnicity))
xs.nz.f <- subset(xs.nz.f, ethnicity != "Other")

clist <- list("sm.bs(age, df = 4)" = rbind(1, 0),
             "sm.bs(age, df = 3)" = rbind(0, 1),
             "ethnicity"           = diag(2),
             "(Intercept)"         = diag(2))
fit1 <- vglm(babies ~ sm.bs(age, df = 4) + sm.bs(age, df = 3) + ethnicity,
            zipoissonff(zero = NULL), xs.nz.f,
            constraints = clist, trace = TRUE)
Fit1 <- rcim(Qvar(fit1, "ethnicity", which.linpred = 1),
            uninormal("explink", imethod = 1), maxit = 99, trace = TRUE)
Fit2 <- rcim(Qvar(fit1, "ethnicity", which.linpred = 2),
            uninormal("explink", imethod = 1), maxit = 99, trace = TRUE)

## End(Not run)
## Not run: par(mfrow = c(1, 2))
qvplot(Fit1, scol = "blue", pch = 16, main = expression(eta[1]),
       slwd = 1.5, las = 1, length.arrows = 0.07)
qvplot(Fit2, scol = "blue", pch = 16, main = expression(eta[2]),
       slwd = 1.5, las = 1, length.arrows = 0.07)

## End(Not run)

```

qvar

Quasi-variances Extraction Function

Description

Takes a [rcim](#) fit of the appropriate format and returns either the quasi-variances or quasi-standard errors.

Usage

```
qvar(object, se = FALSE, ...)
```

Arguments

object	A rcim object that has family function uninormal with the explink link. See below for an example.
se	Logical. If FALSE then the quasi-variances are returned, else the square root of them, called quasi-standard errors.
...	Currently unused.

Details

This simple function is ad hoc and simply is equivalent to computing the quasi-variances by `diag(predict(fit1)[, c(TRUE, FALSE)]) / 2`. This function is for convenience only. Serious users of quasi-variances ought to understand why and how this function works.

Value

A vector of quasi-variances or quasi-standard errors.

Author(s)

T. W. Yee.

See Also

[rcim](#), [uninormal](#), [explink](#), [Qvar](#), [ships](#).

Examples

```
data("ships", package = "MASS")
Shipmodel <- vglm(incidents ~ type + year + period,
                 poissonff, offset = log(service),
                 data = ships, subset = (service > 0))

# Easiest form of input
fit1 = rcim(Qvar(Shipmodel, "type"), uninormal("explink"), maxit=99)
qvar(fit1)           # Quasi-variances
```

```

qvar(fit1, se = TRUE) # Quasi-standard errors

# Manually compute them:
(quasiVar <- exp(diag(fitted(fit1))) / 2) # Version 1
(quasiVar <- diag(predict(fit1)[, c(TRUE, FALSE)]) / 2) # Version 2
(quasiSE <- sqrt(quasiVar))

## Not run: qvplot(fit1, col = "green", lwd = 3, scol = "blue",
  slwd = 2, las = 1)
## End(Not run)

```

R2latvar

*R-squared for Latent Variable Models***Description**

R-squared goodness of fit for latent variable models, such as cumulative link models. Some software such as Stata call the quantity the McKelvey–Zavoina R-squared, which was proposed in their 1975 paper for cumulative probit models.

Usage

```
R2latvar(object)
```

Arguments

object A [cumulative](#) or [binomialff](#) fit using [vglm](#). Only a few selected link functions are currently permitted: [logitlink](#), [probitlink](#), [clogloglink](#). For models with more than one linear predictor, a parallelism assumption is needed also, i.e., the constraint matrices must be a 1-column matrix of 1s (except for the intercept). The model is assumed to have an intercept term.

Details

Models such as the proportional odds model have a latent variable interpretation (see, e.g., Section 6.2.6 of Agresti (2018), Section 14.4.1.1 of Yee (2015), Section 5.2.2 of McCullagh and Nelder (1989)). It is possible to summarize the predictive power of the model by computing R^2 on the transformed scale, e.g., on a standard normal distribution for a [probitlink](#) link. For more details see Section 6.3.7 of Agresti (2018).

Value

The R^2 value. Approximately, that amount is the variability in the latent variable of the model explained by all the explanatory variables. Then taking the positive square-root gives an approximate multiple correlation R .

Author(s)

Thomas W. Yee

References

- Agresti, A. (2018). *An Introduction to Categorical Data Analysis, 3rd ed.*, New York: John Wiley & Sons.
- McKelvey, R. D. and W. Zavoina (1975). A statistical model for the analysis of ordinal level dependent variables. *The Journal of Mathematical Sociology*, **4**, 103–120.

See Also

[vglm](#), [cumulative](#), [propodds](#), [logitlink](#), [probitlink](#), [clogloglink](#), [summary.lm](#).

Examples

```
pneumo <- transform(pneumo, let = log(exposure.time))
(fit <- vglm(cbind(normal, mild, severe) ~ let, propodds, data = pneumo))
R2latvar(fit)
```

Rank	<i>Rank</i>
------	-------------

Description

Returns the rank of reduced-rank regression-type models in the VGAM package.

Usage

```
Rank(object, ...)
Rank.rrvglm(object, ...)
Rank.qrrvglm(object, ...)
Rank.rrvgam(object, ...)
```

Arguments

object	Some VGAM object, for example, having class <code>rrvglm-class</code> . The class <code>vglm-class</code> is not included since this is not based on reduced-rank regression.
...	Other possible arguments fed into the function later (used for added flexibility for the future).

Details

Regression models based on reduced-rank regression have a quantity called the *rank*, which is 1 or 2 or 3 etc. The smaller the value the more dimension reduction, so that there are fewer parameters. This function was not called `rank()` to avoid conflict with [rank](#).

Value

Returns an integer value, provided the rank of the model makes sense.

Note

This function has not been defined for VGLMs yet. It might refer to the rank of the VL model matrix, but for now this function should not be applied to [vglm](#) fits.

Author(s)

T. W. Yee.

See Also

RR-VGLMs are described in [rrvglm-class](#); QRR-VGLMs are described in [qrrvglm-class](#).

Examples

```
pneumo <- transform(pneumo, let = log(exposure.time),
                    x3 = runif(nrow(pneumo)))
(fit1 <- rrvglm(cbind(normal, mild, severe) ~ let + x3,
               acat, data = pneumo))
coef(fit1, matrix = TRUE)
constraints(fit1)
Rank(fit1)
```

Rayleigh

Rayleigh Distribution

Description

Density, distribution function, quantile function and random generation for the Rayleigh distribution with parameter a .

Usage

```
drayleigh(x, scale = 1, log = FALSE)
prayleigh(q, scale = 1, lower.tail = TRUE, log.p = FALSE)
qrayleigh(p, scale = 1, lower.tail = TRUE, log.p = FALSE)
rrayleigh(n, scale = 1)
```

Arguments

<code>x, q</code>	vector of quantiles.
<code>p</code>	vector of probabilities.
<code>n</code>	number of observations. Fed into runif .
<code>scale</code>	the scale parameter b .
<code>log</code>	Logical. If <code>log = TRUE</code> then the logarithm of the density is returned.
<code>lower.tail, log.p</code>	Same meaning as in pnorm or qnorm .

Details

See [rayleigh](#), the **VGAM** family function for estimating the scale parameter b by maximum likelihood estimation, for the formula of the probability density function and range restrictions on the parameter b .

Value

`drayleigh` gives the density, `prayleigh` gives the distribution function, `qrayleigh` gives the quantile function, and `rrayleigh` generates random deviates.

Note

The Rayleigh distribution is related to the Maxwell distribution.

Author(s)

T. W. Yee and Kai Huang

References

Forbes, C., Evans, M., Hastings, N. and Peacock, B. (2011). *Statistical Distributions*, Hoboken, NJ, USA: John Wiley and Sons, Fourth edition.

See Also

[rayleigh](#), [maxwell](#).

Examples

```
## Not run: Scale <- 2; x <- seq(-1, 8, by = 0.1)
plot(x, drayleigh(x, scale = Scale), type = "l", ylim = c(0,1),
     las = 1, ylab = "",
     main = "Rayleigh density divided into 10 equal areas; red = CDF")
abline(h = 0, col = "blue", lty = 2)
qq <- qrayleigh(seq(0.1, 0.9, by = 0.1), scale = Scale)
lines(qq, drayleigh(qq, scale = Scale), col = 2, lty = 3, type = "h")
lines(x, prayleigh(x, scale = Scale), col = "red")
## End(Not run)
```

rayleigh

Rayleigh Regression Family Function

Description

Estimating the parameter of the Rayleigh distribution by maximum likelihood estimation. Right-censoring is allowed.

Usage

```
rayleigh(lscale = "loglink", nrfs = 1/3 + 0.01,
         oim.mean = TRUE, zero = NULL, parallel = FALSE,
         type.fitted = c("mean", "percentiles", "Qlink"),
         percentiles = 50)
cens.rayleigh(lscale = "loglink", oim = TRUE)
```

Arguments

lscale	Parameter link function applied to the scale parameter b . See Links for more choices. A log link is the default because b is positive.
nrfs	Numeric, of length one, with value in $[0, 1]$. Weighting factor between Newton-Raphson and Fisher scoring. The value 0 means pure Newton-Raphson, while 1 means pure Fisher scoring. The default value uses a mixture of the two algorithms, and retaining positive-definite working weights.
oim.mean	Logical, used only for intercept-only models. TRUE means the mean of the OIM elements are used as working weights. If TRUE then this argument has top priority for working out the working weights. FALSE means use another algorithm.
oim	Logical. For censored data only, TRUE means the Newton-Raphson algorithm, and FALSE means Fisher scoring.
zero, parallel	Details at CommonVGAMffArguments .
type.fitted, percentiles	See CommonVGAMffArguments for information. Using "Qlink" is for quantile-links in VGAMextra .

Details

The Rayleigh distribution, which is used in physics, has a probability density function that can be written

$$f(y) = y \exp(-0.5(y/b)^2)/b^2$$

for $y > 0$ and $b > 0$. The mean of Y is $b\sqrt{\pi/2}$ (returned as the fitted values) and its variance is $b^2(4 - \pi)/2$.

The **VGAM** family function `cens.rayleigh` handles right-censored data (the true value is greater than the observed value). To indicate which type of censoring, input `extra = list(rightcensored = vec2)` where `vec2` is a logical vector the same length as the response. If the component of this list is missing then the logical values are taken to be FALSE. The fitted object has this component stored in the `extra` slot.

The **VGAM** family function `rayleigh` handles multiple responses.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as `vglm`, `rrvglm` and `vgam`.

Warning

The theory behind the argument `oim` is not fully complete.

Note

The `poisson.points` family function is more general so that if `ostatistic = 1` and `dimension = 2` then it coincides with `rayleigh`. Other related distributions are the Maxwell and Weibull distributions.

Author(s)

T. W. Yee

References

Forbes, C., Evans, M., Hastings, N. and Peacock, B. (2011). *Statistical Distributions*, Hoboken, NJ, USA: John Wiley and Sons, Fourth edition.

See Also

`Rayleigh`, `genrayleigh`, `riceff`, `maxwell`, `weibullR`, `poisson.points`, `simulate.vlm`.

Examples

```
nn <- 1000; Scale <- exp(2)
rdata <- data.frame(ystar = rrayleigh(nn, scale = Scale))
fit <- vglm(ystar ~ 1, rayleigh, data = rdata, trace = TRUE)
head(fitted(fit))
with(rdata, mean(ystar))
coef(fit, matrix = TRUE)
Coef(fit)

# Censored data
rdata <- transform(rdata, U = runif(nn, 5, 15))
rdata <- transform(rdata, y = pmin(U, ystar))
## Not run: par(mfrow = c(1, 2))
hist(with(rdata, ystar)); hist(with(rdata, y))
## End(Not run)
extra <- with(rdata, list(rightcensored = ystar > U))
fit <- vglm(y ~ 1, cens.rayleigh, data = rdata, trace = TRUE,
           extra = extra, crit = "coef")
table(fit@extra$rightcen)
coef(fit, matrix = TRUE)
head(fitted(fit))
```

Description

Rearrange the rows and columns of the input so that the first row and first column are baseline. This function is for rank-zero row-column interaction models (RCIMs; i.e., general main effects models).

Usage

```
Rcim(mat, rbaseline = 1, cbaseline = 1)
```

Arguments

`mat` Matrix, of dimension r by c . It is best that it is labelled with row and column names.

`rbaseline, cbaseline` Numeric (row number of the matrix `mat`) or character (matching a row name of `mat`) that the user wants as the row baseline or reference level. Similarly `cbaseline` for the column.

Details

This is a data preprocessing function for `rcim`. For rank-zero row-column interaction models this function establishes the baseline (or reference) levels of the matrix response with respect to the row and columns—these become the new first row and column.

Value

Matrix of the same dimension as the input, with `rbaseline` and `cbaseline` specifying the first rows and columns. The default is no change in `mat`.

Note

This function is similar to `moffset`; see `moffset` for information about the differences. If numeric, the arguments `rbaseline` and `cbaseline` differ from arguments `roffset` and `coffset` in `moffset` by 1 (when elements of the matrix agree).

Author(s)

Alfian F. Hadi and T. W. Yee.

See Also

[moffset](#), [rcim](#), [plotrcim0](#).

Examples

```
(alcoff.e <- moffset(alcoff, roffset = "6", postfix = "*"))
(aa <- Rcim(alcoff, rbaseline = "11", cbaseline = "Sun"))
(bb <- moffset(alcoff, "11", "Sun", postfix = "*"))
aa - bb # Note the difference!
```

rcqo

*Constrained Quadratic Ordination***Description**

Random generation for constrained quadratic ordination (CQO).

Usage

```
rcqo(n, p, S, Rank = 1,
     family = c("poisson", "negbinomial", "binomial-poisson",
                "Binomial-negbinomial", "ordinal-poisson",
                "Ordinal-negbinomial", "gamma2"),
     eq.maximums = FALSE, eq.tolerances = TRUE, es.optimums = FALSE,
     lo.abundance = if (eq.maximums) hi.abundance else 10,
     hi.abundance = 100, sd.latvar = head(1.5/2^(0:3), Rank),
     sd.optimums = ifelse(es.optimums, 1.5/Rank, 1) *
                   ifelse(scale.latvar, sd.latvar, 1),
     sd.tolerances = 0.25, Kvector = 1, Shape = 1,
     sqrt.arg = FALSE, log.arg = FALSE, rhox = 0.5, breaks = 4,
     seed = NULL, optimums1.arg = NULL, Crow1positive = TRUE,
     xmat = NULL, scale.latvar = TRUE)
```

Arguments

n	Number of sites. It is denoted by n below.
p	Number of environmental variables, including an intercept term. It is denoted by p below. Must be no less than $1 + R$ in value.
S	Number of species. It is denoted by S below.
Rank	The rank or the number of latent variables or true dimension of the data on the reduced space. This must be either 1, 2, 3 or 4. It is denoted by R .
family	What type of species data is to be returned. The first choice is the default. If binomial then a 0 means absence and 1 means presence. If ordinal then the breaks argument is passed into the breaks argument of <code>cut</code> . Note that either the Poisson or negative binomial distributions are used to generate binomial and ordinal data, and that an upper-case choice is used for the negative binomial distribution (this makes it easier for the user). If "gamma2" then this is the 2-parameter gamma distribution.
eq.maximums	Logical. Does each species have the same maximum? See arguments <code>lo.abundance</code> and <code>hi.abundance</code> .
eq.tolerances	Logical. Does each species have the same tolerance? If TRUE then the common value is 1 along every latent variable, i.e., all species' tolerance matrices are the order- R identity matrix.

es.optimums	Logical. Do the species have equally spaced optimums? If TRUE then the quantity $S^{1/R}$ must be an integer with value 2 or more. That is, there has to be an appropriate number of species in total. This is so that a grid of optimum values is possible in R -dimensional latent variable space in order to place the species' optimums. Also see the argument sd.tolerances.
lo.abundance, hi.abundance	Numeric. These are recycled to a vector of length S . The species have a maximum between lo.abundance and hi.abundance. That is, at their optimal environment, the mean abundance of each species is between the two componentwise values. If eq.maximums is TRUE then lo.abundance and hi.abundance must have the same values. If eq.maximums is FALSE then the logarithm of the maximums are uniformly distributed between $\log(\text{lo.abundance})$ and $\log(\text{hi.abundance})$.
sd.latvar	Numeric, of length R (recycled if necessary). Site scores along each latent variable have these standard deviation values. This must be a decreasing sequence of values because the first ordination axis contains the greatest spread of the species' site scores, followed by the second axis, followed by the third axis, etc.
sd.optimums	Numeric, of length R (recycled if necessary). If es.optimums = FALSE then, for the r th latent variable axis, the optimums of the species are generated from a normal distribution centered about 0. If es.optimums = TRUE then the S optimums are equally spaced about 0 along every latent variable axis. Regardless of the value of es.optimums, the optimums are then scaled to give standard deviation sd.optimums[r].
sd.tolerances	Logical. If eq.tolerances = FALSE then, for the r th latent variable, the species' tolerances are chosen from a normal distribution with mean 1 and standard deviation sd.tolerances[r]. However, the first species y1 has its tolerance matrix set equal to the order- R identity matrix. All tolerance matrices for all species are diagonal in this function. This argument is ignored if eq.tolerances is TRUE, otherwise it is recycled to length R if necessary.
Kvector	A vector of positive k values (recycled to length S if necessary) for the negative binomial distribution (see negbinomial for details). Note that a natural default value does not exist, however the default value here is probably a realistic one, and that for large values of μ one has $Var(Y) = \mu^2/k$ approximately.
Shape	A vector of positive λ values (recycled to length S if necessary) for the 2-parameter gamma distribution (see gamma2 for details). Note that a natural default value does not exist, however the default value here is probably a realistic one, and that $Var(Y) = \mu^2/\lambda$.
sqrt.arg	Logical. Take the square-root of the negative binomial counts? Assigning sqrt.arg = TRUE when family="negbinomial" means that the resulting species data can be considered very crudely to be approximately Poisson distributed. They will not integers in general but much easier (less numerical problems) to estimate using something like <code>cqo(..., family="poissonff")</code> .
log.arg	Logical. Take the logarithm of the gamma random variates? Assigning log.arg = TRUE when family="gamma2" means that the resulting species data can be considered very crudely to be approximately Gaussian distributed about its (quadratic) mean.
rhox	Numeric, less than 1 in absolute value. The correlation between the environmental variables. The correlation matrix is a matrix of 1's along the diagonal and

	rhox in the off-diagonals. Note that each environmental variable is normally distributed with mean 0. The standard deviation of each environmental variable is chosen so that the site scores have the determined standard deviation, as given by argument <code>sd.latvar</code> .
<code>breaks</code>	If <code>family</code> is assigned an ordinal value then this argument is used to define the cutpoints. It is fed into the <code>breaks</code> argument of <code>cut</code> .
<code>seed</code>	If given, it is passed into <code>set.seed</code> . This argument can be used to obtain reproducible results. If set, the value is saved as the "seed" attribute of the returned value. The default will not change the random generator state, and return <code>.Random.seed</code> as "seed" attribute.
<code>optimums1.arg</code>	If assigned and <code>Rank = 1</code> then these are the explicit optimums. Recycled to length <code>S</code> .
<code>Crow1positive</code>	See <code>qrrvglm.control</code> for details.
<code>xmat</code>	The n by $p - 1$ environmental matrix can be inputted.
<code>scale.latvar</code>	Logical. If <code>FALSE</code> the argument <code>sd.latvar</code> is ignored and no scaling of the latent variable values is performed.

Details

This function generates data coming from a constrained quadratic ordination (CQO) model. In particular, data coming from a *species packing model* can be generated with this function. The species packing model states that species have equal tolerances, equal maximums, and optimums which are uniformly distributed over the latent variable space. This can be achieved by assigning the arguments `es.optimums = TRUE`, `eq.maximums = TRUE`, `eq.tolerances = TRUE`.

At present, the Poisson and negative binomial abundances are generated first using `lo.abundance` and `hi.abundance`, and if `family` is binomial or ordinal then it is converted into these forms.

In CQO theory the n by p matrix X is partitioned into two parts X_1 and X_2 . The matrix X_2 contains the 'real' environmental variables whereas the variables in X_1 are just for adjustment purposes; they contain the intercept terms and other variables that one wants to adjust for when (primarily) looking at the variables in X_2 . This function has X_1 only being a matrix of ones, i.e., containing an intercept only.

Value

A n by $p - 1 + S$ data frame with components and attributes. In the following the attributes are labelled with double quotes.

`x2, x3, x4, . . . , xp`

The environmental variables. This makes up the n by $p - 1$ X_2 matrix. Note that `x1` is not present; it is effectively a vector of ones since it corresponds to an intercept term when `cqo` is applied to the data.

`y1, y2, x3, . . . , yS`

The species data. This makes up the n by S matrix Y . This will be of the form described by the argument `family`.

"concoefficients"

The $p - 1$ by R matrix of constrained coefficients (or canonical coefficients). These are also known as weights or loadings.

"formula"	The formula involving the species and environmental variable names. This can be used directly in the formula argument of <code>cqo</code> .
"log.maximums"	The S -vector of species' maximums, on a log scale. These are uniformly distributed between $\log(\text{lo.abundance})$ and $\log(\text{hi.abundance})$.
"latvar"	The n by R matrix of site scores. Each successive column (latent variable) has sample standard deviation equal to successive values of <code>sd.latvar</code> .
"eta"	The linear/additive predictor value.
"optimums"	The S by R matrix of species' optimums.
"tolerances"	The S by R matrix of species' tolerances. These are the square root of the diagonal elements of the tolerance matrices (recall that all tolerance matrices are restricted to being diagonal in this function).

Other attributes are "break", "family", "Rank", "lo.abundance", "hi.abundance", "eq.tolerances", "eq.maximums", "seed" as used.

Note

This function is under development and is not finished yet. There may be a few bugs.

Yet to do: add an argument that allows absences to be equal to the first level if ordinal data is requested.

Author(s)

T. W. Yee

References

Yee, T. W. (2004). A new technique for maximum-likelihood canonical Gaussian ordination. *Ecological Monographs*, **74**, 685–701.

Yee, T. W. (2006). Constrained additive ordination. *Ecology*, **87**, 203–213.

ter Braak, C. J. F. and Prentice, I. C. (1988). A theory of gradient analysis. *Advances in Ecological Research*, **18**, 271–317.

See Also

`cqo`, `qrrvglm.control`, `cut`, `binomialff`, `poissonff`, `negbinomial`, `gamma2`.

Examples

```
## Not run:
# Example 1: Species packing model:
n <- 100; p <- 5; S <- 5
mydata <- rcqo(n, p, S, es.opt = TRUE, eq.max = TRUE)
names(mydata)
(myform <- attr(mydata, "formula"))
fit <- cqo(myform, poissonff, mydata, Bestof = 3) # eq.tol = TRUE
matplot(attr(mydata, "latvar"), mydata[, -(1:(p-1))], col = 1:S)
persp(fit, col = 1:S, add = TRUE)
```

```

lvplot(fit, lcol = 1:S, y = TRUE, pcol = 1:S) # Same plot as above

# Compare the fitted model with the 'truth'
concoef(fit) # The fitted model
attr(mydata, "concoefficients") # The 'truth'

c(apply(attr(mydata, "latvar"), 2, sd),
  apply(latvar(fit), 2, sd)) # Both values should be approx equal

# Example 2: negative binomial data fitted using a Poisson model:
n <- 200; p <- 5; S <- 5
mydata <- rcqo(n, p, S, fam = "negbin", sqrt = TRUE)
myform <- attr(mydata, "formula")
fit <- cqo(myform, fam = poissonff, dat = mydata) # I.tol = TRUE,
lvplot(fit, lcol = 1:S, y = TRUE, pcol = 1:S)
# Compare the fitted model with the 'truth'
concoef(fit) # The fitted model
attr(mydata, "concoefficients") # The 'truth'

## End(Not run)

```

rdiric

The Dirichlet distribution

Description

Generates Dirichlet random variates.

Usage

```
rdiric(n, shape, dimension = NULL, is.matrix.shape = FALSE)
```

Arguments

n	number of observations. Note it has two meanings, see <code>is.matrix.shape</code> below.
shape	the shape parameters. These must be positive. If dimension is specified, values are recycled if necessary to length dimension.
dimension	the dimension of the distribution. If dimension is not numeric then it is taken to be <code>length(shape)</code> (or <code>ncol(shape)</code> if <code>is.matrix.shape == TRUE</code>).
<code>is.matrix.shape</code>	Logical. If TRUE then shape must be a matrix, and then n is no longer the number of rows of the answer but the answer has <code>n * nrow(shape)</code> rows. If FALSE (the default) then shape is a vector and each of the n rows of the answer have shape as its shape parameters.

Details

This function is based on a relationship between the gamma and Dirichlet distribution. Random gamma variates are generated, and then Dirichlet random variates are formed from these.

Value

A n by dimension matrix of Dirichlet random variates. Each element is positive, and each row will sum to unity. If shape has names then these will become the column names of the answer.

Author(s)

Thomas W. Yee

References

Lange, K. (2002). *Mathematical and Statistical Methods for Genetic Analysis*, 2nd ed. New York: Springer-Verlag.

See Also

[dirichlet](#) is a **VGAM** family function for fitting a Dirichlet distribution to data.

Examples

```
ddata <- data.frame(rdiric(n = 1000, shape = c(y1 = 3, y2 = 1, y3 = 4)))
fit <- vglm(cbind(y1, y2, y3) ~ 1, dirichlet, data = ddata, trace = TRUE)
Coef(fit)
coef(fit, matrix = TRUE)
```

 rec.exp1

Upper Record Values from a 1-parameter Exponential Distribution

Description

Maximum likelihood estimation of the rate parameter of a 1-parameter exponential distribution when the observations are upper record values.

Usage

```
rec.exp1(lrate = "loglink", irate = NULL, imethod = 1)
```

Arguments

lrate	Link function applied to the rate parameter. See Links for more choices.
irate	Numeric. Optional initial values for the rate. The default value NULL means they are computed internally, with the help of imethod.
imethod	Integer, either 1 or 2 or 3. Initial method, three algorithms are implemented. Choose the another value if convergence fails, or use irate.

Details

The response must be a vector or one-column matrix with strictly increasing values.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Note

By default, this family function has the intercept-only MLE as the initial value, therefore convergence may only take one iteration. Fisher scoring is used.

Author(s)

T. W. Yee

References

Arnold, B. C. and Balakrishnan, N. and Nagaraja, H. N. (1998). *Records*, New York: John Wiley & Sons.

See Also

[exponential](#).

Examples

```
rawy <- rexp(n <- 10000, rate = exp(1))
y <- unique(cummax(rawy)) # Keep only the records

length(y) / y[length(y)] # MLE of rate

fit <- vglm(y ~ 1, rec.exp1, trace = TRUE)
coef(fit, matrix = TRUE)
Coef(fit)
```

rec.normal

Upper Record Values from a Univariate Normal Distribution

Description

Maximum likelihood estimation of the two parameters of a univariate normal distribution when the observations are upper record values.

Usage

```
rec.normal(lmean = "identitylink", lsd = "loglink",
           imean = NULL, isd = NULL, imethod = 1, zero = NULL)
```


Arguments

lmean, lsd	Link functions applied to the mean and sd parameters. See Links for more choices.
imean, isd	Numeric. Optional initial values for the mean and sd. The default value NULL means they are computed internally, with the help of imethod.
imethod	Integer, either 1 or 2 or 3. Initial method, three algorithms are implemented. Choose the another value if convergence fails, or use imean and/or isd.
zero	Can be an integer vector, containing the value 1 or 2. If so, the mean or standard deviation respectively are modelled as an intercept only. Usually, setting zero = 2 will be used, if used at all. The default value NULL means both linear/additive predictors are modelled as functions of the explanatory variables. See CommonVGAMffArguments for more information.

Details

The response must be a vector or one-column matrix with strictly increasing values.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Note

This family function tries to solve a difficult problem, and the larger the data set the better. Convergence failure can commonly occur, and convergence may be very slow, so set `maxit = 200`, `trace = TRUE`, say. Inputting good initial values are advised.

This family function uses the BFGS quasi-Newton update formula for the working weight matrices. Consequently the estimated variance-covariance matrix may be inaccurate or simply wrong! The standard errors must be therefore treated with caution; these are computed in functions such as `vcov()` and `summary()`.

Author(s)

T. W. Yee

References

Arnold, B. C. and Balakrishnan, N. and Nagaraja, H. N. (1998). *Records*, New York: John Wiley & Sons.

See Also

[uninormal](#), [double.cens.normal](#).

Examples

```

nn <- 10000; mymean <- 100
# First value is reference value or trivial record
Rdata <- data.frame(rawy = c(mymean, rnorm(nn, mymean, exp(3))))
# Keep only observations that are records:
rdata <- data.frame(y = unique(cummax(with(Rdata, rawy))))

fit <- vglm(y ~ 1, rec.normal, rdata, trace = TRUE, maxit = 200)
coef(fit, matrix = TRUE)
Coef(fit)
summary(fit)

```

reciprocallink

Reciprocal Link Function

Description

Computes the reciprocal transformation, including its inverse and the first two derivatives.

Usage

```

reciprocallink(theta, bvalue = NULL, inverse = FALSE, deriv = 0,
               short = TRUE, tag = FALSE)
negreciprocallink(theta, bvalue = NULL, inverse = FALSE, deriv = 0,
                  short = TRUE, tag = FALSE)

```

Arguments

theta Numeric or character. See below for further details.
bvalue See [Links](#).
inverse, deriv, short, tag
 Details at [Links](#).

Details

The `reciprocallink` link function is a special case of the power link function. Numerical values of θ close to 0 result in Inf , $-\text{Inf}$, NA or NaN.

The `negreciprocallink` link function computes the negative reciprocal, i.e., $-1/\theta$.

Value

For `reciprocallink`: for $\text{deriv} = 0$, the reciprocal of θ , i.e., $1/\theta$ when $\text{inverse} = \text{FALSE}$, and if $\text{inverse} = \text{TRUE}$ then $1/\theta$. For $\text{deriv} = 1$, then the function returns $d\theta / d\eta$ as a function of θ if $\text{inverse} = \text{FALSE}$, else if $\text{inverse} = \text{TRUE}$ then it returns the reciprocal.

Note

Numerical instability may occur when θ is close to 0.

Author(s)

Thomas W. Yee

References

McCullagh, P. and Nelder, J. A. (1989). *Generalized Linear Models*, 2nd ed. London: Chapman & Hall.

See Also

[identitylink](#), [powerlink](#).

Examples

```
reciprocallink(1:5)
reciprocallink(1:5, inverse = TRUE, deriv = 2)
negreciprocallink(1:5)
negreciprocallink(1:5, inverse = TRUE, deriv = 2)

x <- (-3):3
reciprocallink(x) # Has Inf
reciprocallink(x, bvalue = .Machine$double.eps) # Has no Inf
```

residualsvglm

Residuals for a VGLM fit

Description

Residuals for a vector generalized linear model (VGLM) object.

Usage

```
residualsvglm(object, type = c("working", "pearson", "response",
  "deviance", "ldot", "stdres", "rquantile"), matrix.arg = TRUE)
```

Arguments

object	Object of class "vglm", i.e., a vglm fit.
type	The value of this argument can be abbreviated. The type of residuals to be returned. The default is the first one: working residuals corresponding to the IRLS algorithm. These are defined for all models. They are sometimes added to VGAM plots of estimated component functions (see plotvgam). Pearson residuals for GLMs, when squared and summed over the data set, total to the Pearson chi-squared statistic. For VGLMs, Pearson residuals involve the working weight matrices and the score vectors. Under certain limiting conditions, Pearson residuals have 0 means and identity matrix as the variance-covariance matrix.

Response residuals are simply the difference between the observed values and the fitted values. Both have to be of the same dimension, hence not all families have response residuals defined.

Deviance residuals are only defined for models with a deviance function. They tend to GLMs mainly. This function returns a NULL for those models whose deviance is undefined.

Randomized quantile residuals (RQRs) (Dunn and Smyth, 1996) are based on the p-type function being fed into `qnorm`. For example, for the default `exponential` it is `qnorm(pexp(y, rate = 1 / fitted(object)))`. So one should expect these residuals to have a standard normal distribution if the model and data agree well. If the distribution is discrete then *randomized* values are returned; see `runif` and `set.seed`. For example, for the default `poissonff` it is `qnorm(runif(length(y), ppois(y - 1, mu), ppois(y, mu)))` where `mu` is the fitted mean. The following excerpts comes from their writings. They highly recommend quantile residuals for discrete distributions since plots using deviance and Pearson residuals may contain distracting patterns. Four replications of the quantile residuals are recommended with discrete distributions because they have a random component. Any features not preserved across all four sets of residuals are considered artifacts of the randomization. This type of residual is continuous even for discrete distributions; for both discrete and continuous distributions, the quantile residuals have an exact standard normal distribution.

The choice "ldot" should not be used currently.

Standardized residuals are currently only defined for 2 types of models: (i) GLMs (`poissonff`, `binomialff`); (ii) those fitted to a two-way table of counts, e.g., `cumulative`, `acat`, `multinomial`, `sratio`, `cratio`. For (ii), they are defined in Section 2.4.5 of Agresti (2018) and are also the output from the "stdres" component of `chisq.test`. For the test of independence they are a useful type of residual. Their formula is $(\text{observed} - \text{expected}) / \sqrt{V}$, where V is the residual cell variance (also see Agresti, 2007, section 2.4.5). When an independence null hypothesis is true, each standardized residual (corresponding to a cell in the table) has a large-sample standard normal distribution. Currently this function merely extracts the table of counts from `object` and then computes the standardized residuals like `chisq.test`.

`matrix.arg`

Logical, which applies when if the pre-processed answer is a vector or a 1-column matrix. If TRUE then the value returned will be a matrix, else a vector.

Details

This function returns various kinds of residuals, sometimes depending on the specific type of model having been fitted. Section 3.7 of Yee (2015) gives some details on several types of residuals defined for the VGLM class.

Standardized residuals for GLMs are described in Section 4.5.6 of Agresti (2013) as the ratio of the raw (response) residuals divided by their standard error. They involve the generalized hat matrix evaluated at the final IRLS iteration. When applied to the LM, standardized residuals for GLMs simplify to `rstandard`. For GLMs they are basically the Pearson residual divided by the square root of 1 minus the leverage.

Value

If that residual type is undefined or inappropriate or not yet implemented, then NULL is returned, otherwise a matrix or vector of residuals is returned.

Warning

This function may change in the future, especially those whose definitions may change.

References

- Agresti, A. (2007). *An Introduction to Categorical Data Analysis, 2nd ed.*, New York: John Wiley & Sons. Page 38.
- Agresti, A. (2013). *Categorical Data Analysis, 3rd ed.*, New York: John Wiley & Sons.
- Agresti, A. (2018). *An Introduction to Categorical Data Analysis, 3rd ed.*, New York: John Wiley & Sons.
- Dunn, P. K. and Smyth, G. K. (1996). Randomized quantile residuals. *Journal of Computational and Graphical Statistics*, **5**, 236–244.

See Also

[resid](#), [vglm](#), [chisq.test](#), [hatvalues](#).

Examples

```
pneumo <- transform(pneumo, let = log(exposure.time))
fit <- vglm(cbind(normal, mild, severe) ~ let, propodds, pneumo)
resid(fit) # Same as having type = "working" (the default)
resid(fit, type = "response")
resid(fit, type = "pearson")
resid(fit, type = "stdres") # Test for independence
```

rhobitlink

Rhobit Link Function

Description

Computes the rhobit link transformation, including its inverse and the first two derivatives.

Usage

```
rhobitlink(theta, bminvalue = NULL, bmaxvalue = NULL,
           inverse = FALSE, deriv = 0, short = TRUE, tag = FALSE)
```

Arguments

theta Numeric or character. See below for further details.
 bminvalue, bmaxvalue Optional boundary values, e.g., values of theta which are less than or equal to -1 can be replaced by bminvalue before computing the link function value. And values of theta which are greater than or equal to 1 can be replaced by bmaxvalue before computing the link function value. See [Links](#).
 inverse, deriv, short, tag Details at [Links](#).

Details

The rhobitlink link function is commonly used for parameters that lie between -1 and 1 . Numerical values of theta close to -1 or 1 or out of range result in Inf, -Inf, NA or NaN.

Value

For $\text{deriv} = 0$, the rhobit of theta, i.e., $\log((1 + \text{theta})/(1 - \text{theta}))$ when $\text{inverse} = \text{FALSE}$, and if $\text{inverse} = \text{TRUE}$ then $(\exp(\text{theta}) - 1)/(\exp(\text{theta}) + 1)$.

For $\text{deriv} = 1$, then the function returns $d \text{ eta} / d \text{ theta}$ as a function of theta if $\text{inverse} = \text{FALSE}$, else if $\text{inverse} = \text{TRUE}$ then it returns the reciprocal.

Note

Numerical instability may occur when theta is close to -1 or 1 . One way of overcoming this is to use bminvalue, etc.

The correlation parameter of a standard bivariate normal distribution lies between -1 and 1 , therefore this function can be used for modelling this parameter as a function of explanatory variables.

The link function rhobitlink is very similar to [fisherzlink](#), e.g., just twice the value of [fisherzlink](#).

Author(s)

Thomas W. Yee

See Also

[Links](#), [binom2.rho](#), [fisherz](#).

Examples

```
theta <- seq(-0.99, 0.99, by = 0.01)
y <- rhobitlink(theta)
## Not run:
plot(theta, y, type = "l", ylab = "", main = "rhobitlink(theta)")
abline(v = 0, h = 0, lty = 2)

## End(Not run)

x <- c(seq(-1.02, -0.98, by = 0.01), seq(0.97, 1.02, by = 0.01))
```

```
rhobitlink(x) # Has NAs
rhobitlink(x, bminvalue = -1 + .Machine$double.eps,
           bmaxvalue = 1 - .Machine$double.eps) # Has no NAs
```

Rice

The Rice Distribution

Description

Density, distribution function, quantile function and random generation for the Rician distribution.

Usage

```
dprice(x, sigma, vee, log = FALSE)
prprice(q, sigma, vee, lower.tail = TRUE, log.p = FALSE, ...)
qrprice(p, sigma, vee, lower.tail = TRUE, log.p = FALSE, ...)
rrprice(n, sigma, vee)
```

Arguments

x, q	vector of quantiles.
p	vector of probabilities.
n	number of observations. Same as in runif .
vee, sigma	See riceff .
...	Other arguments such as <code>lower.tail</code> .
<code>lower.tail</code> , <code>log.p</code>	Same meaning as in pnorm or qnorm .
log	Logical. If <code>log = TRUE</code> then the logarithm of the density is returned.

Details

See [riceff](#), the **VGAM** family function for estimating the two parameters, for the formula of the probability density function and other details.

Formulas for `prprice()` and `qrprice()` are based on the Marcum-Q function.

Value

`dprice` gives the density, `prprice` gives the distribution function, `qrprice` gives the quantile function, and `rrprice` generates random deviates.

Author(s)

T. W. Yee and Kai Huang

See Also

[riceff](#).

Examples

```
## Not run: x <- seq(0.01, 7, len = 201)
plot(x, drice(x, vee = 0, sigma = 1), type = "n", las = 1,
      ylab = "",
      main = "Density of Rice distribution for various v values")
sigma <- 1; vee <- c(0, 0.5, 1, 2, 4)
for (ii in 1:length(vee))
  lines(x, drice(x, vee = vee[ii], sigma), col = ii)
legend(x = 5, y = 0.6, legend = as.character(vee),
       col = 1:length(vee), lty = 1)

x <- seq(0, 4, by = 0.01); vee <- 1; sigma <- 1
probs <- seq(0.05, 0.95, by = 0.05)
plot(x, drice(x, vee = vee, sigma = sigma), type = "l",
      main = "Blue is density, orange is CDF", col = "blue",
      ylim = c(0, 1), sub = "Red are 5, 10, ..., 95 percentiles",
      las = 1, ylab = "", cex.main = 0.9)
abline(h = 0:1, col = "black", lty = 2)
Q <- qrice(probs, sigma, vee = vee)
lines(Q, drice(qrice(probs, sigma, vee = vee),
               sigma, vee = vee), col = "red", lty = 3, type = "h")
lines(x, price(x, sigma, vee = vee), type = "l", col = "orange")
lines(Q, drice(Q, sigma, vee = vee), col = "red", lty = 3, type = "h")
lines(Q, price(Q, sigma, vee = vee), col = "red", lty = 3, type = "h")
abline(h = probs, col = "red", lty = 3)
max(abs(price(Q, sigma, vee = vee) - probs)) # Should be 0

## End(Not run)
```

riceff

Rice Distribution Family Function

Description

Estimates the two parameters of a Rice distribution by maximum likelihood estimation.

Usage

```
riceff(lsigma = "loglink", lvee = "loglink", isigma = NULL,
       ivee = NULL, nsimEIM = 100, zero = NULL, nowarning = FALSE)
```

Arguments

nowarning	Logical. Suppress a warning? Ignored for VGAM 0.9-7 and higher.
lvee, lsigma	Link functions for the v and σ parameters. See Links for more choices and for general information.
ivee, isigma	Optional initial values for the parameters. If convergence failure occurs (this VGAM family function seems to require good initial values) try using these arguments. See CommonVGAMffArguments for more information.
nsimEIM, zero	See CommonVGAMffArguments for information.

Details

The Rician distribution has density function

$$f(y; v, \sigma) = \frac{y}{\sigma^2} \exp(-(y^2 + v^2)/(2\sigma^2)) I_0(yv/\sigma^2)$$

where $y > 0$, $v > 0$, $\sigma > 0$ and I_0 is the modified Bessel function of the first kind with order zero. When $v = 0$ the Rice distribution reduces to a Rayleigh distribution. The mean is $\sigma \sqrt{\pi/2} \exp(z/2) ((1 - z)I_0(-z/2) - zI_1(-z/2))$ (returned as the fitted values) where $z = -v^2/(2\sigma^2)$. Simulated Fisher scoring is implemented.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Note

Convergence problems may occur for data where $v = 0$; if so, use [rayleigh](#) or possibly use an [identity](#) link.

When v is large (greater than 3, say) then the mean is approximately v and the standard deviation is approximately σ .

Author(s)

T. W. Yee

References

Rice, S. O. (1945). Mathematical Analysis of Random Noise. *Bell System Technical Journal*, **24**, 46–156.

See Also

[drice](#), [rayleigh](#), [bessell](#), [simulate.vlm](#).

Examples

```
## Not run: sigma <- exp(1); vee <- exp(2)
rdata <- data.frame(y = rrice(n <- 1000, sigma, vee = vee))
fit <- vglm(y ~ 1, riceff, data = rdata, trace = TRUE, crit = "c")
c(with(rdata, mean(y)), fitted(fit)[1])
coef(fit, matrix = TRUE)
Coef(fit)
summary(fit)

## End(Not run)
```

`rigff`*Reciprocal Inverse Gaussian distribution*

Description

Estimation of the parameters of a reciprocal inverse Gaussian distribution.

Usage

```
rigff(lmu = "identitylink", llambda = "loglink", imu = NULL,
      ilambda = 1)
```

Arguments

`lmu`, `llambda` Link functions for `mu` and `lambda`. See [Links](#) for more choices.
`imu`, `ilambda` Initial values for `mu` and `lambda`. A NULL means a value is computed internally.

Details

See Jorgensen (1997) for details.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Note

This distribution is potentially useful for dispersion modelling.

Author(s)

T. W. Yee

References

Jorgensen, B. (1997). *The Theory of Dispersion Models*. London: Chapman & Hall

See Also

[simplex](#).

Examples

```
rdata <- data.frame(y = rchisq(100, df = 14)) # Not 'proper' data!!
fit <- vglm(y ~ 1, rigff, rdata, trace = TRUE)
fit <- vglm(y ~ 1, rigff, rdata, trace = TRUE, crit = "c")
summary(fit)
```

rlplot.gevff *Return Level Plot for GEV Fits*

Description

A return level plot is constructed for a GEV-type model.

Usage

```
rlplot.gevff(object, show.plot = TRUE,
  probability = c((1:9)/100, (1:9)/10, 0.95, 0.99, 0.995, 0.999),
  add.arg = FALSE, xlab = if(log.arg) "Return Period (log-scale)" else
  "Return Period", ylab = "Return Level",
  main = "Return Level Plot",
  pch = par()$pch, pcol.arg = par()$col, pcex = par()$cex,
  llty.arg = par()$lty, lcol.arg = par()$col, llwd.arg = par()$lwd,
  slty.arg = par()$lty, scol.arg = par()$col, slwd.arg = par()$lwd,
  ylim = NULL, log.arg = TRUE, CI = TRUE, epsilon = 1e-05, ...)
```

Arguments

object	A VGAM extremes model of the GEV-type, produced by <code>vglm</code> with a family function either "gev" or "gevff".
show.plot	Logical. Plot it? If FALSE no plot will be done.
probability	Numeric vector of probabilities used.
add.arg	Logical. Add the plot to an existing plot?
xlab	Caption for the x-axis. See <code>par</code> .
ylab	Caption for the y-axis. See <code>par</code> .
main	Title of the plot. See <code>title</code> .
pch	Plotting character. See <code>par</code> .
pcol.arg	Color of the points. See the <code>col</code> argument of <code>par</code> .
pcex	Character expansion of the points. See the <code>cex</code> argument of <code>par</code> .
llty.arg	Line type. Line type. See the <code>lty</code> argument of <code>par</code> .
lcol.arg	Color of the lines. See the <code>col</code> argument of <code>par</code> .
llwd.arg	Line width. See the <code>lwd</code> argument of <code>par</code> .
slty.arg, scol.arg, slwd.arg	Corresponding arguments for the lines used for the confidence intervals. Used only if CI=TRUE.
ylim	Limits for the y-axis. Numeric of length 2.
log.arg	Logical. If TRUE then log="" otherwise log="x". This changes the labelling of the x-axis only.
CI	Logical. Add in a 95 percent confidence interval?

epsilon	Numeric, close to zero. Used for the finite-difference approximation to the first derivatives with respect to each parameter. If too small, numerical problems will occur.
...	Arguments passed into the plot function when setting up the entire plot. Useful arguments here include sub and las.

Details

A return level plot plots z_p versus $\log(y_p)$. It is linear if the shape parameter $\xi = 0$. If $\xi < 0$ then the plot is convex with asymptotic limit as p approaches zero at $\mu - \sigma/\xi$. And if $\xi > 0$ then the plot is concave and has no finite bound. Here, $G(z_p) = 1 - p$ where $0 < p < 1$ (p corresponds to the argument probability) and G is the cumulative distribution function of the GEV distribution. The quantity z_p is known as the *return level* associated with the *return period* $1/p$. For many applications, this means z_p is exceeded by the annual maximum in any particular year with probability p .

The points in the plot are the actual data.

Value

In the post slot of the object is a list called r1plot with list components

yp	$-\log(\text{probability})$, which is used on the x-axis.
zp	values which are used for the y-axis
lower, upper	lower and upper confidence limits for the 95 percent confidence intervals evaluated at the values of probability (if CI=TRUE).

Note

The confidence intervals are approximate, being based on finite-difference approximations to derivatives.

Author(s)

T. W. Yee

References

Coles, S. (2001). *An Introduction to Statistical Modeling of Extreme Values*. London: Springer-Verlag.

See Also

[gevff](#).

Examples

```

gdata <- data.frame(y = rgev(n <- 100, scale = 2, shape = -0.1))
fit <- vglm(y ~ 1, gevff, data = gdata, trace = TRUE)

# Identity link for all parameters:
fit2 <- vglm(y ~ 1, gevff(lshape = identitylink, lscale = identitylink,
                          iscale = 10), data = gdata, trace = TRUE)
coef(fit2, matrix = TRUE)
## Not run:
par(mfrow = c(1, 2))
rlplot(fit) -> i1
rlplot(fit2, pcol = "darkorange", lcol = "blue", log.arg = FALSE,
        scol = "darkgreen", slty = "dashed", las = 1) -> i2
range(i2$post$rlplot$upper - i1$post$rlplot$upper) # Should be near 0
range(i2$post$rlplot$lower - i1$post$rlplot$lower) # Should be near 0

## End(Not run)

```

rootogram4	<i>Rootograms (S4 generic) for Assessing Goodness of Fit of Probability Models</i>
------------	--

Description

A graphical technique for comparing the observed and fitted counts from a probability model, on a square root scale.

Usage

```

rootogram4(object, ...)
rootogram4vglm(object, newdata = NULL, breaks = NULL, max = NULL,
               xlab = NULL, main = NULL, width = NULL, ...)

```

Arguments

object	an object of class "vglm". This should include "vgam" because "vglm" handles both VGLM and VGAM objects.
newdata	Data upon which to base the calculations. The default is the one used to fit the model.
breaks	numeric. Breaks for the histogram intervals.
max	maximum count displayed. If an error message occurs regarding running out of memory then use this argument; it might occur with a very long tailed distribution such as gaitdzeta .
xlab, main	graphical parameters.
width	numeric. Widths of the histogram bars.
...	any additional arguments to <code>rootogram.default</code> and <code>plot.rootogram</code> in countreg . Probably the most useful of these are <code>style = c("hanging", "standing", "suspended")</code> and <code>scale = c("sqrt", "raw")</code> .

Details

Rootograms are a useful graphical technique for comparing the observed counts with the expected counts given a probability model.

This S4 implementation is based very heavily on rootogram coming from **countreg**. This package is primarily written by A. Zeileis and C. Kleiber. That package is currently on R-Forge but not CRAN, and it is based on S3. Since **VGAM** is written using S4, it was necessary to define an S4 generic function called `rootogram4()` which dispatches appropriately for S4 objects.

Currently, only a selected number of **VGAM** family functions are implemented. Over time, hopefully more and more will be completed.

Value

See rootogram in **countreg**; an object of class "rootogram0" inheriting from "data.frame" with about 8 variables.

Warning

This function is rudimentary and based totally on the implementation in **countreg**.

Note

The function names used coming from **countreg** have been renamed slightly to avoid conflict.

Author(s)

Package **countreg** is primarily written by A. Zeileis and C. Kleiber. Function `rootogram4()` is based very heavily on **countreg**. T. W. Yee wrote code to unpack variables from many various models and feed them into the appropriate d-type function.

References

Friendly, M. and Meyer, D. (2016). *Discrete Data Analysis with R: Visualization and Modeling Techniques for Categorical and Count Data*, Boca Raton, FL, USA: Chapman & Hall/CRC Press.

Kleiber, C. and Zeileis, A. (2016) "Visualizing Count Data Regressions Using Rootograms." *The American Statistician*, **70**(3), 296–303. doi:10.1080/00031305.2016.1173590.

Tukey, J. W. (1977) *Exploratory Data Analysis*, Reading, MA, USA: Addison-Wesley.

See Also

[vglm](#), [vgam](#), [glm](#), [zipoisson](#), [zapoisson](#), rootogram in **countreg**.

Examples

```
## Not run:
data("hspider", package = "VGAM") # Count responses
hs.p  <- vglm(Pardlugu ~ CoveHerb, poissonff, data = hspider)
hs.nb <- vglm(Pardlugu ~ CoveHerb, negbinomial, data = hspider)
hs.zip <- vglm(Pardlugu ~ CoveHerb, zipoisson, data = hspider)
hs.zap <- vglm(Pardlugu ~ CoveHerb, zapoisson, data = hspider)
```

```
opar <- par(mfrow = c(2, 2)) # Plot the rootograms
rootogram4(hs.p, max = 15, main = "poissonff")
rootogram4(hs.nb, max = 15, main = "negbinomial")
rootogram4(hs.zip, max = 15, main = "zipoisson")
rootogram4(hs.zap, max = 15, main = "zapoisson")
par(opar)

## End(Not run)
```

round2

Rounding of Numbers to Base 2

Description

'round2' works like 'round' but the rounding has base 2 under consideration so that bits (binary digits) beyond a certain threshold are zeroed.

Usage

```
round2(x, digits10 = 0)
```

Arguments

x	Same as round .
digits10	Same as digits in round . The "10" is to emphasize the usual base 10 used by humans.

Details

round2() is intended to allow reliable and safe for == comparisons provided both sides have the function applied to the same value of digits10. Internally a numeric has its binary representation (bits) past a certain point set to all 0s, while retaining a certain degree of accuracy. Algorithmically, x is multiplied by 2^{exponent} and then rounded, and then divided by 2^{exponent}. The value of exponent is approximately 3 * digits10 when digits10 is positive. If digits10 is negative then what is returned is round(x, digits10). The value of exponent guarantees that x has been rounded to at least digits10 decimal places (often around digits10 + 1 for safety).

Value

Something similar to [round](#).

Author(s)

T. W. Yee.

See Also

[round](#), [tobit](#).

Examples

```

set.seed(1); x <- sort(rcauchy(10))
x3 <- round2(x, 3)
x3 == round2(x, 3) # Supposed to be reliable (all TRUE)
rbind(x, x3) # Comparison
(x3[1] * 2^(0:9)) / 2^(0:9)
print((x3[1] * 2^(0:11)), digits = 14)

# Round to approx 1 d.p.
x1 <- round2(x, 1)
x1 == round2(x, 1) # Supposed to be reliable (all TRUE)
rbind(x, x1)
x1[8] == 0.75 # 3/4
print((x1[1] * 2^(0:11)), digits = 9)
seq(31) / 32

```

rrar

Nested Reduced-rank Autoregressive Models for Multiple Time Series

Description

Estimates the parameters of a nested reduced-rank autoregressive model for multiple time series.

Usage

```
rrar(Ranks = 1, coefstart = NULL)
```

Arguments

Ranks	Vector of integers: the ranks of the model. Each value must be at least one and no more than M , where M is the number of response variables in the time series. The length of Ranks is the <i>lag</i> , which is often denoted by the symbol L in the literature.
coefstart	Optional numerical vector of initial values for the coefficients. By default, the family function chooses these automatically.

Details

Full details are given in Ahn and Reinsel (1988). Convergence may be very slow, so setting `maxits = 50`, say, may help. If convergence is not obtained, you might like to try inputting different initial values.

Setting `trace = TRUE` in `vglm` is useful for monitoring the progress at each iteration.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as `vglm` and `vgam`.

Note

This family function should be used within `vglm` and not with `rrvglm` because it does not fit into the RR-VGLM framework exactly. Instead, the reduced-rank model is formulated as a VGLM!

A methods function `Coef.rrar`, say, has yet to be written. It would return the quantities Ak_1 , C , D , $\hat{\omega}$, Φ , etc. as slots, and then `show.Coeff.rrar` would also need to be written.

Author(s)

T. W. Yee

References

Ahn, S. and Reinsel, G. C. (1988). Nested reduced-rank autoregressive models for multiple time series. *Journal of the American Statistical Association*, **83**, 849–856.

See Also

[vglm](#), [grain.us](#).

Examples

```
## Not run:
year <- seq(1961 + 1/12, 1972 + 10/12, by = 1/12)
par(mar = c(4, 4, 2, 2) + 0.1, mfrow = c(2, 2))
for (ii in 1:4) {
  plot(year, grain.us[, ii], main = names(grain.us)[ii], las = 1,
        type = "l", xlab = "", ylab = "", col = "blue")
  points(year, grain.us[, ii], pch = "*", col = "blue")
}
apply(grain.us, 2, mean) # mu vector
cgrain <- scale(grain.us, scale = FALSE) # Center the time series only
fit <- vglm(cgrain ~ 1, rrar(Ranks = c(4, 1)), trace = TRUE)
summary(fit)

print(fit@misc$Ak1, digits = 2)
print(fit@misc$Cmatrices, digits = 3)
print(fit@misc$Dmatrices, digits = 3)
print(fit@misc$omegahat, digits = 3)
print(fit@misc$Phimatrices, digits = 2)

par(mar = c(4, 4, 2, 2) + 0.1, mfrow = c(4, 1))
for (ii in 1:4) {
  plot(year, fit@misc$Z[, ii], main = paste("Z", ii, sep = ""),
        type = "l", xlab = "", ylab = "", las = 1, col = "blue")
  points(year, fit@misc$Z[, ii], pch = "*", col = "blue")
}

## End(Not run)
```

rrvglm	<i>Fitting Reduced-Rank Vector Generalized Linear Models (RR-VGLMs) and Doubly Constrained RR-VGLMs (DRR-VGLMs)</i>
--------	---

Description

A *reduced-rank vector generalized linear model* (RR-VGLM) is fitted. RR-VGLMs are VGLMs but some of the constraint matrices are estimated. *Doubly constrained* RR-VGLMs (DRR-VGLMs) can also be fitted, and these provide structure for the two other outer product matrices.

Usage

```
rrvglm(formula, family = stop("'family' is unassigned"),
       data = list(), weights = NULL, subset = NULL,
       na.action = na.fail, etastart = NULL, mustart = NULL,
       coefstart = NULL, control = rrvglm.control(...),
       offset = NULL, method = "rrvglm.fit", model = FALSE,
       x.arg = TRUE, y.arg = TRUE, contrasts = NULL,
       constraints = NULL, extra = NULL, qr.arg = FALSE,
       smart = TRUE, ...)
```

Arguments

formula, family	See vglm .
weights, data	See vglm .
subset, na.action	See vglm .
etastart, mustart, coefstart	See vglm .
control	a list of parameters for controlling the fitting process. See rrvglm.control for details.
offset, model, contrasts	See vglm .
method	the method to be used in fitting the model. The default (and presently only) method <code>rrvglm.fit</code> uses iteratively reweighted least squares (IRLS).
x.arg, y.arg	logical values indicating whether the model matrix and response vector/matrix used in the fitting process should be assigned in the x and y slots. Note the model matrix is the LM model matrix; to get the VGLM model matrix type <code>model.matrix(vglmfit)</code> where <code>vglmfit</code> is a <code>vglm</code> object.
constraints	See vglm .
extra, smart, qr.arg	See vglm .
...	further arguments passed into rrvglm.control .

Details

In this documentation, M is the number of linear predictors. For RR-VGLMs, the central formula is given by

$$\eta = B_1^T x_1 + A\nu$$

where x_1 is a vector (usually just a 1 for an intercept), x_2 is another vector of explanatory variables, and $\nu = C^T x_2$ is an R -vector of latent variables. Here, η is a vector of linear predictors, e.g., the m th element is $\eta_m = \log(E[Y_m])$ for the m th Poisson response. The dimension of η is M by definition. The matrices B_1 , A and C are estimated from the data, i.e., contain the regression coefficients. For ecologists, the central formula represents a *constrained linear ordination* (CLO) since it is linear in the latent variables. It means that the response is a monotonically increasing or decreasing function of the latent variables.

For identifiability it is common to enforce *corner constraints* on \mathbf{A} : by default, for RR-VGLMs, the top R by R submatrix is fixed to be the order- R identity matrix and the remainder of \mathbf{A} is estimated. And by default, for DRR-VGLMs, there is also an order- R identity matrix embedded in \mathbf{A} because the RRR must be *separable* (this is so that any existing structure in \mathbf{A} is preserved).

The underlying algorithm of RR-VGLMs is iteratively reweighted least squares (IRLS) with an optimizing algorithm applied within each IRLS iteration (e.g., alternating algorithm).

In theory, any **VGAM** family function that works for `vglm` and `vgam` should work for `rrvglm` too. The function that actually does the work is `rrvglm.fit`; it is essentially `vglm.fit` with some extra code.

Value

For RR-VGLMs, an object of class "rrvglm", which has the the same slots as a "vglm" object. The only difference is that the some of the constraint matrices are estimates rather than known. But **VGAM** stores the models the same internally. The slots of "vglm" objects are described in [vglm-class](#).

For DRR-VGLMs, an object of class "drrvglm".

Note

The arguments of `rrvglm` are in general the same as those of `vglm` but with some extras in `rrvglm.control`.

The smart prediction (`smartpred`) library is packed with the **VGAM** library.

In an example below, a rank-1 *stereotype* (reduced-rank multinomial logit) model of Anderson (1984) is fitted to some car data. The reduced-rank regression is performed, adjusting for two covariates. Setting a trivial constraint matrix (`diag(M)`) for the latent variable variables in x_2 avoids a warning message when it is overwritten by a (common) estimated constraint matrix. It shows that German cars tend to be more expensive than American cars, given a car of fixed weight and width.

If `fit <- rrvglm(..., data = mydata)` then `summary(fit)` requires corner constraints and no missing values in `mydata`. Sometimes the estimated variance-covariance matrix of the parameters is not positive-definite; if this occurs, try refitting the model with a different value for `Index.corner`.

For *constrained quadratic ordination* (CQO) see `cqo` for more details about QRR-VGLMs.

With multiple binary responses, one must use `binomialff(multiple.responses = TRUE)` to indicate that the response is a matrix with one response per column. Otherwise, it is interpreted as a single binary response variable.

To fit DRR-VGLMs see the arguments `H.A.thy` and `H.C` in `rrvglm.control`. DRR-VGLMs provide structure to the **A** and **C** matrices via constraint matrices. So instead of them being general unstructured matrices, one can make specified elements to be identically equal to 0, for example. This gives greater control over what is modelled as a latent variable, e.g., in a health study, if one subset of the covariates are physical variables and the remainder are psychological variables then a rank-2 model might have each latent variable a linear combination of each of the types of variables separately.

Incidentally before I forget, since `Corner = TRUE`, then the differences between the `@H.A.thy` and `@H.A.alt` slots are due to `Index.corner`, which specifies which rows of **A** are not estimated. However, in the alternating algorithm, it is more efficient to estimate the entire **A**, bar (effectively) rows `str0`, and then normalize it. In contrast, optimizing over the subset of **A** to be estimated is slow.

In the `@misc` slot are logical components `is.drrvglm` and `is.rrvglm`. Only one is `TRUE`. If `is.rrvglm` then (full) corner constraints are used. If `is.drrvglm` then *restricted corner constraints* (RCCs) are used and the reduced rank regression (RRR) must be *separable*. The case `is.rrvglm` means that `H.A.thy` is a vector("list", Rank) with `H.A.thy[[r]] <- diag(M)` assigned to all $r = 1, \dots, R$. Because DRR-VGLMs are implemented only for separable problems, this means that all columns of `H.A.thy[[s]]` are orthogonal to all columns from `H.A.thy[[t]]`, for all s and t . DRR-VGLMs are proposed in Yee et al. (2024) in the context of GAITD regression for heaped and seeped survey data.

Author(s)

Thomas W. Yee

References

- Yee, T. W. and Hastie, T. J. (2003). Reduced-rank vector generalized linear models. *Statistical Modelling*, **3**, 15–41.
- Yee, T. W. (2004). A new technique for maximum-likelihood canonical Gaussian ordination. *Ecological Monographs*, **74**, 685–701.
- Anderson, J. A. (1984). Regression and ordered categorical variables. *Journal of the Royal Statistical Society, Series B, Methodological*, **46**, 1–30.
- Yee, T. W. (2014). Reduced-rank vector generalized linear models with two linear predictors. *Computational Statistics and Data Analysis*, **71**, 889–902.
- Yee, T. W., Frigau, L. and Ma, C. (2024). Heaping and seeping, GAITD regression and doubly constrained reduced rank vector generalized linear models, in smoking studies. *In preparation*.

See Also

`rrvglm.control`, `summary.drrvglm`, `lvplot.rrvglm` (same as `biplot.rrvglm`), `rrvglm-class`, `grc`, `cqo`, `vglmff-class`, `vglm`, `vglm-class`, `smartpred`, `rrvglm.fit`. Special family functions include `negbinomial`, `zipoisson` and `zinegbinomial`. (see Yee (2014) and what was formerly in **COZIGAM**). Methods functions include `Coef.rrvglm`, `calibrate.rrvglm`, etc. Data include `crashi`.

Examples

```

## Not run:
# Example 1: RR NB with  $\text{Var}(Y) = \mu + \delta_1 * \mu^{\delta_2}$ 
nn <- 1000      # Number of observations
delta1 <- 3.0   # Specify this
delta2 <- 1.5   # Specify this; should be greater than 1
a21 <- 2 - delta2
mydata <- data.frame(x2 = runif(nn), x3 = runif(nn))
mydata <- transform(mydata, mu = exp(2 + 3 * x2 + 0 * x3))
mydata <- transform(mydata,
  y2 = rnbinom(nn, mu = mu, size = (1/delta1)*mu^a21))
plot(y2 ~ x2, mydata, pch = "+", col = 'blue', las = 1,
  main = paste0("Var(Y) = mu + ", delta1, " * mu^", delta2))
rrnb2 <- rrvglm(y2 ~ x2 + x3, negbinomial(zero = NULL),
  data = mydata, trace = TRUE)

a21.hat <- (Coef(rrnb2)@A)["loglink(size)", 1]
beta11.hat <- Coef(rrnb2)@B1["(Intercept)", "loglink(mu)"]
beta21.hat <- Coef(rrnb2)@B1["(Intercept)", "loglink(size)"]
(delta1.hat <- exp(a21.hat * beta11.hat - beta21.hat))
(delta2.hat <- 2 - a21.hat)
# delta1.hat:
# exp(a21.hat * predict(rrnb2)[1,1] - predict(rrnb2)[1,2])
summary(rrnb2)

# Obtain a 95 percent CI for delta2:
se.a21.hat <- sqrt(vcov(rrnb2)["I(latvar.mat)", "I(latvar.mat)"])
ci.a21 <- a21.hat + c(-1, 1) * 1.96 * se.a21.hat
(ci.delta2 <- 2 - rev(ci.a21)) # The 95 percent CI

Confint(rrnb2) # Quick way to get it

# Plot the abundances and fitted values vs the latent variable
plot(y2 ~ latvar(rrnb2), data = mydata, col = "blue",
  xlab = "Latent variable", las = 1)
ooo <- order(latvar(rrnb2))
lines(fitted(rrnb2)[ooo] ~ latvar(rrnb2)[ooo], col = "red")

# Example 2: stereotype model (RR multinomial logit model)
data(car.all)
scar <- subset(car.all,
  is.element(Country, c("Germany", "USA", "Japan", "Korea")))
fcols <- c(13,14,18:20,22:26,29:31,33,34,36) # These are factors
scar[, -fcols] <- scale(scar[, -fcols]) # Stdize all numerical vars
ones <- CM.ones(3) # matrix(1, 3, 1)
clist <- list("(Intercept)" = diag(3), Width = ones, Weight = ones,
  Disp. = diag(3), Tank = diag(3), Price = diag(3),
  Frt.Leg.Room = diag(3))
set.seed(111)
fit <- rrvglm(Country ~ Width + Weight + Disp. + Tank +
  Price + Frt.Leg.Room,
  multinomial, data = scar, Rank = 2, trace = TRUE,

```

```

constraints = clist, noRRR = ~ 1 + Width + Weight,
# Uncor = TRUE, Corner = FALSE, # orig.
Index.corner = c(1, 3), # Less correlation
Bestof = 3)
fit@misc$deviance # A history of the fits
Coef(fit)
biplot(fit, chull = TRUE, scores = TRUE, clty = 2, Ccex = 2,
       ccol = "blue", scol = "orange", Ccol = "darkgreen",
       Clwd = 2, main = "1=Germany, 2=Japan, 3=Korea, 4=USA")

## End(Not run)

```

rrvglm-class

Class "rrvglm"

Description

Reduced-rank vector generalized linear models.

Objects from the Class

Objects can be created by calls to [rrvglm](#).

Slots

extra: Object of class "list"; the extra argument on entry to `vglm`. This contains any extra information that might be needed by the family function.

family: Object of class "vglmff". The family function.

iter: Object of class "numeric". The number of IRLS iterations used.

predictors: Object of class "matrix" with M columns which holds the M linear predictors.

assign: Object of class "list", from class "vlm". This named list gives information matching the columns and the (LM) model matrix terms.

call: Object of class "call", from class "vlm". The matched call.

coefficients: Object of class "numeric", from class "vlm". A named vector of coefficients.

constraints: Object of class "list", from class "vlm". A named list of constraint matrices used in the fitting.

contrasts: Object of class "list", from class "vlm". The contrasts used (if any).

control: Object of class "list", from class "vlm". A list of parameters for controlling the fitting process. See [vglm.control](#) for details.

criterion: Object of class "list", from class "vlm". List of convergence criterion evaluated at the final IRLS iteration.

df.residual: Object of class "numeric", from class "vlm". The residual degrees of freedom.

df.total: Object of class "numeric", from class "vlm". The total degrees of freedom.

dispersion: Object of class "numeric", from class "vlm". The scaling parameter.

effects: Object of class "numeric", from class "v1m". The effects.
fitted.values: Object of class "matrix", from class "v1m". The fitted values. This is usually the mean but may be quantiles, or the location parameter, e.g., in the Cauchy model.
misc: Object of class "list", from class "v1m". A named list to hold miscellaneous parameters.
model: Object of class "data.frame", from class "v1m". The model frame.
na.action: Object of class "list", from class "v1m". A list holding information about missing values.
offset: Object of class "matrix", from class "v1m". If non-zero, a M -column matrix of offsets.
post: Object of class "list", from class "v1m" where post-analysis results may be put.
preplot: Object of class "list", from class "v1m" used by `plotvgam`; the plotting parameters may be put here.
prior.weights: Object of class "matrix", from class "v1m" holding the initially supplied weights.
qr: Object of class "list", from class "v1m". QR decomposition at the final iteration.
R: Object of class "matrix", from class "v1m". The **R** matrix in the QR decomposition used in the fitting.
rank: Object of class "integer", from class "v1m". Numerical rank of the fitted model.
residuals: Object of class "matrix", from class "v1m". The *working* residuals at the final IRLS iteration.
ResSS: Object of class "numeric", from class "v1m". Residual sum of squares at the final IRLS iteration with the adjusted dependent vectors and weight matrices.
smart.prediction: Object of class "list", from class "v1m". A list of data-dependent parameters (if any) that are used by smart prediction.
terms: Object of class "list", from class "v1m". The `terms` object used.
weights: Object of class "matrix", from class "v1m". The weight matrices at the final IRLS iteration. This is in matrix-band form.
x: Object of class "matrix", from class "v1m". The model matrix (LM, not VGLM).
xlevels: Object of class "list", from class "v1m". The levels of the factors, if any, used in fitting.
y: Object of class "matrix", from class "v1m". The response, in matrix form.
Xm2: Object of class "matrix", from class "v1m". See `vg1m-class`.
Ym2: Object of class "matrix", from class "v1m". See `vg1m-class`.
callxm2: Object of class "call", from class "v1m". The matched call for argument `form2`.
A.est, C.est: Object of class "matrix". The estimates of **A** and **C**.

Extends

Class "vg1m", directly. Class "v1m", by class "vg1m".

Methods

biplot signature(`x = "rrvglm"`): biplot.

Coef signature(`object = "rrvglm"`): more detailed coefficients giving **A**, B_1 , **C**, etc.

biplot signature(`object = "rrvglm"`): biplot.

print signature(`x = "rrvglm"`): short summary of the object.

summary signature(`object = "rrvglm"`): a more detailed summary of the object.

Note

Two new slots for "rrvglm" were added compared to "vglm" objects, for **VGAM** 1.1-10. They are A.est and C.est.

Author(s)

Thomas W. Yee

References

Yee, T. W. and Hastie, T. J. (2003). Reduced-rank vector generalized linear models. *Statistical Modelling*, **3**, 15–41.

See Also

[rrvglm](#), [lvplot.rrvglm](#), [vglmff-class](#).

Examples

```
## Not run: # Rank-1 stereotype model of Anderson (1984)
pneumo <- transform(pneumo, let = log(exposure.time),
                    x3 = runif(nrow(pneumo))) # Noise
fit <- rrvglm(cbind(normal, mild, severe) ~ let + x3,
             multinomial, data = pneumo, Rank = 1)
Coef(fit)

## End(Not run)
```

rrvglm.control

Control Function for rrvglm()

Description

Algorithmic constants and parameters for running `rrvglm` are set using this function. Doubly constrained RR-VGLMs (DRR-VGLMs) are also catered for.

Usage

```
rrvglm.control(Rank = 1, Corner = TRUE,
              Index.corner = head(setdiff(seq(length(str0) +
              Rank), str0), Rank), noRRR = ~ 1, str0 = NULL,
              CrowIpositive = NULL, trace = FALSE, Bestof = 1,
              H.A.thy = list(), H.C = list(),
              Ainit = NULL, Cinit = NULL, sd.Cinit = 0.02,
              Algorithm = "alternating", Etamat.colmax = 10,
              noWarning = FALSE, Use.Init.Poisson.Q0 = FALSE,
              checkwz = TRUE, Check.rank = TRUE, Check.cm.rank = TRUE,
              wzepsilon = .Machine$double.eps^0.75, ...)
```


Arguments

Rank	The numerical rank R of the model. Must be an element from the set $\{1, 2, \dots, \min(M, p_2)\}$. Here, the vector of explanatory variables \mathbf{x} is partitioned into $(\mathbf{x}_1, \mathbf{x}_2)$, which is of dimension $p_1 + p_2$. The variables making up \mathbf{x}_1 are given by the terms in noRRR argument, and the rest of the terms comprise \mathbf{x}_2 .
Corner	Logical indicating whether corner constraints are to be used. Strongly recommended as the <i>only</i> method for fitting RR-VGLMs and DRR-VGLMs. This is one method for ensuring a unique solution and the availability of standard errors. If TRUE then Index.corner specifies the position of the corner constraints. Also used by DRR-VGLMs, and this is known as RCCs (see below).
Index.corner	Specifies the R rows of the corner constraints, e.g., they hold an order- R identity matrix for RR-VGLMs. The default are the first R positive integers that exclude str0. For DRR-VGLMs one needs to have (restricted) corner constraints. Then argument str0 will be ignored. If there are structural zeros then they should be built into the constraint matrices (see H.A.thy).
noRRR	Formula giving terms that are <i>not</i> to be included in the reduced-rank regression. That is, noRRR specifies which explanatory variables are in the \mathbf{x}_1 vector of rrvglm, and the rest go into \mathbf{x}_2 . I will write \mathbf{x}_1 as x_1 and \mathbf{x}_2 as x_2 hereafter. The x_1 variables constitute the \mathbf{B}_1 matrix in Yee and Hastie (2003). Those x_2 variables which are subject to the reduced-rank regression correspond to the \mathbf{B}_2 matrix. Set noRRR = NULL for the reduced-rank regression to be applied to every explanatory variable including the intercept.
str0	Integer vector specifying which rows of the estimated constraint matrices (\mathbf{A}) are to be all zeros. These are called <i>structural zeros</i> . Must not have any common value with Index.corner, and be a subset of the vector $1:M$. The default, str0 = NULL, means no structural zero rows at all. This argument is ignored by DRR-VGLMs.
Crow1positive	Currently this argument has no effect. In the future, it may be a logical vector of length Rank (recycled if necessary): are the elements of the first row of \mathbf{C} positive? Because of corner constraints (diag(Rank)) this argument currently must be NULL. Having a logical vector would mean some of the diagonals might be 1 and others -1 .
trace	Logical indicating if output should be produced for each iteration.
Bestof	Integer. The best of Bestof models fitted is returned. This argument helps guard against local solutions by (hopefully) finding the global solution from many fits. The argument works only when the function generates its own initial value for \mathbf{C} , i.e., when \mathbf{C} is <i>not</i> passed in as initial values.
H.A.thy, H.C	Lists. DRR-VGLMs are <i>Doubly constrained</i> RR-VGLMs where \mathbf{A} has Rank constraint matrices (one for each column) in a list called H.A.thy, and \mathbf{C} has a constraint matrix for each row, i.e., for each explanatory variable making up it. The class "drrvglm" may arise for such models. So H.C should be a named list of p_2 constraint matrices, each one for a different row of \mathbf{C} , i.e., p_2 is the number of variables making up the latent variable. Note that if H.C has <code>names</code> then matching is done with that, and the components of H.C are reordered if

they are not sorted according to the terms in formula. If they are not named, then their *order* is used, for example, `H.C[[1]]` and `H.C[[2]]` are taken as the constraint matrices for the first two variables of the latent variable(s).

Algorithm	Character string indicating what algorithm is to be used. The default is the first one. The choice "derivative" has been withdrawn in VGAM 1.1-10 .
Ainit, Cinit	Initial A and C matrices which may speed up convergence. They must be of the correct dimension.
sd.Cinit	Standard deviation of the initial values for the elements of C . These are normally distributed with mean zero. This argument is used only if <code>Use.Init.Poisson.Q0 = FALSE</code> .
Etamat.colmax	Positive integer, no smaller than Rank. Controls the amount of memory used by <code>.Init.Poisson.Q0()</code> . It is the maximum number of columns allowed for the pseudo-response and its weights. In general, the larger the value, the better the initial value. Used only if <code>Use.Init.Poisson.Q0 = TRUE</code> .
Use.Init.Poisson.Q0	Logical indicating whether the <code>.Init.Poisson.Q0()</code> should be used to obtain initial values for the C . The function uses a new method that can work well if the data are Poisson counts coming from an equal-tolerances QRR-VGLM (CQO). This option is less realistic for RR-VGLMs compared to QRR-VGLMs.
checkwz	logical indicating whether the diagonal elements of the working weight matrices should be checked whether they are sufficiently positive, i.e., greater than <code>wzepsilon</code> . If not, any values less than <code>wzepsilon</code> are replaced with this value.
noWarning, Check.rank, Check.cm.rank	Same as <code>vglm.control</code> . Ignored for VGAM 0.9-7 and higher.
wzepsilon	Small positive number used to test whether the diagonals of the working weight matrices are sufficiently positive.
...	Variables in ... are passed into <code>vglm.control</code> . If the derivative algorithm is used then ... are also passed into <code>rrvglm.optim.control</code> ; and if the alternating algorithm is used then ... are also passed into <code>valt0.control</code> .

In the above, R is the Rank and M is the number of linear predictors.

Details

VGAM supported three normalizations to ensure a unique solution. But currently, only corner constraints will work with summary of RR-VGLM and DRR-VGLM objects. **Update during late-2023/early-2024:** with ongoing work implementing the "drrvglm" class, there may be disruption and changes to other normalizations. However, corner constraints should be fully supported and have the greatest priority.

Value

A list with components matching the input names. Some error checking is done, but not much.

Note

In **VGAM** 1.1-11 and higher, the following arguments are no longer supported: `Wmat`, `Norrr`, `Svd.arg`, `Uncorrelated.latvar`, `scaleA`. Users should use corner constraints only.

The arguments in this function begin with an upper case letter to help avoid interference with those of [vglm.control](#).

In the example below a rank-1 *stereotype* model (Anderson, 1984) is fitted, however, the intercepts are completely unconstrained rather than sorted.

Author(s)

Thomas W. Yee

References

Yee, T. W. and Hastie, T. J. (2003). Reduced-rank vector generalized linear models. *Statistical Modelling*, **3**, 15–41.

See Also

[rrvglm](#), [rrvglm-class](#), [summary.drrvglm](#), [rrvglm.optim.control](#), [vglm](#), [vglm.control](#), [TypicalVGAMfamilyFunction](#), [CM.qnorm](#), [cqq](#).

Examples

```
## Not run:
set.seed(111)
pneumo <- transform(pneumo, let = log(exposure.time),
                   x3 = runif(nrow(pneumo))) # Unrelated
fit <- rrvglm(cbind(normal, mild, severe) ~ let + x3,
             multinomial, pneumo, Rank = 1, Index.corner = 2)
constraints(fit)
vcov(fit)
summary(fit)

## End(Not run)
```

`rrvglm.optim.control` *Control Function for rrvglm() Calling optim()*

Description

Algorithmic constants and parameters for running `optim` within `rrvglm` are set using this function.

Usage

```
rrvglm.optim.control(Fnscale = 1, Maxit = 100,
                   Switch.optimizer = 3, Abstol = -Inf,
                   Reltol = sqrt(.Machine$double.eps), ...)
```

Arguments

<code>Fnscale</code>	Passed into <code>optim</code> as <code>fnscale</code> .
<code>Maxit</code>	Passed into <code>optim</code> as <code>maxit</code> .
<code>Switch.optimizer</code>	Iteration number when the "Nelder-Mead" method of <code>optim</code> is switched to the quasi-Newton "BFGS" method. Assigning <code>Switch.optimizer</code> a negative number means always BFGS, while assigning <code>Switch.optimizer</code> a value greater than <code>maxits</code> means always use Nelder-Mead.
<code>Abstol</code>	Passed into <code>optim</code> as <code>abstol</code> .
<code>Reltol</code>	Passed into <code>optim</code> as <code>reltol</code> .
<code>...</code>	Ignored.

Details

See [optim](#) for more details.

Value

A list with components equal to the arguments.

Note

The transition between optimization methods may be unstable, so users may have to vary the value of `Switch.optimizer`.

Practical experience with `Switch.optimizer` shows that setting it to too large a value may lead to a local solution, whereas setting it to a low value will obtain the global solution. It appears that, if BFGS kicks in too late when the Nelder-Mead algorithm is starting to converge to a local solution, then switching to BFGS will not be sufficient to bypass convergence to that local solution.

Author(s)

Thomas W. Yee

See Also

[rrvglm.control](#), [optim](#).

ruge	<i>Rutherford-Geiger Polonium Data</i>
------	--

Description

Decay counts of polonium recorded by Rutherford and Geiger (1910).

Usage

```
data(ruge)
```

Format

This data frame contains the following columns:

counts a numeric vector, counts or frequencies

number a numeric vector, the number of decays

Details

These are the radioactive decay counts of polonium recorded by Rutherford and Geiger (1910) representing the number of scintillations in 2608 1/8 minute intervals. For example, there were 57 frequencies of zero counts. The counts can be thought of as being approximately Poisson distributed.

Source

Rutherford, E. and Geiger, H. (1910) The Probability Variations in the Distribution of alpha Particles, *Philosophical Magazine*, **20**, 698–704.

Examples

```
lambdahat <- with(ruge, weighted.mean(number, w = counts))
(N <- with(ruge, sum(counts)))
with(ruge, cbind(number, counts,
                 fitted = round(N * dpois(number, lambdahat))))
```

s

*Defining Smooths in VGAM Formulas***Description**

s is used in the definition of (vector) smooth terms within `vgam` formulas. This corresponds to 1st-generation VGAMs that use backfitting for their estimation. The effective degrees of freedom is prespecified.

Usage

```
s(x, df = 4, spar = 0, ...)
```

Arguments

x	covariate (abscissae) to be smoothed. Note that x must be a <i>single</i> variable and not a function of a variable. For example, <code>s(x)</code> is fine but <code>s(log(x))</code> will fail. In this case, let <code>logx <- log(x)</code> (in the data frame), say, and then use <code>s(logx)</code> . At this stage bivariate smoothers (x would be a two-column matrix) are not implemented.
df	numerical vector of length r . Effective degrees of freedom: must lie between 1 (linear fit) and n (interpolation). Thus one could say that $df-1$ is the <i>effective nonlinear degrees of freedom</i> (ENDF) of the smooth. Recycling of values will be used if <code>df</code> is not of length r . If <code>spar</code> is positive then this argument is ignored. Thus <code>s()</code> means that the effective degrees of freedom is prespecified. If it is known that the component function(s) are more wiggly than usual then try increasing the value of this argument.
spar	numerical vector of length r . Positive smoothing parameters (after scaling) . Larger values mean more smoothing so that the solution approaches a linear fit for that component function. A zero value means that <code>df</code> is used. Recycling of values will be used if <code>spar</code> is not of length r .
...	Ignored for now.

Details

In this help file M is the number of additive predictors and r is the number of component functions to be estimated (so that r is an element from the set $\{1,2,\dots,M\}$). Also, if n is the number of *distinct* abscissae, then `s` will fail if $n < 7$.

`s`, which is symbolic and does not perform any smoothing itself, only handles a single covariate. Note that `s` works in `vgam` only. It has no effect in `vglm` (actually, it is similar to the identity function `I` so that `s(x2)` is the same as `x2` in the LM model matrix). It differs from the `s()` of the `gam` package and the `s` of the `mgcv` package; they should not be mixed together. Also, terms involving `s` should be simple additive terms, and not involving interactions and nesting etc. For example, `myfactor:s(x2)` is not a good idea.

Value

A vector with attributes that are (only) used by `vgam`.

Note

The vector cubic smoothing spline which `s()` represents is computationally demanding for large M . The cost is approximately $O(nM^3)$ where n is the number of unique abscissae.

Currently a bug relating to the use of `s()` is that only constraint matrices whose columns are orthogonal are handled correctly. If any `s()` term has a constraint matrix that does not satisfy this condition then a warning is issued. See [is.buggy](#) for more information.

A more modern alternative to using `s` with `vgam` is to use `sm.os` or `sm.ps`. This does not require backfitting and allows automatic smoothing parameter selection. However, this alternative should only be used when the sample size is reasonably large (> 500 , say). These are called Generation-2 VGAMs.

Another alternative to using `s` with `vgam` is `bs` and/or `ns` with `vglm`. The latter implements half-stepping, which is helpful if convergence is difficult.

Author(s)

Thomas W. Yee

References

Yee, T. W. and Wild, C. J. (1996). Vector generalized additive models. *Journal of the Royal Statistical Society, Series B, Methodological*, **58**, 481–493.

See Also

[vgam](#), [is.buggy](#), [sm.os](#), [sm.ps](#), [vsmooth.spline](#).

Examples

```
# Nonparametric logistic regression
fit1 <- vgam(agaaus ~ s(altitude, df = 2), binomialff, data = hunua)
## Not run: plot(fit1, se = TRUE)

# Bivariate logistic model with artificial data
nn <- 300
bdata <- data.frame(x1 = runif(nn), x2 = runif(nn))
bdata <- transform(bdata,
  y1 = rbinom(nn, size = 1, prob = logitlink(sin(2 * x2), inverse = TRUE)),
  y2 = rbinom(nn, size = 1, prob = logitlink(sin(2 * x2), inverse = TRUE)))
fit2 <- vgam(cbind(y1, y2) ~ x1 + s(x2, 3), trace = TRUE,
  binom2.or(exchangeable = TRUE), data = bdata)
coef(fit2, matrix = TRUE) # Hard to interpret
## Not run: plot(fit2, se = TRUE, which.term = 2, scol = "blue")
```

 sc.studentt2

Scaled Student t Distribution with 2 df Family Function

Description

Estimates the location and scale parameters of a scaled Student t distribution with 2 degrees of freedom, by maximum likelihood estimation.

Usage

```
sc.studentt2(percentile = 50, llocation = "identitylink",
             lscale = "loglink", ilocation = NULL, iscale = NULL,
             imethod = 1, zero = "scale")
```

Arguments

`percentile` A numerical vector containing values between 0 and 100, which are the quantiles and expectiles. They will be returned as ‘fitted values’.

`llocation`, `lscale`

See [Links](#) for more choices, and [CommonVGAMffArguments](#).

`ilocation`, `iscale`, `imethod`, `zero`

See [CommonVGAMffArguments](#) for details.

Details

Koenker (1993) solved for the distribution whose quantiles are equal to its expectiles. Its canonical form has mean and mode at 0, and has a heavy tail (in fact, its variance is infinite).

The standard (“canonical”) form of this distribution can be endowed with a location and scale parameter. The standard form has a density that can be written as

$$f(z) = 2/(4 + z^2)^{3/2}$$

for real y . Then $z = (y - a)/b$ for location and scale parameters a and $b > 0$. The mean of Y is a . By default, $\eta_1 = a$ and $\eta_2 = \log(b)$. The expectiles/quantiles corresponding to `percentile` are returned as the fitted values; in particular, `percentile = 50` corresponds to the mean (0.5 expectile) and median (0.5 quantile).

Note that if Y has a standard `dsc.t2` then $Y = \sqrt{2}T_2$ where T_2 has a Student-t distribution with 2 degrees of freedom. The two parameters here can also be estimated using `studentt2` by specifying `df = 2` and making an adjustment for the scale parameter, however, this **VGAM** family function is more efficient since the EIM is known (Fisher scoring is implemented.)

Value

An object of class “`vglmff`” (see [vglmff-class](#)). The object is used by modelling functions such as `vglm`, `rrvglm` and `vgam`.

Author(s)

T. W. Yee

ReferencesKoenker, R. (1993). When are expectiles percentiles? (solution) *Econometric Theory*, **9**, 526–527.**See Also**[dsc.t2](#), [studentt2](#).**Examples**

```

set.seed(123); nn <- 1000
kdata <- data.frame(x2 = sort(runif(nn)))
kdata <- transform(kdata, mylocat = 1 + 3 * x2,
                  myscale = 1)
kdata <- transform(kdata, y = rsc.t2(nn, loc = mylocat, scale = myscale))
fit <- vglm(y ~ x2, sc.studentt2(perc = c(1, 50, 99)), data = kdata)
fit2 <- vglm(y ~ x2, studentt2(df = 2), data = kdata) # 'same' as fit

coef(fit, matrix = TRUE)
head(fitted(fit))
head(predict(fit))

# Nice plot of the results
## Not run: plot(y ~ x2, data = kdata, col = "blue", las = 1,
               sub = paste("n =", nn),
               main = "Fitted quantiles/expectiles using the sc.studentt2() distribution")
matplot(with(kdata, x2), fitted(fit), add = TRUE, type = "l", lwd = 3)
legend("bottomright", lty = 1:3, lwd = 3, legend = colnames(fitted(fit)),
      col = 1:3)
## End(Not run)

fit@extra$percentile # Sample quantiles

```

score.stat

*Rao's Score Test Statistics Evaluated at the Null Values***Description**

Generic function that computes Rao's score test statistics evaluated at the null values.

Usage

```

score.stat(object, ...)
score.stat.vlm(object, values0 = 0, subset = NULL, omit1s = TRUE,
              all.out = FALSE, orig.SE = FALSE, iterate.SE = TRUE,
              iterate.score = TRUE, trace = FALSE, ...)

```

Arguments

object, values0, subset	Same as in wald.stat.vlm .
omit1s, all.out	Same as in wald.stat.vlm .
orig.SE, iterate.SE	Same as in wald.stat.vlm .
iterate.score	Logical. The score vector is evaluated at one value of values0 and at other regression coefficient values. These other values may be either the MLE obtained from the original object (FALSE), else at values obtained by further IRLS iterations—this argument enables that choice.
trace	Same as in wald.stat.vlm .
...	Ignored for now.

Details

The (Rao) *score test* (also known as the *Lagrange multiplier test* in econometrics) is a third general method for hypothesis testing under a likelihood-based framework (the others are the likelihood ratio test and Wald test; see [lrt.stat](#) and [wald.stat](#)). Asymptotically, the three tests are equivalent. The Wald test is not invariant to parameterization, and the usual Wald test statistics computed at the estimates make it vulnerable to the Hauck-Donner effect (HDE; see [hdeff](#)). This function is similar to [wald.stat](#) in that one coefficient is set to 0 (by default) and the *other* coefficients are iterated by IRLS to get their MLE subject to this constraint. The SE is almost always based on the expected information matrix (EIM) rather than the OIM, and for some models the EIM and OIM coincide.

Value

By default the signed square root of the Rao score statistics are returned. If `all.out = TRUE` then a list is returned with the following components: `score.stat` the score statistic, `SE0` the standard error of that coefficient, `values0` the null values. Approximately, the default score statistics output are standard normal random variates if each null hypothesis is true.

Altogether, by the eight combinations of `iterate.SE`, `iterate.score` and `orig.SE`, there are six different variants of the Rao score statistic that can be returned because the score vector has 2 and the SEs have 3 subvariants.

Warning

See [wald.stat.vlm](#).

Author(s)

Thomas W. Yee

See Also

[wald.stat](#), [lrt.stat](#), [summaryvglm](#), [summary.glm](#), [anova.vglm](#), [vglm](#), [hdeff](#).

Examples

```

set.seed(1)
pneumo <- transform(pneumo, let = log(exposure.time),
                    x3 = rnorm(nrow(pneumo)))
(pfit <- vglm(cbind(normal, mild, severe) ~ let + x3, propodds, pneumo))
score.stat(pfit) # No HDE here; should be similar to the next line:
coef(summary(pfit))[, "z value"] # Wald statistics computed at the MLE
summary(pfit, score0 = TRUE)

```

seglines

*Hauck-Donner Effects: Segmented Lines Plot***Description**

Plots the piecewise segmented curve made up of Wald statistics versus estimates, using a colour code for the HDE severity.

Usage

```

seglines(x, y, dy, ddy, lwd = 2, cex = 2, plot.it = TRUE,
         add.legend = TRUE, cex.legend = 1,
         position.legend = "topleft", eta0 = NA, COPS0 = NA,
         lty.table = c("solid", "dashed", "solid", "dashed",
                       "solid", "dashed", "solid"),
         col.table = rainbow.sky[-5], pch.table = 7:1,
         severity.table = c("None", "Faint", "Weak", "Moderate",
                            "Strong", "Extreme", "Undetermined"), FYI = FALSE, ...)

```

Arguments

<code>x, y, dy, ddy</code>	Same as hdeffsev .
<code>lwd, cex</code>	Graphical parameters: line width, and character expansion.
<code>plot.it</code>	Logical, plot it? If FALSE then the other graphical arguments are ignored.
<code>add.legend, position.legend</code>	Logical and character; add a legend? The other argument is fed into legend .
<code>cex.legend</code>	Self-explanatory.
<code>severity.table, eta0, COPS0</code>	Same as hdeffsev .
<code>lty.table, col.table, pch.table</code>	Graphical parameters for the 7 different types of segments. Usually users should not assign anything to these arguments. Setting <code>pch.table = NULL</code> will suppress pch symbols from the legend.
<code>FYI, ...</code>	Should be ignored.

Details

This function was written to complement [hdeffsev](#) and is rough-and-ready. It plots the signed Wald statistics as a function of the estimates, and uses a colour-code to indicate the severity of the Hauck-Donner effect (HDE). This can be obtained from its first two derivatives.

Value

This function returns the severity of the HDE, possibly invisibly.

Note

This function is likely to change in the short future because it is experimental and far from complete.

Author(s)

Thomas W. Yee.

See Also

[hdeff](#), [hdeffsev](#).

Examples

```
deg <- 4 # myfun is a function that approximates the HDE
myfun <- function(x, deriv = 0) switch(as.character(deriv),
  '0' = x^deg * exp(-x),
  '1' = (deg * x^(deg-1) - x^deg) * exp(-x),
  '2' = (deg * (deg-1) * x^(deg-2) - 2*deg * x^(deg-1) +
    x^deg) * exp(-x))
## Not run:
curve(myfun, 0, 10, col = "white")
xgrid <- seq(0, 10, length = 101)
seglines(xgrid, myfun(xgrid), myfun(xgrid, deriv = 1),
  COPS0 = 2,
  myfun(xgrid, deriv = 2), pch.table = NULL,
  position = "bottom")

## End(Not run)
```

Select

Select Variables for a Formula Response or the RHS of a Formula

Description

Select variables from a data frame whose names begin with a certain character string.

Usage

```
Select(data = list(), prefix = "y",
        lhs = NULL, rhs = NULL, rhs2 = NULL, rhs3 = NULL,
        as.character = FALSE, as.formula.arg = FALSE, tilde = TRUE,
        exclude = NULL, sort.arg = TRUE)
```

Arguments

<code>data</code>	A data frame or a matrix.
<code>prefix</code>	A vector of character strings, or a logical. If a character then the variables chosen from <code>data</code> begin with the value of <code>prefix</code> . If a logical then only <code>TRUE</code> is accepted and all the variables in <code>data</code> are chosen.
<code>lhs</code>	A character string. The response of a formula.
<code>rhs</code>	A character string. Included as part of the RHS a formula. Set <code>rhs = "0"</code> to suppress the intercept.
<code>rhs2, rhs3</code>	Same as <code>rhs</code> but appended to its RHS, i.e., <code>paste0(rhs, " + ", rhs2, " + ", rhs3)</code> . If used, <code>rhs</code> should be used first, and then possibly <code>rhs2</code> and then possibly <code>rhs3</code> .
<code>as.character</code>	Logical. Return the answer as a character string?
<code>as.formula.arg</code>	Logical. Is the answer a formula?
<code>tilde</code>	Logical. If <code>as.character</code> and <code>as.formula.arg</code> are both <code>TRUE</code> then include the tilde in the formula?
<code>exclude</code>	Vector of character strings. Exclude these variables explicitly.
<code>sort.arg</code>	Logical. Sort the variables?

Details

This is meant as a utility function to avoid manually: (i) making a `cbind` call to construct a big matrix response, and (ii) constructing a formula involving a lot of terms. The savings can be made because the variables of interest begin with some prefix, e.g., with the character "y".

Value

If `as.character = FALSE` and `as.formula.arg = FALSE` then a matrix such as `cbind(y1, y2, y3)`.
 If `as.character = TRUE` and `as.formula.arg = FALSE` then a character string such as `"cbind(y1, y2, y3)"`.

If `as.character = FALSE` and `as.formula.arg = TRUE` then a [formula](#) such as `lhs ~ y1 + y2 + y3`.
 If `as.character = TRUE` and `as.formula.arg = TRUE` then a character string such as `"lhs ~ y1 + y2 + y3"`. See the examples below. By default, if no variables beginning the the value of `prefix` is found then a `NULL` is returned. Setting `prefix = " "` is a way of selecting no variables.

Note

This function is a bit experimental at this stage and may change in the short future. Some of its utility may be better achieved using [subset](#) and its `select` argument, e.g., `subset(pdata, TRUE, select = y01:y10)`.

For some models such as `posbernoulli.t` the order of the variables in the `xij` argument is crucial, therefore care must be taken with the argument `sort.arg`. In some instances, it may be good to rename variables `y1` to `y01`, `y2` to `y02`, etc. when there are variables such as `y14`.

Currently `subsetcol()` and `Select()` are identical. One of these functions might be withdrawn in the future.

Author(s)

T. W. Yee.

See Also

`vglm`, `cbind`, `subset`, `formula`, `fill1`.

Examples

```
Pneumo <- pneumo
colnames(Pneumo) <- c("y1", "y2", "y3", "x2") # The "y" variables are response
Pneumo$x1 <- 1; Pneumo$x3 <- 3; Pneumo$x <- 0; Pneumo$x4 <- 4 # Add these

Select(data = Pneumo) # Same as with(Pneumo, cbind(y1, y2, y3))
Select(Pneumo, "x")
Select(Pneumo, "x", sort = FALSE, as.char = TRUE)
Select(Pneumo, "x", exclude = "x1")
Select(Pneumo, "x", exclude = "x1", as.char = TRUE)
Select(Pneumo, c("x", "y"))
Select(Pneumo, "z") # Now returns a NULL
Select(Pneumo, " ") # Now returns a NULL
Select(Pneumo, prefix = TRUE, as.formula = TRUE)
Select(Pneumo, "x", exclude = c("x3", "x1"), as.formula = TRUE,
      lhs = "cbind(y1, y2, y3)", rhs = "0")
Select(Pneumo, "x", exclude = "x1", as.formula = TRUE, as.char = TRUE,
      lhs = "cbind(y1, y2, y3)", rhs = "0")

# Now a 'real' example:
Huggins89table1 <- transform(Huggins89table1, x3.tij = t01)
tab1 <- subset(Huggins89table1,
              rowSums(Select(Huggins89table1, "y")) > 0)

# Same as
# subset(Huggins89table1, y1 + y2 + y3 + y4 + y5 + y6 + y7 + y8 + y9 + y10 > 0)

# Long way to do it:
fit.th <-
  vglm(cbind(y01, y02, y03, y04, y05, y06, y07, y08, y09, y10) ~ x2 + x3.tij,
       xij = list(x3.tij ~ t01 + t02 + t03 + t04 + t05 + t06 + t07 + t08 +
                 t09 + t10 - 1),
       posbernoulli.t(parallel.t = TRUE ~ x2 + x3.tij),
       data = tab1, trace = TRUE,
       form2 = ~ x2 + x3.tij + t01 + t02 + t03 + t04 + t05 + t06 + t07 + t08 +
               t09 + t10)

# Short way to do it:
Fit.th <- vglm(Select(tab1, "y") ~ x2 + x3.tij,
```

```
xij = list(Select(tab1, "t", as.formula = TRUE,
                 sort = FALSE, lhs = "x3.tij", rhs = "0")),
posbernoulli.t(parallel.t = TRUE ~ x2 + x3.tij),
data = tab1, trace = TRUE,
form2 = Select(tab1, prefix = TRUE, as.formula = TRUE))
```

seq2binomial

*The Two-stage Sequential Binomial Distribution Family Function***Description**

Estimation of the probabilities of a two-stage binomial distribution.

Usage

```
seq2binomial(lprob1 = "logitlink", lprob2 = "logitlink",
             iprob1 = NULL, iprob2 = NULL,
             parallel = FALSE, zero = NULL)
```

Arguments

- lprob1, lprob2 Parameter link functions applied to the two probabilities, called p and q below. See [Links](#) for more choices.
- iprob1, iprob2 Optional initial value for the first and second probabilities respectively. A NULL means a value is obtained in the initialize slot.
- parallel, zero Details at [Links](#). If parallel = TRUE then the constraint also applies to the intercept. See [CommonVGAMffArguments](#) for details.

Details

This **VGAM** family function fits the model described by Crowder and Sweeting (1989) which is described as follows. Each of m spores has a probability p of germinating. Of the y_1 spores that germinate, each has a probability q of bending in a particular direction. Let y_2 be the number that bend in the specified direction. The probability model for this data is $P(y_1, y_2) =$

$$\binom{m}{y_1} p^{y_1} (1-p)^{m-y_1} \binom{y_1}{y_2} q^{y_2} (1-q)^{y_1-y_2}$$

for $0 < p < 1$, $0 < q < 1$, $y_1 = 1, \dots, m$ and $y_2 = 1, \dots, y_1$. Here, p is prob1, q is prob2.

Although the Authors refer to this as the *bivariate binomial* model, I have named it the *(two-stage) sequential binomial* model. Fisher scoring is used.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Note

The response must be a two-column matrix of sample proportions corresponding to y_1 and y_2 . The m values should be inputted with the weights argument of `vglm` and `vgam`. The fitted value is a two-column matrix of estimated probabilities p and q . A common form of error is when there are no trials for y_1 , e.g., if `mvector` below has some values which are zero.

Author(s)

Thomas W. Yee

References

Crowder, M. and Sweeting, T. (1989). Bayesian inference for a bivariate binomial distribution. *Biometrika*, **76**, 599–603.

See Also

[binomialff](#), [cfibrosis](#).

Examples

```
sdata <- data.frame(mvector = round(rnorm(nn <- 100, m = 10, sd = 2)),
                   x2 = runif(nn))
sdata <- transform(sdata, prob1 = logitlink(+2 - x2, inverse = TRUE),
                  prob2 = logitlink(-2 + x2, inverse = TRUE))
sdata <- transform(sdata, successes1 = rbinom(nn, size = mvector, prob = prob1))
sdata <- transform(sdata, successes2 = rbinom(nn, size = successes1, prob = prob2))
sdata <- transform(sdata, y1 = successes1 / mvector)
sdata <- transform(sdata, y2 = successes2 / successes1)
fit <- vglm(cbind(y1, y2) ~ x2, seq2binomial, weight = mvector,
           data = sdata, trace = TRUE)

coef(fit)
coef(fit, matrix = TRUE)
head(fitted(fit))
head(depvar(fit))
head(weights(fit, type = "prior")) # Same as with(sdata, mvector)
# Number of first successes:
head(depvar(fit)[, 1] * c(weights(fit, type = "prior")))
# Number of second successes:
head(depvar(fit)[, 2] * c(weights(fit, type = "prior"))) *
      depvar(fit)[, 1])
```

setup.smart

Smart Prediction Setup

Description

Sets up smart prediction in one of two modes: "write" and "read".

Usage

```
setup.smart(mode.arg, smart.prediction = NULL, max.smart = 30)
```

Arguments

mode.arg	mode.arg must be "write" or "read". If in "read" mode then smart.prediction must be assigned the data structure .smart.prediction that was created while fitting. This is stored in object@smart.prediction or object\$smart.prediction where object is the name of the fitted object.
smart.prediction	If in "read" mode then smart.prediction must be assigned the list of data dependent parameters, which is stored on the fitted object. Otherwise, smart.prediction is ignored.
max.smart	max.smart is the initial length of the list .smart.prediction. It is not important because .smart.prediction is made larger if needed.

Details

This function is only required by programmers writing a modelling function such as `lm` and `glm`, or a prediction functions of such, e.g., `predict.lm`. The function `setup.smart` operates by mimicking the operations of a first-in first-out stack (better known as a *queue*).

Value

Nothing is returned.

Side Effects

In "write" mode `.smart.prediction` in `smartpredenv` is assigned an empty list with `max.smart` components. In "read" mode `.smart.prediction` in `smartpredenv` is assigned `smart.prediction`. Then `.smart.prediction.counter` in `smartpredenv` is assigned the value 0, and `.smart.prediction.mode` and `.max.smart` are written to `smartpredenv` too.

See Also

`lm`, `predict.lm`.

Examples

```
## Not run:
setup.smart("write") # Put at the beginning of lm

## End(Not run)

## Not run: # Put at the beginning of predict.lm
setup.smart("read", smart.prediction = object$smart.prediction)

## End(Not run)
```

Simplex

*Simplex Distribution***Description**

Density function, and random generation for the simplex distribution.

Usage

```
dsimplex(x, mu = 0.5, dispersion = 1, log = FALSE)
rsimplex(n, mu = 0.5, dispersion = 1)
```

Arguments

x	Vector of quantiles. The support of the distribution is the interval (0, 1).
mu, dispersion	Mean and dispersion parameters. The former lies in the interval (0, 1) and the latter is positive.
n, log	Same usage as runif .

Details

The **VGAM** family function [simplex](#) fits this model; see that online help for more information. For `rsimplex()` the rejection method is used; it may be very slow if the density is highly peaked, and will fail if the density asymptotes at the boundary.

Value

`dsimplex(x)` gives the density function, `rsimplex(n)` gives n random variates.

Author(s)

T. W. Yee

See Also

[simplex](#).

Examples

```
sigma <- c(4, 2, 1) # Dispersion parameter
mymu <- c(0.1, 0.5, 0.7); xxx <- seq(0, 1, len = 501)
## Not run: par(mfrow = c(3, 3)) # Figure 2.1 of Song (2007)
for (iii in 1:3)
  for (jjj in 1:3) {
    plot(xxx, dsimplex(xxx, mymu[jjj], sigma[iii]),
         type = "l", col = "blue", xlab = "", ylab = "", main =
         paste("mu = ", mymu[jjj], ", sigma = ", sigma[iii], sep = "")) }
## End(Not run)
```

simplex

*Simplex Distribution Family Function***Description**

The two parameters of the univariate standard simplex distribution are estimated by full maximum likelihood estimation.

Usage

```
simplex(lmu = "logitlink", lsigma = "loglink", imu = NULL, isigma = NULL,
       imethod = 1, ishrinkage = 0.95, zero = "sigma")
```

Arguments

`lmu, lsigma` Link function for μ and σ . See [Links](#) for more choices.

`imu, isigma` Optional initial values for μ and σ . A NULL means a value is obtained internally.

`imethod, ishrinkage, zero`
See [CommonVGAMffArguments](#) for information.

Details

The probability density function can be written

$$f(y; \mu, \sigma) = [2\pi\sigma^2(y(1-y))^3]^{-0.5} \exp[-0.5(y-\mu)^2/(\sigma^2y(1-y)\mu^2(1-\mu)^2)]$$

for $0 < y < 1$, $0 < \mu < 1$, and $\sigma > 0$. The mean of Y is μ (called `mu`, and returned as the fitted values).

The second parameter, `sigma`, of this standard simplex distribution is known as the dispersion parameter. The unit variance function is $V(\mu) = \mu^3(1-\mu)^3$. Fisher scoring is applied to both parameters.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Note

This distribution is potentially useful for dispersion modelling. Numerical problems may occur when `mu` is very close to 0 or 1.

Author(s)

T. W. Yee

References

- Jorgensen, B. (1997). *The Theory of Dispersion Models*. London: Chapman & Hall
- Song, P. X.-K. (2007). *Correlated Data Analysis: Modeling, Analytics, and Applications*. Springer.

See Also

[dsimplex](#), [dirichlet](#), [rigff](#), [binomialff](#).

Examples

```
sdata <- data.frame(x2 = runif(nn <- 1000))
sdata <- transform(sdata, eta1 = 1 + 2 * x2,
                  eta2 = 1 - 2 * x2)
sdata <- transform(sdata, y = rsimplex(nn, mu = logitlink(eta1, inverse = TRUE),
                                     dispersion = exp(eta2)))
(fit <- vglm(y ~ x2, simplex(zero = NULL), data = sdata, trace = TRUE))
coef(fit, matrix = TRUE)
summary(fit)
```

simulate.vlm

Simulate Responses for VGLMs and VGAMs

Description

Simulate one or more responses from the distribution corresponding to a fitted model object.

Usage

```
## S3 method for class 'vlm'
simulate(object, nsim = 1, seed = NULL, ...)
```

Arguments

object	an object representing a fitted model. Usually an object of class <code>vglm-class</code> or <code>vgam-class</code> .
nsim, seed	Same as <code>simulate</code> .
...	additional optional arguments.

Details

This is a methods function for `simulate` and hopefully should behave in a very similar manner. Only **VGAM** family functions with a `simslot` slot have been implemented for `simulate`.

Value

Similar to `simulate`. Note that many **VGAM** family functions can handle multiple responses. This can result in a longer data frame with more rows (`nsim` multiplied by `n` rather than the ordinary `n`). In the future an argument may be available so that there is always `n` rows no matter how many responses were inputted.

Warning

With multiple response and/or multivariate responses, the order of the elements may differ. For some **VGAM** families, the order is $n \times N \times F$, where n is the sample size, N is `nsim` and F is `ncol(fitted(vglmObject))`. For other **VGAM** families, the order is $n \times F \times N$. An example of each is given below.

See Also

Currently the **VGAM** family functions with a `simslot` slot are: [alaplace1](#), [alaplace2](#), [betabinomial](#), [betabinomialff](#), [betaR](#), [betaff](#), [biamhcop](#), [bifrankcop](#), [bilogistic](#), [binomialff](#), [binormal](#), [binormalcop](#), [biclaytoncop](#), [cauchy](#), [cauchy1](#), [chisq](#), [dirichlet](#), [dagum](#), [erlang](#), [exponential](#), [bifgmcop](#), [fisk](#), [gamma1](#), [gamma2](#), [gammaR](#), [gengamma.stacy](#), [geometric](#), [gompertz](#), [gumbelII](#), [hzeta](#), [inv.lomax](#), [inv.paralogistic](#), [kumar](#), [lgamma1](#), [lgamma3](#), [lindley](#), [lino](#), [logff](#), [logistic1](#), [logistic](#), [lognormal](#), [lomax](#), [makeham](#), [negbinomial](#), [negbinomial.size](#), [paralogistic](#), [perks](#), [poissonff](#), [posnegbinomial](#), [posnormal](#), [pospoisson](#), [polya](#), [polyaR](#), [posbinomial](#), [rayleigh](#), [riceff](#), [simplex](#), [sinmad](#), [slash](#), [studentt](#), [studentt2](#), [studentt3](#), [triangle](#), [uninormal](#), [yulesimon](#), [zageometric](#), [zageometricff](#), [zanegbinomial](#), [zanegbinomialff](#), [zapoisson](#), [zapoissonff](#), [zigeometric](#), [zigeometricff](#), [zinegbinomial](#), [zipf](#), [zipoisson](#), [zipoissonff](#). Also, categorical family functions: [acat](#), [cratio](#), [sratio](#), [cumulative](#), [multinomial](#).

See also [RNG](#) about random number generation in R, [vglm](#), [vgam](#) for model fitting.

Examples

```
nn <- 10; mysize <- 20; set.seed(123)
bdata <- data.frame(x2 = rnorm(nn))
bdata <- transform(bdata,
  y1 = rbinom(nn, size = mysize, p = logitlink(1+x2, inverse = TRUE)),
  y2 = rbinom(nn, size = mysize, p = logitlink(1+x2, inverse = TRUE)),
  f1 = factor(as.numeric(rbinom(nn, size = 1,
    p = logitlink(1+x2, inverse = TRUE))))))
(fit1 <- vglm(cbind(y1, aaa = mysize - y1) ~ x2, # Matrix response (2-colns)
  binomialff, data = bdata))
(fit2 <- vglm(f1 ~ x2, binomialff, model = TRUE, data = bdata)) # Factor response

set.seed(123); simulate(fit1, nsim = 8)
set.seed(123); c(simulate(fit2, nsim = 3)) # Use c() when model = TRUE

# An n x N x F example
set.seed(123); n <- 100
bdata <- data.frame(x2 = runif(n), x3 = runif(n))
bdata <- transform(bdata, y1 = rnorm(n, 1 + 2 * x2),
  y2 = rnorm(n, 3 + 4 * x2))
fit1 <- vglm(cbind(y1, y2) ~ x2, binormal(eq.sd = TRUE), data = bdata)
nsim <- 1000 # Number of simulations for each observation
my.sims <- simulate(fit1, nsim = nsim)
dim(my.sims) # A data frame
aaa <- array(unlist(my.sims), c(n, nsim, ncol(fitted(fit1)))) # n by N by F
summary(rowMeans(aaa[, , 1]) - fitted(fit1)[, 1]) # Should be all 0s
summary(rowMeans(aaa[, , 2]) - fitted(fit1)[, 2]) # Should be all 0s
```

```

# An n x F x N example
n <- 100; set.seed(111); nsim <- 1000
zdata <- data.frame(x2 = runif(n))
zdata <- transform(zdata, lambda1 = loglink(-0.5 + 2 * x2, inverse = TRUE),
                  lambda2 = loglink( 0.5 + 2 * x2, inverse = TRUE),
                  pstr01 = logitlink( 0,          inverse = TRUE),
                  pstr02 = logitlink(-1.0,       inverse = TRUE))
zdata <- transform(zdata, y1 = rzipois(n, lambda = lambda1, pstr0 = pstr01),
                  y2 = rzipois(n, lambda = lambda2, pstr0 = pstr02))
zip.fit <- vglm(cbind(y1, y2) ~ x2, zipoissonff, data = zdata, crit = "coef")
my.sims <- simulate(zip.fit, nsim = nsim)
dim(my.sims) # A data frame
aaa <- array(unlist(my.sims), c(n, ncol(fitted(zip.fit)), nsim)) # n by F by N
summary(rowMeans(aaa[, 1, ]) - fitted(zip.fit)[, 1]) # Should be all 0s
summary(rowMeans(aaa[, 2, ]) - fitted(zip.fit)[, 2]) # Should be all 0s

```

Sinmad

*The Singh-Maddala Distribution***Description**

Density, distribution function, quantile function and random generation for the Singh-Maddala distribution with shape parameters a and q , and scale parameter $scale$.

Usage

```

dsinmad(x, scale = 1, shape1.a, shape3.q, log = FALSE)
psinmad(q, scale = 1, shape1.a, shape3.q, lower.tail = TRUE, log.p = FALSE)
qsinmad(p, scale = 1, shape1.a, shape3.q, lower.tail = TRUE, log.p = FALSE)
rsinmad(n, scale = 1, shape1.a, shape3.q)

```

Arguments

<code>x, q</code>	vector of quantiles.
<code>p</code>	vector of probabilities.
<code>n</code>	number of observations. If <code>length(n) > 1</code> , the length is taken to be the number required.
<code>shape1.a, shape3.q</code>	shape parameters.
<code>scale</code>	scale parameter.
<code>log</code>	Logical. If <code>log = TRUE</code> then the logarithm of the density is returned.
<code>lower.tail, log.p</code>	Same meaning as in pnorm or qnorm .

Details

See [sinmad](#), which is the **VGAM** family function for estimating the parameters by maximum likelihood estimation.

Value

dsinmad gives the density, psinmad gives the distribution function, qsinmad gives the quantile function, and rsinmad generates random deviates.

Note

The Singh-Maddala distribution is a special case of the 4-parameter generalized beta II distribution.

Author(s)

T. W. Yee and Kai Huang

References

Kleiber, C. and Kotz, S. (2003). *Statistical Size Distributions in Economics and Actuarial Sciences*, Hoboken, NJ, USA: Wiley-Interscience.

See Also

[sinmad](#), [genbetaII](#).

Examples

```
sdata <- data.frame(y = rsinmad(n = 3000, scale = exp(2),
                             shape1 = exp(1), shape3 = exp(1)))
fit <- vglm(y ~ 1, sinmad(lss = FALSE, ishape1.a = 2.1), data = sdata,
           trace = TRUE, crit = "coef")
coef(fit, matrix = TRUE)
Coef(fit)
```

sinmad

Singh-Maddala Distribution Family Function

Description

Maximum likelihood estimation of the 3-parameter Singh-Maddala distribution.

Usage

```
sinmad(lscale = "loglink", lshape1.a = "loglink", lshape3.q = "loglink",
       iscale = NULL, ishape1.a = NULL, ishape3.q = NULL, imethod = 1,
       lss = TRUE, gscale = exp(-5:5), gshape1.a = exp(-5:5),
       gshape3.q = exp(-5:5), probs.y = c(0.25, 0.5, 0.75),
       zero = "shape")
```

Arguments

- `lss` See [CommonVGAMffArguments](#) for important information.
- `lshape1.a`, `lscale`, `lshape3.q`
Parameter link functions applied to the (positive) parameters a , $scale$, and q . See [Links](#) for more choices.
- `iscale`, `ishape1.a`, `ishape3.q`, `imethod`, `zero`
See [CommonVGAMffArguments](#) for information. For `imethod = 2` a good initial value for `ishape3.q` is needed to obtain good estimates for the other parameters.
- `gscale`, `gshape1.a`, `gshape3.q`
See [CommonVGAMffArguments](#) for information.
- `probs.y` See [CommonVGAMffArguments](#) for information.

Details

The 3-parameter Singh-Maddala distribution is the 4-parameter generalized beta II distribution with shape parameter $p = 1$. It is known under various other names, such as the Burr XII (or just the Burr distribution), Pareto IV, beta-P, and generalized log-logistic distribution. More details can be found in Kleiber and Kotz (2003).

Some distributions which are special cases of the 3-parameter Singh-Maddala are the Lomax ($a = 1$), Fisk ($q = 1$), and paralogistic ($a = q$).

The Singh-Maddala distribution has density

$$f(y) = aqy^{a-1}/[b^a\{1 + (y/b)^a\}^{1+q}]$$

for $a > 0$, $b > 0$, $q > 0$, $y \geq 0$. Here, b is the scale parameter $scale$, and the others are shape parameters. The cumulative distribution function is

$$F(y) = 1 - [1 + (y/b)^a]^{-q}.$$

The mean is

$$E(Y) = b\Gamma(1 + 1/a)\Gamma(q - 1/a)/\Gamma(q)$$

provided $-a < 1 < aq$; these are returned as the fitted values. This family function handles multiple responses.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Note

See the notes in [genbetaII](#).

Author(s)

T. W. Yee

References

Kleiber, C. and Kotz, S. (2003). *Statistical Size Distributions in Economics and Actuarial Sciences*, Hoboken, NJ, USA: Wiley-Interscience.

See Also

[Sinmad](#), [genbetaII](#), [betaII](#), [dagum](#), [fisk](#), [inv.lomax](#), [lomax](#), [paralogistic](#), [inv.paralogistic](#), [simulate.vlm](#).

Examples

```
## Not run:
sdata <- data.frame(y = rsinmad(n = 1000, shape1 = exp(1),
                             scale = exp(2), shape3 = exp(0)))
fit <- vglm(y ~ 1, sinmad(lss = FALSE), sdata, trace = TRUE)
fit <- vglm(y ~ 1, sinmad(lss = FALSE, ishape1.a = exp(1)),
           sdata, trace = TRUE)
coef(fit, matrix = TRUE)
Coef(fit)
summary(fit)

# Harder problem (has the shape3.q parameter going to infinity)

set.seed(3)
sdata <- data.frame(y1 = rbeta(1000, 6, 6))
# hist(with(sdata, y1))
if (FALSE) {
# These struggle
  fit1 <- vglm(y1 ~ 1, sinmad(lss = FALSE), sdata, trace = TRUE)
  fit1 <- vglm(y1 ~ 1, sinmad(lss = FALSE), sdata, trace = TRUE,
              crit = "coef")
  Coef(fit1)
}
# Try this remedy:
fit2 <- vglm(y1 ~ 1, data = sdata, trace = TRUE, stepsize = 0.05, maxit = 99,
            sinmad(lss = FALSE, ishape3.q = 3, lshape3.q = "logloglink"))

coef(fit2, matrix = TRUE)
Coef(fit2)
## End(Not run)
```

Description

Density and random generation for the Skellam distribution.

Usage

```
dskellam(x, mu1, mu2, log = FALSE)
rskellam(n, mu1, mu2)
```

Arguments

x	vector of quantiles.
n	number of observations. Same as runif .
mu1, mu2	See skellam
.	
log	Logical; if TRUE, the logarithm is returned.

Details

See [skellam](#), the **VGAM** family function for estimating the parameters, for the formula of the probability density function and other details.

Value

`dskellam` gives the density, and `rskellam` generates random deviates.

Warning

Numerical problems may occur for data if μ_1 and/or μ_2 are large. The normal approximation for this case has not been implemented yet.

See Also

[skellam](#), [dpois](#).

Examples

```
## Not run: mu1 <- 1; mu2 <- 2; x <- (-7):7
plot(x, dskellam(x, mu1, mu2), type = "h", las = 1, col = "blue",
      main = paste("Density of Skellam distribution with mu1 = ", mu1,
                  " and mu2 = ", mu2, sep = ""))
## End(Not run)
```

skellam

*Skellam Distribution Family Function***Description**

Estimates the two parameters of a Skellam distribution by maximum likelihood estimation.

Usage

```
skellam(lmu1 = "loglink", lmu2 = "loglink", imu1 = NULL, imu2 = NULL,
        nsimEIM = 100, parallel = FALSE, zero = NULL)
```

Arguments

`lmu1, lmu2` Link functions for the μ_1 and μ_2 parameters. See [Links](#) for more choices and for general information.

`imu1, imu2` Optional initial values for the parameters. See [CommonVGAMffArguments](#) for more information. If convergence failure occurs (this **VGAM** family function seems to require good initial values) try using these arguments.

`nsimEIM, parallel, zero` See [CommonVGAMffArguments](#) for information. In particular, setting `parallel=TRUE` will constrain the two means to be equal.

Details

The Skellam distribution models the difference between two independent Poisson distributions (with means μ_j , say). It has density function

$$f(y; \mu_1, \mu_2) = \left(\frac{\mu_1}{\mu_2}\right)^{y/2} \exp(-\mu_1 - \mu_2) I_{|y|}(2\sqrt{\mu_1\mu_2})$$

where y is an integer, $\mu_1 > 0$, $\mu_2 > 0$. Here, I_v is the modified Bessel function of the first kind with order v .

The mean is $\mu_1 - \mu_2$ (returned as the fitted values), and the variance is $\mu_1 + \mu_2$. Simulated Fisher scoring is implemented.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Warning

This **VGAM** family function seems fragile and very sensitive to the initial values. Use very cautiously!!

Note

Numerical problems may occur for data if μ_1 and/or μ_2 are large.

References

Skellam, J. G. (1946). The frequency distribution of the difference between two Poisson variates belonging to different populations. *Journal of the Royal Statistical Society, Series A*, **109**, 296.

See Also

[dskellam](#), [dpois](#), [poissonff](#).

Examples

```
## Not run:
sdata <- data.frame(x2 = runif(nn <- 1000))
sdata <- transform(sdata, mu1 = exp(1 + x2), mu2 = exp(1 + x2))
sdata <- transform(sdata, y = rskellam(nn, mu1, mu2))
fit1 <- vglm(y ~ x2, skellam, data = sdata, trace = TRUE, crit = "coef")
fit2 <- vglm(y ~ x2, skellam(parallel = TRUE), data = sdata, trace = TRUE)
coef(fit1, matrix = TRUE)
coef(fit2, matrix = TRUE)
summary(fit1)
# Likelihood ratio test for equal means:
pchisq(2 * (logLik(fit1) - logLik(fit2)),
      df = df.residual(fit2) - df.residual(fit1), lower.tail = FALSE)
lrtest(fit1, fit2) # Alternative

## End(Not run)
```

 skewnorm

Skew-Normal Distribution

Description

Density and random generation for the univariate skew-normal distribution.

Usage

```
dskewnorm(x, location = 0, scale = 1, shape = 0, log = FALSE)
rskewnorm(n, location = 0, scale = 1, shape = 0)
```

Arguments

x	vector of quantiles.
n	number of observations. Same as runif .
location	The location parameter ξ . A vector.
scale	The scale parameter ω . A positive vector.

shape	The shape parameter. It is called α in skewnormal .
log	Logical. If log=TRUE then the logarithm of the density is returned.

Details

See [skewnormal](#), which currently only estimates the shape parameter. More generally here, $Z = \xi + \omega Y$ where Y has a standard skew-normal distribution (see [skewnormal](#)), ξ is the location parameter and ω is the scale parameter.

Value

dskewnorm gives the density, rskewnorm generates random deviates.

Note

The default values of all three parameters corresponds to the skew-normal being the standard normal distribution.

Author(s)

T. W. Yee

References

<http://tango.stat.unipd.it/SN>.

See Also

[skewnormal](#).

Examples

```
## Not run: N <- 200 # Grid resolution
shape <- 7; x <- seq(-4, 4, len = N)
plot(x, dskewnorm(x, shape = shape), type = "l", col = "blue", las = 1,
      ylab = "", lty = 1, lwd = 2)
abline(v = 0, h = 0, col = "grey")
lines(x, dnorm(x), col = "orange", lty = 2, lwd = 2)
legend("topleft", leg = c(paste("Blue = dskewnorm(x, ", shape, ")"),
                          "Orange = standard normal density"), lty = 1:2, lwd = 2,
      col = c("blue", "orange"))
## End(Not run)
```

skewnormal

*Univariate Skew-Normal Distribution Family Function***Description**

Maximum likelihood estimation of the shape parameter of a univariate skew-normal distribution.

Usage

```
skewnormal(lshape = "identitylink", ishape = NULL, nsimEIM = NULL)
```

Arguments

lshape, ishape, nsimEIM

See [Links](#) and [CommonVGAMffArguments](#).

Details

The univariate skew-normal distribution has a density function that can be written

$$f(y) = 2\phi(y)\Phi(\alpha y)$$

where α is the shape parameter. Here, ϕ is the standard normal density and Φ its cumulative distribution function. When $\alpha = 0$ the result is a standard normal distribution. When $\alpha = 1$ it models the distribution of the maximum of two independent standard normal variates. When the absolute value of the shape parameter increases the skewness of the distribution increases. The limit as the shape parameter tends to positive infinity results in the folded normal distribution or half-normal distribution. When the shape parameter changes its sign, the density is reflected about $y = 0$.

The mean of the distribution is $\mu = \alpha\sqrt{2/(\pi(1 + \alpha^2))}$ and these are returned as the fitted values. The variance of the distribution is $1 - \mu^2$. The Newton-Raphson algorithm is used unless the nsimEIM argument is used.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Warning

It is well known that the EIM of Azzalini's skew-normal distribution is singular for skewness parameter tending to zero, and thus produces influential problems.

Note

It is a good idea to use several different initial values to ensure that the global solution is obtained. This family function will be modified (hopefully soon) to handle a location and scale parameter too.

Author(s)

Thomas W. Yee

References

Azzalini, A. A. (1985). A class of distributions which include the normal. *Scandinavian Journal of Statistics*, **12**, 171–178.

Azzalini, A. and Capitanio, A. (1999). Statistical applications of the multivariate skew-normal distribution. *Journal of the Royal Statistical Society, Series B, Methodological*, **61**, 579–602.

See Also

[skewnorm](#), [uninormal](#), [foldnormal](#).

Examples

```
sdata <- data.frame(y1 = rskewnorm(nn <- 1000, shape = 5))
fit1 <- vglm(y1 ~ 1, skewnormal, data = sdata, trace = TRUE)
coef(fit1, matrix = TRUE)
head(fitted(fit1), 1)
with(sdata, mean(y1))
## Not run: with(sdata, hist(y1, prob = TRUE))
x <- with(sdata, seq(min(y1), max(y1), len = 200))
with(sdata, lines(x, dskewnorm(x, shape = Coef(fit1)), col = "blue"))
## End(Not run)

sdata <- data.frame(x2 = runif(nn))
sdata <- transform(sdata, y2 = rskewnorm(nn, shape = 1 + 2*x2))
fit2 <- vglm(y2 ~ x2, skewnormal, data = sdata, trace = TRUE, crit = "coef")
summary(fit2)
```

Slash

Slash Distribution

Description

Density function, distribution function, and random generation for the slash distribution.

Usage

```
dslash(x, mu = 0, sigma = 1, log = FALSE,
       smallno = .Machine$double.eps*1000)
pslash(q, mu = 0, sigma = 1, very.negative = -10000,
       lower.tail = TRUE, log.p = FALSE)
rslash(n, mu = 0, sigma = 1)
```

Arguments

<code>x, q</code>	vector of quantiles.
<code>n</code>	Same as <code>runif</code> .
<code>mu, sigma</code>	the mean and standard deviation of the univariate normal distribution.
<code>log</code>	Logical. If TRUE then the logarithm of the density is returned.
<code>very.negative</code>	Numeric, of length 1. A large negative value. For $(q-\mu)/\sigma$ values less than this, the value 0 is returned because <code>integrate</code> tends to fail. A warning is issued. Similarly, if $(q-\mu)/\sigma$ is greater than <code>abs(very.negative)</code> then 1 is returned with a warning.
<code>smallno</code>	See <code>slash</code> .
<code>lower.tail, log.p</code>	Same meaning as in <code>pnorm</code> or <code>qnorm</code> .

Details

See `slash`, the **VGAM** family function for estimating the two parameters by maximum likelihood estimation, for the formula of the probability density function and other details.

Function `pslash` uses a `for ()` loop and `integrate`, meaning it's very slow. It may also be inaccurate for extreme values of `q`, and returns with 1 or 0 values when too extreme compared to `very.negative`.

Value

`dslash` gives the density, and `pslash` gives the distribution function, `rslash` generates random deviates.

Note

`pslash` is very slow.

Author(s)

Thomas W. Yee and C. S. Chee

See Also

`slash`.

Examples

```
## Not run:
curve(dslash, col = "blue", ylab = "f(x)", -5, 5, ylim = c(0, 0.4), las = 1,
      main = "Standard slash, normal and Cauchy densities", lwd = 2)
curve(dnorm, col = "black", lty = 2, lwd = 2, add = TRUE)
curve(dcauchy, col = "orange", lty = 3, lwd = 2, add = TRUE)
legend("topleft", c("slash", "normal", "Cauchy"), lty = 1:3,
      col = c("blue", "black", "orange"), lwd = 2)
```



```
curve(pslash, col = "blue", -5, 5, ylim = 0:1)
pslash(c(-Inf, -20000, 20000, Inf)) # Gives a warning

## End(Not run)
```

 slash

Slash Distribution Family Function

Description

Estimates the two parameters of the slash distribution by maximum likelihood estimation.

Usage

```
slash(lmu = "identitylink", lsigma = "loglink",
      imu = NULL, isigma = NULL, gprobs.y = ppoints(8), nsimEIM = 250,
      zero = NULL, smallno = .Machine$double.eps*1000)
```

Arguments

lmu, lsigma	Parameter link functions applied to the μ and σ parameters, respectively. See Links for more choices.
imu, isigma	Initial values. A NULL means an initial value is chosen internally. See CommonVGAMffArguments for more information.
gprobs.y	Used to compute the initial values for mu. This argument is fed into the probs argument of quantile to construct a grid, which is used to evaluate the log-likelihood. This must have values between 0 and 1.
nsimEIM, zero	See CommonVGAMffArguments for information.
smallno	Small positive number, used to test for the singularity.

Details

The standard slash distribution is the distribution of the ratio of a standard normal variable to an independent standard uniform(0,1) variable. It is mainly of use in simulation studies. One of its properties is that it has heavy tails, similar to those of the Cauchy.

The general slash distribution can be obtained by replacing the univariate normal variable by a general normal $N(\mu, \sigma)$ random variable. It has a density that can be written as

$$f(y) = \begin{cases} 1/(2\sigma\sqrt{2\pi}) & \text{if } y = \mu, \\ 1 - \exp(-(((y - \mu)/\sigma)^2)/2))/(\sqrt{2\pi}\sigma((y - \mu)/\sigma)^2) & \text{if } y \neq \mu. \end{cases}$$

where μ and σ are the mean and standard deviation of the univariate normal distribution respectively.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Note

Fisher scoring using simulation is used. Convergence is often quite slow. Numerical problems may occur.

Author(s)

T. W. Yee and C. S. Chee

References

Johnson, N. L. and Kotz, S. and Balakrishnan, N. (1994). *Continuous Univariate Distributions*, 2nd edition, Volume 1, New York: Wiley.

Kafadar, K. (1982). A Biweight Approach to the One-Sample Problem *Journal of the American Statistical Association*, **77**, 416–424.

See Also

[rslash](#), [simulate.vlm](#).

Examples

```
## Not run:
sdata <- data.frame(y = rslash(n = 1000, mu = 4, sigma = exp(2)))
fit <- vglm(y ~ 1, slash, data = sdata, trace = TRUE)
coef(fit, matrix = TRUE)
Coef(fit)
summary(fit)

## End(Not run)
```

sloglink

Square root–Log Link Mixtures

Description

Computes some square root–log mixture link transformations, including their inverse and the first few derivatives.

Usage

```
sloglink(theta, bvalue = NULL, taumix.log = 1, tol = 1e-13,
  nmax = 99, inverse = FALSE, deriv = 0, short = TRUE,
  tag = FALSE, c10 = c(2, -2))
lcsloglink(theta, bvalue = NULL, pmix.log = 0.01, tol = 1e-13,
  nmax = 99, inverse = FALSE, deriv = 0, short = TRUE,
  tag = FALSE, c10 = c(2, -2))
```

Arguments

theta	Numeric or character. See below for further details.
bvalue	See Links .
taumix.log	Numeric, of length 1. Mixing parameter directed at loglink . Then $1 - \exp(-\text{taumix.log} * \text{theta})$ is used to weight sqrtlink . Thus a 0 value will result in loglink , and a very large numeric such as 1e4 should be roughly equivalent to sqrtlink over almost all of the parameter space.
pmix.log	Numeric, of length 1. Mixing probability assigned to loglink . Then $1 - \text{pmix.log}$ is used to weight sqrtlink . Thus a 0 value will result in sqrtlink and 1 is equivalent to loglink .
tol, nmax	Arguments fed into a function implementing a vectorized bisection method.
inverse, deriv, short, tag	Details at Links .
c10	See sqrtlink and loglink .

Details

For general information see [alogitlink](#).

Value

The following holds for the *linear combination* (LC) variant. For `deriv = 0`, $(1 - \text{pmix.log}) * \text{sqrtlink}(\mu, c10 = c10) + \text{pmix.log} * \text{loglink}(\mu)$ when `inverse = FALSE`, and if `inverse = TRUE` then a nonlinear equation is solved for μ , given η passed in as `theta`. For `deriv = 1`, then the function returns $d \eta / d \theta$ as a function of θ if `inverse = FALSE`, else if `inverse = TRUE` then it returns the reciprocal.

Warning

The default values for `taumix.log` and `pmix.log` may change in the future. The name and order of the arguments may change too.

Author(s)

Thomas W. Yee

See Also

[alogitlink](#), [sqrtlink](#), [loglink](#), [Links](#), [poissonff](#), [hdeff](#).

Examples

```
mu <- seq(0.01, 3, length = 10)
sloglink(mu)
max(abs(sloglink(sloglink(mu), inv = TRUE) - mu)) # 0?
```

sm.os

*Defining O'Sullivan Spline Smooths in VGAM Formulas***Description**

This function represents an O-spline smooth term in a `vgam` formula and confers automatic smoothing parameter selection.

Usage

```
sm.os(x, ..., ninknots = 6, spar = -1, o.order = 2,
      alg.ninknots = c("s", ".nknots.smspl")[1], all.knots = FALSE,
      ridge.adj = 1e-5, spillover = 0.01, maxspar = 1e12,
      outer.ok = FALSE, fixspar = FALSE)
```

Arguments

<code>x</code>	covariate (abscissae) to be smoothed. Also called the regressor. If the <code>xij</code> facility is used then these covariates are inputted via the <code>...</code> argument.
<code>...</code>	Used to accommodate the other $M - 1$ covariates when the <code>xij</code> facility is used. See Section 3.4.4 of Yee (2015) for something very similar. This argument, found in the second argument, means that the other argument names must be fully specified if used, e.g., <code>outer.ok</code> and not <code>outer</code> . See the example below. In the example below, the term in the main formula is <code>sm.os(gcost.air, gcost.trn, gcost.bus)</code> and one might be tempted to use something like <code>sm.os(gcost)</code> to represent that <code>xij</code> term. However, this is not recommended because <code>sm.os(gcost)</code> might not have the same number of columns as <code>sm.os(gcost.air, gcost.trn, gcost.bus)</code> etc. That is, it is best to select one of the diagonal elements of the block matrix to represent that term.
<code>ninknots</code>	numeric, the number of <i>interior</i> knots, called K below. The default is to use this value. If you want <code>alg.ninknots</code> to operate then assign <code>NULL</code> to this argument.
<code>alg.ninknots</code>	character. The algorithm used to determine the number of interior knots. Only used when <code>all.knots = FALSE</code> and <code>ninknots = NULL</code> . Note that <code>".nknots.smspl"</code> corresponds to the default of smooth.spline . The value <code>"s"</code> corresponds to the same algorithm as <code>s</code> .
<code>all.knots</code>	logical. If <code>TRUE</code> then all distinct points in <code>x</code> are used as the interior knots. If <code>FALSE</code> (default) then a subset of <code>x[]</code> is used, specifically <code>x[j]</code> where the <code>ninknots</code> indices are quantiles that are evenly spaced with respect to the argument <code>probs</code> —see quantile . If <code>all.knots = FALSE</code> and <code>ninknots = NULL</code> then the argument <code>alg.ninknots</code> is used to compute <code>ninknots</code> .
<code>spar, maxspar</code>	<code>spar</code> is a vector of smoothing parameters. Negative values mean that magic will choose initial values in order to do the optimization at each P-IRLS iteration. Positive values mean that they are used as initial values for magic . If <code>fixspar = TRUE</code> then <code>spar</code> should be assigned a vector of positive values (but having values less than <code>maxspar</code>); then the smoothing parameters will be fixed and magic will not be used.

o.order	The order of the O’Sullivan penalized spline. Any one value from 1:4 is acceptable. The degree of the spline is $2 * o.order - 1$, so that cubic splines are the default. Setting o.order = 1 results in a linear spline which is a piecewise linear function.
ridge.adj	small positive number to stabilize linear dependencies among B-spline bases.
spillover	small and positive proportion of the range used on the outside of the boundary values. This defines the endpoints a and b that cover the data x_i , i.e., we are interested in the interval $[a, b]$ which contains all the abscissae. The interior knots are strictly inside (a, b) .
outer.ok	Fed into the argument (by the same name) of splineDesign .
fixspar	logical. If TRUE then spar should be a vector with positive values and the smoothing parameters are fixed at those values. If FALSE then spar contains the initial values for the smoothing parameters, and magic is called to determine (hopefully) some good values for the smoothing parameters.

Details

This function is currently used by [vgam](#) to allow automatic smoothing parameter selection based on O-splines to minimize an UBRE quantity. In contrast, [s](#) operates by having a prespecified amount of smoothing, e.g., its `df` argument. When the sample size is reasonably large this function is recommended over [s](#) also because backfitting is not required. This function therefore allows 2nd-generation VGAMs to be fitted (called G2-VGAMs, or Penalized-VGAMs).

This function should only be used with [vgam](#). This function uses [quantile](#) to choose the knots, whereas [sm.ps](#) chooses equally-spaced knots. As Wand and Ormerod (2008) write, in most situations the differences will be minor, but it is possible for problems to arise for either strategy by constructing certain regression functions and predictor variable distributions. Any differences between O-splines and P-splines tend to be at the boundaries. O-splines have *natural boundary constraints* so that the solution is linear beyond the boundary knots.

Some arguments in decreasing order of precedence are: `all.knots`, `niknots`, `alg.niknots`.

Unlike [s](#), which is symbolic and does not perform any smoothing itself, this function does compute the penalized spline when used by [vgam](#)—it creates the appropriate columns of the model matrix. When this function is used within [vgam](#), automatic smoothing parameter selection is implemented by calling [magic](#) after the necessary link-ups are done.

By default this function centres the component function. This function is also *smart*; it can be used for smart prediction (Section 18.6 of Yee (2015)). Automatic smoothing parameter selection is performed using *performance-oriented iteration* whereby an optimization problem is solved at each IRLS iteration.

This function works better when the sample size is large, e.g., when in the hundreds, say.

Value

A matrix with attributes that are (only) used by [vgam](#). The number of rows of the matrix is `length(x)`. The number of columns is a function of the number of interior knots K and the order of the O-spline m : $K + 2m - 1$. In code, this is `niknots + 2 * o.order - 1`, or using [sm.ps](#)-like arguments, `ps.int + degree - 1` (where `ps.int` should be more generally interpreted as the number of intervals. The formula is the same as [sm.ps](#).) It transpires then that [sm.os](#) and [sm.ps](#) are very similar.

Warning

Being introduced into **VGAM** for the first time, this function (and those associated with it) should be used cautiously. Not all options are fully working or have been tested yet, and there are bound to be some bugs lurking around.

Note

This function is currently under development and may change in the future.

One might try using this function with `vglm` so as to fit a regression spline, however, the default value of `niknots` will probably be too high for most data sets.

Author(s)

T. W. Yee, with some of the essential R code coming from the appendix of Wand and Ormerod (2008).

References

Wand, M. P. and Ormerod, J. T. (2008). On semiparametric regression with O'Sullivan penalized splines. *Australian and New Zealand Journal of Statistics*, **50**(2): 179–198.

See Also

[vgam](#), [sm.ps](#), [s](#), [smartpred](#), [is.smart](#), [summarypvgam](#), [smooth.spline](#), [splineDesign](#), [bs](#), [magic](#).

Examples

```
sm.os(runif(20))

## Not run:
data("TravelMode", package = "AER") # Need to install "AER" first
air.df <- subset(TravelMode, mode == "air") # Form 4 smaller data frames
bus.df <- subset(TravelMode, mode == "bus")
trn.df <- subset(TravelMode, mode == "train")
car.df <- subset(TravelMode, mode == "car")
TravelMode2 <- data.frame(income      = air.df$income,
                          wait.air   = air.df$wait - car.df$wait,
                          wait.trn  = trn.df$wait - car.df$wait,
                          wait.bus   = bus.df$wait - car.df$wait,
                          gcost.air  = air.df$gcost - car.df$gcost,
                          gcost.trn  = trn.df$gcost - car.df$gcost,
                          gcost.bus  = bus.df$gcost - car.df$gcost,
                          wait       = air.df$wait) # Value is unimportant
TravelMode2$mode <- subset(TravelMode, choice == "yes")$mode # The response
TravelMode2 <- transform(TravelMode2, incom.air = income, incom.trn = 0,
                          incom.bus = 0)

set.seed(1)
TravelMode2 <- transform(TravelMode2,
                          junkx2 = runif(nrow(TravelMode2)))

tfit2 <-
```

```

vgam(mode ~ sm.os(gcost.air, gcost.trn, gcost.bus) + ns(junkx2, 4) +
      sm.os(incom.air, incom.trn, incom.bus) + wait ,
      crit = "coef",
      multinomial(parallel = FALSE ~ 1), data = TravelMode2,
      xij = list(sm.os(gcost.air, gcost.trn, gcost.bus) ~
                 sm.os(gcost.air, gcost.trn, gcost.bus) +
                 sm.os(gcost.trn, gcost.bus, gcost.air) +
                 sm.os(gcost.bus, gcost.air, gcost.trn),
                 sm.os(incom.air, incom.trn, incom.bus) ~
                 sm.os(incom.air, incom.trn, incom.bus) +
                 sm.os(incom.trn, incom.bus, incom.air) +
                 sm.os(incom.bus, incom.air, incom.trn),
                 wait ~ wait.air + wait.trn + wait.bus),
      form2 = ~ sm.os(gcost.air, gcost.trn, gcost.bus) +
                sm.os(gcost.trn, gcost.bus, gcost.air) +
                sm.os(gcost.bus, gcost.air, gcost.trn) +
                wait +
                sm.os(incom.air, incom.trn, incom.bus) +
                sm.os(incom.trn, incom.bus, incom.air) +
                sm.os(incom.bus, incom.air, incom.trn) +
                junkx2 + ns(junkx2, 4) +
                incom.air + incom.trn + incom.bus +
                gcost.air + gcost.trn + gcost.bus +
                wait.air + wait.trn + wait.bus)
par(mfrow = c(2, 2))
plot(tfit2, se = TRUE, lcol = "orange", scol = "blue", ylim = c(-4, 4))
summary(tfit2)

## End(Not run)

```

sm.ps

Defining Penalized Spline Smoother in VGAM Formulas

Description

This function represents a P-spline smooth term in a vgam formula and confers automatic smoothing parameter selection.

Usage

```

sm.ps(x, ..., ps.int = NULL, spar = -1, degree = 3, p.order = 2,
      ridge.adj = 1e-5, spillover = 0.01, maxspar = 1e12,
      outer.ok = FALSE, mux = NULL, fixspar = FALSE)

```

Arguments

x, ... See [sm.os](#).

ps.int	the number of equally-spaced B-spline intervals. Note that the number of knots is equal to $\text{ps.int} + 2 * \text{degree} + 1$. The default, signified by NULL, means that the maximum of the value 7 and degree is chosen. This usually means 6 interior knots for big data sets. However, if this is too high compared to the length of x, then some adjustment is made. In the case where mux is assigned a numerical value (suggestions: some value between 1 and 2) then $\text{ceiling}(\text{mux} * \log(\text{length}(\text{unique}(\text{x.index}))))$ is used, where x.index is the combined data. No matter what, the above is not guaranteed to work on every data set. This argument may change in the future. See also argument mux.
spar, maxspar	See sm.os .
mux	numeric. If given, then this argument multiplies $\log(\text{length}(\text{unique}(\text{x})))$ to obtain ps.int. If ps.int is given then this argument is ignored.
degree	degree of B-spline basis. Usually this will be 2 or 3; and the values 1 or 4 might possibly be used.
p.order	order of difference penalty (0 is the ridge penalty).
ridge.adj, spillover	See sm.os .
outer.ok, fixspar	See sm.os .

Details

This function can be used by [vgam](#) to allow automatic smoothing parameter selection based on P-splines and minimizing an UBRE quantity.

This function should only be used with [vgam](#) and is an alternative to [sm.os](#); see that function for some details that also apply here.

Value

A matrix with attributes that are (only) used by [vgam](#). The number of rows of the matrix is $\text{length}(\text{x})$ and the number of columns is $\text{ps.int} + \text{degree} - 1$. The latter is because the function is centred.

Warning

See [sm.os](#).

Note

This function is currently under development and may change in the future. In particular, the default for ps.int is subject to change.

Author(s)

B. D. Marx wrote the original function. Subsequent edits were made by T. W. Yee and C. Somchit.

References

Eilers, P. H. C. and Marx, B. D. (1996). Flexible smoothing with B-splines and penalties (with comments and rejoinder). *Statistical Science*, **11**(2): 89–121.

See Also

[sm.os](#), [s](#), [vgam](#), [smartpred](#), [is.smart](#), [summarypvgam](#), [splineDesign](#), [bs](#), [magic](#).

Examples

```
sm.ps(runif(20))
sm.ps(runif(20), ps.int = 5)

## Not run:
data("TravelMode", package = "AER") # Need to install "AER" first
air.df <- subset(TravelMode, mode == "air") # Form 4 smaller data frames
bus.df <- subset(TravelMode, mode == "bus")
trn.df <- subset(TravelMode, mode == "train")
car.df <- subset(TravelMode, mode == "car")
TravelMode2 <- data.frame(income      = air.df$income,
                          wait.air   = air.df$wait - car.df$wait,
                          wait.trn   = trn.df$wait - car.df$wait,
                          wait.bus   = bus.df$wait - car.df$wait,
                          gcost.air  = air.df$gcost - car.df$gcost,
                          gcost.trn  = trn.df$gcost - car.df$gcost,
                          gcost.bus  = bus.df$gcost - car.df$gcost,
                          wait       = air.df$wait) # Value is unimportant
TravelMode2$mode <- subset(TravelMode, choice == "yes")$mode # The response
TravelMode2 <- transform(TravelMode2, incom.air = income, incom.trn = 0,
                          incom.bus = 0)

set.seed(1)
TravelMode2 <- transform(TravelMode2,
                          junkx2 = runif(nrow(TravelMode2)))

tfit2 <-
  vgam(mode ~ sm.ps(gcost.air, gcost.trn, gcost.bus) + ns(junkx2, 4) +
        sm.ps(incom.air, incom.trn, incom.bus) + wait ,
        crit = "coef",
        multinomial(parallel = FALSE ~ 1), data = TravelMode2,
        xij = list(sm.ps(gcost.air, gcost.trn, gcost.bus) ~
                    sm.ps(gcost.air, gcost.trn, gcost.bus) +
                    sm.ps(gcost.trn, gcost.bus, gcost.air) +
                    sm.ps(gcost.bus, gcost.air, gcost.trn),
                    sm.ps(incom.air, incom.trn, incom.bus) ~
                    sm.ps(incom.air, incom.trn, incom.bus) +
                    sm.ps(incom.trn, incom.bus, incom.air) +
                    sm.ps(incom.bus, incom.air, incom.trn),
                    wait ~ wait.air + wait.trn + wait.bus),
        form2 = ~ sm.ps(gcost.air, gcost.trn, gcost.bus) +
                  sm.ps(gcost.trn, gcost.bus, gcost.air) +
                  sm.ps(gcost.bus, gcost.air, gcost.trn) +
                  wait +
```

```

sm.ps(incom.air, incom.trn, incom.bus) +
sm.ps(incom.trn, incom.bus, incom.air) +
sm.ps(incom.bus, incom.air, incom.trn) +
junkx2 + ns(junkx2, 4) +
incom.air + incom.trn + incom.bus +
gcost.air + gcost.trn + gcost.bus +
wait.air + wait.trn + wait.bus)
par(mfrow = c(2, 2))
plot(tfit2, se = TRUE, lcol = "orange", scol = "blue", ylim = c(-4, 4))
summary(tfit2)

## End(Not run)

```

smart.expression	<i>S Expression for Smart Functions</i>
------------------	---

Description

smart.expression is an S expression for a smart function to call itself. It is best if you go through it line by line, but most users will not need to know anything about it. It requires the primary argument of the smart function to be called "x".

The list component match.call must be assigned the value of match.call() in the smart function; this is so that the smart function can call itself later.

See Also

[match.call](#).

Examples

```
print(sm.min2)
```

smart.mode.is	<i>Determine What Mode the Smart Prediction is In</i>
---------------	---

Description

Determine which of three modes the smart prediction is currently in.

Usage

```
smart.mode.is(mode.arg = NULL)
```

Arguments

mode.arg a character string, either "read", "write" or "neutral".

Details

Smart functions such as [bs](#) and [poly](#) need to know what mode smart prediction is in. If it is in "write" mode then the parameters are saved to `.smart.prediction` using [put.smart](#). If in "read" mode then the parameters are read in using [get.smart](#). If in "neutral" mode then the smart function behaves like an ordinary function.

Value

If `mode.arg` is given, then either TRUE or FALSE is returned. If `mode.arg` is not given, then the mode ("neutral", "read" or "write") is returned. Usually, the mode is "neutral".

See Also

[put.smart](#), [bs](#), [poly](#).

Examples

```
print(sm.min1)
smart.mode.is() # Returns "neutral"
smart.mode.is(smart.mode.is()) # Returns TRUE
```

 smartpred

Smart Prediction

Description

Data-dependent parameters in formula terms can cause problems in when predicting. The **smart-pred** package saves data-dependent parameters on the object so that the bug is fixed. The [lm](#) and [glm](#) functions have been fixed properly. Note that the **VGAM** package by T. W. Yee automatically comes with smart prediction.

Usage

```
sm.bs(x, df = NULL, knots = NULL, degree = 3, intercept = FALSE,
      Boundary.knots = range(x))
sm.ns(x, df = NULL, knots = NULL, intercept = FALSE,
      Boundary.knots = range(x))
sm.poly(x, ..., degree = 1, coefs = NULL, raw = FALSE)
sm.scale(x, center = TRUE, scale = TRUE)
```

Arguments

`x` The `x` argument is actually common to them all.
`df, knots, intercept, Boundary.knots`
 See [bs](#) and/or [ns](#).
`degree, ..., coefs, raw`
 See [poly](#).
`center, scale` See [scale](#).

Details

R version 1.6.0 introduced a partial fix for the prediction problem because it does not work all the time, e.g., for terms such as `I(poly(x, 3))`, `poly(c(scale(x)), 3)`, `bs(scale(x), 3)`, `scale(scale(x))`. See the examples below. Smart prediction, however, will always work.

The basic idea is that the functions in the formula are now smart, and the modelling functions make use of these smart functions. Smart prediction works in two ways: using `smart.expression`, or using a combination of `put.smart` and `get.smart`.

Value

The usual value returned by `bs`, `ns`, `poly` and `scale`. When used with functions such as `vglm` the data-dependent parameters are saved on one slot component called `smart.prediction`.

Side Effects

The variables `.max.smart`, `.smart.prediction` and `.smart.prediction.counter` are created while the model is being fitted. They are created in a new environment called `smartpredenv`. These variables are deleted after the model has been fitted. However, if there is an error in the model fitting function or the fitting model is killed (e.g., by typing control-C) then these variables will be left in `smartpredenv`. At the beginning of model fitting, these variables are deleted if present in `smartpredenv`.

During prediction, the variables `.smart.prediction` and `.smart.prediction.counter` are reconstructed and read by the smart functions when the model frame is re-evaluated. After prediction, these variables are deleted.

If the modelling function is used with argument `smart = FALSE` (e.g., `vglm(..., smart = FALSE)`) then smart prediction will not be used, and the results should match with the original R functions.

WARNING

The functions `bs`, `ns`, `poly` and `scale` are now left alone (from 2014-05 onwards) and no longer smart. They work via safe prediction. The smart versions of these functions have been renamed and they begin with "sm."

The functions `predict.bs` and `predict.ns` are not smart. That is because they operate on objects that contain attributes only and do not have list components or slots. The function `predict.poly` is not smart.

Author(s)

T. W. Yee and T. J. Hastie

See Also

`get.smart.prediction`, `get.smart`, `put.smart`, `smart.expression`, `smart.mode.is`, `setup.smart`, `wrapup.smart`. For `vgam` in **VGAM**, `sm.ps` is important. Commonly used data-dependent functions include `scale`, `poly`, `bs`, `ns`. In R, the functions `bs` and `ns` are in the **splines** package, and this library is automatically loaded in because it contains compiled code that `bs` and `ns` call.

The functions `vglm`, `vgam`, `rrvglm` and `cqo` in T. W. Yee's **VGAM** package are examples of modelling functions that employ smart prediction.

Examples

```

# Create some data first
n <- 20
set.seed(86) # For reproducibility of the random numbers
ldata <- data.frame(x2 = sort(runif(n)), y = sort(runif(n)))
library("splines") # To get ns() in R

# This will work for R 1.6.0 and later
fit <- lm(y ~ ns(x2, df = 5), data = ldata)
## Not run:
plot(y ~ x2, data = ldata)
lines(fitted(fit) ~ x2, data = ldata)
new.ldata <- data.frame(x2 = seq(0, 1, len = n))
points(predict(fit, new.ldata) ~ x2, new.ldata, type = "b", col = 2, err = -1)

## End(Not run)

# The following fails for R 1.6.x and later. It can be
# made to work with smart prediction provided
# ns is changed to sm.ns and scale is changed to sm.scale:
fit1 <- lm(y ~ ns(scale(x2), df = 5), data = ldata)
## Not run:
plot(y ~ x2, data = ldata, main = "Safe prediction fails")
lines(fitted(fit1) ~ x2, data = ldata)
points(predict(fit1, new.ldata) ~ x2, new.ldata, type = "b", col = 2, err = -1)

## End(Not run)

# Fit the above using smart prediction
## Not run:
library("VGAM") # The following requires the VGAM package to be loaded
fit2 <- vglm(y ~ sm.ns(sm.scale(x2), df = 5), uninormal, data = ldata)
fit2@smart.prediction
plot(y ~ x2, data = ldata, main = "Smart prediction")
lines(fitted(fit2) ~ x2, data = ldata)
points(predict(fit2, new.ldata, type = "response") ~ x2, data = new.ldata,
       type = "b", col = 2, err = -1)

## End(Not run)

```

specials

Special Values or Quantities in a Fitted Object

Description

Return any special values or quantities in a fitted object, and in particular in a VGLM fit

Usage

```

specials(object, ...)
specialsvglm(object, ...)

```

Arguments

object an object of class "vglm" whose family function begins with "gait".
 ... any additional arguments, to future-proof this function.

Details

This extractor function was motivated by GAITD regression (Yee and Ma, 2024) where the values from three disjoint sets are referred to as *special*. More generally, S4 methods functions can be written so that `specials()` will work on any S4 object, where what is called special depends on the methodology at hand.

Value

Returns any 'special' values or quantities associated with a fitted regression model. This is often something simple such as a list or a vector.

References

Yee, T. W. and Ma, C. (2024). Generally altered, inflated, truncated and deflated regression. *Statistical Science*, **39** (in press).

See Also

[vglm](#), [vglm-class](#), [inflated](#), [altered](#), [truncated](#), [Gaitdpois](#), [gaitdpoisson](#).

Examples

```
abdata <- data.frame(y = 0:7, w = c(182, 41, 12, 2, 2, 0, 0, 1))
fit1 <- vglm(y ~ 1, gaitdpoisson(a.mix = 0), data = abdata,
            weight = w, subset = w > 0)
specials(fit1)
```

 spikeplot

Spike Plot

Description

Produces a spike plot of a numeric vector.

Usage

```
spikeplot(x, freq = FALSE, as.table = FALSE, col = par("col"),
          lty = par("lty"), lwd = par("lwd"), lend = par("lend"),
          type = "h", xlab = deparse1(substitute(x)), ylab = NULL,
          capped = FALSE, cex = sqrt(lwd) / 2, pch = 19, pcol = col, scol = NULL,
          slty = NULL, slwd = NULL, new.plot = TRUE, offset.x = 0, ymux = 1, ...)
```

Arguments

<code>x</code>	Numeric, passed into <code>table</code> .
<code>freq</code>	Logical. If TRUE then the y-axis measures the frequencies, else the sample proportions. Intended to be as <code>hist</code> .
<code>as.table</code>	Logical. If TRUE then the call to <code>plot</code> is closer to <code>plot(table(x), ...)</code> , meaning the labelling differs from <code>as.table = FALSE</code> . The default is to convert <code>table(x)</code> into a numeric vector which is then passed into <code>plot</code> so that the labelling is more uniform along the x-axis.
<code>col, type, lty, lwd</code>	See <code>par</code> .
<code>lend, xlab, ylab</code>	See <code>par</code> .
<code>capped, cex, pch, pcol</code>	First argument is logical. If TRUE then the others argument are used to place points at the top using <code>points</code> with <code>pcol</code> being its colour. See <code>par</code> .
<code>scol, slty, slwd</code>	Similar to <code>col, lty</code> and <code>lwd</code> but apply to some selected values. The input may be a named list such as <code>scol = list("green" = c(2, 4, 6), "blue" = 5)</code> , <code>slty = list("dashed" = c(2, 4, 6), "dotted" = 5)</code> , <code>slwd = list("2" = c(2, 4, 6), "3" = 5)</code> , else a named vector such as <code>scol = c("green" = 2, "green" = 4, "green" = 6, "blue" = 5)</code> , <code>slty = c("dashed" = 2, "dashed" = 4, "dashed" = 6, "dotted" = 5)</code> , <code>slwd = c("2" = 2, "2" = 4, "2" = 6, "3" = 5)</code> . The three arguments are ignored if <code>as.table = TRUE</code> .
<code>new.plot, offset.x</code>	Logical and numeric. Add to an existing plot? If so, set <code>new.plot = FALSE</code> and it is useful for the spikes to be shifted by some amount <code>offset.x</code> .
<code>ymux</code>	Numeric, y-multiplier. The response is multiplied by <code>ymux</code> . This can be useful when plotting subsets side-by-side so that the constituent proportions add up to the overall proportion.
<code>...</code>	Additional graphical arguments passed into an ordinary <code>plot</code> , for example, <code>xlim, las, main</code> .

Details

Heaping is a very commonly occurring phenomenon in retrospective self-reported survey data. Also known as *digit preference* data, it is often characterized by an excess of multiples of 10 or 5 upon rounding. For this type of data this simple function is meant to be convenient for plotting the frequencies or sample proportions of a vector `x` representing a discrete random variable. This type of plot is known as a *spike plot* in STATA circles. If `table(x)` works then this function should hopefully work. The default for `type` means that any heaping and *seeping* should easily be seen. If such features exist then *GAITD regression* is potentially useful—see `gaitdpoisson` etc. Currently missing values are ignored totally because `table(x)` is used without further arguments; this might change in the future.

Value

Returns invisibly `table(x)`.

Author(s)

T. W. Yee.

See Also

[table](#), [plot](#), [par](#), [deparse1](#), [dgaitdplot](#), [plotdgaitd](#), [gaitdpoisson](#).

Examples

```
## Not run:
spikeplot(with(marital.nz, age), col = "pink2", lwd = 2)

## End(Not run)
```

sqrtlink

Square Root and Folded Square Root Link Functions

Description

Computes the square root and folded square root transformations, including their inverse and their first two derivatives.

Usage

```
foldsqrtlink(theta, min = 0, max = 1, mux = sqrt(2),
             inverse = FALSE, deriv = 0, short = TRUE, tag = FALSE)
sqrtlink(theta, inverse = FALSE, deriv = 0, short = TRUE,
         tag = FALSE, c10 = c(2, -2))
```

Arguments

theta Numeric or character. See below for further details.

min, max, mux These are called L , U and K below.

inverse, deriv, short, tag
 Details at [Links](#).

c10 Numeric, 2-vector $c(c1, c0)$ for a linear transformation. The plain link is multiplied by $c1$ and then $c0$ is added so that $c1 = 1:0$ is simply [sqrt](#). The default is intended to match [lcsloglink](#) for [poissonff](#) at λ (theta) equal to 1.

Details

The folded square root link function can be applied to parameters that lie between L and U inclusive. Numerical values of theta out of range result in NA or NaN.

More general information can be found at [alogitlink](#).

Value

For `foldsqrtlink` with `deriv = 0`: $K(\sqrt{\theta - L} - \sqrt{U - \theta})$ or $\text{mux} * (\text{sqrt}(\text{theta} - \text{min}) - \text{sqrt}(\text{max} - \text{theta}))$ when `inverse = FALSE`, and if `inverse = TRUE` then some more complicated function that returns a NA unless `theta` is between $-\text{mux} * \text{sqrt}(\text{max} - \text{min})$ and $\text{mux} * \text{sqrt}(\text{max} - \text{min})$.

For `sqrtlink` with `deriv = 0` and `c10 = 1:0`: $\sqrt{\theta}$ when `inverse = FALSE`, and if `inverse = TRUE` then the square is returned.

For `deriv = 1`, then the function returns $d \text{ eta} / d \text{ theta}$ as a function of `theta` if `inverse = FALSE`, else if `inverse = TRUE` then it returns the reciprocal.

Note

For `foldsqrtlink`, the default has, if `theta` is 0 or 1, the link function value is $-\text{sqrt}(2)$ and $+\text{sqrt}(2)$ respectively. These are finite values, therefore one cannot use this link function for general modelling of probabilities because of numerical problem, e.g., with `binomialff`, `cumulative`. See the example below.

Author(s)

Thomas W. Yee

See Also

[Links](#), [poissonff](#), [sloglink](#), [hdiff](#).

Examples

```
p <- seq(0.01, 0.99, by = 0.01)
foldsqrtlink(p)
max(abs(foldsqrtlink(foldsqrtlink(p), inverse = TRUE) - p)) # 0

p <- c(seq(-0.02, 0.02, by = 0.01), seq(0.97, 1.02, by = 0.01))
foldsqrtlink(p) # Has NAs

## Not run:
p <- seq(0.01, 0.99, by = 0.01)
par(mfrow = c(2, 2), lwd = (mylwd <- 2))
y <- seq(-4, 4, length = 100)
for (d in 0:1) {
  matplot(p, cbind( logitlink(p, deriv = d),
                   foldsqrtlink(p, deriv = d)),
          col = "blue", ylab = "transformation",
          main = ifelse(d == 0, "Some probability links",
                       "First derivative"), type = "n", las = 1)
  lines(p, logitlink(p, deriv = d), col = "green")
  lines(p, probitlink(p, deriv = d), col = "blue")
  lines(p, clogloglink(p, deriv = d), col = "red")
  lines(p, foldsqrtlink(p, deriv = d), col = "tan")
  if (d == 0) {
    abline(v = 0.5, h = 0, lty = "dashed")
    legend(0, 4.5, c("logitlink", "probitlink",
```

```

        "clogloglink", "foldsqrtlink"),
        lwd = 2, col = c("green", "blue",
                        "red", "tan"))
    } else
      abline(v = 0.5, lty = "dashed")
  }

for (d in 0) {
  matplot(y,
          cbind( logitlink(y, deriv = d, inverse = TRUE),
                 foldsqrtlink(y, deriv = d, inverse = TRUE)),
          type = "n", col = "blue", xlab = "transformation",
          ylab = "p", lwd = 2, las = 1, main = if (d == 0)
            "Some inverse probability link functions" else
            "First derivative")
  lines(y, logitlink(y, deriv=d, inverse=TRUE), col="green")
  lines(y, probitlink(y, deriv=d, inverse=TRUE), col="blue")
  lines(y, clogloglink(y, deriv=d, inverse=TRUE), col="red")
  lines(y, foldsqrtlink(y, deriv=d, inverse=TRUE), col="tan")
  if (d == 0) {
    abline(h = 0.5, v = 0, lty = "dashed")
    legend(-4, 1, c("logitlink", "probitlink",
                  "clogloglink", "foldsqrtlink"), lwd = 2,
          col = c("green", "blue", "red", "tan"))
  }
}
}
par(lwd = 1)

## End(Not run)

# This is lucky to converge
fit.h <- vglm(agaau ~ sm.bs(altitude),
             binomialff(foldsqrtlink(mux = 5)),
             hunua, trace = TRUE)

## Not run:
plotvgam(fit.h, se = TRUE, lcol = "orange", scol = "orange",
         main = "Orange is Hunua, Blue is Waitakere")
## End(Not run)
head(predict(fit.h, hunua, type = "response"))

## Not run:
# The following fails.
pneumo <- transform(pneumo, let = log(exposure.time))
fit <- vglm(cbind(normal, mild, severe) ~ let,
           cumulative(foldsqrtlink(mux = 10), par = TRUE, rev = TRUE),
           data = pneumo, trace = TRUE, maxit = 200)
## End(Not run)

```

Description

Fits a stopping ratio logit/probit/cloglog/cauchit/... regression model to an ordered (preferably) factor response.

Usage

```
sratio(link = "logitlink", parallel = FALSE, reverse = FALSE,
       zero = NULL, ynames = FALSE, Thresh = NULL, Trev = reverse,
       Tref = if (Trev) "M" else 1, whitespace = FALSE)
```

Arguments

link	Link function applied to the M stopping ratio probabilities. See Links for more choices.
parallel	A logical, or formula specifying which terms have equal/unequal coefficients.
reverse	Logical. By default, the stopping ratios used are $\eta_j = \text{logit}(P[Y = j Y \geq j])$ for $j = 1, \dots, M$. If reverse is TRUE, then $\eta_j = \text{logit}(P[Y = j+1 Y \leq j+1])$ will be used.
ynames	See multinomial for information.
zero	Can be an integer-valued vector specifying which linear/additive predictors are modelled as intercepts only. The values must be from the set $\{1, 2, \dots, M\}$. The default value means none are modelled as intercept-only terms. See CommonVGAMffArguments for information.
Thresh, Trev, Tref	See cumulative for information. These arguments apply to ordinal categorical regression models.
whitespace	See CommonVGAMffArguments for information.

Details

In this help file the response Y is assumed to be a factor with ordered values $1, 2, \dots, M + 1$, so that M is the number of linear/additive predictors η_j .

There are a number of definitions for the *continuation ratio* in the literature. To make life easier, in the **VGAM** package, we use *continuation* ratios (see [cratio](#)) and *stopping* ratios. Continuation ratios deal with quantities such as $\text{logitlink}(P[Y > j | Y \geq j])$.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), [rrvglm](#) and [vgam](#).

Warning

No check is made to verify that the response is ordinal if the response is a matrix; see [ordered](#).

Boersch-Supan (2021) considers a sparse data set (called [budworm](#)) and the numerical problems encountered when fitting models such as [cratio](#), [sratio](#), [cumulative](#). Although improvements to links such as [clogloglink](#) have been made, currently these family functions have not been properly adapted to handle sparse data as well as they could.

Note

The response should be either a matrix of counts (with row sums that are all positive), or a factor. In both cases, the y slot returned by `vglm/vgam/rrvglm` is the matrix of counts.

For a nominal (unordered) factor response, the multinomial logit model (`multinomial`) is more appropriate.

Here is an example of the usage of the `parallel` argument. If there are covariates x_1 , x_2 and x_3 , then `parallel = TRUE ~ x1 + x2 - 1` and `parallel = FALSE ~ x3` are equivalent. This would constrain the regression coefficients for x_1 and x_2 to be equal; those of the intercepts and x_3 would be different.

Author(s)

Thomas W. Yee

References

- Agresti, A. (2013). *Categorical Data Analysis*, 3rd ed. Hoboken, NJ, USA: Wiley.
- Boersch-Supan, P. H. (2021). Modeling insect phenology using ordinal regression and continuation ratio models. *ReScience C*, **7.1**, 1–14. doi:10.18637/jss.v032.i10.
- McCullagh, P. and Nelder, J. A. (1989). *Generalized Linear Models*, 2nd ed. London: Chapman & Hall.
- Tutz, G. (2012). *Regression for Categorical Data*, Cambridge: Cambridge University Press.
- Yee, T. W. (2010). The **VGAM** package for categorical data analysis. *Journal of Statistical Software*, **32**, 1–34. doi:10.18637/jss.v032.i10.

See Also

[cratio](#), [acat](#), [cumulative](#), [multinomial](#), [CM.equid](#), [CommonVGAMffArguments](#), [margeff](#), [pneumo](#), [budworm](#), [logitlink](#), [probitlink](#), [clogloglink](#), [cauchitlink](#).

Examples

```
pneumo <- transform(pneumo, let = log(exposure.time))
(fit <- vglm(cbind(normal, mild, severe) ~ let,
            sratio(parallel = TRUE), data = pneumo))
coef(fit, matrix = TRUE)
constraints(fit)
predict(fit)
predict(fit, untransform = TRUE)
```

step4 *Choose a model by AIC in a Stepwise Algorithm*

Description

Select a formula-based model by AIC.

Usage

```
step4(object, ...)
step4vglm(object, scope, direction = c("both", "backward", "forward"),
          trace = 1, keep = NULL, steps = 1000, k = 2, ...)
```

Arguments

object	an object of class "vglm". This is used as the initial model in the stepwise search.
scope	See step .
direction	See step .
trace, keep	See step .
steps, k	See step .
...	any additional arguments to extractAIC.vglm , drop1.vglm and add1.vglm .

Details

This function is a direct adaptation of [step](#) for [vglm-class](#) objects. Since [step](#) is not generic, the name `step4()` was adopted and it *is* generic, as well as being S4 rather than S3. It is the intent that this function should work as similar as possible to [step](#).

Internally, the methods function for [vglm-class](#) objects calls [add1.vglm](#) and [drop1.vglm](#) repeatedly.

Value

The results are placed in the post slot of the stepwise-selected model that is returned. There are up to two additional components. There is an "anova" component corresponding to the steps taken in the search, as well as a "keep" component if the `keep=` argument was supplied in the call.

Warning

In general, the same warnings in [drop1.glm](#) and [drop1.vglm](#) apply here.

This function

See Also

[add1.vglm](#), [drop1.vglm](#), [vglm](#), [trim.constraints](#), [add1.glm](#), [drop1.glm](#), [backPain2](#), [step](#), [update](#).

Examples

```
data("backPain2", package = "VGAM")
summary(backPain2)
fit1 <- vglm(pain ~ x2 + x3 + x4 + x2:x3 + x2:x4 + x3:x4,
             propodds, data = backPain2)
spom1 <- step4(fit1)
summary(spom1)
spom1$post$anova
```

studentt

Student t Distribution

Description

Estimating the parameters of a Student t distribution.

Usage

```
studentt (ldf = "logloglink", idf = NULL, tol1 = 0.1, imethod = 1)
studentt2(df = Inf, llocation = "identitylink", lscale = "loglink",
          ilocation = NULL, iscale = NULL, imethod = 1, zero = "scale")
studentt3(llocation = "identitylink", lscale = "loglink",
          ldf = "logloglink", ilocation = NULL, iscale = NULL,
          idf = NULL, imethod = 1, zero = c("scale", "df"))
```

Arguments

`llocation`, `lscale`, `ldf`
 Parameter link functions for each parameter, e.g., for degrees of freedom ν . See [Links](#) for more choices. The defaults ensures the parameters are in range. A [loglog](#) link keeps the degrees of freedom greater than unity; see below.

`ilocation`, `iscale`, `idf`
 Optional initial values. If given, the values must be in range. The default is to compute an initial value internally.

`tol1`
 A positive value, the tolerance for testing whether an initial value is 1. Best to leave this argument alone.

`df`
 Numeric, user-specified degrees of freedom. It may be of length equal to the number of columns of a response matrix.

`imethod`, `zero` See [CommonVGAMffArguments](#).

Details

The Student t density function is

$$f(y; \nu) = \frac{\Gamma((\nu + 1)/2)}{\sqrt{\nu\pi}\Gamma(\nu/2)} \left(1 + \frac{y^2}{\nu}\right)^{-(\nu+1)/2}$$

for all real y . Then $E(Y) = 0$ if $\nu > 1$ (returned as the fitted values), and $Var(Y) = \nu/(\nu - 2)$ for $\nu > 2$. When $\nu = 1$ then the Student t -distribution corresponds to the standard Cauchy distribution, [cauchy1](#). When $\nu = 2$ with a scale parameter of `sqrt(2)` then the Student t -distribution corresponds to the standard (Koenker) distribution, [sc.studentt2](#). The degrees of freedom can be treated as a parameter to be estimated, and as a real and not an integer. The Student t distribution is used for a variety of reasons in statistics, including robust regression.

Let $Y = (T - \mu)/\sigma$ where μ and σ are the location and scale parameters respectively. Then `studentt3` estimates the location, scale and degrees of freedom parameters. And `studentt2` estimates the location, scale parameters for a user-specified degrees of freedom, `df`. And `studentt` estimates the degrees of freedom parameter only. The fitted values are the location parameters. By default the linear/additive predictors are $(\mu, \log(\sigma), \log \log(\nu))^T$ or subsets thereof.

In general convergence can be slow, especially when there are covariates.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Note

`studentt3()` and `studentt2()` can handle multiple responses.

Practical experience has shown reasonably good initial values are required. If convergence failure occurs try using arguments such as `idf`. Local solutions are also possible, especially when the degrees of freedom is close to unity or the scale parameter is close to zero.

A standard normal distribution corresponds to a t distribution with infinite degrees of freedom. Consequently, if the data is close to normal, there may be convergence problems; best to use [uninormal](#) instead.

Author(s)

T. W. Yee

References

Student (1908). The probable error of a mean. *Biometrika*, **6**, 1–25.

Zhu, D. and Galbraith, J. W. (2010). A generalized asymmetric Student- t distribution with application to financial econometrics. *Journal of Econometrics*, **157**, 297–305.

See Also

[uninormal](#), [cauchy1](#), [logistic](#), [huber2](#), [sc.studentt2](#), [TDist](#), [simulate.vlm](#).

Examples

```
tdata <- data.frame(x2 = runif(nn <- 1000))
tdata <- transform(tdata, y1 = rt(nn, df = exp(exp(0.5 - x2))),
                  y2 = rt(nn, df = exp(exp(0.5 - x2))))
fit1 <- vglm(y1 ~ x2, studentt, data = tdata, trace = TRUE)
coef(fit1, matrix = TRUE)
```

```
# df inputted into studentt2() not quite right:
fit2 <- vglm(y1 ~ x2, studentt2(df = exp(exp(0.5))), tdata)
coef(fit2, matrix = TRUE)

fit3 <- vglm(cbind(y1, y2) ~ x2, studentt3, tdata, trace = TRUE)
coef(fit3, matrix = TRUE)
```

summary.drrvglm	<i>Summarizing Reduced Rank Vector Generalized Linear Model (RR-VGLM) and Doubly constrained RR-VGLM Fits</i>
-----------------	---

Description

These functions are all [methods](#) for class "drrvglm" or "summary.drrvglm" objects, or for class "rrvglm" or "summary.rrvglm" objects.

Usage

```
## S3 method for class 'drrvglm'
summary(object, correlation = FALSE, dispersion = NULL,
  digits = NULL, numerical = TRUE, h.step = 0.005, omit123 = FALSE,
  omit13 = FALSE, fixA = FALSE, presid = FALSE,
  signif.stars = getOption("show.signif.stars"),
  nopredictors = FALSE, eval0 = TRUE, ...)
## S3 method for class 'summary.drrvglm'
show(x, digits = NULL,
  quote = TRUE, prefix = "", signif.stars = NULL)
## S3 method for class 'rrvglm'
summary(object, correlation = FALSE, dispersion = NULL,
  digits = NULL, numerical = TRUE, h.step = 0.005, omit123 = FALSE,
  omit13 = FALSE, fixA = TRUE, presid = FALSE,
  signif.stars = getOption("show.signif.stars"),
  nopredictors = FALSE, upgrade = FALSE, ...)
## S3 method for class 'summary.rrvglm'
show(x, digits = NULL,
  quote = TRUE, prefix = "", signif.stars = NULL)
```

Arguments

object	an object of class "drrvglm" or "rrvglm", a result of a call to rrvglm .
x	an object of class "summary.drrvglm" or "summary.rrvglm", a result of a call to summary.drrvglm or summary.rrvglm .
dispersion	used mainly for GLMs. Not really implemented in VGAM so should not be used.
correlation	See summaryvglm .
digits	See summaryvglm .

signif.stars	See summaryvglm .
presid,quote	See summaryvglm .
nopredictors	See summaryvglm .
upgrade	Logical. Upgrade object to <code>drrvglm-class</code> ? Treating the object as a DRR-VGLM has advantages since the framework is larger. The code for ordinary RR-VGLMs was written a long time ago so it is a good idea to check that both give the same answer.
numerical	Logical, use a finite difference approximation for partial derivatives? If FALSE then theoretical formulas are used (however this option may no longer be implemented).
h.step	Numeric, positive and close to 0. If <code>numerical</code> then this is the forward step for each finite difference approximation. That is, it plays the role of h in $(f(x+h) - f(x))/h$ for some function f . If the overall variance-covariance matrix is not positive-definite, varying this argument can make a difference, e.g., increasing it to 0.01 is recommended.
fixA	Logical, if TRUE then the largest block matrix is for B1 and C , else it is for A and B1 . This should not make any difference because both estimates of B1 should be extremely similar, including the SEs.
omit13	Logical, if TRUE then the (1,3) block matrix is set to O . That is, A and C are assumed to asymptotically uncorrelated. Setting this TRUE is an option when V (see below) is not positive-definite. If this fails, another option that is often better is to set <code>omit123 = TRUE</code> .
omit123	Logical. If TRUE then <i>two</i> block matrices are set to O (blocks (1,2) and (1,3), else blocks (1,3) and (2,3), depending on <code>fixA</code>), This will almost surely result in an overall variance-covariance matrix that is positive-definite, however, the SEs will be biased. This argument is more extreme than <code>omit13</code> .
prefix	See summaryvglm .
eval0	Logical. Check if V is positive-definite? That is, all its eigenvalues are positive.
...	Logical argument check. 2 might work here. If TRUE then some quantities are printed out, for checking and debugging.

Details

Most of this document concerns DRR-VGLMs but also apply equally well to RR-VGLMs as a special case.

The overall variance-covariance matrix The overall variance-covariance matrix (called **V** below) is computed. Since the parameters comprise the elements of the matrices **A**, **B1** and **C** (called here block matrices 1, 2, 3 respectively), and an alternating algorithm is used for estimation, then there are two overlapping submodels that are fitted within an IRLS algorithm. These have blocks 1 and 2, and 2 and 3, so that **B1** is common to both. They are combined into one large overall variance-covariance matrix. Argument `fixA` specifies which submodel the **B1** block is taken from. Block (1,3) is the most difficult to compute and numerical approximations based on first derivatives are used by default for this.

Sometimes the computed **V** is not positive-definite. If so, then the standard errors will be NA. To avoid this problem, try varying `h.step` or refitting the model with a different `Index.corner`.

Argument `omit13` and `omit123` can also be used to give approximate answers. If \mathbf{V} is not positive-definite then this may indicate that the model does not fit the data very well, e.g., Rank is not a good value. Potentially, there are many ways why the model may be ill-conditioned. Try several options and set `trace = TRUE` to monitor convergence—this is informative about how well the model and data agree.

How can one fit an ordinary RR-VGLM as a DRR-VGLM? If one uses corner constraints (default) then one should input `H.A` as a list containing Rank `diag(M)` matrices—one for each column of `A`. Then since `Corner = TRUE` by default, then `object@H.A.alt` has certain columns deleted due to corner constraints. In contrast, `object@H.A.thy` is the `H.A` that was inputted. FYI, the `alt` suffix indicates the alternating algorithm, while the suffix `thy` stands for *theory*.

Value

`summarydrrvglm` returns an object of class "summary.drrvglm".

Warning

DRR-VGLMs are a recent development so it will take some time to get things totally ironed out. RR-VGLMs were developed a long time ago and are more well-established, however they have only recently been documented here.

Note

Note that `vcov` methods exist for `rrvglm-class` and `drrvglm-class` objects.

Sometimes this function can take a long time and this is because numerical derivatives are computed.

Author(s)

T. W. Yee.

References

Chapter 5 of: Yee, T. W. (2015). *Vector Generalized Linear and Additive Models: With an Implementation in R*. New York, USA: *Springer*. Sections 5.2.2 and 5.3 are particularly relevant.

See Also

[rrvglm](#), [rrvglm.control](#), [vcovdrrvglm](#), [CM.free](#), [summaryvglm](#), [summary.rrvglm-class](#), [summary.drrvglm-class](#).

Examples

```
## Not run: # Fit a rank-1 RR-VGLM as a DRR-VGLM.
set.seed(1); n <- 1000; S <- 6 # S must be even
myrank <- 1
rdata <- data.frame(x1 = runif(n), x2 = runif(n),
                   x3 = runif(n), x4 = runif(n))
dval <- ncol(rdata) # Number of covariates
# Involves x1, x2, ... a rank-1 model:
ymatrix <- with(rdata,
```

```

matrix(rpois(n*S, exp(3 + x1 - 0.5*x2)), n, S))
H.C <- vector("list", dval) # Ordinary "rrvglm"
for (i in 1:dval) H.C[[i]] <- CM.free(myrank)
names(H.C) <- paste0("x", 1:dval)
H.A <- list(CM.free(S)) # rank-1

rfit1 <- rrvglm(ymatrix ~ x1 + x2 + x3 + x4,
               poissonff, rdata, trace = TRUE)
class(rfit1)
dfit1 <- rrvglm(ymatrix ~ x1 + x2 + x3 + x4,
               poissonff, rdata, trace = TRUE,
               H.A = H.A, # drrvglm
               H.C = H.C) # drrvglm
class(dfit1)
Coef(rfit1) # The RR-VGLM is the same as
Coef(dfit1) # the DRR-VGLM.
max(abs(predict(rfit1) - predict(dfit1))) # 0
abs(logLik(rfit1) - logLik(dfit1)) # 0
summary(rfit1)
summary(dfit1)

## End(Not run)

```

summarypvgam

Summarizing Penalized Vector Generalized Additive Model Fits

Description

These functions are all [methods](#) for class "pvgam" or `summary.pvgam` objects.

Usage

```

summarypvgam(object, dispersion = NULL, digits = options()$digits - 2,
              presid = TRUE)
## S3 method for class 'summary.pvgam'
show(x, quote = TRUE, prefix = "", digits = options()$digits -
      2, signif.stars = getOption("show.signif.stars"))

```

Arguments

`object` an object of class "pvgam", which is the result of a call to [vgam](#) with at least one [sm.os](#) or [sm.ps](#) term.

`x` an object of class "summary.pvgam", which is the result of a call to `summarypvgam()`.

`dispersion`, `digits`, `presid`
See [summaryvglm](#).

`quote`, `prefix`, `signif.stars`
See [summaryvglm](#).

Details

This methods function reports a summary more similar to `summary.gam` from **mgcv** than `summary.gam()` from **gam**. It applies to G2-VGAMs using `sm.os` and O-splines, else `sm.ps` and P-splines. In particular, the hypothesis test for whether each `sm.os` or `sm.ps` term can be deleted follows quite closely to `summary.gam`. The p-values from this type of test tend to be biased downwards (too small) and corresponds to `p.type = 5`. It is hoped in the short future that improved p-values be implemented, somewhat like the default of `summary.gam`. This methods function was adapted from `summary.gam`.

Value

`summarypvgam` returns an object of class "summary.pvgam"; see `summary.pvgam-class`.

Warning

See `sm.os`.

See Also

`vgam`, `summaryvgam`, `summary.pvgam-class`, `sm.os`, `sm.ps`, `summary.glm`, `summary.lm`, `summary.gam` from **mgcv**, `summaryvgam` for G1-VGAMs.

Examples

```
## Not run:
hfit2 <- vgam(agaaus ~ sm.os(altitude), binomialff, data = hunua)
coef(hfit2, matrix = TRUE)
summary(hfit2)

## End(Not run)
```

summaryvgam

Summarizing Vector Generalized Additive Model Fits

Description

These functions are all `methods` for class `vgam` or `summary.vgam` objects.

Usage

```
summaryvgam(object, dispersion = NULL, digits = options()$digits - 2,
             presid = TRUE, nopredictors = FALSE)
## S3 method for class 'summary.vgam'
show(x, quote = TRUE, prefix = "",
     digits = options()$digits-2, nopredictors = NULL)
```

Arguments

object an object of class "vglm", which is the result of a call to `vglm` with at least one `s` term.

x an object of class "summary.vglm", which is the result of a call to `summaryvglm()`.

dispersion, digits, presid
 See `summaryvglm`.

quote, prefix, nopredictors
 See `summaryvglm`.

Details

This methods function reports a summary more similar to `summary.gam()` from **gam** than `summary.gam` from **mgcv**. It applies to G1-VGAMs using `s` and vector backfitting. In particular, an approximate score test for *linearity* is conducted for each `s` term—see Section 4.3.4 of Yee (2015) for details. The p-values from this type of test tend to be biased upwards (too large).

Value

`summaryvglm` returns an object of class "summary.vglm"; see `summary.vglm-class`.

See Also

`vglm`, `summary.glm`, `summary.lm`, `summary.gam` from **mgcv**, `summarypvglm` for P-VGAMs.

Examples

```
hfit <- vglm(agaaus ~ s(altitude, df = 2), binomialff, data = hunua)
summary(hfit)
summary(hfit)@anova # Table for (approximate) testing of linearity
```

summaryvglm

Summarizing Vector Generalized Linear Model Fits

Description

These functions are all `methods` for class `vglm` or `summary.vglm` objects.

Usage

```
summaryvglm(object, correlation = FALSE, dispersion = NULL,
            digits = NULL, presid = FALSE,
            HDEtest = TRUE, hde.NA = TRUE, threshold.hde = 0.001,
            signif.stars = getOption("show.signif.stars"),
            nopredictors = FALSE,
            lrt0.arg = FALSE, score0.arg = FALSE, wald0.arg = FALSE,
            values0 = 0, subset = NULL, omit1s = TRUE,
            ...)
```

```
## S3 method for class 'summary.vglm'
show(x, digits = max(3L, getOption("digits") - 3L),
      quote = TRUE, prefix = "", presid = length(x@pearson.resid) > 0,
      HDEtest = TRUE, hde.NA = TRUE, threshold.hde = 0.001,
      signif.stars = NULL, nopredictors = NULL,
      top.half.only = FALSE, ...)
```

Arguments

object	an object of class "vglm", usually, a result of a call to vglm .
x	an object of class "summary.vglm", usually, a result of a call to summaryvglm() .
dispersion	used mainly for GLMs. See summary.glm .
correlation	logical; if TRUE, the correlation matrix of the estimated parameters is returned and printed.
digits	the number of significant digits to use when printing.
signif.stars	logical; if TRUE, 'significance stars' are printed for each coefficient.
presid	Pearson residuals; print out some summary statistics of these?
HDEtest	logical; if TRUE (the default) then a test for the HDE is performed, else all arguments related to the HDE are ignored.
hde.NA	logical; if a test for the Hauck-Donner effect is done (for each coefficient) and it is affirmative should that Wald test p-value be replaced by an NA? The default is to do so. Setting <code>hde.NA = FALSE</code> will print the p-value even though it will be biased upwards. Also see argument <code>threshold.hde</code> .
threshold.hde	numeric; used if <code>hde.NA = TRUE</code> and is present for some coefficients. Only p-values greater than this argument will be replaced by an NA, the reason being that small p-values will already be statistically significant. Hence setting <code>threshold.hde = 0</code> will print out a NA if the HDE is present.
quote	Fed into <code>print()</code> .
nopredictors	logical; if TRUE the names of the linear predictors are not printed out. The default is that they are.
<code>lrt0.arg</code> , <code>score0.arg</code> , <code>wald0.arg</code>	Logical. If <code>lrt0.arg = TRUE</code> then the other arguments are passed into lrt.stat.vlm and the equivalent of the so-called Wald table is outputted. Similarly, if <code>score0.arg = TRUE</code> then the other arguments are passed into score.stat.vlm and the equivalent of the so-called Wald table is outputted. Similarly, if <code>wald0.arg = TRUE</code> then the other arguments are passed into wald.stat.vlm and the Wald table corresponding to that is outputted. See details below. Setting any of these will result in further IRLS iterations being performed, therefore may be computationally expensive.
<code>values0</code> , <code>subset</code> , <code>omit1s</code>	These arguments are used if any of the <code>lrt0.arg</code> , <code>score0.arg</code> , <code>wald0.arg</code> arguments are used. They are passed into the appropriate function, such as wald.stat.vlm .
top.half.only	logical; if TRUE then only print out the top half of the usual output. Used for P-VGAMs.

prefix	Not used.
...	Not used.

Details

Originally, `summaryvglm()` was written to be very similar to `summary.glm`, however now there are a quite a few more options available. By default, `show.summary.vglm()` tries to be smart about formatting the coefficients, standard errors, etc. and additionally gives ‘significance stars’ if `signif.stars` is TRUE. The coefficients component of the result gives the estimated coefficients and their estimated standard errors, together with their ratio. This third column is labelled z value regardless of whether the dispersion is estimated or known (or fixed by the family). A fourth column gives the two-tailed p-value corresponding to the z ratio based on a Normal reference distribution. In general, the t distribution is not used, but the normal distribution is.

Correlations are printed to two decimal places (or symbolically): to see the actual correlations print `summary(object)@correlation` directly.

The Hauck-Donner effect (HDE) is tested for almost all models; see `hdeff.vglm` for details. Arguments `hde.NA` and `threshold.hde` here are meant to give some control of the output if this aberration of the Wald statistic occurs (so that the p-value is biased upwards). If the HDE is present then using `lrt.stat.vlm` to get a more accurate p-value is a good alternative as p-values based on the likelihood ratio test (LRT) tend to be more accurate than Wald tests and do not suffer from the HDE. Alternatively, if the HDE is present then using `wald0.arg = TRUE` will compute Wald statistics that are HDE-free; see `wald.stat`.

The arguments `lrt0.arg` and `score0.arg` enable the so-called Wald table to be replaced by the equivalent LRT and Rao score test table; see `lrt.stat.vlm`, `score.stat`. Further IRLS iterations are performed for both of these, hence the computational cost might be significant.

It is possible for programmers to write a methods function to print out extra quantities when `summary(vglmObject)` is called. The generic function is `summaryvglmS4VGAM()`, and one can use the S4 function `setMethod` to compute the quantities needed. Also needed is the generic function `showsummaryvglmS4VGAM()` to actually print the quantities out.

Value

`summaryvglm` returns an object of class “`summary.vglm`”; see `summary.vglm-class`.

Warning

Currently the SE column is deleted when `lrt0 = TRUE` because SEs are not so meaningful with the LRT. In the future an SE column may be inserted (with NA values) so that it has 4-column output like the other tests. In the meantime, the columns of this matrix should be accessed by name and not number.

Author(s)

T. W. Yee.

See Also

`vglm`, `confintvglm`, `vcovvglm`, `summary.rrvglm`, `summary.glm`, `summary.lm`, `summary`, `hdeff.vglm`, `lrt.stat.vlm`, `score.stat`, `wald.stat`.

Examples

```
## For examples see example(glm)
pneumo <- transform(pneumo, let = log(exposure.time))
(afit <- vglm(cbind(normal, mild, severe) ~ let, acat, data = pneumo))
coef(afit, matrix = TRUE)
summary(afit) # Might suffer from the Hauck-Donner effect
coef(summary(afit))
summary(afit, lrt0 = TRUE, score0 = TRUE, wald0 = TRUE)
```

 SURff

Seemingly Unrelated Regressions Family Function

Description

Fits a system of seemingly unrelated regressions.

Usage

```
SURff(mle.normal = FALSE,
      divisor = c("n", "n-max(pj,pk)", "sqrt((n-pj)*(n-pk)")),
      parallel = FALSE, Varcov = NULL, matrix.arg = FALSE)
```

Arguments

<code>mle.normal</code>	Logical. If TRUE then the MLE, assuming multivariate normal errors, is computed; the effect is just to add a loglikelihood slot to the returned object. Then it results in the <i>maximum likelihood estimator</i> .
<code>divisor</code>	Character, partial matching allowed and the first choice is the default. The divisor for the estimate of the covariances. If "n" then the estimate will be biased. If the others then the estimate will be unbiased for some elements. If <code>mle.normal = TRUE</code> and this argument is not "n" then a warning or an error will result.
<code>parallel</code>	See CommonVGAMffArguments . If <code>parallel = TRUE</code> then the constraint applies to the intercept too.
<code>Varcov</code>	Numeric. This may be assigned a variance-covariance of the errors. If <code>matrix.arg</code> then this is a $M \times M$ matrix. If <code>!matrix.arg</code> then this is a $M \times M$ matrix in matrix-band format (a vector with at least M and at most $M*(M+1)/2$ elements).
<code>matrix.arg</code>	Logical. Of single length.

Details

Proposed by Zellner (1962), the basic seemingly unrelated regressions (SUR) model is a set of LMs ($M > 1$ of them) tied together at the error term level. Each LM's model matrix may potentially have its own set of predictor variables.

Zellner's efficient (ZEF) estimator (also known as *Zellner's two-stage Aitken estimator*) can be obtained by setting `maxit = 1` (and possibly `divisor = "sqrt"` or `divisor = "n-max"`).

The default value of `maxit` (in `vglm.control`) probably means *iterative GLS* (IGLS) estimator is computed because IRLS will probably iterate to convergence. IGLS means, at each iteration, the residuals are used to estimate the error variance-covariance matrix, and then the matrix is used in the GLS. The IGLS estimator is also known as *Zellner's iterative Aitken estimator*, or IZEF.

Value

An object of class "vglmff" (see `vglmff-class`). The object is used by modelling functions such as `vglm` and `vgam`.

Warning

The default convergence criterion may be a little loose. Try setting `epsilon = 1e-11`, especially with `mle.normal = TRUE`.

Note

The fitted object has slot `@extra$ncols.X.lm` which is a M vector with the number of parameters for each LM. Also, `@misc$values.divisor` is the M -vector of divisor values.

Constraint matrices are needed in order to specify which response variables that each term on the RHS of the formula is a regressor for. See the `constraints` argument of `vglm` for more information.

Author(s)

T. W. Yee.

References

Zellner, A. (1962). An Efficient Method of Estimating Seemingly Unrelated Regressions and Tests for Aggregation Bias. *J. Amer. Statist. Assoc.*, **57**(298), 348–368.

Kmenta, J. and Gilbert, R. F. (1968). Small Sample Properties of Alternative Estimators of Seemingly Unrelated Regressions. *J. Amer. Statist. Assoc.*, **63**(324), 1180–1200.

See Also

`uninormal`, `gew`.

Examples

```
# Obtain some of the results of p.1199 of Kmenta and Gilbert (1968)
clist <- list("(Intercept)" = diag(2),
             "capital.g"   = rbind(1, 0),
             "value.g"     = rbind(1, 0),
             "capital.w"   = rbind(0, 1),
             "value.w"     = rbind(0, 1))
zef1 <- vglm(cbind(invest.g, invest.w) ~
            capital.g + value.g + capital.w + value.w,
            SURff(divisor = "sqrt"), maxit = 1,
            data = gew, trace = TRUE, constraints = clist)
```

```

round(coef(zef1, matrix = TRUE), digits = 4) # ZEF
zef1@extra$ncols.X.lm
zef1@misc$divisor
zef1@misc$values.divisor
round(sqrt(diag(vcov(zef1))), digits = 4) # SEs
nobs(zef1, type = "lm")
df.residual(zef1, type = "lm")

mle1 <- vglm(cbind(invest.g, invest.w) ~
  capital.g + value.g + capital.w + value.w,
  SURff(mle.normal = TRUE),
  epsilon = 1e-11,
  data = gew, trace = TRUE, constraints = clist)
round(coef(mle1, matrix = TRUE), digits = 4) # MLE
round(sqrt(diag(vcov(mle1))), digits = 4) # SEs

```

SurvS4

*Create a Survival Object***Description**

Create a survival object, usually used as a response variable in a model formula.

Usage

```
SurvS4(time, time2, event, type =, origin = 0)
is.SurvS4(x)
```

Arguments

time	for right censored data, this is the follow up time. For interval data, the first argument is the starting time for the interval.
x	any R object.
event	The status indicator, normally 0=alive, 1=dead. Other choices are TRUE/FALSE (TRUE = death) or 1/2 (2=death). For interval censored data, the status indicator is 0=right censored, 1=event at time, 2=left censored, 3=interval censored. Although unusual, the event indicator can be omitted, in which case all subjects are assumed to have an event.
time2	ending time of the interval for interval censored or counting process data only. Intervals are assumed to be open on the left and closed on the right, (start, end]. For counting process data, event indicates whether an event occurred at the end of the interval.
type	character string specifying the type of censoring. Possible values are "right", "left", "counting", "interval", or "interval2". The default is "right" or "counting" depending on whether the time2 argument is absent or present, respectively.

`origin` for counting process data, the hazard function origin. This is most often used in conjunction with a model containing time dependent strata in order to align the subjects properly when they cross over from one strata to another.

Details

Typical usages are

```
SurvS4(time, event)
SurvS4(time, time2, event, type=, origin=0)
```

In theory it is possible to represent interval censored data without a third column containing the explicit status. Exact, right censored, left censored and interval censored observation would be represented as intervals of (a,a), (a, infinity), (-infinity,b), and (a,b) respectively; each specifying the interval within which the event is known to have occurred.

If `type = "interval2"` then the representation given above is assumed, with NA taking the place of infinity. If `type="interval"` event must be given. If event is 0, 1, or 2, the relevant information is assumed to be contained in time, the value in time2 is ignored, and the second column of the result will contain a placeholder.

Presently, the only methods allowing interval censored data are the parametric models computed by `survreg`, so the distinction between open and closed intervals is unimportant. The distinction is important for counting process data and the Cox model.

The function tries to distinguish between the use of 0/1 and 1/2 coding for left and right censored data using `if (max(status)==2)`. If 1/2 coding is used and all the subjects are censored, it will guess wrong. Use 0/1 coding in this case.

Value

An object of class `SurvS4` (formerly `Surv`). There are methods for `print`, `is.na`, and subscripting survival objects. `SurvS4` objects are implemented as a matrix of 2 or 3 columns.

In the case of `is.SurvS4`, a logical value TRUE if x inherits from class "SurvS4", otherwise a FALSE.

Note

The purpose of having `SurvS4` in **VGAM** is so that the same input can be fed into `vglm` as functions in **survival** such as `survreg`. The class name has been changed from "Surv" to "SurvS4"; see [SurvS4-class](#).

The format `J+` is interpreted in **VGAM** as $\geq J$. If `type="interval"` then these should not be used in **VGAM**: `(L,U-]` or `(L,U+]`.

Author(s)

The code and documentation comes from **survival**. Slight modifications have been made for conversion to S4 by T. W. Yee. Also, for "interval" data, `as.character.SurvS4()` has been modified to print intervals of the form `(start, end]` and not `[start, end]` as previously. (This makes a difference for discrete data, such as for `cens.poisson`). All **VGAM** family functions beginning with "cen" require the packaging function `Surv` to format the input.

See Also

[SurvS4-class](#), [cens.poisson](#), [survreg](#), [leukemia](#).

Examples

```
with(leukemia, SurvS4(time, status))
class(with(leukemia, SurvS4(time, status)))
```

SurvS4-class	<i>Class "SurvS4"</i>
--------------	-----------------------

Description

S4 version of the Surv class.

Objects from the Class

A virtual Class: No objects may be created from it.

Extends

Class "Surv", directly. Class "[matrix](#)", directly. Class "[oldClass](#)", by class "Surv", distance 2. Class "[structure](#)", by class "matrix", distance 2. Class "[array](#)", by class "matrix", distance 2. Class "[vector](#)", by class "matrix", distance 3, with explicit coerce. Class "[vector](#)", by class "matrix", distance 4, with explicit coerce.

Methods

`show` signature(object = "SurvS4"): ...

Warning

This code has not been thoroughly tested.

Note

The purpose of having [SurvS4](#) in **VGAM** is so that the same input can be fed into [vglm](#) as functions in **survival** such as [survreg](#).

Author(s)

T. W. Yee.

References

See [survival](#).

See Also

[SurvS4](#).

Examples

```
showClass("SurvS4")
```

TIC

Takeuchi's Information Criterion

Description

Calculates the Takeuchi information criterion for a fitted model object for which a log-likelihood value has been obtained.

Usage

```
TIC(object, ...)
TICvlm(object, ...)
```

Arguments

`object` A **VGAM** object having class `vglm-class`.
`...` Other possible arguments fed into `logLik` in order to compute the log-likelihood.

Details

The following formula is used for VGLMs: $-2\log\text{-likelihood} + 2\text{trace}(VK)$, where V is the inverse of the EIM from the fitted model, and K is the outer product of the score vectors. Both V and K are order- p . *VLM* matrices. One has V equal to `vcov(object)`, and K is computed by taking the outer product of the output from the `deriv` slot multiplied by the large VLM matrix and then taking their sum. Hence for the huge majority of models, the penalty is computed at the MLE and is empirical in nature. Theoretically, if the fitted model is the true model then AIC equals TIC.

When there are prior weights the score vectors are divided by the square root of these, because $(a_i U_i / \sqrt{a_i})^2 = a_i U_i^2$.

This code relies on the log-likelihood being defined, and computed, for the object. When comparing fitted objects, the smaller the TIC, the better the fit. The log-likelihood and hence the TIC is only defined up to an additive constant.

Currently any estimated scale parameter (in GLM parlance) is ignored by treating its value as unity. Also, currently this function is written only for `vglm` objects and not `vgam` or `rrvglm`, etc., objects.

Value

Returns a numeric TIC value.

Warning

This code has not been double-checked. The general applicability of TIC for the VGLM/VGAM classes has not been developed fully. In particular, TIC should not be run on some **VGAM** family functions because of violation of certain regularity conditions, etc.

Some authors note that quite large sample sizes are needed for this IC to work reasonably well.

Note

TIC has not been defined for RR-VGLMs, QRR-VGLMs, etc., yet.

See [AICvglm](#) about models such as `posbernoulli.tb` that require `posbinomial(omit.constant = TRUE)`.

Author(s)

T. W. Yee.

References

Takeuchi, K. (1976). Distribution of informational statistics and a criterion of model fitting. (In Japanese). *Suri-Kagaku* (Mathematic Sciences), **153**, 12–18.

Burnham, K. P. and Anderson, D. R. (2002). *Model Selection and Multi-Model Inference: A Practical Information-Theoretic Approach*, 2nd ed. New York, USA: Springer.

See Also

VGLMs are described in [vglm-class](#); [AIC](#), [AICvglm](#), [BICvglm](#).

Examples

```
pneumo <- transform(pneumo, let = log(exposure.time))
(fit1 <- vglm(cbind(normal, mild, severe) ~ let,
             cumulative(parallel = TRUE, reverse = TRUE), data = pneumo))
coef(fit1, matrix = TRUE)
TIC(fit1)
(fit2 <- vglm(cbind(normal, mild, severe) ~ let,
             cumulative(parallel = FALSE, reverse = TRUE), data = pneumo))
coef(fit2, matrix = TRUE)
TIC(fit2)
```

Description

Density, distribution function, quantile function and random generation for the Tobit model.

Usage

```
dtobit(x, mean = 0, sd = 1, Lower = 0, Upper = Inf, log = FALSE)
ptobit(q, mean = 0, sd = 1, Lower = 0, Upper = Inf,
       lower.tail = TRUE, log.p = FALSE)
qtobit(p, mean = 0, sd = 1, Lower = 0, Upper = Inf,
       lower.tail = TRUE, log.p = FALSE)
rtobit(n, mean = 0, sd = 1, Lower = 0, Upper = Inf)
```

Arguments

x, q	vector of quantiles.
p	vector of probabilities.
n	number of observations. If <code>length(n) > 1</code> then the length is taken to be the number required.
Lower, Upper	vector of lower and upper thresholds.
mean, sd, lower.tail, log, log.p	see rnorm .

Details

See [tobit](#), the **VGAM** family function for estimating the parameters, for details. Note that the density at Lower and Upper is the the area to the left and right of those points. Thus there are two spikes (but less in value); see the example below. Consequently, `dtobit(Lower) + dtobit(Upper) + the area in between` equals unity.

Value

`dtobit` gives the density, `ptobit` gives the distribution function, `qtobit` gives the quantile function, and `rtobit` generates random deviates.

Author(s)

T. W. Yee

See Also

[tobit](#), [rnorm](#).

Examples

```
mu <- 0.5; x <- seq(-2, 4, by = 0.01)
Lower <- -1; Upper <- 2.0

integrate(dtobit, lower = Lower, upper = Upper,
          mean = mu, Lower = Lower, Upper = Upper)$value +
dtobit(Lower, mean = mu, Lower = Lower, Upper = Upper) +
dtobit(Upper, mean = mu, Lower = Lower, Upper = Upper) # Adds to 1

## Not run:
```

```

plot(x, ptobit(x, m = mu, Lower = Lower, Upper = Upper),
     type = "l", ylim = 0:1, las = 1, col = "orange",
     ylab = paste("ptobit(m = ", mu, ", sd = 1, Lower =", Lower,
                  ", Upper =", Upper, ")"),
     main = "Orange is the CDF; blue is density",
     sub = "Purple lines are the 10,20,...,90 percentiles")
abline(h = 0)
lines(x, dtobit(x, m = mu, L = Lower, U = Upper), col = "blue")

probs <- seq(0.1, 0.9, by = 0.1)
Q <- qtobit(probs, m = mu, Lower = Lower, Upper = Upper)
lines(Q, ptobit(Q, m = mu, Lower = Lower, Upper = Upper),
      col = "purple", lty = "dashed", type = "h")
lines(Q, dtobit(Q, m = mu, Lower = Lower, Upper = Upper),
      col = "darkgreen", lty = "dashed", type = "h")
abline(h = probs, col = "purple", lty = "dashed")
max(abs(ptobit(Q, mu, L = Lower, U = Upper) - probs)) # Should be 0

epts <- c(Lower, Upper) # Endpoints have a spike (not quite, actually)
lines(epts, dtobit(epts, m = mu, Lower = Lower, Upper = Upper),
      col = "blue", lwd = 3, type = "h")

## End(Not run)

```

tobit

Tobit Regression

Description

Fits a Tobit regression model.

Usage

```

tobit(Lower = 0, Upper = Inf, lmu = "identitylink",
      lsd = "loglink", imu = NULL, isd = NULL,
      type.fitted = c("uncensored", "censored", "mean.obs"),
      byrow.arg = FALSE, imethod = 1, zero = "sd")

```

Arguments

Lower	Numeric. It is the value L described below. Any value of the linear model $x_i^T \beta$ that is less than this lowerbound is assigned this value. Hence this should be the smallest possible value in the response variable. May be a vector (see below for more information).
Upper	Numeric. It is the value U described below. Any value of the linear model $x_i^T \beta$ that is greater than this upperbound is assigned this value. Hence this should be the largest possible value in the response variable. May be a vector (see below for more information).

lmu, lsd	Parameter link functions for the mean and standard deviation parameters. See Links for more choices. The standard deviation is a positive quantity, therefore a log link is its default.
imu, isd, byrow.arg	See CommonVGAMffArguments for information.
type.fitted	Type of fitted value returned. The first choice is default and is the ordinary uncensored or unbounded linear model. If "censored" then the fitted values in the interval $[L, U]$. If "mean.obs" then the mean of the observations is returned; this is a doubly truncated normal distribution augmented by point masses at the truncation points (see dtobit). See CommonVGAMffArguments for more information.
imethod	Initialization method. Either 1 or 2 or 3, this specifies some methods for obtaining initial values for the parameters. See CommonVGAMffArguments for information.
zero	A vector, e.g., containing the value 1 or 2. If so, the mean or standard deviation respectively are modelled as an intercept-only. Setting zero = NULL means both linear/additive predictors are modelled as functions of the explanatory variables. See CommonVGAMffArguments for more information.

Details

The Tobit model can be written

$$y_i^* = x_i^T \beta + \varepsilon_i$$

where the $e_i \sim N(0, \sigma^2)$ independently and $i = 1, \dots, n$. However, we measure $y_i = y_i^*$ only if $y_i^* > L$ and $y_i^* < U$ for some cutpoints L and U . Otherwise we let $y_i = L$ or $y_i = U$, whatever is closer. The Tobit model is thus a multiple linear regression but with censored responses if it is below or above certain cutpoints.

The defaults for Lower and Upper and lmu correspond to the *standard* Tobit model. Fisher scoring is used for the standard and nonstandard models. By default, the mean $x_i^T \beta$ is the first linear/additive predictor, and the log of the standard deviation is the second linear/additive predictor. The Fisher information matrix for uncensored data is diagonal. The fitted values are the estimates of $x_i^T \beta$.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Warning

If values of the response and Lower and/or Upper are not integers then there is the danger that the value is wrongly interpreted as uncensored. For example, if the first 10 values of the response were `runif(10)` and Lower was assigned these value then testing `y[1:10] == Lower[1:10]` is numerically fraught. Currently, if any `y < Lower` or `y > Upper` then a warning is issued. The function [round2](#) may be useful.

Note

The response can be a matrix. If so, then Lower and Upper are recycled into a matrix with the number of columns equal to the number of responses, and the recycling is done row-wise *if* `byrow.arg = TRUE`. The default order is as `matrix`, which is `byrow.arg = FALSE`. For example, these are returned in `fit4@misc$Lower` and `fit4@misc$Upper` below.

If there is no censoring then `uninormal` is recommended instead. Any value of the response less than Lower or greater than Upper will be assigned the value Lower and Upper respectively, and a warning will be issued. The fitted object has components `censoredL` and `censoredU` in the extra slot which specifies whether observations are censored in that direction. The function `cens.normal` is an alternative to `tobit()`.

When obtaining initial values, if the algorithm would otherwise want to fit an underdetermined system of equations, then it uses the entire data set instead. This might result in rather poor quality initial values, and consequently, monitoring convergence is advised.

Author(s)

Thomas W. Yee

References

Tobin, J. (1958). Estimation of relationships for limited dependent variables. *Econometrica* **26**, 24–36.

See Also

`rtobit`, `cens.normal`, `uninormal`, `double.cens.normal`, `posnormal`, `CommonVGAMffArguments`, `round2`, `mills.ratio`, `margeff`, `rnorm`.

Examples

```
# Here, fit1 is a standard Tobit model and fit2 is nonstandard
tdata <- data.frame(x2 = seq(-1, 1, length = (nn <- 100)))
set.seed(1)
Lower <- 1; Upper <- 4 # For the nonstandard Tobit model
tdata <- transform(tdata,
                  Lower.vec = rnorm(nn, Lower, 0.5),
                  Upper.vec = rnorm(nn, Upper, 0.5))
meanfun1 <- function(x) 0 + 2*x
meanfun2 <- function(x) 2 + 2*x
meanfun3 <- function(x) 3 + 2*x
tdata <- transform(tdata,
                  y1 = rtobit(nn, mean = meanfun1(x2)), # Standard Tobit model
                  y2 = rtobit(nn, mean = meanfun2(x2), Lower = Lower, Upper = Upper),
                  y3 = rtobit(nn, mean = meanfun3(x2), Lower = Lower.vec,
                              Upper = Upper.vec),
                  y4 = rtobit(nn, mean = meanfun3(x2), Lower = Lower.vec,
                              Upper = Upper.vec))
with(tdata, table(y1 == 0)) # How many censored values?
with(tdata, table(y2 == Lower | y2 == Upper)) # Ditto
with(tdata, table(attr(y2, "cenL")))
```

```

with(tdata, table(attr(y2, "cenU")))

fit1 <- vglm(y1 ~ x2, tobit, data = tdata, trace = TRUE)
coef(fit1, matrix = TRUE)
summary(fit1)

fit2 <- vglm(y2 ~ x2,
            tobit(Lower = Lower, Upper = Upper, type.f = "cens"),
            data = tdata, trace = TRUE)
table(fit2@extra$censoredL)
table(fit2@extra$censoredU)
coef(fit2, matrix = TRUE)

fit3 <- vglm(y3 ~ x2, tobit(Lower = with(tdata, Lower.vec),
                           Upper = with(tdata, Upper.vec),
                           type.f = "cens"),
            data = tdata, trace = TRUE)
table(fit3@extra$censoredL)
table(fit3@extra$censoredU)
coef(fit3, matrix = TRUE)

# fit4 is fit3 but with type.fitted = "uncen".
fit4 <- vglm(cbind(y3, y4) ~ x2,
            tobit(Lower = rep(with(tdata, Lower.vec), each = 2),
                  Upper = rep(with(tdata, Upper.vec), each = 2),
                  byrow.arg = TRUE),
            data = tdata, crit = "coeff", trace = TRUE)
head(fit4@extra$censoredL) # A matrix
head(fit4@extra$censoredU) # A matrix
head(fit4@misc$Lower)     # A matrix
head(fit4@misc$Upper)    # A matrix
coef(fit4, matrix = TRUE)

## Not run: # Plot fit1--fit4
par(mfrow = c(2, 2))

plot(y1 ~ x2, tdata, las = 1, main = "Standard Tobit model",
     col = as.numeric(attr(y1, "cenL")) + 3,
     pch = as.numeric(attr(y1, "cenL")) + 1)
legend(x = "topleft", leg = c("censored", "uncensored"),
      pch = c(2, 1), col = c("blue", "green"))
legend(-1.0, 2.5, c("Truth", "Estimate", "Naive"), lwd = 2,
      col = c("purple", "orange", "black"), lty = c(1, 2, 2))
lines(meanfun1(x2) ~ x2, tdata, col = "purple", lwd = 2)
lines(fitted(fit1) ~ x2, tdata, col = "orange", lwd = 2, lty = 2)
lines(fitted(lm(y1 ~ x2, tdata)) ~ x2, tdata, col = "black",
      lty = 2, lwd = 2) # This is simplest but wrong!

plot(y2 ~ x2, data = tdata, las = 1, main = "Tobit model",
     col = as.numeric(attr(y2, "cenL")) + 3 +
           as.numeric(attr(y2, "cenU")),
     pch = as.numeric(attr(y2, "cenL")) + 1 +
           as.numeric(attr(y2, "cenU")))

```

```

legend(x = "topleft", leg = c("censored", "uncensored"),
      pch = c(2, 1), col = c("blue", "green"))
legend(-1.0, 3.5, c("Truth", "Estimate", "Naive"), lwd = 2,
      col = c("purple", "orange", "black"), lty = c(1, 2, 2))
lines(meanfun2(x2) ~ x2, tdata, col = "purple", lwd = 2)
lines(fitted(fit2) ~ x2, tdata, col = "orange", lwd = 2, lty = 2)
lines(fitted(lm(y2 ~ x2, tdata)) ~ x2, tdata, col = "black",
      lty = 2, lwd = 2) # This is simplest but wrong!

plot(y3 ~ x2, data = tdata, las = 1,
     main = "Tobit model with nonconstant censor levels",
     col = as.numeric(attr(y3, "cenL")) + 2 +
           as.numeric(attr(y3, "cenU")) * 2),
     pch = as.numeric(attr(y3, "cenL")) + 1 +
           as.numeric(attr(y3, "cenU")) * 2))
legend(x = "topleft", pch = c(2, 3, 1), col = c(3, 4, 2),
      leg = c("censoredL", "censoredU", "uncensored"))
legend(-1.0, 3.5, c("Truth", "Estimate", "Naive"), lwd = 2,
      col = c("purple", "orange", "black"), lty = c(1, 2, 2))
lines(meanfun3(x2) ~ x2, tdata, col = "purple", lwd = 2)
lines(fitted(fit3) ~ x2, tdata, col = "orange", lwd = 2, lty = 2)
lines(fitted(lm(y3 ~ x2, tdata)) ~ x2, tdata, col = "black",
      lty = 2, lwd = 2) # This is simplest but wrong!

plot(y3 ~ x2, data = tdata, las = 1,
     main = "Tobit model with nonconstant censor levels",
     col = as.numeric(attr(y3, "cenL")) + 2 +
           as.numeric(attr(y3, "cenU")) * 2),
     pch = as.numeric(attr(y3, "cenL")) + 1 +
           as.numeric(attr(y3, "cenU")) * 2))
legend(x = "topleft", pch = c(2, 3, 1), col = c(3, 4, 2),
      leg = c("censoredL", "censoredU", "uncensored"))
legend(-1.0, 3.5, c("Truth", "Estimate", "Naive"), lwd = 2,
      col = c("purple", "orange", "black"), lty = c(1, 2, 2))
lines(meanfun3(x2) ~ x2, data = tdata, col = "purple", lwd = 2)
lines(fitted(fit4)[, 1] ~ x2, tdata, col="orange", lwd = 2, lty = 2)
lines(fitted(lm(y3 ~ x2, tdata)) ~ x2, data = tdata, col = "black",
      lty = 2, lwd = 2) # This is simplest but wrong!

## End(Not run)

```

Tol

Tolerances

Description

Generic function for the *tolerances* of a model.

Usage

```
Tol(object, ...)
```

Arguments

object	An object for which the computation or extraction of a tolerance or tolerances is meaningful.
...	Other arguments fed into the specific methods function of the model. Sometimes they are fed into the methods function for Coef .

Details

Different models can define an optimum in different ways. Many models have no such notion or definition.

Tolerances occur in quadratic ordination, i.e., CQO and UQO. They have ecological meaning because a high tolerance for a species means the species can survive over a large environmental range (stenoecous species), whereas a small tolerance means the species' niche is small (eurycous species). Mathematically, the tolerance is like the variance of a normal distribution.

Value

The value returned depends specifically on the methods function invoked. For a [cqo](#) binomial or Poisson fit, this function returns a $R \times R \times S$ array, where R is the rank and S is the number of species. Each tolerance matrix ought to be positive-definite, and for a rank-1 fit, taking the square root of each tolerance matrix results in each species' tolerance (like a standard deviation).

Warning

There is a direct inverse relationship between the scaling of the latent variables (site scores) and the tolerances. One normalization is for the latent variables to have unit variance. Another normalization is for all the tolerances to be unit. These two normalization cannot simultaneously hold in general. For rank- $R > 1$ models it becomes more complicated because the latent variables are also uncorrelated. An important argument when fitting quadratic ordination models is whether eq. tolerances is TRUE or FALSE. See Yee (2004) for details.

Note

Tolerances are undefined for 'linear' and additive ordination models. They are well-defined for quadratic ordination models.

Author(s)

Thomas W. Yee

References

- Yee, T. W. (2004). A new technique for maximum-likelihood canonical Gaussian ordination. *Ecological Monographs*, **74**, 685–701.
- Yee, T. W. (2006). Constrained additive ordination. *Ecology*, **87**, 203–213.

See Also

Tol.qrrvglm. [Max](#), [Opt](#), [cqo](#), [rcim](#) for UQO.

Examples

```
## Not run:
set.seed(111) # This leads to the global solution
hspider[,1:6] <- scale(hspider[, 1:6]) # Standardized environmental vars
p1 <- cqo(cbind(Alopacce, Alopcone, Alopfabr, Arctlute, Arctperi,
              Auloalbi, Pardlugu, Pardmont, Pardnigr, Pardpull,
              Trocterr, Zoraspin) ~
          WaterCon + BareSand + FallTwig + CoveMoss + CoveHerb + ReflLux,
          poissonff, data = hspider, Crowlpositive = FALSE)

Tol(p1)

## End(Not run)
```

Topple

The Topp-Leone Distribution

Description

Density, distribution function, quantile function and random generation for the Topp-Leone distribution.

Usage

```
dtopple(x, shape, log = FALSE)
ptopple(q, shape, lower.tail = TRUE, log.p = FALSE)
qtopple(p, shape)
rtopple(n, shape)
```

Arguments

x, q, p, n	Same as Uniform .
shape	the (shape) parameter, which lies in (0, 1).
log	Logical. If log = TRUE then the logarithm of the density is returned.
lower.tail, log.p	Same meaning as in pnorm or qnorm .

Details

See [topple](#), the **VGAM** family function for estimating the (shape) parameter s by maximum likelihood estimation, for the formula of the probability density function.

Value

dtopple gives the density, ptopple gives the distribution function, qtopple gives the quantile function, and rtopple generates random deviates.

Note

The Topp-Leone distribution is related to the triangle distribution.

Author(s)

T. W. Yee

References

Topp, C. W. and F. C. Leone (1955). A family of J-shaped frequency functions. *Journal of the American Statistical Association*, **50**, 209–219.

See Also

[topple](#), [Triangle](#).

Examples

```
## Not run: shape <- 0.7; x <- seq(0.02, 0.999, length = 300)
plot(x, dtopple(x, shape = shape), type = "l", col = "blue",
     main = "Blue is density, orange is CDF", ylab = "", las = 1,
     sub = "Purple lines are the 10,20,...,90 percentiles")
abline(h = 0, col = "blue", lty = 2)
lines(x, ptopple(x, shape = shape), type = "l", col = "orange")
probs <- seq(0.1, 0.9, by = 0.1)
Q <- qtopple(probs, shape = shape)
lines(Q, dtopple(Q, shape), col = "purple", lty = 3, type = "h")
lines(Q, ptopple(Q, shape), col = "purple", lty = 3, type = "h")
abline(h = probs, col = "purple", lty = 3)
max(abs(ptopple(Q, shape) - probs)) # Should be zero

## End(Not run)
```

topple

Topp-Leone Distribution Family Function

Description

Estimating the parameter of the Topp-Leone distribution by maximum likelihood estimation.

Usage

```
topple(lshape = "logitlink", zero = NULL, gshape = ppoints(8),
      parallel = FALSE, percentiles = 50,
      type.fitted = c("mean", "percentiles", "Qlink"))
```

Arguments

- lshape, gshape Details at [CommonVGAMffArguments](#). The CIA link is [loglink](#), for shape approaching unity.
- zero, parallel Details at [CommonVGAMffArguments](#).
- type.fitted, percentiles
See [CommonVGAMffArguments](#) for information. Using "Qlink" is for quantile-links in **VGAMextra**.

Details

The Topples distribution has a probability density function that can be written

$$f(y; s) = 2s(1 - y)[y(2 - y)]^{s-1}$$

for $0 < y < 1$ and shape parameter $0 < s < 1$. The mean of Y is $1 - 4^s[\Gamma(1 + s)]^2/\Gamma(2 + 2s)$ (returned as the fitted values).

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Note

Fisher-scoring and Newton-Raphson are the same here. A related distribution is the triangle distribution. This **VGAM** family function handles multiple responses.

Author(s)

T. W. Yee

References

Topp, C. W. and F. C. Leone (1955). A family of J-shaped frequency functions. *Journal of the American Statistical Association*, **50**, 209–219.

See Also

[Topleft](#), [Triangle](#).

Examples

```
tdata <- data.frame(y = rtopple(1000, logitlink(1, inverse = TRUE)))
tfit <- vglm(y ~ 1, topple, tdata, trace = TRUE, crit = "coef")
coef(tfit, matrix = TRUE)
Coef(tfit)
```

toxop

Toxoplasmosis Data

Description

Toxoplasmosis data in 34 cities in El Salvador.

Usage

```
data(toxop)
```

Format

A data frame with 34 observations on the following 4 variables.

rainfall a numeric vector; the amount of rainfall in each city.

ssize a numeric vector; sample size.

cityNo a numeric vector; the city number.

positive a numeric vector; the number of subjects testing positive for the disease.

Details

See the references for details.

Source

See the references for details.

References

Efron, B. (1978). Regression and ANOVA With zero-one data: measures of residual variation. *Journal of the American Statistical Association*, **73**, 113–121.

Efron, B. (1986). Double exponential families and their use in generalized linear regression. *Journal of the American Statistical Association*, **81**, 709–721.

See Also

[double.expbinoial](#).

Examples

```
## Not run: with(toxop, plot(rainfall, positive/ssize, col = "blue"))
plot(toxop, col = "blue")
## End(Not run)
```

Description

Density, distribution function, quantile function and random generation for the Triangle distribution with parameter θ .

Usage

```
dtriangle(x, theta, lower = 0, upper = 1, log = FALSE)
ptriangle(q, theta, lower = 0, upper = 1, lower.tail = TRUE, log.p = FALSE)
qttriangle(p, theta, lower = 0, upper = 1, lower.tail = TRUE, log.p = FALSE)
rttriangle(n, theta, lower = 0, upper = 1)
```

Arguments

<code>x, q</code>	vector of quantiles.
<code>p</code>	vector of probabilities.
<code>n</code>	number of observations. Same as runif .
<code>theta</code>	the θ parameter which lies between lower and upper.
<code>lower, upper</code>	lower and upper limits of the distribution. Must be finite.
<code>log</code>	Logical. If <code>log = TRUE</code> then the logarithm of the density is returned.
<code>lower.tail, log.p</code>	Same meaning as in pnorm or qnorm .

Details

See [triangle](#), the **VGAM** family function for estimating the parameter θ by maximum likelihood estimation, however the regular conditions do not hold so the algorithm crawls to the solution if lucky.

Value

`dtriangle` gives the density, `ptriangle` gives the distribution function, `qttriangle` gives the quantile function, and `rttriangle` generates random deviates.

Author(s)

T. W. Yee and Kai Huang

See Also

[triangle](#), [topple](#).

Examples

```
## Not run: x <- seq(-0.1, 1.1, by = 0.01); theta <- 0.75
plot(x, dtriangle(x, theta = theta), type = "l", col = "blue", las = 1,
     main = "Blue is density, orange is the CDF",
     sub = "Purple lines are the 10,20,...,90 percentiles",
     ylim = c(0,2), ylab = "")
abline(h = 0, col = "blue", lty = 2)
lines(x, ptriangle(x, theta = theta), col = "orange")
probs <- seq(0.1, 0.9, by = 0.1)
Q <- qtriangle(probs, theta = theta)
lines(Q, dtriangle(Q, theta = theta), col = "purple", lty = 3, type = "h")
ptriangle(Q, theta = theta) - probs # Should be all zero
abline(h = probs, col = "purple", lty = 3)
## End(Not run)
```

trim.constraints

Trimmed Constraint Matrices

Description

Deletes statistically nonsignificant regression coefficients via their constraint matrices, for future refitting.

Usage

```
trim.constraints(object, sig.level = 0.05, max.num = Inf,
                intercepts = TRUE, ...)
```

Arguments

object	Some VGAM object, especially having class <code>vglmff-class</code> . It has not yet been tested on non- <code>vglm</code> objects.
sig.level	Significance levels, with values in $[0, 1]$. Columns of constraint matrices whose p-values are larger than this argument are deleted. With terms that generate more than one column of the "lm" model matrix, all p-values must be greater than this argument for deletion. This argument is recycled to the total number of regression coefficients of object.
max.num	Numeric, positive and integer-valued. Maximum number of regression coefficients allowable for deletion. This allows one to limit the number of deleted coefficients. For example, if <code>max.num = 1</code> then only the largest p-value is used for the deletion, provided it is larger than <code>sig.level</code> . The default is to delete all those coefficients whose p-values are greater than <code>sig.level</code> . With a finite value, this argument will probably not work properly when there are terms that generate more than one column of the LM model matrix. Having a value greater than unity might be unsuitable in the presence of multicollinearity because all correlated variables might be eliminated at once.

intercepts	Logical. Trim the intercept term? If FALSE then the constraint matrix for the "(Intercept)" term is left unchanged.
...	Unused but for provision in the future.

Details

This utility function is intended to simplify an existing `vglm` object having variables (terms) that affect unnecessary parameters. Suppose the explanatory variables in the formula includes a simple numeric covariate called `x2`. This variable will affect every linear predictor if `zero = NULL` in the **VGAM** family function. This situation may correspond to the constraint matrices having unnecessary columns because their regression coefficients are statistically nonsignificant. This function attempts to delete those columns and return a possibly simplified list of constraint matrices that can make refitting a simpler model easy to do. P-values obtained from `summaryvglm` (with `HDEtest = FALSE` for increased speed) are compared to `sig.level` to test for statistical significance.

For terms that generate more than one column of the "lm" model matrix, such as `bs` and `poly`, the column is deleted if all regression coefficients are statistically nonsignificant. Incidentally, users should instead use `sm.bs`, `sm.ns`, `sm.poly`, etc., for smart and safe prediction.

One can think of this function as facilitating *backward elimination* for variable selection, especially if `max.num = 1` and $M = 1$, however usually more than one regression coefficient is deleted here by default.

Value

A list of possibly simpler constraint matrices that can be fed back into the model using the `constraints` argument (usually `zero = NULL` is needed to avoid a warning). Consequently, they are required to be of the "term"-type. After the model is refitted, applying `summaryvglm` should result in regression coefficients that are 'all' statistically significant.

Warning

This function has not been tested thoroughly. One extreme is that a term is totally deleted because none of its regression coefficients are needed, and that situation has not yet been finalized. Ideally, object only contains terms where at least one regression coefficient has a p-value less than `sig.level`. For ordered factors and other situations, deleting certain columns may not make sense and destroy interpretability.

As stated above, `max.num` may not work properly when there are terms that generate more than one column of the LM model matrix. However, this limitation may change in the future.

Note

This function is experimental and may be replaced by some other function in the future. This function does not use S4 object oriented programming but may be converted to such in the future.

Author(s)

T. W. Yee

See Also

[constraints](#), [vglm](#), [summaryvglm](#), [model.matrixvglm](#), [drop1.vglm](#), [step4vglm](#), [sm.bs](#), [sm.ns](#), [sm.poly](#).

Examples

```
## Not run: data("xs.nz", package = "VGAMdata")
fit1 <-
  vglm(cbind(worry, worrier) ~ bs(age) + sex + ethnicity + cat + dog,
        binom2.or(zero = NULL), data = xs.nz, trace = TRUE)
summary(fit1, HDEtest = FALSE) # 'cat' is not significant at all
dim(constraints(fit1, matrix = TRUE))
(tclist1 <- trim.constraints(fit1)) # No 'cat'
fit2 <- # Delete 'cat' manually from the formula:
  vglm(cbind(worry, worrier) ~ bs(age) + sex + ethnicity + dog,
        binom2.or(zero = NULL), data = xs.nz,
        constraints = tclist1, trace = TRUE)
summary(fit2, HDEtest = FALSE) # A simplified model
dim(constraints(fit2, matrix = TRUE)) # Fewer regression coefficients

## End(Not run)
```

Trinorm

*Trivariate Normal Distribution Density and Random Variates***Description**

Density and random generation for the trivariate normal distribution distribution.

Usage

```
dtrinorm(x1, x2, x3, mean1 = 0, mean2 = 0, mean3 = 0,
          var1 = 1, var2 = 1, var3 = 1,
          cov12 = 0, cov23 = 0, cov13 = 0, log = FALSE)
rtrinorm(n, mean1 = 0, mean2 = 0, mean3 = 0,
          var1 = 1, var2 = 1, var3 = 1,
          cov12 = 0, cov23 = 0, cov13 = 0)
```

Arguments

`x1, x2, x3` vector of quantiles.
`mean1, mean2, mean3` vectors of means.
`var1, var2, var3` vectors of variances.
`cov12, cov23, cov13` vectors of covariances.
`n` number of observations. Same as [rnorm](#).
`log` Logical. If `log = TRUE` then the logarithm of the density is returned.

Details

The default arguments correspond to the standard trivariate normal distribution with correlation parameters equal to 0, which corresponds to three independent standard normal distributions. Let sd_1 (say) be $\sqrt{\text{var1}}$ and written σ_1 , etc. Then the general formula for each correlation coefficient is of the form $\rho_{12} = \text{cov}_{12}/(\sigma_1\sigma_2)$, and similarly for the two others. Thus if the var arguments are left alone then the cov can be inputted with ρ s.

Value

`dtrinorm` gives the density, `rtrinorm` generates random deviates (n by 3 matrix).

Warning

`dtrinorm()`'s arguments might change in the future! It's safest to use the full argument names to future-proof possible changes!

Note

For `rtrinorm()`, if the i th variance-covariance matrix is not positive-definite then the i th row is all NAs.

See Also

[pnorm](#), [trinormal](#), [uninormal](#), [binormal](#), [rbinorm](#).

Examples

```
## Not run: nn <- 1000
tdata <- data.frame(x2 = sort(runif(nn)))
tdata <- transform(tdata, mean1 = 1 + 2 * x2,
                  mean2 = 3 + 1 * x2, mean3 = 4,
                  var1 = exp( 1), var2 = exp( 1), var3 = exp( 1),
                  rho12 = rhobitlink( 1, inverse = TRUE),
                  rho23 = rhobitlink( 1, inverse = TRUE),
                  rho13 = rhobitlink(-1, inverse = TRUE))
ymat <- with(tdata, rtrinorm(nn, mean1, mean2, mean3,
                             var1, var2, var3,
                             sqrt(var1)*sqrt(var1)*rho12,
                             sqrt(var2)*sqrt(var3)*rho23,
                             sqrt(var1)*sqrt(var3)*rho13))

pairs(ymat, col = "blue")

## End(Not run)
```

trinormal

*Trivariate Normal Distribution Family Function***Description**

Maximum likelihood estimation of the nine parameters of a trivariate normal distribution.

Usage

```
trinormal(zero = c("sd", "rho"), eq.mean = FALSE,
  eq.sd = FALSE, eq.cor = FALSE,
  lmean1 = "identitylink", lmean2 = "identitylink",
  lmean3 = "identitylink",
  lsd1 = "loglink", lsd2 = "loglink", lsd3 = "loglink",
  lrho12 = "rhobitlink", lrho23 = "rhobitlink", lrho13 = "rhobitlink",
  imean1 = NULL, imean2 = NULL, imean3 = NULL,
  isd1 = NULL, isd2 = NULL, isd3 = NULL,
  irho12 = NULL, irho23 = NULL, irho13 = NULL, imethod = 1)
```

Arguments

`lmean1`, `lmean2`, `lmean3`, `lsd1`, `lsd2`, `lsd3`

Link functions applied to the means and standard deviations. See [Links](#) for more choices. Being positive quantities, a log link is the default for the standard deviations.

`lrho12`, `lrho23`, `lrho13`

Link functions applied to the correlation parameters. See [Links](#) for more choices. By default the correlation parameters are allowed to have a value between -1 and 1, but that may be problematic when `eq.cor = TRUE` because they should have a value between -0.5 and 1.

`imean1`, `imean2`, `imean3`, `isd1`, `isd2`, `isd3`

See [CommonVGAMffArguments](#) for more information.

`irho12`, `irho23`, `irho13`, `imethod`, `zero`

See [CommonVGAMffArguments](#) for more information.

`eq.mean`, `eq.sd`, `eq.cor`

Logical. Constrain the means or the standard deviations or correlation parameters to be equal?

Details

For the trivariate normal distribution, this fits a linear model (LM) to the means, and by default, the other parameters are intercept-only. The response should be a three-column matrix. The three correlation parameters are prefixed by rho, and the default gives them values between -1 and 1 however, this may be problematic when the correlation parameters are constrained to be equal, etc.. The fitted means are returned as the fitted values, which is in the form of a three-column matrix. Fisher scoring is implemented.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Warning

The default parameterization does not make the estimated variance-covariance matrix positive-definite. In order for the variance-covariance matrix to be positive-definite the quantity $1 - \rho_{12}^2 - \rho_{13}^2 - \rho_{23}^2 + 2 * \rho_{12} * \rho_{13} * \rho_{23}$ must be positive, and if `eq.cor = TRUE` then this means that the rhos must be between -0.5 and 1.

Author(s)

T. W. Yee

See Also

[uninormal](#), [binormal](#), [rtrinorm](#).

Examples

```
## Not run:  set.seed(123); nn <- 1000
tdata <- data.frame(x2 = runif(nn), x3 = runif(nn))
tdata <- transform(tdata, y1 = rnorm(nn, 1 + 2 * x2),
                  y2 = rnorm(nn, 3 + 4 * x2),
                  y3 = rnorm(nn, 4 + 5 * x2))
fit1 <- vglm(cbind(y1, y2, y3) ~ x2, data = tdata,
            trinormal(eq.sd = TRUE, eq.cor = TRUE), trace = TRUE)
coef(fit1, matrix = TRUE)
constraints(fit1)
summary(fit1)
# Try this when eq.sd = TRUE, eq.cor = TRUE:
fit2 <-
  vglm(cbind(y1, y2, y3) ~ x2, data = tdata, stepsize = 0.25,
      trinormal(eq.sd = TRUE, eq.cor = TRUE,
                lrho12 = extlogitlink(min = -0.5),
                lrho23 = extlogitlink(min = -0.5),
                lrho13 = extlogitlink(min = -0.5)), trace = TRUE)
coef(fit2, matrix = TRUE)

## End(Not run)
```

trplot

Trajectory Plot

Description

Generic function for a trajectory plot.

Usage

```
trplot(object, ...)
```

Arguments

object	An object for which a trajectory plot is meaningful.
...	Other arguments fed into the specific methods function of the model. They usually are graphical parameters, and sometimes they are fed into the methods function for Coef .

Details

Trajectory plots can be defined in different ways for different models. Many models have no such notion or definition.

For quadratic and additive ordination models they plot the fitted values of two species against each other (more than two is theoretically possible, but not implemented in this software yet).

Value

The value returned depends specifically on the methods function invoked.

Author(s)

Thomas W. Yee

References

Yee, T. W. (2020). On constrained and unconstrained quadratic ordination. *Manuscript in preparation*.

See Also

[trplot.qrrvglm](#), [persprrvglm](#), [lvplot](#).

Examples

```
## Not run: set.seed(123)
hspider[, 1:6] <- scale(hspider[, 1:6]) # Stdze environ. vars
p1cqo <- cqc(cbind(Alopacce, Alopcune, Alopfabr, Arctlute,
                 Arctperi, Auloalbi, Pardlugu, Pardmont,
                 Pardnigr, Pardpull, Trocterr, Zoraspin) ~
            WaterCon + BareSand + FallTwig +
            CoveMoss + CoveHerb + ReflLux,
            poissonff, data = hspider, Crow1positive = FALSE)

nos <- ncol(depvar(p1cqo))
clr <- 1:nos # OR (1:(nos+1))[-7] to omit yellow

trplot(p1cqo, which.species = 1:3, log = "xy", lwd = 2,
       col = c("blue", "orange", "green"), label = TRUE) -> ii
```

```

legend(0.00005, 0.3, paste(ii$species[, 1], ii$species[, 2],
                           sep = " and "),
       lwd = 2, lty = 1, col = c("blue", "orange", "green"))
abline(a = 0, b = 1, lty = "dashed", col = "grey")
## End(Not run)

```

trplot.qrrvglm

Trajectory plot for QRR-VGLMs

Description

Produces a trajectory plot for *quadratic reduced-rank vector generalized linear models* (QRR-VGLMs). It is only applicable for rank-1 models with argument `noRRR = ~ 1`.

Usage

```

trplot.qrrvglm(object, which.species = NULL, add = FALSE,
               show.plot = TRUE,
               label.sites = FALSE, sitenames = rownames(object@y),
               axes.equal = TRUE, cex = par()$cex,
               col = 1:(nos * (nos - 1)/2), log = "",
               lty = rep_len(par()$lty, nos * (nos - 1)/2),
               lwd = rep_len(par()$lwd, nos * (nos - 1)/2),
               tcol = rep_len(par()$col, nos * (nos - 1)/2),
               xlab = NULL, ylab = NULL,
               main = "", type = "b", check.ok = TRUE, ...)

```

Arguments

<code>object</code>	Object of class "qrrvglm", i.e., a CQO object.
<code>which.species</code>	Integer or character vector specifying the species to be plotted. If integer, these are the columns of the response matrix. If character, these must match exactly with the species' names. The default is to use all species.
<code>add</code>	Logical. Add to an existing plot? If FALSE (default), a new plot is made.
<code>show.plot</code>	Logical. Plot it?
<code>label.sites</code>	Logical. If TRUE, the points on the curves/trajectories are labelled with the <code>sitenames</code> .
<code>sitenames</code>	Character vector. The names of the sites.
<code>axes.equal</code>	Logical. If TRUE, the x- and y-axes will be on the same scale.
<code>cex</code>	Character expansion of the labelling of the site names. Used only if <code>label.sites</code> is TRUE. See the <code>cex</code> argument in par .
<code>col</code>	Color of the lines. See the <code>col</code> argument in par . Here, <code>nos</code> is the number of species.
<code>log</code>	Character, specifying which (if any) of the x- and y-axes are to be on a logarithmic scale. See the <code>log</code> argument in par .

lty	Line type. See the lty argument of par .
lwd	Line width. See the lwd argument of par .
tcol	Color of the text for the site names. See the col argument in par . Used only if label.sites is TRUE.
xlab	Character caption for the x-axis. By default, a suitable caption is found. See the xlab argument in plot or title .
ylab	Character caption for the y-axis. By default, a suitable caption is found. See the xlab argument in plot or title .
main	Character, giving the title of the plot. See the main argument in plot or title .
type	Character, giving the type of plot. A common option is to use type="l" for lines only. See the type argument of plot .
check.ok	Logical. Whether a check is performed to see that noRRR = ~ 1 was used. It doesn't make sense to have a trace plot unless this is so.
...	Arguments passed into the plot function when setting up the entire plot. Useful arguments here include xlim and ylim.

Details

A trajectory plot plots the fitted values of a 'second' species against a 'first' species. The argument which.species must therefore contain at least two species. By default, all of the species that were fitted in object are plotted. With more than a few species the resulting plot will be very congested, and so it is recommended that only a few species be selected for plotting.

In the above, M is the number of species selected for plotting, so there will be $M(M - 1)/2$ curves/trajectories in total.

A trajectory plot will be fitted only if noRRR = ~ 1 because otherwise the trajectory will not be a smooth function of the latent variables.

Value

A list with the following components.

species.names	A matrix of characters giving the 'first' and 'second' species. The number of different combinations of species is given by the number of rows. This is useful for creating a legend.
sitenames	A character vector of site names, sorted by the latent variable (from low to high).

Note

Plotting the axes on a log scale is often a good idea. The use of xlim and ylim to control the axis limits is also a good idea, so as to limit the extent of the curves at low abundances or probabilities. Setting label.sites = TRUE is a good idea only if the number of sites is small, otherwise there is too much clutter.

Author(s)

Thomas W. Yee

References

Yee, T. W. (2020). On constrained and unconstrained quadratic ordination. *Manuscript in preparation*.

See Also

[cgo](#), [par](#), [title](#).

Examples

```
## Not run: set.seed(111) # Leads to the global solution
# hspider[,1:6] <- scale(hspider[,1:6]) # Stdize the environ vars
p1 <- cgo(cbind(Alopacce, Alopcone, Alopfabr, Arctlute,
               Arctperi, Auloalbi, Pardlugu, Pardmont,
               Pardnigr, Pardpull, Trocterr, Zoraspin) ~
          WaterCon + BareSand + FallTwig + CoveMoss +
          CoveHerb + ReflLux,
          poissonff, data = hspider, trace = FALSE)

trplot(p1, which.species = 1:3, log = "xy", type = "b", lty = 1,
       main = "Trajectory plot of three hunting spiders species",
       col = c("blue", "red", "green"), lwd = 2, label = TRUE) -> ii
legend(0.00005, 0.3, lwd = 2, lty = 1,
      col = c("blue", "red", "green"),
      with(ii, paste(species.names[,1], species.names[,2],
                    sep = " and ")))
abline(a = 0, b = 1, lty = "dashed", col = "grey") # Ref. line

## End(Not run)
```

Trunc

Truncated Values for the GT-Expansion Method

Description

Given the minimum and maximum values in a response variable, and a positive multiplier, returns the truncated values for generally-truncated regression

Usage

```
Trunc(Range, mux = 2, location = 0, omits = TRUE)
```

Arguments

Range	Numeric, of length 2 containing the minimum and maximum (in that order) of the untransformed data. Alternatively, if <code>length(Range) > 2</code> then it is assumed that the entire untransformed data is passed in so that range is applied.
mux	Numeric, the multiplier. A positive integer.

location	Numeric, the location parameter, allows a shift to the right.
omits	Logical. The default is to return the truncated values (those being omitted). If FALSE then the multiples are returned.

Details

Generally-truncated regression can handle underdispersion with respect to some parent or base distribution such as the Poisson. Yee and Ma (2023) call this the *GT-Expansion* (GTE) method, which is a special case of the GT-location-scale (GT-LS) method. This is a utility function to help make life easier. It is assumed that the response is a count variable.

Value

A vector of values to be fed into the truncate argument of a **VGAM** family function such as [gaitdpoisson](#). If `mux = 1` then the function will return a NULL rather than `integer(0)`.

Author(s)

T. W. Yee

See Also

[gaitdpoisson](#), [gaitdlog](#), [gaitdzeta](#), [range](#), [setdiff](#), [goffset](#).

Examples

```
Trunc(c(1, 8), 2)

## Not run:
set.seed(1) # The following example is based on the normal
mymean <- 20; m.truth <- 3 # approximation to the Poisson.
gdata <- data.frame(y1 = round(rnorm((nn <- 1000), mymean,
                                sd = sqrt(mymean / m.truth))))
org1 <- with(gdata, range(y1)) # Original range of the raw data
m.max <- 5 # Try multipliers 1:m.max
logliks <- numeric(m.max)
names(logliks) <- as.character(1:m.max)
for (i in 1:m.max) {
  logliks[i] <- logLik(vglm(i * y1 ~ offset(rep(log(i), nn)),
    gaitdpoisson(truncate = Trunc(org1, i)), data = gdata))
}
sort(logliks, decreasing = TRUE) # Best to worst
par(mfrow = c(1, 2))
plot(with(gdata, table(y1))) # Underdispersed wrt Poisson
plot(logliks, col = "blue", type = "b", xlab = "Multiplier")
## End(Not run)
```

Truncpareto

The Truncated Pareto Distribution

Description

Density, distribution function, quantile function and random generation for the upper truncated Pareto(I) distribution with parameters lower, upper and shape.

Usage

```
dtruncpareto(x, lower, upper, shape, log = FALSE)
ptruncpareto(q, lower, upper, shape, lower.tail = TRUE, log.p = FALSE)
qtruncpareto(p, lower, upper, shape)
rtruncpareto(n, lower, upper, shape)
```

Arguments

x, q	vector of quantiles.
p	vector of probabilities.
n, log	Same meaning as runif .
lower, upper, shape	the lower, upper and shape (k) parameters. If necessary, values are recycled.
lower.tail, log.p	Same meaning as in pnorm or qnorm .

Details

See [truncpareto](#), the **VGAM** family function for estimating the parameter k by maximum likelihood estimation, for the formula of the probability density function and the range restrictions imposed on the parameters.

Value

`dtruncpareto` gives the density, `ptruncpareto` gives the distribution function, `qtruncpareto` gives the quantile function, and `rtruncpareto` generates random deviates.

Author(s)

T. W. Yee and Kai Huang

References

Aban, I. B., Meerschaert, M. M. and Panorska, A. K. (2006). Parameter estimation for the truncated Pareto distribution, *Journal of the American Statistical Association*, **101**(473), 270–277.

See Also

[truncpareto](#).

Examples

```

lower <- 3; upper <- 8; kay <- exp(0.5)
## Not run: xx <- seq(lower - 0.5, upper + 0.5, len = 401)
plot(xx, dtruncpareto(xx, low = lower, upp = upper, shape = kay),
      main = "Truncated Pareto density split into 10 equal areas",
      type = "l", ylim = 0:1, xlab = "x")
abline(h = 0, col = "blue", lty = 2)
qq <- qtruncpareto(seq(0.1, 0.9, by = 0.1), low = lower, upp = upper,
                  shape = kay)
lines(qq, dtruncpareto(qq, low = lower, upp = upper, shape = kay),
      col = "purple", lty = 3, type = "h")
lines(xx, ptruncpareto(xx, low = lower, upp = upper, shape = kay),
      col = "orange")
## End(Not run)
pp <- seq(0.1, 0.9, by = 0.1)
qq <- qtruncpareto(pp, lower = lower, upper = upper, shape = kay)

ptruncpareto(qq, lower = lower, upper = upper, shape = kay)
qtruncpareto(ptruncpareto(qq, lower = lower, upper = upper, shape = kay),
             lower = lower, upper = upper, shape = kay) - qq # Should be all 0

```

truncweibull

*Truncated Weibull Distribution Family Function***Description**

Maximum likelihood estimation of the 2-parameter Weibull distribution with lower truncation. No observations should be censored.

Usage

```

truncweibull(lower.limit = 1e-5,
             lAlpha = "loglink", lBetaa = "loglink",
             iAlpha = NULL, iBetaa = NULL,
             nrfs = 1, probs.y = c(0.2, 0.5, 0.8),
             imethod = 1, zero = "Betaa")

```

Arguments

lower.limit Positive lower truncation limits. Recycled to the same dimension as the response, going across rows first. The default, being close to 0, should mean effectively the same results as [weibullR](#) if there are no response values that are smaller.

lAlpha, lBetaa Parameter link functions applied to the (positive) parameters Alpha (called α below) and (positive) Betaa (called β below). See [Links](#) for more choices.

iAlpha, iBetaa See [CommonVGAMffArguments](#).

imethod, nrfs, zero, probs.y Details at [weibullR](#) and [CommonVGAMffArguments](#).

Details

MLE of the two parameters of the Weibull distribution are computed, subject to lower truncation. That is, all response values are greater than `lower.limit`, element-wise. For a particular observation this is any known positive value. This function is currently based directly on Wingo (1989) and his parameterization is used (it differs from `weibullR`.) In particular, $\beta = a$ and $\alpha = (1/b)^a$ where a and b are as in `weibullR` and `dweibull`.

Upon fitting the extra slot has a component called `lower.limit` which is of the same dimension as the response. The fitted values are the mean, which are computed using `pgamma.deriv` and `pgamma.deriv.unscaled`.

Value

An object of class "vglmff" (see `vglmff-class`). The object is used by modelling functions such as `vglm`, and `vgam`.

Warning

This function may be converted to the same parameterization as `weibullR` at any time. Yet to do: one element of the EIM may be wrong (due to two interpretations of a formula; but it seems to work). Convergence is slower than usual and this may imply something is wrong; use argument `maxit`. In fact, it's probably because `pgamma.deriv.unscaled` is inaccurate at $q = 1$ and $q = 2$. Also, convergence should be monitored, especially if the truncation means that a large proportion of the data is lost compared to an ordinary Weibull distribution.

Note

More improvements need to be made, e.g., initial values are currently based on no truncation. This **VGAM** family function handles multiple responses.

Author(s)

T. W. Yee

References

Wingo, D. R. (1989). The left-truncated Weibull distribution: theory and computation. *Statistical Papers*, **30**(1), 39–48.

See Also

`weibullR`, `dweibull`, `pgamma.deriv`, `pgamma.deriv.unscaled`.

Examples

```
## Not run:
nn <- 5000; prop.lost <- 0.40 # Proportion lost to truncation
wdata <- data.frame(x2 = runif(nn)) # Complete Weibull data
wdata <- transform(wdata,
  Betaa = exp(1)) # > 2 okay (satisfies regularity conds)
wdata <- transform(wdata, Alpha = exp(0.5 - 1 * x2))
```



```

wdata <- transform(wdata, Shape = Betaa,
#               aaa = Betaa,
#               bbb = 1 / Alpha^(1 / Betaa),
#               Scale = 1 / Alpha^(1 / Betaa))
wdata <- transform(wdata, y2 = rweibull(nn, Shape, scale = Scale))
summary(wdata)

# Proportion lost:
lower.limit2 <- with(wdata, quantile(y2, prob = prop.lost))
# Smaller due to truncation:
wdata <- subset(wdata, y2 > lower.limit2)

fit1 <- vglm(y2 ~ x2, maxit = 100, trace = TRUE,
            truncweibull(lower.limit = lower.limit2), wdata)
coef(fit1, matrix = TRUE)
summary(fit1)
vcov(fit1)
head(fit1@extra$lower.limit)

## End(Not run)

```

ucberk

University California Berkeley Graduate Admissions

Description

University California Berkeley Graduate Admissions: counts cross-classified by acceptance/rejection and gender, for the six largest departments.

Usage

```
data(ucberk)
```

Format

A data frame with 6 departmental groups with the following 5 columns.

m.deny Counts of men denied admission.

m.admit Counts of men admitted.

w.deny Counts of women denied admission.

w.admit Counts of women admitted.

dept Department (the six largest), called A, B, ..., F.

Details

From Bickel et al. (1975), the data consists of applications for admission to graduate study at the University of California, Berkeley, for the fall 1973 quarter. In the admissions cycle for that quarter, the Graduate Division at Berkeley received approximately 15,000 applications, some of which were later withdrawn or transferred to a different proposed entry quarter by the applicants. Of the applications finally remaining for the fall 1973 cycle 12,763 were sufficiently complete to permit a decision. There were about 101 graduate department and interdepartmental graduate majors. There were 8442 male applicants and 4321 female applicants. About 44 percent of the males and about 35 percent of the females were admitted. The data are well-known for illustrating Simpson's paradox.

References

Bickel, P. J., Hammel, E. A. and O'Connell, J. W. (1975). Sex bias in graduate admissions: data from Berkeley. *Science*, **187**(4175): 398–404.

Freedman, D., Pisani, R. and Purves, R. (1998). Chapter 2 of *Statistics*, 3rd. ed., W. W. Norton & Company.

Examples

```
summary(ucberk)
```

 uninormal

Univariate Normal Distribution

Description

Maximum likelihood estimation of the two parameters of a univariate normal distribution.

Usage

```
uninormal(lmean = "identitylink", lsd = "loglink", lvar =
  "loglink", var.arg = FALSE, imethod = 1, isd = NULL,
  parallel = FALSE, vfl = FALSE, Form2 = NULL,
  smallno = 1e-05, zero = if (var.arg)
  "var" else "sd")
```

Arguments

<code>lmean, lsd, lvar</code>	Link functions applied to the mean and standard deviation/variance. See Links for more choices. Being positive quantities, a log link is the default for the standard deviation and variance (see <code>var.arg</code>).
<code>var.arg</code>	Logical. If TRUE then the second parameter is the variance and <code>lsd</code> and <code>esd</code> are ignored, else the standard deviation is used and <code>lvar</code> and <code>evvar</code> are ignored.
<code>smallno</code>	Numeric, positive but close to 0. Used specifically for quasi-variances; if the link for the mean is explink then any non-positive value of η is replaced by this quantity (hopefully, temporarily and only during early iterations).

imethod, parallel, isd, zero
 See [CommonVGAMffArguments](#) for information. If `lmean = loglink` then try `imethod = 2`. If `parallel = TRUE` then the parallelism constraint is not applied to the intercept.

vf1, Form2 See [CommonVGAMffArguments](#) for information.

Details

This fits a linear model (LM) as the first linear/additive predictor. So, by default, this is just the mean. By default, the log of the standard deviation is the second linear/additive predictor. The Fisher information matrix is diagonal. This **VGAM** family function can handle multiple responses.

Value

An object of class "vg1mff" (see [vg1mff-class](#)). The object is used by modelling functions such as [vg1m](#), and [vgam](#).

Warning

`gaussianff()` was deprecated but has been brought back into **VGAM** nominally. It should be called Mickey Mouse. It gives a warning and calls [uninormal](#) instead (hopefully all the arguments should pass in correctly). Users should avoid calling `gaussianff()`; use `glm` with [gaussian](#) instead. It is dangerous to treat what is an [uninormal](#) fit as a `gaussianff()` object.

Note

Yet to do: allow an argument such as `eq.sd` that enables the standard deviations to be the same. Also, this function used to be called `normal1()` too, but it has been decommissioned.

Author(s)

T. W. Yee

References

Forbes, C., Evans, M., Hastings, N. and Peacock, B. (2011). *Statistical Distributions*, Hoboken, NJ, USA: John Wiley and Sons, Fourth edition.

See Also

[posnormal](#), [mix2normal](#), [ordsup](#), [N1binomial](#), [N1poisson](#), [Qvar](#), [tobit](#), [cens.normal](#), [foldnormal](#), [skewnormal](#), [double.cens.normal](#), [SURff](#), [AR1](#), [huber2](#), [studentt](#), [binormal](#), [trinormal](#), [dnorm](#), [simulate.vlm](#), [hdeff.vglm](#).

Examples

```

udata <- data.frame(x2 = rnorm(nn <- 200))
udata <- transform(udata,
  y1 = rnorm(nn, m = 1 - 3*x2, sd = exp(1 + 0.2*x2)),
  y2a = rnorm(nn, m = 1 + 2*x2, sd = exp(1 + 2.0*x2)^0.5),
  y2b = rnorm(nn, m = 1 + 2*x2, sd = exp(1 + 2.0*x2)^0.5))

```

```

fit1 <- vglm(y1 ~ x2, uninormal(zero = NULL), udata, trace = TRUE)
coef(fit1, matrix = TRUE)
fit2 <- vglm(cbind(y2a, y2b) ~ x2, data = udata, trace = TRUE,
             uninormal(var = TRUE, parallel = TRUE ~ x2,
                       zero = NULL))
coef(fit2, matrix = TRUE)

# Generate data from N(mu=theta=10, sigma=theta) and estimate theta.
theta <- 10
udata <- data.frame(y3 = rnorm(100, m = theta, sd = theta))
fit3a <- vglm(y3 ~ 1, uninormal(lsd = "identitylink"), data = udata,
             constraints = list("Intercept" = rbind(1, 1)))
fit3b <- vglm(y3 ~ 1, uninormal(lsd = "identitylink",
                               parallel = TRUE ~ 1, zero = NULL), udata)
coef(fit3a, matrix = TRUE)
coef(fit3b, matrix = TRUE) # Same as fit3a

```

UtilitiesVGAM

Utility Functions for the VGAM Package

Description

A set of common utility functions used by **VGAM** family functions.

Usage

```

param.names(string, S = 1, skip1 = FALSE, sep = "")
dimm(M, hbw = M)
interleave.VGAM(.M, M1, inverse = FALSE)

```

Arguments

string	Character. Name of the parameter.
M, .M	Numeric. The total number of linear/additive predictors, called M . By total, it is meant summed over the number of responses. Often, M is the total number of parameters to be estimated (but this is not the same as the number of regression coefficients, unless the RHS of the formula is an intercept-only). The use of $.M$ is unfortunate, but it is a compromise solution to what is presented in Yee (2015). Ideally, $.M$ should be just M .
M1	Numeric. The number of linear/additive predictors for one response, called M_1 . This argument used to be called M , but is now renamed properly.
inverse	Logical. Useful for the inverse function of <code>interleave.VGAM()</code> .
S	Numeric. The number of responses.
skip1, sep	The former is logical; should one skip (or omit) "1" when $S = 1$? The latter is the same argument as paste .
hbw	Numeric. The half-bandwidth, which measures the number of bands emanating from the central diagonal band.

Details

See Yee (2015) for some details about some of these functions.

Value

For `param.names()`, this function returns the parameter names for S responses, i.e., `string` is returned unchanged if $S = 1$, else `paste(string, 1:S, sep = "")`.

For `dimm()`, this function returns the number of elements to be stored for each of the working weight matrices. They are represented as columns in the matrix `wz` in e.g., `vglm.fit()`. See the *matrix-band* format described in Section 18.3.5 of Yee (2015).

For `interleave.VGAM()`, this function returns a reordering of the linear/additive predictors depending on the number of responses. The arguments presented in Table 18.5 may not be valid in your version of Yee (2015).

Author(s)

T. W. Yee. Victor Miranda added the `inverse` argument to `interleave.VGAM()`.

References

Yee, T. W. (2015). *Vector Generalized Linear and Additive Models: With an Implementation in R*. New York, USA: Springer.

See Also

[CommonVGAMffArguments](#), [VGAM-package](#).

Examples

```
param.names("shape", 1) # "shape"
param.names("shape", 3) # c("shape1", "shape2", "shape3")

dimm(3, hbw = 1) # Diagonal matrix; the 3 elements need storage.
dimm(3) # A general 3 x 3 symmetric matrix has 6 unique elements.
dimm(3, hbw = 2) # Tridiagonal matrix; the 3-3 element is 0 and unneeded.

M1 <- 2; ncoly <- 3; M <- ncoly * M1
mynames1 <- param.names("location", ncoly)
mynames2 <- param.names("scale", ncoly)
(parameters.names <- c(mynames1, mynames2)[interleave.VGAM(M, M1 = M1)])
# The following is/was in Yee (2015) and has a poor/deceptive style:
(parameters.names <- c(mynames1, mynames2)[interleave.VGAM(M, M = M1)])
parameters.names[interleave.VGAM(M, M1 = M1, inverse = TRUE)]
```

V1

V1 Flying-Bombs Hits in London

Description

A small count data set. During WWII V1 flying-bombs were fired from sites in France (Pas-de-Calais) and Dutch coasts towards London. The number of hits per square grid around London were recorded.

Usage

`data(V1)`

Format

A data frame with the following variables.

hits Values between 0 and 4, and 7. Actually, the 7 is really imputed from the paper (it was recorded as "5 and over").

ofreq Observed frequency, i.e., the number of grids with that many hits.

Details

The data concerns 576 square grids each of 0.25 square kms about south London. The area was selected comprising 144 square kms over which the basic probability function of the distribution was very nearly constant. V1s, which were one type of flying-bomb, were a "Vergeltungswaffen" or vengeance weapon fired during the summer of 1944 at London. The V1s were informally called Buzz Bombs or Doodlebugs, and they were pulse-jet-powered with a warhead of 850 kg of explosives. Over 9500 were launched at London, and many were shot down by artillery and the RAF. Over the period considered the total number of bombs within the area was 537.

It was asserted that the bombs tended to be grouped in clusters. However, a basic Poisson analysis shows this is not the case. Their guidance system being rather primitive, the data is consistent with a Poisson distribution (random).

Compared to Clarke (1946), the more modern analysis of Shaw and Shaw (2019). shows a higher density of hits in south London, hence the distribution is not really uniform over the entire region.

Source

Clarke, R. D. (1946). An application of the Poisson distribution. *Journal of the Institute of Actuaries*, **72**(3), 481.

References

Shaw, L. P. and Shaw, L. F. (2019). The flying bomb and the actuary. *Significance*, **16**(5): 12–17.

See Also

[V2](#), [poissonff](#).

Examples

```
V1
mean(with(V1, rep(hits, times = ofreq)))
var(with(V1, rep(hits, times = ofreq)))
sum(with(V1, rep(hits, times = ofreq)))
## Not run: barplot(with(V1, ofreq),
                    names.arg = as.character(with(V1, hits)),
                    main = "London V1 buzz bomb hits",
                    col = "lightblue", las = 1,
                    ylab = "Frequency", xlab = "Hits")
## End(Not run)
```

V2

V2 Missile Hits in London

Description

A small count data set. During WWII V2 missiles were fired from the continent mainly towards London. The number of hits per square grid around London were recorded.

Usage

```
data(V2)
```

Format

A data frame with the following variables.

hits Values between 0 and 3.

ofreq Observed frequency, i.e., the number of grids with that many hits.

Details

The data concerns 408 square grids each of 0.25 square kms about south London (south of the River Thames). They were picked in a rectangular region of 102 square kilometres where the density of hits were roughly uniformly distributed. The data is somewhat comparable to [V1](#) albeit is a smaller data set.

Source

Shaw, L. P. and Shaw, L. F. (2019). The flying bomb and the actuary. *Significance*, **16**(5): 12–17.

See Also

[V1](#), [poissonff](#).

Examples

```
V2
mean(with(V2, rep(hits, times = ofreq)))
var(with(V2, rep(hits, times = ofreq)))
sum(with(V2, rep(hits, times = ofreq)))
## Not run: barplot(with(V2, ofreq),
  names.arg = as.character(with(V2, hits)),
  main = "London V2 rocket hits",
  col = "lightgreen", las = 1,
  ylab = "Frequency", xlab = "Hits")
## End(Not run)
```

vcovvlm	<i>Calculate Variance-Covariance Matrix for a Fitted VLM or RR-VGLM or DRR-VGLM or QRR-VGLM Object</i>
---------	--

Description

Returns the variance-covariance matrix of the parameters of a fitted `vlm-class` or `rrvglm-class` or `drrvglm-class` object.

Usage

```
vcovvlm(object, dispersion = NULL, untransform = FALSE,
  complete = TRUE, ...)
vcovrrvglm(object, ...)
vcovdrrvglm(object, ...)
vcovqrrvglm(object, ...)
```

Arguments

object	A fitted model object, having class <code>vlm-class</code> or <code>rrvglm-class</code> or <code>drrvglm-class</code> or <code>qrrvglm-class</code> or a superclass of such. The former includes a <code>vglm</code> object.
dispersion	Numerical. This argument should not be used as VGAM will be phasing out dispersion parameters. Formerly, a value may be specified, else it is estimated for quasi-GLMs (e.g., method of moments). For almost all other types of VGLMs it is usually unity. The value is multiplied by the raw variance-covariance matrix.
untransform	logical. For intercept-only models with trivial constraints; if set TRUE then the parameter link function is inverted to give the answer for the untransformed/raw parameter.
complete	An argument currently ignored. Added only so that <code>linearHypothesis()</code> in car can be called.
...	Same as <code>vcov</code> .

Details

This methods function is based on the QR decomposition of the (large) VLM model matrix and working weight matrices. Currently `vcovvlm` operates on the fundamental `vlm-class` objects because pretty well all modelling functions in **VGAM** inherit from this. Currently `vcovrrvglm` is not entirely reliable because the elements of the **A–C** part of the matrix sometimes cannot be computed very accurately, so that the entire matrix is not positive-definite.

For "qrrvglm" objects, `vcovqrrvglm` is currently working with Rank = 1 objects or when `I.tolerances = TRUE`. Then the answer is conditional given **C**. The code is based on `model.matrixqrrvglm` so that the dimnames are the same.

Value

Same as `vcov`.

Note

For some models inflated standard errors can occur, such as parameter estimates near the boundary of the parameter space. Detection for this is available for some models using `hdeff.vglm`, which tests for an Hauck-Donner effect (HDE) for each regression coefficient. If the HDE is present, using `lrt.stat.vlm` should return more accurate p-values.

Author(s)

Thomas W. Yee

See Also

`confintvglm`, `summaryvglm`, `vcov`, `hdeff.vglm`, `lrt.stat.vlm`, `model.matrixqrrvglm`.

Examples

```
## Not run:
ndata <- data.frame(x2 = runif(nn <- 300))
ndata <- transform(ndata, y1 = rnbino(m, mu = exp(3+x2), exp(1)),
                  y2 = rnbino(m, mu = exp(2-x2), exp(0)))
fit1 <- vglm(cbind(y1, y2) ~ x2, negbinomial, ndata, trace = TRUE)
fit2 <- rrvglm(y1 ~ x2, negbinomial(zero = NULL), data = ndata)
coef(fit1, matrix = TRUE)
vcov(fit1)
vcov(fit2)

## End(Not run)
```

 venice

Venice Maximum Sea Levels Data

Description

Some sea levels data sets recorded at Venice, Italy.

Usage

```
data(venice)
data(venice90)
```

Format

`venice` is a data frame with 51 observations on the following 11 variables. It concerns the maximum heights of sea levels between 1931 and 1981.

year a numeric vector.

r1,r2,r3,r4,r5,r6,r7,r8,r9,r10 numeric vectors; r1 is the highest recorded value, r2 is the second highest recorded value, etc.

`venice90` is a data frame with 455 observations on the following 7 variables.

year, month, day, hour numeric vectors; actual time of the recording.

sealevel numeric; sea level.

ohour numeric; number of hours since the midnight of 31 Dec 1939 and 1 Jan 1940.

Year numeric vector; approximate year as a real number. The formula is `start.year + ohour / (365.26 * 24)` where `start.year` is 1940. One can treat `Year` as continuous whereas `year` can be treated as both continuous and discrete.

Details

Sea levels are in cm. For `venice90`, the value 0 corresponds to a fixed reference point (e.g., the mean sea level in 1897 at an old palace of Venice). Clearly since the relative (perceived) mean sea level has been increasing in trend over time (more than an overall 0.4 m increase by 2010), therefore the value 0 is (now) a very low and unusual measurement.

For `venice`, in 1935 only the top six values were recorded.

For `venice90`, this is a subset of a data set provided by Paolo Pirazzoli consisting of hourly sea levels from 1940 to 2009. Values greater than 90 cm were extracted, and then declustered (each cluster provides no more than one value, and each value is at least 24 hours apart). Thus the values are more likely to be independent. Of the original $(2009-1940+1)*365.26*24$ values about 7 percent of these comprise `venice90`.

Yet to do: check for consistency between the data sets. Some external data sets elsewhere have some extremes recorded at times not exactly on the hour.

Source

Pirazzoli, P. (1982) Maree estreme a Venezia (periodo 1872–1981). *Acqua Aria*, **10**, 1023–1039.
Thanks to Paolo Pirazzoli and Alberto Tomasin for the `venice90` data.

References

Smith, R. L. (1986). Extreme value theory based on the r largest annual events. *Journal of Hydrology*, **86**, 27–43.
Battistin, D. and Canestrelli, P. (2006). *La serie storica delle maree a Venezia, 1872–2004* (in Italian), Comune di Venezia. Istituzione Centro Previsione e Segnalazioni Maree.

See Also

[guplot](#), [gev](#), [gpd](#).

Examples

```
## Not run:
matplot(venice[["year"]], venice[, -1], xlab = "Year",
        ylab = "Sea level (cm)", type = "l")

ymat <- as.matrix(venice[, paste("r", 1:10, sep = "")])
fit1 <- vgam(ymat ~ s(year, df = 3), gumbel(R = 365, mpv = TRUE),
           venice, trace = TRUE, na.action = na.pass)
head(fitted(fit1))

par(mfrow = c(2, 1), xpd = TRUE)
plot(fit1, se = TRUE, lcol = "blue", llwd = 2, slty = "dashed")

par(mfrow = c(1,1), bty = "l", xpd = TRUE, las = 1)
qtplot(fit1, mpv = TRUE, lcol = c(1, 2, 5), tcol = c(1, 2, 5),
       llwd = 2, pcol = "blue", tadj = 0.1)

plot(sealevel ~ Year, data = venice90, type = "h", col = "blue")
summary(venice90)
dim(venice90)
round(100 * nrow(venice90)/((2009-1940+1)*365.26*24), dig = 3)

## End(Not run)
```

Description

Fit a vector generalized additive model (VGAM). Both 1st-generation VGAMs (based on backfitting) and 2nd-generation VGAMs (based on P-splines, with automatic smoothing parameter selection) are implemented. This is a large class of models that includes generalized additive models (GAMs) and vector generalized linear models (VGLMs) as special cases.

Usage

```
vgam(formula,
      family = stop("argument 'family' needs to be assigned"),
      data = list(), weights = NULL, subset = NULL,
      na.action, etastart = NULL, mustart = NULL,
      coefstart = NULL, control = vgam.control(...),
      offset = NULL, method = "vgam.fit", model = FALSE,
      x.arg = TRUE, y.arg = TRUE, contrasts = NULL,
      constraints = NULL, extra = list(), form2 = NULL,
      qr.arg = FALSE, smart = TRUE, ...)
```

Arguments

- | | |
|--|---|
| formula | a symbolic description of the model to be fit. The RHS of the formula is applied to each linear/additive predictor, and should include at least one sm.os term or sm.ps term or s term. Mixing both together is not allowed. Different variables in each linear/additive predictor can be chosen by specifying constraint matrices. |
| family | Same as for vglm . |
| data | an optional data frame containing the variables in the model. By default the variables are taken from <code>environment(formula)</code> , typically the environment from which <code>vgam</code> is called. |
| weights, subset, na.action | Same as for vglm . Note that <code>subset</code> may be unreliable and to get around this problem it is best to use subset to create a new smaller data frame and feed in the smaller data frame. See below for an example. This is a bug that needs fixing. |
| etastart, mustart, coefstart | Same as for vglm . |
| control | a list of parameters for controlling the fitting process. See vgam.control for details. |
| method | the method to be used in fitting the model. The default (and presently only) method <code>vgam.fit</code> uses iteratively reweighted least squares (IRLS). |
| constraints, model, offset | Same as for vglm . |
| x.arg, y.arg | logical values indicating whether the model matrix and response vector/matrix used in the fitting process should be assigned in the <code>x</code> and <code>y</code> slots. Note the model matrix is the LM model matrix; to get the VGAM model matrix type <code>model.matrix(vgamfit)</code> where <code>vgamfit</code> is a <code>vgam</code> object. |
| contrasts, extra, form2, qr.arg, smart | Same as for vglm . |
| ... | further arguments passed into vgam.control . |

Details

A vector generalized additive model (VGAM) is loosely defined as a statistical model that is a function of M additive predictors. The central formula is given by

$$\eta_j = \sum_{k=1}^p f_{(j)k}(x_k)$$

where x_k is the k th explanatory variable (almost always $x_1 = 1$ for the intercept term), and $f_{(j)k}$ are smooth functions of x_k that are estimated by smoothers. The first term in the summation is just the intercept. Currently two types of smoothers are implemented: `s` represents the older and more traditional one, called a *vector (cubic smoothing spline) smoother* and is based on Yee and Wild (1996); it is more similar to the R package `gam`. The newer one is represented by `sm.os` and `sm.ps`, and these are based on O-splines and P-splines—they allow automatic smoothing parameter selection; it is more similar to the R package `mgcv`.

In the above, $j = 1, \dots, M$ where M is finite. If all the functions are constrained to be linear then the resulting model is a vector generalized linear model (VGLM). VGLMs are best fitted with `vglm`.

Vector (cubic smoothing spline) smoothers are represented by `s()` (see `s`). Local regression via `lo()` is *not* supported. The results of `vgam` will differ from the `gam()` (in the `gam`) because `vgam()` uses a different knot selection algorithm. In general, fewer knots are chosen because the computation becomes expensive when the number of additive predictors M is large.

Second-generation VGAMs are based on the O-splines and P-splines. The latter is due to Eilers and Marx (1996). Backfitting is not required, and estimation is performed using IRLS. The function `sm.os` represents a *smart* implementation of O-splines. The function `sm.ps` represents a *smart* implementation of P-splines. Written G2-VGAMs or P-VGAMs, this methodology should not be used unless the sample size is reasonably large. Usually an UBRE predictive criterion is optimized (at each IRLS iteration) because the scale parameter for VGAMs is usually assumed to be known. This search for optimal smoothing parameters does not always converge, and neither is it totally reliable. G2-VGAMs implicitly set `criterion = "coefficients"` so that convergence occurs when the change in the regression coefficients between 2 IRLS iterations is sufficiently small. Otherwise the search for the optimal smoothing parameters might cause the log-likelihood to decrease between 2 IRLS iterations. Currently *outer iteration* is implemented, by default, rather than *performance iteration* because the latter is more easy to converge to a local solution; see Wood (2004) for details. One can use *performance iteration* by setting `Maxit.outer = 1` in `vgam.control`.

The underlying algorithm of VGAMs is IRLS. First-generation VGAMs (called G1-VGAMs) are estimated by modified vector backfitting using vector splines. O-splines are used as the basis functions for the vector (smoothing) splines, which are a lower dimensional version of natural B-splines. The function `vgam.fit()` actually does the work. The smoothing code is based on F. O'Sullivan's BART code.

A closely related methodology based on VGAMs called *constrained additive ordination* (CAO) first forms a linear combination of the explanatory variables (called *latent variables*) and then fits a GAM to these. This is implemented in the function `cao` for a very limited choice of family functions.

Value

For G1-VGAMs and G2-VGAMs, an object of class "vgam" or "pvgam" respectively (see `vgam-class` and `pvgam-class` for further information).

WARNING

For G1-VGAMs, currently `vgam` can only handle constraint matrices `cmat`, say, such that `crossprod(cmat)` is diagonal. It can be detected by `is.buggy`. VGAMs with constraint matrices that have non-orthogonal columns should be fitted with `sm.os` or `sm.ps` terms instead of `s`.

See warnings in `vglm.control`.

Note

This function can fit a wide variety of statistical models. Some of these are harder to fit than others because of inherent numerical difficulties associated with some of them. Successful model fitting benefits from cumulative experience. Varying the values of arguments in the **VGAM** family function itself is a good first step if difficulties arise, especially if initial values can be inputted. A second, more general step, is to vary the values of arguments in `vgam.control`. A third step is to make use of arguments such as `etastart`, `coefstart` and `mustart`.

Some **VGAM** family functions end in "ff" to avoid interference with other functions, e.g., `binomialff`, `poissonff`. This is because **VGAM** family functions are incompatible with `glm` (and also `gam()` in `gam` and `gam` in `mgcv`).

The smart prediction (`smartpred`) library is packed with the **VGAM** library.

The theory behind the scaling parameter is currently being made more rigorous, but it should give the same value as the scale parameter for GLMs.

Author(s)

Thomas W. Yee

References

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- Yee, T. W. and Wild, C. J. (1996). Vector generalized additive models. *Journal of the Royal Statistical Society, Series B, Methodological*, **58**, 481–493.
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- Yee, T. W. (2015). Vector Generalized Linear and Additive Models: With an Implementation in R. New York, USA: *Springer*.
- Yee, T. W. (2016). Comments on “Smoothing parameter and model selection for general smooth models” by Wood, S. N. and Pya, N. and Safken, N., *J. Amer. Statist. Assoc.*, **110**(516).

See Also

`is.buggy`, `vgam.control`, `vgam-class`, `vglmff-class`, `plotvgam`, `summaryvgam`, `summarypvgam`, `sm.os`, `sm.ps`, `s`, `magic`, `vglm`, `vsmooth.spline`, `cao`.

Examples

```

# Nonparametric proportional odds model
pneumo <- transform(pneumo, let = log(exposure.time))
vgam(cbind(normal, mild, severe) ~ s(let),
      cumulative(parallel = TRUE), data = pneumo, trace = TRUE)

# Nonparametric logistic regression
hfit <- vgam(agaas ~ s(altitude, df = 2), binomialff, hunua)
## Not run: plot(hfit, se = TRUE)
phfit <- predict(hfit, type = "terms", raw = TRUE, se = TRUE)
names(phfit)
head(phfit$fitted)
head(phfit$se.fit)
phfit$df
phfit$sigma

## Not run: # Fit two species simultaneously
hfit2 <- vgam(cbind(agaas, kniexc) ~ s(altitude, df = c(2, 3)),
              binomialff(multiple.responses = TRUE), data = hunua)
coef(hfit2, matrix = TRUE) # Not really interpretable
plot(hfit2, se = TRUE, overlay = TRUE, lcol = 3:4, scol = 3:4)
ooo <- with(hunua, order(altitude))
with(hunua, matplot(altitude[ooo], fitted(hfit2)[ooo,],
                   ylim = c(0, 0.8), las = 1, type = "l", lwd = 2,
                   xlab = "Altitude (m)", ylab = "Probability of presence",
                   main = "Two plant species' response curves"))
with(hunua, rug(altitude))

# The 'subset' argument does not work here. Use subset() instead.
set.seed(1)
zdata <- data.frame(x2 = runif(nn <- 500))
zdata <- transform(zdata, y = rbinom(nn, 1, 0.5))
zdata <- transform(zdata, subS = runif(nn) < 0.7)
sub.zdata <- subset(zdata, subS) # Use this instead
if (FALSE)
  fit4a <- vgam(cbind(y, y) ~ s(x2, df = 2),
                binomialff(multiple.responses = TRUE),
                data = zdata, subset = subS) # This fails!!!
fit4b <- vgam(cbind(y, y) ~ s(x2, df = 2),
              binomialff(multiple.responses = TRUE),
              data = sub.zdata) # This succeeds!!!
fit4c <- vgam(cbind(y, y) ~ sm.os(x2),
              binomialff(multiple.responses = TRUE),
              data = sub.zdata) # This succeeds!!!

par(mfrow = c(2, 2))
plot(fit4b, se = TRUE, shade = TRUE, shcol = "pink")
plot(fit4c, se = TRUE, shade = TRUE, shcol = "pink")

## End(Not run)

```

 vgam-class

 Class "vgam"

Description

Vector generalized additive models.

Objects from the Class

Objects can be created by calls of the form `vgam(...)`.

Slots

`nl.chisq`: Object of class "numeric". Nonlinear chi-squared values.

`nl.df`: Object of class "numeric". Nonlinear chi-squared degrees of freedom values.

`spar`: Object of class "numeric" containing the (scaled) smoothing parameters.

`s.xargument`: Object of class "character" holding the variable name of any `s()` terms.

`var`: Object of class "matrix" holding approximate pointwise standard error information.

`Bspline`: Object of class "list" holding the scaled (internal and boundary) knots, and the fitted B-spline coefficients. These are used for prediction.

`extra`: Object of class "list"; the extra argument on entry to `vglm`. This contains any extra information that might be needed by the family function.

`family`: Object of class "vglmff". The family function.

`iter`: Object of class "numeric". The number of IRLS iterations used.

`predictors`: Object of class "matrix" with M columns which holds the M linear predictors.

`assign`: Object of class "list", from class "vlm". This named list gives information matching the columns and the (LM) model matrix terms.

`call`: Object of class "call", from class "vlm". The matched call.

`coefficients`: Object of class "numeric", from class "vlm". A named vector of coefficients.

`constraints`: Object of class "list", from class "vlm". A named list of constraint matrices used in the fitting.

`contrasts`: Object of class "list", from class "vlm". The contrasts used (if any).

`control`: Object of class "list", from class "vlm". A list of parameters for controlling the fitting process. See [vglm.control](#) for details.

`criterion`: Object of class "list", from class "vlm". List of convergence criterion evaluated at the final IRLS iteration.

`df.residual`: Object of class "numeric", from class "vlm". The residual degrees of freedom.

`df.total`: Object of class "numeric", from class "vlm". The total degrees of freedom.

`dispersion`: Object of class "numeric", from class "vlm". The scaling parameter.

`effects`: Object of class "numeric", from class "vlm". The effects.

fitted.values: Object of class "matrix", from class "vlm". The fitted values. This is usually the mean but may be quantiles, or the location parameter, e.g., in the Cauchy model.

misc: Object of class "list", from class "vlm". A named list to hold miscellaneous parameters.

model: Object of class "data.frame", from class "vlm". The model frame.

na.action: Object of class "list", from class "vlm". A list holding information about missing values.

offset: Object of class "matrix", from class "vlm". If non-zero, a M -column matrix of offsets.

post: Object of class "list", from class "vlm" where post-analysis results may be put.

preplot: Object of class "list", from class "vlm" used by `plotvgam`; the plotting parameters may be put here.

prior.weights: Object of class "matrix", from class "vlm" holding the initially supplied weights.

qr: Object of class "list", from class "vlm". QR decomposition at the final iteration.

R: Object of class "matrix", from class "vlm". The **R** matrix in the QR decomposition used in the fitting.

rank: Object of class "integer", from class "vlm". Numerical rank of the fitted model.

residuals: Object of class "matrix", from class "vlm". The *working* residuals at the final IRLS iteration.

ResSS: Object of class "numeric", from class "vlm". Residual sum of squares at the final IRLS iteration with the adjusted dependent vectors and weight matrices.

smart.prediction: Object of class "list", from class "vlm". A list of data-dependent parameters (if any) that are used by smart prediction.

terms: Object of class "list", from class "vlm". The `terms` object used.

weights: Object of class "matrix", from class "vlm". The weight matrices at the final IRLS iteration. This is in matrix-band form.

x: Object of class "matrix", from class "vlm". The model matrix (LM, not VGLM).

xlevels: Object of class "list", from class "vlm". The levels of the factors, if any, used in fitting.

y: Object of class "matrix", from class "vlm". The response, in matrix form.

Xm2: Object of class "matrix", from class "vlm". See `vglm-class`.

Ym2: Object of class "matrix", from class "vlm". See `vglm-class`.

callxm2: Object of class "call", from class "vlm". The matched call for argument `form2`.

Extends

Class "vglm", directly. Class "vlm", by class "vglm".

Methods

cdf signature(object = "vglm"): cumulative distribution function. Useful for quantile regression and extreme value data models.

deplot signature(object = "vglm"): density plot. Useful for quantile regression models.

deviance signature(object = "vglm"): deviance of the model (where applicable).

- plot** signature(x = "vglm"): diagnostic plots.
- predict** signature(object = "vglm"): extract the additive predictors or predict the additive predictors at a new data frame.
- print** signature(x = "vglm"): short summary of the object.
- qtplot** signature(object = "vglm"): quantile plot (only applicable to some models).
- resid** signature(object = "vglm"): residuals. There are various types of these.
- residuals** signature(object = "vglm"): residuals. Shorthand for resid.
- rlplot** signature(object = "vglm"): return level plot. Useful for extreme value data models.
- summary** signature(object = "vglm"): a more detailed summary of the object.

Note

VGAMs have all the slots that `vglm` objects have (`vglm-class`), plus the first few slots described in the section above.

Author(s)

Thomas W. Yee

References

Yee, T. W. and Wild, C. J. (1996). Vector generalized additive models. *Journal of the Royal Statistical Society, Series B, Methodological*, **58**, 481–493.

See Also

[vgam.control](#), [vglm](#), [s](#), [vglm-class](#), [vglmff-class](#).

Examples

```
# Fit a nonparametric proportional odds model
pneumo <- transform(pneumo, let = log(exposure.time))
vgam(cbind(normal, mild, severe) ~ s(let),
      cumulative(parallel = TRUE), data = pneumo)
```

vgam.control

Control Function for vgam()

Description

Algorithmic constants and parameters for running `vgam` are set using this function.

Usage

```
vgam.control(all.knots = FALSE, bf.epsilon = 1e-07, bf.maxit = 30,
             checkwz=TRUE, Check.rank = TRUE, Check.cm.rank = TRUE,
             criterion = names(.min.criterion.VGAM),
             epsilon = 1e-07, maxit = 30, Maxit.outer = 10,
             noWarning = FALSE,
             na.action = na.fail,
             nk = NULL, save.weights = FALSE, se.fit = TRUE,
             trace = FALSE, wzepsilon = .Machine$double.eps^0.75,
             xij = NULL, gamma.arg = 1, ...)
```

Arguments

- | | |
|---------------------------|--|
| all.knots | logical indicating if all distinct points of the smoothing variables are to be used as knots. By default, all.knots=TRUE for $n \leq 40$, and for $n > 40$, the number of knots is approximately $40 + (n - 40)^{0.25}$. This increases very slowly with n so that the number of knots is approximately between 50 and 60 for large n . |
| bf.epsilon | tolerance used by the modified vector backfitting algorithm for testing convergence. Must be a positive number. |
| bf.maxit | maximum number of iterations allowed in the modified vector backfitting algorithm. Must be a positive integer. |
| checkwz | logical indicating whether the diagonal elements of the working weight matrices should be checked whether they are sufficiently positive, i.e., greater than wzepsilon. If not, any values less than wzepsilon are replaced with this value. |
| Check.rank, Check.cm.rank | See vglm.control . |
| criterion | character variable describing what criterion is to be used to test for convergence. The possibilities are listed in <code>.min.criterion.VGAM</code> , but most family functions only implement a few of these. |
| epsilon | positive convergence tolerance epsilon. Roughly speaking, the Newton-Raphson/Fisher-scoring/local-scoring iterations are assumed to have converged when two successive criterion values are within epsilon of each other. |
| maxit | maximum number of Newton-Raphson/Fisher-scoring/local-scoring iterations allowed. |
| Maxit.outer | maximum number of outer iterations allowed when there are sm.os or sm.ps terms. See vgam for a little information about the default <i>outer iteration</i> . Note that one can use <i>performance iteration</i> by setting <code>Maxit.outer = 1</code> ; then the smoothing parameters will be automatically chosen at each IRLS iteration (some specific programming allows this).

Note that gam uses outer iteration by default. However, magic is only invoked for the Gaussian family, so the results of gam may differ substantially from sm.os and sm.ps in general. |
| na.action | how to handle missing values. Unlike the SPLUS <code>gam</code> function, vgam cannot handle NAs when smoothing. |

nk	vector of length d containing positive integers. where d be the number of s terms in the formula. Recycling is used if necessary. The i th value is the number of B-spline coefficients to be estimated for each component function of the i th $s()$ term. nk differs from the number of knots by some constant. If specified, nk overrides the automatic knot selection procedure.
save.weights	logical indicating whether the weights slot of a "vglm" object will be saved on the object. If not, it will be reconstructed when needed, e.g., <code>summary</code> .
se.fit	logical indicating whether approximate pointwise standard errors are to be saved on the object. If TRUE, then these can be plotted with <code>plot(..., se = TRUE)</code> .
trace	logical indicating if output should be produced for each iteration.
wzepsilon	Small positive number used to test whether the diagonals of the working weight matrices are sufficiently positive.
noWarning	Same as <code>vglm.control</code> .
xij	Same as <code>vglm.control</code> .
gamma.arg	Numeric; same as <code>gamma</code> in <code>magic</code> . Inflation factor for optimizing the UBRE/GCV criterion. If given, a suggested value is 1.4 to help avoid overfitting, based on the work of Gu and co-workers (values between 1.2 and 1.4 appeared reasonable, based on simulations). A warning may be given if the value is deemed out-of-range.
...	other parameters that may be picked up from control functions that are specific to the VGAM family function.

Details

Most of the control parameters are used within `vgam.fit` and you will have to look at that to understand the full details. Many of the control parameters are used in a similar manner by `vglm.fit` (`vglm`) because the algorithm (IRLS) is very similar.

Setting `save.weights=FALSE` is useful for some models because the weights slot of the object is often the largest and so less memory is used to store the object. However, for some **VGAM** family function, it is necessary to set `save.weights=TRUE` because the weights slot cannot be reconstructed later.

Value

A list with components matching the input names. A little error checking is done, but not much. The list is assigned to the control slot of `vgam` objects.

Warning

See `vglm.control`.

Note

`vgam` does not implement half-stepsizing, therefore parametric models should be fitted with `vglm`. Also, `vgam` is slower than `vglm` too.

Author(s)

Thomas W. Yee

References

Yee, T. W. and Wild, C. J. (1996). Vector generalized additive models. *Journal of the Royal Statistical Society, Series B, Methodological*, **58**, 481–493.

See Also

[vgam](#), [vglm.control](#), [vsmooth.spline](#), [vglm](#).

Examples

```
pneumo <- transform(pneumo, let = log(exposure.time))
vgam(cbind(normal, mild, severe) ~ s(let, df = 2), multinomial,
     data = pneumo, trace = TRUE, eps = 1e-4, maxit = 10)
```

vglm

*Fitting Vector Generalized Linear Models***Description**

vglm fits vector generalized linear models (VGLMs). This very large class of models includes generalized linear models (GLMs) as a special case.

Usage

```
vglm(formula,
     family = stop("argument 'family' needs to be assigned"),
     data = list(), weights = NULL, subset = NULL,
     na.action, etastart = NULL, mustart = NULL,
     coefstart = NULL, control = vglm.control(...), offset = NULL,
     method = "vglm.fit", model = FALSE, x.arg = TRUE, y.arg = TRUE,
     contrasts = NULL, constraints = NULL, extra = list(),
     form2 = NULL, qr.arg = TRUE, smart = TRUE, ...)
```

Arguments

formula	a symbolic description of the model to be fit. The RHS of the formula is applied to each linear predictor. The effect of different variables in each linear predictor can be controlled by specifying constraint matrices—see constraints below.
family	a function of class "vglmff" (see vglmff-class) describing what statistical model is to be fitted. This is called a “VGAM family function”. See CommonVGAMffArguments for general information about many types of arguments found in this type of function. The argument name “family” is used loosely and for the ease of existing glm users; there is no concept of a formal “error distribution” for VGLMs.

	Possibly the argument name should be better "model" but unfortunately that name has already been taken.
data	an optional data frame containing the variables in the model. By default the variables are taken from <code>environment(formula)</code> , typically the environment from which <code>vglm</code> is called.
weights	an optional vector or matrix of (prior fixed and known) weights to be used in the fitting process. If the VGAM family function handles multiple responses ($Q > 1$ of them, say) then weights can be a matrix with Q columns. Each column matches the respective response. If it is a vector (the usually case) then it is recycled into a matrix with Q columns. The values of weights must be positive; try setting a very small value such as $1.0e-8$ to effectively delete an observation. Currently the <code>weights</code> argument supports sampling weights from complex sampling designs via svyVGAM . Some details can be found at https://CRAN.R-project.org/package=svyVGAM .
subset	an optional logical vector specifying a subset of observations to be used in the fitting process.
na.action	a function which indicates what should happen when the data contain NAs. The default is set by the <code>na.action</code> setting of <code>options</code> , and is <code>na.fail</code> if that is unset. The "factory-fresh" default is <code>na.omit</code> which is known as <i>complete case analysis</i> and applied to both sides of the formula.
etastart	optional starting values for the linear predictors. It is a M -column matrix with the same number of rows as the response. If $M = 1$ then it may be a vector. Note that <code>etastart</code> and the output of <code>predict(fit)</code> should be comparable. Here, <code>fit</code> is the fitted object. Almost all VGAM family functions are self-starting.
mustart	optional starting values for the fitted values. It can be a vector or a matrix; if a matrix, then it has the same number of rows as the response. Usually <code>mustart</code> and the output of <code>fitted(fit)</code> should be comparable. Most family functions do not make use of this argument because it is not possible to compute all M columns of η from μ .
coefstart	optional starting values for the coefficient vector. The length and order must match that of <code>coef(fit)</code> .
control	a list of parameters for controlling the fitting process. See <code>vglm.control</code> for details.
offset	a vector or M -column matrix of offset values. These are <i>a priori</i> known and are added to the linear/additive predictors during fitting.
method	the method to be used in fitting the model. The default (and presently only) method <code>vglm.fit()</code> uses iteratively reweighted least squares (IRLS).
model	a logical value indicating whether the <i>model frame</i> should be assigned in the <code>model</code> slot.
x.arg, y.arg	logical values indicating whether the LM matrix and response vector/matrix used in the fitting process should be assigned in the <code>x</code> and <code>y</code> slots. Note that the model matrix is the LM matrix; to get the VGLM matrix type <code>model.matrix(vglmfit)</code> where <code>vglmfit</code> is a <code>vglm</code> object.
contrasts	an optional list. See the <code>contrasts.arg</code> of <code>model.matrix.default</code> .

constraints	<p>an optional <code>list</code> of constraint matrices. The components of the list must be named (labelled) with the term it corresponds to (and it must match in character format <i>exactly</i>—see below). There are two types of input: "lm"-type and "vlm"-type. The former is a subset of the latter. The former has a matrix for each term of the LM matrix. The latter has a matrix for each column of the big VLM matrix. After fitting, the <code>constraints</code> extractor function may be applied; it returns the "vlm"-type list of constraint matrices by default. If "lm"-type are returned by <code>constraints</code> then these can be fed into this argument and it should give the same model as before.</p> <p>If the <code>constraints</code> argument is used then the family function's zero argument (if it exists) needs to be set to <code>NULL</code>. This avoids what could be a probable contradiction. Sometimes setting other arguments related to constraint matrices to <code>FALSE</code> is also a good idea, e.g., <code>parallel = FALSE</code>, <code>exchangeable = FALSE</code>.</p> <p>Properties: each constraint matrix must have M rows, and be of full-column rank. By default, constraint matrices are the M by M identity matrix unless arguments in the family function itself override these values, e.g., <code>parallel</code> (see <code>CommonVGAMffArguments</code>). If <code>constraints</code> is used then it must contain <i>all</i> the terms; an incomplete list is not accepted.</p> <p>As mentioned above, the labelling of each constraint matrix must match exactly, e.g., <code>list("s(x2, df=3)"=diag(2))</code> will fail as <code>as.character(~ s(x2, df=3))</code> produces white spaces: "s(x2, df = 3)". Thus <code>list("s(x2, df = 3)" = diag(2))</code> is needed. See Example 6 below. More details are given in Yee (2015; Section 3.3.1.3) which is on p.101. Note that the label for the intercept is "(Intercept)".</p>
extra	an optional list with any extra information that might be needed by the VGAM family function.
form2	the second (optional) formula. If argument <code>xij</code> is used (see <code>vglm.control</code>) then <code>form2</code> needs to have <i>all</i> terms in the model. Also, some VGAM family functions such as <code>micmen</code> use this argument to input the regressor variable. If given, the slots <code>@Xm2</code> and <code>@Ym2</code> may be assigned. Note that smart prediction applies to terms in <code>form2</code> too.
qr.arg	logical value indicating whether the slot <code>qr</code> , which returns the QR decomposition of the VLM model matrix, is returned on the object.
smart	logical value indicating whether smart prediction (<code>smartpred</code>) will be used.
...	further arguments passed into <code>vglm.control</code> .

Details

A vector generalized linear model (VGLM) is loosely defined as a statistical model that is a function of M linear predictors and can be estimated by Fisher scoring. The central formula is given by

$$\eta_j = \beta_j^T x$$

where x is a vector of explanatory variables (sometimes just a 1 for an intercept), and β_j is a vector of regression coefficients to be estimated. Here, $j = 1, \dots, M$, where M is finite. Then one can write $\eta = (\eta_1, \dots, \eta_M)^T$ as a vector of linear predictors.

Most users will find `vglm` similar in flavour to `glm`. The function `vglm.fit` actually does the work.

Value

An object of class "vglm", which has the following slots. Some of these may not be assigned to save space, and will be recreated if necessary later.

extra	the list extra at the end of fitting.
family	the family function (of class "vglmff").
iter	the number of IRLS iterations used.
predictors	a M -column matrix of linear predictors.
assign	a named list which matches the columns and the (LM) model matrix terms.
call	the matched call.
coefficients	a named vector of coefficients.
constraints	a named list of constraint matrices used in the fitting.
contrasts	the contrasts used (if any).
control	list of control parameter used in the fitting.
criterion	list of convergence criterion evaluated at the final IRLS iteration.
df.residual	the residual degrees of freedom.
df.total	the total degrees of freedom.
dispersion	the scaling parameter.
effects	the effects.
fitted.values	the fitted values, as a matrix. This is often the mean but may be quantiles, or the location parameter, e.g., in the Cauchy model.
misc	a list to hold miscellaneous parameters.
model	the model frame.
na.action	a list holding information about missing values.
offset	if non-zero, a M -column matrix of offsets.
post	a list where post-analysis results may be put.
preplot	used by <code>plotvglm</code> , the plotting parameters may be put here.
prior.weights	initially supplied weights (the <code>weights</code> argument). Also see <code>weightsvglm</code> .
qr	the QR decomposition used in the fitting.
R	the R matrix in the QR decomposition used in the fitting.
rank	numerical rank of the fitted model.
residuals	the <i>working</i> residuals at the final IRLS iteration.
ResSS	residual sum of squares at the final IRLS iteration with the adjusted dependent vectors and weight matrices.
smart.prediction	a list of data-dependent parameters (if any) that are used by smart prediction.
terms	the <code>terms</code> object used.
weights	the working weight matrices at the final IRLS iteration. This is in matrix-band form.

x	the model matrix (linear model LM, not VGLM).
xlevels	the levels of the factors, if any, used in fitting.
y	the response, in matrix form.

This slot information is repeated at [vglm-class](#).

WARNING

See warnings in [vglm.control](#). Also, see warnings under `weights` above regarding sampling weights from complex sampling designs.

Note

This function can fit a wide variety of statistical models. Some of these are harder to fit than others because of inherent numerical difficulties associated with some of them. Successful model fitting benefits from cumulative experience. Varying the values of arguments in the **VGAM** family function itself is a good first step if difficulties arise, especially if initial values can be inputted. A second, more general step, is to vary the values of arguments in [vglm.control](#). A third step is to make use of arguments such as `etastart`, `coefstart` and `mustart`.

Some **VGAM** family functions end in "ff" to avoid interference with other functions, e.g., [binomialff](#), [poissonff](#). This is because **VGAM** family functions are incompatible with `glm` (and also `gam()` in `gam` and `gam` in the `mgcv` library).

The smart prediction ([smartpred](#)) library is incorporated within the **VGAM** library.

The theory behind the scaling parameter is currently being made more rigorous, but it should give the same value as the scale parameter for GLMs.

In Example 5 below, the `xij` argument to illustrate covariates that are specific to a linear predictor. Here, `lop/rop` are the ocular pressures of the left/right eye (artificial data). Variables `leye` and `reye` might be the presence/absence of a particular disease on the LHS/RHS eye respectively. See [vglm.control](#) and [fill1](#) for more details and examples.

Author(s)

Thomas W. Yee

References

- Yee, T. W. (2015). *Vector Generalized Linear and Additive Models: With an Implementation in R*. New York, USA: *Springer*.
- Yee, T. W. and Hastie, T. J. (2003). Reduced-rank vector generalized linear models. *Statistical Modelling*, **3**, 15–41.
- Yee, T. W. and Wild, C. J. (1996). Vector generalized additive models. *Journal of the Royal Statistical Society, Series B, Methodological*, **58**, 481–493.
- Yee, T. W. (2014). Reduced-rank vector generalized linear models with two linear predictors. *Computational Statistics and Data Analysis*, **71**, 889–902.
- Yee, T. W. (2008). The VGAM Package. *R News*, **8**, 28–39.

See Also

[vglm.control](#), [vglm-class](#), [vglmff-class](#), [smartpred](#), [vglm.fit](#), [fill1](#), [rrvglm](#), [vgam](#). Methods functions include [add1.vglm](#), [anova.vglm](#), [AICv1m](#), [coefv1m](#), [confintvglm](#), [constraints.v1m](#), [drop1.vglm](#), [fittedv1m](#), [hatvaluesv1m](#), [hdeff.vglm](#), [Influence.vglm](#), [linkfunv1m](#), [lrt.stat.v1m](#), [score.stat.v1m](#), [wald.stat.v1m](#), [nobs.v1m](#), [npred.v1m](#), [plotvglm](#), [predictvglm](#), [residualsvglm](#), [step4vglm](#), [summaryvglm](#), [lrtest_vglm](#), [update](#), [TypicalVGAMfamilyFunction](#), etc.

Examples

```
# Example 1. See help(glm)
(d.AD <- data.frame(treatment = gl(3, 3),
                   outcome = gl(3, 1, 9),
                   counts = c(18,17,15,20,10,20,25,13,12)))
vglm.D93 <- vglm(counts ~ outcome + treatment, poissonff,
                data = d.AD, trace = TRUE)
summary(vglm.D93)

# Example 2. Multinomial logit model
pneumo <- transform(pneumo, let = log(exposure.time))
vglm(cbind(normal, mild, severe) ~ let, multinomial, pneumo)

# Example 3. Proportional odds model
fit3 <- vglm(cbind(normal, mild, severe) ~ let, propodds, pneumo)
coef(fit3, matrix = TRUE)
constraints(fit3)
model.matrix(fit3, type = "lm") # LM model matrix
model.matrix(fit3)             # Larger VGLM (or VLM) matrix

# Example 4. Bivariate logistic model
fit4 <- vglm(cbind(nBnW, nBW, BnW, BW) ~ age, binom2.or, coalminers)
coef(fit4, matrix = TRUE)
deivar(fit4) # Response are proportions
weights(fit4, type = "prior")

# Example 5. The use of the xij argument (simple case).
# The constraint matrix for 'op' has one column.
nn <- 1000
eyesdat <- round(data.frame(lop = runif(nn),
                           rop = runif(nn),
                           op = runif(nn)), digits = 2)
eyesdat <- transform(eyesdat, eta1 = -1 + 2 * lop,
                    eta2 = -1 + 2 * rop)
eyesdat <- transform(eyesdat,
                    leye = rbinom(nn, 1, prob = logitlink(eta1, inv = TRUE)),
                    reye = rbinom(nn, 1, prob = logitlink(eta2, inv = TRUE)))
head(eyesdat)
fit5 <- vglm(cbind(leye, reye) ~ op,
            binom2.or(exchangeable = TRUE, zero = 3),
```

```

      data = eyesdat, trace = TRUE,
      xij = list(op ~ lop + rop + fill1(lop)),
      form2 = ~ op + lop + rop + fill1(lop))
coef(fit5)
coef(fit5, matrix = TRUE)
constraints(fit5)
fit5@control$xij
head(model.matrix(fit5))

# Example 6. The use of the 'constraints' argument.
as.character(~ bs(year,df=3)) # Get the white spaces right
clist <- list("(Intercept)" = diag(3),
              "bs(year, df = 3)" = rbind(1, 0, 0))
fit1 <- vglm(r1 ~ bs(year,df=3), gev(zero = NULL),
             data = venice, constraints = clist, trace = TRUE)
coef(fit1, matrix = TRUE) # Check

```

vglm-class

Class "vglm"

Description

Vector generalized linear models.

Objects from the Class

Objects can be created by calls of the form `vglm(...)`.

Slots

In the following, M is the number of linear predictors.

extra: Object of class "list"; the extra argument on entry to `vglm`. This contains any extra information that might be needed by the family function.

family: Object of class "vglmff". The family function.

iter: Object of class "numeric". The number of IRLS iterations used.

predictors: Object of class "matrix" with M columns which holds the M linear predictors.

assign: Object of class "list", from class "vglm". This named list gives information matching the columns and the (LM) model matrix terms.

call: Object of class "call", from class "vglm". The matched call.

coefficients: Object of class "numeric", from class "vglm". A named vector of coefficients.

constraints: Object of class "list", from class "vglm". A named list of constraint matrices used in the fitting.

contrasts: Object of class "list", from class "vglm". The contrasts used (if any).

control: Object of class "list", from class "vlm". A list of parameters for controlling the fitting process. See [vglm.control](#) for details.

criterion: Object of class "list", from class "vlm". List of convergence criterion evaluated at the final IRLS iteration.

df.residual: Object of class "numeric", from class "vlm". The residual degrees of freedom.

df.total: Object of class "numeric", from class "vlm". The total degrees of freedom.

dispersion: Object of class "numeric", from class "vlm". The scaling parameter.

effects: Object of class "numeric", from class "vlm". The effects.

fitted.values: Object of class "matrix", from class "vlm". The fitted values.

misc: Object of class "list", from class "vlm". A named list to hold miscellaneous parameters.

model: Object of class "data.frame", from class "vlm". The model frame.

na.action: Object of class "list", from class "vlm". A list holding information about missing values.

offset: Object of class "matrix", from class "vlm". If non-zero, a M -column matrix of offsets.

post: Object of class "list", from class "vlm" where post-analysis results may be put.

preplot: Object of class "list", from class "vlm" used by [plotvgam](#); the plotting parameters may be put here.

prior.weights: Object of class "matrix", from class "vlm" holding the initially supplied weights.

qr: Object of class "list", from class "vlm". QR decomposition at the final iteration.

R: Object of class "matrix", from class "vlm". The **R** matrix in the QR decomposition used in the fitting.

rank: Object of class "integer", from class "vlm". Numerical rank of the fitted model.

residuals: Object of class "matrix", from class "vlm". The *working* residuals at the final IRLS iteration.

ResSS: Object of class "numeric", from class "vlm". Residual sum of squares at the final IRLS iteration with the adjusted dependent vectors and weight matrices.

smart.prediction: Object of class "list", from class "vlm". A list of data-dependent parameters (if any) that are used by smart prediction.

terms: Object of class "list", from class "vlm". The [terms](#) object used.

weights: Object of class "matrix", from class "vlm". The weight matrices at the final IRLS iteration. This is in matrix-band form.

x: Object of class "matrix", from class "vlm". The model matrix (LM, not VGLM).

xlevels: Object of class "list", from class "vlm". The levels of the factors, if any, used in fitting.

y: Object of class "matrix", from class "vlm". The response, in matrix form.

Xm2: Object of class "matrix", from class "vlm". See [vglm-class](#)).

Ym2: Object of class "matrix", from class "vlm". See [vglm-class](#)).

callXm2: Object of class "call", from class "vlm". The matched call for argument form2.

Extends

Class "vlm", directly.

Methods

cdf signature(object = "vglm"): cumulative distribution function. Applicable to, e.g., quantile regression and extreme value data models.

deplot signature(object = "vglm"): Applicable to, e.g., quantile regression.

deviance signature(object = "vglm"): deviance of the model (where applicable).

plot signature(x = "vglm"): diagnostic plots.

predict signature(object = "vglm"): extract the linear predictors or predict the linear predictors at a new data frame.

print signature(x = "vglm"): short summary of the object.

qplot signature(object = "vglm"): quantile plot (only applicable to some models).

resid signature(object = "vglm"): residuals. There are various types of these.

residuals signature(object = "vglm"): residuals. Shorthand for resid.

rlplot signature(object = "vglm"): return level plot. Useful for extreme value data models.

summary signature(object = "vglm"): a more detailed summary of the object.

Author(s)

Thomas W. Yee

References

Yee, T. W. and Hastie, T. J. (2003). Reduced-rank vector generalized linear models. *Statistical Modelling*, **3**, 15–41.

Yee, T. W. and Wild, C. J. (1996). Vector generalized additive models. *Journal of the Royal Statistical Society, Series B, Methodological*, **58**, 481–493.

See Also

[vglm](#), [vglmff-class](#), [vgam-class](#).

Examples

```
# Multinomial logit model
pneumo <- transform(pneumo, let = log(exposure.time))
vglm(cbind(normal, mild, severe) ~ let, multinomial, data = pneumo)
```

vglm.control	<i>Control Function for vglm()</i>
--------------	------------------------------------

Description

Algorithmic constants and parameters for running `vglm` are set using this function.

Usage

```
vglm.control(checkwz = TRUE, Check.rank = TRUE, Check.cm.rank = TRUE,
             criterion = names(.min.criterion.VGAM),
             epsilon = 1e-07, half.stepsizing = TRUE,
             maxit = 30, noWarning = FALSE,
             stepsize = 1, save.weights = FALSE,
             trace = FALSE, wzepsilon = .Machine$double.eps^0.75,
             xij = NULL, ...)
```

Arguments

<code>checkwz</code>	logical indicating whether the diagonal elements of the working weight matrices should be checked whether they are sufficiently positive, i.e., greater than <code>wzepsilon</code> . If not, any values less than <code>wzepsilon</code> are replaced with this value.
<code>Check.rank</code>	logical indicating whether the rank of the VLM matrix should be checked. If this is not of full column rank then the results are not to be trusted. The default is to give an error message if the VLM matrix is not of full column rank.
<code>Check.cm.rank</code>	logical indicating whether the rank of each constraint matrix should be checked. If this is not of full column rank then an error will occur. Under no circumstances should any constraint matrix have a rank less than the number of columns.
<code>criterion</code>	character variable describing what criterion is to be used to test for convergence. The possibilities are listed in <code>.min.criterion.VGAM</code> , but most family functions only implement a few of these.
<code>epsilon</code>	positive convergence tolerance <code>epsilon</code> . Roughly speaking, the Newton-Raphson/Fisher-scoring iterations are assumed to have converged when two successive <code>criterion</code> values are within <code>epsilon</code> of each other.
<code>half.stepsizing</code>	logical indicating if half-stepsizing is allowed. For example, in maximizing a log-likelihood, if the next iteration has a log-likelihood that is less than the current value of the log-likelihood, then a half step will be taken. If the log-likelihood is still less than at the current position, a quarter-step will be taken etc. Eventually a step will be taken so that an improvement is made to the convergence criterion. <code>half.stepsizing</code> is ignored if <code>criterion == "coefficients"</code> .
<code>maxit</code>	maximum number of (usually Fisher-scoring) iterations allowed. Sometimes Newton-Raphson is used.
<code>noWarning</code>	logical indicating whether to suppress a warning if convergence is not obtained within <code>maxit</code> iterations. This is ignored if <code>maxit = 1</code> is set.

stepsize	usual step size to be taken between each Newton-Raphson/Fisher-scoring iteration. It should be a value between 0 and 1, where a value of unity corresponds to an ordinary step. A value of 0.5 means half-steps are taken. Setting a value near zero will cause convergence to be generally slow but may help increase the chances of successful convergence for some family functions.
save.weights	logical indicating whether the weights slot of a "vglm" object will be saved on the object. If not, it will be reconstructed when needed, e.g., summary. Some family functions have save.weights = TRUE and others have save.weights = FALSE in their control functions.
trace	logical indicating if output should be produced for each iteration. Setting trace = TRUE is recommended in general because VGAM fits a very broad variety of models and distributions, and for some of them, convergence is intrinsically more difficult. Monitoring convergence can help check that the solution is reasonable or that a problem has occurred. It may suggest better initial values are needed, the making of invalid assumptions, or that the model is inappropriate for the data, etc.
wzepsilon	small positive number used to test whether the diagonals of the working weight matrices are sufficiently positive.
xij	A list of formulas. Each formula has a RHS giving M terms making up a covariate-dependent term (whose name is the response). That is, it creates a variable that takes on different values for each linear/additive predictor, e.g., the ocular pressure of each eye. The M terms must be unique; use <code>fill1</code> , <code>fill2</code> , <code>fill3</code> , etc. if necessary. Each formula should have a response which is taken as the name of that variable, and the M terms are enumerated in sequential order. Each of the M terms multiply each successive row of the constraint matrix. When <code>xij</code> is used, the use of <code>form2</code> is also required to give every term used by the model. A formula or a list of formulas. The function <code>Select</code> can be used to select variables beginning with the same character string.
...	other parameters that may be picked up from control functions that are specific to the VGAM family function.

Details

Most of the control parameters are used within `vglm.fit` and you will have to look at that to understand the full details.

Setting `save.weights = FALSE` is useful for some models because the `weights` slot of the object is the largest and so less memory is used to store the object. However, for some **VGAM** family function, it is necessary to set `save.weights = TRUE` because the `weights` slot cannot be reconstructed later.

Value

A list with components matching the input names. A little error checking is done, but not much. The list is assigned to the `control` slot of `vglm` objects.

Warning

For some applications the default convergence criterion should be tightened. Setting something like `criterion = "coef"`, `epsilon = 1e-09` is one way to achieve this, and also add `trace = TRUE` to monitor the convergence. Setting `maxit` to some higher number is usually not needed, and needing to do so suggests something is wrong, e.g., an ill-conditioned model, over-fitting or under-fitting.

Note

Reiterating from above, setting `trace = TRUE` is recommended in general.

In Example 2 below there are two covariates that have linear/additive predictor specific values. These are handled using the `xij` argument.

Author(s)

Thomas W. Yee

References

Yee, T. W. and Hastie, T. J. (2003). Reduced-rank vector generalized linear models. *Statistical Modelling*, **3**, 15–41.

See Also

[vglm](#), [TypicalVGAMfamilyFunction](#), [fill1](#). The author's homepage has further documentation about the `xij` argument; see also [Select](#).

Examples

```
# Example 1.
pneumo <- transform(pneumo, let = log(exposure.time))
vglm(cbind(normal, mild, severe) ~ let, multinomial, pneumo,
     crit = "coef", step = 0.5, trace = TRUE, epsil = 1e-8,
     maxit = 40)

# Example 2. The use of the xij argument (simple case).
ymat <- rdiric(n <- 1000, shape = rep(exp(2), len = 4))
mydat <- data.frame(x1 = runif(n), x2 = runif(n), x3 = runif(n),
                  x4 = runif(n),
                  z1 = runif(n), z2 = runif(n), z3 = runif(n),
                  z4 = runif(n))
mydat <- transform(mydat, X = x1, Z = z1)
mydat <- round(mydat, digits = 2)
fit2 <- vglm(ymat ~ X + Z,
            dirichlet(parallel = TRUE), mydat, trace = TRUE,
            xij = list(Z ~ z1 + z2 + z3 + z4,
                      X ~ x1 + x2 + x3 + x4),
            form2 = ~ Z + z1 + z2 + z3 + z4 +
                      X + x1 + x2 + x3 + x4)
head(model.matrix(fit2, type = "lm")) # LM model matrix
head(model.matrix(fit2, type = "vlm")) # Big VLM model matrix
```



```

coef(fit2)
coef(fit2, matrix = TRUE)
max(abs(predict(fit2)-predict(fit2, new = mydat))) # Predicts ok
summary(fit2)
## Not run:
# plotvgam(fit2, se = TRUE, xlab = "x1", which.term = 1) # Bug!
# plotvgam(fit2, se = TRUE, xlab = "z1", which.term = 2) # Bug!
plotvgam(fit2, xlab = "x1") # Correct
plotvgam(fit2, xlab = "z1") # Correct

## End(Not run)

# Example 3. The use of the xij argument (complex case).
set.seed(123)
coalminers <-
  transform(coalminers,
            Age = (age - 42) / 5,
            dum1 = round(runif(nrow(coalminers)), digits = 2),
            dum2 = round(runif(nrow(coalminers)), digits = 2),
            dum3 = round(runif(nrow(coalminers)), digits = 2),
            dumm = round(runif(nrow(coalminers)), digits = 2))
BS <- function(x, ..., df = 3)
  sm.bs(c(x,...), df = df)[1:length(x),,drop = FALSE]
NS <- function(x, ..., df = 3)
  sm.ns(c(x,...), df = df)[1:length(x),,drop = FALSE]

# Equivalently...
BS <- function(x, ..., df = 3)
  head(sm.bs(c(x,...), df = df), length(x), drop = FALSE)
NS <- function(x, ..., df = 3)
  head(sm.ns(c(x,...), df = df), length(x), drop = FALSE)

fit3 <- vglm(cbind(nBnW,nBW,BnW,BW) ~ Age + NS(dum1, dum2),
            fam = binom2.or(exchangeable = TRUE, zero = 3),
            xij = list(NS(dum1, dum2) ~ NS(dum1, dum2) +
                      NS(dum2, dum1) +
                      fill1(NS( dum1))),
            form2 = ~ NS(dum1, dum2) + NS(dum2, dum1) +
                    fill1(NS(dum1)) +
                    dum1 + dum2 + dum3 + Age + age + dumm,
            data = coalminers, trace = TRUE)
head(model.matrix(fit3, type = "lm")) # LM model matrix
head(model.matrix(fit3, type = "vlm")) # Big VLM model matrix
coef(fit3)
coef(fit3, matrix = TRUE)
## Not run:
plotvgam(fit3, se = TRUE, lcol = 2, scol = 4, xlab = "dum1")

## End(Not run)

```

vglmff-class

Class "vglmff"

Description

Family functions for the **VGAM** package

Objects from the Class

Objects can be created by calls of the form `new("vglmff", ...)`.

Slots

In the following, M is the number of linear/additive predictors.

start1: Object of class "expression" to insert code at a special position (the very start) in `vglm.fit` or `vgam.fit`.

blurb: Object of class "character" giving a small description of the model. Important arguments such as parameter link functions can be expressed here.

charfun: Object of class "function" which returns the characteristic function or variance function (usually for some GLMs only). The former uses a dummy variable x . Both use the linear/additive predictors. The function must have arguments `function(x, eta, extra = NULL, varfun = FALSE)`. The `eta` and `extra` arguments are used to obtain the parameter values. If `varfun = TRUE` then the function returns the variance function, else the characteristic function (default). Note that one should check that the `infos` slot has a list component called `charfun` which is `TRUE` before attempting to use this slot. This is an easier way to test that this slot is operable.

constraints: Object of class "expression" which sets up any constraint matrices defined by arguments in the family function. A zero argument is always fed into `cm.zero.vgam`, whereas other constraints are fed into `cm.vgam`.

deviance: Object of class "function" returning the deviance of the model. This slot is optional. If present, the function must have arguments `function(mu, y, w, residuals = FALSE, eta, extra = NULL)`. Deviance residuals are returned if `residuals = TRUE`.

rqresslot: Object of class "function" returning the randomized quantile residuals of the distribution. This slot is optional. If present, the function must have arguments `function(mu, y, w, eta, extra = NULL)`.

fini1: Object of class "expression" to insert code at a special position in `vglm.fit` or `vgam.fit`. This code is evaluated immediately after the fitting.

first: Object of class "expression" to insert code at a special position in `vglm` or `vgam`.

infos: Object of class "function" which returns a list with components such as $M1$. At present only a very few **VGAM** family functions have this feature implemented. Those that do do not require specifying the $M1$ argument when used with `rcim`.

initialize: Object of class "expression" used to perform error checking (especially for the variable y) and obtain starting values for the model. In general, `etastart` or `mustart` are assigned values based on the variables y , x and w .

- linkinv:** Object of class "function" which returns the fitted values, given the linear/additive predictors. The function must have arguments `function(eta, extra = NULL)`.
- last:** Object of class "expression" to insert code at a special position (at the very end) of `vglm.fit()` or `vgam.fit()`. This code is evaluated after the fitting. The list `misc` is often assigned components in this slot, which becomes the `misc` slot on the fitted object.
- linkfun:** Object of class "function" which, given the fitted values, returns the linear/additive predictors. If present, the function must have arguments `function(mu, extra = NULL)`. Most **VGAM** family functions do not have a `linkfun` function. They largely are for classical exponential families, i.e., GLMs.
- loglikelihood:** Object of class "function" returning the log-likelihood of the model. This slot is optional. If present, the function must have arguments `function(mu, y, w, residuals = FALSE, eta, extra = NULL)`. The argument `residuals` can be ignored because log-likelihood residuals aren't defined.
- middle1:** Object of class "expression" to insert code at a special position in `vglm.fit` or `vgam.fit`.
- middle2:** Object of class "expression" to insert code at a special position in `vglm.fit` or `vgam.fit`.
- simslot:** Object of class "function" to allow `simulate` to work.
- hadof:** Object of class "function"; experimental.
- summary.dispersion:** Object of class "logical" indicating whether the general VGLM formula (based on a residual sum of squares) can be used for computing the scaling/dispersion parameter. It is `TRUE` for most models except for nonlinear regression models.
- vfamily:** Object of class "character" giving class information about the family function. Although not developed at this stage, more flexible classes are planned in the future. For example, family functions `sratio`, `cratio`, `cumulative`, and `acat` all operate on categorical data, therefore will have a special class called "VGAMcat", say. Then if `fit` was a `vglm` object, then `coef(fit)` would print out the `vglm` coefficients plus "VGAMcat" information as well.
- deriv:** Object of class "expression" which returns a M -column matrix of first derivatives of the log-likelihood function with respect to the linear/additive predictors, i.e., the score vector. In Yee and Wild (1996) this is the \mathbf{d}_i vector. Thus each row of the matrix returned by this slot is such a vector.
- weight:** Object of class "expression" which returns the second derivatives of the log-likelihood function with respect to the linear/additive predictors. This can be either the observed or expected information matrix, i.e., Newton-Raphson or Fisher-scoring respectively. In Yee and Wild (1996) this is the \mathbf{W}_i matrix. Thus each row of the matrix returned by this slot is such a matrix. Like the `weights` slot of `vglm/vgam`, it is stored in *matrix-band* form, whereby the first M columns of the matrix are the diagonals, followed by the upper-diagonal band, followed by the band above that, etc. In this case, there can be up to $M(M + 1)$ columns, with the last column corresponding to the $(1, M)$ elements of the weight matrices.
- validfitted, validparams:** Functions that test that the fitted values and all parameters are within range. These functions can issue a warning if violations are detected.

Methods

print `signature(x = "vglmff")`: short summary of the family function.

Warning

VGAM family functions are not compatible with `glm`, nor `gam()` (from either **gam** or **mgcv**).

Note

With link functions etc., one must use `substitute` to embed the options into the code. There are two different forms: `eval(substitute(expression({...}), list(...)))` for expressions, and `eval(substitute(function(...) { ... }, list(...)))` for functions.

The extra argument in `linkinv`, `linkfun`, `deviance`, `loglikelihood`, etc. matches with the argument `extra` in `vglm`, `vgam` and `rrvglm`. This allows input to be fed into all slots of a **VGAM** family function.

The expression derivative is evaluated immediately prior to `weight`, so there is provision for re-use of variables etc. Programmers must be careful to choose variable names that do not interfere with `vglm.fit`, `vgam.fit()` etc.

Programmers of **VGAM** family functions are encouraged to keep to previous conventions regarding the naming of arguments, e.g., `link` is the argument for parameter link functions, `zero` for allowing some of the linear/additive predictors to be an intercept term only, etc.

In general, Fisher-scoring is recommended over Newton-Raphson where tractable. Although usually slightly slower in convergence, the weight matrices from using the expected information are positive-definite over a larger parameter space.

Author(s)

Thomas W. Yee

References

Yee, T. W. and Wild, C. J. (1996). Vector generalized additive models. *Journal of the Royal Statistical Society, Series B, Methodological*, **58**, 481–493.

See Also

`vglm`, `vgam`, `rrvglm`, `rcim`.

Examples

```
cratio()  
cratio(link = "clogloglink")  
cratio(link = "clogloglink", reverse = TRUE)
```

vonmises	<i>von Mises Distribution Family Function</i>
----------	---

Description

Estimates the location and scale parameters of the von Mises distribution by maximum likelihood estimation.

Usage

```
vonmises(llocation = extlogitlink(min = 0, max = 2*pi),
         lscale = "loglink", ilocation = NULL, iscale = NULL,
         imethod = 1, zero = NULL)
```

Arguments

llocation, lscale	Parameter link functions applied to the location a parameter and scale parameter k , respectively. See Links for more choices. For k , a log link is the default because the parameter is positive.
ilocation	Initial value for the location a parameter. By default, an initial value is chosen internally using <code>imethod</code> . Assigning a value will override the argument <code>imethod</code> .
iscale	Initial value for the scale k parameter. By default, an initial value is chosen internally using <code>imethod</code> . Assigning a value will override the argument <code>imethod</code> .
imethod	An integer with value 1 or 2 which specifies the initialization method. If failure to converge occurs try the other value, or else specify a value for <code>ilocation</code> and <code>iscale</code> .
zero	An integer-valued vector specifying which linear/additive predictors are modelled as intercepts only. The default is none of them. If used, one can choose one value from the set {1,2}. See CommonVGAMffArguments for more information.

Details

The (two-parameter) von Mises is the most commonly used distribution in practice for circular data. It has a density that can be written as

$$f(y; a, k) = \frac{\exp[k \cos(y - a)]}{2\pi I_0(k)}$$

where $0 \leq y < 2\pi$, $k > 0$ is the scale parameter, a is the location parameter, and $I_0(k)$ is the modified Bessel function of order 0 evaluated at k . The mean of Y (which is the fitted value) is a and the circular variance is $1 - I_1(k)/I_0(k)$ where $I_1(k)$ is the modified Bessel function of order 1. By default, $\eta_1 = \log(a/(2\pi - a))$ and $\eta_2 = \log(k)$ for this family function.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), [rrvglm](#) and [vgam](#).

Warning

Numerically, the von Mises can be difficult to fit because of a log-likelihood having multiple maximums. The user is therefore encouraged to try different starting values, i.e., make use of `ilocat` and `iscale`.

Note

The response and the fitted values are scaled so that $0 \leq y < 2\pi$. The linear/additive predictors are left alone. Fisher scoring is used.

Author(s)

T. W. Yee

References

Forbes, C., Evans, M., Hastings, N. and Peacock, B. (2011). *Statistical Distributions*, Hoboken, NJ, USA: John Wiley and Sons, Fourth edition.

See Also

[Bessel](#), [cardioid](#).

CircStats and **circular** currently have a lot more R functions for circular data than the **VGAM** package.

Examples

```
vdata <- data.frame(x2 = runif(nn <- 1000))
vdata <- transform(vdata,
                  y = rnorm(nn, 2+x2, exp(0.2))) # Bad data!!
fit <- vglm(y ~ x2, vonmises(zero = 2), vdata, trace = TRUE)
coef(fit, matrix = TRUE)
Coef(fit)
with(vdata, range(y)) # Original data
range(depvar(fit))    # Processed data is in [0,2*pi)
```

Description

[plot](#) and [pairs](#) methods for objects of class "profile", but renamed as `vplot` and `vpairs`.

Usage

```
vplot.profile(x, ...)  
vpairs.profile(x, colours = 2:3, ...)
```

Arguments

<code>x</code>	an object inheriting from class "profile".
<code>colours</code>	Colours to be used for the mean curves conditional on x and y respectively.
<code>...</code>	arguments passed to or from other methods.

Details

See [profile.glm](#) for details.

Author(s)

T. W. Yee adapted this function from [profile.glm](#), written originally by D. M. Bates and W. N. Venables. (For S in 1996.)

See Also

[profilevglm](#), [confintvglm](#), [lrt.stat](#), [profile.glm](#), [profile.nls](#).

Examples

```
pneumo <- transform(pneumo, let = log(exposure.time))  
fit1 <- vglm(cbind(normal, mild, severe) ~ let, acat,  
            trace = TRUE, data = pneumo)  
pfit1 <- profile(fit1, trace = FALSE)  
## Not run:  
vplot.profile(pfit1)  
vpairs.profile(pfit1)  
  
## End(Not run)
```

vsmooth.spline

*Vector Cubic Smoothing Spline***Description**

Fits a vector cubic smoothing spline.

Usage

```
vsmooth.spline(x, y, w = NULL, df = rep(5, M), spar = NULL,
               i.constraint = diag(M),
               x.constraint = diag(M),
               constraints = list("(Intercepts)" = i.constraint,
                                 x = x.constraint),
               all.knots = FALSE, var.arg = FALSE, scale.w = TRUE,
               nk = NULL, control.spar = list())
```

Arguments

- | | |
|-----------|--|
| x | A vector, matrix or a list. If a list, the x component is used. If a matrix, the first column is used. x may also be a complex vector, in which case the real part is used, and the imaginary part is used for the response. In this help file, n is the number of unique values of x. |
| y | A vector, matrix or a list. If a list, the y component is used. If a matrix, all but the first column is used. In this help file, M is the number of columns of y if there are no constraints on the functions. |
| w | The weight matrices or the number of observations. If the weight matrices, then this must be a n-row matrix with the elements in matrix-band form (see iam). If a vector, then these are the number of observations. By default, w is the M by M identity matrix, denoted by <code>matrix(1, n, M)</code> . |
| df | Numerical vector containing the degrees of freedom for each component function (smooth). If necessary, the vector is recycled to have length equal to the number of component functions to be estimated (M if there are no constraints), which equals the number of columns of the x-constraint matrix. A value of 2 means a linear fit, and each element of df should lie between 2 and n. The larger the values of df the more wiggly the smooths. |
| spar | Numerical vector containing the non-negative smoothing parameters for each component function (smooth). If necessary, the vector is recycled to have length equal to the number of component functions to be estimated (M if there are no constraints), which equals the number of columns of the x-constraint matrix. A value of zero means the smooth goes through the data and hence is wiggly. A value of Inf may be assigned, meaning the smooth will be linear. By default, the NULL value of spar means df is used to determine the smoothing parameters. |
| all.knots | Logical. If TRUE then each distinct value of x will be a knot. By default, only a subset of the unique values of x are used; typically, the number of knots is $O(n^{0.25})$ for n large, but if $n \leq 40$ then all the unique values of x are used. |

<code>i.constraint</code>	A M-row constraint matrix for the intercepts. It must be of full column rank. By default, the constraint matrix for the intercepts is the M by M identity matrix, meaning no constraints.
<code>x.constraint</code>	A M-row constraint matrix for x. It must be of full column rank. By default, the constraint matrix for the intercepts is the M by M identity matrix, meaning no constraints.
<code>constraints</code>	An alternative to specifying <code>i.constraint</code> and <code>x.constraint</code> , this is a list with two components corresponding to the intercept and x respectively. They must both be a M-row constraint matrix with full column rank.
<code>var.arg</code>	Logical: return the pointwise variances of the fit? Currently, this corresponds only to the nonlinear part of the fit, and may be wrong.
<code>scale.w</code>	Logical. By default, the weights w are scaled so that the diagonal elements have mean 1.
<code>nk</code>	Number of knots. If used, this argument overrides <code>all.knots</code> , and must lie between 6 and <code>n+2</code> inclusive.
<code>control.spar</code>	See smooth.spline .

Details

The algorithm implemented is detailed in Yee (2000). It involves decomposing the component functions into a linear and nonlinear part, and using B-splines. The cost of the computation is $O(nM^3)$.

The argument `spar` contains *scaled* smoothing parameters.

Value

An object of class "vsmooth.spline" (see `vsmooth.spline-class`).

WARNING

See [vgam](#) for information about an important bug.

Note

This function is quite similar to [smooth.spline](#) but offers less functionality. For example, cross validation is not implemented here. For $M = 1$, the results will be generally different, mainly due to the different way the knots are selected.

The vector cubic smoothing spline which `s()` represents is computationally demanding for large M . The cost is approximately $O(nM^3)$ where n is the number of unique abscissae.

Yet to be done: return the *unscaled* smoothing parameters.

Author(s)

Thomas W. Yee

References

Yee, T. W. (2000). Vector Splines and Other Vector Smoothers. Pages 529–534. In: Bethlehem, J. G. and van der Heijde, P. G. M. *Proceedings in Computational Statistics COMPSTAT 2000*. Heidelberg: Physica-Verlag.

See Also

vsmooth.spline-class, plot.vsmooth.spline, predict.vsmooth.spline, iam, [sm.os](#), [s](#), [smooth.spline](#).

Examples

```
nn <- 20; x <- 2 + 5*(nn:1)/nn
x[2:4] <- x[5:7] # Allow duplication
y1 <- sin(x) + rnorm(nn, sd = 0.13)
y2 <- cos(x) + rnorm(nn, sd = 0.13)
y3 <- 1 + sin(x) + rnorm(nn, sd = 0.13) # For constraints
y <- cbind(y1, y2, y3)
ww <- cbind(rep(3, nn), 4, (1:nn)/nn)

(fit <- vsmooth.spline(x, y, w = ww, df = 5))
## Not run:
plot(fit) # The 1st & 3rd functions dont differ by a constant

## End(Not run)

mat <- matrix(c(1,0,1, 0,1,0), 3, 2)
(fit2 <- vsmooth.spline(x, y, w = ww, df = 5, i.constr = mat,
                       x.constr = mat))
# The 1st and 3rd functions do differ by a constant:
mycols <- c("orange", "blue", "orange")
## Not run: plot(fit2, lcol = mycols, pcol = mycols, las = 1)

p <- predict(fit, x = model.matrix(fit, type = "lm"), deriv = 0)
max(abs(depvar(fit) - with(p, y))) # Should be 0

par(mfrow = c(3, 1))
ux <- seq(1, 8, len = 100)
for (dd in 1:3) {
  pp <- predict(fit, x = ux, deriv = dd)
  ## Not run:
  with(pp, matplot(x, y, type = "l", main = paste("deriv =", dd),
                  lwd = 2, ylab = "", cex.axis = 1.5,
                  cex.lab = 1.5, cex.main = 1.5))
  ## End(Not run)
}
```

Description

The waitakere data frame has 579 rows and 18 columns. Altitude is explanatory, and there are binary responses (presence/absence = 1/0 respectively) for 17 plant species.

Usage

```
data(waitakere)
```

Format

This data frame contains the following columns:

agaaus Agathis australis, or Kauri
beitaw Beilschmiedia tawa, or Tawa
corlae Corynocarpus laevigatus
cyadea Cyathea dealbata
cyamed Cyathea medullaris
daccup Dacrydium cupressinum
dacdac Dacrycarpus dacrydioides
eladen Elaeocarpus dentatus
hedarb Hedycarya arborea
hohpop Species name unknown
kniexc Knightia excelsa, or Rewarewa
kuneri Kunzea ericoides
lepsco Leptospermum scoparium
metrob Metrosideros robusta
neslan Nestegis lanceolata
rhosap Rhopalostylis sapida
vitluc Vitex lucens, or Puriri
altitude meters above sea level

Details

These were collected from the Waitakere Ranges, a small forest in northern Auckland, New Zealand. At 579 sites in the forest, the presence/absence of 17 plant species was recorded, as well as the altitude. Each site was of area size $200m^2$.

Source

Dr Neil Mitchell, University of Auckland.

See Also

[hunua](#).

Examples

```
fit <- vgam(agaau ~ s(altitude, df = 2), binomialff, waitakere)
head(predict(fit, waitakere, type = "response"))
## Not run: plot(fit, se = TRUE, lcol = "orange", scol = "blue")
```

wald.stat

Wald Test Statistics Evaluated at the Null Values

Description

Generic function that computes Wald test statistics evaluated at the null values (consequently they do not suffer from the Hauck-Donner effect).

Usage

```
wald.stat(object, ...)
wald.stat.vlm(object, values0 = 0, subset = NULL, omit1s = TRUE,
              all.out = FALSE, orig.SE = FALSE, iterate.SE = TRUE,
              trace = FALSE, ...)
```

Arguments

object	A vglm fit.
values0	Numeric vector. The null values corresponding to the null hypotheses. Recycled if necessary.
subset	Same as in hdeff .
omit1s	Logical. Does one omit the intercepts? Because the default would be to test that each intercept is equal to 0, which often does not make sense or is unimportant, the intercepts are not tested by default. If they are tested then each linear predictor must have at least one coefficient (from another variable) to be estimated.
all.out	Logical. If TRUE then a list is returned containing various quantities such as the SEs, instead of just the Wald statistics.
orig.SE	Logical. If TRUE then the standard errors are computed at the MLE (of the original object). In practice, the (usual or unmodified) Wald statistics etc. are extracted from <code>summary(object)</code> because it was computed there. These may suffer from the HDE since <i>all</i> the SEs are evaluated at the MLE of the original object. If TRUE then argument <code>iterate.SE</code> may be ignored or overwritten. If <code>orig.SE = FALSE</code> then the <i>k</i> th SE uses the <i>k</i> th value of <code>values0</code> in its computation and <code>iterate.SE</code> specifies the choice of the other coefficients. This argument was previously called as <code>.summary</code> because if TRUE then the Wald statistics are the same as <code>summary(glm())</code> . For one-parameter models setting <code>orig.SE = FALSE</code> results in what is called the <i>null Wald</i> (NW) statistic by some people, e.g., Laskar and King (1997) and Goh and King (1999). The NW statistic does not suffer from the HDE.

<code>iterate.SE</code>	Logical, for the standard error computations. If TRUE then IRLS iterations are performed to get MLEs of the <i>other</i> regression coefficients, subject to one coefficient being equal to the appropriate <code>values0</code> value. If FALSE then the other regression coefficients have values obtained at the original fit. It is recommended that a TRUE be used as the answer tends to be more accurate. If the large (VLM) model matrix only has one column and <code>iterate.SE = TRUE</code> then an error will occur because there are no <i>other</i> regression coefficients to estimate.
<code>trace</code>	Logical. If TRUE then some output is produced as the IRLS iterations proceed. The value NULL means to use the trace value of the fitted object; see <code>vglm.control</code> .
<code>...</code>	Ignored for now.

Details

By default, [`summaryvglm`](#) and most regression modelling functions such as [`summary.glm`](#) compute all the standard errors (SEs) of the estimates at the MLE and not at 0. This corresponds to `orig.SE = TRUE` and it is vulnerable to the Hauck-Donner effect (HDE; see [`hdeff`](#)). One solution is to compute the SEs at 0 (or more generally, at the values of the argument `values0`). This function does that. The two variants of Wald statistics are asymptotically equivalent; however in small samples there can be an appreciable difference, and the difference can be large if the estimates are near to the boundary of the parameter space.

None of the tests here are joint, hence the degrees of freedom is always unity. For a factor with more than 2 levels one can use [`anova.vglm`](#) to test for the significance of the factor. If `orig.SE = FALSE` and `iterate.SE = FALSE` then one retains the MLEs of the original fit for the values of the other coefficients, and replaces one coefficient at a time by the value 0 (or whatever specified by `values0`). One alternative would be to recompute the MLEs of the other coefficients after replacing one of the values; this is the default because `iterate.SE = TRUE` and `orig.SE = FALSE`. Just like with the original IRLS iterations, the iterations here are not guaranteed to converge.

Almost all **VGAM** family functions use the EIM and not the OIM; this affects the resulting standard errors. Also, regularity conditions are assumed for the Wald, likelihood ratio and score tests; some **VGAM** family functions such as [`alaplace1`](#) are experimental and do not satisfy such conditions, therefore naive inference is hazardous.

The default output of this function can be seen by setting `wald0.arg = TRUE` in [`summaryvglm`](#).

Value

By default the signed square root of the Wald statistics whose SEs are computed at one each of the null values. If `all.out = TRUE` then a list is returned with the following components: `wald.stat` the Wald statistic, `SE0` the standard error of that coefficient, `values0` the null values. Approximately, the default Wald statistics output are standard normal random variates if each null hypothesis is true.

Altogether, by the four combinations of `iterate.SE` and `orig.SE`, there are three different variants of the Wald statistic that can be returned.

Warning

This function has been tested but not thoroughly. Convergence failure is possible for some models applied to certain data sets; it is a good idea to set `trace = TRUE` to monitor convergence. For

example, for a particular explanatory variable, the estimated regression coefficients of a non-parallel cumulative logit model (see [cumulative](#)) are ordered, and perturbing one coefficient might disrupt the order and create numerical problems.

Author(s)

Thomas W. Yee

References

- Laskar, M. R. and M. L. King (1997). Modified Wald test for regression disturbances. *Economics Letters*, **56**, 5–11.
- Goh, K.-L. and M. L. King (1999). A correction for local biasedness of the Wald and null Wald tests. *Oxford Bulletin of Economics and Statistics* **61**, 435–450.

See Also

[lrt.stat](#), [score.stat](#), [summaryvglm](#), [summary.glm](#), [anova.vglm](#), [vglm](#), [hdeff](#), [hdeffsev](#).

Examples

```
set.seed(1)
pneumo <- transform(pneumo, let = log(exposure.time),
                    x3 = rnorm(nrow(pneumo)))
(fit <- vglm(cbind(normal, mild, severe) ~ let + x3, propodds, pneumo))
wald.stat(fit) # No HDE here
summary(fit, wald0 = TRUE) # See them here
coef(summary(fit)) # Usual Wald statistics evaluated at the MLE
wald.stat(fit, orig.SE = TRUE) # Same as previous line
```

waldff

Wald Distribution Family Function

Description

Estimates the parameter of the standard Wald distribution by maximum likelihood estimation.

Usage

```
waldff(llambda = "loglink", ilambda = NULL)
```

Arguments

llambda, ilambda

See [CommonVGAMffArguments](#) for information.

Details

The standard Wald distribution is a special case of the inverse Gaussian distribution with $\mu = 1$. It has a density that can be written as

$$f(y; \lambda) = \sqrt{\lambda/(2\pi y^3)} \exp(-\lambda(y-1)^2/(2y))$$

where $y > 0$ and $\lambda > 0$. The mean of Y is 1 (returned as the fitted values) and its variance is $1/\lambda$. By default, $\eta = \log(\lambda)$.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Note

The **VGAM** family function [inv.gaussianff](#) estimates the location parameter μ too.

Author(s)

T. W. Yee

References

Johnson, N. L. and Kotz, S. and Balakrishnan, N. (1994). *Continuous Univariate Distributions*, 2nd edition, Volume 1, New York: Wiley.

See Also

[inv.gaussianff](#), [rinv.gaussian](#).

Examples

```
wdata <- data.frame(y = rinv.gaussian(1000, mu = 1, exp(1)))
wfit <- vglm(y ~ 1, waldff(ilambda = 0.2), wdata, trace = TRUE)
coef(wfit, matrix = TRUE)
Coef(wfit)
summary(wfit)
```

weibull.mean

Weibull Distribution Family Function, Parameterized by the Mean

Description

Maximum likelihood estimation of the 2-parameter Weibull distribution. The mean is one of the parameters. No observations should be censored.

Usage

```
weibull.mean(lmean = "loglink", lshape = "loglink",
             imean = NULL, ishape = NULL, probs.y = c(0.2, 0.5, 0.8),
             imethod = 1, zero = "shape")
```

Arguments

`lmean, lshape` Parameter link functions applied to the (positive) mean parameter (called *mu* below) and (positive) shape parameter (called *a* below). See [Links](#) for more choices.

`imean, ishape` Optional initial values for the mean and shape parameters.

`imethod, zero, probs.y` Details at [CommonVGAMffArguments](#).

Details

See [weibullR](#) for most of the details for this family function too. The mean of Y is $b\Gamma(1 + 1/a)$ (returned as the fitted values), and this is the first parameter (a [loglink](#) link is the default because it is positive). The other parameter is the positive shape parameter a , also having a default [loglink](#) link.

This **VGAM** family function currently does not handle censored data. Fisher scoring is used to estimate the two parameters. Although the expected information matrices used here are valid in all regions of the parameter space, the regularity conditions for maximum likelihood estimation are satisfied only if $a > 2$ (according to Kleiber and Kotz (2003)). If this is violated then a warning message is issued. One can enforce $a > 2$ by choosing `lshape = logofflink(offset = -2)`. Common values of the shape parameter lie between 0.5 and 3.5.

Value

An object of class `"vglmff"` (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

Note

See [weibullR](#) for more details. This **VGAM** family function handles multiple responses.

Author(s)

T. W. Yee

See Also

[weibullR](#), [dweibull](#), [truncweibull](#), [gev](#), [lognormal](#), [expexpff](#), [maxwell](#), [rayleigh](#), [gumbelIII](#).

Examples

```
## Not run:
wdata <- data.frame(x2 = runif(nn <- 1000)) # Complete data
wdata <- transform(wdata, mu      = exp(-1 + 1 * x2),
                  x3      = rnorm(nn),
                  shape1 = exp(1),
                  shape2 = exp(2))

wdata <- transform(wdata,
  y1 = rweibull(nn, shape1, scale = mu / gamma(1 + 1/shape1)),
  y2 = rweibull(nn, shape2, scale = mu / gamma(1 + 1/shape2)))
fit <- vglm(cbind(y1, y2) ~ x2 + x3, weibull.mean, wdata,
  trace = TRUE)
coef(fit, matrix = TRUE)
sqrt(diag(vcov(fit))) # SEs
summary(fit, presid = FALSE)
## End(Not run)
```

weibullR

*Weibull Distribution Family Function***Description**

Maximum likelihood estimation of the 2-parameter Weibull distribution. No observations should be censored.

Usage

```
weibullR(lscale = "loglink", lshape = "loglink",
  iscale = NULL,  ishape = NULL, lss = TRUE, nrfs = 1,
  probs.y = c(0.2, 0.5, 0.8), imethod = 1, zero = "shape")
```

Arguments

lshape, lscale Parameter link functions applied to the (positive) shape parameter (called *a* below) and (positive) scale parameter (called *b* below). See [Links](#) for more choices.

ishape, iscale Optional initial values for the shape and scale parameters.

nrfs Currently this argument is ignored. Numeric, of length one, with value in $[0, 1]$. Weighting factor between Newton-Raphson and Fisher scoring. The value 0 means pure Newton-Raphson, while 1 means pure Fisher scoring. The default value uses a mixture of the two algorithms, and retaining positive-definite working weights.

imethod Initialization method used if there are censored observations. Currently only the values 1 and 2 are allowed.

zero, probs.y, lss Details at [CommonVGAMffArguments](#).

Details

The Weibull density for a response Y is

$$f(y; a, b) = ay^{a-1} \exp[-(y/b)^a]/(b^a)$$

for $a > 0$, $b > 0$, $y > 0$. The cumulative distribution function is

$$F(y; a, b) = 1 - \exp[-(y/b)^a].$$

The mean of Y is $b\Gamma(1 + 1/a)$ (returned as the fitted values), and the mode is at $b(1 - 1/a)^{1/a}$ when $a > 1$. The density is unbounded for $a < 1$. The k th moment about the origin is $E(Y^k) = b^k \Gamma(1 + k/a)$. The hazard function is at^{a-1}/b^a .

This **VGAM** family function currently does not handle censored data. Fisher scoring is used to estimate the two parameters. Although the expected information matrices used here are valid in all regions of the parameter space, the regularity conditions for maximum likelihood estimation are satisfied only if $a > 2$ (according to Kleiber and Kotz (2003)). If this is violated then a warning message is issued. One can enforce $a > 2$ by choosing `lshape = logofflink(offset = -2)`. Common values of the shape parameter lie between 0.5 and 3.5.

Summarized in Harper et al. (2011), for inference, there are 4 cases to consider. If $a \leq 1$ then the MLEs are not consistent (and the smallest observation becomes a hyperefficient solution for the location parameter in the 3-parameter case). If $1 < a < 2$ then MLEs exist but are not asymptotically normal. If $a = 2$ then the MLEs exist and are normal and asymptotically efficient but with a slower convergence rate than when $a > 2$. If $a > 2$ then MLEs have classical asymptotic properties.

The 3-parameter (location is the third parameter) Weibull can be estimated by maximizing a profile log-likelihood (see, e.g., Harper et al. (2011) and Lawless (2003)), else try `gev` which is a better parameterization.

Value

An object of class "vglmff" (see `vglmff-class`). The object is used by modelling functions such as `vglm`, and `vgam`.

Warning

This function is under development to handle other censoring situations. The version of this function which will handle censored data will be called `cerweibull()`. It is currently being written and will use `SurvS4` as input. It should be released in later versions of **VGAM**.

If the shape parameter is less than two then misleading inference may result, e.g., in the summary and `vcov` of the object.

Note

Successful convergence depends on having reasonably good initial values. If the initial values chosen by this function are not good, make use the two initial value arguments.

This **VGAM** family function handles multiple responses.

The Weibull distribution is often an alternative to the lognormal distribution. The inverse Weibull distribution, which is that of $1/Y$ where Y has a Weibull(a, b) distribution, is known as the log-Gompertz distribution.

There are problems implementing the three-parameter Weibull distribution. These are because the classical regularity conditions for the asymptotic properties of the MLEs are not satisfied because the support of the distribution depends on one of the parameters.

Other related distributions are the Maxwell and Rayleigh distributions.

Author(s)

T. W. Yee

References

Kleiber, C. and Kotz, S. (2003). *Statistical Size Distributions in Economics and Actuarial Sciences*, Hoboken, NJ, USA: Wiley-Interscience.

Johnson, N. L. and Kotz, S. and Balakrishnan, N. (1994). *Continuous Univariate Distributions*, 2nd edition, Volume 1, New York: Wiley.

Lawless, J. F. (2003). *Statistical Models and Methods for Lifetime Data*, 2nd ed. Hoboken, NJ, USA: John Wiley & Sons.

Rinne, Horst. (2009). *The Weibull Distribution: A Handbook*. Boca Raton, FL, USA: CRC Press.

Gupta, R. D. and Kundu, D. (2006). On the comparison of Fisher information of the Weibull and GE distributions, *Journal of Statistical Planning and Inference*, **136**, 3130–3144.

Harper, W. V. and Eschenbach, T. G. and James, T. R. (2011). Concerns about Maximum Likelihood Estimation for the Three-Parameter Weibull Distribution: Case Study of Statistical Software, *The American Statistician*, **65**(1), 44–54.

Smith, R. L. (1985). Maximum likelihood estimation in a class of nonregular cases. *Biometrika*, **72**, 67–90.

Smith, R. L. and Naylor, J. C. (1987). A comparison of maximum likelihood and Bayesian estimators for the three-parameter Weibull distribution. *Applied Statistics*, **36**, 358–369.

See Also

[weibull.mean](#), [dweibull](#), [truncweibull](#), [gev](#), [lognormal](#), [expexpff](#), [maxwell](#), [rayleigh](#), [gumbelII](#).

Examples

```
wdata <- data.frame(x2 = runif(nn <- 1000)) # Complete data
wdata <- transform(wdata,
  y1 = rweibull(nn, exp(1), scale = exp(-2 + x2)),
  y2 = rweibull(nn, exp(2), scale = exp( 1 - x2)))
fit <- vglm(cbind(y1, y2) ~ x2, weibullR, wdata, trace = TRUE)
coef(fit, matrix = TRUE)
vcov(fit)
summary(fit)
```

weightsvglm

Prior and Working Weights of a VGLM fit

Description

Returns either the prior weights or working weights of a VGLM object.

Usage

```
weightsvglm(object, type = c("prior", "working"),
            matrix.arg = TRUE, ignore.slot = FALSE,
            deriv.arg = FALSE, ...)
```

Arguments

object	a model object from the VGAM R package that inherits from a <i>vector generalized linear model</i> (VGLM), e.g., a model of class "vglm".
type	Character, which type of weight is to be returned? The default is the first one.
matrix.arg	Logical, whether the answer is returned as a matrix. If not, it will be a vector.
ignore.slot	Logical. If TRUE then object@weights is ignored even if it has been assigned, and the long calculation for object@weights is repeated. This may give a slightly different answer because of the final IRLS step at convergence may or may not assign the latest value of quantities such as the mean and weights.
deriv.arg	Logical. If TRUE then a list with components deriv and weights is returned. See below for more details.
...	Currently ignored.

Details

Prior weights are usually inputted with the weights argument in functions such as `vglm` and `vgam`. It may refer to frequencies of the individual data or be weight matrices specified beforehand.

Working weights are used by the IRLS algorithm. They correspond to the second derivatives of the log-likelihood function with respect to the linear predictors. The working weights correspond to positive-definite weight matrices and are returned in *matrix-band* form, e.g., the first M columns correspond to the diagonals, etc.

If one wants to perturb the linear predictors then the `fitted.values` slots should be assigned to the object before calling this function. The reason is that, for some family functions, the variable μ is used directly as one of the parameter estimates, without recomputing it from η .

Value

If `type = "working"` and `deriv = TRUE` then a list is returned with the two components described below. Otherwise the prior or working weights are returned depending on the value of `type`.

`deriv` Typically the first derivative of the log-likelihood with respect to the linear predictors. For example, this is the variable `deriv.mu` in `vglm.fit()`, or equivalently, the matrix returned in the "deriv" slot of a **VGAM** family function.

`weights` The working weights.

Note

This function is intended to be similar to `weights.glm` (see [glm](#)).

Author(s)

Thomas W. Yee

See Also

[glm](#), [vglmff-class](#), [vglm](#).

Examples

```
pneumo <- transform(pneumo, let = log(exposure.time))
(fit <- vglm(cbind(normal, mild, severe) ~ let,
            cumulative(parallel = TRUE, reverse = TRUE), pneumo))
depvar(fit) # These are sample proportions
weights(fit, type = "prior", matrix = FALSE) # No. of observations

# Look at the working residuals
nn <- nrow(model.matrix(fit, type = "lm"))
M <- ncol(predict(fit))

wwt <- weights(fit, type="working", deriv=TRUE) # Matrix-band format
wz <- m2a(wwt$weights, M = M) # In array format
wzinv <- array(apply(wz, 3, solve), c(M, M, nn))
wresid <- matrix(NA, nn, M) # Working residuals
for (ii in 1:nn)
  wresid[ii, ] <- wzinv[, , ii, drop = TRUE] %*% wwt$deriv[ii, ]
max(abs(c(resid(fit, type = "work")) - c(wresid))) # Should be 0

(zedd <- predict(fit) + wresid) # Adjusted dependent vector
```

wine

Bitterness in Wine Data

Description

This oenological data frame concerns the amount of bitterness in 78 bottles of white wine.

Usage

```
data(wine)
```

Format

A data frame with 4 rows on the following 7 variables.

temp temperature, with levels cold and warm.

contact whether contact of the juice with the skin was allowed or avoided, for a specified period.
Two levels: no or yes.

bitter1, bitter2, bitter3, bitter4, bitter5 numeric vectors, the counts. The order is none to most intense.

Details

The data set comes from Randall (1989) and concerns a factorial experiment for investigating factors that affect the bitterness of white wines. There are two factors in the experiment: temperature at the time of crushing the grapes and contact of the juice with the skin. Two bottles of wine were fermented for each of the treatment combinations. A panel of 9 judges were selected and trained for the ability to detect bitterness. Thus there were 72 bottles in total. Originally, the bitterness of the wine were taken on a continuous scale in the interval from 0 (none) to 100 (intense) but later they were grouped using equal lengths into five ordered categories 1, 2, 3, 4 and 5.

Source

Christensen, R. H. B. (2013) Analysis of ordinal data with cumulative link models—estimation with the R-package **ordinal**. R Package Version 2013.9-30. <https://CRAN.R-project.org/package=ordinal>.

Randall, J. H. (1989). The analysis of sensory data by generalized linear model. *Biometrical Journal* **31**(7), 781–793.

Kosmidis, I. (2014). Improved estimation in cumulative link models. *Journal of the Royal Statistical Society, Series B, Methodological*, **76**(1): 169–196.

Examples

```
wine
summary(wine)
```

```
wrapup.smart
```

Cleans Up After Smart Prediction

Description

wrapup.smart deletes any variables used by smart prediction. Needed by both the modelling function and the prediction function.

Usage

```
wrapup.smart()
```

Details

The variables to be deleted are `.smart.prediction`, `.smart.prediction.counter`, and `.smart.prediction.mode`. The function `wrapup.smart` is useful in R because these variables are held in `smartpredenv`.

See Also

[setup.smart](#).

Examples

```
## Not run: # Place this inside modelling functions such as lm, glm, vglm.
wrapup.smart() # Put at the end of lm

## End(Not run)
```

 yeo.johnson

Yeo-Johnson Transformation

Description

Computes the Yeo-Johnson transformation, which is a normalizing transformation.

Usage

```
yeo.johnson(y, lambda, derivative = 0,
            epsilon = sqrt(.Machine$double.eps), inverse = FALSE)
```

Arguments

<code>y</code>	Numeric, a vector or matrix.
<code>lambda</code>	Numeric. It is recycled to the same length as <code>y</code> if necessary.
<code>derivative</code>	Non-negative integer. The default is the ordinary function evaluation, otherwise the derivative with respect to <code>lambda</code> .
<code>epsilon</code>	Numeric and positive value. The tolerance given to values of <code>lambda</code> when comparing it to 0 or 2.
<code>inverse</code>	Logical. Return the inverse transformation?

Details

The Yeo-Johnson transformation can be thought of as an extension of the Box-Cox transformation. It handles both positive and negative values, whereas the Box-Cox transformation only handles positive values. Both can be used to transform the data so as to improve normality. They can be used to perform LMS quantile regression.

Value

The Yeo-Johnson transformation or its inverse, or its derivatives with respect to `lambda`, of `y`.

Note

If `inverse = TRUE` then the argument `derivative = 0` is required.

Author(s)

Thomas W. Yee

References

Yeo, I.-K. and Johnson, R. A. (2000). A new family of power transformations to improve normality or symmetry. *Biometrika*, **87**, 954–959.

Yee, T. W. (2004). Quantile regression via vector generalized additive models. *Statistics in Medicine*, **23**, 2295–2315.

See Also

[lms.yjn](#), [boxcox](#).

Examples

```
y <- seq(-4, 4, len = (nn <- 200))
ltry <- c(0, 0.5, 1, 1.5, 2) # Try these values of lambda
lltry <- length(ltry)
psi <- matrix(NA_real_, nn, lltry)
for (ii in 1:lltry)
  psi[, ii] <- yeo.johnson(y, lambda = ltry[ii])

## Not run:
matplot(y, psi, type = "l", ylim = c(-4, 4), lwd = 2,
        lty = 1:lltry, col = 1:lltry, las = 1,
        ylab = "Yeo-Johnson transformation",
        main = "Yeo-Johnson transformation with some lambda values")
abline(v = 0, h = 0)
legend(x = 1, y = -0.5, lty = 1:lltry, legend = as.character(ltry),
       lwd = 2, col = 1:lltry)
## End(Not run)
```

Yules

Yule-Simon Distribution

Description

Density, distribution function, quantile function and random generation for the Yule-Simon distribution.

Usage

```
dyules(x, shape, log = FALSE)
pyules(q, shape, lower.tail = TRUE, log.p = FALSE)
qyules(p, shape)
ryules(n, shape)
```

Arguments

x, q, p, n	Same meaning as in Normal .
shape	See yulesimon .
log, lower.tail, log.p	Same meaning as in pnorm or qnorm .

Details

See [yulesimon](#), the **VGAM** family function for estimating the parameter, for the formula of the probability density function and other details.

Value

dyules gives the density, pyules gives the distribution function, qyules gives the quantile function, and ryules generates random deviates.

Note

Numerical problems may occur with qyules() when p is very close to 1.

Author(s)

T. W. Yee

See Also

[yulesimon](#).

Examples

```
dyules(1:20, 2.1)
ryules(20, 2.1)

round(1000 * dyules(1:8, 2))
table(ryules(1000, 2))

## Not run: x <- 0:6
plot(x, dyules(x, shape = 2.2), type = "h", las = 1, col = "blue")

## End(Not run)
```

yulesimon

Yule-Simon Family Function

Description

Estimating the shape parameter of the Yule-Simon distribution.

Usage

```
yulesimon(lshape = "loglink", ishape = NULL, nsimEIM = 200,
          zero = NULL)
```

Arguments

lshape	Link function for the shape parameter, called ρ below. See Links for more choices and for general information.
ishape	Optional initial value for the (positive) parameter. See CommonVGAMffArguments for more information. The default is to obtain an initial value internally. Use this argument if the default fails.
nsimEIM, zero	See CommonVGAMffArguments for more information.

Details

The probability function is

$$f(y; \rho) = \rho * \text{beta}(y, \rho + 1),$$

where the parameter $\rho > 0$, *beta* is the [beta](#) function, and $y = 1, 2, \dots$. The function [dyules](#) computes this probability function. The mean of Y , which is returned as fitted values, is $\rho/(\rho - 1)$ provided $\rho > 1$. The variance of Y is $\rho^2/((\rho - 1)^2(\rho - 2))$ provided $\rho > 2$.

The distribution was named after Udny Yule and Herbert A. Simon. Simon originally called it the Yule distribution. This family function can handle multiple responses.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Author(s)

T. W. Yee

References

Simon, H. A. (1955). On a class of skew distribution functions. *Biometrika*, **42**, 425–440.

See Also

[ryules](#), [simulate.vlm](#).

Examples

```

ydata <- data.frame(x2 = runif(nn <- 1000))
ydata <- transform(ydata, y = ryules(nn, shape = exp(1.5 - x2)))
with(ydata, table(y))
fit <- vglm(y ~ x2, yulesimon, data = ydata, trace = TRUE)
coef(fit, matrix = TRUE)
summary(fit)

```

Zabinom

*Zero-Altered Binomial Distribution***Description**

Density, distribution function, quantile function and random generation for the zero-altered binomial distribution with parameter $pobs0$.

Usage

```

dzabinom(x, size, prob, pobs0 = 0, log = FALSE)
pzabinom(q, size, prob, pobs0 = 0)
qzabinom(p, size, prob, pobs0 = 0)
rzabinom(n, size, prob, pobs0 = 0)

```

Arguments

x, q	vector of quantiles.
p	vector of probabilities.
n	number of observations. If $\text{length}(n) > 1$ then the length is taken to be the number required.
size, prob, log	Parameters from the ordinary binomial distribution (see dbinom).
pobs0	Probability of (an observed) zero, called <i>pobs0</i> . The default value of $pobs0 = 0$ corresponds to the response having a positive binomial distribution.

Details

The probability function of Y is 0 with probability $pobs0$, else a positive binomial(size, prob) distribution.

Value

`dzabinom` gives the density and `pzabinom` gives the distribution function, `qzabinom` gives the quantile function, and `rzabinom` generates random deviates.

Note

The argument $pobs0$ is recycled to the required length, and must have values which lie in the interval $[0, 1]$.

Author(s)

T. W. Yee

See Also[zibinomial](#), [Gaitdbinom](#).**Examples**

```

size <- 10; prob <- 0.15; pobs0 <- 0.05; x <- (-1):7
dzabinom(x, size = size, prob = prob, pobs0 = pobs0)
table(rzabinom(100, size = size, prob = prob, pobs0 = pobs0))

## Not run: x <- 0:10
barplot(rbind(dzabinom(x, size = size, prob = prob, pobs0 = pobs0),
             dbinom(x, size = size, prob = prob)),
        beside = TRUE, col = c("blue", "orange"), cex.main = 0.7, las = 1,
        ylab = "Probability", names.arg = as.character(x),
        main = paste("ZAB(size = ", size, ", prob = ", prob, ", pobs0 = ", pobs0,
                    ") [blue] vs", " Binom(size = ", size, ", prob = ", prob,
                    ") [orange] densities", sep = ""))
## End(Not run)

```

zabinomial

*Zero-Altered Binomial Distribution***Description**

Fits a zero-altered binomial distribution based on a conditional model involving a Bernoulli distribution and a positive-binomial distribution.

Usage

```

zabinomial(lpobs0 = "logitlink", lprob = "logitlink",
           type.fitted = c("mean", "prob", "pobs0"),
           ipobs0 = NULL, iprob = NULL, imethod = 1, zero = NULL)
zabinomialff(lprob = "logitlink", lonempobs0 = "logitlink",
            type.fitted = c("mean", "prob", "pobs0", "onempobs0"),
            iprob = NULL, ionempobs0 = NULL, imethod = 1, zero = "onempobs0")

```

Arguments

lprob	Parameter link function applied to the probability parameter of the binomial distribution. See Links for more choices.
lpobs0	Link function for the parameter p_0 , called pobs0 here. See Links for more choices.
type.fitted	See CommonVGAMffArguments and fittedv1m for information.

iprob, ipobs0 See [CommonVGAMffArguments](#).
 lonempobs0, ionempobs0
 Corresponding argument for the other parameterization. See details below.
 imethod, zero See [CommonVGAMffArguments](#).

Details

The response Y is zero with probability p_0 , else Y has a positive-binomial distribution with probability $1 - p_0$. Thus $0 < p_0 < 1$, which may be modelled as a function of the covariates. The zero-altered binomial distribution differs from the zero-inflated binomial distribution in that the former has zeros coming from one source, whereas the latter has zeros coming from the binomial distribution too. The zero-inflated binomial distribution is implemented in [zibinomial](#). Some people call the zero-altered binomial a *hurdle* model.

The input is currently a vector or one-column matrix. By default, the two linear/additive predictors for `zabinomial()` are $(\text{logit}(p_0), \log(p))^T$.

The **VGAM** family function `zabinomialff()` has a few changes compared to `zabinomial()`. These are: (i) the order of the linear/additive predictors is switched so the binomial probability comes first; (ii) argument `onempobs0` is now 1 minus the probability of an observed 0, i.e., the probability of the positive binomial distribution, i.e., `onempobs0` is `1-pobs0`; (iii) argument `zero` has a new default so that the `onempobs0` is intercept-only by default. Now `zabinomialff()` is generally recommended over `zabinomial()`. Both functions implement Fisher scoring and neither can handle multiple responses.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as `vglm`, and `vgam`.

The `fitted.values` slot of the fitted object, which should be extracted by the generic function `fitted`, returns the mean μ (default) which is given by

$$\mu = (1 - p_0)\mu_b / [1 - (1 - \mu_b)^N]$$

where μ_b is the usual binomial mean. If `type.fitted = "pobs0"` then p_0 is returned.

Note

The response should be a two-column matrix of counts, with first column giving the number of successes.

Note this family function allows p_0 to be modelled as functions of the covariates by having `zero = NULL`. It is a conditional model, not a mixture model.

These family functions effectively combine [posbinomial](#) and [binomialff](#) into one family function.

Author(s)

T. W. Yee

See Also

[dzabinom](#), [zibinomial](#), [posbinomial](#), [spikeplot](#), [binomialff](#), [dbinom](#), [CommonVGAMffArguments](#).

Examples

```

zdata <- data.frame(x2 = runif(nn <- 1000))
zdata <- transform(zdata, size = 10,
                  prob = logitlink(-2 + 3*x2, inverse = TRUE),
                  pobs0 = logitlink(-1 + 2*x2, inverse = TRUE))
zdata <- transform(zdata,
                  y1 = rzabinom(nn, size = size, prob = prob, pobs0 = pobs0))
with(zdata, table(y1))

zfit <- vglm(cbind(y1, size - y1) ~ x2, zabinomial(zero = NULL),
            data = zdata, trace = TRUE)
coef(zfit, matrix = TRUE)
head(fitted(zfit))
head(predict(zfit))
summary(zfit)

```

Zageom

Zero-Altered Geometric Distribution

Description

Density, distribution function, quantile function and random generation for the zero-altered geometric distribution with parameter $pobs0$.

Usage

```

dzageom(x, prob, pobs0 = 0, log = FALSE)
pzageom(q, prob, pobs0 = 0)
qzageom(p, prob, pobs0 = 0)
rzageom(n, prob, pobs0 = 0)

```

Arguments

x, q	vector of quantiles.
p	vector of probabilities.
n	number of observations. If $\text{length}(n) > 1$ then the length is taken to be the number required.
prob, log	Parameters from the ordinary geometric distribution (see dgeom).
pobs0	Probability of (an observed) zero, called <i>pobs0</i> . The default value of $pobs0 = 0$ corresponds to the response having a positive geometric distribution.

Details

The probability function of Y is 0 with probability `pobs0`, else a positive geometric(`prob`) distribution.

Value

`dzageom` gives the density and `pzageom` gives the distribution function, `qzageom` gives the quantile function, and `rzageom` generates random deviates.

Note

The argument `pobs0` is recycled to the required length, and must have values which lie in the interval $[0, 1]$.

Author(s)

T. W. Yee

See Also

[zageometric](#), [zigeometric](#), [rposgeom](#).

Examples

```
prob <- 0.35; pobs0 <- 0.05; x <- (-1):7
dzageom(x, prob = prob, pobs0 = pobs0)
table(rzageom(100, prob = prob, pobs0 = pobs0))

## Not run: x <- 0:10
barplot(rbind(dzageom(x, prob = prob, pobs0 = pobs0),
             dgeom(x, prob = prob)), las = 1,
        beside = TRUE, col = c("blue", "orange"), cex.main = 0.7,
        ylab = "Probability", names.arg = as.character(x),
        main = paste("ZAG(prob = ", prob, ", pobs0 = ", pobs0,
                    ") [blue] vs", " Geometric(prob = ", prob,
                    ") [orange] densities", sep = ""))
## End(Not run)
```

zageometric

Zero-Altered Geometric Distribution

Description

Fits a zero-altered geometric distribution based on a conditional model involving a Bernoulli distribution and a positive-geometric distribution.

Usage

```
zageometric(lpobs0 = "logitlink", lprob = "logitlink",
            type.fitted = c("mean", "prob", "pobs0", "onempobs0"),
            imethod = 1, ipobs0 = NULL, iprob = NULL, zero = NULL)
zageometricff(lprob = "logitlink", lonempobs0 = "logitlink",
              type.fitted = c("mean", "prob", "pobs0", "onempobs0"),
              imethod = 1, iprob = NULL, ionempobs0 = NULL, zero = "onempobs0")
```

Arguments

lpobs0	Link function for the parameter p_0 or ϕ , called pobs0 or phi here. See Links for more choices.
lprob	Parameter link function applied to the probability of success, called prob or p . See Links for more choices.
type.fitted	See CommonVGAMffArguments and fittedvglm for information.
ipobs0, iprob	Optional initial values for the parameters. If given, they must be in range. For multi-column responses, these are recycled sideways.
lonempobs0, ionempobs0	Corresponding argument for the other parameterization. See details below.
zero, imethod	See CommonVGAMffArguments .

Details

The response Y is zero with probability p_0 , or Y has a positive-geometric distribution with probability $1 - p_0$. Thus $0 < p_0 < 1$, which is modelled as a function of the covariates. The zero-altered geometric distribution differs from the zero-inflated geometric distribution in that the former has zeros coming from one source, whereas the latter has zeros coming from the geometric distribution too. The zero-inflated geometric distribution is implemented in the **VGAM** package. Some people call the zero-altered geometric a *hurdle* model.

The input can be a matrix (multiple responses). By default, the two linear/additive predictors of zageometric are $(\text{logit}(\phi), \text{logit}(p))^T$.

The **VGAM** family function zageometricff() has a few changes compared to zageometric(). These are: (i) the order of the linear/additive predictors is switched so the geometric probability comes first; (ii) argument onempobs0 is now 1 minus the probability of an observed 0, i.e., the probability of the positive geometric distribution, i.e., onempobs0 is $1 - p_{obs0}$; (iii) argument zero has a new default so that the pobs0 is intercept-only by default. Now zageometricff() is generally recommended over zageometric(). Both functions implement Fisher scoring and can handle multiple responses.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), and [vgam](#).

The fitted.values slot of the fitted object, which should be extracted by the generic function fitted, returns the mean μ (default) which is given by

$$\mu = (1 - \phi)/p.$$

If `type.fitted = "pobs0"` then p_0 is returned.

Warning

Convergence for this **VGAM** family function seems to depend quite strongly on providing good initial values.

Inference obtained from `summary.vglm` and `summary.vgam` may or may not be correct. In particular, the p-values, standard errors and degrees of freedom may need adjustment. Use simulation on artificial data to check that these are reasonable.

Note

Note this family function allows p_0 to be modelled as functions of the covariates. It is a conditional model, not a mixture model.

This family function effectively combines `binomialff` and `posgeometric()` and `geometric` into one family function. However, `posgeometric()` is not written because it is trivially related to `geometric`.

Author(s)

T. W. Yee

See Also

[dzageom](#), [geometric](#), [zigeometric](#), [spikeplot](#), [dgeom](#), [CommonVGAMffArguments](#), [simulate.vlm](#).

Examples

```
zdata <- data.frame(x2 = runif(nn <- 1000))
zdata <- transform(zdata, pobs0 = logitlink(-1 + 2*x2, inverse = TRUE),
                  prob = logitlink(-2 + 3*x2, inverse = TRUE))
zdata <- transform(zdata, y1 = rzageom(nn, prob = prob, pobs0 = pobs0),
                  y2 = rzageom(nn, prob = prob, pobs0 = pobs0))
with(zdata, table(y1))

fit <- vglm(cbind(y1, y2) ~ x2, zageometric, data = zdata, trace = TRUE)
coef(fit, matrix = TRUE)
head(fitted(fit))
head(predict(fit))
summary(fit)
```

Description

Density, distribution function, quantile function and random generation for the zero-altered negative binomial distribution with parameter p_{obs0} .

Usage

```
dzanegbin(x, size, munb, pobs0 = 0, log = FALSE)
pzanegbin(q, size, munb, pobs0 = 0)
qzanegbin(p, size, munb, pobs0 = 0)
rzanegbin(n, size, munb, pobs0 = 0)
```

Arguments

x, q	vector of quantiles.
p	vector of probabilities.
n	number of observations. If <code>length(n) > 1</code> then the length is taken to be the number required.
size, munb, log	Parameters from the ordinary negative binomial distribution (see dnbinom). Some arguments have been renamed slightly.
pobs0	Probability of zero, called <i>pobs0</i> . The default value of <code>pobs0 = 0</code> corresponds to the response having a positive negative binomial distribution.

Details

The probability function of Y is 0 with probability `pobs0`, else a positive negative binomial(μ_{nb} , size) distribution.

Value

`dzanegbin` gives the density and `pzanegbin` gives the distribution function, `qzanegbin` gives the quantile function, and `rzanegbin` generates random deviates.

Note

The argument `pobs0` is recycled to the required length, and must have values which lie in the interval $[0, 1]$.

Author(s)

T. W. Yee

See Also

[Gaitdnbinom](#), [zanegbinomial](#).

Examples

```
munb <- 3; size <- 4; pobs0 <- 0.3; x <- (-1):7
dzanegbin(x, munb = munb, size = size, pobs0 = pobs0)
table(rzanegbin(100, munb = munb, size = size, pobs0 = pobs0))

## Not run: x <- 0:10
barplot(rbind(dzanegbin(x, munb = munb, size = size, pobs0 = pobs0),
              dnbinom(x, mu = munb, size = size)),
```

```

beside = TRUE, col = c("blue", "green"), cex.main = 0.7,
ylab = "Probability", names.arg = as.character(x), las = 1,
main = paste0("ZANB(munb = ", munb, ", size = ", size, ",
  pobs0 = ", pobs0,
  ") [blue] vs", " NB(mu = ", munb, ", size = ", size,
  ") [green] densities"))
## End(Not run)

```

zanegbinomial

*Zero-Altered Negative Binomial Distribution***Description**

Fits a zero-altered negative binomial distribution based on a conditional model involving a binomial distribution and a positive-negative binomial distribution.

Usage

```

zanegbinomial(zero = "size", type.fitted = c("mean", "munb", "pobs0"),
  mds.min = 1e-3, nsimEIM = 500, cutoff.prob = 0.999,
  eps.trig = 1e-7, max.support = 4000, max.chunk.MB = 30,
  lpobs0 = "logitlink", lmunb = "loglink", lsize = "loglink",
  imethod = 1, ipobs0 = NULL,
  imunb = NULL, iprobs.y = NULL, gprobs.y = (0:9)/10,
  isize = NULL, gsize.mux = exp(c(-30, -20, -15, -10, -6:3)))
zanegbinomialff(lmunb = "loglink", lsize = "loglink", lonempobs0 = "logitlink",
  type.fitted = c("mean", "munb", "pobs0", "onempobs0"),
  isize = NULL, ionempobs0 = NULL, zero = c("size",
  "onempobs0"), mds.min = 1e-3, iprobs.y = NULL, gprobs.y = (0:9)/10,
  cutoff.prob = 0.999, eps.trig = 1e-7, max.support = 4000,
  max.chunk.MB = 30, gsize.mux = exp(c(-30, -20, -15, -10, -6:3)),
  imethod = 1, imunb = NULL,
  nsimEIM = 500)

```

Arguments

lpobs0	Link function for the parameter p_0 , called pobs0 here. See Links for more choices.
lmunb	Link function applied to the munb parameter, which is the mean μ_{nb} of an ordinary negative binomial distribution. See Links for more choices.
lsize	Parameter link function applied to the reciprocal of the dispersion parameter, called k. That is, as k increases, the variance of the response decreases. See Links for more choices.
type.fitted	See CommonVGAMffArguments and fittedvlm for information.
lonempobs0, ionempobs0	Corresponding argument for the other parameterization. See details below.

<code>ipobs0, imunb, isize</code>	Optional initial values for p_0 and μ_{nb} and k . If given then it is okay to give one value for each response/species by inputting a vector whose length is the number of columns of the response matrix.
<code>zero</code>	Specifies which of the three linear predictors are modelled as intercept-only. All parameters can be modelled as a function of the explanatory variables by setting <code>zero = NULL</code> (not recommended). A negative value means that the value is recycled, e.g., setting <code>-3</code> means all k are intercept-only for <code>zanegbinomial</code> . See CommonVGAMffArguments for more information.
<code>nsimEIM, imethod</code>	See CommonVGAMffArguments .
<code>iprbs.y, gsize.mux, gprbs.y</code>	See negbinomial .
<code>cutoff.prob, eps.trig</code>	See negbinomial .
<code>mbs.min, max.support, max.chunk.MB</code>	See negbinomial .

Details

The response Y is zero with probability p_0 , or Y has a positive-negative binomial distribution with probability $1 - p_0$. Thus $0 < p_0 < 1$, which is modelled as a function of the covariates. The zero-altered negative binomial distribution differs from the zero-inflated negative binomial distribution in that the former has zeros coming from one source, whereas the latter has zeros coming from the negative binomial distribution too. The zero-inflated negative binomial distribution is implemented in the **VGAM** package. Some people call the zero-altered negative binomial a *hurdle* model.

For one response/species, by default, the three linear/additive predictors for `zanegbinomial()` are $(\text{logit}(p_0), \log(\mu_{nb}), \log(k))^T$. This vector is recycled for multiple species.

The **VGAM** family function `zanegbinomialff()` has a few changes compared to `zanegbinomial()`. These are: (i) the order of the linear/additive predictors is switched so the negative binomial mean comes first; (ii) argument `onempobs0` is now 1 minus the probability of an observed 0, i.e., the probability of the positive negative binomial distribution, i.e., `onempobs0` is $1 - p_0$; (iii) argument `zero` has a new default so that the `pobs0` is intercept-only by default. Now `zanegbinomialff()` is generally recommended over `zanegbinomial()`. Both functions implement Fisher scoring and can handle multiple responses.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as `vglm`, and `vgam`.

The `fitted.values` slot of the fitted object, which should be extracted by the generic function `fitted`, returns the mean μ (default) which is given by

$$\mu = (1 - p_0)\mu_{nb}/[1 - (k/(k + \mu_{nb}))^k].$$

If `type.fitted = "pobs0"` then p_0 is returned.

Warning

This family function is fragile; it inherits the same difficulties as [posnegbinomial](#). Convergence for this **VGAM** family function seems to depend quite strongly on providing good initial values.

This **VGAM** family function is computationally expensive and usually runs slowly; setting `trace = TRUE` is useful for monitoring convergence.

Inference obtained from `summary.vglm` and `summary.vgam` may or may not be correct. In particular, the p-values, standard errors and degrees of freedom may need adjustment. Use simulation on artificial data to check that these are reasonable.

Note

Note this family function allows p_0 to be modelled as functions of the covariates provided zero is set correctly. It is a conditional model, not a mixture model. Simulated Fisher scoring is the algorithm.

This family function effectively combines [posnegbinomial](#) and [binomialff](#) into one family function.

This family function can handle multiple responses, e.g., more than one species.

Author(s)

T. W. Yee

References

Welsh, A. H., Cunningham, R. B., Donnelly, C. F. and Lindenmayer, D. B. (1996). Modelling the abundances of rare species: statistical models for counts with extra zeros. *Ecological Modelling*, **88**, 297–308.

Yee, T. W. (2014). Reduced-rank vector generalized linear models with two linear predictors. *Computational Statistics and Data Analysis*, **71**, 889–902.

See Also

[gaitdnbinomial](#), [posnegbinomial](#), [Gaitdnbinom](#), [negbinomial](#), [binomialff](#), [zinegbinomial](#), [zipoisson](#), [spikeplot](#), [dnbinom](#), [CommonVGAMffArguments](#), [simulate.vlm](#).

Examples

```
## Not run:
zdata <- data.frame(x2 = runif(nn <- 2000))
zdata <- transform(zdata, pobs0 = logitlink(-1 + 2*x2, inverse = TRUE))
zdata <- transform(zdata,
  y1 = rzanegbin(nn, munb = exp(0+2*x2), size = exp(1), pobs0 = pobs0),
  y2 = rzanegbin(nn, munb = exp(1+2*x2), size = exp(1), pobs0 = pobs0))
with(zdata, table(y1))
with(zdata, table(y2))

fit <- vglm(cbind(y1, y2) ~ x2, zanebinomial, data = zdata, trace = TRUE)
coef(fit, matrix = TRUE)
```

```

head(fitted(fit))
head(predict(fit))

## End(Not run)

```

Zapois

Zero-Altered Poisson Distribution

Description

Density, distribution function, quantile function and random generation for the zero-altered Poisson distribution with parameter $pobs0$.

Usage

```

dzapois(x, lambda, pobs0 = 0, log = FALSE)
pzapois(q, lambda, pobs0 = 0)
qzapois(p, lambda, pobs0 = 0)
rzapois(n, lambda, pobs0 = 0)

```

Arguments

x, q	vector of quantiles.
p	vector of probabilities.
n	number of observations. If $\text{length}(n) > 1$ then the length is taken to be the number required.
lambda	Vector of positive means.
pobs0	Probability of zero, called <i>pobs0</i> . The default value of $pobs0 = 0$ corresponds to the response having a positive Poisson distribution.
log	Logical. Return the logarithm of the answer?

Details

The probability function of Y is 0 with probability $pobs0$, else a positive $Poisson(\lambda)$.

Value

`dzapois` gives the density, `pzapois` gives the distribution function, `qzapois` gives the quantile function, and `rzapois` generates random deviates.

Note

The argument $pobs0$ is recycled to the required length, and must have values which lie in the interval $[0, 1]$.

Author(s)

T. W. Yee

See Also[zapoisson](#), [Gaitdpois](#), [dzipois](#).**Examples**

```
lambda <- 3; pobs0 <- 0.2; x <- (-1):7
(ii <- dzapois(x, lambda, pobs0))
max(abs(cumsum(ii) - pzapois(x, lambda, pobs0))) # Should be 0
table(rzapois(100, lambda, pobs0))
table(qzapois(runif(100), lambda, pobs0))
round(dzapois(0:10, lambda, pobs0) * 100) # Should be similar

## Not run: x <- 0:10
barplot(rbind(dzapois(x, lambda, pobs0), dpois(x, lambda)),
        beside = TRUE, col = c("blue", "green"), las = 1,
        main = paste0("ZAP(", lambda, ", ", pobs0 = ", pobs0, ") [blue]",
                      "vs Poisson(", lambda, ") [green] densities"),
        names.arg = as.character(x), ylab = "Probability")
## End(Not run)
```

zapoisson

*Zero-Altered Poisson Distribution***Description**

Fits a zero-altered Poisson distribution based on a conditional model involving a Bernoulli distribution and a positive-Poisson distribution.

Usage

```
zapoisson(lpobs0 = "logitlink", llambda = "loglink", type.fitted =
  c("mean", "lambda", "pobs0", "onempobs0"), imethod = 1,
  ipobs0 = NULL, ilambda = NULL, ishrinkage = 0.95, probs.y = 0.35,
  zero = NULL)
zapoissonff(llambda = "loglink", lonempobs0 = "logitlink", type.fitted =
  c("mean", "lambda", "pobs0", "onempobs0"), imethod = 1,
  ilambda = NULL, ionempobs0 = NULL, ishrinkage = 0.95,
  probs.y = 0.35, zero = "onempobs0")
```

Arguments

lpobs0 Link function for the parameter p_0 , called `pobs0` here. See [Links](#) for more choices.

llambda Link function for the usual λ parameter. See [Links](#) for more choices.

type.fitted	See CommonVGAMffArguments and fittedvglm for information.
lonempobs0	Corresponding argument for the other parameterization. See details below.
imethod, ipobs0, ionempobs0, ilambda, ishrinkage	See CommonVGAMffArguments for information.
probs.y, zero	See CommonVGAMffArguments for information.

Details

The response Y is zero with probability p_0 , else Y has a positive-Poisson(λ) distribution with probability $1 - p_0$. Thus $0 < p_0 < 1$, which is modelled as a function of the covariates. The zero-altered Poisson distribution differs from the zero-inflated Poisson distribution in that the former has zeros coming from one source, whereas the latter has zeros coming from the Poisson distribution too. Some people call the zero-altered Poisson a *hurdle* model.

For one response/species, by default, the two linear/additive predictors for `zapoisson()` are $(\text{logit}(p_0), \log(\lambda))^T$.

The **VGAM** family function `zapoissonff()` has a few changes compared to `zapoisson()`. These are: (i) the order of the linear/additive predictors is switched so the Poisson mean comes first; (ii) argument `onempobs0` is now 1 minus the probability of an observed 0, i.e., the probability of the positive Poisson distribution, i.e., `onempobs0` is $1 - p_0$; (iii) argument `zero` has a new default so that the `onempobs0` is intercept-only by default. Now `zapoissonff()` is generally recommended over `zapoisson()`. Both functions implement Fisher scoring and can handle multiple responses.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as `vglm`, and `vgam`.

The `fitted.values` slot of the fitted object, which should be extracted by the generic function `fitted`, returns the mean μ (default) which is given by

$$\mu = (1 - p_0)\lambda / [1 - \exp(-\lambda)].$$

If `type.fitted = "pobs0"` then p_0 is returned.

Note

There are subtle differences between this family function and `zipoisson` and `yip88`. In particular, `zipoisson` is a *mixture* model whereas `zapoisson()` and `yip88` are *conditional* models.

Note this family function allows p_0 to be modelled as functions of the covariates.

This family function effectively combines `pospoisson` and `binomialff` into one family function. This family function can handle multiple responses, e.g., more than one species.

It is recommended that `Gaitdpois` be used, e.g., `rgaitdpois(nn, lambda, probs.mlm = pobs0, a.mlm = 0)` instead of `rzapois(nn, lambda, pobs0 = pobs0)`.

Author(s)

T. W. Yee

References

- Welsh, A. H., Cunningham, R. B., Donnelly, C. F. and Lindenmayer, D. B. (1996). Modelling the abundances of rare species: statistical models for counts with extra zeros. *Ecological Modelling*, **88**, 297–308.
- Angers, J-F. and Biswas, A. (2003). A Bayesian analysis of zero-inflated generalized Poisson model. *Computational Statistics & Data Analysis*, **42**, 37–46.
- Yee, T. W. (2014). Reduced-rank vector generalized linear models with two linear predictors. *Computational Statistics and Data Analysis*, **71**, 889–902.

See Also

[Gaitdpois](#), [rzapois](#), [zipoisson](#), [gaitdpoisson](#), [pospoisson](#), [posnegbinomial](#), [spikeplot](#), [binomialff](#), [CommonVGAMffArguments](#), [simulate.vlm](#).

Examples

```
zdata <- data.frame(x2 = runif(nn <- 1000))
zdata <- transform(zdata, pobs0 = logitlink(-1 + 1*x2, inverse = TRUE),
                  lambda = loglink(-0.5 + 2*x2, inverse = TRUE))
zdata <- transform(zdata, y = rgaitdpois(nn, lambda, pobs.mlm = pobs0,
                                       a.mlm = 0))

with(zdata, table(y))
fit <- vglm(y ~ x2, zapoisson, data = zdata, trace = TRUE)
fit <- vglm(y ~ x2, zapoisson, data = zdata, trace = TRUE, crit = "coef")
head(fitted(fit))
head(predict(fit))
head(predict(fit, untransform = TRUE))
coef(fit, matrix = TRUE)
summary(fit)

# Another example -----
# Data from Angers and Biswas (2003)
abdata <- data.frame(y = 0:7, w = c(182, 41, 12, 2, 2, 0, 0, 1))
abdata <- subset(abdata, w > 0)
Abdata <- data.frame(yy = with(abdata, rep(y, w)))
fit3 <- vglm(yy ~ 1, zapoisson, data = Abdata, trace = TRUE, crit = "coef")
coef(fit3, matrix = TRUE)
Coef(fit3) # Estimate lambda (they get 0.6997 with SE 0.1520)
head(fitted(fit3), 1)
with(Abdata, mean(yy)) # Compare this with fitted(fit3)
```

Description

The zero argument allows users to conveniently model certain linear/additive predictors as intercept-only.

Details

Often a certain parameter needs to be modelled simply while other parameters in the model may be more complex, for example, the λ parameter in LMS-Box-Cox quantile regression should be modelled more simply compared to its μ parameter. Another example is the ξ parameter in a GEV distribution which should be modelled simpler than its μ parameter. Using the zero argument allows this to be fitted conveniently without having to input all the constraint matrices explicitly.

The zero argument can be assigned an integer vector from the set $\{1:M\}$ where M is the number of linear/additive predictors. Full details about constraint matrices can be found in the references. See [CommonVGAMffArguments](#) for more information.

Value

Nothing is returned. It is simply a convenient argument for constraining certain linear/additive predictors to be an intercept only.

Warning

The use of other arguments may conflict with the zero argument. For example, using constraints to input constraint matrices may conflict with the zero argument. Another example is the argument `parallel`. In general users should not assume any particular order of precedence when there is potential conflict of definition. Currently no checking for consistency is made.

The argument zero may be renamed in the future to something better.

Side Effects

The argument creates the appropriate constraint matrices internally.

Note

In all **VGAM** family functions `zero = NULL` means none of the linear/additive predictors are modelled as intercepts-only. Almost all **VGAM** family function have `zero = NULL` as the default, but there are some exceptions, e.g., [binom2.or](#).

Typing something like `coef(fit, matrix = TRUE)` is a useful way to ensure that the zero argument has worked as expected.

Author(s)

T. W. Yee

References

Yee, T. W. and Wild, C. J. (1996). Vector generalized additive models. *Journal of the Royal Statistical Society, Series B, Methodological*, **58**, 481–493.

Yee, T. W. and Hastie, T. J. (2003). Reduced-rank vector generalized linear models. *Statistical Modelling*, **3**, 15–41.

See Also

[CommonVGAMffArguments](#), [constraints](#).

Examples

```
args(multinomial)
args(binom2.or)
args(gpd)

#LMS quantile regression example
fit <- vglm(BMI ~ sm.bs(age, df = 4), lms.bcg(zero = c(1, 3)),
           data = bmi.nz, trace = TRUE)
coef(fit, matrix = TRUE)
```

Zeta

*The Zeta Distribution***Description**

Density, distribution function, quantile function and random generation for the zeta distribution.

Usage

```
dzeta(x, shape, log = FALSE)
pzeta(q, shape, lower.tail = TRUE)
qzeta(p, shape)
rzeta(n, shape)
```

Arguments

`x`, `q`, `p`, `n` Same as [Poisson](#).
`shape` The positive shape parameter s .
`lower.tail`, `log` Same meaning as in [Normal](#).

Details

The density function of the zeta distribution is given by

$$y^{-s-1}/\zeta(s+1)$$

where $s > 0$, $y = 1, 2, \dots$, and ζ is Riemann's zeta function.

Value

`dzeta` gives the density, `pzeta` gives the distribution function, `qzeta` gives the quantile function, and `rzeta` generates random deviates.

Note

`qzeta()` runs slower and slower as `shape` approaches 0 and `shape` approaches 1. The **VGAM** family function `zetaaff` estimates the shape parameter s .

Author(s)

T. W. Yee

References

Johnson N. L., Kotz S., and Balakrishnan N. (1993). *Univariate Discrete Distributions*, 2nd ed. New York: Wiley.

See Also

[zeta](#), [zetaff](#), [Oazeta](#), [Oizeta](#), [Otzeta](#).

Examples

```
dzeta(1:20, shape = 2)
myshape <- 0.5
max(abs(pzeta(1:200, myshape) -
      cumsum(1/(1:200)^(1+myshape)) / zeta(myshape+1))) # Should be 0

## Not run: plot(1:6, dzeta(1:6, 2), type = "h", las = 1,
                col = "orange", ylab = "Probability",
                main = "zeta probability function; orange: shape = 2; blue: shape = 1")
points(0.10 + 1:6, dzeta(1:6, 1), type = "h", col = "blue")
## End(Not run)
```

zeta

Riemann's (and the Hurwitz) Zeta Function, With Derivatives

Description

Computes Riemann's zeta function and its first two derivatives. Also can compute the Hurwitz zeta function.

Usage

```
zeta(x, deriv = 0, shift = 1)
```

Arguments

x	A complex-valued vector/matrix whose real values must be ≥ 1 . Otherwise, x may be real. It is called <i>s</i> below. If <i>deriv</i> is 1 or 2 then x must be real and positive.
deriv	An integer equalling 0 or 1 or 2, which is the order of the derivative. The default means it is computed ordinarily.
shift	Positive and numeric, called <i>A</i> below. Allows for the Hurwitz zeta to be returned. The default corresponds to the Riemann formula.

Details

The (Riemann) formula for real s is

$$\sum_{n=1}^{\infty} 1/n^s.$$

While the usual definition involves an infinite series that converges when the real part of the argument is > 1 , more efficient methods have been devised to compute the value. In particular, this function uses Euler–Maclaurin summation. Theoretically, the zeta function can be computed over the whole complex plane because of analytic continuation.

The (Riemann) formula used here for analytic continuation is

$$\zeta(s) = 2^s \pi^{s-1} \sin(\pi s/2) \Gamma(1-s) \zeta(1-s).$$

This is actually one of several formulas, but this one was discovered by Riemann himself and is called the *functional equation*.

The Hurwitz zeta function for real $s > 0$ is

$$\sum_{n=0}^{\infty} 1/(A+n)^s.$$

where $0 < A$ is known here as the *shift*. Since $A = 1$ by default, this function will therefore return Riemann’s zeta function by default. Currently derivatives are unavailable.

Value

The default is a vector/matrix of computed values of Riemann’s zeta function. If *shift* contains values not equal to 1, then this is Hurwitz’s zeta function.

Warning

This function has not been fully tested, especially the derivatives. In particular, analytic continuation does not work here for complex x with $\text{Re}(x) < 1$ because currently the [gamma](#) function does not handle complex arguments.

Note

Estimation of the parameter of the zeta distribution can be achieved with [zetaff](#).

Author(s)

T. W. Yee, with the help of Garry J. Tee.

References

- Riemann, B. (1859). Ueber die Anzahl der Primzahlen unter einer gegebenen Grosse. *Monatsberichte der Berliner Akademie*, November 1859.
- Edwards, H. M. (1974). *Riemann’s Zeta Function*. Academic Press: New York.
- Markman, B. (1965). The Riemann zeta function. *BIT*, **5**, 138–141.
- Abramowitz, M. and Stegun, I. A. (1972). *Handbook of Mathematical Functions with Formulas, Graphs, and Mathematical Tables*, New York: Dover Publications Inc.

See Also

[zetaff](#), [Zeta](#), [oazeta](#), [oizeta](#), [otzeta](#), [lerch](#), [gamma](#).

Examples

```
zeta(2:10)

## Not run:
curve(zeta, -13, 0.8, xlim = c(-12, 10), ylim = c(-1, 4), col = "orange",
      las = 1, main = expression({zeta}(x)))
curve(zeta, 1.2, 12, add = TRUE, col = "orange")
abline(v = 0, h = c(0, 1), lty = "dashed", col = "gray")

curve(zeta, -14, -0.4, col = "orange", main = expression({zeta}(x)))
abline(v = 0, h = 0, lty = "dashed", col = "gray") # Close up plot

x <- seq(0.04, 0.8, len = 100) # Plot of the first derivative
plot(x, zeta(x, deriv = 1), type = "l", las = 1, col = "blue",
     xlim = c(0.04, 3), ylim = c(-6, 0), main = "zeta'(x)")
x <- seq(1.2, 3, len = 100)
lines(x, zeta(x, deriv = 1), col = "blue")
abline(v = 0, h = 0, lty = "dashed", col = "gray")
## End(Not run)

zeta(2) - pi^2 / 6      # Should be 0
zeta(4) - pi^4 / 90    # Should be 0
zeta(6) - pi^6 / 945   # Should be 0
zeta(8) - pi^8 / 9450  # Should be 0
zeta(0, deriv = 1) + 0.5 * log(2*pi) # Should be 0
gamma0 <- 0.5772156649
gamma1 <- -0.07281584548
zeta(0, deriv = 2) -
  gamma1 + 0.5 * (log(2*pi))^2 + pi^2/24 - gamma0^2 / 2 # Should be 0
zeta(0.5, deriv = 1) + 3.92264613 # Should be 0
zeta(2.0, deriv = 1) + 0.93754825431 # Should be 0
```

zetaff

Zeta Distribution Family Function

Description

Estimates the parameter of the zeta distribution.

Usage

```
zetaff(lshape = "loglink", ishape = NULL, gshape = 1 + exp(-seq(7)),
      zero = NULL)
```

Arguments

lshape, ishape, zero

These arguments apply to the (positive) parameter p . See [Links](#) for more choices. Choosing `loglog` constrains $p > 1$, but may fail if the maximum likelihood estimate is less than one. See [CommonVGAMffArguments](#) for more information.

gshape

See [CommonVGAMffArguments](#) for more information.

Details

In this long tailed distribution the response must be a positive integer. The probability function for a response Y is

$$P(Y = y) = 1/[y^{p+1}\zeta(p+1)], \quad p > 0, \quad y = 1, 2, \dots$$

where ζ is Riemann's zeta function. The parameter p is positive, therefore a log link is the default. The mean of Y is $\mu = \zeta(p)/\zeta(p+1)$ (provided $p > 1$) and these are the fitted values. The variance of Y is $\zeta(p-1)/\zeta(p+1) - \mu^2$ provided $p > 2$.

It appears that good initial values are needed for successful convergence. If convergence is not obtained, try several values ranging from values near 0 to values about 10 or more.

Multiple responses are handled.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as `vglm`, and `vgam`.

Note

The [zeta](#) function may be used to compute values of the zeta function.

Author(s)

T. W. Yee

References

pp.527– of Chapter 11 of Johnson N. L., Kemp, A. W. and Kotz S. (2005). *Univariate Discrete Distributions*, 3rd edition, Hoboken, New Jersey: Wiley.

Knight, K. (2000). *Mathematical Statistics*. Boca Raton, FL, USA: Chapman & Hall/CRC Press.

See Also

[zeta](#), [Zeta](#), [gaitdzeta](#), [oazeta](#), [oizeta](#), [otzeta](#), [diffzeta](#), [hzeta](#), [zipf](#).

Examples

```

zdata <- data.frame(y = 1:5, w = c(63, 14, 5, 1, 2)) # Knight, p.304
fit <- vglm(y ~ 1, zetaff, data = zdata, trace = TRUE, weight = w, crit = "c")
(phat <- Coef(fit)) # 1.682557
with(zdata, cbind(round(dzeta(y, phat) * sum(w), 1), w))

with(zdata, weighted.mean(y, w))
fitted(fit, matrix = FALSE)
predict(fit)

# The following should be zero at the MLE:
with(zdata, mean(log(rep(y, w))) + zeta(1+phat, deriv = 1) / zeta(1+phat))

```

Zibinom

*Zero-Inflated Binomial Distribution***Description**

Density, distribution function, quantile function and random generation for the zero-inflated binomial distribution with parameter $pstr0$.

Usage

```

dzibinom(x, size, prob, pstr0 = 0, log = FALSE)
pzibinom(q, size, prob, pstr0 = 0)
qzibinom(p, size, prob, pstr0 = 0)
rzibinom(n, size, prob, pstr0 = 0)

```

Arguments

x, q	vector of quantiles.
p	vector of probabilities.
size	number of trials. It is the N symbol in the formula given in zibinomial .
prob	probability of success on each trial.
n	Same as in runif .
log	Same as pbinom .
pstr0	Probability of a structural zero (i.e., ignoring the binomial distribution), called ϕ . The default value of $\phi = 0$ corresponds to the response having an ordinary binomial distribution.

Details

The probability function of Y is 0 with probability ϕ , and $Binomial(size, prob)$ with probability $1 - \phi$. Thus

$$P(Y = 0) = \phi + (1 - \phi)P(W = 0)$$

where W is distributed $Binomial(size, prob)$.

Value

dzibinom gives the density, pzibinom gives the distribution function, qzibinom gives the quantile function, and rzibinom generates random deviates.

Note

The argument `pstr0` is recycled to the required length, and must have values which lie in the interval $[0, 1]$.

These functions actually allow for *zero-deflation*. That is, the resulting probability of a zero count is *less than* the nominal value of the parent distribution. See [Zipois](#) for more information.

Author(s)

T. W. Yee

See Also

[zibinomial](#), [Gaitdbinom](#), [Binomial](#).

Examples

```
prob <- 0.2; size <- 10; pstr0 <- 0.5
(ii <- dzibinom(0:size, size, prob, pstr0 = pstr0))
max(abs(cumsum(ii) - pzibinom(0:size, size, prob, pstr0 = pstr0))) # 0?
table(rzibinom(100, size, prob, pstr0 = pstr0))

table(qzibinom(runif(100), size, prob, pstr0 = pstr0))
round(dzibinom(0:10, size, prob, pstr0 = pstr0) * 100) # Similar?

## Not run: x <- 0:size
barplot(rbind(dzibinom(x, size, prob, pstr0 = pstr0),
             dbinom(x, size, prob)),
        beside = TRUE, col = c("blue", "green"), ylab = "Probability",
        main = paste0("ZIB(", size, ", ", prob, ", ", pstr0 = ", pstr0, ")",
                      "(blue) vs Binomial(", size, ", ", prob, ") (green)"),
        names.arg = as.character(x), las = 1, lwd = 2)
## End(Not run)
```

zibinomial

Zero-Inflated Binomial Distribution Family Function

Description

Fits a zero-inflated binomial distribution by maximum likelihood estimation.

Usage

```
zibinomial(lpstr0 = "logitlink", lprob = "logitlink",
           type.fitted = c("mean", "prob", "pobs0", "pstr0", "onempstr0"),
           ipstr0 = NULL, zero = NULL, multiple.responses = FALSE,
           imethod = 1)
zibinomialff(lprob = "logitlink", lonempstr0 = "logitlink",
            type.fitted = c("mean", "prob", "pobs0", "pstr0", "onempstr0"),
            ionempstr0 = NULL, zero = "onempstr0",
            multiple.responses = FALSE, imethod = 1)
```

Arguments

`lpstr0, lprob` Link functions for the parameter ϕ and the usual binomial probability μ parameter. See [Links](#) for more choices. For the zero-deflated model see below.

`type.fitted` See [CommonVGAMffArguments](#) and [fittedv1m](#).

`ipstr0` Optional initial values for ϕ , whose values must lie between 0 and 1. The default is to compute an initial value internally. If a vector then recycling is used.

`lonempstr0, ionempstr0` Corresponding arguments for the other parameterization. See details below.

`multiple.responses` Logical. Currently it must be FALSE to mean the function does not handle multiple responses. This is to remain compatible with the same argument in [binomialff](#).

`zero, imethod` See [CommonVGAMffArguments](#) for information. Argument zero changed its default value for version 0.9-2.

Details

These functions are based on

$$P(Y = 0) = \phi + (1 - \phi)(1 - \mu)^N,$$

for $y = 0$, and

$$P(Y = y) = (1 - \phi) \binom{N}{Ny} \mu^{Ny} (1 - \mu)^{N(1-y)}.$$

for $y = 1/N, 2/N, \dots, 1$. That is, the response is a sample proportion out of N trials, and the argument size in [rzibinom](#) is N here. The parameter ϕ is the probability of a structural zero, and it satisfies $0 < \phi < 1$. The mean of Y is $E(Y) = (1 - \phi)\mu$ and these are returned as the fitted values by default. By default, the two linear/additive predictors for [zibinomial\(\)](#) are $(\text{logit}(\phi), \text{logit}(\mu))^T$.

The **VGAM** family function [zibinomialff\(\)](#) has a few changes compared to [zibinomial\(\)](#). These are: (i) the order of the linear/additive predictors is switched so the binomial probability comes first; (ii) argument `onempstr0` is now 1 minus the probability of a structural zero, i.e., the probability of the parent (binomial) component, i.e., `onempstr0` is `1-pstr0`; (iii) argument `zero` has a new default so that the `onempstr0` is intercept-only by default. Now [zibinomialff\(\)](#) is generally recommended over [zibinomial\(\)](#). Both functions implement Fisher scoring.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Warning

Numerical problems can occur. Half-stepping is not uncommon. If failure to converge occurs, make use of the argument `ipstr0` or `ionempstr0`, or `imethod`.

Note

The response variable must have one of the formats described by [binomialff](#), e.g., a factor or two column matrix or a vector of sample proportions with the `weights` argument specifying the values of N .

To work well, one needs large values of N and $\mu > 0$, i.e., the larger N and μ are, the better. If $N = 1$ then the model is unidentifiable since the number of parameters is excessive.

Setting `stepsize = 0.5`, say, may aid convergence.

Estimated probabilities of a structural zero and an observed zero are returned, as in [zipoisson](#).

The *zero-deflated* binomial distribution might be fitted by setting `lpstr0 = identitylink`, albeit, not entirely reliably. See [zipoisson](#) for information that can be applied here. Else try the zero-altered binomial distribution (see [zabinomial](#)).

Author(s)

T. W. Yee

References

Welsh, A. H., Lindenmayer, D. B. and Donnelly, C. F. (2013). Fitting and interpreting occupancy models. *PLOS One*, **8**, 1–21.

See Also

[rzibinom](#), [binomialff](#), [posbinomial](#), [spikeplot](#), [Binomial](#).

Examples

```
size <- 10 # Number of trials; N in the notation above
nn <- 200
zdata <- data.frame(pstr0 = logitlink( 0, inverse = TRUE), # 0.50
                  mubin = logitlink(-1, inverse = TRUE), # Mean of usual binomial
                  sv     = rep(size, length = nn))
zdata <- transform(zdata,
                  y = rzibinom(nn, size = sv, prob = mubin, pstr0 = pstr0))
with(zdata, table(y))
fit <- vglm(cbind(y, sv - y) ~ 1, zibinomialff, data = zdata, trace = TRUE)
fit <- vglm(cbind(y, sv - y) ~ 1, zibinomialff, data = zdata, trace = TRUE,
           stepsize = 0.5)
```

```

coef(fit, matrix = TRUE)
Coef(fit) # Useful for intercept-only models
head(fitted(fit, type = "pobs0")) # Estimate of P(Y = 0)
head(fitted(fit))
with(zdata, mean(y)) # Compare this with fitted(fit)
summary(fit)

```

Zigeom

Zero-Inflated Geometric Distribution

Description

Density, and random generation for the zero-inflated geometric distribution with parameter `pstr0`.

Usage

```

dzigeom(x, prob, pstr0 = 0, log = FALSE)
pzigeom(q, prob, pstr0 = 0)
qzigeom(p, prob, pstr0 = 0)
rzigeom(n, prob, pstr0 = 0)

```

Arguments

<code>x, q</code>	vector of quantiles.
<code>p</code>	vector of probabilities.
<code>prob</code>	see dgeom .
<code>n</code>	Same as in runif .
<code>pstr0</code>	Probability of structural zero (ignoring the geometric distribution), called ϕ . The default value corresponds to the response having an ordinary geometric distribution.
<code>log</code>	Logical. Return the logarithm of the answer?

Details

The probability function of Y is 0 with probability ϕ , and *geometric*(*prob*) with probability $1 - \phi$. Thus

$$P(Y = 0) = \phi + (1 - \phi)P(W = 0)$$

where W is distributed *geometric*(*prob*).

Value

`dzigeom` gives the density, `pzigeom` gives the distribution function, `qzigeom` gives the quantile function, and `rzigeom` generates random deviates.

Note

The argument `pstr0` is recycled to the required length, and must have values which lie in the interval $[0, 1]$.

These functions actually allow for *zero-deflation*. That is, the resulting probability of a zero count is *less than* the nominal value of the parent distribution. See [Zipois](#) for more information.

Author(s)

T. W. Yee

See Also

[zigeometric](#), [dgeom](#).

Examples

```
prob <- 0.5; pstr0 <- 0.2; x <- (-1):20
(ii <- dzigeom(x, prob, pstr0))
max(abs(cumsum(ii) - pzigeom(x, prob, pstr0))) # Should be 0
table(rzigeom(1000, prob, pstr0))

## Not run: x <- 0:10
barplot(rbind(dzigeom(x, prob, pstr0), dgeom(x, prob)),
        beside = TRUE, col = c("blue", "orange"),
        ylab = "P[Y = y]", xlab = "y", las = 1,
        main = paste0("zigeometric(", prob, ", pstr0 = ", pstr0,
                      ") (blue) vs", " geometric(", prob, ") (orange)"),
        names.arg = as.character(x))
## End(Not run)
```

zigeometric

Zero-Inflated Geometric Distribution Family Function

Description

Fits a zero-inflated geometric distribution by maximum likelihood estimation.

Usage

```
zigeometric(lpstr0 = "logitlink", lprob = "logitlink",
            type.fitted = c("mean", "prob", "pobs0", "pstr0", "onempstr0"),
            ipstr0 = NULL, iprob = NULL,
            imethod = 1, bias.red = 0.5, zero = NULL)
zigeometricff(lprob = "logitlink", lonempstr0 = "logitlink",
             type.fitted = c("mean", "prob", "pobs0", "pstr0", "onempstr0"),
             iprob = NULL, ionempstr0 = NULL,
             imethod = 1, bias.red = 0.5, zero = "onempstr0")
```

Arguments

lpstr0, lprob	Link functions for the parameters ϕ and p (prob). The usual geometric probability parameter is the latter. The probability of a structural zero is the former. See Links for more choices. For the zero-deflated model see below.
lonempstr0, ionempstr0	Corresponding arguments for the other parameterization. See details below.
bias.red	A constant used in the initialization process of pstr0. It should lie between 0 and 1, with 1 having no effect.
type.fitted	See CommonVGAMffArguments and fittedvlm for information.
ipstr0, iprob	See CommonVGAMffArguments for information.
zero, imethod	See CommonVGAMffArguments for information.

Details

Function `zigeometric()` is based on

$$P(Y = 0) = \phi + (1 - \phi)p,$$

for $y = 0$, and

$$P(Y = y) = (1 - \phi)p(1 - p)^y.$$

for $y = 1, 2, \dots$. The parameter ϕ satisfies $0 < \phi < 1$. The mean of Y is $E(Y) = (1 - \phi)p/(1 - p)$ and these are returned as the fitted values by default. By default, the two linear/additive predictors are $(\text{logit}(\phi), \text{logit}(p))^T$. Multiple responses are handled.

Estimated probabilities of a structural zero and an observed zero can be returned, as in [zipoisson](#); see [fittedvlm](#) for information.

The **VGAM** family function `zigeometricff()` has a few changes compared to `zigeometric()`. These are: (i) the order of the linear/additive predictors is switched so the geometric probability comes first; (ii) argument `onempstr0` is now 1 minus the probability of a structural zero, i.e., the probability of the parent (geometric) component, i.e., `onempstr0` is `1-pstr0`; (iii) argument `zero` has a new default so that the `onempstr0` is intercept-only by default. Now `zigeometricff()` is generally recommended over `zigeometric()`. Both functions implement Fisher scoring and can handle multiple responses.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Note

The zero-deflated geometric distribution might be fitted by setting `lpstr0 = identitylink`, albeit, not entirely reliably. See [zipoisson](#) for information that can be applied here. Else try the zero-altered geometric distribution (see [zageometric](#)).

Author(s)

T. W. Yee

See Also

[rzigeom](#), [geometric](#), [zageometric](#), [spikeplot](#), [rgeom](#), [simulate.vlm](#).

Examples

```

gdata <- data.frame(x2 = runif(nn <- 1000) - 0.5)
gdata <- transform(gdata, x3 = runif(nn) - 0.5,
                   x4 = runif(nn) - 0.5)
gdata <- transform(gdata, eta1 = 1.0 - 1.0 * x2 + 2.0 * x3,
                   eta2 = -1.0,
                   eta3 = 0.5)
gdata <- transform(gdata, prob1 = logitlink(eta1, inverse = TRUE),
                   prob2 = logitlink(eta2, inverse = TRUE),
                   prob3 = logitlink(eta3, inverse = TRUE))
gdata <- transform(gdata, y1 = rzigeom(nn, prob1, pstr0 = prob3),
                   y2 = rzigeom(nn, prob2, pstr0 = prob3),
                   y3 = rzigeom(nn, prob2, pstr0 = prob3))

with(gdata, table(y1))
with(gdata, table(y2))
with(gdata, table(y3))
head(gdata)

fit1 <- vglm(y1 ~ x2 + x3 + x4, zigeometric(zero = 1), data = gdata, trace = TRUE)
coef(fit1, matrix = TRUE)
head(fitted(fit1, type = "pstr0"))

fit2 <- vglm(cbind(y2, y3) ~ 1, zigeometric(zero = 1), data = gdata, trace = TRUE)
coef(fit2, matrix = TRUE)
summary(fit2)

```

Zinegbin

Zero-Inflated Negative Binomial Distribution

Description

Density, distribution function, quantile function and random generation for the zero-inflated negative binomial distribution with parameter $pstr0$.

Usage

```

dzinegbin(x, size, prob = NULL, munb = NULL, pstr0 = 0, log = FALSE)
pzinegbin(q, size, prob = NULL, munb = NULL, pstr0 = 0)
qzinegbin(p, size, prob = NULL, munb = NULL, pstr0 = 0)
rzinegbin(n, size, prob = NULL, munb = NULL, pstr0 = 0)

```

Arguments

x, q	vector of quantiles.
p	vector of probabilities.
n	Same as in runif .
size, prob, munb, log	Arguments matching dnbinom . The argument munb corresponds to mu in dnbinom and has been renamed to emphasize the fact that it is the mean of the negative binomial <i>component</i> .
pstr0	Probability of structural zero (i.e., ignoring the negative binomial distribution), called ϕ .

Details

The probability function of Y is 0 with probability ϕ , and a negative binomial distribution with probability $1 - \phi$. Thus

$$P(Y = 0) = \phi + (1 - \phi)P(W = 0)$$

where W is distributed as a negative binomial distribution (see [rnbinom](#).) See [negbinomial](#), a **VGAM** family function, for the formula of the probability density function and other details of the negative binomial distribution.

Value

[dzinegbin](#) gives the density, [pzinegbin](#) gives the distribution function, [qzinegbin](#) gives the quantile function, and [rzinegbin](#) generates random deviates.

Note

The argument `pstr0` is recycled to the required length, and must have values which lie in the interval $[0, 1]$.

These functions actually allow for *zero-deflation*. That is, the resulting probability of a zero count is *less than* the nominal value of the parent distribution. See [Zipois](#) for more information.

Author(s)

T. W. Yee

See Also

[zinegbinomial](#), [rnbinom](#), [rzipois](#).

Examples

```
munb <- 3; pstr0 <- 0.2; size <- k <- 10; x <- 0:10
(ii <- dzinegbin(x, pstr0 = pstr0, mu = munb, size = k))
max(abs(cumsum(ii) - pzinegbin(x, pstr0 = pstr0, mu = munb, size = k)))
table(rzinegbin(100, pstr0 = pstr0, mu = munb, size = k))

table(qzinegbin(runif(1000), pstr0 = pstr0, mu = munb, size = k))
```



```

round(dzinegbin(x, pstr0 = pstr0, mu = munb, size = k) * 1000) # Similar?

## Not run: barplot(rbind(dzinegbin(x, pstr0 = pstr0, mu = munb, size = k),
                        dnbinom(x, mu = munb, size = k)), las = 1,
                  beside = TRUE, col = c("blue", "green"), ylab = "Probability",
                  main = paste("ZINB(mu = ", munb, ", k = ", k, ", pstr0 = ", pstr0,
                              ") (blue) vs NB(mu = ", munb,
                              ", size = ", k, ") (green)", sep = ""),
                  names.arg = as.character(x))
## End(Not run)

```

zinegbinomial

Zero-Inflated Negative Binomial Distribution Family Function

Description

Fits a zero-inflated negative binomial distribution by full maximum likelihood estimation.

Usage

```

zinegbinomial(zero = "size",
              type.fitted = c("mean", "munb", "pobs0", "pstr0",
                              "onempstr0"),
              mds.min = 1e-3, nsimEIM = 500, cutoff.prob = 0.999,
              eps.trig = 1e-7, max.support = 4000, max.chunk.MB = 30,
              lpstr0 = "logitlink", lmunb = "loglink", lsize = "loglink",
              imethod = 1, ipstr0 = NULL, imunb = NULL,
              iprobs.y = NULL, isize = NULL,
              gprobs.y = (0:9)/10,
              gsize.mux = exp(c(-30, -20, -15, -10, -6:3)))
zinegbinomialff(lmunb = "loglink", lsize = "loglink", lonempstr0 = "logitlink",
                type.fitted = c("mean", "munb", "pobs0", "pstr0",
                                "onempstr0"), imunb = NULL, isize = NULL, ionempstr0 =
                NULL, zero = c("size", "onempstr0"), imethod = 1,
                iprobs.y = NULL, cutoff.prob = 0.999,
                eps.trig = 1e-7, max.support = 4000, max.chunk.MB = 30,
                gprobs.y = (0:9)/10, gsize.mux = exp((-12:6)/2),
                mds.min = 1e-3, nsimEIM = 500)

```

Arguments

lpstr0, lmunb, lsize

Link functions for the parameters ϕ , the mean and k ; see [negbinomial](#) for details, and [Links](#) for more choices. For the zero-deflated model see below.

type.fitted See [CommonVGAMffArguments](#) and [fittedvlm](#) for more information.

ipstr0, isize, imunb

Optional initial values for ϕ and k and μ . The default is to compute an initial value internally for both. If a vector then recycling is used.

lonempstr0, ionempstr0	Corresponding arguments for the other parameterization. See details below.
imethod	An integer with value 1 or 2 or 3 which specifies the initialization method for the mean parameter. If failure to converge occurs try another value. See CommonVGAMffArguments for more information.
zero	Specifies which linear/additive predictors are to be modelled as intercept-only. They can be such that their absolute values are either 1 or 2 or 3. The default is the ϕ and k parameters (both for each response). See CommonVGAMffArguments for more information.
nsimEIM	See CommonVGAMffArguments for information.
iprob.y, cutoff.prob, max.support, max.chunk.MB	See negbinomial and/or posnegbinomial for details.
mds.min, eps.trig	See negbinomial for details.
gprob.y, gsize.mux	These arguments relate to grid searching in the initialization process. See negbinomial and/or posnegbinomial for details.

Details

These functions are based on

$$P(Y = 0) = \phi + (1 - \phi)(k/(k + \mu))^k,$$

and for $y = 1, 2, \dots$,

$$P(Y = y) = (1 - \phi) \text{dnbinom}(y, \mu, k).$$

The parameter ϕ satisfies $0 < \phi < 1$. The mean of Y is $(1 - \phi)\mu$ (returned as the fitted values). By default, the three linear/additive predictors for `zinegbinomial()` are $(\text{logit}(\phi), \log(\mu), \log(k))^T$. See [negbinomial](#), another **VGAM** family function, for the formula of the probability density function and other details of the negative binomial distribution.

Independent multiple responses are handled. If so then arguments `ipstr0` and `isize` may be vectors with length equal to the number of responses.

The **VGAM** family function `zinegbinomialff()` has a few changes compared to `zinegbinomial()`. These are: (i) the order of the linear/additive predictors is switched so the NB mean comes first; (ii) `onempstr0` is now 1 minus the probability of a structural 0, i.e., the probability of the parent (NB) component, i.e., `onempstr0` is `1-pstr0`; (iii) argument `zero` has a new default so that the `onempstr0` is intercept-only by default. Now `zinegbinomialff()` is generally recommended over `zinegbinomial()`. Both functions implement Fisher scoring and can handle multiple responses.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as `vglm`, and `vgam`.

Warning

This model can be difficult to fit to data, and this family function is fragile. The model is especially difficult to fit reliably when the estimated k parameter is very large (so the model approaches a zero-inflated Poisson distribution) or much less than 1 (and gets more difficult as it approaches 0). Numerical problems can also occur, e.g., when the probability of a zero is actually less than, and not more than, the nominal probability of zero. Similarly, numerical problems can occur if there is little or no 0-inflation, or when the sample size is small. Half-stepping is not uncommon. Successful convergence is sensitive to the initial values, therefore if failure to converge occurs, try using combinations of arguments `stepsize` (in `vglm.control`), `imethod`, `imunb`, `ipstr0`, `isize`, and/or `zero` if there are explanatory variables. Else try fitting an ordinary [negbinomial](#) model or a [zipoisson](#) model.

This **VGAM** family function can be computationally expensive and can run slowly; setting `trace = TRUE` is useful for monitoring convergence.

Note

Estimated probabilities of a structural zero and an observed zero can be returned, as in [zipoisson](#); see [fittedvglm](#) for more information.

If k is large then the use of **VGAM** family function [zipoisson](#) is probably preferable. This follows because the Poisson is the limiting distribution of a negative binomial as k tends to infinity.

The *zero-deflated* negative binomial distribution might be fitted by setting `lpstr0 = identitylink`, albeit, not entirely reliably. See [zipoisson](#) for information that can be applied here. Else try the zero-altered negative binomial distribution (see [zanegbinomial](#)).

Author(s)

T. W. Yee

See Also

[gaitdnbinomial](#), [Zinegbin](#), [negbinomial](#), [spikeplot](#), [rpois](#), [CommonVGAMffArguments](#).

Examples

```
## Not run:
# Example 1
ndata <- data.frame(x2 = runif(nn <- 1000))
ndata <- transform(ndata, pstr0 = logitlink(-0.5 + 1 * x2, inverse = TRUE),
                  munb = exp( 3 + 1 * x2),
                  size = exp( 0 + 2 * x2))
ndata <- transform(ndata,
                  y1 = rzinegbin(nn, mu = munb, size = size, pstr0 = pstr0))
with(ndata, table(y1)["0"] / sum(table(y1)))
nfit <- vglm(y1 ~ x2, zinegbinomial(zero = NULL), data = ndata)
coef(nfit, matrix = TRUE)
summary(nfit)
head(cbind(fitted(nfit), with(ndata, (1 - pstr0) * munb)))
round(vcov(nfit), 3)
```

```

# Example 2: RR-ZINB could also be called a COZIVGLM-ZINB-2
ndata <- data.frame(x2 = runif(nn <- 2000))
ndata <- transform(ndata, x3 = runif(nn))
ndata <- transform(ndata, eta1 =      3  + 1  * x2 + 2 * x3)
ndata <- transform(ndata, pstr0 = logitlink(-1.5 + 0.5 * eta1, inverse = TRUE),
                    munb = exp(eta1),
                    size = exp(4))

ndata <- transform(ndata,
                  y1 = rzinegbin(nn, pstr0 = pstr0, mu = munb, size = size))
with(ndata, table(y1)["0"] / sum(table(y1)))
rrzinb <- rrvglm(y1 ~ x2 + x3, zinegbinomial(zero = NULL), data = ndata,
                Index.corner = 2, str0 = 3, trace = TRUE)
coef(rrzinb, matrix = TRUE)
Coef(rrzinb)

## End(Not run)

```

zipebcom

Exchangeable Bivariate cloglog Odds-ratio Model From a Zero-inflated Poisson Distribution

Description

Fits an exchangeable bivariate odds-ratio model to two binary responses with a complementary log-log link. The data are assumed to come from a zero-inflated Poisson distribution that has been converted to presence/absence.

Usage

```

zipebcom(lmu12 = "clogloglink", lphi12 = "logitlink", loratio = "loglink",
         imu12 = NULL, iphi12 = NULL, ioratio = NULL,
         zero = c("phi12", "oratio"), tol = 0.001, addRidge = 0.001)

```

Arguments

lmu12, imu12	Link function, extra argument and optional initial values for the first (and second) marginal probabilities. Argument lmu12 should be left alone. Argument imu12 may be of length 2 (one element for each response).
lphi12	Link function applied to the ϕ parameter of the zero-inflated Poisson distribution (see zipoisson). See Links for more choices.
loratio	Link function applied to the odds ratio. See Links for more choices.
iphi12, ioratio	Optional initial values for ϕ and the odds ratio. See CommonVGAMffArguments for more details. In general, good initial values (especially for iphi12) are often required, therefore use these arguments if convergence failure occurs. If inputted, the value of iphi12 cannot be more than the sample proportions of zeros in either response.

zero	Which linear/additive predictor is modelled as an intercept only? A NULL means none. The default has both ϕ and the odds ratio as not being modelled as a function of the explanatory variables (apart from an intercept). See CommonVGAMffArguments for information.
tol	Tolerance for testing independence. Should be some small positive numerical value.
addRidge	Some small positive numerical value. The first two diagonal elements of the working weight matrices are multiplied by 1+addRidge to make it diagonally dominant, therefore positive-definite.

Details

This **VGAM** family function fits an exchangeable bivariate odds ratio model ([binom2.or](#)) with a [clogloglink](#) link. The data are assumed to come from a zero-inflated Poisson (ZIP) distribution that has been converted to presence/absence. Explicitly, the default model is

$$\text{cloglog}[P(Y_j = 1)/(1 - \phi)] = \eta_1, \quad j = 1, 2$$

for the (exchangeable) marginals, and

$$\text{logit}[\phi] = \eta_2,$$

for the mixing parameter, and

$$\log[P(Y_{00} = 1)P(Y_{11} = 1)/(P(Y_{01} = 1)P(Y_{10} = 1))] = \eta_3,$$

specifies the dependency between the two responses. Here, the responses equal 1 for a success and a 0 for a failure, and the odds ratio is often written $\psi = p_{00}p_{11}/(p_{10}p_{01})$. We have $p_{10} = p_{01}$ because of the exchangeability.

The second linear/additive predictor models the ϕ parameter (see [zipoisson](#)). The third linear/additive predictor is the same as [binom2.or](#), viz., the log odds ratio.

Suppose a dataset1 comes from a Poisson distribution that has been converted to presence/absence, and that both marginal probabilities are the same (exchangeable). Then [binom2.or\("clogloglink", exch=TRUE\)](#) is appropriate. Now suppose a dataset2 comes from a *zero-inflated* Poisson distribution. The first linear/additive predictor of [zipebcom\(\)](#) applied to dataset2 is the same as that of [binom2.or\("clogloglink", exch=TRUE\)](#) applied to dataset1. That is, the ϕ has been taken care of by [zipebcom\(\)](#) so that it is just like the simpler [binom2.or](#).

Note that, for η_1 , $\mu_{12} = \text{prob}_{12} / (1 - \phi_{12})$ where prob_{12} is the probability of a 1 under the ZIP model. Here, μ_{12} correspond to μ_1 and μ_2 in the [binom2.or](#)-Poisson model.

If $\phi = 0$ then [zipebcom\(\)](#) should be equivalent to [binom2.or\("clogloglink", exch=TRUE\)](#). Full details are given in Yee and Dirnbock (2009).

The leading 2×2 submatrix of the expected information matrix (EIM) is of rank-1, not 2! This is due to the fact that the parameters corresponding to the first two linear/additive predictors are unidentifiable. The quick fix around this problem is to use the [addRidge](#) adjustment. The model is fitted by maximum likelihood estimation since the full likelihood is specified. Fisher scoring is implemented.

The default models η_2 and η_3 as single parameters only, but this can be circumvented by setting [zero=NULL](#) in order to model the ϕ and odds ratio as a function of all the explanatory variables.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

When fitted, the `fitted.values` slot of the object contains the four joint probabilities, labelled as $(Y_1, Y_2) = (0,0), (0,1), (1,0), (1,1)$, respectively. These estimated probabilities should be extracted with the `fitted` generic function.

Warning

The fact that the EIM is not of full rank may mean the model is naturally ill-conditioned. Not sure whether there are any negative consequences wrt theory. For now it is certainly safer to fit [binom2.or](#) to bivariate binary responses.

Note

The "12" in the argument names reinforce the user about the exchangeability assumption. The name of this **VGAM** family function stands for *zero-inflated Poisson exchangeable bivariate complementary log-log odds-ratio model* or ZIP-EBCOM.

See [binom2.or](#) for details that are pertinent to this **VGAM** family function too. Even better initial values are usually needed here.

The `xij` (see [vglm.control](#)) argument enables environmental variables with different values at the two time points to be entered into an exchangeable [binom2.or](#) model. See the author's webpage for sample code.

References

Yee, T. W. and Dirnbock, T. (2009). Models for analysing species' presence/absence data at two time points. *Journal of Theoretical Biology*, **259**(4), 684–694.

See Also

[binom2.or](#), [zipoisson](#), [clogloglink](#), [CommonVGAMffArguments](#).

Examples

```
## Not run:
zdata <- data.frame(x2 = seq(0, 1, len = (nsites <- 2000)))
zdata <- transform(zdata, eta1 = -3 + 5 * x2,
                  phi1 = logitlink(-1, inverse = TRUE),
                  oratio = exp(2))
zdata <- transform(zdata, mu12 = clogloglink(eta1, inverse = TRUE) * (1-phi1))
tmat <- with(zdata, rbinom2.or(nsites, mu1 = mu12, oratio = oratio, exch = TRUE))
zdata <- transform(zdata, ybin1 = tmat[, 1], ybin2 = tmat[, 2])

with(zdata, table(ybin1, ybin2)) / nsites # For interest only
# Various plots of the data, for interest only
par(mfrow = c(2, 2))
plot(jitter(ybin1) ~ x2, data = zdata, col = "blue")
```

```

plot(jitter(ybin2) ~ jitter(ybin1), data = zdata, col = "blue")

plot(mu12 ~ x2, data = zdata, col = "blue", type = "l", ylim = 0:1,
     ylab = "Probability", main = "Marginal probability and phi")
with(zdata, abline(h = phi1[1], col = "red", lty = "dashed"))

tmat2 <- with(zdata, dbinom2.or(mu1 = mu12, oratio = oratio, exch = TRUE))
with(zdata, matplot(x2, tmat2, col = 1:4, type = "l", ylim = 0:1,
                   ylab = "Probability", main = "Joint probabilities"))

# Now fit the model to the data.
fit <- vglm(cbind(ybin1, ybin2) ~ x2, zipfcom, data = zdata, trace = TRUE)
coef(fit, matrix = TRUE)
summary(fit)
vcov(fit)
## End(Not run)

```

Zipf

The Zipf Distribution

Description

Density, distribution function, quantile function and random generation for the Zipf distribution.

Usage

```

dzipf(x, N, shape, log = FALSE)
pzipf(q, N, shape, log.p = FALSE)
qzipf(p, N, shape)
rzipf(n, N, shape)

```

Arguments

x, q, p, n	Same as Poisson .
N, shape	the number of elements, and the exponent characterizing the distribution. See zipf for more details.
log, log.p	Same meaning as in Normal .

Details

This is a finite version of the zeta distribution. See [zetaff](#) for more details. In general, these functions runs slower and slower as N increases.

Value

`dzipf` gives the density, `pzipf` gives the cumulative distribution function, `qzipf` gives the quantile function, and `rzipf` generates random deviates.

Author(s)

T. W. Yee

See Also[zipf](#), [Zipfmb](#).**Examples**

```

N <- 10; shape <- 0.5; y <- 1:N
proby <- dzipf(y, N = N, shape = shape)
## Not run: plot(proby ~ y, type = "h", col = "blue",
  ylim = c(0, 0.2), ylab = "Probability", lwd = 2, las = 1,
  main = paste0("Zipf(N = ", N, ", shape = ", shape, ")"))
## End(Not run)
sum(proby) # Should be 1
max(abs(cumsum(proby) - pzipf(y, N = N, shape = shape))) # 0?

```

zipf

*Zipf Distribution Family Function***Description**

Estimates the parameter of the Zipf distribution.

Usage

```
zipf(N = NULL, lshape = "loglink", ishape = NULL)
```

Arguments

N	Number of elements, an integer satisfying $1 < N < \text{Inf}$. The default is to use the maximum value of the response. If given, N must be no less than the largest response value. If $N = \text{Inf}$ and $s > 1$ then this is the zeta distribution (use zetaff instead).
lshape	Parameter link function applied to the (positive) shape parameter s . See Links for more choices.
ishape	Optional initial value for the parameter s . The default is to choose an initial value internally. If convergence failure occurs use this argument to input a value.

Details

The probability function for a response Y is

$$P(Y = y) = y^{-s} / \sum_{i=1}^N i^{-s}, \quad s > 0, \quad y = 1, 2, \dots, N,$$

where s is the exponent characterizing the distribution. The mean of Y , which are returned as the fitted values, is $\mu = H_{N,s-1}/H_{N,s}$ where $H_{n,m} = \sum_{i=1}^n i^{-m}$ is the n th generalized harmonic number.

Zipf's law is an experimental law which is often applied to the study of the frequency of words in a corpus of natural language utterances. It states that the frequency of any word is inversely proportional to its rank in the frequency table. For example, "the" and "of" are first two most common words, and Zipf's law states that "the" is twice as common as "of". Many other natural phenomena conform to Zipf's law.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#) and [vgam](#).

Note

Upon convergence, the N is stored as @misc\$N.

Author(s)

T. W. Yee

References

pp.526– of Chapter 11 of Johnson N. L., Kemp, A. W. and Kotz S. (2005). *Univariate Discrete Distributions*, 3rd edition, Hoboken, New Jersey, USA: Wiley.

See Also

[dzipf](#), [zetaff](#), [simulate.vlm](#).

Examples

```
zdata <- data.frame(y = 1:5, ofreq = c(63, 14, 5, 1, 2))
zfit <- vglm(y ~ 1, zipf, data = zdata, trace = TRUE, weight = ofreq)
zfit <- vglm(y ~ 1, zipf(lshape = "identitylink", ishape = 3.4), data = zdata,
  trace = TRUE, weight = ofreq, crit = "coef")
zfit@misc$N
(shape.hat <- Coef(zfit))
with(zdata, weighted.mean(y, ofreq))
fitted(zfit, matrix = FALSE)
```

 Zipfmb

The Zipf-Mandelbrot Distribution

Description

Density, distribution function, quantile function and random generation for the Mandelbrot distribution.

Usage

```
dzipfmb(x, shape, start = 1, log = FALSE)
pzipfmb(q, shape, start = 1, lower.tail = TRUE, log.p = FALSE)
qzipfmb(p, shape, start = 1)
rzipfmb(n, shape, start = 1)
```

Arguments

x	vector of (non-negative integer) quantiles.
q	vector of quantiles.
p	vector of probabilities.
n	number of random values to return.
shape	vector of positive shape parameter.
start	integer, the minimum value of the support of the distribution.
log, log.p	logical; if TRUE, probabilities p are given as log(p)
lower.tail	logical; if TRUE (default), probabilities are P[X <= x], otherwise, P[X > x].

Details

The probability mass function of the Zipf-Mandelbrot distribution is given by

$$\Pr(Y = y; s) = \frac{s \Gamma(y_{min})}{\Gamma(y_{min} - s)} \cdot \frac{\Gamma(y - s)}{\Gamma(y + 1)}$$

where $0 \leq b < 1$ and the starting value start being by default 1.

Value

dzipfmb gives the density, pzipfmb gives the distribution function, qzipfmb gives the quantile function, and rzipfmb generates random deviates.

Author(s)

M. Chou, with edits by T. W. Yee.

References

Mandelbrot, B. (1961). On the theory of word frequencies and on related Markovian models of discourse. In R. Jakobson, *Structure of Language and its Mathematical Aspects*, pp. 190–219, Providence, RI, USA. American Mathematical Society.

Moreno-Sanchez, I. and Font-Clos, F. and Corral, A. (2016). Large-Scale Analysis of Zipf's Law in English Texts. *PLoS ONE*, **11**(1), 1–19.

See Also

[Zipf](#).

Examples

```
aa <- 1:10
(pp <- pzipfmb(aa, shape = 0.5, start = 1))
cumsum(dzipfmb(aa, shape = 0.5, start = 1)) # Should be same
qzipfmb(pp, shape = 0.5, start = 1) - aa # Should be all 0s

rdiffzeta(30, 0.5)

## Not run: x <- 1:10
plot(x, dzipfmb(x, shape = 0.5), type = "h", ylim = 0:1,
     sub = "shape=0.5", las = 1, col = "blue", ylab = "Probability",
     main = "Zipf-Mandelbrot distribution: blue=PMF; orange=CDF")
lines(x+0.1, pzipfmb(x, shape = 0.5), col = "red", lty = 3, type = "h")

## End(Not run)
```

Zipois

Zero-Inflated Poisson Distribution

Description

Density, distribution function, quantile function and random generation for the zero-inflated and zero-deflated Poisson distribution with parameter $pstr0$.

Usage

```
dzipois(x, lambda, pstr0 = 0, log = FALSE)
pzipois(q, lambda, pstr0 = 0)
qzipois(p, lambda, pstr0 = 0)
rzipois(n, lambda, pstr0 = 0)
```

Arguments

x, q	vector of quantiles.
p	vector of probabilities.
n	number of observations. Must be a single positive integer.
lambda	Vector of positive means.
pstr0	Probability of a structural zero (i.e., ignoring the Poisson distribution), called ϕ . The default value of $\phi = 0$ corresponds to the response having an ordinary Poisson distribution. If ϕ lies in $(0, 1)$ then this is known as the zero-inflated Poisson (ZIP) distribution. This argument may be negative to allow for 0-deflation, hence its interpretation as a probability ceases.
log	Logical. Return the logarithm of the answer?

Details

The probability function of Y is 0 with probability ϕ , and $Poisson(\lambda)$ with probability $1 - \phi$. Thus

$$P(Y = 0) = \phi + (1 - \phi)P(W = 0)$$

where W is distributed $Poisson(\lambda)$.

Value

`dzipois` gives the density, `pzipois` gives the distribution function, `qzipois` gives the quantile function, and `rzipois` generates random deviates.

Note

The argument `pstr0` is recycled to the required length, and must have values which lie in the interval $[0, 1]$.

These functions actually allow for the *zero-deflated Poisson* (ZDP) distribution. Here, `pstr0` is also permitted to lie in the interval $[-1/\exp(1), 0]$. The resulting probability of a zero count is *less than* the nominal Poisson value, and the use of `pstr0` to stand for the probability of a structural zero loses its meaning. When `pstr0` equals $-1/\exp(1)$ this corresponds to the positive-Poisson distribution (e.g., see [Gaitdpois](#)), also called the zero-truncated Poisson or ZTP.

The *zero-modified Poisson* (ZMP) is a combination of the ZIP and ZDP and ZTP distributions. The family function

Author(s)

T. W. Yee

See Also

[zipoisson](#), [Gaitdpois](#), [dpois](#), [rzinegbin](#).

Examples

```

lambda <- 3; pstr0 <- 0.2; x <- (-1):7
(ii <- dzipois(x, lambda, pstr0 = pstr0))
max(abs(cumsum(ii) - pzipois(x, lambda, pstr0 = pstr0))) # 0?
table(rzipois(100, lambda, pstr0 = pstr0))

table(qzipois(runif(100), lambda, pstr0))
round(dzipois(0:10, lambda, pstr0 = pstr0) * 100) # Similar?

## Not run: x <- 0:10
par(mfrow = c(2, 1)) # Zero-inflated Poisson
barplot(rbind(dzipois(x, lambda, pstr0 = pstr0), dpois(x, lambda)),
        beside = TRUE, col = c("blue", "orange"),
        main = paste0("ZIP(", lambda,
                      ", pstr0 = ", pstr0, ") (blue) vs",
                      " Poisson(", lambda, ") (orange)"),
        names.arg = as.character(x))

deflat.limit <- -1 / expm1(lambda) # Zero-deflated Poisson
newpstr0 <- round(deflat.limit / 1.5, 3)
barplot(rbind(dzipois(x, lambda, pstr0 = newpstr0),
              dpois(x, lambda)),
        beside = TRUE, col = c("blue", "orange"),
        main = paste0("ZDP(", lambda, ", pstr0 = ", newpstr0, ")",
                      " (blue) vs Poisson(", lambda, ") (orange)"),
        names.arg = as.character(x))
## End(Not run)

```

zipoisson

Zero-Inflated Poisson Distribution Family Function

Description

Fits a zero-inflated or zero-deflated Poisson distribution by full maximum likelihood estimation.

Usage

```

zipoisson(lpstr0 = "logitlink", llambda = "loglink", type.fitted =
  c("mean", "lambda", "pobs0", "pstr0", "onempstr0"),
  ipstr0 = NULL, ilambda = NULL, gpstr0 = NULL, imethod = 1,
  ishrinkage = 0.95, probs.y = 0.35, parallel = FALSE, zero = NULL)
zipoissonff(llambda = "loglink", lonempstr0 = "logitlink",
  type.fitted = c("mean", "lambda", "pobs0", "pstr0", "onempstr0"),
  ilambda = NULL, ionempstr0 = NULL, gonempstr0 = NULL,
  imethod = 1, ishrinkage = 0.95, probs.y = 0.35, zero = "onempstr0")

```

Arguments

<code>lpstr0, llambda</code>	Link function for the parameter ϕ and the usual λ parameter. See Links for more choices; see CommonVGAMffArguments for more information. For the zero-deflated model see below.
<code>ipstr0, ilambda</code>	Optional initial values for ϕ , whose values must lie between 0 and 1. Optional initial values for λ , whose values must be positive. The defaults are to compute an initial value internally for each. If a vector then recycling is used.
<code>lonempstr0, ionempstr0</code>	Corresponding arguments for the other parameterization. See details below.
<code>type.fitted</code>	Character. The type of fitted value to be returned. The first choice (the expected value) is the default. The estimated probability of an observed 0 is an alternative, else the estimated probability of a structural 0, or one minus the estimated probability of a structural 0. See CommonVGAMffArguments and fittedvlm for more information.
<code>imethod</code>	An integer with value 1 or 2 which specifies the initialization method for λ . If failure to converge occurs try another value and/or else specify a value for <code>ishrinkage</code> and/or else specify a value for <code>ipstr0</code> . See CommonVGAMffArguments for more information.
<code>ishrinkage</code>	How much shrinkage is used when initializing λ . The value must be between 0 and 1 inclusive, and a value of 0 means the individual response values are used, and a value of 1 means the median or mean is used. This argument is used in conjunction with <code>imethod</code> . See CommonVGAMffArguments for more information.
<code>zero</code>	Specifies which linear/additive predictors are to be modelled as intercept-only. If given, the value can be either 1 or 2, and the default is none of them. Setting <code>zero = 1</code> makes ϕ a single parameter. See CommonVGAMffArguments for more information.
<code>gpstr0, gonempstr0, probs.y</code>	Details at CommonVGAMffArguments .
<code>parallel</code>	Details at CommonVGAMffArguments , but unlikely to be practically used actually.

Details

These models are a mixture of a Poisson distribution and the value 0; it has value 0 with probability ϕ else is $\text{Poisson}(\lambda)$ distributed. Thus there are two sources for zero values, and ϕ is the probability of a *structural zero*. The model for `zipoisson()` can be written

$$P(Y = 0) = \phi + (1 - \phi) \exp(-\lambda),$$

and for $y = 1, 2, \dots$,

$$P(Y = y) = (1 - \phi) \exp(-\lambda) \lambda^y / y!.$$

Here, the parameter ϕ satisfies $0 < \phi < 1$. The mean of Y is $(1 - \phi)\lambda$ and these are returned as the fitted values, by default. The variance of Y is $(1 - \phi)\lambda(1 + \phi\lambda)$. By default, the two linear/additive predictors of `zipoisson()` are $(\text{logit}(\phi), \log(\lambda))^T$.

The **VGAM** family function `zipoissonff()` has a few changes compared to `zipoisson()`. These are: (i) the order of the linear/additive predictors is switched so the Poisson mean comes first;

(ii) `onempstr0` is now 1 minus the probability of a structural 0, i.e., the probability of the parent (Poisson) component, i.e., `onempstr0` is $1 - \text{pstr0}$; (iii) argument `zero` has a new default so that the `onempstr0` is intercept-only by default. Now `zipoissonff()` is generally recommended over `zipoisson()` (and definitely recommended over [yip88](#)). Both functions implement Fisher scoring and can handle multiple responses.

Both family functions can fit the zero-*modified* Poisson (ZMP), which is a combination of the ZIP and zero-*deflated* Poisson (ZDP); see [Zipois](#) for some details and the example below. The key is to set the link function to be `identitylink`. However, problems might occur when iterations get close to or go past the boundary of the parameter space, especially when there are covariates. The PMF of the ZMP is best written not as above but in terms of `onempstr0` which may be greater than unity; when using `pstr0` the above PMF is negative for non-zero values.

Value

An object of class "vglmff" (see [vglmff-class](#)). The object is used by modelling functions such as [vglm](#), [rrvglm](#) and [vgam](#).

Warning

Numerical problems can occur, e.g., when the probability of zero is actually less than, not more than, the nominal probability of zero. For example, in the Angers and Biswas (2003) data below, replacing 182 by 1 results in nonconvergence. Half-stepping is not uncommon. If failure to converge occurs, try using combinations of `imethod`, `ishrinkage`, `ipstr0`, and/or `zipoisson(zero = 1)` if there are explanatory variables. The default for `zipoissonff()` is to model the structural zero probability as an intercept-only.

Note

This family function can be used to estimate the 0-*deflated* model, hence `pstr0` is not to be interpreted as a probability. One should set, e.g., `lpstr0 = "identitylink"`. Likewise, the functions in [Zipois](#) can handle the zero-deflated Poisson distribution too. Although the iterations might fall outside the parameter space, the `validparams` slot should keep them inside. A (somewhat) similar alternative for zero-deflation is to try the zero-altered Poisson model (see [zapoisson](#)).

The use of this **VGAM** family function with [rrvglm](#) can result in a so-called COZIGAM or COZIGLM. That is, a reduced-rank zero-inflated Poisson model (RR-ZIP) is a constrained zero-inflated generalized linear model. See what used to be **COZIGAM** on CRAN. A RR-ZINB model can also be fitted easily; see [zinegbinomial](#). Jargon-wise, a COZIGLM might be better described as COZIVGLM-ZIP.

Author(s)

T. W. Yee

References

- Thas, O. and Rayner, J. C. W. (2005). Smooth tests for the zero-inflated Poisson distribution. *Biometrics*, **61**, 808–815.
- Data: Angers, J-F. and Biswas, A. (2003). A Bayesian analysis of zero-inflated generalized Poisson model. *Computational Statistics & Data Analysis*, **42**, 37–46.

Cameron, A. C. and Trivedi, P. K. (1998). *Regression Analysis of Count Data*. Cambridge University Press: Cambridge.

M'Kendrick, A. G. (1925). Applications of mathematics to medical problems. *Proc. Edinb. Math. Soc.*, **44**, 98–130.

Yee, T. W. (2014). Reduced-rank vector generalized linear models with two linear predictors. *Computational Statistics and Data Analysis*, **71**, 889–902.

See Also

[gaitdpoisson](#), [zipoisson](#), [Zipois](#), [yip88](#), [spikeplot](#), [lpossums](#), [rrvglm](#), [negbinomial](#), [zipebcom](#), [rpois](#), [simulate.vlm](#), [hdeff.vglm](#).

Examples

```
# Example 1: simulated ZIP data
zdata <- data.frame(x2 = runif(nn <- 1000))
zdata <- transform(zdata,
  pstr01 = logitlink(-0.5 + 1*x2, inverse = TRUE),
  pstr02 = logitlink( 0.5 - 1*x2, inverse = TRUE),
  Ps01   = logitlink(-0.5      , inverse = TRUE),
  Ps02   = logitlink( 0.5      , inverse = TRUE),
  lambda1 = loglink(-0.5 + 2*x2, inverse = TRUE),
  lambda2 = loglink( 0.5 + 2*x2, inverse = TRUE))
zdata <- transform(zdata, y1 = rzipois(nn, lambda1, pstr0 = Ps01),
  y2 = rzipois(nn, lambda2, pstr0 = Ps02))

with(zdata, table(y1)) # Eyeball the data
with(zdata, table(y2))
fit1 <- vglm(y1 ~ x2, zipoisson(zero = 1), zdata, crit = "coef")
fit2 <- vglm(y2 ~ x2, zipoisson(zero = 1), zdata, crit = "coef")
coef(fit1, matrix = TRUE) # Should agree with the above values
coef(fit2, matrix = TRUE) # Should agree with the above values

# Fit all two simultaneously, using a different parameterization:
fit12 <- vglm(cbind(y1, y2) ~ x2, zipoissonff, zdata, crit = "coef")
coef(fit12, matrix = TRUE) # Should agree with the above values

# For the first observation compute the probability that y1 is
# due to a structural zero.
(fitted(fit1, type = "pstr0") / fitted(fit1, type = "pobs0"))[1]

# Example 2: McKendrick (1925). From 223 Indian village households
cholera <- data.frame(ncases = 0:4, # Number of cholera cases,
  wfreq = c(168, 32, 16, 6, 1)) # Frequencies
fit <- vglm(ncases ~ 1, zipoisson, wei = wfreq, cholera)
coef(fit, matrix = TRUE)
with(cholera, cbind(actual = wfreq,
  fitted = round(dzipois(ncases, Coef(fit)[2],
    pstr0 = Coef(fit)[1]) *
    sum(wfreq), digits = 2)))
```



```

# Example 3: data from Angers and Biswas (2003)
abdata <- data.frame(y = 0:7, w = c(182, 41, 12, 2, 2, 0, 0, 1))
abdata <- subset(abdata, w > 0)
fit3 <- vglm(y ~ 1, zipoisson(lpstr0 = probitlink, ipstr0 = 0.8),
            data = abdata, weight = w, trace = TRUE)
fitted(fit3, type = "pobs0") # Estimate of P(Y = 0)
coef(fit3, matrix = TRUE)
Coef(fit3) # Estimate of pstr0 and lambda
fitted(fit3)
with(abdata, weighted.mean(y, w)) # Compare this with fitted(fit)
summary(fit3)

# Example 4: zero-deflated (ZDP) model for intercept-only data
zdata <- transform(zdata, lambda3 = loglink(0.0, inverse = TRUE))
zdata <- transform(zdata, deflat.limit=-1/expm1(lambda3)) # Bndy
# The 'pstr0' parameter is negative and in parameter space:
# Not too near the boundary:
zdata <- transform(zdata, usepstr0 = deflat.limit / 2)
zdata <- transform(zdata,
                  y3 = rzipois(nn, lambda3, pstr0 = usepstr0))
head(zdata)
with(zdata, table(y3)) # A lot of deflation
fit4 <- vglm(y3 ~ 1, data = zdata, trace = TRUE, crit = "coef",
            zipoisson(lpstr0 = "identitylink"))
coef(fit4, matrix = TRUE)
# Check how accurate it was:
zdata[1, "usepstr0"] # Answer
coef(fit4)[1] # Estimate
Coef(fit4)
vcov(fit4) # Is positive-definite

# Example 5: RR-ZIP
set.seed(123)
rrzip <- rrvglm(Alopacce ~ sm.bs(WaterCon, df = 3),
               zipoisson(zero = NULL),
               data = hspider, trace = TRUE, Index.corner = 2)
coef(rrzip, matrix = TRUE)
Coef(rrzip)
summary(rrzip)
## Not run: plotvgam(rrzip, lcol = "blue")

```

Description

Density, distribution function, and random generation for the zero/one-inflated beta distribution.

Usage

```

dzoabeta(x, shape1, shape2, pobs0 = 0, pobs1 = 0, log = FALSE,
         tol = .Machine$double.eps)
pzoabeta(q, shape1, shape2, pobs0 = 0, pobs1 = 0,
         lower.tail = TRUE, log.p = FALSE, tol = .Machine$double.eps)
qzoabeta(p, shape1, shape2, pobs0 = 0, pobs1 = 0,
         lower.tail = TRUE, log.p = FALSE, tol = .Machine$double.eps)
rzoabeta(n, shape1, shape2, pobs0 = 0, pobs1 = 0,
         tol = .Machine$double.eps)

```

Arguments

`x, q, p, n` Same as [Beta](#).

`pobs0, pobs1` vector of probabilities that 0 and 1 are observed (ω_0 and ω_1).

`shape1, shape2` Same as [Beta](#). They are called a and b in [beta](#) respectively.

`lower.tail, log, log.p`
Same as [Beta](#).

`tol` Numeric, tolerance for testing equality with 0 and 1.

Details

This distribution is a mixture of a discrete distribution with a continuous distribution. The cumulative distribution function of Y is

$$F(y) = (1 - \omega_0 - \omega_1)B(y) + \omega_0 \times I[0 \leq y] + \omega_1 \times I[1 \leq y]$$

where $B(y)$ is the cumulative distribution function of the beta distribution with the same shape parameters ([pbeta](#)), ω_0 is the inflated probability at 0 and ω_1 is the inflated probability at 1. The default values of ω_j mean that these functions behave like the ordinary [Beta](#) when only the essential arguments are inputted.

Value

`dzoabeta` gives the density, `pzoabeta` gives the distribution function, `qzoabeta` gives the quantile, and `rzoabeta` generates random deviates.

Author(s)

Xiangjie Xue and T. W. Yee

See Also

[zoabetaR](#), [beta](#), [betaR](#), [Betabinom](#).

Examples

```
## Not run:
N <- 1000; y <- rzoabeta(N, 2, 3, 0.2, 0.2)
hist(y, probability = TRUE, border = "blue", las = 1,
     main = "Blue = 0- and 1-altered; orange = ordinary beta")
sum(y == 0) / N # Proportion of 0s
sum(y == 1) / N # Proportion of 1s
Ngrid <- 1000
lines(seq(0, 1, length = Ngrid),
      dbeta(seq(0, 1, length = Ngrid), 2, 3), col = "orange")
lines(seq(0, 1, length = Ngrid), col = "blue",
      dzoabeta(seq(0, 1, length = Ngrid), 2, 3, 0.2, 0.2))

## End(Not run)
```

zoabetaR

*Zero- and One-Inflated Beta Distribution Family Function***Description**

Estimation of the shape parameters of the two-parameter beta distribution plus the probabilities of a 0 and/or a 1.

Usage

```
zoabetaR(lshape1 = "loglink", lshape2 = "loglink", lpobs0 = "logitlink",
         lpobs1 = "logitlink", ishape1 = NULL, ishape2 = NULL, trim = 0.05,
         type.fitted = c("mean", "pobs0", "pobs1", "beta.mean"),
         parallel.shape = FALSE, parallel.pobs = FALSE, zero = NULL)
```

Arguments

`lshape1`, `lshape2`, `lpobs0`, `lpobs1`
 Details at [CommonVGAMffArguments](#). See [Links](#) for more choices.

`ishape1`, `ishape2`
 Details at [CommonVGAMffArguments](#).

`trim`, `zero` Same as `betaR`. See [CommonVGAMffArguments](#) for information.

`parallel.shape`, `parallel.pobs`
 See [CommonVGAMffArguments](#) for more information.

`type.fitted` The choice "beta.mean" mean to return the mean of the beta distribution; the 0s and 1s are ignored. See [CommonVGAMffArguments](#) for more information.

Details

The standard 2-parameter beta distribution has a support on (0,1), however, many datasets have 0 and/or 1 values too. This family function handles 0s and 1s (at least one of them must be present) in the data set by modelling the probability of a 0 by a logistic regression (default link is the logit), and similarly for the probability of a 1. The remaining proportion, $1 - \text{pobs0} - \text{pobs1}$, of the data comes from a standard beta distribution. This family function therefore extends [betaR](#). One has $M = 3$ or $M = 4$ per response. Multiple responses are allowed.

Value

Similar to [betaR](#).

Author(s)

Thomas W. Yee and Xiangjie Xue.

See Also

[Zoabeta](#), [betaR](#), [betaff](#), [Beta](#), [zipoisson](#).

Examples

```
nn <- 1000; set.seed(1)
bdata <- data.frame(x2 = runif(nn))
bdata <- transform(bdata,
  pobs0 = logitlink(-2 + x2, inverse = TRUE),
  pobs1 = logitlink(-2 + x2, inverse = TRUE))
bdata <- transform(bdata,
  y1 = rzoabeta(nn, shape1 = exp(1 + x2), shape2 = exp(2 - x2),
    pobs0 = pobs0, pobs1 = pobs1))
summary(bdata)
fit1 <- vglm(y1 ~ x2, zoabetaR(parallel.pobs = TRUE),
  data = bdata, trace = TRUE)
coef(fit1, matrix = TRUE)
summary(fit1)
```

Index

- * **Constrained additive ordination**
 - cao, 145
 - cao.control, 149
- * **Constraint matrices**
 - CM.equid, 176
 - constraints, 201
- * **DFBETAs**
 - hatvalues, 410
- * **Hauck–Donner effect**
 - hdeff, 412
 - hdeffsev, 416
- * **Penalized regression**
 - vgam, 843
- * **Reduced-Rank Vector Generalized Linear Model**
 - rrvglm, 730
 - rrvglm-class, 734
 - rrvglm.control, 736
 - rrvglm.optim.control, 739
- * **Reduced-rank regression**
 - rrvglm, 730
- * **Vector Generalized Additive Model**
 - vgam, 843
 - vgam-class, 848
 - vgam.control, 850
- * **Vector Generalized Linear Model**
 - vglm, 853
 - vglm-class, 859
 - vglm.control, 862
 - vglmff-class, 866
- * **classes**
 - calibrate-methods, 137
 - Coef.qrrvglm-class, 181
 - Coef.rrvglm-class, 184
 - concoef-methods, 198
 - rrvglm-class, 734
 - SurvS4-class, 804
 - vgam-class, 848
 - vglm-class, 859
 - vglmff-class, 866
- * **datagen**
 - posbernUC, 657
 - rcqo, 706
 - simulate.vlm, 756
- * **datasets**
 - auuc, 51
 - backPain, 53
 - beggs, 54
 - bmi.nz, 125
 - budworm, 134
 - cfibrosis, 167
 - chest.nz, 169
 - chinese.nz, 170
 - coalminers, 177
 - corbet, 204
 - crashes, 212
 - deermice, 225
 - ducklings, 250
 - enzyme, 251
 - finney44, 293
 - flourbeetle, 301
 - gew, 383
 - grain.us, 394
 - hormone, 418
 - hspider, 420
 - Huggins89.t1, 423
 - hunua, 425
 - lake0, 460
 - leukemia, 471
 - lirat, 487
 - lpossums, 521
 - machinists, 536
 - marital.nz, 542
 - melbmaxtemp, 550
 - olympics, 601
 - oxtemp, 607
 - pneumo, 641
 - prats, 671

- prinia, 678
- ruge, 741
- toxop, 817
- ucberk, 833
- V1, 838
- V2, 839
- venice, 842
- waitakere, 874
- wine, 885
- * **distribution**
 - Benford, 56
 - Benini, 57
 - Betabinom, 60
 - Betageom, 72
 - Betanorm, 76
 - Biamhcop, 81
 - Biclaytoncop, 84
 - Bifgmcop, 88
 - bilogis, 95
 - Binom2.or, 98
 - Binom2.rho, 103
 - Binorm, 110
 - Binormcop, 115
 - Biplackett, 116
 - Bisa, 119
 - Bistudentt, 123
 - Bort, 128
 - Card, 152
 - Dagum, 220
 - dAR1, 223
 - dextlogF, 229
 - dgaitdplot, 231
 - dhuber, 234
 - Diffzeta, 236
 - dlogF, 244
 - Expectiles-Exponential, 255
 - Expectiles-Normal, 256
 - Expectiles-sc.t2, 258
 - Expectiles-Uniform, 259
 - expgeom, 266
 - explog, 271
 - exppois, 275
 - Extbetabinom, 278
 - Felix, 286
 - Fisk, 295
 - Foldnorm, 302
 - Frank, 306
 - Frechet, 308
 - Gaitdbinom, 313
 - Gaitdlog, 316
 - Gaitdbinom, 321
 - Gaitdpois, 327
 - Gaitdzeta, 337
 - GenbetaII, 353
 - gengammaUC, 358
 - Genpois0, 360
 - Genpois1, 362
 - genray, 369
 - Gensh, 372
 - gevUC, 381
 - Gompertz, 386
 - gpdUC, 392
 - Gumbel-II, 403
 - gumbelUC, 406
 - Hurea, 427
 - Hzeta, 433
 - Inv.gaussian, 441
 - Inv.lomax, 444
 - Inv.paralogistic, 447
 - Kumar, 458
 - laplaceUC, 465
 - lgammaUC, 475
 - Lindley, 476
 - Lino, 483
 - Log, 496
 - Lomax, 518
 - Makeham, 537
 - Maxwell, 545
 - N1binom, 576
 - N1pois, 579
 - Nakagami, 582
 - Paralogistic, 608
 - Pareto, 611
 - ParetoIV, 614
 - Perks, 619
 - PoissonPoints, 645
 - Polono, 646
 - posbernUC, 657
 - Posgeom, 661
 - Posnorm, 665
 - Rayleigh, 701
 - rcqo, 706
 - rdiric, 710
 - Rice, 719
 - Simplex, 754
 - Sinmad, 758

- Skellam, 761
- skewnorm, 764
- Slash, 767
- spikeplot, 782
- Tobit, 806
- Topple, 814
- Triangle, 818
- Trinorm, 821
- Truncpareto, 830
- UtilitiesVGAM, 836
- Yules, 888
- Zabinom, 891
- Zageom, 894
- Zanegbin, 897
- Zapois, 902
- Zeta, 907
- Zibinom, 912
- Zigeom, 916
- Zinegbin, 919
- Zipf, 927
- Zipfmb, 930
- Zipois, 931
- Zoabeta, 937
- * dplot**
 - dgaitdplot, 231
 - plotvgam.control, 638
 - prplot, 683
- * hplot**
 - biplot-methods, 119
 - deplot.lmscreg, 226
 - dgaitdplot, 231
 - guplot, 408
 - lvplot, 527
 - lvplot.qrrvglm, 529
 - lvplot.rrvglm, 533
 - meplot, 551
 - plotdeplot.lmscreg, 627
 - plotdgaitd.vglm, 629
 - plotqrrvglm, 630
 - plotqtplot.lmscreg, 631
 - plotrcim0, 633
 - plotvgam, 635
 - plotvglm, 639
 - prplot, 683
 - qtplot.gumbel, 691
 - qtplot.lmscreg, 693
 - rlplot.gevff, 723
 - spikeplot, 782
 - trplot, 824
 - trplot.qrrvglm, 826
 - vplot.profile, 871
- * htest**
 - anova.vglm, 40
 - cops, 203
 - hdeff, 412
 - hdeffsev, 416
 - lrt.stat, 524
 - lrtest, 526
 - score.stat, 745
 - wald.stat, 876
- * manip**
 - iam, 435
- * math**
 - alogitlink, 27
 - asinlink, 49
 - bell, 55
 - cauchitlink, 156
 - clogloglink, 174
 - erf, 252
 - expint, 268
 - explink, 270
 - fisherzlink, 294
 - identitylink, 437
 - kendall.tau, 455
 - lambertW, 462
 - lerch, 469
 - logclink, 498
 - logitlink, 504
 - logitoffsetlink, 507
 - loglink, 512
 - logloglink, 514
 - logofflink, 517
 - mills.ratio, 555
 - multilogitlink, 570
 - nbcanlink, 585
 - ordpoisson, 603
 - pgamma.deriv, 624
 - pgamma.deriv.unscaled, 626
 - powerlink, 670
 - probitlink, 679
 - reciprocallink, 714
 - rhobitlink, 717
 - round2, 727
 - sloglink, 770
 - sqrtlink, 784
 - zeta, 908

*** methods**

calibrate-methods, 137
 concoef-methods, 198

*** models**

A1A2A3, 17
 AA.Aa.aa, 18
 AB.Ab.aB.ab, 20
 ABO, 21
 acat, 22
 add1.vglm, 24
 AICvlm, 25
 alogitlink, 27
 altered, 29
 amlbinomial, 31
 amlexponential, 33
 amlnormal, 35
 amlpoisson, 38
 AR1, 42
 asinlink, 49
 aux.posbernoulli.t, 52
 benini1, 59
 betabinomial, 64
 betabinomialff, 67
 betaff, 70
 betageometric, 73
 betaII, 75
 betaprime, 78
 betaR, 79
 biamhcop, 82
 biclaytoncop, 85
 BICvlm, 87
 bifgmcop, 89
 bifgmexp, 90
 bifrankcop, 92
 bigumbelIexp, 93
 bilogistic, 96
 binom2.or, 100
 binom2.rho, 105
 binomialff, 107
 binormal, 111
 binormalcop, 113
 biplackettcop, 117
 bisa, 121
 bistudentt, 124
 borel.tanner, 126
 Brat, 129
 brat, 130
 bratt, 132

calibrate, 135
 calibrate-methods, 137
 calibrate.qrrvglm, 137
 calibrate.qrrvglm.control, 141
 calibrate.rrvglm, 142
 calibrate.rrvglm.control, 144
 cao, 145
 cao.control, 149
 cardioid, 154
 cauchitlink, 156
 cauchy, 158
 cdf.lmscreg, 160
 cens.gumbel, 161
 cens.normal, 163
 cens.poisson, 165
 cgo, 168
 chisq, 171
 clo, 172
 clogloglink, 174
 CM.equid, 176
 Coef, 178
 Coef.qrrvglm, 179
 Coef.rrvglm, 183
 Coef.vlm, 185
 coefvgam, 186
 coefvlm, 187
 CommonVGAMffArguments, 188
 concoef, 197
 concoef-methods, 198
 confintvglm, 199
 constraints, 201
 cops, 203
 cqo, 205
 cratio, 214
 cumulative, 216
 dagum, 222
 depvar, 228
 df.residual, 230
 diffzeta, 237
 dirichlet, 238
 dirmul.old, 240
 dirmultinomial, 242
 double.cens.normal, 245
 double.expbinoial, 247
 eCDF, 250
 erlang, 253
 expexpff, 262
 expexpff1, 264

expgeometric, 267
explink, 270
explogff, 272
exponential, 274
exppoisson, 277
extbetabinomial, 280
extlogF1, 282
familyname, 285
felix, 287
fff, 288
fill1, 290
fisherzlink, 294
fisk, 296
fittedv1m, 298
fix.crossing, 300
foldnormal, 303
formulav1m, 305
frechet, 309
freund61, 311
gaitdlog, 318
gaitdnbinomial, 323
gaitdpoisson, 331
gaitdzeta, 340
gamma1, 342
gamma2, 344
gammaff.mm, 346
gammahyperbola, 348
gammaR, 349
garma, 351
genbetaII, 354
gengamma.stacy, 356
genpoisson0, 363
genpoisson1, 366
genpoisson2, 367
genrayleigh, 370
gensh, 373
geometric, 375
get.smart, 376
get.smart.prediction, 377
gev, 378
gompertz, 387
gpd, 389
grc, 395
gumbel, 400
gumbelIII, 404
has.interceptv1m, 409
hdeff, 412
hdeffsev, 416
huber2, 422
hurea, 428
hyperg, 430
hypersecant, 431
hzeta, 434
identitylink, 437
Influence, 438
inv.binomial, 439
inv.gaussianff, 442
inv.lomax, 445
inv.paralogistic, 448
is.buggy, 450
is.crossing, 451
is.parallel, 452
is.smart, 453
is.zero, 454
KLD, 456
kumar, 459
laplace, 463
latvar, 466
leipnik, 468
levy, 472
lgamma1, 473
lindley, 478
linkfun, 479
Links, 480
lino, 485
lms.bcg, 488
lms.bcn, 490
lms.yjn, 493
logclink, 498
logF, 499
logff, 501
logistic, 502
logitlink, 504
logitoffsetlink, 507
loglik.v1m, 508
loglinb2, 509
loglinb3, 511
loglink, 512
logloglink, 514
lognormal, 515
logofflink, 517
lomax, 520
lqnorm, 522
lrt.stat, 524
makeham, 538
margeff, 540

- Max, 543
- maxwell, 546
- mccullagh89, 547
- meangaitd, 549
- micmen, 553
- mix2exp, 556
- mix2normal, 558
- mix2poisson, 560
- MNSs, 562
- model.framevln, 564
- model.matrixqrrvglm, 565
- model.matrixvln, 566
- multilogitlink, 570
- multinomial, 571
- N1binomial, 577
- N1poisson, 580
- nakagami, 583
- nbcanlink, 585
- negbinomial, 587
- negbinomial.size, 594
- normal.vcm, 596
- nparam.vln, 599
- Opt, 602
- ordpoisson, 603
- ordsup, 606
- paralogistic, 609
- paretoff, 612
- paretoIV, 616
- perks, 620
- perspqrrvglm, 622
- plotvgam, 635
- plotvglm, 639
- poisson.points, 642
- poissonff, 643
- posbernoulli.b, 648
- posbernoulli.t, 651
- posbernoulli.tb, 654
- posbinomial, 659
- posnegbinomial, 662
- posnormal, 667
- pospoisson, 669
- powerlink, 670
- predictqrrvglm, 673
- predictvglm, 674
- prentice74, 676
- probitlink, 679
- profilevglm, 681
- propodds, 682
- put.smart, 685
- qrrvglm.control, 686
- Qvar, 694
- qvar, 698
- R2latvar, 699
- Rank, 700
- rayleigh, 702
- rec.exp1, 711
- rec.normal, 712
- reciprocallink, 714
- residualsvglm, 715
- rhobitlink, 717
- riceff, 720
- rigff, 722
- rootogram4, 725
- rrar, 728
- rrvglm, 730
- rrvglm.control, 736
- rrvglm.optim.control, 739
- s, 742
- sc.studentt2, 744
- score.stat, 745
- seglines, 747
- Select, 748
- seq2binomial, 751
- setup.smart, 752
- simplex, 755
- simulate.vln, 756
- sinmad, 759
- skellam, 763
- skewnormal, 766
- slash, 769
- sloglink, 770
- sm.os, 772
- sm.ps, 775
- smart.expression, 778
- smart.mode.is, 778
- smartpred, 779
- specials, 781
- sqrtlink, 784
- sratio, 786
- step4, 789
- studentt, 790
- summary.drrvglm, 792
- summarypvglm, 795
- summaryvgam, 796
- summaryvglm, 797
- SURff, 800

- TIC, 805
- tobit, 808
- Tol, 812
- topple, 815
- trim.constraints, 819
- trinormal, 823
- Trunc, 828
- truncweibull, 831
- uninormal, 834
- UtilitiesVGAM, 836
- vcovvml, 840
- vgam, 843
- vgam-class, 848
- VGAM-package, 13
- vgam.control, 850
- vglm, 853
- vglm.control, 862
- vonnises, 869
- wald.stat, 876
- waldff, 878
- weibull.mean, 879
- weibullR, 881
- weightsvglm, 884
- wrapup.smart, 886
- yeo.johnson, 887
- yulesimon, 890
- zabinomial, 892
- zageometric, 895
- zanegbinomial, 899
- zapoisson, 903
- zero, 905
- zetaff, 910
- zibinomial, 913
- zigeometric, 917
- zinegbinomial, 921
- zipebcom, 924
- zipf, 928
- zipoisson, 933
- * **nonlinear**
 - calibrate.qrrvglm, 137
 - calibrate.qrrvglm.control, 141
 - Coef.qrrvglm, 179
 - Coef.qrrvglm-class, 181
 - lvplot.qrrvglm, 529
 - model.matrixqrrvglm, 565
 - persprrvglm, 622
 - plotqrrvglm, 630
 - predictqrrvglm, 673
 - qrrvglm.control, 686
 - trplot.qrrvglm, 826
- * **optimize**
 - calibrate.qrrvglm.control, 141
 - calibrate.rrvglm.control, 144
 - cao.control, 149
 - qrrvglm.control, 686
 - rrvglm.control, 736
 - rrvglm.optim.control, 739
 - vgam.control, 850
 - vglm.control, 862
- * **package**
 - VGAM-package, 13
- * **programming**
 - get.smart, 376
 - get.smart.prediction, 377
 - iam, 435
 - is.smart, 453
 - put.smart, 685
 - setup.smart, 752
 - smart.expression, 778
 - smart.mode.is, 778
 - smartpred, 779
 - UtilitiesVGAM, 836
 - wrapup.smart, 886
 - zero, 905
- * **regression**
 - A1A2A3, 17
 - AA.Aa.aa, 18
 - AB.Ab.aB.ab, 20
 - ABO, 21
 - acat, 22
 - AICvml, 25
 - alogitlink, 27
 - amlbinomial, 31
 - amlexponential, 33
 - amlnormal, 35
 - amlpoisson, 38
 - AR1, 42
 - asinlink, 49
 - aux.posbernoulli.t, 52
 - benini1, 59
 - betabinomial, 64
 - betabinomialff, 67
 - betaff, 70
 - betageometric, 73
 - betaII, 75
 - betaprime, 78

- betaR, 79
- biamhcop, 82
- biclaytoncop, 85
- BICv1m, 87
- bifgmcop, 89
- bifgmexp, 90
- bifrankcop, 92
- bigumbelIexp, 93
- bilogistic, 96
- binom2.or, 100
- binom2.rho, 105
- binomialff, 107
- binormal, 111
- binormalcop, 113
- biplackettcop, 117
- biplot-methods, 119
- bisa, 121
- bistudentt, 124
- borel.tanner, 126
- Brat, 129
- brat, 130
- bratt, 132
- calibrate, 135
- calibrate-methods, 137
- calibrate.qrrvglm, 137
- calibrate.qrrvglm.control, 141
- calibrate.rrvglm, 142
- calibrate.rrvglm.control, 144
- cao, 145
- cao.control, 149
- cardioid, 154
- cauchitlink, 156
- cauchy, 158
- cdf.lmscreg, 160
- cens.gumbel, 161
- cens.normal, 163
- cens.poisson, 165
- cgo, 168
- chisq, 171
- clo, 172
- clogloglink, 174
- Coef, 178
- Coef.qrrvglm, 179
- Coef.rrvglm, 183
- Coef.v1m, 185
- coefvgam, 186
- coefv1m, 187
- concoef, 197
- concoef-methods, 198
- confintvglm, 199
- constraints, 201
- cops, 203
- cqo, 205
- cratio, 214
- cumulative, 216
- dagum, 222
- devar, 228
- df.residual, 230
- diffzeta, 237
- dirichlet, 238
- dirmul.old, 240
- dirmultinomial, 242
- double.cens.normal, 245
- double.expbinomial, 247
- eCDF, 250
- erlang, 253
- expexpff, 262
- expexpff1, 264
- expgeometric, 267
- explink, 270
- explogff, 272
- exponential, 274
- exppoisson, 277
- extbetabinomial, 280
- extlogF1, 282
- familyname, 285
- felix, 287
- fff, 288
- fill1, 290
- fisherzlink, 294
- fisk, 296
- fittedv1m, 298
- fix.crossing, 300
- foldnormal, 303
- formulav1m, 305
- frechet, 309
- freund61, 311
- gaitdlog, 318
- gaitdnbinomial, 323
- gaitdpoisson, 331
- gaitdzeta, 340
- gamma1, 342
- gamma2, 344
- gammaff.mm, 346
- gammahyperbola, 348
- gammaR, 349

garma, 351
genbetaII, 354
gengamma.stacy, 356
genpoisson0, 363
genpoisson1, 366
genpoisson2, 367
genrayleigh, 370
gensh, 373
geometric, 375
get.smart, 376
get.smart.prediction, 377
gev, 378
gompertz, 387
gpd, 389
grc, 395
gumbel, 400
gumbelII, 404
guplot, 408
has.interceptv1m, 409
hatvalues, 410
hdeff, 412
hdeffsev, 416
huber2, 422
hurea, 428
hyperg, 430
hypersecant, 431
hzeta, 434
identitylink, 437
inv.binomial, 439
inv.gaussianff, 442
inv.lomax, 445
inv.paralogistic, 448
is.buggy, 450
is.crossing, 451
is.parallel, 452
is.smart, 453
is.zero, 454
KLD, 456
kumar, 459
laplace, 463
latvar, 466
leipnik, 468
levy, 472
lgamma1, 473
lindley, 478
linkfun, 479
lino, 485
lms.bcg, 488
lms.bcn, 490
lms.yjn, 493
logclink, 498
logF, 499
logff, 501
logistic, 502
logitlink, 504
logitoffsetlink, 507
logLik.vlm, 508
loglinb2, 509
loglinb3, 511
loglink, 512
logloglink, 514
lognormal, 515
logofflink, 517
lomax, 520
lqnorm, 522
lrt.stat, 524
lvplot, 527
lvplot.qrrvglm, 529
lvplot.rrvglm, 533
makeham, 538
margeff, 540
Max, 543
maxwell, 546
mccullagh89, 547
meangaitd, 549
meplot, 551
micmen, 553
mix2exp, 556
mix2normal, 558
mix2poisson, 560
MNSs, 562
multilogitlink, 570
multinomial, 571
N1binomial, 577
N1poisson, 580
nakagami, 583
nbcnlink, 585
negbinomial, 587
negbinomial.size, 594
normal.vcm, 596
nparam.vlm, 599
Opt, 602
ordpoisson, 603
ordsup, 606
paralogistic, 609
paretoff, 612

paretoIV, 616
 perks, 620
 persprrvglm, 622
 plotdeplot.lmscreg, 627
 plotdgaitd.vglm, 629
 plotqrrvglm, 630
 plotqtplot.lmscreg, 631
 plotrcim0, 633
 plotvgam, 635
 plotvgam.control, 638
 plotvglm, 639
 poisson.points, 642
 poissonff, 643
 posbernoulli.b, 648
 posbernoulli.t, 651
 posbernoulli.tb, 654
 posbinomial, 659
 posnegbinomial, 662
 posnormal, 667
 pospoisson, 669
 powerlink, 670
 predictqrrvglm, 673
 predictvglm, 674
 prentice74, 676
 probitlink, 679
 profilevglm, 681
 propodds, 682
 prplot, 683
 put.smart, 685
 qrrvglm.control, 686
 qtplot.gumbel, 691
 qtplot.lmscreg, 693
 Qvar, 694
 qvar, 698
 R2latvar, 699
 Rank, 700
 rayleigh, 702
 rec.exp1, 711
 rec.normal, 712
 reciprocallink, 714
 residualsvglm, 715
 rhobitlink, 717
 riceff, 720
 rigff, 722
 rlplot.gevff, 723
 rrar, 728
 rrvglm, 730
 rrvglm.optim.control, 739
 s, 742
 sc.studentt2, 744
 score.stat, 745
 seglines, 747
 Select, 748
 seq2binomial, 751
 setup.smart, 752
 simplex, 755
 sinmad, 759
 skellam, 763
 skewnormal, 766
 slash, 769
 sloglink, 770
 sm.os, 772
 sm.ps, 775
 smart.expression, 778
 smart.mode.is, 778
 smartpred, 779
 sqrtlink, 784
 sratio, 786
 studentt, 790
 summary.drrvglm, 792
 summarypvgam, 795
 summaryvgam, 796
 summaryvglm, 797
 SURff, 800
 TIC, 805
 tobit, 808
 Tol, 812
 topple, 815
 trim.constraints, 819
 trinormal, 823
 trplot, 824
 trplot.qrrvglm, 826
 Trunc, 828
 truncweibull, 831
 uninormal, 834
 UtilitiesVGAM, 836
 vcovvglm, 840
 vgam, 843
 vgam-class, 848
 VGAM-package, 13
 vglm, 853
 vonmises, 869
 vplot.profile, 871
 vsmooth.spline, 872
 wald.stat, 876
 waldff, 878

- weibull.mean, 879
- weibullR, 881
- weightsvglm, 884
- wrapup.smart, 886
- yeo.johnson, 887
- yulesimon, 890
- zabinomial, 892
- zageometric, 895
- zanegbinomial, 899
- zapoisson, 903
- zero, 905
- zetaff, 910
- zibinomial, 913
- zigeometric, 917
- zinegbinomial, 921
- zipebcom, 924
- zipf, 928
- zipoisson, 933
- zoabetaR, 939
- * **smooth**
 - plotvgam, 635
 - s, 742
 - sm.os, 772
 - sm.ps, 775
 - vgam, 843
 - vgam-class, 848
 - vsmooth.spline, 872
- * **survival**
 - SurvS4, 802
- * **ts**
 - rrar, 728
- * **utilities**
 - CM.equid, 176
 - model.matrixqrrvglm, 565
 - Trunc, 828
 - UtilitiesVGAM, 836
- .Machine, 361
- .Random.seed, 708
- [.SurvS4 (SurvS4), 802
- A1A2A3, 17, 19, 20, 22, 563
- AA.Aa.aa, 18, 18, 20, 22, 563
- AB.Ab.aB.ab, 18, 19, 20, 22, 563
- ABO, 18–20, 21, 563
- acat, 22, 177, 191, 193, 215, 218, 219, 452, 541, 542, 574, 716, 757, 788, 867
- add1.glm, 24, 25, 789
- add1.vglm, 24, 25, 41, 567, 789, 858
- AIC, 27, 88, 509, 806
- AICc, vglm-method (AICv1m), 25
- AICdrrvglm (AICv1m), 25
- AICqrrvglm (AICv1m), 25
- AICrrvgam (AICv1m), 25
- AICrrvglm (AICv1m), 25
- AICvgam (AICv1m), 25
- AICv1m, 25, 87, 88, 600, 653, 659, 660, 806, 858
- alaplace1, 32, 34, 37, 39, 283–285, 492, 757, 877
- alaplace2, 397, 464, 757
- alclevels (crashes), 212
- alcoff, 398, 568, 569
- alcoff (crashes), 212
- alogitlink, 27, 49, 109, 415, 505, 771, 784
- altered, 29, 782
- amlbinomial, 31, 34, 36, 37, 39, 109
- amlexponential, 32, 33, 34, 37, 275, 489
- amlnormal, 32, 34, 35, 39, 257, 492, 495
- amlpoisson, 32, 34, 36, 37, 38, 645
- anova, 526
- anova.glm, 40, 41
- anova.vglm, 25, 40, 230, 509, 525, 527, 746, 858, 877, 878
- AR1, 42, 46–48, 224, 835
- AR1EIM, 43, 44, 45
- arima.sim, 44
- array, 804
- as.character.SurvS4 (SurvS4), 802
- as.data.frame.SurvS4 (SurvS4), 802
- asinlink, 28, 49, 109, 415, 505
- auuc, 51, 398
- aux.posbernoulli.t, 52, 649
- backPain, 53
- backPain2, 25, 41, 789
- backPain2 (backPain), 53
- beggs, 54
- bell, 55, 462
- bellff, 55
- Benford, 56
- Benini, 57, 60
- benini1, 58, 59, 614
- Bessel, 870
- besselI, 721
- Beta, 61–63, 71, 72, 79, 81, 938, 940
- beta, 61, 65, 70, 72, 77, 80, 485, 500, 890, 938
- Betabinom, 60, 62, 66, 69, 279, 938

- betabinomial, *62, 63, 64, 68, 69, 109, 174, 243, 244, 280–282, 672, 757*
- betabinomialff, *62, 63, 66, 67, 71, 81, 241, 244, 281, 282, 672, 757*
- betaff, *66, 69, 70, 72, 74, 76, 79–81, 356, 460, 757, 940*
- Betageom, *72*
- betageometric, *71, 73, 81, 376*
- betaII, *71, 75, 81, 223, 298, 356, 446, 449, 521, 610, 761*
- Betanorm, *76*
- betaprime, *71, 78, 81*
- betaR, *71, 79, 757, 938–940*
- Biamhcop, *81*
- biamhcop, *81, 82, 82, 95, 96, 757*
- BIC, *88*
- Biclaytoncop, *84*
- biclaytoncop, *84, 85, 757*
- BICvgam (BICv1m), *87*
- BICv1m, *27, 87, 653, 659, 660, 806*
- Bifgmcop, *88*
- bifgmcop, *83, 88, 89, 89, 91, 93, 757*
- bifgmexp, *90, 90, 94*
- bifrankcop, *90, 92, 117, 119, 307, 757*
- bigamma.mckay, *345–347, 350*
- bigumbellexp, *83, 91, 93*
- bilogis, *95, 95*
- bilogistic, *96, 96, 504, 757*
- Binom2.or, *98*
- binom2.or, *99, 100, 101, 102, 106, 107, 510, 512, 906, 925, 926*
- Binom2.rho, *103*
- binom2.Rho (binom2.rho), *105*
- binom2.rho, *101, 102, 104, 105, 510, 718*
- Binomial, *279, 314, 913, 915*
- binomial, *108, 109*
- binomialff, *28, 31, 49, 65–69, 102, 107, 107, 131, 133, 138, 139, 143, 148, 175, 206, 210, 248, 281, 282, 302, 334, 412–414, 416, 417, 431, 578, 645, 649, 660, 690, 699, 709, 716, 752, 756, 757, 785, 846, 857, 893, 894, 897, 901, 904, 905, 914, 915*
- Binorm, *110*
- binormal, *84, 111, 111, 116, 125, 757, 822, 824, 835*
- binormalcop, *84, 113, 115, 116, 456, 578, 581, 757*
- Binormcop, *115*
- Biplackett, *116*
- biplackettcop, *117, 117*
- biplot, qrrvglm-method (biplot-methods), *119*
- biplot, rrvglm-method (biplot-methods), *119*
- biplot-methods, *119*
- biplot.rrvglm, *535, 732*
- biplot.rrvglm (lvplot.rrvglm), *533*
- Bisa, *119*
- bisa, *120, 121, 443*
- Bistudentt, *123*
- bistudentt, *113, 123, 124*
- bmi.nz, *37, 125, 489, 495, 543*
- borel.tanner, *126, 128, 288, 412*
- Bort, *128*
- boxcox, *888*
- Brat, *129, 131, 133*
- brat, *129, 130, 130, 131–133*
- bratt, *129–131, 132*
- bs, *283, 291, 380, 390, 450, 451, 527, 541, 743, 774, 777, 779, 780, 820*
- budworm, *24, 134, 215, 219, 787, 788*
- calibrate, *135, 140, 144, 676*
- calibrate, Coef.qrrvglm-method (calibrate-methods), *137*
- calibrate, qrrvglm-method (calibrate-methods), *137*
- calibrate, rrvgam-method (calibrate-methods), *137*
- calibrate, rrvglm-method (calibrate-methods), *137*
- calibrate-methods, *137*
- calibrate.qrrvglm, *136, 137, 141–144, 148, 210, 674*
- calibrate.qrrvglm.control, *138, 140, 141, 145*
- calibrate.rrvglm, *136, 140, 142, 144, 145, 732*
- calibrate.rrvglm.control, *143, 144*
- cao, *13, 108, 109, 140, 145, 149, 152, 175, 207, 210, 430, 482, 527, 592, 593, 644, 645, 690, 845, 846*
- cao.control, *146–148, 149*
- Card, *152*
- cardioid, *153, 154, 870*
- cauchit, *159*

- cauchitlink, *28, 49, 101, 102, 156, 174, 175, 215, 219, 351, 481, 505, 597, 680, 788*
- Cauchy, *157, 159*
- cauchy, *157, 158, 757*
- cauchy1, *156, 157, 757, 791*
- cauchy1 (cauchy), *158*
- cbind, *201, 749, 750*
- cdf.lmscreg, *160, 489, 492, 495*
- cennormal (cens.normal), *163*
- cens.gumbel, *161, 402*
- cens.normal, *163, 246, 810, 835*
- cens.poisson, *165, 556, 645, 803, 804*
- cens.rayleigh (rayleigh), *702*
- cfibrosis, *167, 752*
- cgo, *168*
- chest.nz, *169, 543*
- chinese.nz, *170*
- chisq, *171, 358, 757*
- chisq.test, *716, 717*
- Chisquare, *172*
- choose, *240, 243*
- clo, *172*
- cloglink, *64, 66, 174, 280–282*
- cloglink (clogloglink), *174*
- clogloglink, *28, 49, 101, 102, 108, 147, 157, 174, 175, 206, 215–217, 219, 270, 351, 481, 499, 505, 597, 680, 699, 700, 787, 788, 925, 926*
- CM.equid, *24, 176, 177, 193, 215, 219, 574, 788*
- CM.free, *177, 794*
- CM.free (CM.equid), *176*
- CM.ones, *177, 219*
- CM.ones (CM.equid), *176*
- CM.qlogis, *177*
- CM.qlogis (CM.equid), *176*
- CM.qnorm, *177, 193, 217, 739*
- CM.qnorm (CM.equid), *176*
- CM.symm0, *14*
- CM.symm0 (CM.equid), *176*
- CM.symm1, *177, 217*
- CM.symm1 (CM.equid), *176*
- coalminers, *102, 107, 177*
- Coef, *147, 178, 186, 466, 527, 543, 602, 813, 825*
- coef, *13, 178, 179, 185–188, 198, 199, 204*
- coef, vgam-method (coefvgam), *186*
- Coef.qrrvglm, *141, 142, 179, 179, 180, 182, 210, 529, 531, 532, 622, 624, 673, 690*
- Coef.qrrvglm-class, *181*
- Coef.rrvglm, *145, 179, 183, 185, 535, 732*
- Coef.rrvglm-class, *184*
- Coef.vlm, *179, 185*
- coefficients, vgam-method (coefvgam), *186*
- coefvglm, *186, 188*
- coefvglm, *186, 187, 187, 204, 858*
- colMeans, *655*
- colnames, *333, 397*
- CommonVGAMffArguments, *16, 19, 21, 23, 24, 31, 33, 36, 38, 43, 59, 64, 67, 70, 73, 75, 78, 80, 82, 85, 89, 92, 97, 100, 105, 107, 108, 112, 113, 118, 121, 122, 124, 126, 129, 146, 154, 158, 161, 162, 164, 165, 171, 177, 188, 202, 206, 214–217, 219, 222, 238, 240, 242, 245, 247, 248, 254, 262, 263, 265, 267, 272, 274, 277, 280, 283, 284, 287, 289, 297, 298, 303, 304, 309, 311, 319, 320, 325, 327, 332–334, 337, 341–346, 349, 354, 356, 357, 364, 366, 368, 371, 373–375, 379, 381, 387, 389, 390, 400, 405, 422, 423, 428, 434, 439, 442, 445, 448, 455, 459, 463, 474, 478, 485, 488, 490, 492, 494, 500, 501, 503, 504, 510, 511, 516, 520, 522, 539, 546, 548, 554, 557, 559, 561, 571–574, 577, 580, 584, 588–590, 593, 594, 596, 604, 609, 610, 612, 620, 644, 648, 651, 654, 658, 659, 663, 664, 667, 669, 677, 703, 713, 720, 744, 751, 755, 760, 763, 766, 769, 787, 788, 790, 800, 809, 810, 816, 823, 831, 835, 837, 853, 855, 869, 878, 880, 881, 890, 892–894, 896, 897, 899–901, 904–906, 911, 914, 918, 921–926, 934, 939*
- concoef, *147, 197, 197*
- concoef, cao-method (concoef-methods), *198*
- concoef, Coef.cao-method (concoef-methods), *198*
- concoef, Coef.qrrvglm-method

- (concoef-methods), 198
- concoef, Coef.rrvglm-method
(concoef-methods), 198
- concoef, qrrvglm-method
(concoef-methods), 198
- concoef, rrvglm-method
(concoef-methods), 198
- concoef-method (concoef-methods), 198
- concoef-methods, 198
- confint, 138, 199, 200
- confint.default, 199
- confint.lm, 200
- confintrrvglm (confintvglm), 199
- confintvgam (confintvglm), 199
- confintvglm, 199, 415, 525, 606, 681, 682, 799, 841, 858, 871
- constraints, 201, 216, 217, 219, 333, 452, 453, 455, 821, 855, 906
- constraints.vlm, 567, 858
- cops, 203
- cops, vglm-method (cops), 203
- copsvglm (cops), 203
- cor, 456
- corbet, 204, 664
- coxph, 413
- cqo, 13, 16, 108, 109, 140, 146–148, 168, 169, 173, 175, 181, 182, 205, 344, 397, 398, 421, 430, 482, 527, 532, 565, 566, 588, 590, 592, 593, 624, 631, 644, 645, 674, 686, 690, 708, 709, 731, 732, 739, 780, 813, 828
- crashbc (crashes), 212
- crashes, 212
- crashf (crashes), 212
- crashi, 398, 569, 732
- crashi (crashes), 212
- crashmc (crashes), 212
- crashp (crashes), 212
- crashtr (crashes), 212
- cratio, 24, 177, 191, 193, 214, 218, 219, 452, 541, 542, 574, 716, 757, 787, 788, 867
- cumsum, 552
- cumulative, 23, 24, 87, 176, 177, 191, 193, 199, 203, 214, 215, 216, 286, 412, 452, 504, 541, 542, 572–574, 606, 607, 641, 682–684, 699, 700, 716, 757, 785, 787, 788, 867, 878
- cut, 604, 706, 708, 709
- D, 481
- Dagum, 220, 223
- dagum, 76, 220, 221, 222, 298, 356, 446, 449, 521, 610, 757, 761
- dalap, 229, 283, 285
- dAR1, 44, 223
- data.frame, 564
- dbenf (Benford), 56
- dbenini (Benini), 57
- dbetabinom (Betabinom), 60
- dbetageom (Betageom), 72
- dbetanorm (Betanorm), 76
- dbiamhcop (Biamhcop), 81
- dbiclaytoncop, 86
- dbiclaytoncop (Biclaytoncop), 84
- dbifgmcop (Bifgmcop), 88
- dbifrankcop (Frank), 306
- dbilogis (bilogis), 95
- dbinom, 61, 315, 891, 894
- dbinom2.or (Binom2.or), 98
- dbinom2.rho (Binom2.rho), 103
- dbinorm (Binorm), 110
- dbinormcop (Binormcop), 115
- dbiplackcop (Biplackett), 116
- dbisa (Bisa), 119
- dbistudentt, 125
- dbistudentt (Bistudentt), 123
- dbort (Bort), 128
- dcard (Card), 152
- ddagum (Dagum), 220
- ddiffzeta (Diffzeta), 236
- deermice, 52, 225, 649, 653, 656
- deexp, 34, 257, 261
- deexp (Expectiles-Exponential), 255
- deflated (altered), 29
- denorm, 32, 37, 256, 261, 492
- denorm (Expectiles-Normal), 256
- deparse1, 784
- deplot.lmscreg, 161, 226, 489, 492, 495, 628, 694
- depvar, 179, 228
- deriv3, 23, 24
- deunif, 255–257
- deunif (Expectiles-Uniform), 259
- deviance, 230
- dexp, 256
- dexpgeom, 268

- dexpgeom (expgeom), 266
dexplog, 273
dexplog (explog), 271
dexppois, 278
dexppois (exppois), 275
dextbetabinom (Extbetabinom), 278
dextlogF, 229, 245, 285
df.residual, 230
df.residual_vlm (df.residual), 230
dfbeta (hatvalues), 410
dfbetavlm (hatvalues), 410
dfelix, 288
dfelix (Felix), 286
dfisk (Fisk), 295
dfoldnorm (Foldnorm), 302
dfrechet (Frechet), 308
dgaitdbinom (Gaitdbinom), 313
dgaitdlog, 233
dgaitdlog (Gaitdlog), 316
dgaitdnbinom, 234
dgaitdnbinom (Gaitdnbinom), 321
dgaitdplot, 231, 314, 317, 327, 330, 339, 549, 550, 629, 784
dgaitdpois, 233
dgaitdpois (Gaitdpois), 327
dgaitdzeta (Gaitdzeta), 337
dgamma, 343
dgenbetaII, 356
dgenbetaII (GenbetaII), 353
dengamma.stacy (engammaUC), 358
dgenpois0 (Genpois0), 360
dgenpois1 (Genpois1), 362
dgenpois2 (Genpois1), 362
dgenray, 371
dgenray (genray), 369
dgensh (Gensh), 372
dgeom, 894, 897, 916, 917
dgev, 379
dgev (gevUC), 381
dgomperz, 388, 407
dgomperz (Gompertz), 386
dgp, 389, 391
dgp (gpdUC), 392
dgumbel, 386, 404
dgumbel (gumbelUC), 406
dgumbelII, 402, 405
dgumbelII (Gumbel-II), 403
dhuber, 234
dhurea, 429
dhurea (Hurea), 427
dhyper, 430, 431
dhzeta, 435
dhzeta (Hzeta), 433
Diffzeta, 236, 238
diffzeta, 236, 237, 237, 911
digamma, 677
dimm (UtilitiesVGAM), 836
dinv.gaussian (Inv.gaussian), 441
dinv.lomax (Inv.lomax), 444
dinv.paralogistic (Inv.paralogistic), 447
dirichlet, 238, 241, 244, 250, 436, 574, 711, 756, 757
dirmul.old, 240, 244
dirmultinomial, 66, 69, 239, 241, 242, 282, 574
dkumar, 460
dkumar (Kumar), 458
dlaplace (laplaceUC), 465
dlgamma (lgammaUC), 475
dlind, 479
dlind (Lindley), 476
dlino (Lino), 483
dlog, 316, 317
dlog (Log), 496
dlogF, 244, 500
dlogis, 503
dlomax (Lomax), 518
dmakeham, 386, 539, 540
dmakeham (Makeham), 537
dmaxwell (Maxwell), 545
dN1binom (N1binom), 576
dN1pois (N1pois), 579
dnaka (Nakagami), 582
dnbinom, 322, 898, 901, 920
dnorm, 257, 305, 667, 835
do.call, 233
double.cens.normal, 164, 245, 713, 810, 835
double.exbinomial, 109, 247, 817
dparalogistic (Paralogistic), 608
dpareto (Pareto), 611
dparetoI (ParetoIV), 614
dparetoII (ParetoIV), 614
dparetoIII (ParetoIV), 614
dparetoIV (ParetoIV), 614

- dperks, [621](#)
 dperks (Perks), [619](#)
 dpois, [139](#), [330](#), [361](#), [365](#), [561](#), [581](#), [646](#), [762](#),
 [764](#), [932](#)
 dpois.points (PoissonPoints), [645](#)
 dpolono (Polono), [646](#)
 dposbern, [649](#), [653](#)
 dposbern (posbernUC), [657](#)
 dposgeom (Posgeom), [661](#)
 dposnorm (Posnorm), [665](#)
 drayleigh (Rayleigh), [701](#)
 drice, [721](#)
 drice (Rice), [719](#)
 drop1.glm, [24](#), [25](#), [789](#)
 drop1.vglm, [27](#), [41](#), [567](#), [789](#), [821](#), [858](#)
 drop1.vglm (add1.vglm), [24](#)
 dsc.t2, [261](#), [744](#), [745](#)
 dsc.t2 (Expectiles-sc.t2), [258](#)
 dsimplex, [756](#)
 dsimplex (Simplex), [754](#)
 dsinmad (Sinmad), [758](#)
 dskellam, [764](#)
 dskellam (Skellam), [761](#)
 dskewnorm (skewnorm), [764](#)
 dslash (Slash), [767](#)
 dt, [123](#), [259](#), [503](#)
 dtobit, [809](#)
 dtobit (Tobit), [806](#)
 dtopple (Topleft), [814](#)
 dtriangle (Triangle), [818](#)
 dtrinorm (Trinorm), [821](#)
 dtruncpareto (Truncpareto), [830](#)
 ducklings, [250](#)
 dunif, [260](#), [261](#)
 dweibull, [832](#), [880](#), [883](#)
 dyules, [890](#)
 dyules (Yules), [888](#)
 dzabinom, [894](#)
 dzabinom (Zabinom), [891](#)
 dzageom, [897](#)
 dzageom (Zageom), [894](#)
 dzanegbin (Zanegbin), [897](#)
 dzapois (Zapois), [902](#)
 dzeta, [338](#), [339](#)
 dzeta (Zeta), [907](#)
 dzibinom (Zibinom), [912](#)
 dzigeom (Zigeom), [916](#)
 dzinegbin (Zinegbin), [919](#)
 dzipf, [929](#)
 dzipf (Zipf), [927](#)
 dzipfmb (Zipfmb), [930](#)
 dzipois, [903](#)
 dzipois (Zipois), [931](#)
 dzoabeta, [71](#)
 dzoabeta (Zoabeta), [937](#)
 dzoibetabinom (Betabinom), [60](#)
 eCDF, [250](#), [285](#), [492](#)
 edhuber (dhuber), [234](#)
 eexp (Expectiles-Exponential), [255](#)
 enorm (Expectiles-Normal), [256](#)
 enzyme, [251](#), [555](#)
 erf, [252](#)
 erfc (erf), [252](#)
 erlang, [253](#), [412](#), [757](#)
 eunif (Expectiles-Uniform), [259](#)
 exp, [269](#), [462](#), [498](#)
 Expectiles-Exponential, [255](#)
 Expectiles-Normal, [256](#)
 Expectiles-sc.t2, [258](#)
 Expectiles-Uniform, [259](#)
 expexpff, [262](#), [264](#), [265](#), [345](#), [350](#), [880](#), [883](#)
 expexpff1, [263](#), [264](#)
 expexpint (expint), [268](#)
 expgeom, [266](#)
 expgeometric, [266](#), [267](#), [275](#), [376](#)
 expint, [268](#)
 explink, [270](#), [513](#), [695](#), [696](#), [698](#), [834](#)
 explog, [271](#)
 explogff, [271](#), [272](#), [272](#), [275](#), [502](#)
 expm1, [497](#), [498](#)
 Exponential, [275](#), [393](#)
 exponential, [33](#), [34](#), [254](#), [266](#), [268](#), [272](#), [273](#),
 [274](#), [278](#), [312](#), [349](#), [464](#), [558](#), [712](#),
 [716](#), [757](#)
 exppois, [275](#)
 exppoisson, [276](#), [277](#)
 Extbetabinom, [63](#), [278](#), [282](#)
 extbetabinomial, [65](#), [66](#), [68](#), [69](#), [279](#), [280](#),
 [517](#), [518](#)
 extlogF1, [32](#), [34](#), [37](#), [39](#), [229](#), [251](#), [282](#), [300](#),
 [451](#), [452](#), [492](#), [500](#)
 extlogitlink, [70](#), [71](#), [155](#), [381](#), [432](#), [481](#)
 extlogitlink (logitlink), [504](#)
 extractAIC.vglm, [25](#), [27](#), [789](#)
 familyname, [285](#)

- FDist, 289
 Felix, 286
 felix, 127, 286, 287, 287, 412
 fff, 288
 fill1, 218, 225, 290, 573, 574, 656, 750, 857,
 858, 863, 864
 finney44, 293
 fisherz, 718
 fisherzlink, 107, 294, 481, 597, 718
 Fisk, 295, 298
 fisk, 76, 223, 296, 296, 355, 356, 446, 449,
 521, 610, 757, 761
 fitted, 299
 fitted.values.vlm(fittedvlm), 298
 fittedvlm, 192, 298, 674, 858, 892, 896, 899,
 904, 914, 918, 921, 923, 934
 fix.crossing, 283, 285, 300, 452
 flourbeetle, 301
 Foldnorm, 302
 foldnormal, 302, 303, 303, 767, 835
 foldsqrtlink, 481
 foldsqrtlink(sqrtlink), 784
 format.SurvS4(SurvS4), 802
 formula, 305, 306, 526, 749, 750
 formula.vlm(formulavlm), 305
 formulavlm, 305, 410
 Frank, 306
 Frechet, 308
 frechet, 308, 309, 309, 381
 freund61, 275, 311

 Gaitdbinom, 313, 323, 330, 892, 913
 Gaitdlog, 315, 316, 319, 320, 323, 330, 497
 gaitdlog, 30, 233, 317, 318, 327, 333, 337,
 342, 385, 502, 829
 Gaitdnbinom, 231, 234, 315, 321, 327, 330,
 898, 901
 gaitdnbinomial, 323, 323, 337, 385, 457,
 593, 664, 901, 923
 Gaitdpois, 30, 231, 232, 234, 314, 315, 317,
 320, 322, 323, 325, 327, 334, 335,
 337–339, 342, 550, 574, 670, 782,
 903–905, 932
 gaitdpoisson, 30, 232–234, 319, 320,
 324–327, 330, 331, 333, 340–342,
 385, 457, 550, 571, 574, 629, 645,
 670, 782–784, 829, 905, 936
 Gaitdzeta, 315, 317, 323, 330, 337, 341, 342
 gaitdzeta, 30, 320, 327, 337, 339, 340, 385,
 725, 829, 911
 gam, 846, 851, 857
 gamma, 254, 343, 344, 346, 349, 909, 910
 gamma1, 342, 345, 350, 358, 474, 757
 gamma2, 148, 206, 210, 344, 350, 358, 585,
 688, 690, 707, 709, 757
 GammaDist, 345
 gammaff, 347
 gammaff.mm, 343, 345, 346, 350
 gammahyperbola, 348
 gammaR, 254, 263, 343, 345, 349, 479, 757
 garma, 351
 gaussian, 835
 gaussianff(uninormal), 834
 GenbetaII, 353
 genbetaII, 71, 76, 81, 221, 223, 296–298,
 353, 354, 354, 445, 446, 448, 449,
 486, 519, 521, 609, 610, 759–761
 gengamma.stacy, 356, 359, 474, 678, 757
 gengammaUC, 358
 Genpois0, 360, 362, 363, 365
 Genpois1, 360, 361, 362, 367
 Genpois2, 369
 Genpois2(Genpois1), 362
 genpoisson0, 360, 361, 363, 364, 366–369,
 591, 593, 645
 genpoisson1, 364, 365, 366, 369, 591, 593,
 645
 genpoisson2, 191, 364, 365, 367, 367, 591,
 593, 645
 genray, 369
 genrayleigh, 369, 370, 370, 704
 Gensh, 372, 374
 gensh, 373, 373, 432, 504
 Geometric, 376
 geometric, 72, 74, 266, 268, 375, 757, 897,
 919
 get.smart, 376, 377, 378, 685, 779, 780
 get.smart.prediction, 377, 377, 780
 gev, 163, 310, 378, 382, 391, 401, 402, 405,
 407, 409, 843, 880, 882, 883
 gevff, 382, 402, 437, 724
 gevff(gev), 378
 gevUC, 381
 gew, 383, 801
 glm, 13, 16, 188, 377, 482, 567, 592, 726, 753,
 779, 835, 846, 853, 855, 857, 868,

- 885
 goffset, 320, 327, 337, 342, 384, 829
 Gompertz, 386
 gompertz, 386, 387, 540, 757
 gpd, 275, 381, 389, 393, 552, 553, 614, 618, 843
 gpdUC, 392
 grain.us, 394, 729
 grc, 54, 213, 395, 601, 732
 grep, 190
 gumbel, 162, 163, 381, 400, 405, 407, 409, 692, 693
 Gumbel-II, 403
 gumbelff, 163, 381, 407, 409, 692
 gumbelff (gumbel), 400
 gumbelII, 403, 404, 404, 757, 880, 883
 gumbelUC, 406
 guplot, 163, 381, 402, 408, 843

 has.intercept (has.interceptvml), 409
 has.interceptvml, 306, 409
 hatplot (hatvalues), 410
 hatvalues, 410, 717
 hatvaluesvml, 858
 hatvaluesvml (hatvalues), 410
 hdeff, 14, 16, 28, 49, 204, 336, 412, 414, 416–418, 514, 525, 746, 748, 771, 785, 876–878
 hdeff.vglm, 108, 109, 219, 645, 799, 835, 841, 858, 936
 hdeffsev, 414, 415, 416, 747, 748, 878
 hdeffsev2 (hdeffsev), 416
 hist, 557, 561, 783
 hormone, 418
 hspider, 210, 420
 huber1 (huber2), 422
 huber2, 235, 422, 791, 835
 Huggins89.t1, 423, 653, 656
 Huggins89table1, 653, 656
 Huggins89table1 (Huggins89.t1), 423
 hunua, 425, 512, 875
 Hurea, 427
 hurea, 427, 428, 428, 429
 hyperg, 430
 hypersecant, 245, 373, 374, 431
 hypersecant01 (hypersecant), 431
 Hzeta, 433, 435
 hzeta, 433, 434, 434, 757, 911

 I, 742
 iam, 435
 identity, 597, 644, 721
 identitylink, 192, 437, 481, 597, 606, 715, 935
 inflated, 782
 inflated (altered), 29
 Influence, 438
 influence.measures, 412
 Influence.vglm, 858
 integrate, 428, 647, 768
 interleave.VGAM (UtilitiesVGAM), 836
 inv.binomial, 439, 593, 645
 Inv.gaussian, 441, 443
 inv.gaussianff, 122, 441, 442, 442, 879
 Inv.lomax, 444
 inv.lomax, 76, 223, 298, 356, 444, 445, 445, 446, 449, 521, 610, 757, 761
 Inv.paralogistic, 447, 449
 inv.paralogistic, 76, 223, 298, 356, 446–448, 448, 521, 610, 757, 761
 iris, 574
 is.altered (altered), 29
 is.buggy, 450, 743, 846
 is.crossing, 283, 285, 300, 451
 is.deflated (altered), 29
 is.inflated (altered), 29
 is.na.SurvS4 (SurvS4), 802
 is.parallel, 202, 452
 is.smart, 453, 774, 777
 is.SurvS4 (SurvS4), 802
 is.truncated (altered), 29
 is.zero, 202, 454

 kendall.tau, 86, 114, 455
 KLD, 327, 337, 456
 KLDvglm (KLD), 456
 Kumar, 458
 kumar, 71, 81, 458, 459, 757

 lake0, 460
 lambertW, 462, 513, 537, 538
 laplace, 275, 423, 463, 465, 466
 laplaceUC, 465
 latvar, 148, 466, 467, 528
 lbeta, 61
 lcalogitlink, 28, 49
 lcalogitlink (alogitlink), 27
 lcsloglink, 784

- lcsloglink (sloglink), 770
- legend, 747
- leipnik, 468, 549
- lerch, 469, 910
- leukemia, 471, 804
- levy, 472
- lfactorial, 361
- lgamma, 430, 474, 678
- lgamma1, 343, 473, 475, 476, 757
- lgamma3, 476, 677, 678, 757
- lgamma3 (lgamma1), 473
- lgammaUC, 475
- Lindley, 476
- lindley, 343, 412, 477, 478, 757
- linkfun, 479, 480, 482
- linkfunv1m, 858
- linkfunv1m (linkfun), 479
- Links, 16, 17, 19–22, 27, 28, 42, 49, 59, 64, 67, 73, 75, 78, 80, 82, 85, 89, 91, 92, 94, 97, 100, 105, 107, 109, 112, 113, 118, 121, 124, 126, 154, 156, 158, 162, 163, 165, 166, 174, 175, 189, 195, 214, 216, 219, 222, 238, 240, 242, 245, 247, 254, 262, 264, 267, 270, 271, 274, 277, 280, 283, 287, 289, 294, 295, 297, 303, 309, 311, 319, 332, 341, 343, 344, 346, 348, 349, 354, 356, 364, 366, 368, 371, 373, 375, 378, 387, 389, 400, 404, 422, 430, 432, 434, 437–439, 442, 445, 448, 459, 463, 468, 472, 473, 478, 480, 485, 490, 498–501, 503–505, 507, 508, 513–518, 520, 522, 539, 546, 548, 554, 557, 558, 561, 563, 570, 571, 577, 580, 584–586, 588, 609, 612, 617, 620, 642, 644, 645, 662, 667, 669–671, 677, 679, 680, 703, 711, 713, 714, 718, 720, 722, 744, 751, 755, 760, 763, 766, 769, 771, 784, 785, 787, 790, 809, 823, 831, 834, 869, 880, 881, 890, 892, 896, 899, 903, 911, 914, 918, 921, 924, 928, 934, 939
- Lino, 483, 486
- lino, 356, 484, 485, 757
- lirat, 66, 69, 282, 487
- list, 855
- lm, 16, 36, 188, 230, 377, 378, 567, 598, 753, 779
- lm.influence, 439
- lms.bcg, 34, 161, 227, 488, 492, 495, 694
- lms.bcn, 37, 161, 227, 251, 257, 283, 285, 300, 452, 488, 489, 490, 494, 495, 551, 694
- lms.yjn, 161, 226, 227, 489, 492, 493, 632, 693, 694, 888
- lms.yjn2 (lms.yjn), 493
- Log, 496, 502
- log, 269, 462, 498, 501, 502, 513
- log10, 301
- log1mexp, 497
- log1p, 497, 498
- log1pexp (log1mexp), 497
- log1plink, 64, 66, 517
- log1plink (logofflink), 517
- logclink, 481, 498, 513
- loge, 482
- logF, 245, 285, 499
- logff, 57, 320, 496, 497, 500, 501, 664, 757
- logffMlink, 319, 320, 385
- Logistic, 373, 505
- logistic, 98, 373, 374, 502, 505, 757, 791
- logistic1, 219, 503, 505, 757
- logistic1 (logistic), 502
- logit, 482
- logitlink, 28, 49, 102, 108, 147, 157, 174, 175, 192, 206, 215, 219, 283, 295, 334, 351, 365, 438, 481, 482, 501, 504, 504, 507, 508, 513, 570, 571, 597, 606, 680, 699, 700, 788
- logitoffsetlink, 175, 505, 507
- logLik, 526
- logLik.v1m, 508
- loglinb2, 102, 107, 509, 512
- loglinb3, 102, 510, 511
- loglink, 157, 192, 271, 283, 351, 438, 481, 482, 486, 499, 501, 502, 505, 512, 514, 515, 517, 518, 541, 586, 597, 644, 645, 671, 771, 816, 880
- loglog, 435, 514, 790, 911
- loglog (logloglink), 514
- logloglink, 481, 499, 513, 514, 588, 597
- loglogloglink, 481
- loglogloglink (logloglink), 514
- logneglink (loglink), 512
- Lognormal, 516, 646

- lognormal, 358, 515, 647, 757, 880, 883
 logofflink, 378, 389, 481, 499, 502, 513,
 515, 517, 549, 597
 Lomax, 518, 521
 lomax, 76, 223, 298, 356, 446, 449, 519, 520,
 610, 757, 761
 lpossums, 521, 936
 lqnorm, 522
 lrt.stat, 415, 417, 524, 682, 746, 871, 878
 lrt.stat.vlm, 41, 200, 414, 527, 798, 799,
 841, 858
 lrtest, 41, 525, 526
 lrtest.vglm, 858
 lrtest.vglm(lrtest), 526
 lv, 467
 lv(latvar), 466
 lvplot, 119, 147, 209, 467, 527, 532, 535, 825
 lvplot.qrrvglm, 180, 181, 210, 528, 529,
 623, 624, 630, 631, 690
 lvplot.rrvglm, 533, 535, 732, 736

 machinists, 536
 magic, 13, 772–774, 777, 846, 851, 852
 Makeham, 537
 makeham, 388, 537, 538, 538, 757
 mapply, 361
 margeff, 24, 215, 219, 540, 574, 593, 645,
 664, 788, 810
 marital.nz, 542
 match.call, 778
 Math.SurvS4 (SurvS4), 802
 matrix, 99, 104, 194, 254, 329, 346, 413, 594,
 804, 810
 Max, 147, 148, 543, 603, 813
 Maxwell, 545, 547, 646
 maxwell, 545, 546, 546, 642, 643, 702, 704,
 880, 883
 mccullagh89, 469, 547
 meangaitd, 234, 327, 337, 549
 median, 464
 medpolish, 397, 398
 melbmaxtemp, 550
 meplot, 391, 551
 methods, 792, 795–797
 micmen, 252, 553, 855
 mills.ratio, 165, 166, 555, 810
 mills.ratio2 (mills.ratio), 555
 mix2exp, 275, 556
 mix2normal, 558, 562, 835

 mix2poisson, 558, 560, 560, 645
 MNSs, 18–20, 22, 562
 model.frame, 564
 model.frame.vlm, 564, 567
 model.matrix, 228, 567
 model.matrix.default, 146, 854
 model.matrix.qrrvglm, 210, 565, 841
 model.matrix.vlm, 564, 566, 566, 821
 moffset, 398, 568, 569, 635, 705
 multilogitlink, 194, 195, 234, 327, 337,
 479, 480, 505, 570, 574, 597
 multinom, 574
 Multinomial, 574
 multinomial, 14, 22–24, 131, 143, 176, 177,
 193, 195, 214–216, 218, 219, 239,
 241, 243, 244, 291, 315, 317, 323,
 327, 330, 332, 337, 339, 398, 436,
 452, 534, 541, 542, 570, 571, 571,
 682, 716, 757, 787, 788

 N1binom, 576
 N1binomial, 109, 576, 577, 577, 581, 835
 N1pois, 579
 N1poisson, 578–580, 580, 645, 835
 na.omit, 854
 Nakagami, 582
 nakagami, 582, 583, 583
 names, 737
 nbcanlink, 585, 586, 591–593, 595
 nef.hs (hypersecant), 431
 NegBinomial, 324, 593
 negbinomial, 148, 189, 192, 206, 210, 325,
 327, 362, 365, 367, 369, 375, 376,
 397, 440, 514, 536, 541, 542, 586,
 587, 594, 595, 644, 645, 647, 663,
 664, 688, 690, 707, 709, 732, 757,
 900, 901, 920–923, 936
 negbinomial.size, 586, 593, 594, 757
 negidentitylink, 481
 negidentitylink (identitylink), 437
 negloglink, 345, 350, 481, 588, 590, 617
 negloglink (loglink), 512
 negreciprocallink, 481
 negreciprocallink (reciprocallink), 714
 nobs.vlm, 858
 Normal, 77, 372, 556, 560, 680, 889, 907, 927
 normal.vcm, 194, 414, 571, 596
 nparam (nparam.vlm), 599
 nparam.vlm, 599

- npred.vlm, 858
 ns, 291, 380, 390, 450, 451, 527, 743, 779, 780

 Oalog, 317
 oalog, 319, 320, 502
 Oazeta, 339, 908
 oazeta, 341, 342, 910, 911
 offset, 385
 Oilog, 317, 497
 oilog, 319, 320, 502
 Oizeta, 237, 339, 908
 oizeta, 341, 342, 910, 911
 oldClass, 804
 olym08, 398
 olym08 (olympics), 601
 olym12, 398
 olym12 (olympics), 601
 olympics, 601
 Ops.SurvS4 (SurvS4), 802
 Opt, 147, 148, 544, 602, 813
 optim, 138, 140, 141, 144, 146, 150, 687, 688, 690, 740
 optimize, 203, 204
 options, 146, 206, 854
 order, 182
 ordered, 23, 215, 217, 605, 683, 787
 ordpoisson, 603, 645
 ordsup, 219, 606, 835
 Otlog, 317, 497
 otlog, 319, 320, 502
 otpospoisson, 670
 Otzeta, 339, 908
 otzeta, 341, 342, 910, 911
 oxtemp, 381, 607

 pairs, 871
 par, 232, 233, 408, 411, 529–532, 534, 535, 552, 622–624, 627, 628, 630, 632, 634, 640, 684, 692, 723, 783, 784, 826–828
 Paralogistic, 608, 610
 paralogistic, 76, 223, 298, 356, 446, 449, 521, 608, 609, 609, 757, 761
 param.names (UtilitiesVGAM), 836
 Pareto, 611, 614, 616
 paretoff, 391, 611, 612, 612, 618
 ParetoI (ParetoIV), 614
 ParetoII (ParetoIV), 614
 paretoII (paretoIV), 616
 ParetoIII (ParetoIV), 614
 paretoIII (paretoIV), 616
 ParetoIV, 612, 614, 618
 paretoIV, 613–616, 616
 paste, 836
 pbenf (Benford), 56
 pbenini (Benini), 57
 pbeta, 938
 pbetabinom (Betabinom), 60
 pbetabinom.ab, 62
 pbetageom (Betageom), 72
 pbetanorm (Betanorm), 76
 pbiamhcop (Biamhcop), 81
 pbifgmcop (Bifgmcop), 88
 pbifrankcop (Frank), 306
 pbilogis (bilogis), 95
 pbinom, 315, 912
 pbinorm, 107, 113, 114
 pbinorm (Binorm), 110
 pbinormcop (Binormcop), 115
 pbiplackcop (Biplackett), 116
 pbisa, 122
 pbisa (Bisa), 119
 pcard (Card), 152
 pchisq, 525
 pdagum (Dagum), 220
 pdiffzeta (Diffzeta), 236
 peexp (Expectiles-Exponential), 255
 penorm (Expectiles-Normal), 256
 Perks, 619
 perks, 619, 620, 757
 persp, 623, 624
 persprrvglm, 210, 532, 622, 825
 peunif (Expectiles-Uniform), 259
 pexp, 255
 pexpgeom (expgeom), 266
 pexplog (explog), 271
 pexppois (exppois), 275
 pextbetabinom (Extbetabinom), 278
 pfisk (Fisk), 295
 pfoldnorm (Foldnorm), 302
 pfrechet (Frechet), 308
 pgaitdbinom (Gaitdbinom), 313
 pgaitdlog (Gaitdlog), 316
 pgaitdnbinom (Gaitdnbinom), 321
 pgaitdpois (Gaitdpois), 327
 pgaitdzeta (Gaitdzeta), 337
 pgamma, 625–627

- pgamma.deriv, 624, 626, 627, 832
 pgamma.deriv.unscaled, 625, 626, 832
 pgengamma.stacy (gengammaUC), 358
 pgenpois0 (Genpois0), 360
 pgenpois1 (Genpois1), 362
 pgenpois2 (Genpois1), 362
 pgenray (genray), 369
 pgensh (Gensh), 372
 pgev (gevUC), 381
 pgompertz (Gompertz), 386
 pgpd (gpdUC), 392
 pgumbel, 175
 pgumbel (gumbelUC), 406
 pgumbelII (Gumbel-II), 403
 phuber (dhuber), 234
 phzeta (Hzeta), 433
 pinv.gaussian (Inv.gaussian), 441
 pinv.lomax (Inv.lomax), 444
 pinv.paralogistic (Inv.paralogistic),
 447
 pkumar (Kumar), 458
 plaplace (laplaceUC), 465
 plgamma (lgammaUC), 475
 plind (Lindley), 476
 plino (Lino), 483
 plog, 236, 317
 plog (Log), 496
 plomax (Lomax), 518
 plot, 232, 408, 552, 622, 623, 640, 783, 784,
 827, 871
 plot.default, 232, 233, 634
 plot.vgam (plotvgam), 635
 plot.window, 634
 plotdeplot.lmscreg, 226, 227, 627
 plotdgaitd, 234, 327, 337, 784
 plotdgaitd (plotdgaitd.vglm), 629
 plotdgaitd.vglm, 629
 plotqrrvglm, 630
 plotqtplot.lmscreg, 631, 693, 694
 plotrcim0, 398, 569, 633, 705
 plotvgam, 635, 639, 640, 715, 735, 846, 849,
 856, 860
 plotvgam.control, 637, 638, 640
 plotvglm, 637, 639, 858
 pmakeham, 537
 pmakeham (Makeham), 537
 pmaxwell (Maxwell), 545
 pnaka (Nakagami), 582
 pnbinom, 322
 pneumo, 24, 215, 218, 219, 572, 641, 788
 pnorm, 56, 58, 110, 111, 114, 120, 121, 153,
 220, 235, 252, 253, 257, 276, 295,
 302, 304, 359, 369, 386, 403, 433,
 444, 447, 458, 465, 475, 477, 484,
 518, 537, 545, 578, 581, 582, 608,
 611, 615, 619, 667, 701, 719, 758,
 768, 814, 818, 822, 830, 889
 pnorm2 (Binorm), 110
 points, 783
 Poisson, 328, 330, 360, 365, 367, 369, 645,
 907, 927
 poisson, 278, 645
 poisson.points, 547, 642, 645, 646, 704
 poissonff, 38, 127, 138, 139, 143, 148, 165,
 166, 204, 206, 210, 275, 334, 337,
 362, 365, 367, 369, 398, 412–414,
 416, 440, 536, 541, 542, 562, 581,
 593, 595, 604, 605, 643, 643, 647,
 670, 690, 709, 716, 757, 764, 771,
 784, 785, 838, 839, 846, 857
 PoissonPoints, 645
 polf, 605, 645
 Polono, 646
 poly, 541, 779, 780, 820
 polya, 757
 polya (negbinomial), 587
 polyaR, 757
 polyaR (negbinomial), 587
 posbernoulli.b, 225, 513, 648, 652, 653,
 656–658, 660
 posbernoulli.t, 52, 225, 648, 649, 651, 652,
 654–660, 750
 posbernoulli.tb, 26, 648, 649, 652, 653,
 654, 658–660, 806
 posbernUC, 657
 Posbinom, 315
 posbinomial, 26, 109, 193, 649, 652, 653,
 656, 659, 757, 893, 894, 915
 Posgeom, 661
 Posnegbin, 323
 posnegbinomial, 541, 593, 662, 670, 757,
 901, 905, 922
 Posnorm, 665
 posnormal, 666, 667, 757, 810, 835
 Pospois, 329, 330, 670
 pospoisson, 337, 645, 664, 669, 757, 904, 905

- powerlink, [438](#), [481](#), [670](#), [715](#)
pparalogistic (Paralogistic), [608](#)
ppareto (Pareto), [611](#)
pparetoI (ParetoIV), [614](#)
pparetoII (ParetoIV), [614](#)
pparetoIII (ParetoIV), [614](#)
pparetoIV (ParetoIV), [614](#)
pperks (Perks), [619](#)
ppoints, [593](#)
ppois, [165](#), [330](#)
ppolono (Polono), [646](#)
pposgeom (Posgeom), [661](#)
pposnorm (Posnorm), [665](#)
prats, [671](#)
prayleigh (Rayleigh), [701](#)
predict, [13](#), [135](#), [136](#), [676](#)
predict.bs, [780](#)
predict.lm, [753](#)
predict.poly, [780](#)
predictqrrvglm, [210](#), [673](#)
predictvglm, [299](#), [564](#), [567](#), [673](#), [674](#), [858](#)
prentice74, [358](#), [474](#), [476](#), [676](#)
price (Rice), [719](#)
prinia, [649](#), [653](#), [656](#), [678](#)
probitlink, [28](#), [49](#), [101](#), [102](#), [157](#), [174](#), [175](#),
[177](#), [215](#), [216](#), [219](#), [302](#), [351](#), [438](#),
[481](#), [505](#), [597](#), [606](#), [679](#), [699](#), [700](#),
[788](#)
profile, [682](#)
profile.glm, [199](#), [200](#), [681](#), [682](#), [871](#)
profile.nls, [871](#)
profilevglm, [199](#), [415](#), [525](#), [681](#), [871](#)
propodds, [203](#), [218](#), [219](#), [286](#), [452](#), [542](#), [607](#),
[682](#), [684](#), [700](#)
prplot, [219](#), [683](#)
psc.t2 (Expectiles-sc.t2), [258](#)
psinmad (Sinmad), [758](#)
pslash, [768](#)
pslash (Slash), [767](#)
pt, [125](#), [258](#)
ptobit (Tobit), [806](#)
ptopple (Topple), [814](#)
ptriangle (Triangle), [818](#)
ptruncpareto (Truncpareto), [830](#)
punif, [260](#), [308](#), [382](#), [393](#), [406](#)
put.smart, [685](#), [779](#), [780](#)
pyules (Yules), [888](#)
pzabinom (Zabinom), [891](#)
pzageom (Zageom), [894](#)
pzanegbin (Zanegbin), [897](#)
pzapois (Zapois), [902](#)
pzeta, [339](#)
pzeta (Zeta), [907](#)
pzibinom (Zibinom), [912](#)
pzigeom (Zigeom), [916](#)
pzinegbin (Zinegbin), [919](#)
pzipf (Zipf), [927](#)
pzipfmb (Zipfmb), [930](#)
pzipois (Zipois), [931](#)
pzoabeta (Zoabeta), [937](#)
pzoibetabinom (Betabinom), [60](#)
qbenf (Benford), [56](#)
qbenini (Benini), [57](#)
qbetanorm (Betanorm), [76](#)
qbinom, [315](#)
qbisa (Bisa), [119](#)
qcard (Card), [152](#)
qdagum (Dagum), [220](#)
qdiffzeta (Diffzeta), [236](#)
qeexp (Expectiles-Exponential), [255](#)
qenorm (Expectiles-Normal), [256](#)
qeunif (Expectiles-Uniform), [259](#)
qexp, [255](#)
qexpgeom (expgeom), [266](#)
qexplog (explog), [271](#)
qexppois (exppois), [275](#)
qextbetabinom (Extbetabinom), [278](#)
qfisk (Fisk), [295](#)
qfoldnorm (Foldnorm), [302](#)
qfrechet (Frechet), [308](#)
qgaitdbinom (Gaitdbinom), [313](#)
qgaitdlog (Gaitdlog), [316](#)
qgaitdnbinom (Gaitdnbinom), [321](#)
qgaitdpois (Gaitdpois), [327](#)
qgaitdzeta (Gaitdzeta), [337](#)
qgamma, [343](#)
qgengamma.stacy (gengammaUC), [358](#)
qgenpois0 (Genpois0), [360](#)
qgenpois1 (Genpois1), [362](#)
qgenpois2 (Genpois1), [362](#)
qgenray (genray), [369](#)
qgensh (Gensh), [372](#)
qgev (gevUC), [381](#)
qgompertz (Gompertz), [386](#)
qgpd (gpdUC), [392](#)
qgumbel (gumbelUC), [406](#)

- qgumbelII (Gumbel-II), 403
 qhuber (dhuber), 234
 qhzeta (Hzeta), 433
 qinv.lomax (Inv.lomax), 444
 qinv.paralogistic (Inv.paralogistic), 447
 qkumar (Kumar), 458
 qlaplace (laplaceUC), 465
 qlgamma (lgammaUC), 475
 qlino (Lino), 483
 qllog, 317
 qllog (Log), 496
 qllogis, 177
 qlomax (Lomax), 518
 qmakeham (Makeham), 537
 qmaxwell (Maxwell), 545
 qnaka (Nakagami), 582
 qnbinom, 322, 588
 qnorm, 56, 58, 120, 153, 177, 220, 235, 257, 276, 295, 302, 359, 369, 386, 403, 433, 444, 447, 458, 465, 475, 477, 484, 518, 537, 545, 581, 582, 608, 611, 615, 619, 701, 716, 719, 758, 768, 814, 818, 830, 889
 qparalogistic (Paralogistic), 608
 qpareto (Pareto), 611
 qparetoI (ParetoIV), 614
 qparetoII (ParetoIV), 614
 qparetoIII (ParetoIV), 614
 qparetoIV, 617
 qparetoIV (ParetoIV), 614
 qperks (Perks), 619
 qpois, 330
 qposgeom (Posgeom), 661
 qposnorm (Posnorm), 665
 qrayleigh (Rayleigh), 701
 qrice (Rice), 719
 qrrvglm.control, 150, 151, 206–210, 344, 566, 588, 686, 708, 709
 qsc.t2 (Expectiles-sc.t2), 258
 qsinmad (Sinmad), 758
 qt, 258
 qtobit (Tobit), 806
 qtopple (Topple), 814
 qtplot.gumbel, 691
 qtplot.gumbelff (qtplot.gumbel), 691
 qtplot.lmscreg, 161, 227, 489, 492, 495, 633, 693
 qtriangle (Triangle), 818
 qtruncpareto (Truncpareto), 830
 quantile, 190, 192, 557, 559, 561, 590, 769, 772, 773
 quasipoisson, 365, 367, 369, 590, 592, 593
 qunif, 260, 308, 382, 393, 406
 Qvar, 271, 398, 694, 698, 835
 qvar, 696, 698
 qyules (Yules), 888
 qzabinom (Zabinom), 891
 qzageom (Zageom), 894
 qzanegbin (Zanegbin), 897
 qzapois (Zapois), 902
 qzeta, 339
 qzeta (Zeta), 907
 qzibinom (Zibinom), 912
 qzigeom (Zigeom), 916
 qzinegbin (Zinegbin), 919
 qzipf (Zipf), 927
 qzipfmb (Zipfmb), 930
 qzipois (Zipois), 931
 qzoabeta (Zoabeta), 937
 R2latvar, 219, 683, 699
 range, 828, 829
 Rank, 700
 rank, 700
 Rayleigh, 546, 701, 704
 rayleigh, 358, 370, 371, 546, 547, 585, 642, 643, 702, 702, 704, 721, 757, 880, 883
 rbell, 55
 rbenf (Benford), 56
 rbenini (Benini), 57
 rbetabinom (Betabinom), 60
 rbetageom, 71, 74, 81, 376
 rbetageom (Betageom), 72
 rbetanorm, 71, 81
 rbetanorm (Betanorm), 76
 rbiamhcop, 83
 rbiamhcop (Biamhcop), 81
 rbiclaytoncop, 86
 rbiclaytoncop (Biclaytoncop), 84
 rbifgmcop, 90
 rbifgmcop (Bifgmcop), 88
 rbifrankcop, 92, 93
 rbifrankcop (Frank), 306
 rbilogis, 83, 98
 rbilogis (bilogis), 95

- rbinom, [315](#), [577](#)
- rbinom2.or, [101](#), [102](#)
- rbinom2.or (Binom2.or), [98](#)
- rbinom2.rho, [107](#)
- rbinom2.rho (Binom2.rho), [103](#)
- rbinorm, [822](#)
- rbinorm (Binorm), [110](#)
- rbinormcop, [114](#)
- rbinormcop (Binormcop), [115](#)
- rbiplackcop, [118](#), [119](#)
- rbiplackcop (Biplackett), [116](#)
- rbisa (Bisa), [119](#)
- rbort, [127](#), [128](#)
- rbort (Bort), [128](#)
- rcard, [155](#)
- rcard (Card), [152](#)
- Rcim, [398](#), [569](#), [635](#), [704](#)
- rcim, [13](#), [16](#), [54](#), [213](#), [270](#), [271](#), [568](#), [569](#), [634](#), [635](#), [694–696](#), [698](#), [705](#), [813](#), [866](#), [868](#)
- rcim (grc), [395](#)
- rcqo, [210](#), [690](#), [706](#)
- rdagum (Dagum), [220](#)
- rdiffzeta (Diffzeta), [236](#)
- rdiric, [239](#), [710](#)
- rec.exp1, [711](#)
- rec.normal, [414](#), [712](#)
- reciprocal, [351](#)
- reciprocallynk, [481](#), [588](#), [590](#), [714](#)
- reexp (Expectiles-Exponential), [255](#)
- renorm (Expectiles-Normal), [256](#)
- resid, [717](#)
- residualsvglm, [109](#), [336](#), [337](#), [645](#), [715](#), [858](#)
- reunif (Expectiles-Uniform), [259](#)
- rexp, [255](#), [274](#), [558](#)
- rexpgeom (expgeom), [266](#)
- rexplog (explog), [271](#)
- rexp Pois (exppois), [275](#)
- rextbetabinom (Extbetabinom), [278](#)
- rfisk (Fisk), [295](#)
- rfoldnorm, [305](#)
- rfoldnorm (Foldnorm), [302](#)
- rfrechet, [310](#)
- rfrechet (Frechet), [308](#)
- rgaitdbinom (Gaitdbinom), [313](#)
- rgaitdlog (Gaitdlog), [316](#)
- rgaitdnbinom (Gaitdnbinom), [321](#)
- rgaitdpois (Gaitdpois), [327](#)
- rgaitdzeta (Gaitdzeta), [337](#)
- rgamma, [254](#), [343](#), [345](#), [350](#)
- rgengamma.stacy, [358](#)
- rgengamma.stacy (gengammaUC), [358](#)
- rgenpois0 (Genpois0), [360](#)
- rgenpois1 (Genpois1), [362](#)
- rgenpois2 (Genpois1), [362](#)
- rgenray (genray), [369](#)
- rgensh (Gensh), [372](#)
- rgeom, [661](#), [919](#)
- rgev, [381](#)
- rgev (gevUC), [381](#)
- rgompertz (Gompertz), [386](#)
- rgpd, [391](#)
- rgpd (gpdUC), [392](#)
- rgumbel, [163](#), [402](#)
- rgumbel (gumbelUC), [406](#)
- rgumbelII (Gumbel-II), [403](#)
- rhitlink, [106](#), [107](#), [112–114](#), [280](#), [294](#), [295](#), [365](#), [481](#), [549](#), [597](#), [717](#)
- rhuber, [422](#), [423](#)
- rhuber (dhuber), [234](#)
- rhzeta (Hzeta), [433](#)
- Rice, [719](#)
- riceff, [704](#), [719](#), [720](#), [757](#)
- rigff, [722](#), [756](#)
- rinv.gaussian, [879](#)
- rinv.gaussian (Inv.gaussian), [441](#)
- rinv.lomax (Inv.lomax), [444](#)
- rinv.paralogistic (Inv.paralogistic), [447](#)
- rkumar (Kumar), [458](#)
- rlaplace, [464](#)
- rlaplace (laplaceUC), [465](#)
- rlgamma, [474](#)
- rlgamma (lgammaUC), [475](#)
- rlind (Lindley), [476](#)
- rlino (Lino), [483](#)
- rlog, [317](#)
- rlog (Log), [496](#)
- rlogis, [95](#), [504](#)
- rlomax (Lomax), [518](#)
- rlplot.gev (rlplot.gevff), [723](#)
- rlplot.gevff, [381](#), [723](#)
- rmakeham (Makeham), [537](#)
- rmaxwell (Maxwell), [545](#)
- rN1binom, [578](#)
- rN1binom (N1binom), [576](#)

- rN1pois, *581*
- rN1pois (N1pois), *579*
- rnaka, *585*
- rnaka (Nakagami), *582*
- rnbinom, *322, 588, 591, 592, 594, 595, 664, 920*
- RNG, *757*
- rnorm, *84, 110, 115, 257, 302, 576, 577, 579, 580, 666, 807, 810, 821*
- rootogram4, *320, 327, 337, 342, 725*
- rootogram4vglm (rootogram4), *725*
- Round, *31, 33, 36, 38*
- round, *727*
- round2, *727, 809, 810*
- rownames, *397*
- rowSums, *333*
- rparalogistic (Paralogistic), *608*
- rpareto (Pareto), *611*
- rparetoI (ParetoIV), *614*
- rparetoII (ParetoIV), *614*
- rparetoIII (ParetoIV), *614*
- rparetoIV (ParetoIV), *614*
- rperks (Perks), *619*
- rpois, *274, 330, 562, 580, 923, 936*
- rpois.points (PoissonPoints), *645*
- rpolono (Polono), *646*
- rposbern, *649, 653*
- rposbern (posbernUC), *657*
- rposgeom, *895*
- rposgeom (Posgeom), *661*
- rposnorm (Posnorm), *665*
- rrar, *728*
- rrayleigh (Rayleigh), *701*
- rrice (Rice), *719*
- rrvglm, *13, 16, 23, 54, 78, 80, 97, 108, 109, 143–145, 147, 155, 172, 173, 177, 183–185, 210, 213, 215, 217, 218, 239, 241, 243, 320, 326, 335, 341, 395, 397, 398, 419, 430, 443, 468, 482, 488, 491, 503, 510, 512, 527, 535, 547, 548, 573, 574, 591–593, 595, 636, 642, 644, 645, 663, 669, 703, 729, 730, 734, 736, 737, 739, 744, 780, 787, 792, 794, 805, 858, 868, 870, 935, 936*
- rrvglm-class, *734*
- rrvglm.control, *209, 396–398, 535, 690, 730–732, 736, 740, 794*
- rrvglm.optim.control, *738, 739, 739*
- rsc.t2 (Expectiles-sc.t2), *258*
- rsimplex (Simplex), *754*
- rsinmad, *193*
- rsinmad (Sinmad), *758*
- rskellam (Skellam), *761*
- rskewnorm (skewnorm), *764*
- rslash, *770*
- rslash (Slash), *767*
- rstandard, *716*
- rtobit, *810*
- rtobit (Tobit), *806*
- rtopple (Topple), *814*
- rtriangle (Triangle), *818*
- rtrnorm, *824*
- rtrnorm (Trinorm), *821*
- rtruncpareto (Truncpareto), *830*
- ruge, *645, 741*
- runif, *58, 61, 72, 77, 81, 88, 98, 103, 116, 120, 153, 236, 258, 260, 278, 279, 307, 308, 358, 360, 362, 386, 403, 433, 465, 475, 477, 484, 496, 537, 582, 611, 615, 619, 657, 661, 701, 716, 719, 754, 762, 764, 768, 818, 830, 912, 916, 920*
- rweibull, *193*
- ryules, *890*
- ryules (Yules), *888*
- rzabinom (Zabinom), *891*
- rzageom (Zageom), *894*
- rzaneqbin (Zaneqbin), *897*
- rzapois, *905*
- rzapois (Zapois), *902*
- rzeta, *339*
- rzeta (Zeta), *907*
- rzibinom, *914, 915*
- rzibinom (Zibinom), *912*
- rzigeom, *919*
- rzigeom (Zigeom), *916*
- rzineqbin, *932*
- rzineqbin (Zineqbin), *919*
- rzipf (Zipf), *927*
- rzipfmb (Zipfmb), *930*
- rzipois, *920*
- rzipois (Zipois), *931*
- rzoabeta (Zoabeta), *937*
- rzoibetabinom (Betabinom), *60*
- s, *14, 187, 291, 380, 390, 391, 450, 742, 742,*

- 772–774, 777, 797, 844–846, 850, 852, 874
- sc.studentt2, 258, 259, 744, 791
- scale, 147, 634, 690, 779, 780
- score.stat, 415, 525, 745, 799, 878
- score.stat.vlm, 41, 414, 527, 798, 858
- seglines, 418, 747
- Select, 291, 398, 653, 656, 748, 863, 864
- seq2binomial, 109, 168, 751
- set.seed, 148, 208, 210, 708, 716
- setdiff, 829
- setMethod, 541, 799
- setOldClass, 165
- setup.smart, 752, 780, 887
- ships, 696, 698
- show, SurvS4-method (SurvS4-class), 804
- show.summary.drrvglm (summary.drrvglm), 792
- show.summary.pvglm (summary.pvglm), 795
- show.summary.rrvglm (summary.rrvglm), 792
- show.summary.vglm (summary.vglm), 796
- show.summary.vglm (summary.vglm), 797
- show.SurvS4 (SurvS4), 802
- Simplex, 754
- simplex, 109, 239, 722, 754, 755, 757
- simulate, 756, 867
- simulate.vlm, 66, 69, 71, 81, 83, 90, 93, 109, 159, 223, 254, 275, 298, 320, 327, 337, 342, 343, 345, 350, 358, 376, 388, 434, 435, 446, 449, 460, 479, 502, 504, 516, 521, 540, 593, 595, 621, 645, 660, 664, 670, 704, 721, 756, 761, 770, 791, 835, 890, 897, 901, 905, 919, 929, 936
- Sinmad, 758, 761
- sinmad, 76, 193, 223, 298, 355, 356, 446, 449, 521, 610, 757–759, 759
- Skellam, 761
- skellam, 645, 762, 763
- skewnorm, 764, 767
- skewnormal, 305, 765, 766, 835
- Slash, 767
- slash, 757, 768, 769
- sloglink, 28, 49, 770, 785
- sm.bs, 377, 820, 821
- sm.bs (smartpred), 779
- sm.ns, 820, 821
- sm.ns (smartpred), 779
- sm.os, 13, 14, 743, 772, 773, 775–777, 795, 796, 844–846, 851, 874
- sm.poly, 377, 820, 821
- sm.poly (smartpred), 779
- sm.ps, 13, 14, 450, 743, 773, 774, 775, 780, 795, 796, 844–846, 851
- sm.scale (smartpred), 779
- smart.expression, 778, 780
- smart.mode.is, 778, 780
- smartpred, 147, 207, 564, 567, 675, 676, 731, 732, 774, 777, 779, 846, 855, 857, 858
- smooth.spline, 772, 774, 873, 874
- specials, 327, 330, 337, 781
- specialsvglm, 30
- specialsvglm (specials), 781
- spikeplot, 234, 320, 327, 330, 334, 335, 337, 342, 629, 782, 894, 897, 901, 905, 915, 919, 923, 936
- splineDesign, 773, 774, 777
- sqrt, 784
- sqrtlink, 49, 481, 771, 784
- sratio, 24, 177, 191, 193, 215, 217–219, 452, 541, 542, 574, 716, 757, 786, 787, 867
- stat.anova, 41
- stats:dt, 372
- step, 789
- step4, 789
- step4vglm, 25, 567, 821, 858
- step4vglm (step4), 789
- stop, 41, 234
- structure, 804
- studentt, 159, 503, 757, 790, 835
- studentt2, 744, 745, 757
- studentt2 (studentt), 790
- studentt3, 757
- studentt3 (studentt), 790
- subset, 749, 750, 844
- subsetcol (Select), 748
- sum, 508
- summary, 147, 799
- summary.drrvglm, 732, 739, 792, 792
- summary.gam, 796, 797
- summary.glm, 746, 796–799, 877, 878
- summary.lm, 700, 796, 797, 799
- summary.rrvglm, 792, 799

- summary.rrvglm (summary.drrvglm), 792
 Summary.SurvS4 (SurvS4), 802
 summarypvgam, 774, 777, 795, 797, 846
 summaryvgam, 796, 796, 846
 summaryvglm, 14, 200, 414, 415, 525, 746,
 792–795, 797, 797, 820, 821, 841,
 858, 877, 878
 SURff, 384, 800, 835
 survreg, 803, 804
 SurvS4, 165, 166, 802, 804, 805, 882
 SurvS4-class, 804

 table, 783, 784
 TDist, 791
 term.names (formulavlm), 305
 term.namesvml (formulavlm), 305
 terms, 735, 849, 856, 860
 TIC, 805
 TICvml, 27
 TICvml (TIC), 805
 title, 232, 622–624, 692, 723, 827, 828
 Tobit, 806
 tobit, 164, 246, 556, 668, 727, 807, 808, 835
 Tol, 544, 603, 812
 Topleft, 814, 816
 topple, 412, 814, 815, 815, 818
 toxop, 248, 817
 trap0, 148, 210, 461
 Triangle, 815, 816, 818
 triangle, 757, 818
 trigamma, 589
 trim.constraints, 25, 202, 567, 789, 819
 Trinorm, 821
 trinormal, 113, 822, 823, 835
 trplot, 528, 824
 trplot.qrrvglm, 210, 825, 826
 Trunc, 320, 327, 337, 342, 385, 828
 truncated, 782
 truncated (altered), 29
 truncgeometric (geometric), 375
 Truncpareto, 614, 830
 truncpareto, 830
 truncpareto (paretoff), 612
 truncweibull, 627, 831, 880, 883
 TypicalVGAMfamilyFunction, 16, 482, 739,
 858, 864
 TypicalVGAMfamilyFunction
 (CommonVGAMffArguments), 188
 TypicalVGAMlink (Links), 480

 ucberk, 833
 Uniform, 427, 545, 814
 uninormal, 43, 111, 113, 164, 172, 246, 270,
 271, 280, 304, 305, 366, 397, 412,
 419, 423, 516, 523, 560, 578, 581,
 596, 598, 606, 607, 668, 680, 694,
 696, 698, 713, 757, 767, 791, 801,
 810, 822, 824, 834, 835
 uniroot, 36, 138, 140, 144, 302, 303, 582
 update, 25, 41, 526, 789, 858
 uqo, 207, 210
 uqo (grc), 395
 UtilitiesVGAM, 195, 836

 V1, 645, 838, 839
 V2, 645, 838, 839
 valt0.control, 738
 vcov, 794, 840, 841
 vcovdrrvglm, 794
 vcovdrrvglm (vcovvml), 840
 vcovqrrvglm, 210, 566, 841
 vcovqrrvglm (vcovvml), 840
 vcovrrvglm, 841
 vcovrrvglm (vcovvml), 840
 vcovvml, 200, 799, 840, 841
 vector, 804
 venice, 163, 381, 402, 409, 842
 venice90 (venice), 842
 VGAM (VGAM-package), 13
 vgam, 13, 16, 18–21, 23, 31, 34, 36, 38, 43, 60,
 71, 74, 75, 78, 80, 83, 85, 90–92, 94,
 97, 101, 106, 108, 112, 114, 118,
 122, 124, 127, 147, 155, 159, 160,
 162, 164, 165, 172, 186, 187, 194,
 195, 202, 215, 217, 218, 223, 226,
 237, 239, 241, 243, 246, 254, 262,
 264, 268, 273, 274, 277, 284, 286,
 288, 289, 291, 297, 304, 310, 312,
 320, 326, 335, 341, 343, 345, 347,
 348, 350, 355, 357, 364, 367, 368,
 371, 374, 375, 379, 380, 388, 390,
 391, 401, 405, 422, 429, 430, 432,
 435, 440, 443, 446, 449, 450, 460,
 464, 467, 468, 472, 474, 478, 482,
 486, 488, 491, 494, 500, 501, 503,
 510, 512, 516, 520, 523, 527, 528,
 539, 541, 547, 548, 554, 557, 559,
 561, 563, 573, 578, 581, 584, 591,
 595, 597, 600, 604, 610, 613, 618,

- 621, 632, 636, 637, 642, 644, 649,
652, 655, 659, 663, 668, 669, 677,
682, 692, 693, 703, 712, 713, 721,
722, 726, 728, 731, 742–744, 751,
752, 755, 757, 760, 763, 766, 769,
773, 774, 776, 777, 780, 787, 791,
795–797, 801, 805, 809, 816, 824,
832, 835, 843, 850–853, 858, 866,
868, 870, 873, 879, 880, 882, 884,
890, 893, 896, 900, 904, 911, 915,
918, 922, 926, 929, 935
- vgam-class, 848
- VGAM-package, 13
- vgam.control, 149, 844–846, 850, 850
- vglm, 13, 16, 18–21, 23–25, 30, 31, 34, 36, 38,
40, 41, 43, 46, 60, 65, 68, 71, 74, 75,
78, 80, 83, 85, 87, 90–92, 94, 97,
101, 106, 108, 112, 114, 118, 122,
124, 127, 131, 133, 155, 159, 160,
162, 164, 165, 172, 176, 179, 187,
188, 191, 192, 194, 195, 199, 202,
203, 209, 210, 215, 217, 218, 223,
226, 228, 230, 237, 239, 241, 243,
246, 248, 251, 254, 262, 264, 268,
273, 274, 277, 281, 284, 286,
288–291, 297, 300, 304, 310, 312,
320, 326, 335, 341, 343, 345, 347,
348, 350–352, 355, 357, 364, 367,
368, 371, 374, 375, 379, 380, 385,
388, 390, 391, 395, 397, 401, 405,
410–412, 415, 422, 429, 430, 432,
435, 439, 440, 443, 446, 449–455,
460, 464, 467, 468, 472, 474, 478,
480, 482, 486, 488, 491, 494, 500,
501, 503, 510, 512, 516, 520, 523,
525–528, 539, 541, 542, 547, 548,
554, 557, 559, 561, 563, 567, 573,
578, 581, 584, 591, 592, 595, 597,
604, 606, 610, 613, 618, 621, 632,
636, 637, 640, 642, 644, 649, 652,
655, 659, 663, 668, 669, 674–677,
682, 692–696, 699–701, 703, 712,
713, 715, 717, 721–723, 726,
728–732, 739, 742–744, 746,
750–752, 755, 757, 760, 763, 766,
769, 774, 780, 782, 787, 789, 791,
798, 799, 801, 803–805, 809, 816,
820, 821, 824, 832, 835, 840,
844–846, 850, 852, 853, 853, 861,
864, 866, 868, 870, 876, 878–880,
882, 884, 885, 890, 893, 896, 900,
904, 911, 915, 918, 922, 926, 929,
935
- vglm-class, 859
- vglm.control, 44, 102, 199, 218, 285, 291,
300, 355, 356, 382, 397, 573, 637,
681, 690, 734, 738, 739, 801, 846,
848, 851–855, 857, 858, 860, 862,
877, 923, 926
- vglmff-class, 866
- vonmises, 155, 869
- vpairs.profile (vplot.profile), 871
- vplot.profile, 871
- vsmooth.spline, 490, 743, 846, 853, 872
- waitakere, 426, 874
- wald.stat, 200, 415, 525, 746, 799, 876
- wald.stat.vlm, 41, 414, 524, 525, 527, 746,
798, 858
- waldff, 442, 443, 878
- weibull.mean, 879, 883
- weibullR, 193, 263, 358, 381, 404, 405, 704,
831, 832, 880, 881
- weightsvglm, 144, 856, 884
- wine, 885
- wrapup.smart, 780, 886
- xs.nz, 572
- yeo.johnson, 887
- yip88, 904, 935, 936
- Yules, 888
- yulesimon, 757, 889, 890
- Zabinom, 891
- zabinomial, 892, 915
- zabinomialff (zabinomial), 892
- Zageom, 894
- zageometric, 376, 661, 757, 895, 895, 918,
919
- zageometricff, 757
- zageometricff (zageometric), 895
- Zanegbin, 897
- zanegbinomial, 664, 757, 898, 899, 923
- zanegbinomialff, 757
- zanegbinomialff (zanegbinomial), 899
- Zapois, 329, 330, 902

zapoisson, [334](#), [337](#), [670](#), [726](#), [757](#), [903](#), [903](#),
[935](#), [936](#)
zapoissonff, [334](#), [757](#)
zapoissonff (zapoisson), [903](#)
zero, [202](#), [905](#)
Zeta, [907](#), [910](#), [911](#)
zeta, [238](#), [434](#), [435](#), [470](#), [908](#), [908](#), [911](#)
zetaff, [237](#), [238](#), [342](#), [434](#), [435](#), [907–910](#),
[910](#), [927–929](#)
zetaffMlink, [341](#), [342](#), [385](#)
Zibinom, [912](#)
zibinomial, [109](#), [892–894](#), [912](#), [913](#), [913](#)
zibinomialff (zibinomial), [913](#)
Zigeom, [916](#)
zigeometric, [376](#), [661](#), [757](#), [895](#), [897](#), [917](#),
[917](#)
zigeometricff, [757](#)
zigeometricff (zigeometric), [917](#)
Zinegbin, [919](#), [923](#)
zinegbinomial, [190](#), [593](#), [732](#), [757](#), [901](#), [920](#),
[921](#), [935](#)
zinegbinomialff (zinegbinomial), [921](#)
ziP, [335](#)
zipebcom, [102](#), [924](#), [936](#)
Zipf, [927](#), [931](#)
zipf, [237](#), [238](#), [757](#), [911](#), [927](#), [928](#), [928](#)
Zipfmb, [928](#), [930](#)
Zipois, [328–330](#), [913](#), [917](#), [920](#), [931](#), [935](#), [936](#)
zipoisson, [30](#), [334](#), [337](#), [395](#), [397](#), [412](#), [645](#),
[669](#), [670](#), [726](#), [732](#), [757](#), [901](#), [904](#),
[905](#), [915](#), [918](#), [923–926](#), [932](#), [933](#),
[940](#)
zipoissonff, [334](#), [395](#), [397](#), [412](#), [522](#), [757](#)
zipoissonff (zipoisson), [933](#)
Zoabeta, [63](#), [937](#), [940](#)
zoabetaR, [938](#), [939](#)